

# EE4.10 Probability & Stochastic Processes Answers (2015)

B — Book work  
E — New example  
T — New theory

1. a) The pdf of  $X$  is

$$f_X(x) = \frac{1}{2\pi} \quad x \in [-\pi, \pi]$$

i)  $y = g(x) = x^3 \quad x = y^{1/3} \quad y \in [-\pi^3, \pi^3] \quad [1E]$

$$g'(x) = 3x^2 = 3y^{2/3} \quad [1E]$$

$$f_Y(y) = \frac{1}{|g'(x)|} f_X(x) = \frac{1}{3y^{2/3}} f_X(x)$$

$$= \frac{1}{6\pi y^{2/3}} \quad y \in [-\pi^3, \pi^3] \quad [1E]$$

ii)  $y = g(x) = x^4 \quad x_1 = y^{1/4} \quad y \in [0, \pi^4] \quad [1E]$

$$g'(x) = 4x^3 \quad x_2 = -y^{1/4}$$

$$f_Y(y) = \frac{1}{|g'(x_1)|} f_X(x_1) + \frac{1}{|g'(x_2)|} f_X(x_2) \quad [1E]$$

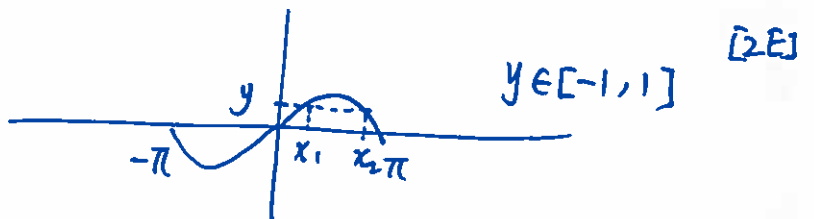
$$= \frac{1}{4y^{3/4}} [f_X(x_1) + f_X(x_2)]$$

$$= \frac{1}{4\pi y^{3/4}} \quad y \in [0, \pi^4] \quad [1E]$$

iii)  $y = g(x) = \sin(x) \quad x_1 = \sin^{-1}(y) \quad x_2 = \pi - \sin^{-1}(y)$

$$g'(x) = \cos x$$

$$= \sqrt{1-y^2}$$



$$f_Y(y) = \frac{1}{|g'(x_1)|} f_X(x_1) + \frac{1}{|g'(x_2)|} f_X(x_2) \quad [2E]$$

$$= \frac{1}{\sqrt{1-y^2}} \left( \frac{1}{2\pi} + \frac{1}{2\pi} \right)$$

$$= \frac{1}{\pi \sqrt{1-y^2}} \quad y \in [-1, 1]$$

b) i) Let  $X' = 2X$

Then  $f_{X'}(x') = \frac{1}{2} f_X(x) = \frac{1}{2} e^{-x'/2} \quad x' > 0 \quad [2E]$

$$\begin{aligned} f_Z(z) &= \int_0^z f_{X'}(z-y) f_Y(y) dy \\ &= \int_0^z \frac{1}{2} e^{-(z-y)/2} e^{-y} dy \\ &= \frac{1}{2} e^{-z/2} \int_0^z e^{-y/2} dy \\ &= \frac{1}{2} e^{-z/2} \cdot 2(1 - e^{-z/2}) \\ &= e^{-z/2} - e^{-z} \quad z > 0 \end{aligned} \quad [3E]$$

$$\begin{aligned} \text{ii) } F_Z(z) &= P\{\min(X, Y) \leq z\} \\ &= 1 - P\{\overset{X}{\cancel{X}} > z, Y > z\} \\ &= 1 - (1 - F_X(z))(1 - F_Y(z)) \\ &= F_X(z) + F_Y(z) - F_X(z)F_Y(z) \end{aligned} \quad [2E]$$

$$f_Z(z) = f_X(z) + f_Y(z) - f_X(z)F_Y(z) - F_X(z)f_Y(z)$$

We have  $f_X(z) = f_Y(z) = e^{-z} \quad z > 0 \quad [1E]$

$$F_Y(z) = F_X(z) = 1 - e^{-z} \quad z > 0$$

Thus,

$$\begin{aligned} f_Z(z) &= 2e^{-z} - 2e^{-z}(1 - e^{-z}) \\ &= 2e^{-2z} \quad z > 0 \end{aligned} \quad [2E]$$

iii)  $Z = \max(X, Y)$

$$F_Z(z) = P\{\max(X, Y) \leq z\}$$

[2E]

$$= P\{X \leq z, Y \leq z\}$$

$$= F_X(z) F_Y(z)$$

$$f_Z(z) = f_X(z) F_Y(z) + F_X(z) f_Y(z)$$

[3E]

$$= 2 e^{-z} (1 - e^{-z}) \quad z > 0$$

2. a) Let  $\bar{X}$  denote the average.

The joint density

$$f(X, c) = c^n e^{-cn(\bar{X} - x_0)} \quad [3E]$$

has maximum if

$$\frac{\partial f(X, c)}{\partial c} = 0 \Rightarrow \hat{c} = \frac{1}{\bar{X} - x_0} \quad [2E]$$

obviously,  $\bar{X} = 9$  in this problem. So

$$\hat{c} = \frac{1}{9-5} = \frac{1}{4} \quad [3E]$$

b) From the Wiener-Hopf equation,

$$c = R^T r \quad [2B]$$

$$\sigma^2 = r_0 - r^T R^T r$$

i) When  $n=1$ , we have

$$R = 1 \quad [2E]$$

$$r = r_1 = 0.643$$

Thus,

$$c_1 = r_1 = 0.643$$

$$\sigma^2 = 1 - r_1^2 = 1 - 0.643^2 = 0.587 \quad [3E]$$

ii) When  $n=2$ , we have

$$R = \begin{bmatrix} 1 & 0.643 \\ 0.643 & 1 \end{bmatrix}$$

[2E]

$$r = \begin{bmatrix} -0.055 \\ 0.643 \end{bmatrix}$$

Thus,

$$\begin{aligned} C = \begin{bmatrix} c_1 \\ c_2 \end{bmatrix} &= R^{-1} r \\ &= \begin{bmatrix} 1 & 0.643 \\ 0.643 & 1 \end{bmatrix}^{-1} \begin{bmatrix} -0.055 \\ 0.643 \end{bmatrix} \quad [4E] \\ &= \begin{bmatrix} -0.797 \\ 1.154 \end{bmatrix} \end{aligned}$$

$$\sigma^2 = 1 - \begin{bmatrix} -0.055 & 0.643 \end{bmatrix} \begin{bmatrix} -0.797 \\ 1.154 \end{bmatrix} \quad [4E]$$

$$= 0.214$$

3. a) i)  $X(n) = \cos(nu)$

$$E[X(n)] = E[\cos(nu)] = 0 \quad [2E]$$

$$\begin{aligned} E[X^2(n)] &= E[\cos^2(nu)] \\ &= E\left[\frac{1}{2}(1 + \cos(2nu))\right] = \frac{1}{2} \end{aligned}$$

$$\begin{aligned} E[X(m)X(n)] &= E[\cos(mu)\cos(nu)] \quad [3E] \\ &= E\left[\frac{1}{2}(\cos((m+n)u) + \cos((m-n)u))\right] \\ &= 0 \quad \text{if } m \neq n \end{aligned}$$

Therefore,  $X(n)$  is wide-sense stationary.

ii) Here, the answer is not unique.

For example, one may check

$$\begin{aligned} E[X(m)X(n)X(r)] &= E[\cos(mu)\cos(nu)\cos(ru)] \\ &= \frac{1}{2} E[(\cos((m+n)u) + \cos((m-n)u))\cos(ru)] \\ &= \frac{1}{4} E[\cos((m+n+r)u) + \cos((m+n-r)u) \\ &\quad + \cos((m-n+r)u) + \cos((m-n-r)u)] \\ &= \frac{1}{4} [\delta(m+n+r) + \delta(m+n-r) \quad [3T] \\ &\quad + \delta(m-n+r) + \delta(m-n-r)] \end{aligned}$$

where  $\delta(n) = 1$  if  $n = 0$

$\delta(n) = 0$  if  $n \neq 0$

Consider two cases  $(m, n, r) = (1, 2, 3), (2, 3, 4)$ . [2T]

They take different values  $\frac{1}{4}$  and  $0$ .

So it doesn't satisfy the definition of strict-sense stationarity (which would require the same values).

b)

i) This is the same as <sup>average</sup> the time when the third patient arrives.

$$E[T_3] = \frac{3}{\lambda} = \frac{3}{0.1} = 30 \text{ minutes} \quad [3E]$$

ii) This means that the number of patients arrived in the first hour is less than three

$$\begin{aligned} P(N(t) < 3) & \quad t = 60 \text{ minutes} \\ &= P(N(t) = 0) + P(N(t) = 1) + P(N(t) = 2) \quad [3E] \\ &= e^{-60/10} + \left(\frac{60}{10}\right) e^{-60/10} + \frac{1}{2} \left(\frac{60}{10}\right)^2 e^{-60/10} \\ &= 25 \cdot e^{-6} \quad [3E] \\ &= 0.062 \end{aligned}$$

Recall Poisson:  $P(N(t) = k) = e^{-\lambda t} \frac{(\lambda t)^k}{k!} \quad k = 0, 1, 2, \dots$

iii) Poisson process is memoryless. So this probability is given by

$$\begin{aligned} & P(N(t_1) \geq 2) \cdot P(N(t_2 - t_1) \leq 2) \quad \begin{array}{l} t_1 = 60 \text{ minutes} \\ t_2 = 120 \end{array} \\ &= [1 - P(N(t_1) < 2)] P(N(t_2 - t_1) \leq 2) \quad [3E] \\ &= \{1 - [P(N(t_1) = 0) + P(N(t_1) = 1)]\} P(N(t_2 - t_1) \leq 2) \\ &= [1 - (e^{-6} + 6e^{-6})] \cdot 25 \cdot e^{-6} \quad (\text{from ii}) \quad [3E] \\ &= 0.061 \end{aligned}$$

4. a) i) For any sequence of states  $i_0, i_1, \dots$

$$\begin{aligned}
 & P(Y_{r+1} = i_{r+1} \mid Y_r = i_r, Y_{r-1} = i_{r-1}, \dots, Y_0 = i_0) \\
 &= \frac{P(Y_{r+1} = i_{r+1}, Y_r = i_r, \dots, Y_0 = i_0)}{P(Y_r = i_r, \dots, Y_0 = i_0)} \quad [2T] \\
 &= \frac{\prod_{s=0}^r P_{i_s, i_{s+1}}(n_{s+1} - n_s)}{\prod_{s=0}^{r-1} P_{i_s, i_{s+1}}(n_{s+1} - n_s)} \\
 &= P_{i_r, i_{r+1}}(n_{r+1} - n_r) \quad [2T] \\
 &= P(Y_{r+1} = i_{r+1} \mid Y_r = i_r)
 \end{aligned}$$

ii) The transition matrix is given by

$$P^2 = \begin{pmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} & 0 \end{pmatrix} \begin{pmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} & 0 \end{pmatrix} = \begin{pmatrix} \frac{1}{2} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{1}{2} & \frac{1}{4} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{pmatrix} \quad [3E]$$

b) Let  $S_{n+1} = S_n + Z_{n+1}$  where

$$P(Z_{n+1} = 1) = p, \quad P(Z_{n+1} = -1) = q.$$

Thus,

$$\begin{aligned}
 E[Y_{n+1} \mid Y_n, \dots, Y_0] &= E\left[\left(\frac{q}{p}\right)^{S_{n+1}} \mid S_n, S_{n-1}, \dots, S_0\right] \quad [2B] \\
 &= E\left[\left(\frac{q}{p}\right)^{S_n + Z_{n+1}} \mid S_n\right] \quad \text{Markov} \\
 &= \left(\frac{q}{p}\right)^{S_n} \left[\frac{q}{p} \cdot p + \left(\frac{q}{p}\right)^{-1} \cdot q\right] \\
 &= \left(\frac{q}{p}\right)^{S_n} \\
 &= Y_n \quad [2B]
 \end{aligned}$$

i.e.,  $\{Y_n\}$  is a martingale.



ii) Let  $T$  denote the stopping time.

$$E[Y_T] = E[Y_0] = \left(\frac{q}{p}\right)^i$$

Since  $Y_0 = i$ .

We also have

$$E[Y_T] = p_i \left(\frac{q}{p}\right)^0 + (1-p_i) \left(\frac{q}{p}\right)^N \quad [2B]$$

$$= p_i + (1-p_i) \left(\frac{q}{p}\right)^N$$

Thus

$$p_i = \frac{1 - \left(\frac{q}{p}\right)^{N-i}}{1 - \left(\frac{q}{p}\right)^N} \quad \text{if } \frac{q}{p} \neq 1 \quad [2B]$$

$$p_i = 1 - \frac{i}{N} \quad \text{if } p = q = \frac{1}{2}$$

c) From the iteration  $T_k = 1 + p T_{k+1} + q T_{k-1}$ , we have

$$p (T_{i+1} - T_i) = q (T_i - T_{i-1}) - 1 \quad [2T]$$

Let  $M_{i+1} = T_{i+1} - T_i$ , we obtain iteration

$$M_{i+1} = \frac{q}{p} M_i - \frac{1}{p}$$

$$= \begin{cases} \left(\frac{q}{p}\right)^i M_1 - \frac{1 - (q/p)^i}{p - q} & p \neq q \\ M_1 - \frac{i}{p} & p = q \end{cases} \quad [2T]$$

This gives

$$T_i = \sum_{k=0}^{i-1} M_{k+1}$$

$$= \begin{cases} \left(M_1 + \frac{1}{p-q}\right) \sum_{k=0}^{i-1} \left(\frac{q}{p}\right)^k - \frac{i}{p-q} & p \neq q \\ i M_1 - \frac{i(i-1)}{2p} & p = q \end{cases} \quad [2T]$$

We yet need to determine  $M_1$  from initial conditions

$$T_0 = 0$$

$$T_N = 0$$

Which gives

$$M_1 + \frac{1}{p-q} = \frac{N}{p-q} \cdot \frac{1 - q/p}{1 - (q/p)^N} \quad p \neq q$$

Thus

$$\begin{aligned} T_i &= \frac{N}{p-q} \cdot \frac{1 - q/p}{1 - (q/p)^N} \cdot \frac{1 - (q/p)^i}{1 - q/p} - \frac{i}{p-q} \quad [2T] \\ &= \left\{ \frac{N}{p-q} \cdot \frac{(1 - q/p)^i}{(1 - q/p)^N} - \frac{i}{p-q} \right\} \quad p \neq q \end{aligned}$$

Similarly, if  $p = q$

$$T_N = 0 = NM_1 - \frac{N(N-1)}{2p} = 0$$

$$M_1 = \frac{N-1}{2p}$$

$$\begin{aligned} T_i &= \frac{i(N-1)}{2p} - \frac{i(i-1)}{2p} \quad [2T] \\ &= \frac{i(N-i)}{2p} \quad p = \frac{1}{2} \\ &= i(N-i) \end{aligned}$$