#### 1. Solution:

If we denote:

D={diseased}, H={healthy}, and +={positive}, then from the conditions, we know that:

$$P(D) = 10^{-5}$$
 ,  $P(H) = 1 - 10^{-5}$ 

$$P(+|D| = 0.99$$
 ,  $P(+|H| = 0.01$ 

Based on these, we can use Bayes' Formula to calculate

$$P(D \mid +) = \frac{P(+ \mid D)P(D)}{P(+ \mid D)P(D) + P(+ \mid H)P(H)}$$
$$= \frac{0.99 \times 10^{-5}}{0.99 \times 10^{-5} + 0.01 \times (1 - 10^{-5})} = 9.89 \times 10^{-4}$$

The value is pretty small

## Extension:

How about the probability of diseased, if two independent tests are both positive, i.e.  $P(D \mid +, +)$ ?

$$P(D|+,+) = \frac{P(D)P(+,+|D)}{P(D)P(+,+|D)+P(H)P(+,+|H)}$$
$$= \frac{P(D)P(+|D)P(+|D)}{P(D)P(+|D)P(+|D)+P(H)P(+|H)P(+|H)}$$

### 2. Solution:

1) If we denote there is an open road between A and B as  ${\cal O}_{_{AB}}$  , and similarly  ${\cal O}_{_{BC}}$  and  ${\cal O}_{_{AC}}$  , then

$$P(O_{AC}) = P(O_{AB} \cap O_{BC})$$
  
=  $P(O_{AB}) \cdot P(O_{BC})$  (independence)

Furthermore,  $P(O_{AB}) = 1 - P(\text{two roads are blocked}) = 1 - P^2$ 

So, 
$$P(O_{AC}) = (1 - P^2)^2$$

2)

Direct road 
$$p$$
  $P(O_{AC} | \text{Direct road blocked}) = (1 - p^2)^2$ 

Direct road  $p$   $P(O_{AC} | \text{Direct road open}) = 1$ 

$$P(O_{AC}) = P(O_{AC} | \text{Direct road blocked}) \cdot p + P(O_{AC} | \text{Direct road open}) \cdot (1 - p)$$
$$= (1 - p^2)^2 \cdot p + (1 - p)$$

#### 3. Solution:

We denote:

A: exactly one Ace, and KK: exactly two Kings

We want to calculate:  $P(A \mid KK) = \frac{P(A \cap KK)}{P(KK)}$ ,

where 
$$P(A \cap KK) = \frac{\binom{4}{1}\binom{4}{2}\binom{44}{10}}{\binom{52}{13}}$$
, and  $P(KK) = \frac{\binom{4}{2}\binom{48}{11}}{\binom{52}{13}}$ .

So, 
$$P(A|KK) \approx 0.44$$

## 4. Solution:

We denote:

S: their sum is 7

X: the score of the first dice is X, where X=1, 2, 3, 4, 5, 6

We want to show that  $P(S \cap X) = P(S) \cdot P(X)$  (so, S and X are independent)

$$P(S \cap X) = \frac{1}{36}$$
 for any X

$$P(S) = \sum_{X=1}^{6} P(S \mid X) P(X) = \sum_{X=1}^{6} \frac{1}{6} \cdot \frac{1}{6} = \frac{1}{6}$$

$$P(X) = \frac{1}{6}$$
 for any X

So, 
$$P(S \cap X) = P(S) \cdot P(X)$$
 is true.

# 5. Solution:

If we denote the label of the k-th card as  $L_{\iota}$  , then

"the label of the k-th card dealt is the largest of the first k cards dealt" can be described as

$$\{L_{\iota} > L_{r}, \text{ for any } 1 \leq r \leq k-1\}$$

"it is also the largest in the pack" means  $L_k = m$ 

What we want to calculate is  $P\left(L_k = m \middle| L_k > L_r, \text{ for any } 1 \leq r \leq k-1\right)$ 

$$=\frac{P\left(L_{k}=m,\,\mathrm{and}\,\,L_{k}>L_{r},\,\mathrm{for\,\,any}\,\,1\leq r\leq k-1\right)}{P\left(L_{k}>L_{r},\,\mathrm{for\,\,any}\,\,1\leq r\leq k-1\right)}$$
 
$$1/m\qquad k$$

$$=\frac{1/m}{1/k}=\frac{k}{m}$$

## 6. We denote:

 $A_i$ : Professor i leaves with his own hat

 $A_i^c$ : Professor i leaves without his own hat

So, what we want to calculate is  $P\left(A_1^c\cap\cdots\cap A_n^c\right)$  , which equals to  $1-P\left(A_1\cup\cdots\cup A_n\right)$  , where

$$P\left(A_{1} \cup \cdots \cup A_{n}\right) = P\left(A_{1}\right) + P\left(A_{2}\right) + \cdots + P\left(A_{n}\right)$$

$$-P\left(A_{1}A_{2}\right) - P\left(A_{1}A_{3}\right) - \cdots - P\left(A_{n-1}A_{n}\right)$$

$$+P\left(A_{1}A_{2}A_{3}\right) + \cdots + P\left(A_{n-2}A_{n-1}A_{n}\right)$$

$$-\cdots$$

$$+\left(-1\right)^{n+1}P\left(A_{1} \cdots A_{n}\right)$$

$$= \frac{\left(n-1\right)!}{n!} + \frac{\left(n-1\right)!}{n!} + \cdots + \frac{\left(n-1\right)!}{n!}$$

$$-\frac{\left(n-2\right)!}{n!} - \frac{\left(n-2\right)!}{n!} - \cdots - \frac{\left(n-2\right)!}{n!}$$

$$+ \frac{\left(n-3\right)!}{n!} + \cdots + \frac{\left(n-3\right)!}{n!}$$

$$-\cdots$$

$$+\left(-1\right)^{n+1}\frac{\left(n-n\right)!}{n!}$$
The term  $\frac{\left(n-k\right)!}{n!}$  appears  $\binom{n}{k}$  times, so
$$P\left(A_{1} \cup \cdots \cup A_{n}\right) = \binom{n}{1}\frac{\left(n-1\right)!}{n!} - \binom{n}{2}\frac{\left(n-2\right)!}{n!} + \cdots + \left(-1\right)^{n+1}\binom{n}{n}\frac{\left(n-n\right)!}{n!},$$

$$= \frac{1}{1!} - \frac{1}{2!} + \frac{1}{3!} - \cdots + \frac{\left(-1\right)^{n+1}}{n!}$$

then

$$P\left(A_{1}^{c} \cap \cdots \cap A_{n}^{c}\right) = 1 - \frac{1}{1!} + \frac{1}{2!} - \frac{1}{3!} - \cdots + (-1)^{n} \frac{1}{n!} = e^{-1} \text{ when } n \to \infty.$$

### 7. Solution:

Symmetric random walk: "Gambler's ruin"

We denote:

A: the gambler is finally bankrupted

B: the first toss of coin shows heads (the manager pays the gambler one unit)

k: starting point (with k units money)

From the Total Probability Theorem, we know that

$$P_{k}(A) = P_{k}(A \mid B) P(B) + P_{k}(A \mid B^{c}) P(B^{c}),$$
(\*)

where  $P_k(A)$  means "the gambler finally bankrupted when starting point is k".

 $P_k(A \mid B) = P_{k+1}(A)$  because event B happens means the money gambler has becomes k+1.

Similarly: 
$$P_k(A \mid B^c) = P_{k-1}(A)$$

So, (\*) 
$$\Rightarrow P_k(A) = P_{k+1}(A) \cdot \frac{1}{2} + P_{k-1}(A) \cdot \frac{1}{2}$$
, where  $P(B) = P(B^c) = \frac{1}{2}$ 

With boundary  $P_{_{0}}\left(\,A\,
ight)=1,\;\;P_{_{N}}\left(\,A\,
ight)=\,0$  , we will have

$$\Rightarrow P_k(A) = 1 - \frac{k}{N}$$
.

Details: 
$$P_k(A) = P_{k+1}(A) \cdot \frac{1}{2} + P_{k-1}(A) \cdot \frac{1}{2}$$

$$\Rightarrow 2P_k(A) = P_{k+1}(A) + P_{k-1}(A)$$

$$\Rightarrow P_{k+1}(A) - P_k(A) = P_k(A) - P_{k-1}(A)$$

$$\Rightarrow b_k = b_{k-1}$$
, where  $b_k = P_{k+1}(A) - P_k(A)$ 

$$\Rightarrow P_N(A) = Nb_1 + P_0(A)$$

$$\Rightarrow b_1 = -\frac{1}{N} \Rightarrow P_k(A) = 1 - \frac{k}{N}.$$