

$$Y = \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix}_{n \times 1} \quad X = \begin{bmatrix} X_1, \dots, X_{1,p} \\ \vdots \\ X_{n, \dots, X_{n,p}} \end{bmatrix}_{n \times p} \quad \varepsilon = \begin{bmatrix} \varepsilon_1 \\ \vdots \\ \varepsilon_n \end{bmatrix} \quad \varepsilon \sim N(0, \sigma^2 I_n)$$

$$\beta = \begin{bmatrix} \beta_1 \\ \vdots \\ \beta_p \end{bmatrix}_{p \times 1}$$

$$Y = X\beta + \varepsilon$$

Ridge:

$$\begin{aligned} \hat{\beta}_\lambda &= \frac{1}{n+\lambda} X^T Y = \frac{1}{n+\lambda} X^T (X\beta + \varepsilon) \\ &= \frac{1}{n+\lambda} X^T X \beta + \frac{1}{n+\lambda} X^T \varepsilon \quad (X^T X = n I_p) \\ &= \underbrace{\frac{n}{n+\lambda} \beta}_{\text{constant}} + \underbrace{\frac{1}{n+\lambda} X^T \varepsilon}_{\text{random}} \end{aligned}$$

Bias: (It is a  $\mathbb{R}^p$  vector here)

$$\Rightarrow \mathbb{E} \hat{\beta}_\lambda = \frac{n}{n+\lambda} \beta + \frac{1}{n+\lambda} \mathbb{E}[X^T \varepsilon] \underset{=0}{=} \frac{n}{n+\lambda} \cdot \beta \quad \textcircled{1}$$

$$\text{Bias} = \mathbb{E} \hat{\beta}_\lambda - \beta = \frac{-\lambda}{n+\lambda} \beta \quad (\text{if } \lambda=0 \Rightarrow \text{OLS . 0-bias})$$

Cov: (It is a  $p \times p$  matrix here)

$$\text{Cov}(\hat{\beta}_\lambda) = \text{Cov}\left(\underbrace{\frac{n}{n+\lambda} \beta}_{\text{constant}} + \underbrace{\frac{1}{n+\lambda} X^T \varepsilon}_{\text{R.V.}}\right) = \text{Cov}\left(\frac{1}{n+\lambda} X^T \varepsilon\right)$$

For a random vector  $U \in \mathbb{R}^P$ , the cov-matrix is:

$$\text{cov}(U) = \mathbb{E}[(U - \mathbb{E}U)(U - \mathbb{E}U)^T] \quad (\text{P} \times \text{P} \text{ matrix})$$

$$(U = \frac{1}{n+\lambda} X^T \varepsilon. \text{ here, } \mathbb{E}\left[\frac{1}{n+\lambda} X^T \varepsilon\right] = 0)$$

$$\begin{aligned} \text{cov}\left(\frac{1}{n+\lambda} X^T \varepsilon\right) &= \mathbb{E}\left[\left(\frac{1}{n+\lambda} X^T \varepsilon\right)\left(\frac{1}{n+\lambda} X^T \varepsilon\right)^T\right] \\ &= \frac{1}{(n+\lambda)^2} \mathbb{E}\left[(X^T \varepsilon)(X^T \varepsilon)^T\right] = \frac{1}{(n+\lambda)^2} \mathbb{E}[X^T \varepsilon \varepsilon^T X] \\ &= \frac{1}{(n+\lambda)^2} X^T \cdot \mathbb{E}[\varepsilon \varepsilon^T] X \xrightarrow{\mathbb{E}[\varepsilon \varepsilon^T] = \sigma^2 I} \frac{1}{(n+\lambda)^2} X^T \mathbb{E}[\varepsilon - \mathbb{E}\varepsilon](\varepsilon - \mathbb{E}\varepsilon)^T X \\ &= \frac{1}{(n+\lambda)^2} X^T \text{cov}(\varepsilon) X = \frac{\sigma^2}{(n+\lambda)^2} X^T X = \frac{n\sigma^2}{(n+\lambda)^2} \cdot I_P \quad (2) \end{aligned}$$

$\sigma^2 I_n$  by def

$$\text{Thus, } \text{cov}(\hat{\beta}_\lambda) = \text{cov}\left(\frac{1}{n+\lambda} X^T \varepsilon\right) = \frac{n\sigma^2}{(n+\lambda)^2} \cdot I_P$$

( $\lambda=0 \Rightarrow \text{OLS cov-matrix}$ )

Error: ( $\lambda$  is a scalar here)

$$\mathbb{E} \|\hat{\beta}_\lambda - \beta\|_2^2 = \mathbb{E} (\hat{\beta}_\lambda - \beta)^T (\hat{\beta}_\lambda - \beta)$$

$$= \mathbb{E} \sum_{i=1}^P (\hat{\beta}_{\lambda,i} - \beta_i)^2$$

$$= \sum_{i=1}^P \mathbb{E} [(\hat{\beta}_{\lambda,i} - \beta_i)^2]$$

$$\hat{\beta}_\lambda = \begin{bmatrix} \hat{\beta}_{\lambda,1} \\ \vdots \\ \hat{\beta}_{\lambda,P} \end{bmatrix} \quad \beta = \begin{bmatrix} \beta_1 \\ \vdots \\ \beta_P \end{bmatrix}$$

Using bias-var decomposition:

$$\hat{\beta}_{\lambda,i} - \beta_i = (\hat{\beta}_{\lambda,i} - \mathbb{E}\hat{\beta}_{\lambda,i}) + (\mathbb{E}\hat{\beta}_{\lambda,i} - \beta_i)$$

$$\mathbb{E}[(\hat{\beta}_{\lambda,i} - \beta_i)^2] = \underbrace{\mathbb{E}[(\hat{\beta}_{\lambda,i} - \mathbb{E}\hat{\beta}_{\lambda,i})^2]}_{\text{Var of } \hat{\beta}_{\lambda,i}} + \underbrace{(\mathbb{E}\hat{\beta}_{\lambda,i} - \beta_i)^2}_{\text{Bias of } \hat{\beta}_{\lambda,i}}$$

By the definition of cov-matrix

Var of  $\hat{\beta}_{\lambda,i}$  is the (i,i)-entry of cov-matrix in  $\textcircled{2}$

$$= \frac{h\sigma^2}{(n+\lambda)^2}$$

Bias of  $\hat{\beta}_{\lambda,i}$  is the square of i-th element of bias vector in  $\textcircled{1}$

$$= \frac{\lambda^2}{(n+\lambda)^2} \beta_i^2$$

$$\Rightarrow \mathbb{E}(\|\hat{\beta}_{\lambda} - \beta\|_2^2) = \sum_{i=1}^p \mathbb{E}(\hat{\beta}_{\lambda,i} - \beta_i)^2$$

$$= \sum_{i=1}^p (\text{Var of } \hat{\beta}_{\lambda,i} + \text{Bias of } \hat{\beta}_{\lambda,i})$$

$$= \sum_{i=1}^p \left( \frac{h\sigma^2}{(n+\lambda)^2} + \frac{\lambda^2}{(n+\lambda)^2} \beta_i^2 \right) = \frac{hp\sigma^2}{(n+\lambda)^2} + \frac{\lambda^2}{(n+\lambda)^2} \sum_{i=1}^p \underbrace{\beta_i^2}_{\|\beta\|_2^2}$$

In matrix form, we have

$$\mathbb{E} \| \hat{\beta}_\lambda - \beta \|_2^2 = \mathbb{E} [(\hat{\beta}_\lambda - \beta)^T (\hat{\beta}_\lambda - \beta)]$$

$$\stackrel{(*)}{=} \mathbb{E} [\text{tr} \{ (\hat{\beta}_\lambda - \beta)^T (\hat{\beta}_\lambda - \beta) \}]$$

$$= \mathbb{E} [\text{tr} \{ (\hat{\beta}_\lambda - \beta) (\beta_\lambda - \beta)^T \}]$$

$$= \mathbb{E} [\text{tr} \{ ((\hat{\beta}_\lambda - \mathbb{E} \hat{\beta}_\lambda) + (\mathbb{E} \hat{\beta}_\lambda - \beta)) (\hat{\beta}_\lambda - \mathbb{E} \hat{\beta}_\lambda) + (\mathbb{E} \hat{\beta}_\lambda - \beta))^T \}]$$

$$\stackrel{(\Delta)}{=} \text{tr} \{ \mathbb{E} [(\hat{\beta}_\lambda - \mathbb{E} \hat{\beta}_\lambda) (\hat{\beta}_\lambda - \mathbb{E} \hat{\beta}_\lambda)^T] + (\mathbb{E} \hat{\beta}_\lambda - \beta) (\mathbb{E} \hat{\beta}_\lambda - \beta)^T \}$$

$$\stackrel{(**)}{=} \text{tr} ( \text{Cov}(\hat{\beta}_\lambda) ) + \text{tr} ( (\mathbb{E} \hat{\beta}_\lambda - \beta)^T (\mathbb{E} \hat{\beta}_\lambda - \beta) )$$

$$= \frac{n \rho \sigma^2}{(n \rho \lambda)^2} + \frac{\lambda^2}{(n \rho \lambda)^2} \| \beta \|_2^2$$

In (\*) and (\*\*) we use

$$U^T V = \text{tr} (U^T V) = \text{tr} (V U^T) \quad (U, V \in \mathbb{R}^P)$$

In (\Delta) we use  $\text{tr}(A+B) = \text{tr}(A) + \text{tr}(B)$   
 $A, B \in \mathbb{R}^{P \times P}$  (matrix)

