

is the condition for a Markov chain

$$1. P(X_n = y | X_0 = x_0, \dots, X_{n-1} = x_{n-1}) = P(X_n = y | X_{n-1} = x_{n-1})$$

LHS = RHS ✓ +5

$$X_n = \sum_{i=0}^n \epsilon_i = \sum_{i=0}^{n-1} \epsilon_i + \epsilon_n = X_{n-1} + \epsilon_n$$

$$\text{where } \epsilon_i = \begin{cases} 1 & \text{w.p. } \frac{1}{2} \\ -1 & \text{w.p. } \frac{1}{2} \end{cases}$$

Very well done
submission.

$$\text{LHS} = P(\epsilon_n = y - x_{n-1} | X_0 = x_0, \dots, X_{n-1} = x_{n-1})$$

ϵ_n is independent of X_i where $i \in \{0, \dots, n-1\}$ and

the history, $X_0 = x_0, \dots, X_{n-1} = x_{n-1}$, that is being conditioned has a positive probability.

$$\Rightarrow \text{LHS} = P(\epsilon_n = y - x_{n-1}) = \begin{cases} \frac{1}{2} & |y - x_{n-1}| = 1 \\ 0 & \text{"} \neq 1 \end{cases}$$

Then, apply the same transformation of the new random variable, epsilon ϵ .

$$\text{RHS} = P(\epsilon_n = y - x_{n-1} | X_{n-1} = x_{n-1})$$

due to the same rule as independence and positive probability of the conditioned history

$$\Rightarrow \text{RHS} = P(\epsilon_n = y - x_{n-1}) = \text{LHS} \quad \checkmark 28$$

□

2a) let states $x, y \in \{0, \dots, N\}$

$$\begin{aligned} \text{let } x=y &\Rightarrow P(X_n=x | X_0=y) = P(X_n=x | X_0=x) \\ &= P(X_n=x | X_{n-1}=x+1) \cdot P(X_{n-1}=x+1 | X_{n-2}=x) \\ \text{let } n=2 &= P(X_2=x | X_1=x+1) \cdot P(X_1=x+1 | X_0=x) \\ &> 0 \end{aligned}$$

For the product of positive probabilities is also positive.
This is the case for when $x < y$ and $x > y$. ✓

let $y > x$ $x+n=y$ where $n \in \{1, \dots, N\}$

$$\begin{aligned} P(X_n=x | X_0=y) &= P(X_n=x | X_{n-1}=x+1) \cdot P(X_{n-1}=x+1 | X_{n-2}=x+2) \\ &\quad \dots P(X_1=x+n-1 | X_0=x+n) \\ &= \prod_{k=0}^{n-1} P(X_{n-k}=x+k | X_{n-k-1}=x+1+k) \\ &> 0 \end{aligned}$$

let $x > y$ $x-n=y$ where $n \in \{1, \dots, N\}$

$$\begin{aligned} P(X_n=x | X_0=y) &= P(X_n=x | X_{n-1}=x-1) \cdot P(X_{n-1}=x-1 | X_{n-2}=x-2) \\ &\quad \dots P(X_1=x-n+1 | X_0=x-n) \\ &= \prod_{k=0}^{n-1} P(X_{n-k}=x-k | X_{n-k-1}=x-k-1) \\ &> 0 \end{aligned}$$

□

✓ +10

2b) An aperiodic chain is if all states are aperiodic.

State $x \in S$ is aperiodic if the highest common factor of $\tau(x)$ is equal to 1

~~A counter-example will show that the Markov chain is not aperiodic.~~

~~Let $x \neq 1$~~ \Rightarrow where $\tau(x) = \{n \in \mathbb{N} : p_n(x, x) > 0\}$

$$P(X_n = x | X_0 = x) = P(X_n = x | X_{n-1} = x+1) \cdot P(X_{n-1} = x+1 | X_{n-2} = x)$$
$$= \frac{x+1}{N} \cdot \frac{N-x}{N}$$

$$\begin{aligned} &> 0 \\ \text{or} \quad &= P(X_n = x | X_{n-1} = x-1) \cdot P(X_{n-1} = x-1 | X_{n-2} = x) \\ &= \frac{N-x+1}{N} \cdot \frac{x}{N} \end{aligned}$$

> 0

requires an even, $2n$, number of steps \Rightarrow period is equal to 2.

$$P(X_n = x | X_{n-1} = x) = 0 \quad \forall n \in \mathbb{N}_+$$

✓ +10

□

2c) stationary distribution is a binomial distribution $\text{Bin}(N, \frac{1}{2})$.

$$\Rightarrow \bar{\pi} = P(X_n = N) \quad \text{for } x \in \{0, 1, \dots, N\}$$

$$= \binom{N}{x} \left(\frac{1}{2}\right)^N$$

$$\Rightarrow \bar{\pi} = \bar{\pi} P$$

where $\bar{\pi} = (\pi_0, \pi_1, \dots, \pi_N)$ describes us reaching state $x \in \{0, \dots, N\}$ from any other state.

$$\bar{\pi} = \begin{bmatrix} \pi_0 \\ \pi_1 \\ \pi_2 \\ \vdots \\ \pi_{N-1} \\ \pi_N \end{bmatrix} \begin{bmatrix} 0 & 1 & 0 & \dots & 0 \\ \frac{1}{N} & 0 & \frac{N-1}{N} & \dots & 0 \\ 0 & \frac{2}{N} & 0 & \dots & \vdots \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \vdots & \vdots & \frac{2}{N} & 0 \\ \vdots & \vdots & \vdots & \vdots & 0 & \frac{1}{N} \\ 0 & \dots & \dots & \dots & 1 & 0 \end{bmatrix}$$

$$\begin{bmatrix} \pi_0 \\ \pi_1 \\ \pi_2 \\ \vdots \\ \pi_{N-1} \\ \pi_N \end{bmatrix} = \begin{bmatrix} \pi_1 \cdot \frac{1}{N} \\ \pi_0 + \pi_2 \cdot \frac{2}{N} \\ \pi_1 \cdot \frac{N-1}{N} + \pi_3 \cdot \frac{3}{N} \\ \vdots \\ \pi_{N-2} \cdot \frac{2}{N} + \pi_N \\ \pi_{N-1} \cdot \frac{1}{N} \end{bmatrix}$$

$$\Rightarrow \pi_0 = \pi_1 \cdot \frac{1}{N} \Rightarrow \pi_1 = N \pi_0$$

~~π_1~~

$$\pi_1 = \pi_0 + \pi_2 \cdot \frac{2}{N} \Rightarrow N \pi_0 = \pi_0 + \pi_2 \cdot \frac{2}{N}$$

$$\Rightarrow \pi_0 (N-1) = \pi_2 \cdot \frac{2}{N} \Rightarrow \pi_2 = \frac{N(N-1)}{2} \pi_0$$

each π follows a pattern

$$\pi_x = \frac{N!}{N! (N-x)! x!} \pi_0 \Rightarrow \pi_x = \binom{N}{x} \pi_0$$

lastly find π_0

$$\sum_{x=0}^N \pi_x = 1 \quad : \text{ Law of total probability}$$

$$= \sum_{x=0}^N \binom{N}{x} \pi_0 = 1 \quad : \text{ Substitution from above}$$

$\sum_{x=0}^N {}^N C_x$ follows the pattern of Pascal's Triangle and so does $(1+y)^N$

$$\begin{aligned} \sum_{x=0}^N y^x {}^N C_x &= {}^N C_0 y^0 + {}^N C_1 y^1 + {}^N C_2 y^2 + \dots + {}^N C_N y^N \\ &= (1+y)^N \end{aligned}$$

$$\text{if } y=1 \Rightarrow \sum_{x=0}^N 1^x {}^N C_x = \sum_{x=0}^N {}^N C_x = (1+1)^N = (2)^N$$

$$\Rightarrow \text{substitute } \sum_{x=0}^N \binom{N}{x} = 2^N \text{ into } \sum_{x=0}^N \binom{N}{x} \pi_0 = 1$$

$$2^N \pi_0 = 1 \Rightarrow \pi_0 = \frac{1}{2^N} = \left(\frac{1}{2}\right)^N$$

$$\Rightarrow \pi_x = \binom{N}{x} \left(\frac{1}{2}\right)^N \text{ where } x \in \{0, 1, \dots, N\}$$

□

✓ +30

3.)

Directed graph:

