

EE 046202 - Technion - Unsupervised Learning & Data Analysis

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Tutorial 04 - Classical Methods in Statistical Inference - Hypothesis Testing 2



Agenda

- Recap
- T-Test)
 - Exercise Statistics for Machine Learning
- Pearsons's Chi-Squared (χ^2) Test-Test)
- Uniformly Most Powerful (UMP) Test)
- Which Statistical Test To Use???
- Recommended Videos
- Credits

```
In [1]: # imports for the tutorial
import numpy as np
import pandas as pd
from scipy.stats import t as t_dist
from scipy.stats import chi2
import matplotlib.pyplot as plt
%matplotlib notebook
```



In the previous tutorial we described the framework of Hypothesis Testing. Following is a recap of some of the definitions and a brief summmary:

- 1. Null & Alternative Hypotheses Formulate the null hypothesis H_0 : $\theta \in \Theta_0$ (that the observations are the result of pure chance) and the alternative hypothesis H_1 : $\theta \in \Theta_1$ (that the observations show a real effect combined with a component of chance variation).
- 2. Test Statistic Identify a test statistic that can be used to assess the truth of the null hypothesis. It is a value computed from sample data. The test statistic is used to assess the strength of evidence in support of a null hypothesis.
 - A **statistic** is a real-valued function of the data. For example, the sample mean: $W(X_1,X_2,\ldots,X_n)=rac{X_1+X_2+\ldots+X_n}{n}$ is a statistic.
 - A **test satistic** is a statistic on we which we build our test.
 - **Acceptence Region** A A set $A \subset \mathbb{R}$ is defined to be the set of all possible values of the test statistic for which we accept H_0 .
 - ullet Rejection Region R A set $R=\mathbb{R}-A$ is defined to be the set of all possible values of the test statistic for which we reject H_0 and accept H_1 .
- 3. P-value & Interpretation Compute the P-value, which is the probability that a test statistic, at least as significant as the one observed, would be obtained assuming that the null hypothesis were true. The smaller the P-value, the stronger the evidence against the null hypothesis.
- 4. Significance Level Compare the p-value to an acceptable significance value α (sometimes called an α value, a probability threshold below which the null hypothesis will be rejected. Common values are 5% and 1%.). If $p \le \alpha$ (the observed effect is statistically significant), the null hypothesis is ruled out, and the alternative hypothesis is valid.
 - Commonly used p-values:

P-Value	Wording
p > 0.05	Not Significant
$0.01 \leq p \leq 0.05$	Significant
$0.001 \leq p < 0.01$	Very Significant
p<0.001	Extremely Significant

- Type I Error (False Positive) the incorrect rejection of a true null hypothesis. Usually a type I error leads one to conclude that a supposed effect or relationship exists when in fact it doesn't.
 - For example, a test that shows a patient to have a disease when in fact the patient does not have the disease, a fire alarm going on indicating a fire when in fact there is no fire, or an experiment indicating that a medical treatment should cure a disease when in fact it does not.
 - ullet The chance of **rejecting the null hypthesis** H_0 , **when it is TRUE**, denoted by lpha
 - ightarrow the chance of accepting the null hypthesis H_0 , when it is TRUE is 1-lpha
 - Formally:
 - \circ Denote a test statistic as W
 - $P(\text{Type 1 Error}|\theta) = P(\text{Reject } H_0|\theta) = P(W \in R|\theta), \theta \in \Theta_0$
 - If $P(\text{Type 1 Error}|\theta) \leq \alpha, \forall \theta \in \Theta_0$, we say that the test has significance level α .
- Type II Error (False Negative) the failure to reject a false null hypothesis.
 - For example, a blood test failing to detect the disease it was designed to detect, in a patient who really has the disease; a fire breaking out and the fire alarm does not ring; or a clinical trial of a medical treatment failing to show that the treatment works when really it does.
 - ullet The chance of **not rejecting the null hypothesis** H_0 , **when it is FALSE**, denoted by eta
 - ullet the chance of **rejecting the null hypthesis** H_0 , **when it is FALSE** is 1-eta (also called **power**)
 - Since the alternative hypothesis, H_1 , usually contains more than one value of θ , the probability of type II error is usually a **function of** θ , and denoted by β .
 - Formally: $\beta(\theta) = P(\text{Accept } H_0 | \theta), \text{ for } \theta \in \Theta_1$

When To Use What:

- 2-sided hypothesis testing for the mean: $H_0: \mu=\mu_0, H_1: \mu
eq \mu_0$

Case	Test Statistic	Acceptance Region
$X_i \sim \mathcal{N}(\mu, \sigma), \sigma$ known	$W=rac{\overline{X}-\mu_0}{rac{\sigma}{\sqrt{n}}}$	$ W \leq z_{rac{lpha}{2}}$
n large, X_i non-normal	$W=rac{\overline{X}-\mu_0}{rac{S}{\sqrt{n}}}$	$ W \leq z_{rac{lpha}{2}}$

• 1-sided hypothesis testing for the mean: $H_0: \mu \leq \mu_0, H_1: \mu > \mu_0$

Case	Test Statistic	Acceptance Region
$X_i \sim \mathcal{N}(\mu, \sigma), \sigma$ known	$W=rac{\overline{X}-\mu_0}{rac{\sigma}{\sqrt{n}}}$	$W \leq z_{rac{lpha}{2}}$
n large, X_i non-normal	$W=rac{\overline{X}-\mu_0}{rac{S}{\sqrt{n}}}$	$W \leq z_{rac{lpha}{2}}$

- The only difference is the $\emph{absolute}$ sign on W

All expansions can be found HERE (https://www.probabilitycourse.com/chapter8/8 4 3 hypothesis testing for mean.php).



Exercise - How Much Protein Does a Tuna Fish Contain?

• Note: this is for all of you who work out :)

Assume that the amount of protein in a Tuna-can is a Gaussian random variable with std $\sigma=2.5$ grams (that is, the variance is *known*) and *unknown* expectation μ . We are interested in the following hypothesis:

$$H_0: \mu = 30 \ H_1: \mu = 32$$

- 1. The following test is proposed: take 16 random cans (samples are i.i.d.), if the mean is greater than 30.8 grams, reject H_0 . What is the probability of type 1 error? What is the probability of type 2 error?
- 2. We decided to use a different test: reject H_0 in case $\overline{X}_n > 31.3$. Can you say if type 1 or 2 error probabilities will increase or decrease without further computations?



Section 1

Q: The following test is proposed: take 16 random cans (samples are i.i.d.), if the mean is greater than 30.8 grams, reject H_0 . What is the probability of type 1 error? What is the probability of type 2 error?

Since each sample is Gaussian, and the samples are i.i.d., the empirical average \overline{X}_n of 16 samples is also Gaussian, with the **same** mean, and std of $\frac{\sigma}{\sqrt{16}}$:

$$\overline{X}_n \sim \mathcal{N}(\mu, rac{2.5^2}{16})$$

(Z-test)

- Type-1-error (False Positive) $=P_{H_0}(\overline{X}_n>30.8)=lpha$

$$=P(rac{\overline{X}_n-\mu_{H_0}}{2.5/\sqrt{16}}>rac{30.8-\mu_{H_0}}{2.5/\sqrt{16}}) \ =P(rac{\overline{X}_n-30}{2.5/\sqrt{16}}>rac{30.8-30}{2.5/\sqrt{16}}) \ =P(Z>1.28)=1-P(Z<1.28)=1-\Phi(1.28)=0.1$$

- $\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{x} \exp(\frac{-x^2}{2}) dx$
- Comulative distribution function of a standard normal distribution (https://en.wikipedia.org/wiki/Standard normal table#Cumulative)
- Type-2-error (False Negative) $=P_{H_1}(\overline{X}_n \leq 30.8)$

$$=P(\frac{\overline{X}_n-\mu_{H_1}}{2.5/\sqrt{16}}\leq \frac{30.8-\mu_{H_1}}{2.5/\sqrt{16}})$$

$$=P(\frac{\overline{X}_n-32}{2.5/\sqrt{16}}\leq \frac{30.8-32}{2.5/\sqrt{16}})$$

$$=P(Z\geq 1.92)=1-P(Z\leq 1.92)=1-\Phi(1.92)=0.0274$$

Section 2

Q: We decided to use a different test: reject H_0 in case $\overline{X}_n>31.3$. Can you say if type 1 or 2 error probabilities will increase or decrease without further computations?

We increased the threshold, such that it less likely to reject H_0 . That is, the event in which H_0 is true, and we reject it is more rare. This means that type-1 error must *decrease*, i.e, we will make less False-Positive mistakes. However, the event in which we accept H_0 while H_1 is true is **more likely** (notice the trade-off between the error types). This means that the type-2 error must increase, and we will make more False-Negative mistakes. This demnostrates the inherent **trade-off** between type-1 and type-2 errors.

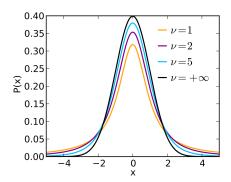


T-Test (Hypothesis Testing of Gaussian with Unknown Mean and Variance)

- ullet The t-test deals with the scenraio where both the mean and variance are unknown.
 - lacktriangledown It is possible to use z-test, however, it is sometimes better to do t-test for this case.
- It is based on **student-t** distribution, which is a heavy-tail distribution.
 - Note that although the student-t distribution may look like a Gaussian, it is different (in the limit of degrees of freedom → ∞, student-t approaches
- Assume we are given data $\{X_i\}_{i=1}^N$, where $X_i\sim \mathcal{N}(\mu,\sigma^2)$. Both μ and σ^2 are unknown.
- Let the empirical mean and variance be:

$$egin{aligned} \overline{X}_N &= rac{1}{N} \sum_{i=1}^N X_i \ & \overline{\sigma}_N^2 &= rac{1}{N-1} \sum_{i=1}^N (X_i - \overline{X}_N)^2 \end{aligned}$$

• Student-t PDF:



• Image Credits: By Skbkekas - Own work, CC BY 3.0 (https://commons.wikimedia.org/w/index.php?curid=9546828)

· Assume we wish to test the following 2-sided null and alternative hypotheses given the data:

$$H_0: \mu=\mu' \ H_1: \mu
eq \mu'$$

ullet The t-test statistic is defined by:

$$t=rac{\overline{X}_N-\mu'}{rac{ar{ar{\sigma}_N}}{\sqrt{N}}}$$

ullet The t-test is |t| < C, that is, we **reject** H_0 if $|t| \geq C$

In general, there are three possible alternative hypotheses and rejection regions:

Alternative Hytpothesis	Rejection Region
$H_a: \mu_1 eq \mu_2$	$ T >t_{n-1,1-\frac{\alpha}{2}}$
$H_a: \mu_1 > \mu_2$	$T>t_{n-1,1-lpha}$
$H_a: \mu_1 < \mu_2$	$T < t_{n-1,lpha}$

More Details (https://www.itl.nist.gov/div898/handbook/eda/section3/eda353.htm)



Exercise - Statistics for Machine Learning

In this exercise we will use the *t*-test to estimate the performance of two classifiers on a single domain. The data is taken from "Evaluating Learning Algorithms - A Classification Perspective", N.Japkowicz, M.Shah.

- **Preview**: Two data samples are matched if they come from repeated observation of the same subject. Given two types of samples, we want to test whether the difference in means between these two samples is *significant*. One way of doing so is by looking at the difference in observed means and standard deviations (the first and second moments of the samples) between these two samples.
- In classification problems, we often try out different classifiers (which are the output of different learning algorithms) and check their performance. Eventually, we wish to choose the best classifier (the one with the lowest classification error).
- Consider two classifiers and their performance on the data. How can we tell if one is better than the other?
- ullet The null hypothesis, H_0 , will be that the classifiers are equivalent and we will see if we can reject H_0 .
- Let \mathcal{X}, \mathcal{Y} be the feature space and the label space accordingly. Let $f: \mathcal{X} \to \mathcal{Y}$ be a classifier. Denote the validation data (the data on which is not seen during training, but is used to tune the hyper-parameters of the algorithms) as $\mathcal{D} = \{x_i, y_i\}_{i=1}^n$.
- Define our performance measure (pm):

$$pm_i(f) = \begin{cases} 1, & f \text{ classified incorrectly } x_i \\ 0, & \text{otherwise} \end{cases}$$

- Note that it is actually the error ("...incorrectly...").
- The *average* performance measure of f:

$$\overline{pm}(f) = rac{1}{n} \sum_{i=1}^n pm_i(f)$$

- Performance difference measures of two classifier f_1 and $f_2\colon$

$$egin{aligned} d_i &= pm_i(f_1) - pm_i(f_2) \ ar{d} &= \overline{pm}(f_1) - \overline{pm}(f_2) \end{aligned}$$

• Define the t-statistics (as we do not know the mean nor the std):

$$t=rac{ar{d}-0}{ ilde{\sigma}_d/\sqrt{n}}$$

$$\tilde{\sigma}_d^2 = \frac{1}{n-1} \sum_{i=1}^n (d_i - \bar{d})^2$$

- $\tilde{\sigma}_d^2 = \frac{1}{n-1} \sum_{i=1}^n (d_i \bar{d})^2$ Recall that our null hypothesis, H_0 is that the classifiers are equivalent, thus, the expected difference between their performance mesures is 0.
- The data: performance measures of 10 cross-validation sets on some given data. The algorithms were repeated for 10 trials, and the performance was measured in each case. The values are emprical risks - the error.

Out[2]:

	Trial No.	Classifiers	f1	f2	f3	f4	f5	f6	f7	f8	f9	f10
0	1	c4.5	0.5000	0.0000	0.3333	0.0000	0.3333	0.3333	0.3333	0.2	0.2	0.2
1	1	RIP	0.3333	0.5000	0.0000	0.0000	0.3333	0.5000	0.1667	0.0	0.2	0.0
2	1	NB	0.1667	0.0000	0.0000	0.0000	0.0000	0.1667	0.0000	0.4	0.0	0.0
3	2	c4.5	0.3333	0.1667	0.1667	0.3333	0.3333	0.0000	0.0000	0.0	0.2	0.2
4	2	RIP	0.3333	0.1667	0.0000	0.0000	0.3333	0.1667	0.0000	0.2	0.2	0.0
5	2	NB	0.3333	0.0000	0.1667	0.1667	0.0000	0.0000	0.0000	0.0	0.0	0.0
6	3	c4.5	0.0000	0.0000	0.3333	0.3333	0.1667	0.3333	0.1667	0.0	0.2	0.2
7	3	RIP	0.0000	0.0000	0.1667	0.3333	0.1667	0.1667	0.1667	0.2	0.4	0.2
8	3	NB	0.0000	0.0000	0.0000	0.0000	0.0000	0.1667	0.0000	0.0	0.0	0.0
9	4	c4.5	0.5000	0.1667	0.3333	0.1667	0.0000	0.1667	0.5000	0.2	0.2	0.4
10	4	RIP	0.3333	0.0000	0.1667	0.1667	0.1667	0.0000	0.6667	0.2	0.2	0.2
11	4	NB	0.3333	0.0000	0.0000	0.1667	0.0000	0.0000	0.0000	0.0	0.0	0.2
12	5	c4.5	0.0000	0.1667	0.3333	0.1667	0.0000	0.1667	0.0000	0.2	0.4	0.2
13	5	RIP	0.3333	0.1667	0.1667	0.1667	0.0000	0.1667	0.0000	0.4	0.0	0.2
14	5	NB	0.0000	0.1667	0.0000	0.0000	0.1667	0.0000	0.0000	0.0	0.0	0.4
15	6	c4.5	0.3333	0.1667	0.1667	0.1667	0.6667	0.1667	0.3333	0.2	0.2	0.0
16	6	RIP	0.0000	0.0000	0.3333	0.0000	0.5000	0.0000	0.0000	0.0	0.2	0.0
17	6	NB	0.0000	0.0000	0.1667	0.0000	0.1667	0.1667	0.0000	0.0	0.0	0.0
18	7	c4.5	0.1667	0.3333	0.1667	0.3333	0.1667	0.1667	0.3333	0.2	0.0	0.2
19	7	RIP	0.1667	0.1667	0.1667	0.3333	0.1667	0.1667	0.3333	0.4	0.0	0.2
20	7	NB	0.1667	0.1667	0.1667	0.0000	0.1667	0.3333	0.3333	0.0	0.2	0.0
21	8	c4.5	0.0000	0.1667	0.1667	0.0000	0.0000	0.0000	0.1667	0.0	0.6	0.4
22	8	RIP	0.1667	0.1667	0.1667	0.0000	0.0000	0.1667	0.0000	0.0	0.2	0.2
23	8	NB	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0	0.2	0.2
24	9	c4.5	0.1667	0.3333	0.1667	0.3333	0.3333	0.3333	0.0000	0.6	0.2	0.2
25	9	RIP	0.0000	0.0000	0.1667	0.0000	0.0000	0.1667	0.0000	0.4	0.0	0.2
26	9	NB	0.0000	0.0000	0.0000	0.1667	0.0000	0.0000	0.0000	0.0	0.0	0.2
27	10	c4.5	0.5000	0.1667	0.1667	0.5000	0.5000	0.3333	0.0000	0.4	0.4	0.0
28	10	RIP	0.5000	0.1667	0.1667	0.1667	0.0000	0.3333	0.3333	0.2	0.2	0.2
29	10	NB	0.5000	0.1667	0.0000	0.0000	0.0000	0.0000	0.1667	0.2	0.2	0.0

In [3]: # for a certain classifier:
 data.where(data['Classifiers'] == 'NB').dropna() # without dropna(), it will show NaN on other rows

Out[3]:

	Trial No.	Classifiers	f1	f2	f3	f4	f5	f6	f7	f8	f9	f10
2	1.0	NB	0.1667	0.0000	0.0000	0.0000	0.0000	0.1667	0.0000	0.4	0.0	0.0
5	2.0	NB	0.3333	0.0000	0.1667	0.1667	0.0000	0.0000	0.0000	0.0	0.0	0.0
8	3.0	NB	0.0000	0.0000	0.0000	0.0000	0.0000	0.1667	0.0000	0.0	0.0	0.0
11	4.0	NB	0.3333	0.0000	0.0000	0.1667	0.0000	0.0000	0.0000	0.0	0.0	0.2
14	5.0	NB	0.0000	0.1667	0.0000	0.0000	0.1667	0.0000	0.0000	0.0	0.0	0.4
17	6.0	NB	0.0000	0.0000	0.1667	0.0000	0.1667	0.1667	0.0000	0.0	0.0	0.0
20	7.0	NB	0.1667	0.1667	0.1667	0.0000	0.1667	0.3333	0.3333	0.0	0.2	0.0
23	8.0	NB	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0	0.2	0.2
26	9.0	NB	0.0000	0.0000	0.0000	0.1667	0.0000	0.0000	0.0000	0.0	0.0	0.2
29	10.0	NB	0.5000	0.1667	0.0000	0.0000	0.0000	0.0000	0.1667	0.2	0.2	0.0

Questions

- 1. Map the problem to the framework of Hypothesis Testing, with a test in the form of a t-test.
- 2. Define d_i as the *average* difference per trial. Calculate d_i between the classifier c4.5 and NB for every one of the 10 given trials.
- 3. Calculate $\bar{d}, \tilde{\sigma}_d^2$ and t.
- 4. For the significane levels of lpha=0.05, do the two classifiers have similar performance?



Section 1

Q: Map the problem to the framework of Hypothesis Testing, with a test in the form of a t-test.

- We ask a question about the *random variable* d_i .
- d_i measures the performance difference of the algorithms when they are trained on a specific realization of the data (which is an RV).
- Our null hypothesis, H_0 , is that $\mu_{d_i}=0$: the two classifiers are similar.
- However, we do not know the variance of the random variable.
- · We can assume that:

$$d_i \sim \mathcal{N}(\mu', \sigma)$$

where the parameters are unknown.

- This assumption makes sense in the *limit of large data* and under the (wrong) assumption that each entry of the table is drawn independently from a Gaussian distribution.
- · Since we do not know the variance, this framework exactly fits the t-test framework.

Section 2

Q: Define d_i as the average difference per trial. Calculate d_i between the classifier c4.5 and NB for every one of the 10 given trials.

- ullet Recall that d_i is the difference in performance measure between the two classifiers.
- We are given 10 partitions of the dataset (for cross-validation) and we also perform 10 trials.
- ullet Thus, d_i is the difference of performance in the i^{th} trial, and we calculate it as the average over cross-validations sets.

$$d_i = rac{1}{10} \sum_{j=1}^{10} p m_{i,j}(c4.5) - p m_{i,j}(NB),$$

where j is the index of the cross-valdiation set.

```
In [4]: # let's calculate all of the d_i
        nb_data = data.where(data['Classifiers'] == 'NB').drop(action == 'Trial No.', 'Classifiers']).mean(axis=1).
        c45_data = data.where(data['Classifiers'] == 'c4.5').dropna().drop(columns=['Trial No.', 'Classifiers']).mean(axis=
        1).values
        d = c45_data - nb_data
        print("differnce \stackrel{-}{\text{in}} performance measure for c4.5 and NB:")
        for i in range(len(d)):
           print("d_{{}} = {})".format(i + 1, d[i]))
       differnce in performance measure for c4.5 and NB:
       d_1 = 0.16998000000000005
       d_3 = 0.15666
       d = 0.19334
       d_5 = 0.0899999999999998
        d_6 = 0.19
       d_7 = 0.053330000000000016
       d_9 = 0.22999
        d_{10} = 0.17333
```

Section 3

Q: Calculate $\bar{d}, \tilde{\sigma}_d^2$ and t.

Let's calculate $\bar{d}, \tilde{\sigma}_d^2$ and t. Recall:

• Define our **performance measure** (pm):

$$pm_i(f) = egin{cases} 1, & f ext{ classified incorrectly } x_i \ 0, & ext{otherwise} \end{cases}$$

• The average performance measure of f:

$$\overline{pm}(f) = rac{1}{n} \sum_{i=1}^n pm_i(f)$$

- Performance $\emph{difference}$ measures of two classifier f_1 and f_2 :

$$egin{aligned} d_i &= pm_i(f_1) - pm_i(f_2) \ \overline{d} &= \overline{pm}(f_1) - \overline{pm}(f_2) \end{aligned}$$

• Define the t-statistics (as we do not know the mean nor the std):

$$t=rac{ar{d}-0}{ ilde{\sigma}_d/\sqrt{n}}$$

• $\tilde{\sigma}_d^2 = \frac{1}{n-1} \sum_{i=1}^n (d_i - \bar{d})^2$

```
In [5]: # Let's calculate the wanted terms
d_avg = d.mean()
print("d_average: ", d_avg)

var_d = (1 / (len(d) - 1)) * np.sum(np.square(d - d_avg))
print("empirical variance of d: ", var_d)
print("empirical std of d: ", np.sqrt(var_d))

t_test = (d_avg - 0) / (np.sqrt(var_d) / np.sqrt(len(d)))
print("t = ", t_test)
```

d_average: 0.1473300000000002
empirical variance of d: 0.0030239951333333338
empirical std of d: 0.054990864089713425
t = 8.472286722254982

Q: For the significane levels of lpha=0.05, do the two classifiers have similar performance?

- After calculating t, all that is left to be done is analyzing the distribution of t, which we know has the form of the **student-t distribution** with n-1 degrees of freedom (note that for $n \to \infty$ the student-t distribution approaches the Normal distribution).
- In our case, n=10, since we averaged 10 RVs.
- For two-tailed alternative hypothesis we can read the value of $t_{n-1,1-rac{lpha}{2}}$ which set the t-test to have a certain size:

$$P(t>t_{n-1,1-rac{lpha}{2}})=rac{lpha}{2}$$

Since we wish to have a test with size lpha=0.05, we read the value:

$$t_{n-1,1-rac{lpha}{2}}=t_{9,1-0.025}$$

```
In [6]: alpha = 0.05
t_value = t_dist.ppf(1 - alpha / 2, df=len(d) - 1) # Percent point function (inverse of cdf - percentiles)
print("t value: ", t_value)
```

t value: 2.2621571627409915

So we got that:

$$t_{n-1,1-\frac{\alpha}{2}}=t_{9,1-0.025}=2.26$$

And the test takes the following form: ${f reject}$ the null hypothesis H_0 if

$$|t|>t_{n-1,rac{lpha}{2}}=2.26$$

Therfore, given the test size, lpha=0.05, we reject H_0 and conclude that NB and c4.5 peform differently.

Section 4 Continued

• For one-sided alternative hypothesis ($\mu_1>\mu_2 o ar d>0$) - we can read the value of $t_{n-1,1-lpha}$ which set the t-test to have a certain size:

$$P(t > t_{n-1,1-lpha}) = lpha$$

Since we wish to have a test with size lpha=0.05, we read the value:

$$t_{n-1,1-\alpha} = t_{9,0.05}$$

```
In [7]: alpha = 0.05
t_value = t_dist.ppf(1-alpha, df=len(d)-1) # Percent point function (inverse of cdf - percentiles)
print("t value: ", t_value)
```

t value: 1.8331129326536335

So we got that:

$$t_{n-1,1-\alpha} = t_{9,1-0,05} = 1.833$$

And the test takes the following form: ${f reject}$ the null hypothesis H_0 if

$$t > t_{n-1,1-\alpha} = 1.833$$

Therfore, given the test size, $\alpha=0.05$, we reject H_0 and conclude that NB performs consistently *better* than c4.5 (since the expected error rate of c4.5 is higher).



Pearson's Chi-Squared (χ^2) Test

- Pearson's chi-squared test is a statistical test applied to sets of categorical data to evaluate how likely it is that any observed difference between the sets arose by chance.
 - Categorical data data consisting of categorical variables (a variable that can take on one of the limited possible values).
 - For example subway usage on specific days:

Average Weekday	Average Saturday	Average Sunday
5,284,295	3,202,388	2,555,814

- By using Pearson's chi-squared test we can accept or reject the null hypothesis, H₀. Usually, a null hypothesis represented in the way of an expected set
 of data, and alternative hypothesis, H₁, as some newly observed data.
- The Chi-Squared Statistic:

$$\chi^2 = \sum_{i=1}^n rac{(O[i] - E[i])^2}{E[i]}$$

- ullet O[i] Observed data (i^{th} category).
- E[i] Expected data (for the i^{th} category).
- ullet We will use the **Chi-squared distribution function** with n-1 degrees of freedom (DOF), where n is the number of categories.
- For a **significance level** α , we reject the null hypothesis, H_0 if the chi-squared statistic is above the **critical value**.
 - The critical value is calculated using the Chi-squared distribution function for a certain significance level.

Example - Dropout/Graduation Rates

Consider the following data about dropout/graduation rates (which will serve as our null hypothesis).

First Year Dropout	Second Year Dropout	Third Year Dropout	Fourth Year Dropout	Graduated
0.2	0.1	0.15	0.05	0.5

Recently, a research was conducted and the observed data from a sample size of N=2000 was:

	First Year Dropout	Second Year Dropout	Third Year Dropout	Fourth Year Dropout	Graduated
Number of People	543	145	210	110	992

Assuming the old data was corerct (H_0) , what would the expected data look like?

	First Year Dropout	Second Year Dropout	Third Year Dropout	Fourth Year Dropout	Graduated	Total
Observed	543	145	210	110	992	2000
Expected	$2000 \cdot 0.2 = 400$	$2000\cdot 0.1=200$	$2000 \cdot 0.15 = 300$	$2000 \cdot 0.05 = 100$	$2000 \cdot 0.5$	2000

We will now perform the Chi-squared test for a significance level lpha=0.05

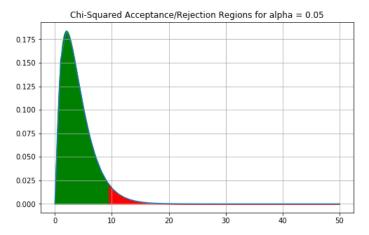
```
In [8]: # chi-squared test
        num_categories = 5
        dof = num_categories - 1 # degrees of freedom
        alpha = 0.05
        observed = [543, 145, 210, 110, 992]
        expected = [400, 200, 300, 100, 1000]
        # compute chi-squared statistic
        chi_square_stat = np.sum((np.array(observed) - np.array(expected))**2 / np.array(expected))
        print("chi-squared statistic value: ", chi_square_stat)
        # compute critical value using the inverse of the CDF
        critical_value = chi2.ppf(q=1-alpha, df=dof)
        print("critical value: ", critical_value)
        if chi_square_stat > critical_value:
            print("H_0 is rejected!")
        else:
            print("H_0 is accepted!")
```

chi-squared statistic value: 94.3115 critical value: 9.487729036781154 H_0 is rejected!

```
In [9]: # plot rejection/acceptance regions
x = np.linspace(0, 50)
fig = plt.figure(figsize=(8, 5))
ax = fig.add_subplot(111)
ax.plot(x, chi2.pdf(x, df=dof))

accept_region_xs = np.arange(0, critical_value, 0.001)
ax.fill_between(accept_region_xs, chi2.pdf(accept_region_xs, df=dof), color='green')
reject_region_xs = np.arange(critical_value, 50, 0.001)
ax.fill_between(reject_region_xs, chi2.pdf(reject_region_xs, df=dof), color='red')
ax.grid()
ax.set_title("Chi-Squared Acceptance/Rejection Regions for alpha = 0.05")
```

Out[9]: Text(0.5, 1.0, 'Chi-Squared Acceptance/Rejection Regions for alpha = 0.05')





Uniformly Most Powerful Test (UMP)

- A Uniformly Most Powerful (UMP) test has the most statistical power from the set of all possible alternate hypotheses of the same size α .
- The UMP doesn't always exist, especially when the test has nuisance variables (variables that are irrelevant to your study but that have to be be accounted for).
- · However, if the UMP does exist, you can use the Neyman-Pearson lemma (NPL) to find it.
- · A UMP test is usually defined in terms of a uniformly most powerful rejection region (UMPCR) (also called a "critical region").
 - A region C of size α is the UMPCR for testing a simple null hypothesis against a set of alternate hypotheses if it is the "best" critical region.
 - The "best" critical region is one that minimizes the probability of making a Type I or a Type II error.
 - In other words, the UMPCR is the region that gives the smallest chance of making a Type I or II error. It is also the region that gives a UMP test
 the largest (or equally largest) power function.
- Formal definition: a test defined by a critical region C of size α is a uniformly most powerful (UMP) test if it is a most powerful test against each simple alternative in the alternative hypothesis H_A . The critical region C is called a uniformly most powerful critical region of size α .
- UMP and the Neyman-Pearson Lemma:
 - The Neyman-Pearson lemma can tell you the best hypothesis test if you have a simple null hypothesis and a simple alternative hypothesis.
 - o If you have multiple hypotheses (also called a composite hypothesis), the NPL can be extended to all individual alternate hypotheses.
 - \circ Composite hypotheses have multiple options for solutions. For example, $\sigma^2 > 8$ is a composite hypothesis because it doesn't specify a value for σ^2
 - The basic idea is that you test each simple hypothesis in turn to see if it is the UMP out of all possibilities.
 - Formally:
 - Let C be a class of tests for testing H_0 : $\theta \in \Theta_0$ versus H_1 : $\theta \in \Theta_1^c$. A test in class C, with power function $\beta(\theta)$, is a uniformly most powerful (UMP) class C test if $\beta(\theta) \geq \beta'(\theta)$ for every $\theta \in \Theta_0^c$ and every $\beta'(\theta)$ that is a power function of a test in class C.
 - In more simple terms, this is really just telling you that the UMP test is the one with the biggest power function (out of all tests of the same size
 α). Θ₀ is the set of all possible values for θ under the null hypothesis.
 - The same statement can be rewritten using the **likelihood ratio test**. Let's say you had two simple hypotheses H_0 : $\theta = \theta_0$ and H_1 : $\theta = \theta_1$. In order to find the most powerful test at a certain α level (with threshold η), you would look for the likelihood-ratio test which rejects the null hypothesis in favor of the alternate hypothesis when:

$$\Lambda(x) = rac{L(x| heta_0)}{L(x| heta_1)} \leq \eta$$

Where

$$P(\Lambda(X) \le \eta | H_0) = \alpha$$

- $\circ \ L$ is the likelihood function.
- Exercise Example (https://newonlinecourses.science.psu.edu/stat414/node/308/)

Which Statistical Test To Use???

TYPE	TEST STATISTIC	DF.	APPLICATION
z	$Z = \frac{\overline{X} - \mu}{\frac{\sigma}{\sqrt{n}}}$	N.A.	Single sample mean. Standard deviation of population is known.
t Test	$t = \frac{\bar{X} - \mu}{\frac{S}{\sqrt{n}}}$	n - 1	Single sample mean. Standard deviation of population unknown.
2 Mean Equal Variance t Test	$t = \frac{\bar{X}_1 - \bar{X}_2}{*S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$	n ₁ + n ₂ - 2	2 sample means. Variances are unknown, but considered equal. ${}^{\star}S_p = \sqrt{\frac{\left(n_1-1\right)S_1^2+\left(n_2-1\right)S_2^2}{n_1+n_2-2}}$
2 Mean Unequal Variance t Test	$t = \frac{\overline{X}_1 - \overline{X}_2}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$		2 sample means. Variances are unknown, but considered unequal. DF is determined from the Welch-Satterthwaite
Paired t Test	$t = \frac{\overline{d}}{s_d / \sqrt{n}}$	n - 1	2 sample means. Data is taken in pairs. A different d is calculated for each pair.
X² σ² Known	$X^2 = \frac{(n-1)S^2}{\sigma^2}$	n-1	Tests sample variance against known variance.
Xs	$X^2 = \sum \frac{(O - E)^2}{E}$	(r-1)(c-1)	Compares observed and expected frequencies of test outcomes.
F	$F = \frac{S_1^2}{S_2^2}$	n, - 1 n ₂ - 1	Tests if two sample variances are equal.

- Image Source (http://isoconsultantpune.com/hypothesis-testing/)
- Video: https://www.youtube.com/watch?v=rullUAN0U3w (https://www.youtube.com/watch?v=rullUAN0U3w)



Recommended Videos



- These videos do not replace the lectures and tutorials.
- Please use these to get a better understanding of the material, and not as an alternative to the written material.

Video By Subject

- t-Test T-Tests: A Matched Pair Made in Heaven (https://www.youtube.com/watch?v=AGh66ZPpOSQ)
- χ^2 -Test Chi-squared Test (https://www.youtube.com/watch?v=WXPBoFDqNVk)



- Examples, exercises and definitions from Introduction to Probability, Statistics and Random Processes (https://probabilitycourse.com/) https://probabilitycourse.com) (<a href="https://probabilitycourse.com"
- Pearson's chi-squared test (https://medium.com/@geekrodion/statistics-pearsons-chi-squared-test-95fe36d74c1c) by Rodion Chachura
- <u>Uniformly Most Powerful (UMP) Test (https://www.statisticshowto.datasciencecentral.com/ump-uniformly-most-powerful/)</u> by Statistics How To
- $\bullet \ \ \textbf{lcons from} \ \underline{\textbf{lcon8.com} \ (\underline{\textbf{https://icons8.com/)}}} \ \ \underline{\textbf{https://icons8.com} \ (\underline{\textbf{https://icons8.com}}}$
- Datasets from Kaggle (https://www.kaggle.com/) https://www.kaggle.com/ (https://www.kaggle.com/)