# KDE and BoxCox

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## **Theory**

#### **Kernel Density Estimation**

A procedure used for estimating a probability density function using the observed data is considered. As with histograms, the procedure is used to provide a graphical representation of the pdf of X using the observed data,  $x_1, x_2, \ldots, x_n$ . The estimate is called the *kernel density estimate*. Kernel density estimation is a nonparametric technique for density estimation in which a known density function (the kernel) is averaged across the observed data points to create a smooth approximation for  $f_X(x)$ . SAS PROC KDE uses a Gaussian density as the kernel, and its assumed variance determines the smoothness of the resulting estimate. See Silverman (1986) for a thorough review and discussion.

#### **Computational Methods**

#### **Univariate Kernel Density Estimates - SAS**

Let  $(X_i, W_i)$ , denote the observed sample of  $X_i$  with specified weight  $W_i$  for i = 1, 2, ..., n. The weighted kernel density estimate of f(x), the density of X, is

$$\hat{f}(x) = \frac{1}{\sum_{i=1}^{n} W_i} \sum_{i=1}^{n} \varphi_h(x - X_i)$$

where h is the bandwidth and

$$\varphi_h(x) = \frac{1}{\sqrt{2\pi}h} \exp\left(-\frac{x^2}{2h^2}\right)$$

is the normal density rescaled by the bandwidth  $(N(0,h^2))$ . If  $h\to 0$  and  $nh\to \infty$ , then the optimal bandwidth is

$$h_{AMISE} = \left[\frac{1}{2\sqrt{\pi}n\int(f'')^2}\right]^{1/5}$$

where  $\frac{\partial^2 f}{\partial x} = f''$ . Since the optimal value is unknown, approximations methods are needed. For a derivation and discussion of these methods, see Silverman (1986) and Jones, Marron, and Sheather (1996).

#### **General Univariate Kernel Density Estimate**

Assume that  $W_i = 1$ , in which case the kernel density estimate for f(x) is,

$$\hat{f}(x) = \frac{1}{nh} \sum_{i=1}^{n} K(\frac{x - X_i}{h})$$

where the kernel, K, satisfies  $\int K(x)dx = 1$  and the smoothing parameter, h, is called the bandwidth. In practice, the kernel is an unimodal function satisfying,  $\int xK(x)dx = 0$ . A popular choice for the *normal* kernel given by

$$K(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right).$$

If one assumes that the underlying density is sufficiently smooth and that the kernel has finite forth moments, then an asymptotic expansion for the bias and variance of a kernel estimate are given by

$$Bias_{asy}\{\hat{f}_h(x)\} = \frac{h^2}{2}\mu_2(K)^2 f''(x)$$

and

$$Var_{asy}\{\hat{f}_h(x)\} = \frac{1}{nh}R(K)f(x)$$

where  $R(K) = \int K^2(y) dy$ ,  $\mu_2(K) = \int y^2 K(y) dy$  and f'' is the second derivative of f.

A widely used criteria for measuring the discrepancy between f and  $\hat{f}$  is the mean integrated squared error (MISE) is given by,

$$MISE(\hat{f}) = E\{ \int (f(y) - \hat{f}(y))^2 dy \}$$
$$= \int Bias(\hat{f}(y))^2 dy + \int Var(\hat{f}(y)) dy.$$

If one assumes that the function f is integrable, then the asymptotic mean integrated squared error (AMISE) is given by

$$AMISE(\hat{f}(h) = \frac{1}{nh}R(K) = \frac{h^2}{4}\mu_2(K)^2R(f'').$$

The bandwidth that minimizes the AMISE is given by

$$h_{AMISE} = \left\{ \frac{R(K)}{\mu_2(K)^2 R(f'')} \right\}^{1/3} n^{-1/3}.$$

#### **Bandwidth Selection**

Several different bandwidth selection methods are available in PROC KDE in the univariate case. Following the recommendations of Jones, Marron, and Sheather (1996), the default method follows a plug-in formula of Sheather and Jones.

This method solves the fixed-point equation

$$h = \left[\frac{R(\varphi)}{nR\left(\hat{f}_{g(h)}^{"}\right)\left(\int x^{2}\varphi(x)dx\right)^{2}}\right]^{1/5}$$

<sup>&</sup>lt;sup>1</sup>Note: the kernel, K(x) is a unimodal density function with expected value,  $E_K(X) = \int xK(x)dx = 0$ .

where  $R(\varphi) = \int \varphi^2(x) dx$  and  $g(h) = C(K)[R(f'')/R(f''')]^{1/7}h^{5/7}$  is the bandwidth for the estimate of  $R(\hat{f}'')$ .

PROC KDE solves this equation by first evaluating it on a grid of values spaced equally on a log scale. The largest two values from this grid that bound a solution are then used as starting values for a bisection algorithm. The simple normal reference rule works by assuming  $\hat{f}$  is Gaussian in the preceding fixed-point equation. This results in

$$h = \hat{\sigma}[4/(3n)]^{1/5}$$
$$= 1.06 \,\hat{\sigma}n^{-1/5}$$

where  $\hat{\sigma}$  is the sample standard deviation.

Alternatively, the bandwidth can be computed using the interquartile range,

$$\begin{array}{rcl} h & = & 1.06 \hat{\sigma} n^{-1/5} \\ & \approx & 1.06 \hat{\sigma} n^{-1/5} \\ & \approx & 1.06 \; (Q/1.34) n^{-1/5} \end{array}$$

Silverman's rule of thumb (Silverman, 1986, Section 3.4.2) is computed as

$$h = 0.9 \min[\hat{\sigma}, Q/1.34] n^{-1/5}$$

The oversmoothed bandwidth is computed as

$$h = 3\hat{\sigma}[1/(70\sqrt{\pi}n)]^{1/5}$$

When you specify a WEIGHT variable, PROC KDE uses weighted versions of  $Q_3$ ,  $Q_1$ , and  $\hat{\sigma}$  in the preceding expressions. The weighted quartiles are computed as weighted order statistics, and the weighted variance

$$\hat{\sigma}^2 = \frac{\sum_{i=1}^n W_i (X_i - \bar{X})^2}{\sum_{i=1}^n W_i}$$

where  $\bar{X} = (\sum_{i=1}^n W_i X_i)/(\sum_{i=1}^n W_i)$  is the weighted sample mean.

#### **Box Cox Transformations**

Suppose that y>0, define the Box-Cox transformation as

$$y_i^{(\lambda)} = \left\{ egin{array}{ll} (y_i^{\lambda} - 1)/\lambda & & when \quad \lambda 
eq 0, \\ \ln y_i & & when \quad \lambda = 0, \end{array} \right.$$

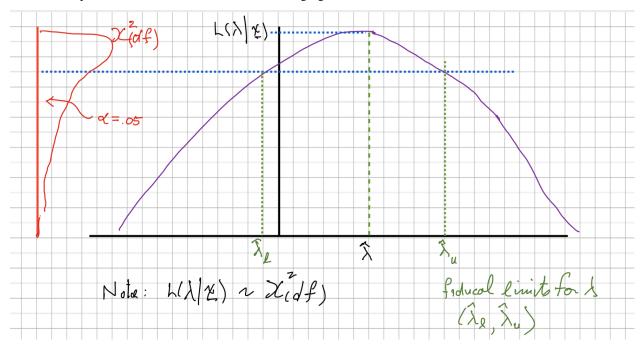
where  $i = 1, 2, \dots, n$ . One determines  $\lambda$  by maximizing

$$-n/2\log[s^{2}(\lambda)] = (\lambda - 1)\sum_{i=1}^{n} \ln(y_{i}) - n/2\log[\hat{\sigma}^{2}(\lambda)],$$

and  $\hat{\sigma}^2(\lambda) = 1/n\vec{y}^{(\lambda)\prime}[I-H]\vec{y}^{(\lambda)}$  i.e., it is the sum of squares for the error term when  $y_i^{(\lambda)}$  is used instead of  $y_i$  and  $\vec{y}^{(\lambda)} = (y_1^{(\lambda)}, y_2^{(\lambda)}, \dots, y_n^{(\lambda)})'$ .

Since, there is not a close form solution to the above maximization, one usually plots  $-n/2\log[s^2(\lambda)]$  vs  $\lambda$ . Another approach is to compute a confidence interval using the fact that  $-n/2\log[s^2(\lambda)] \sim \chi^2(df=1)$ . One can then use any  $\lambda$  which is contained in the confidence interval.

Often the output for the BoxCox results is the following figure.



#### **Comment**

I did not perform the BoxCox method in SAS.

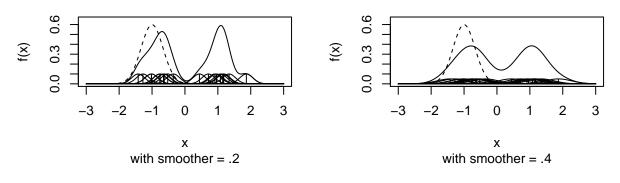
# R Code and Output

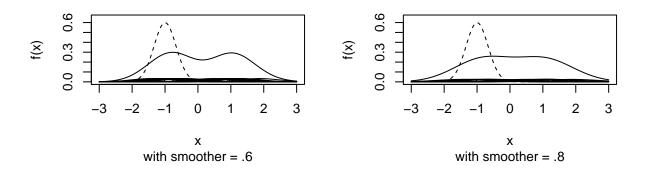
#### **Sheather Simulated KDE Example**

```
ysum = numeric(601)
for (i in 1:n)
  {points(x[i], 1/(n*h*sqrt(2*pi)),type="h")
     x1 = numeric(601)+x[i]
     y = (1/(h*sqrt(2*pi)))*exp(-0.5*((xx-x1)/h)^2)
     ysum = y/n + ysum
     lines(xx,y/n,lty=1)}
lines(xx,ysum,lty=1)
lines(xx,truedensity,lty=2)
}

par(mfrow=c(2,2))
sheather.curve(.2, "Sheather Bimodal Data", "with smoother = .2")
sheather.curve(.4, " ", "with smoother = .4")
sheather.curve(.8, " ", with smoother = .6")
sheather.curve(.8, " ", "with smoother = .8")
```

#### **Sheather Bimodal Data**



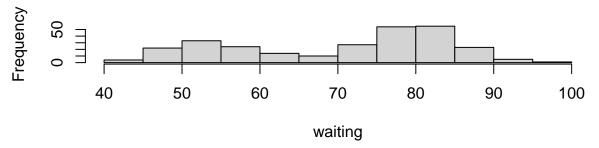


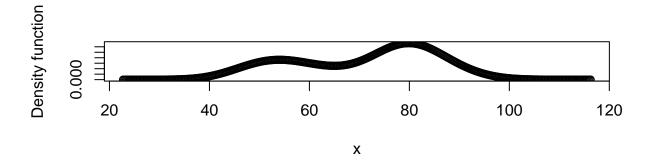
### Old Faithful geyser data

#### **Waiting Time**

```
par(mfrow=c(2,1))
library(KernSmooth)
attach(faithful)
hist(x=waiting)
fhat <- bkde(x=waiting)
plot (fhat, xlab="x", ylab="Density function")</pre>
```

# **Histogram of waiting**

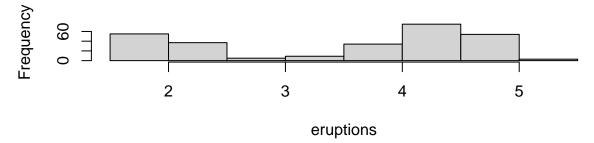


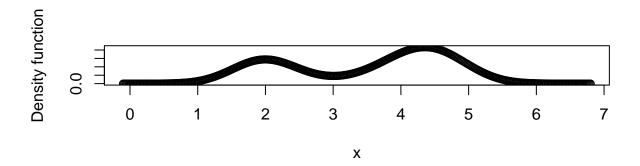


# **Eruption Time**

```
par(mfrow=c(2,1))
hist(x=eruptions)
fhat <- bkde(x=eruptions)
plot (fhat, xlab="x", ylab="Density function")</pre>
```

## **Histogram of eruptions**

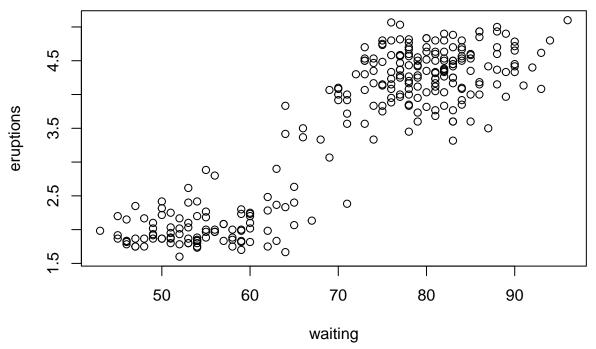




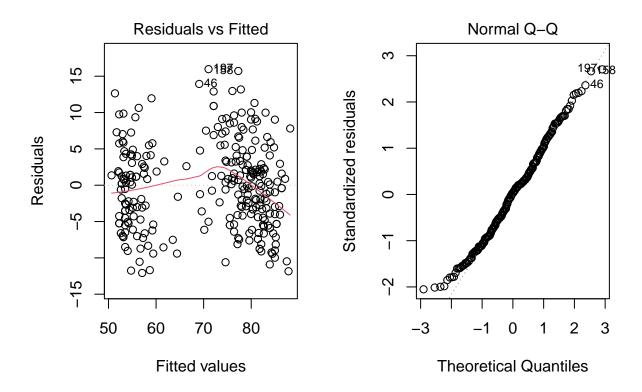
#### Regression

```
mod1 = lm(waiting ~ eruptions, data=faithful)
summary (mod1)
##
## Call:
## lm(formula = waiting ~ eruptions, data = faithful)
##
## Residuals:
       Min
                 1Q
                       Median
##
                                    3Q
## -12.0796 -4.4831
                      0.2122
                                3.9246 15.9719
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) 33.4744
                           1.1549
                                     28.98
                                             <2e-16 ***
## eruptions
              10.7296
                           0.3148
                                     34.09
                                             <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 5.914 on 270 degrees of freedom
## Multiple R-squared: 0.8115, Adjusted R-squared: 0.8108
## F-statistic: 1162 on 1 and 270 DF, p-value: < 2.2e-16
covb = vcov (mod1)
coeff.mod1 = coef(mod1)
covb = vcov (mod1)
covb
```

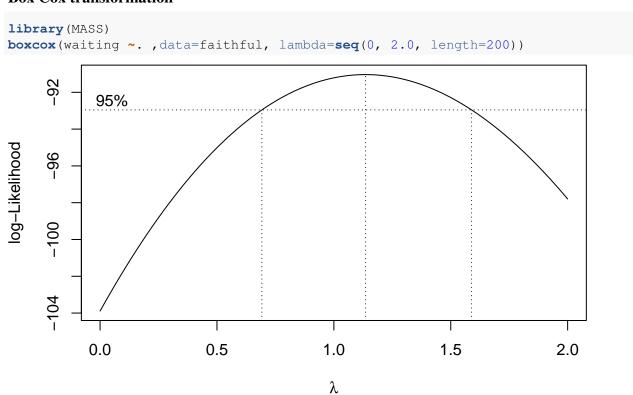
```
(Intercept)
                            eruptions
## (Intercept)
                 1.3337328 -0.34553365
## eruptions
                -0.3455336 0.09906971
pred.per_fat = predict(mod1)
res.per_fat = residuals(mod1)
summary(res.per_fat)
       Min.
            1st Qu.
                       Median
                                   Mean
                                         3rd Qu.
                                                      Max.
## -12.0796
            -4.4831
                       0.2122
                                 0.0000
                                          3.9246
                                                  15.9719
Plots of regression
par (mfrow=c(1,1))
plot (waiting, eruptions)
```



```
par(mfrow=c(1,2))
plot(mod1, which=c(1,2))
```



# **Box Cox transformation**



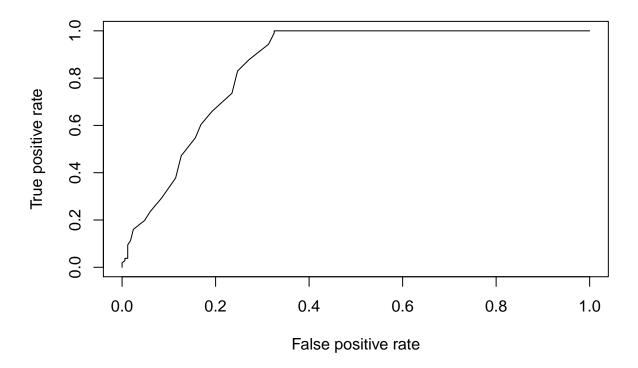
#### **ROC Curves**

#### **ROC** Curves for eruption > 3 minutes

```
library(ROCR)
cut_point=(eruptions > 3)
pred = prediction(waiting, cut_point)
perf=performance(pred, "tpr", "fpr")
plot (perf)
      0.8
True positive rate
      9.0
      0.4
      0.2
      0.0
                           0.2
            0.0
                                         0.4
                                                        0.6
                                                                      8.0
                                                                                     1.0
                                        False positive rate
```

#### **ROC** Curves for eruption > 4.2 minutes

```
library(ROCR)
cut_point=(eruptions > 4.2)
pred = prediction(waiting, cut_point)
perf=performance(pred, "tpr", "fpr")
plot(perf)
```



#### **SAS**

#### **SAS Code**

```
options center nodate pagesize=100 ls=70;
libname LDATA '/home/jacktubbs/my_shared_file_links/jacktubbs/LaTeX/';
/* Simplified LaTeX output that uses plain LaTeX tables */
ods tagsets.simplelatex
file="/home/jacktubbs/my_shared_file_links/jacktubbs/LaTeX/sheather_faithful.tex"
stylesheet="/home/jacktubbs/my_shared_file_links/jacktubbs/LaTeX/sas.sty"
(url="sas");
title 'Sheather KDE Simulated Data';
ods graphics on;
data bimodal; set ldata.bimodal;
run;
proc kde data=bimodal;
  univar x(bwm=.2) x(bwm=0.4) x(bwm=.6) x(bwm=0.8);
run;
title 'Old Faithful Data';
data faithful; set ldata.faithful;
run;
proc sgplot data=faithful;
    scatter y=wait x=duration;
    reg y=wait x=duration/ clm;
```

```
loess y=wait x=duration;
run;

proc sgplot data=faithful;
histogram wait;
density wait;
run;

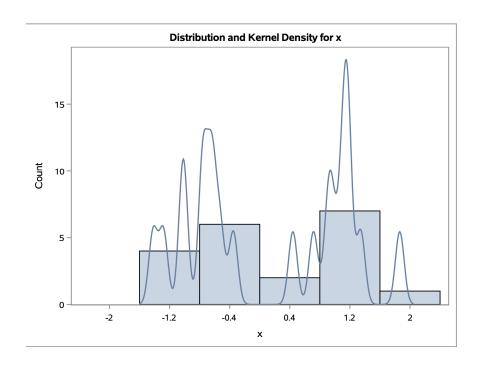
proc kde data=faithful;
univar wait;
run;

proc kde data=faithful;
bivar wait duration ;
run;
quit;
```

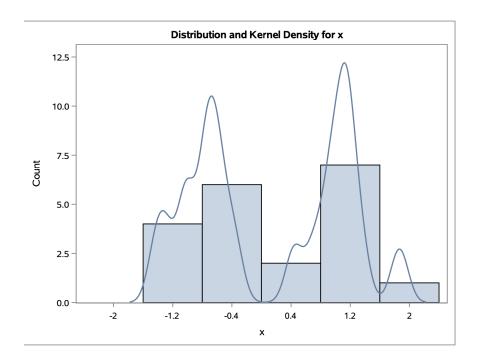
# Output

# Sheather KDE Simulated Data The KDE Procedure

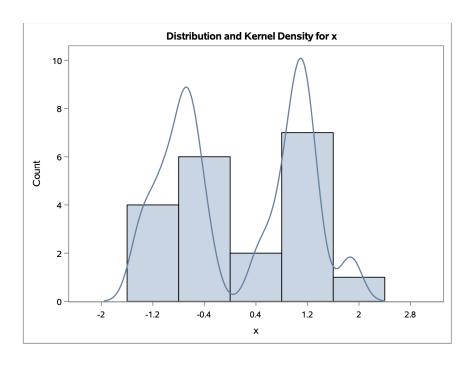
Controls	
	X
Grid Points	401
Lower Grid Limit	-1.601
Upper Grid Limit	2.0459
Bandwidth Multiplier	0.2



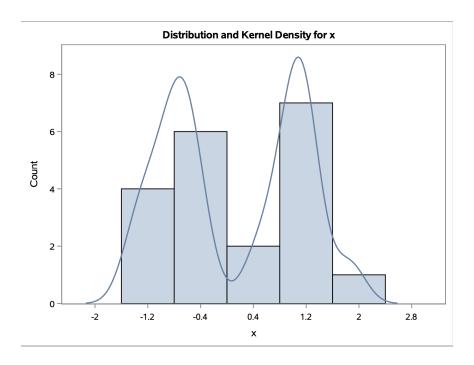
Controls	
	X
Grid Points	401
Lower Grid Limit	-1.601
Upper Grid Limit	2.0459
Bandwidth Multiplier	0.4



Controls	
	X
Grid Points	401
Lower Grid Limit	-1.601
Upper Grid Limit	2.0459
Bandwidth Multiplier	0.6



Controls	
	X
Grid Points	401
Lower Grid Limit	-1.601
Upper Grid Limit	2.0459
Bandwidth Multiplier	0.8



# Old Faithful Data The KDE Procedure

