Random Matrix Theory

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Preliminaries

Let ξ_{ij} , η_{ij} be normal random variables (i.e. Gaussian, mean 0, variance 1).

e.g.
$$\mathbb{P}(\xi_{11} < s) = \int_{-\infty}^{s} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$
.

$$\int_{-\infty}^{\infty} x^2 \cdot \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$
 is the variance.

$$\frac{1}{\sqrt{2\pi}}e^{-x^2/2}$$
 is the Probability Density Function (PDF).

 $\frac{1}{\sqrt{2\pi}}e^{-x^2/2}\,dx$ is the probability measure on our probability space (i.e. totally finite measure space). We build matrices

$$\begin{bmatrix} \xi_{11} & \frac{\xi_{12} + i\eta_{12}}{\sqrt{2}} & \frac{\xi_{13} + i\eta_{13}}{\sqrt{2}} & \cdots \\ \frac{\xi_{21} + i\eta_{21}}{\sqrt{2}} & \xi_{22} & \frac{\xi_{22} + i\eta_{22}}{\sqrt{2}} \\ \frac{\xi_{31} + i\eta_{31}}{\sqrt{2}} & \frac{\xi_{32} + i\eta_{32}}{\sqrt{2}} & \xi_{33} \\ \vdots & & \ddots \end{bmatrix}$$

Computing Random Matrices in Matlab

Gassuain, real valued 1x1 matrix.

randn

Gaussian, real valued 2x2 matrix.

randn(2)

Gaussian, complex valued 2x2 matrix.

Gaussian, complex valued, self-adjoint 2x2 matrix.

Note that appending 'to a matrix takes the conjugate transpose, and matlab reserves i for the imaginary unit.

Producing eigenvalues.

Running tests to see how many hits we get within the interval [0,2].

```
edges=[0,2];
H=zeros(1,length(edges)-1);
trials=10;
for j=1:trials
m = randn(2)+i*randn(2);
l=(m+m')/2;
ev=eig(1);
H=H+histcount(ev,edges)
end
```

Homework

Is the PDF of $\frac{a+b}{2}$ the same as $\frac{\xi_{12}}{\sqrt{2}}$ for normal RVs a,b,ξ_{12} ? i.e. $\mathbb{P}\left(\frac{a+b}{2} < s\right) \stackrel{?}{=} \mathbb{P}\left(\frac{\xi_{12}}{\sqrt{2}} < s\right)$

2x2 Random Matrix

Our matrix L corresponds to eigenvalues λ_1, λ_2 which are random variables determined by $\{\xi_{ij}, \eta_{ij}\}$. Then the number of evaluations in the interval B is given by $\sum_{j=1}^{2} \chi_B(\lambda_j)$. We may take the average by

$$\int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} \sum_{i=1}^{2} \chi_{B}(\lambda_{j}) \frac{1}{\sqrt{2\pi}} e^{-\xi_{11}^{2}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\xi_{22}^{2}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\xi_{12}^{2}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\eta_{12}^{2}} d\xi_{11} d\xi_{22} d\xi_{12} d\eta_{12}.$$

Expected Evaluations

We have that the expectation of the number of evaluations in the interval (a,b) is given by $\int_a^b G(s) \, ds$ where

$$G(s) = e^{-\frac{s^2}{2}} \sum_{\ell=0}^{2} P_{\ell}(s)^2$$

and $P_{\ell}(s)$ is the Hermite polynomial of degree d.

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Differntiability

```
delta = 0.05;
edges=-6:delta:6;
dimensions = 3;
trials = 1000000;

H=zeros(dimensions,trials);

for j=1:trials
m=randn(dimensions)+1i*randn(dimensions);
L=(m+m')/2;
ev=eig(L);
H(:,j) = ev;
end

G = histcounts(H,edges);
plot(edges(1:end-1),G/(trials*delta),'*')
```

Observe that each * in the graph corresponds to the average number of eigenvalues in the interaval (a, b). Therefore, they correspond to $\int_a^b C(\lambda) d\lambda$. We may consider the limit of the expectation of hits in each interval

$$\lim_{\Delta \to 0} \frac{\mathbb{E}(\#(a, a + \Delta))}{\Delta}.$$

```
delta = 0.01;
edges=-6:delta:6;
dimensions = 3;
trials = 1000000;

H=zeros(dimensions,trials);

for j=1:trials
m=randn(dimensions)+1i*randn(dimensions);
L=(m+m')/2;
ev=eig(L);
H(:,j) = ev;
end

G = histcounts(H,edges);
plot(edges(1:end-1),G/(trials*delta),'*')
```

As dimension grows large, we observe that the plot tends to a semi-circle with endpoints about $\pm 2\sqrt{\text{dimension}}$. We therefore want a rescaling by \sqrt{N} where $\dim = N$. Then if $G(\alpha) = \frac{d}{d\alpha}\mathbb{E}(\# \text{ of evals in } (a, \alpha))$, we want

$$\int_{-\infty}^{\infty} G(\alpha) d\alpha = N.$$

Guess: $G(\alpha) \approx cN^{1/2} \cdot \sqrt{A^2 - \alpha^2/N} \cdot \chi_{(-A\sqrt{N},A\sqrt{N})}(\alpha)$. We compute

$$\int_{-A/N}^{A\sqrt{N}} c N^{1/2} \sqrt{A^2 - \alpha^2/N} \, d\alpha \stackrel{\alpha = \sqrt{N}t}{=} c N \int_{-A}^{A} \sqrt{A^2 - t^2} \, dt = \frac{c\pi N A^2}{2}.$$

Choosing A=2 and c such that $\frac{\pi A^2 c}{2}=1$, we get

$$\int_{-\infty}^{\infty} G(\alpha) d\alpha \approx \frac{N^{1/2}}{2\pi} \int_{-\infty}^{\infty} \sqrt{4 - \alpha^2/N} d\alpha = N.$$

Number of Eigenvalues in an Interval

Let B be a subset of \mathbb{R} (typically an interval). Write $n(B) = \#\{\text{evaluations in } B\}$, a random variable. Recall that variance is given by the expectation of the square minus the square of the expectation. That is

$$\operatorname{var}(n(B)) = \mathbb{E}(n(B)^{2}) - (\mathbb{E}(n(B))^{2}.$$

Our ultimate goal is to understand PDF and $\mathbb{P}(n(B)) = \ell$) as (the dimension) $N \to \infty$.

Smallest Scale of Interest

Suppose B = (0, S) and N is large (i.e. $N \to \infty$). How large should we choose s such that $\mathbb{E}(n(B)) = 1$? We compute

$$\int_0^S cN^{1/2} \sqrt{4 - \alpha^2/N} \ d\alpha \stackrel{\alpha = \sqrt{N}t}{=} \int_0^{\frac{S}{\sqrt{N}}} cN \sqrt{4 - t^2} \ dt \approx cN \cdot 2 \frac{S}{\sqrt{N}} = 2cS\sqrt{N}.$$

Sets of size $N^{-1/2}$, the smallest interesting scale, are called the "microscopic scaling regime".

Homework: Largest Scale of Interest

How large should B be to see a fraction of the eigenvalues (on average)? That is, how should we scale a and b such that $\mathbb{E}(n((a,b))) = r \cdot N$ for 0 < r < 1?

Level Repulsion

```
m=randn(2)+sqrt(-1)*randn(2);
L=(m+m')/2;
ev=eig(L);
subplot(2,1,2),plot(real(ev),imag(ev))
xlim([edges(1),edges(end)])
```

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Macroscopic Scaling Regime for Random Matrices

Suppose $a = \alpha \sqrt{N}$ and $b = \beta \sqrt{N}$ such that $\alpha < \beta$, $-2 < \alpha$ and $\beta < 2$. Then

$$\lim_{n\to\infty} \frac{\mathbb{E}(\# \text{ of evaluations in } (\alpha\sqrt{N}, \beta\sqrt{N}))}{N} = \kappa > 0.$$

Recall that we defined $G(b) = \frac{d}{db}\mathbb{E}(\# \text{ of evaluations in } (a,b))$ and

$$G(b)\approx cN^{1/2}\sqrt{A^2-x^2/N}\chi_{[-A\sqrt{N},A\sqrt{n}]}(x).$$

We want that $\int_a^b G(x) dx = \kappa N$.

Spacings

Suppose we have eigenvalues $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_N = \lambda_{\max}$. We can take the spacing $s_i = \lambda_{i+1} - \lambda_i$.

```
m=randn(2)+sqrt(-1)*randn(2);
L=(m+m')/2;
ev=sort(eig(L));
spacing=diff(ev)
0.4839
```

Summary So Far

Given ξ_{ij} and η_{ij} iid RVs with distribution $\frac{1}{\sqrt{2\pi}}e^{-x^2/2}$, we have explored

- The behavior of average $n_N(B)$.
- · Microscopic, macroscopic (and mesoscopic) scaling.
- That $\lambda_{\rm max} \sim 2\sqrt{N}$ Tracy-Widom distribution.
- Eigenvalue repulsion.

Induced Distribution

Let M be our matrix built using random variables. Then $M = F\Lambda F^T$ where

$$\Lambda = \begin{pmatrix} \lambda_1 & 0 & \cdots \\ 0 & \lambda_2 & \\ \vdots & & \ddots \end{pmatrix}, \quad F = \begin{pmatrix} | & | & & | \\ f_{\lambda_1} & f_{\lambda_2} & \cdots & f_{\lambda_N} \\ | & | & & | \end{pmatrix},$$

and $Mf_{\lambda_i} = \lambda_j f_{\lambda_i}$. What we are interested in is the induced joint PDF on $\{\lambda_1, \dots, \lambda_N\}$. We may write explicitly

$$\frac{1}{Z^n}e^{-\frac{1}{2}\sum_{j=1}^N\lambda_j^2}\prod_{1\leq j< k\leq N}(\lambda_k-\lambda_j)^2.$$

Example

Let N = 2 and, suppressing the constant term, write

$$\rho = e^{-\frac{1}{2}(x^2 + y^2)}(x - y)^2.$$

Taking partial derivatives, we have that

$$\rho_x = e^{-\frac{1}{2}(x^2 + y^2)} (x - y)^2 (-x + \frac{2}{x - y})$$

$$\rho_y = e^{-\frac{1}{2}(x^2 + y^2)} (x - y)^2 (-x + \frac{2}{y - x})$$

which implies maxima at $x = \pm 1$ and y = -x.

Example

If N=3,

$$\rho = e^{-\frac{1}{2}(x^2 + y^2 + z^2)}(x - y)^2(x - z)^2(y - z)^2.$$

We may visualize the maxima here by level surfaces (homework).

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Recall: Spectral Theorem

Let $M = F\Lambda F^{\dagger}$ where $F^{\dagger}F = I = FF^{\dagger}$

$$\Lambda = \begin{pmatrix} \lambda_N & 0 & \cdots \\ 0 & \lambda_{N-1} \\ \vdots & & \ddots \\ & & \lambda_1 \end{pmatrix}, \quad F = \begin{pmatrix} | & | & & | \\ f_{\lambda_1} & f_{\lambda_2} & \cdots & f_{\lambda_N} \\ | & | & & | \end{pmatrix},$$

for $\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_N$.

Deriving the Joint PDF

Let n = 2. If

$$F = \begin{pmatrix} | & | \\ V & W \\ | & | \end{pmatrix},$$

then the expectation of eigenvalues may be computed by

$$\begin{split} \mathbb{E}(\mathcal{G}(M)) &= \frac{1}{Z_2^4} \int \cdots \int \mathcal{G}(M(\xi_{11}, \xi_{12}, \xi_{22}, \eta_{12})) x \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\xi_{11}^2 + \xi_{12}^2 + \xi_{22}^2 + \eta_{12}^2)} d\eta_{12} d\xi_{22} d\xi_{12} d\xi_{11} \\ &= \int \mathcal{G}(M(\lambda_1, \lambda_2, V_1, \phi)) x \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\xi_{11}^2 + \xi_{12}^2 + \xi_{22}^2 + \eta_{12}^2)} d\eta_{12} d\xi_{22} d\xi_{12} d\xi_{11}. \end{split}$$

So we need the Jacobian, and therefore a reparameterization using spectral theorem. We want a collection of independent variables which will produce all 2×2 Hermitian matrices. Consider $Mv=\lambda_2 v$ and $||v||^2=|v_1|^2+|v_2|^2=1$, then multiply by $e^{i\eta}$ such that $v_1\in\mathbb{R}_+$. Then $v_2=\sqrt{1-v_1^2}e^{i\theta}$. That is, $0\le v_1\le 1$ and $v_2=\sqrt{1-v_1^2}(\cos\theta+i\sin\theta)$. We want that $|w_1|^2+|w_2|^2=1$ and know that $w\perp v$, so $v_1w_1+\overline{v}_2w_2=0$. As before, we can choose w such that $w_2\in\mathbb{R}_+$.

We want that $|w_1|^2 + |w_2|^2 = 1$ and know that $w \perp v$, so $v_1 w_1 + \overline{v}_2 w_2 = 0$. As before, we can choose w such that $w_2 \in \mathbb{R}_+$. This implies that w_1 and \overline{v}_2 have the same argument, and $w_1 = -|w_1|e^{-i\theta}$. Therefore $e^{-i\theta}(-v_1|w_1| + |v_2|w_2) = 0$, and $v_1|w_1| - |v_2|w_2 = 0$. It follows that

$$v_1^2(1-w_2^2) = w_2^2(1-v_1^2) \iff v_1 = w_2.$$

Therefore, the entire system may be parameterized by v_1 and θ . We write

$$F = \begin{pmatrix} v_1 & -\sqrt{1 - v_1^2} e^{-i\theta} \\ \sqrt{1 - v_1^2} e^{i\theta} & v_1 \end{pmatrix}$$

and

$$M = F\Lambda F^{\dagger} = \begin{pmatrix} v_1 & -\sqrt{1-v_1^2}e^{-i\theta} \\ \sqrt{1-v_1^2}e^{i\theta} & v_1 \end{pmatrix} \begin{pmatrix} \lambda_2 & 0 \\ 0 & \lambda_1 \end{pmatrix} \begin{pmatrix} v_1 & \sqrt{1-v_1^2}e^{-i\theta} \\ -\sqrt{1-v_1^2}e^{i\theta} & v_1 \end{pmatrix}.$$

Therefore

$$M = \begin{pmatrix} \lambda_2 v_1^2 + \lambda_1 (1 - v_1^2) & v_1 \sqrt{1 - v_1^2} e^{-i\theta} (\lambda_2 - \lambda_1) \\ v_1 \sqrt{1 - v_1^2} e^{-i\theta} (\lambda_2 - \lambda_1) & \lambda_2 (1 - v_1^2) + \lambda_1 v_1^2 \end{pmatrix}.$$

Recall, we want $\mathcal{G}(M(\xi)) \rightsquigarrow \mathcal{G}(M(\lambda_2, \lambda_1, \nu_1, \theta))$ and the Jacobian of $M = M(\lambda_2, \lambda_1, \nu_1, \theta)$. After computation, write

$$|\det J = (\lambda_2 - \lambda_1)^2 \det J' = (\lambda_2 - \lambda_1)^2 Q(\nu_1, \theta).$$

We integrate

$$\int \cdots \int \mathcal{G}(M(\xi,\eta_{12})) e^{-\frac{1}{2}(\xi_{11}^2 + \xi_{12}^2 + \xi_{22}^2 + \eta_{12}^2)} \frac{1}{(2\pi)^4} d\xi_{11} d\xi_{12} d\xi_{22} d\eta_{12}$$

which we may think of uas a function of λ_1 and λ_2 alone. So

$$\frac{1}{(2\pi)^2} \int \cdots \int \mathcal{G}(\lambda_1, \lambda_2) e^{-\frac{1}{2} \left[M_{11}^2 + M_{22}^2 + 2 \cdot \text{Re}(M_{12})^2 + 2 \cdot \text{Im}(M_{12})^2\right]} d\xi_{11} d\xi_{12} d\xi_{22} d\eta_{12}$$

where we observe that $M_{11}^2 + M_{22}^2 + 2 \cdot \text{Re}(M_{12})^2 + 2 \cdot \text{Im}(M_{12})^2 = \text{Tr}(M^2)$. It follows that we have

$$\begin{split} \frac{1}{(2\pi)^2} \int \cdots \int \mathcal{G}(\lambda_1, \lambda_2) e^{-\frac{1}{2}(\lambda_1^2 + \lambda_2^2)} \, d\xi_{11} d\xi_{12} d\xi_{22} d\eta_{12} &= \frac{1}{(2\pi)^2} \int_0^{2\pi} \int_0^1 \int \int_{-\infty < \lambda_1 \le \lambda_2 < \infty} \mathcal{G}(\lambda_1, \lambda_2) e^{-\frac{1}{2}(\lambda_1^2 + \lambda_2^2)} (\lambda_2 - \lambda_1)^2 Q(v, \theta) \, d\xi_1 \\ &= \int \int_{-\infty < \lambda_1 \le \lambda_2 < \infty} \mathcal{G}(\lambda_1, \lambda_2) e^{-\frac{1}{2}(\lambda_1^2 + \lambda_2^2)} (\lambda_2 - \lambda_1)^2 \int_0^{2\pi} \int_0^1 \frac{Q(v, \theta)}{(2\pi)^2} \, dv_1 d\theta dv_1 \\ &= c \int \int \mathcal{G}(\lambda_1, \lambda_2) e^{-\frac{1}{2}(\lambda_1^2 + \lambda_2^2)} (\lambda_2 - \lambda_1)^2 \, d\lambda_1 d\lambda_2 \end{split}$$

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Recall: Joint PDF on Evaluation of Eigenvalues

 $\lambda_1 \le \lambda_2 \le \cdots \le \lambda_N$ and PDF $\frac{1}{Z_N} e^{-\frac{1}{2} \sum \lambda_i^2} \prod (\lambda_k - \lambda_j)^2$. This is the Gaussian Unitary Ensemble.

Hermite Polynomials

Write $p_j = \kappa_j^{(j)} x^j + \kappa_{j-1}^{(j)} x^{j-1} + \dots + \kappa_0^{(j)}$ where the superscript is usually supressed. Then

$$\int_{-\infty}^{\infty} p_j(x) p_k(x) e^{-\frac{1}{2}x^2} dx = \delta_{jk}.$$

Observe that $\{e^{-\frac{1}{4}x^2}p_j(x)\}_{j=0}^{\infty}$ forms a basis for $L^2(\mathbb{R})$. For $f \in L^2$, write the truncation $P^{(N)}(f) = \sum_{\ell=0}^{N-1} \left(\int_{\mathbb{R}} f(y)p_\ell(y)e^{-\frac{1}{4}y^2}\right)e^{-\frac{1}{4}y^2}$. Then

$$P^{(N)} = \int_{\mathbb{R}} \left(e^{-\frac{1}{4}(x^2 + y^2)} \sum_{\ell=0}^{N-1} p_{\ell}(x) p_{\ell}(y) \right) f(y) \, dy$$

and we write $K_N(x,y) = e^{-\frac{1}{4}(x^2+y^2)} \sum_{\ell=0}^{N-1} p_{\ell}(x) p_{\ell}(y)$ and $K_N(f) = \int_{\mathbb{R}} K_n(x,y) f(y) \, dy$. We have that

$$\frac{1}{Z_N}e^{-\frac{1}{2}\sum \lambda_i^2} \prod (\lambda_k - \lambda_j)^2 = \det \begin{bmatrix} K_N(\lambda_1, \lambda_1) & \cdots & K_N(\lambda_1, \lambda_N) \\ \vdots & \ddots & \vdots \\ K_N(\lambda_N, \lambda_1) & \cdots & K_N(\lambda_N, \lambda_N) \end{bmatrix}.$$

For N = 2, we see

$$\frac{1}{Z_N}e^{-\frac{1}{2}\sum \lambda_i^2}\prod (\lambda_k-\lambda_j)^2=(K_N(\lambda_1,\lambda_1)K_N(\lambda_2,\lambda_2)-K_N(\lambda_1,\lambda_2)^2).$$

Example Computation

Let I be an interval and consider $\mathbb{E}(\#)$ of evaluations in I). Then

$$\begin{split} &\int_{-\infty<\lambda_1\leq\lambda_2<\infty} \left(\sum_{j=1}^2 \chi_I(\lambda_j)\right) (K_2(\lambda_1,\lambda_1)K_2(\lambda_2,\lambda_2) - K_2(\lambda_1,\lambda_2)^2) \, d\lambda_2\lambda_2 \\ &= \frac{1}{2!} \int \int_{\mathbb{R}^2} (\chi_I(\lambda_1) + \chi(\lambda_2)) (K_2(\lambda_1,\lambda_1)K_2(\lambda_2,\lambda_2) - K_2(\lambda_1,\lambda_2)^2) \, d\lambda_1 d\lambda_2 \\ &= \frac{1}{2!} \int_I \int_{-\infty}^{\infty} (K_2(\lambda_1,\lambda_1)K_2(\lambda_2,\lambda_2) - K_2(\lambda_1,\lambda_2)^2) \, d\lambda_1 d\lambda_2 + \frac{1}{2!} \int_I \int_{-\infty}^{\infty} (K_2(\lambda_1,\lambda_1)K_2(\lambda_2,\lambda_2) - K_2(\lambda_1,\lambda_2)^2) \, d\lambda_2 d\lambda_1. \end{split}$$

Observe that $\int_{-\infty}^{\infty} K_2(\lambda_2, \lambda_2) d\lambda_2 = \int_{-\infty}^{\infty} e^{-\frac{1}{2}\lambda_2^2} (p_0(\lambda_2)^2 + p_1(\lambda_2)^2) d\lambda_2 = 2$. We also compute that

$$\int_{-\infty}^{\infty} K_{2}(\lambda_{1}, \lambda_{2}) K_{2}(\lambda_{2}, \lambda_{1}) d\lambda_{2} = \int_{\mathbb{R}} e^{-\frac{1}{2}(\lambda_{1}^{2} + \lambda_{2}^{2})} \left(\sum_{\ell=0}^{1} p_{\ell}(\lambda_{1}) p_{\ell}(\lambda_{2}) \right) \left(\sum_{\ell'=0}^{1} p_{\ell'}(\lambda_{2}) p_{\ell'}(\lambda_{1}) \right) d\lambda_{2}$$

$$= \sum_{\ell, \ell'=0}^{1} e^{-\frac{1}{2}\lambda_{1}^{2}} p_{\ell}(\lambda_{1}) p_{\ell'}(\lambda_{1}) \int_{\infty}^{\infty} e^{-\frac{1}{2}\lambda_{2}^{2}} p_{\ell}(\lambda_{2}) p_{\ell'}(\lambda_{2}) d\lambda_{2}$$

$$= K_{2}(\lambda_{1}, \lambda_{1}).$$

Returning to the first calculation,

$$\begin{split} &\frac{1}{2!} \int_{I} \int_{-\infty}^{\infty} \left(K_{2}(\lambda_{1}, \lambda_{1}) K_{2}(\lambda_{2}, \lambda_{2}) - K_{2}(\lambda_{1}, \lambda_{2})^{2} \right) d\lambda_{1} d\lambda_{2} + \frac{1}{2!} \int_{I} \int_{-\infty}^{\infty} \left(K_{2}(\lambda_{1}, \lambda_{1}) K_{2}(\lambda_{2}, \lambda_{2}) - K_{2}(\lambda_{1}, \lambda_{2})^{2} \right) d\lambda_{2} d\lambda_{1} \\ &= \frac{1}{2!} \left[\int_{I} (2 - 1) K_{2}(\lambda_{1}, \lambda_{1}) d\lambda_{1} + \int_{I} (2 - 1) K_{2}(\lambda_{2}, \lambda_{2}) d\lambda_{2} \right] \\ &= \int_{I} K_{2}(\lambda_{1}, \lambda_{1}) d\lambda_{1} \end{split}$$

which is the density function for the average number of evaluations in *I*. So $K_2(\lambda, \lambda) = \frac{e^{-\frac{1}{2}\lambda^2}}{\sqrt{2\pi}}(1 + \lambda^2)$.

Question:

What is the probability of having zero evaluations in an interval I? We have an indicator function $(1 - \chi_I(\lambda_1))(1 - \chi_I(\lambda_2))$, so

$$\begin{split} P(\text{no evaluations in } I) &= \frac{1}{2} \int_{\mathbb{R}^2} (1 - \chi_I(\lambda_1))(1 - \chi_I(\lambda_2)) \big[K_2(\lambda_1, \lambda_1) K_2(\lambda_2, \lambda_2) - K_2(\lambda_1, \lambda_2)^2 \big] \, d\lambda_1 d\lambda_2 \\ &= \frac{1}{2} \int_{\mathbb{R}^2} (1 - (\chi_I(\lambda_1) + \chi_I(\lambda_2)) + \chi_I(\lambda_1) \chi_I(\lambda_2)) K_2(\lambda_1, \lambda_1) K_2(\lambda_2, \lambda_2) - K_2(\lambda_1, \lambda_2)^2 \big] \, d\lambda_1 d\lambda_2 \\ &= \frac{1}{2} \big[\int_{\mathbb{R}^2} K_2(\lambda_1, \lambda_1) K_2(\lambda_2, \lambda_2) - K_2(\lambda_1, \lambda_2)^2 \, d\lambda_1 d\lambda_2 \\ &- 2 \int_I \int_{\mathbb{R}} K_2(\lambda_1, \lambda_1) K_2(\lambda_2, \lambda_2) - K_2(\lambda_1, \lambda_2)^2 \, d\lambda_2 d\lambda_1 \\ &+ \int_I \int_I K_2(\lambda_1, \lambda_1) K_2(\lambda_2, \lambda_2) - K_2(\lambda_1, \lambda_2)^2 \big] \\ &= \frac{1}{2} \big[4 - 2 - 2 \int_I K_2(\lambda_1, \lambda_1) \, d\lambda_1 + \int_I \int_I K_2(\lambda_1, \lambda_1) K_2(\lambda_2, \lambda_2) - K_2(\lambda_1, \lambda_2)^2 \, d\lambda_1 \lambda_2 \big] \\ &= 1 - \int_I K_2(\lambda_1, \lambda_1) \, d\lambda_1 + \int_I \int_I \det(\quad)_{2 \times 2} \, d^2 \lambda \\ &= \det(1 - \mathcal{K}_2^{(I)}) \end{split}$$

If $I = (0, \infty)$, then the probability is $\frac{\pi - 2}{4\pi}$.

Fredholm Determinant

Write $H_N(I,t) = \det(1-t\mathcal{K}_N^{(I)})$ where $K_N^{(I)}$ is an integral operator which acts on $L_2(I)$ by

$$\mathcal{K}_N^{(I)}(f) = \int_I K_N(x, y) f(y) \ dy.$$

So the range of $\mathcal{K}_N^{(I)}$ is finite dimensional (i.e. it is a finite rank operator). Then

$$H_N(I,t) = 1 - \int_I K_N(\lambda_1,\lambda_1) d\lambda_1 - \frac{t}{2!} \int_I \int_I \det(\quad)_{2\times 2} d^2\lambda + \dots + \frac{(-t)^j}{j!} \int_I \dots \int_I \det(\quad)_{j\times j} d^j\lambda + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots \int_I \det(\quad)_{N\times N} d^j\lambda + \dots + \frac{(-t)^N}{N!} \int_I \dots + \frac{(-t)^N}{N$$

Then $H_N(I,1)$ is the probability of no evaluations in I, and $H_N^I(I,1)$ is negative the probability of exactly one evaluation in I. So

$$H_N^{(j)}(I,1) = (-1)^j j! P(\text{exactly } j \text{ eigenvalues in } I).$$

April 22, 2025

Recall

$$\frac{1}{z_N}e^{-\frac{1}{2}\sum \lambda_j^2}\prod_{j\leq k}(\lambda_j-\lambda_k)^2=\frac{1}{N!}\det(K_N(\lambda_j,\lambda_k))_{N\times N}$$

For n = 2, we have

$$\mathbb{E}\left(\sum_{j=1}^{2}\chi_{B}(\lambda_{j})\right) = \int_{B} K_{2}(\lambda,\lambda) d\lambda$$

We also have that

$$\mathbb{E}((1-\chi_B(\lambda_1))(1-\chi_B(\lambda_2)) = P(\text{no evaluations})$$

$$= 1 - \lambda_B K_2(\lambda, \lambda) d\lambda + \frac{1}{2} \int_{R} \int_{R} \det(-)_{2\times 2} d\lambda$$

where

$$(\lambda_{1} - \lambda_{2})^{2} = \det\begin{pmatrix} 1 & \lambda_{1} \\ 1 & \lambda_{2} \end{pmatrix} \det\begin{pmatrix} 1 & 1 \\ \lambda_{1} & \lambda_{2} \end{pmatrix}$$

$$q(\lambda_{i}) = \lambda_{i} + c_{0} \det\begin{pmatrix} 1 & q(\lambda_{1}) \\ 1 & q(\lambda_{2}) \end{pmatrix} \det\begin{pmatrix} 1 & 1 \\ q(\lambda_{1}) & q(\lambda_{2}) \end{pmatrix}$$

$$= \frac{1}{(\kappa_{0}^{(0)} \kappa_{1}^{(1)})^{2}} \det\begin{pmatrix} \kappa_{0}^{(0)} & \kappa_{1}^{(1)} q(\lambda_{1}) \\ \kappa_{0}^{(0)} & \kappa_{1}^{(1)} q(\lambda_{2}) \end{pmatrix} \det\begin{pmatrix} \kappa_{0}^{(0)} & \kappa_{0}^{(0)} \\ \kappa_{1}^{(1)} q(\lambda_{1}) & \kappa_{1}^{(1)} q(\lambda_{2}) \end{pmatrix}$$

$$= \frac{1}{\prod_{0}^{1} (\kappa_{i}^{(i)})^{2}} \det\begin{pmatrix} P_{0}(\lambda_{1}) & P_{1}(\lambda_{1}) \\ P_{0}(\lambda_{2}) & P_{1}(\lambda_{2}) \end{pmatrix} \det\begin{pmatrix} P_{0}(\lambda_{1}) & P_{0}(\lambda_{2}) \\ P_{1}(\lambda_{1}) & P_{1}(\lambda_{2}) \end{pmatrix}$$

It follows that

$$\begin{split} e^{-\frac{1}{2}\sum_{1}^{2}\lambda_{j}^{2}}(\lambda_{2}-\lambda_{1})^{2} &= \prod_{j=0}^{1}(\kappa_{j}^{(j)})^{-2} \det \begin{pmatrix} P_{0}(\lambda_{1})e^{-\frac{1}{4}\lambda_{1}^{2}} & P_{1}(\lambda_{1})e^{-\frac{1}{4}\lambda_{1}^{2}} \\ P_{0}(\lambda_{1})e^{-\frac{1}{4}\lambda_{2}^{2}} & P_{1}(\lambda_{1})e^{-\frac{1}{4}\lambda_{2}^{2}} \end{pmatrix} \det \begin{pmatrix} P_{0}(\lambda_{1})e^{-\frac{1}{4}\lambda_{1}^{2}} & P_{0}(\lambda_{1})e^{-\frac{1}{4}\lambda_{2}^{2}} \\ P_{1}(\lambda_{1})e^{-\frac{1}{4}\lambda_{2}^{2}} \end{pmatrix} \\ &= \prod_{j=0}^{1}(\kappa_{j}^{(j)})^{-2} \det (K_{2}(\lambda_{i},\lambda_{j}))_{2\times 2} \end{split}$$

where $K_2(x, y) = e^{-\frac{1}{4}(x^2+y^2)} \sum_{\ell=0}^{1} P_{\ell}(x) P_{\ell}(Y)$. So we have

$$\frac{1}{z_N} \prod_{j=0}^{1} (\kappa_j^{(j)})^{-2} \det \begin{pmatrix} K_2(\lambda_1, \lambda_1) & K_2(\lambda_1, \lambda_2) \\ K_2(\lambda_2, \lambda_1) & K_2(\lambda_2, \lambda_2) \end{pmatrix}$$

and the fact that

$$\frac{1}{z_N \prod_{i=1}^2 (\kappa_i^{(j)})} \int_{\mathbb{R}^2} \left[K_2(\lambda_1, \lambda_1) K_2(\lambda_2, \lambda_2) - K_2(\lambda_1, \lambda_2)^2 \right] d\lambda_1 d\lambda_2 = 1$$

Observe that (to do: fill in these calculations)

$$\int_{\mathbb{R}^2} K_2(\lambda_1, \lambda_1) K_2(\lambda_2, \lambda_2) d\lambda_1 d\lambda_2 = \int_{\mathbb{R}^2} \left(e^{-\frac{1}{4}\lambda_1^2} (P_0(\lambda_1) P_0(\lambda_1) + P_1(\lambda_1) P_1(\lambda_1)) \left(e^{-\frac{1}{4}\lambda_2^2} (P_0(\lambda_2) P_0(\lambda_2) + P_1(\lambda_2) P_1(\lambda_2) \right) \right) d\lambda_1 d\lambda_2$$

So it must be that

$$\frac{1}{z_N(\kappa_0^{(0)})^2(\kappa_1^{(1)})^2}(2) = 1.$$

We conclude that the original joint PDF can be written as $\frac{1}{2!} \det(K_1(\lambda_i, \lambda_j))_{2 \times 2}$.

Vandermonde Determinant

Write

$$\begin{vmatrix} 1 & 1 & \cdots & 1 \\ \lambda_1 & \lambda_2 & \cdots & \lambda_N \\ \lambda_1^2 & \lambda_2^2 & \cdots & \lambda_N^2 \\ \vdots & & & \vdots \\ \lambda_1^{N-1} & \lambda_2^{N-1} & \cdots & \lambda_N^{N-1} \end{vmatrix} = \begin{vmatrix} 1 & 1 & \cdots & 1 \\ \lambda_1 & \lambda_2 & \cdots & \lambda_{N-1} \\ \lambda_1^2 & \lambda_2^2 & \cdots & \lambda_{N-1}^2 \\ \vdots & & & \vdots \\ \lambda_1^{N-2} & \lambda_2^{N-2} & \cdots & \lambda_{N-1}^{N-2} \end{vmatrix} (\lambda_N - \lambda_1)(\lambda_N - \lambda_2) \cdots (\lambda_N - \lambda_{N-1})$$

$$= \prod_{j < k} (\lambda_k - \lambda_j)$$

and observe that this is zero when $\lambda_i = \lambda_j$. Now write

$$\det \begin{pmatrix} 1 & \cdots & \lambda_1^{N-1} \\ \vdots & & \vdots \\ 1 & \cdots & \lambda_N^{N-1} \end{pmatrix} \det \begin{pmatrix} 1 & \cdots & 1 \\ \vdots & & \vdots \\ \lambda_1^{N-1} & \cdots & \lambda_N^{N-1} \end{pmatrix}$$

then by using the multilinearity of the dterminant and adding rows we can write

$$\det\begin{pmatrix} 1 & \cdots & \lambda_1^{N-1} \\ \vdots & & \vdots \\ 1 & \cdots & \lambda_N^{N-1} \end{pmatrix} \det\begin{pmatrix} 1 & \cdots & 1 \\ \lambda_1 + c_0 & \cdots & \lambda_N + c_0 \\ \pi_2(\lambda_1) & \cdots & \pi_2(\lambda_N) \\ \vdots & & \vdots \\ \lambda_1^{N-1} & \cdots & \lambda_N^{N-1} \end{pmatrix}$$

So we can write

$$\det \begin{pmatrix} e^{-\frac{1}{4}\lambda_1^2} P_0(\lambda_1) & \cdots & e^{-\frac{1}{4}\lambda_1^2} P_{N-1}(\lambda_1) \\ \vdots & & & \vdots \\ e^{-\frac{1}{4}\lambda_N^2} P_0(\lambda_N) & \cdots & e^{-\frac{1}{4}\lambda_1^2} P_{N-1}(\lambda_N) \end{pmatrix} \frac{1}{\prod_{i=0}^{N-1} (\kappa_j^{(j)})^2} \det \begin{pmatrix} P_0(\lambda) & \cdots & P_0\lambda_N \\ \vdots & & \vdots \\ P_{N-1}(\lambda_1) & \cdots & P_{N-1}(\lambda_N) \end{pmatrix}$$

Examining the (i, k) entry, we have

$$\frac{1}{\prod (\kappa_{j}^{(j)})^{2}} e^{-\frac{1}{4}(\lambda_{j}^{2} + \lambda_{k}^{2})} (P_{0}(\lambda_{j}) P_{0}(\lambda_{k}) + P_{1}(\lambda_{j}) P_{1}(\lambda_{k}) + \dots + P_{N-1}(\lambda_{j}) P_{N-1}(\lambda_{k})).$$

or

$$\frac{1}{z_N \prod (\kappa_i^{(j)})^2} \det [K_n(\lambda_j, \lambda_k)]_{N \times N}$$

which must integrate across \mathbb{R}^n to exactly 1. From this we conclude that $\frac{N!}{z_N \prod (\kappa_i^{(j)})^2} = 1$.

April 24, 2025

Determinants

$$\int_{\mathbb{R}} \det \begin{pmatrix} K_N(\lambda_1, \lambda_1) & K_N(\lambda_1, \lambda_2) \\ K_N(\lambda_2, \lambda_1) & K_N(\lambda_2, \lambda_2) \end{pmatrix} d\lambda_2 = \int_{R} K_N(\lambda_1, \lambda_1) K_N(\lambda_2, \lambda_2) - K_N(\lambda_1, \lambda_2) K_N(\lambda_2, \lambda_1) d\lambda_2$$

$$= K_N(\lambda_1, \lambda_1) \int_{\mathbb{R}} e^{-\frac{1}{2}\lambda_2^2} P_{\ell}(\lambda_2)^2 d\lambda_2 - 0$$

$$= NK_N(\lambda_1, \lambda_1)$$

We have that $\int_{\mathbb{R}} K_N(\lambda, x) K_N(x, \mu) dx = K_N(\lambda, \mu)$. Then

$$\int_{\mathbb{R}} \begin{vmatrix} K_{11} & K_{12} & K_{13} \\ K_{21} & K_{22} & K_{23} \\ K_{31} & K_{32} & K_{33} \end{vmatrix} d\lambda^{3} = \int_{\mathbb{R}} K_{31} \begin{vmatrix} K_{12} & K_{13} \\ K_{22} & K_{23} \end{vmatrix} - K_{32} \begin{vmatrix} K_{11} & K_{13} \\ K_{21} & K_{23} \end{vmatrix} + K_{33} \begin{vmatrix} K_{11} & K_{12} \\ K_{21} & K_{22} \end{vmatrix} d\lambda^{3}$$

$$= \begin{vmatrix} K_{12} & \int_{\mathbb{R}} K(\lambda_{1}, \lambda_{3}) K(\lambda_{3}, \lambda_{1}) d\lambda^{3} \\ K_{22} & \int_{\mathbb{R}} K(\lambda_{2}, \lambda_{3}) K(\lambda_{3}, \lambda_{2}) d\lambda^{3} \end{vmatrix}$$

$$- \begin{vmatrix} K_{11} & \int_{\mathbb{R}} K(\lambda_{1}, \lambda_{3}) K(\lambda_{3}, \lambda_{2}) d\lambda^{3} \\ K_{21} & \int_{\mathbb{R}} K(\lambda_{2}, \lambda_{3}) K(\lambda_{3}, \lambda_{2}) d\lambda^{3} \end{vmatrix}$$

$$+ \begin{vmatrix} K_{11} & \int_{\mathbb{R}} K(\lambda_{1}, \lambda_{2}) K(\lambda_{3}, \lambda_{3}) d\lambda^{3} \\ K_{21} & \int_{\mathbb{R}} K(\lambda_{2}, \lambda_{2}) K(\lambda_{3}, \lambda_{3}) d\lambda^{3} \end{vmatrix}$$

$$= \begin{vmatrix} K_{12} & K_{11} \\ K_{22} & K_{21} \end{vmatrix} - \begin{vmatrix} K_{11} & K_{12} \\ K_{21} & K_{22} \end{vmatrix} + N \begin{vmatrix} K_{11} & K_{12} \\ K_{21} & K_{22} \end{vmatrix}$$

$$= - \begin{vmatrix} K_{11} & K_{12} \\ K_{21} & K_{22} \end{vmatrix} - \begin{vmatrix} K_{11} & K_{12} \\ K_{21} & K_{22} \end{vmatrix} + N \begin{vmatrix} K_{11} & K_{12} \\ K_{21} & K_{22} \end{vmatrix}$$

$$= (N-2) \det()_{2 \times 2}$$

So we have that

$$\int_{\mathbb{R}} \int_{\mathbb{R}} (N-2) \begin{vmatrix} K_{11} & K_{12} \\ K_{21} & K_{22} \end{vmatrix} d\lambda_2 d\lambda_1 = (N-2)(N-1)N$$

In general, we see that

$$\int_{\mathbb{R}} \begin{vmatrix} K_{11} & \cdots & K_{1j} \\ \vdots & \ddots & \vdots \\ K_{j1} & \cdots & K_{jj} \end{vmatrix} d\lambda_j = (N - (j-1)) \begin{vmatrix} K_{11} & \cdots & K_{1j} \\ \vdots & \ddots & \vdots \\ K_{j1} & \cdots & K_{j-1j-1} \end{vmatrix}$$

or

$$\int_{\mathbb{R}} \begin{vmatrix} K_{11} & \cdots & K_{1j} \\ \vdots & \ddots & \vdots \\ K_{j1} & \cdots & K_{jj} \end{vmatrix} d\lambda_{j} = (-1)^{j+1} K_{j1} \begin{vmatrix} K_{12} & \cdots & K_{1j} \\ \vdots & \ddots & \vdots \\ K_{(j-1)2} & \cdots & K_{(j-1)j} \end{vmatrix} + (-1)^{j+2} K_{j2} \begin{vmatrix} K_{11} & \cdots & K_{1j} \\ \vdots & \ddots & \vdots \\ K_{(j-1)1} & \cdots & K_{(j-1)j} \end{vmatrix} + \cdots$$

$$\cdots + K_{jj} \begin{vmatrix} K_{11} & \cdots & K_{1(j-1)} \\ \vdots & \ddots & \vdots \\ K_{(j-1)1} & \cdots & K_{(j-1)(j-1)} \end{vmatrix} = (-1)^{j+1} \begin{vmatrix} K_{12} & \cdots & K_{1j} \\ \vdots & \ddots & \vdots \\ K_{(j-1)2} & \cdots & K_{(j-1)1} \end{vmatrix} + (-1)^{j+2} K_{j2} \begin{vmatrix} K_{11} & \cdots & K_{12} \\ \vdots & \ddots & \vdots \\ K_{(j-1)1} & \cdots & K_{(j-1)2} \end{vmatrix} + \cdots$$

$$\cdots + N \begin{vmatrix} K_{11} & \cdots & K_{1(j-1)} \\ \vdots & \ddots & \vdots \\ K_{(j-1)1} & \cdots & K_{(j-1)(j-1)} \end{vmatrix}$$

It takes, for example, j-1 column moves to convert the leading matrix into the final matrix. So it picks up a leading -1. In fact, we see that each term save the last will be negative. It follows that we have that the integral may be written

$$(N-(j-1))$$
 $\begin{vmatrix} K_{11} & \cdots & K_{1(j-1)} \\ \vdots & \ddots & \vdots \\ K_{(j-1)1} & \cdots & K_{(j-1)(j-1)} \end{vmatrix}$

Evaluations

Now consider

$$\mathbb{E}(\text{\# of evaluations in }B) = \int_{\mathbb{R}^n} \left(\sum_{j=1}^N \chi_B(\lambda_j) \right)$$
$$= \sum_{j=1}^N \int_{\mathbb{R}^n} \chi_B(\lambda_j) \frac{1}{N!} \det(\quad)_{N \times N} \, d\lambda_N \cdots d\lambda_1$$

With a change of variables where $\mu_{\ell} = \lambda_{\ell}$ for $\ell \in \{1, j\}$ such that $\mu_{j} = \lambda_{1}$ and $\mu_{1} = \lambda_{j}$,

$$\sum_{j=1}^{N} \int_{\mathbb{R}^{n}} \chi_{B}(\mu_{1}) \frac{1}{N!} \det(K_{N}(\mu_{j}, \mu_{k}))_{N \times N} d\mu_{N} \cdots d\mu_{1} = \sum_{j=1}^{N} \int_{\mathbb{R}} \chi_{B}(\mu_{1}) \frac{(N-1)!}{N!} K_{N}(\mu_{1}, \mu_{1}) d\mu_{1}$$

$$= \int_{\mathbb{R}} K_{N}(\mu_{1}, \mu_{1}) d\mu_{1}$$

Variance

Let $n_N(B) = (\# \text{ of evaluations in } B)$ be a random variable. What is the variance? We have that

$$n_N(B) = \sum_{i=1}^N \chi_B(\lambda_j)$$

SO

$$\operatorname{var}(n_{N}(B)) = \mathbb{E}((n_{N}(B))^{2}) - (\mathbb{E}(n_{N}(B)))^{2}$$

$$= \int_{\mathbb{R}^{n}} \left(\sum_{j=1}^{N} \chi_{B}(\lambda_{j})\right)^{2} \frac{1}{N!} \operatorname{det}()_{N \times N} d\lambda^{N} - []^{2}$$

$$= \int_{\mathbb{R}^{n}} \left(\sum_{k} \sum_{j} \chi_{B}(\lambda_{j}) \chi_{B}(\lambda_{k})\right) \frac{1}{N!} \operatorname{det}()_{N \times N} d\lambda^{N} - []^{2}$$

April 29, 2025

Variance

Compute

$$\operatorname{Var}(n_N(B) = \mathbb{E}(n_N(B)^2) - (\mathbb{E}(n_N(B)))^2 = \int_B K_n(\lambda_1, \lambda_1) d\lambda_1 - \int_{B \times B} K_N(\lambda_1, \lambda_2)^2 d\lambda_1 d\lambda_2$$

This follows from $n_N(B) = \sum_{j=1}^N \chi_B(\lambda_1)$, so

$$\mathbb{E}\left(\left(\sum_{i=1}^{n} \chi_{B}(\lambda_{j})\right)^{2}\right) = \mathbb{E}\left(\sum_{j=1}^{n} \sum_{k=1}^{n} \chi_{B}(\lambda_{1})\chi_{B}(\lambda_{2})\right)$$

$$= \mathbb{E}\left(\sum_{j=1}^{n} \chi_{B}(\lambda_{j})\right) + \mathbb{E}\left(\sum_{j\neq k} \chi_{B}(\lambda_{j})\chi_{B}(\lambda_{k})\right)$$

$$= \int_{B} K_{N}(\lambda_{1}, \lambda_{1}) d\lambda_{1} + \sum_{j\neq k} \int_{\mathbb{R}^{n}} \chi_{B}(\lambda_{j})\chi_{B}(\lambda_{k}) \frac{1}{N!} \det(K_{N}(\lambda_{m}, \lambda_{n}))_{N \times N} d^{N} \lambda$$

Then using the same trick as before such that $\lambda_j = \mu_1$ and $\lambda_k = \mu_2$, we rewrite this

$$\int_{B} K_{N}(\lambda_{1},\lambda_{1}) d\lambda_{1} + \sum_{i \neq k} \int_{\mathbb{R}^{n}} \chi_{B}(\mu_{1}) \chi_{B}(\mu_{2}) \frac{1}{N!} \det(K_{N}(\mu_{m},\mu_{n})_{N \times N} d^{N} \mu)$$

Then we have

$$I = \sum_{j \neq j} \int \cdots \int \chi_B(\mu_1) \chi_B(\mu_2) \frac{(1)}{N!} \det()_{(N-1) \times (N-1)} d^{N-1} \mu$$

$$= \sum_{j \neq k} \int_B \int_B \frac{(N-2)!}{N!} \det()_{2 \times 2} d\mu$$

$$= \frac{N!}{N!} \int_B \int_B \left| \frac{K_N(\mu_1, \mu_1) \quad K_N(\mu_1, \mu_2)}{K_N(\mu_2, \mu_1) \quad K_N(\mu_2, \mu_2)} \right| d^2 \mu$$

Then we have

$$\mathbb{E}(n_N(B)^2) = \int_B K_N(\lambda_1, \lambda_1) d\lambda_1 + \int_B \int_B K_N(\lambda_1, \lambda_1) K_N(\lambda_2, \lambda_2) - K_N(\lambda_1, \lambda_2)^2 d^2\lambda$$

as well as

$$\left(\mathbb{E}(n_N(B))\right)^2 = \left(\int_B K_N(\lambda_1, \lambda_1) \, d\lambda_1\right)^2$$

Then, since $\int_B \int_B K_N(\lambda_1, \lambda_1) K_N(\lambda_2, \lambda_2) d^2 \lambda = \left(\int_B K_N(\lambda_1, \lambda_1) d\lambda_1 \right)^2$, the terms cancel and we get the expression we want.

Probability of No Evaluations

Now consider $\prod_{i=1}^{N} (1 - \chi_B(\lambda_i))$ which returns 1 if there are no evaluations in B and 0 otherwise. Therefore

$$\int \prod_{j=1}^{N} (1 - \chi_B(\lambda_j)) \frac{1}{N!} \det(K_N(\lambda_j, \lambda_k)_{N \times N} d^N \lambda$$

is the probability of having zero eigenvalues in B (i.e. the probability that $n_N(B) = 0$). If we use the case where $B = (a, \infty)$, then this returns the probability that the largest eigenvalue is less than a. Consider

$$\sum_{k=1}^{N} \chi_B(\lambda_k) \prod_{\substack{j=1\\j\neq k}}^{N} (1 - \chi_B(\lambda_k))$$

and suppose we have exactly one eigenvalues (λ_3) in B. This returns 1 when we have exactly one eigenvalue in B and 0 otherwise. So

$$\int \sum \chi_B(\lambda_k) \prod_{j=1}^N {k \choose j} (1 - \chi_B(\lambda_1) \frac{1}{N!} \det()_{N \times N} d^N \lambda,$$

where the product skips the k-th term, is the probability $\mathbb{P}\{n_N(B)=1\}$. Now write

$$H(B,t) = \mathbb{E}\left(\prod_{j=1}^{N} (1 - t\chi_B(\lambda_j))\right)$$

which gives $H(B,1) = \mathbb{P}\{n_N(B) = 1\}$. Then the derivative with respect to t,

$$H'(B,t) = \mathbb{E}\left(\sum_{k=1}^{N} (-\chi_B(\lambda_k)) \prod_{j=1}^{N} {k \choose j} (1 - t\chi_B(\lambda_j))\right)$$

so $H'(B,1) = -\mathbb{P}\{n_N(B) = 1\}$, and $H''(B,1) = 2\mathbb{P}\{n_N(B) = 2\}$. It follows that $H^j(B,1) = (-1)^j \cdot j! \cdot \mathbb{P}\{n_N = j\}$. Then $n_N(B)$ is the number of evaluations in B, and this process gives us the number statistics. We compute this fact as follows

$$H''(B,t) = \mathbb{E}\left(\sum_{k=1}^{N} \chi_{B}(\lambda_{k}) \sum_{\ell=1}^{N} {k \choose k} \chi_{B}(\lambda_{\ell}) \prod_{j=1}^{N} {k,\ell \choose j} (1 - t\chi_{B}(\lambda_{j}))\right)$$

$$= \mathbb{E}\left(\sum_{k \neq \ell} \chi_{B}(\lambda_{k}) \chi_{B}(\lambda_{\ell}) \prod_{j=1}^{N} {k,\ell \choose j} (1 - t\chi_{B}(\lambda_{j}))\right)$$

$$= 2! \cdot \mathbb{P}\{n_{N}(B) = 2\} \quad (t = 1)$$

and

$$H^{j}(B,t) = \mathbb{E}\left(\sum_{k_{1}=1}^{N} \sum_{k_{2}=1}^{N} {}^{(k_{1})} \cdots \sum_{k_{j}=1}^{N} {}^{(k_{1},k_{2},\dots,k_{j-1})} \prod_{v=1}^{j} \chi_{B}(\lambda_{k_{i_{v}}}) \prod_{j=1}^{N} {}^{\prime}(1 - t\chi_{B}(\lambda_{j}))\right)$$
$$= (-1)^{j} \cdot j! \cdot \mathbb{P}\{n_{N}(B) = j\} \quad (t = 1)$$

Coming Next

We know that $\mathbb{E}(n_N(B)) = \int_B K_N(\lambda, \lambda) d\lambda$. We will define an integral operator on functions $f \in L^2(B)$

$$\mathcal{K}_{N}^{(B)}(f) = \int_{B} K_{N}(x, y) f(y) \, dy$$
$$= \int_{B} e^{-\frac{1}{4}(x^{2} + y^{2})} \sum_{\ell=0}^{N-1} P_{\ell}(x) P_{\ell}(y) f(y) \, dy$$

We can define the trace of this operator,

$$\operatorname{Tr}(\mathcal{K}_N^{(B)}) = \int_B K_N(\lambda, \lambda) \ d\lambda = \mathbb{E}(n_N(B))$$

Then

$$H(B,t) = \det(1 - t\mathcal{K}_N^{(B)})$$

May 6, 2025

Expectation

Recall that

$$\mathbb{E}\left(\prod_{j=1}^{N}(1-t\chi_{B}(\lambda_{i}))\right) = \mathbb{E}\left(1-t\sum_{j=1}^{N}\chi_{B}(\lambda_{j})+t^{2}\sum_{j\leq k}\chi_{B}(\lambda_{j})\chi_{B}(\lambda_{k})-t^{3}\sum_{j\leq k< m}\chi_{B}(\lambda_{j})\chi_{B}(\lambda_{k})\chi_{B}(\lambda_{k})\chi_{B}(\lambda_{m})\right)$$

$$=1-t\int_{B}K_{N}(\lambda_{1},\lambda_{1}) d\lambda_{1}+t_{2}\sum_{j\leq k}\int\chi_{B}(\lambda_{j})\chi_{B}(\lambda_{k})\frac{1}{N!}\det()_{N\times N} d^{N}\lambda$$

$$=1-t\int_{B}K_{N}(\lambda_{1},\lambda_{1}) d\lambda_{1}+\frac{t^{2}}{2}\int_{B}\int_{B}|_{2\times 2} d^{2}\lambda-\frac{t^{3}}{3!}\int_{B}\int_{B}|_{3\times 3} d^{3}\lambda+\cdots+\int_{B}\cdots\int_{B}|_{N\times N} d^{N}\lambda$$

We claim that

$$\int_{\mathbb{R}^n} \left(\sum_{k_1 < \dots < k_j} \prod_{\ell=1}^j \chi_B(\lambda_{K_\ell}) \frac{\det}{N!} (K_N(\lambda_m, \lambda_n)_{N \times N}) d^N \lambda = \frac{1}{j!} \int_B \dots \int_B \det(K_N(\lambda_m, \lambda_n))_{j \times j} d\lambda_j \dots d\lambda_1 \right)$$

Observing that $\binom{N}{j} \cdot \frac{(N-j)!}{N!} = \frac{N!}{j!(N-j)!} = \frac{1}{j!}$, write

$$\begin{split} \sum_{k_1 < \dots < k_j} \int_{\mathbb{R}^n} \left(\prod_{\ell=1}^j \chi_B(\lambda_\ell) \right) \frac{\det}{N!} ()_{N \times N} \, d\lambda_{j+1} \dots d\lambda_1 &= \sum_{k_1 < \dots < k_j} \int_{\mathbb{R}^{n-1}} \left(\prod \chi_B(\lambda_\ell) \right) \frac{\det}{N!} ()_{(N-1) \times (N-1)} \, d\lambda_{N-1} \dots d\lambda_1 \\ &= \sum_{k_1 < \dots < k_j} \int_{\mathbb{R}^j} \left(\prod_{\ell=1}^j \chi_B(\lambda_\ell) \right) \frac{1 \cdot 2 \cdot 3 \cdot \dots \cdot (N-j)}{N!} \det()_{j \times j} \, d^j \lambda \\ &= \frac{1}{j!} \int \dots \int \det()_{j \times j} \, d^j \lambda \end{split}$$

So we have

$$H(B,t) = 1 + \sum_{j=1}^{N} \frac{(-1)^{j} t^{j}}{j!} \int_{B} \cdots \int_{B} \det \begin{pmatrix} K_{N}(\lambda_{1}, \lambda_{1}) \\ K_{N}(\lambda_{j}, \lambda_{j}) \end{pmatrix} d^{j} \lambda = \det(1 - t \mathcal{K}_{N})$$

So if A is a symmetric, real $N \times N$ matrix with ||A|| < 1,

$$\det(I - tA) = e^{\log \det(I - tA)} = e^{\log \prod_{j=1}^{N} (1 - t\mu_j^A)} = e^{\sum_{j=1}^{N} \log(1 - t\mu_j^A)} = e^{\operatorname{Tr}\log(I - tA)}$$

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Continuing

We want to show that $\det(I - tA) = e^{\operatorname{Tr} \log(I - tA)}$. Write

$$\log(1-x) = \int_0^x -\frac{1}{1-s} \, ds = -\int_0^x \sum_{j=0}^\infty s^j \, ds = -\sum_{j=0}^\infty \frac{x^{j+1}}{j+1} = -\sum_{j=1}^\infty \frac{x^j}{j}.$$

So $\log(I-tA) := -\sum_{j=1}^{\infty} \frac{(tA)^j}{j}$ converges in the matrix norm (for sufficiently small t). Equivalently, $e^{-\sum_{j=1}^{\infty} \frac{(tA)^j}{j}} = I - tA$. Suppose we have an eigenvalue μ of A, then $\log(1-t\mu)$ is an eigenvalue of $\log(I-tA)$. More commonly, this is presented as $\log\det(I-tA) = \operatorname{Tr}\log(I-tA)$. In the end, we have

$$\det(I - tA) = \exp[\operatorname{Tr}(\log(I - tA))] = \exp\left[-\operatorname{Tr}\sum_{j=1}^{\infty} \frac{(tA)^{j}}{j}\right].$$

Trace of an Operator

Given $\mathcal{K}_n(f) = \int_B K_N(x,y) f(y) \, dy$, we define $\mathrm{Tr}(\mathcal{K}_n) = \int_B K_N(x,x) \, dx$. We want to consider the trace of \mathcal{K}_N^j . Taking j=2 as an example,

$$\mathcal{K}_{N}^{2}(f) = \mathcal{K}_{N} \left[\int_{B} K_{N}(x_{1}, x_{2}) f(x_{2}) dx_{2} \right]$$

$$= \int_{B} K_{N}(x, x_{1}) \int_{B} K_{N}(x_{1}, x_{2}) f(x_{2}) dx_{2} dx$$

$$= \int_{B} \left[\int_{B} K_{N}(x, x_{1}) K_{N}(x_{1}, x_{2}) dx \right] f(x_{2}) dx_{2}$$

we note that $\int_B K_N(x,x_1)K_n(x_1,x_2)\ dx$ is our new kernel. Then $\mathrm{Tr}(\mathcal{K}_N^2)=\int_B \int_B K_N(x_1,x_2)K_N(x_2,x_1)\ d^2x$. Therefore \mathcal{K}_N^ℓ has a kernel given by

$$\int_{B}^{\ell-1} \cdots \int_{B} K_{N}(x,x_{1}) \cdots K_{N}(x_{\ell-1},x_{\ell}) dx_{1} \cdots dx_{\ell-1}.$$

with trace given by

$$\operatorname{Tr}(\mathcal{K}_N^{\ell}) = \int_B^{\ell} \cdots \int_B^{\ell} K_N(x_{\ell}, x_1) \cdots K_N(x_{\ell-1}, X_{\ell}) \ d^{\ell} x.$$

Continuing Computation

$$\det(I - tA) = \sum_{k=0}^{\infty} \frac{1}{k!} \left[-\operatorname{Tr} \sum_{j=1}^{\infty} \frac{(tA)^{j}}{j} \right]^{k} = 1 - \left[\sum_{j=1}^{\infty} \frac{t^{j}}{j} \operatorname{Tr}(A^{j}) \right] + \frac{1}{2!} \left[\sum_{j=1}^{\infty} \frac{t^{j}}{j} \operatorname{Tr}(A^{j}) \right]^{2} - \frac{1}{3!} \left[\sum_{j=1}^{\infty} \frac{t^{j}}{j} \operatorname{Tr}(A^{j}) \right]^{3} + \cdots$$

If we are intersetd in $O(t^1)$, we examine the first non-trivial term and have -Tr(A). For $O(t^2)$ we have $-\frac{1}{2}\text{Tr}(A^2) + \frac{1}{2}(\text{Tr}A)^2$, which we compute as

$$\frac{1}{2} \left[\int_{B} K_{N}(\lambda_{1}, \lambda_{1}) d\lambda_{1} \int_{B} K_{N}(\lambda_{2}, \lambda_{2}) d\lambda_{2} - \int_{B} \int_{B} K_{N}(\lambda_{1}, \lambda_{2}), K_{N}(\lambda_{2}, \lambda_{1}) d^{2}\lambda \right]$$

Examining $O(t^3)$, we have $-\frac{1}{3}\operatorname{Tr}(A^3) + \frac{1}{4}(\operatorname{Tr} A)\operatorname{Tr}(A^2) + \frac{1}{4}\operatorname{Tr}(A^2)(\operatorname{Tr} A) - \frac{1}{3!}(\operatorname{Tr} A)^3$ giving

$$-\frac{t^{3}}{3!} \left[(\operatorname{Tr} A)^{3} - \frac{3}{2} (\operatorname{Tr} A) \operatorname{Tr} (A^{2}) - \frac{3}{2} \operatorname{Tr} (A^{2}) (\operatorname{Tr} A) + 2 \operatorname{Tr} (A^{3}) \right]$$

which we compare with

$$-\frac{t^{3}}{3!} \left[\int_{B} \int_{B} \int_{B} \left| \begin{array}{ccc} K_{11} & K_{12} & K_{13} \\ K_{21} & K_{22} & K_{23} \\ K_{31} & K_{32} & K_{33} \end{array} \right| d^{3}\lambda \right] = -\frac{t^{3}}{3!} \left[\int_{B} \int_{B} \int_{B} K_{11}(K_{22}K_{33} - K_{23}K_{32}) - K_{12}(K_{21}K_{33} - K_{23}K_{31}) + K_{13}(K_{21}K_{32} - K_{22}K_{31}) + K_{13}(K_{21}K_{32} - K_{22}K_{31}) \right] d^{3}\lambda d^{3}\lambda$$

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Recall

 $H(B,t) = \mathbb{E}\left(\prod_{j=1}^{N}(1-t\chi_B(\lambda_j))\right) \stackrel{!}{=} \det(1-t\mathcal{K}_N^B)$, the generating function for number statistics.

Observations

We have that $\mathcal{K}_N^B(f) = \int_B K_N(x,y) f(y) \, dy$ where $K_N(xy) = e^{-\frac{x^2+y^2}{4}} \sum_{\ell=0}^{N-1} P_\ell(x) P_\ell(y)$. We know that \mathcal{K}_N^B is bounded and, taking $B = \mathbb{R}$, we see that (1) $||\mathcal{K}_N^B|| < 1$ since $\mathcal{K}_N^\mathbb{R}$ is the projection onto the first N orthonormal functions. Homework: prove this inequality (1).

(2) It follows that $(1 - tK_N)$ is invertible for small t (namely for t = 1 for $N < \infty$). (3) It is also true that K_N^B is a finite-rank operator.

Then we may pick an $L^2(B)$ basis such that $1 - t\mathcal{K}_N$ is realized as a matrix

$$1 - t\mathcal{K}_N = \begin{bmatrix} I - t\mathfrak{K}_N & 0 \\ 0 & I \end{bmatrix}$$

For example, we can orthonormalize $\{e^{-\frac{x^2}{4}}x^j\}_{j=0}^{N-1}$ and take the first N.

Homework: do this. $\int_{B} ? q_{j} ?$.

Now $\det(1 - t\mathcal{K}_N) = e^{\operatorname{Tr}\log(1 - t\mathcal{K}_N)}$ where both trace and log are basis agnostic. Then for \mathcal{K}_N of finite rank,

$$\operatorname{Tr} \mathcal{K}_N = \sum_{j=1}^N (\mathfrak{K}_N)_{jj} = \int K_N(x, x) \, dx$$

It follows that

$$\det(1 - t\mathcal{K}_N) = \exp\left(-\operatorname{Tr}\sum_{j=1}^{\infty} \frac{t^j(\mathcal{K}_N)^j}{j}\right) = \exp\left(-\sum_{j=1}^{\infty} \frac{t^j}{j}\int_{B} \cdots \int_{B} K_N(\lambda_1, \lambda_2) \cdots K_N(\lambda_j, \lambda_1) d^j \lambda\right)$$

Coefficients

So

$$1 - t \int_{B} K_{N}(\lambda_{1}, \lambda_{1}) d\lambda_{1} + \dots + \frac{(-1)^{j}}{j!} t^{j} \int \dots \int \det[K_{N}(\lambda_{m}, \lambda_{n})_{j \times j} d^{j} \lambda + \frac{(-1)^{N} t^{N}}{N!} \int \dots \int \det[K_{N}(\lambda_{1}, \lambda_{1})_{j \times j} d^{j} \lambda + \frac{(-1)^{N} t^{N}}{N!} \int \dots \int_{B} K_{N}(\lambda_{1}, \lambda_{2}) \dots K_{N}(\lambda_{j}, \lambda_{1}) d^{j} \lambda d^{j}$$

Where on the left-hand side we have coefficients

$$\frac{(-1)^n t^n}{n!} \int_{B} \cdots \int_{B} \det [K_N(\lambda, \lambda)]_{n \times n} d^n \lambda$$

and on the right-hand side

$$1 - \left(-\operatorname{Tr}\sum_{j=1}^{\infty} \frac{t^{j} (\mathcal{K}_{N})^{j}}{j}\right) + \frac{1}{2!} \left(-\right)^{2} - \frac{1}{3!} \left(-\right)^{3} + \dots + \frac{\left(-1\right)^{n}}{n!} \left(-\right)^{n}$$

so

$$-\frac{1}{n}\operatorname{Tr}(\mathcal{K}_{N}^{n})+\frac{1}{2!}\left[\frac{\operatorname{Tr}(\mathcal{K}_{N})\operatorname{Tr}(\mathcal{K}_{N}^{n-1})}{1\cdot(n-1)}+\frac{\operatorname{Tr}(\mathcal{K}_{N}^{2})\operatorname{Tr}(\mathcal{K}_{N}^{n-2})}{2\cdot(n-2)}+\cdots+\frac{\operatorname{Tr}(\mathcal{K}_{N}^{n-1})\operatorname{Tr}(\mathcal{K}_{N})}{(n-1)\cdot1}\right]-\cdots$$

Note that the *i*-th power is constructed by the integer partitions $\mu \vdash n$ for $|\mu| = \ell$. Then the *n*-th term is

$$t^{n} \sum_{\mu \vdash n} \frac{(-1)^{\ell}}{\ell!} \frac{\prod_{k=1}^{\ell} (\operatorname{Tr}(\mathcal{K}_{N}^{\mu_{k}}))}{\prod_{k=1}^{\ell} \mu_{k}} \tilde{C}_{n}(\mu)$$

where $\tilde{C}_n(\mu)$ is the number of unordered partitions associated to μ . Returning to the left-hand side,

$$\int \cdots \int \sum_{\sigma \in S_n} \operatorname{sign}(\sigma) K_N(\lambda_1, \lambda_{\sigma(1)}) K_N(\lambda_2, \lambda_{\sigma(2)}) \cdots K_N(\lambda_n, \lambda_{\sigma(n)}) d^n \lambda$$

When $\sigma = id$,

$$\int \cdots \int K_N(\lambda_1,\lambda_1)\cdots K_N(\lambda_n,\lambda_n) d^n \lambda$$

We can consider the integer partition μ representing the cycle structure of σ . Then if $\tilde{\sigma}$ are permutations with cycles of lengths associated to μ ,

$$\sum_{\mu \vdash n} \sum_{\tilde{\sigma}} \operatorname{sign}(\tilde{\sigma}) \prod_{k=1}^{\ell} \operatorname{Tr}(\mathcal{K}_N^{\mu_k}) = \sum_{\mu \vdash n} \prod_{k=1}^{\ell} \left(\operatorname{Tr}(\mathcal{K}_N^{\mu_k}) \right) \sum_{\tilde{\sigma}} \operatorname{sign}(\tilde{\sigma}) = \sum_{\mu \vdash n} \prod_{k=1}^{\ell} \left(\operatorname{Tr}(\mathcal{K}_N^{\mu_k}) \right) C_n(\mu) \operatorname{sign}(\mu)$$

since $sign(\tilde{\sigma}) = sign(\tilde{\tilde{\sigma}})$ if their cycle lengths agree. Then

$$\frac{(-1)^n}{n!}\operatorname{sign}(\mu)C_n(\mu) = \frac{(-1)^{\ell}}{\ell!} \frac{\tilde{C}_n(\mu)}{\prod_{k=1}^{\ell} \mu_k}$$

Returning to the Microscopic Regime

$$\mathbb{P}(\text{no evaluations in }(a,b)) = \det(1-\mathcal{K}_N)$$

on $L^2(a,b)$. Suppose $a=\sqrt{N}\alpha$ and $b=\sqrt{N}\left(\alpha+\frac{S}{N}\right)$. Then $K_N(x,y): x=\alpha\sqrt{N}+\frac{\xi}{\sqrt{N}}, y=\alpha\sqrt{N}+\frac{\eta}{\sqrt{N}}$ for ξ and η in a range between 0 and S. Write

$$\mathcal{K}_N(f) = \int_a^b K_N(x,y) f(y) \ dy = \int_0^S \frac{1}{\sqrt{N}} K_N\left(x = \alpha \sqrt{N} + \frac{\xi}{\sqrt{N}}, y = \alpha \sqrt{N} + \frac{\eta}{\sqrt{N}}\right) f\left(\alpha \sqrt{N} + \frac{\eta}{\sqrt{N}}\right) d\eta$$

so that we may define a new kernel

$$\tilde{K}(\xi,\eta) = \frac{1}{\sqrt{N}}K\left(x = \alpha\sqrt{N} + \frac{\xi}{\sqrt{N}}, y = \alpha\sqrt{N} + \frac{\eta}{\sqrt{N}}\right)$$

which acts on functions in (0,S). We can show that $\mathbb{P}(\text{no evaluations in }(a,b)) \stackrel{?}{=} \det(1-\tilde{\mathcal{K}}_N)$. We may write

$$\frac{1}{\sqrt{N}}K\left(x=\alpha\sqrt{N}+\frac{\xi}{\sqrt{N}},y=\alpha\sqrt{N}+\frac{\eta}{\sqrt{N}}\right)=\Box\frac{\sin(\pi(\xi-\eta))}{\pi(\xi-\eta)}+O\left(\frac{1}{n}\right)$$

 $(\Box \text{ some appropriate constant) and can prove that } \det(1-\tilde{\mathcal{K}}_N) = \det[1-\mathcal{S}]_{L^2(0,S)} \text{ where } \mathcal{S} \text{ has the rewritten kernel.}$