

1 Capturing the Epistemic Uncertainty in Subduction Earthquake Rupture Parameters

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18 **Declaration of interest statement**

19 The authors declare that they have no known competing financial interests or personal
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21 **Abstract**

22 Subduction earthquakes are highly destructive and can trigger devastating tsunamis, highlighting
23 the need for a consistent probabilistic seismic–tsunami hazard assessment. However, current
24 approaches treat both hazards separately: tsunami models use coseismic slip distributions, while
25 seismic models rely mostly on empirical ground-motion prediction equations. This creates
26 inconsistencies because both depend on different representations of the source. Here, we address
27 this inconsistency by systematically quantifying epistemic uncertainties in ground-motion
28 simulations derived from kinematic rupture models based on slip distributions of large
29 subduction earthquakes, enabling a common and physically consistent source representation for
30 both hazards. We use the 2015 Illapel Mw 8.3 earthquake ground-motion in Central Chile as a
31 test study. We test different rupture velocities, rise-time parameterisations, as well as hypocentre
32 locations using two one-dimensional velocity models to capture the epistemic uncertainty of
33 these kinematic rupture parameters and evaluate the impact on ground-motion. Our results show
34 that rupture time, controlled by rupture velocity, is the main source of epistemic uncertainty. We
35 compare the resulting simulations with the observed ground-motion at eight stations. Bias and
36 RMSE decrease when adopting depth-dependent rupture velocities, with 0.3 of shear-wave
37 velocity (V_s) in the shallow domain (<10 km) and 0.5 V_s at depths greater than 15 km. Rupture
38 velocity is then identified as the key parameter controlling ground-motion variability in
39 kinematic source models. For this event, the simulations fit the observations better when the
40 regional velocity model is used, and they more accurately reproduce ground-motion in the
41 frequency band of 0.5–3 Hz.

42 **Keywords:** Epistemic uncertainties, Kinematic rupture propagation, Subduction earthquakes,
43 Coseismic slip distribution, Central Chile.

44 **1. Introduction**

45 Subduction earthquakes pose a severe hazard to coastal populations, causing loss of life and
46 substantial economic damage (Løvholt et al., 2012, 2015). Probabilistic seismic–tsunami hazard
47 assessment, therefore, plays a crucial role in subduction zones. However, it remains a global
48 challenge as seismic and tsunami hazard are often treated as independent phenomena, each based
49 on different rupture representations for the same earthquake scenario (e.g., De Risi and Goda,
50 2016; Park et al., 2017; Goda et al., 2021). In practice, Ground Motion Prediction Equations
51 (GMPEs) are used to estimate Earthquake Intensity Measures (EIMs), while tsunami intensity
52 measures are estimated from numerical inundation models based on coseismic slip distributions,
53 so that both hazards rely on different and often incompatible descriptions of the earthquake
54 source, leading to inconsistencies (Figure 1a).

55 One way to mitigate this problem is to use the same coseismic slip distribution employed in
56 tsunami inundation modelling to generate ground motion and derive EIMs. This ensures a
57 consistent source representation for both hazards (Figure 1b). However, earthquake rupture along
58 subduction interfaces is a complex process that strongly influences ground-motion characteristics
59 (e.g., Yokota et al., 2011; Kurahashi and Irikura, 2013).

60 Translating a static slip model into a ground-motion simulation at a desired frequency, therefore,
61 poses an additional challenge, as it requires defining a propagation rupture model. Fully dynamic
62 rupture models have therefore been developed to simulate the physical processes governing
63 earthquake rupture, incorporating friction laws, stress evolution, and energy dissipation.

64 However, these models are often computationally expensive (e.g., Andrews, 1976; Madariaga et
65 al., 1998; Andrews, 2005; Gabriel et al., 2012, 2013). To reduce this cost, some studies have
66 proposed hybrid approaches, such as pseudo-dynamic models. Such models combine kinematic

67 descriptions with physics-based constraints and dynamic rupture modelling (e.g., Guatteri et al.,
68 2004; Crempien and Archuleta, 2017; Castro-Cruz and Mai, 2025).

69 Although these models better capture the physics of the rupture process, they are rarely used
70 to constrain numerical simulations of tsunami inundation. This is mainly because tsunami
71 modelling is primarily sensitive to the principal vertical deformation of the seafloor above the
72 fault. As a result, a static Okada approach (Okada, 1985) is commonly adopted as a reasonable
73 and computationally efficient approximation for tsunami inundation, thereby reducing the need
74 for fully dynamic rupture simulations.

75 Instead of full dynamic rupture simulations, most tsunami applications then use stochastic slip
76 distribution, which reproduces the variability and spectral characteristics of subduction
77 earthquake slip distributions relevant to tsunami generation (e.g., Goda et al., 2015; Melgar et al.,
78 2016; Cienfuegos et al., 2018; Crempien et al., 2020; Scala et al., 2020; Small and Melgar, 2023;
79 Buenrostro et al., 2026).

80 To describe the space-time evolution of the rupture process needed for ground-motion
81 simulations, a kinematic framework requires defining key rupture parameters, such as rupture
82 time (t_{rup}), rupture velocity (V_r) and rise time (t_0). Such models commonly assume a constant V_r
83 as a fraction of the shear-wave velocity (V_s), typically ranging from 0.8 to 0.9 (e.g., Heaton,
84 1990; Herrero and Bernard, 1994; Ruiz et al., 2011; Venegas-Aravena, 2023).

85 However, most applications that adopt stochastic slip distributions and a kinematic rupture
86 process focus on crustal earthquake scenarios, mostly located in California (e.g., Graves and
87 Pitarka, 2010, 2015, 2016; Pitarka et al., 2022). Some approaches, such as those by Melgar et al.
88 (2016), extend these methods to subduction-zone earthquakes, primarily for tsunami hazard

89 assessments. Their model adopts depth-dependent rupture velocities from Graves and Pitarka
90 (2010, 2015), rise-time scaling from Somerville et al. (1999), and variable hypocentre locations.
91 In general, these methods, called semi-kinematic methods, follow similar procedures to generate
92 stochastic slip distributions. However, the main source of epistemic uncertainty lies in the
93 kinematic rupture parameters, particularly in V_r , t_0 , and hypocentre location, which remain
94 poorly constrained by observations and the impact of these uncertainties is not evaluated.
95 To address this concern, in this study, we quantify epistemic uncertainties in ground-motion
96 simulations using kinematic rupture models from slip distributions. As a case study, we use the
97 2015 Illapel Mw 8.3 earthquake in Central Chile. We explore different formulations for t_{rup} ,
98 which depends on V_r , t_0 and hypocentre location following previous empirical and theoretical
99 studies, as well as our proposed approach (e.g., Heaton, 1990; Somerville et al., 1999; Mai et al.,
100 2005; Di Toro et al., 2011; Graves and Pitarka, 2015 ; Melgar et al., 2016; Melgar and Hayes,
101 2017, 2019; Goldberg and Melgar, 2020; Goldberg et al., 2022). Such parameters are evaluated
102 using two one-dimensional (1D) velocity models representing subduction-zone crustal structures
103 (CRUST 2.0: Bassin et al., 2000; Central Chile: Caballero et al., 2023) to assess their effects on
104 the simulated ground-motion at eight near-field stations.
105 We simulate ground-motions up to 2 Hz and band-pass filter both observed and simulated
106 records between 0.1 and 1 Hz. By comparing observed and simulated geometric-mean response
107 spectra, we evaluate the sensitivity of ground-motion predictions to the assumed rupture process
108 and velocity structure, calculating metrics such as bias, Standard Deviation (SD) of residuals,
109 and Root Mean Square Error (RMSE). Additionally, we test whether the ground-motion of a
110 single event, such as the 2015 Illapel earthquake, is captured by the range of simulations. For
111 this, we analyse the Probability Density Functions (PDFs) from all filtered simulation cases for

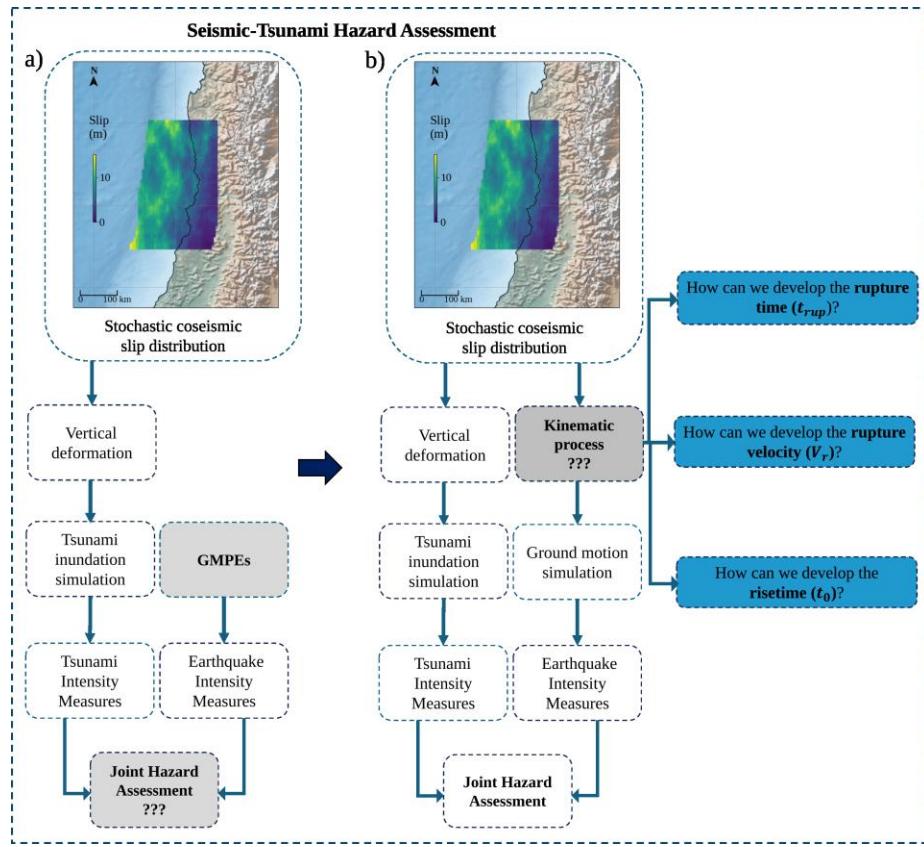
112 several EIMs, including Peak Ground Acceleration (PGA) and Spectral Accelerations (SA) at
113 0.5, 2.0, and 10.0 s. Finally, we perform a frequency reliability analysis of the simulations on the
114 models which are showing the best fit to the data, evaluating both the bias factor and the SD of
115 residuals.

116 Our approach provides a novel framework that links kinematic rupture modelling with epistemic
117 uncertainty quantification, bridging seismic and tsunami hazard assessment by adopting a
118 common source description and reducing epistemic uncertainty by identifying and constraining
119 the rupture parameters that control ground-motion variability, thereby enabling consistent
120 physics-based scenarios for probabilistic hazard analysis in subduction zones.

121 **2. Methodology**

122 *2.1. Kinematic rupture process*

123 We address the inconsistency in conventional seismic-tsunami hazard assessment by proposing a
124 unified framework that incorporates a kinematic rupture process, thereby ensuring consistency
125 between the two hazards (Figure 1). We introduce a workflow through a logic tree approach
126 (Figure 2: Step 1) that evaluates the variability of key kinematic parameters: t_{rup} , V_r , t_0 and
127 hypocentre location using two different 1D velocity models. As a case study, we use the rupture
128 characteristics and ground-motion variability of the 2015 Illapel Mw 8.3 megathrust earthquake
129 in Central Chile.



130

131 Figure 1. Seismic–tsunami hazard assessment: a) conventional approach, where tsunami and
 132 earthquake intensity measures are derived from independent source representations; b) proposed
 133 framework introducing a kinematic rupture process, enabling both hazards to originate from the
 134 same stochastic slip distribution (modified from Buenrostro et al., 2025). The framework
 135 requires defining the key kinematic parameters: rupture time (t_{rup}), rupture velocity (V_r) and rise
 136 time (t_0).

137 *2.1.1. Coseismic slip distribution*

138 We use the coseismic slip distribution derived from the inversion framework presented by the
 139 U.S. Geological Survey (USGS: see Data and Resources) as input for the kinematic process
 140 (Figure 2, Step 1). The finite fault plane extends 300 km along strike and 158.4 km down dip,
 141 with a spatial resolution of 12 km along strike and 8.8 km down dip, resulting in 450 subfaults.

142 Each subfault represents a discrete slip element within the fault plane. We adopt the focal
 143 mechanism parameters from the USGS inversion.
 144 Because we use the final slip distribution at each subfault, the local slip direction is not explicitly
 145 defined. The slip direction varies spatially during rupture and cannot be assumed to be constant
 146 (e.g., Cotton and Campillo, 1995; Guatteri and Cocco, 1996). To account for these local
 147 variations, we estimate the rake for each subfault following the approach of Graves and Pitarka
 148 (2016) and Melgar et al. (2016), with a mean rake angle of 105°, consistent with the overall rake
 149 of the earthquake.

150 *2.1.2 Velocity models*

151 We estimate Green's functions for two 1D velocity models (Figure 2, Step 1) representing
 152 subduction-zone crustal structures: the CRUST 2.0 model (Bassin et al., 2000) and a regional
 153 model for Central Chile (Caballero et al., 2023). Key kinematic parameters: t_{rup} , V_r , t_0 and
 154 hypocentre location defined in the logic tree are evaluated for both models.

155 *2.1.3. Rupture time and rupture velocity*

156 We calculate the t_{rup} using Eq. (1), defined as the time when rupture initiates at the nucleation
 157 point and propagates radially across the fault plane. For each subfault, the rupture time t_{rup}
 158 depends on the distance d_i to the hypocentre and the average rupture velocity $\bar{V}_{r,i}$ along the
 159 propagation path,

$$160 \quad t_{rup,i} = \frac{d_i}{\bar{V}_{r,i}}. \quad (1)$$

161 To capture epistemic uncertainty in rupture propagation, we test five different V_r models, all
 162 defined as a fraction of V_s . These models control the spatiotemporal evolution of the rupture and
 163 allow a systematic evaluation of how V_r influence near-field ground-motion. The first case
 164 adopts a depth-dependent scaling, z , for subduction proposed by Melgar et al. (2016),

165
$$V_r = \begin{cases} 0.56 V_s, & z < 10 \text{ km}, \\ 0.80 V_s, & z > 15 \text{ km}. \end{cases} \quad (2)$$

166 The second case follows a common assumption in kinematic modelling, where V_r is constant
 167 (Heaton, 1990),

168
$$V_r = 0.80 V_s. \quad (3)$$

169 The third case applies the subduction scaling of Goldberg and Melgar (2020),

170
$$V_r = \begin{cases} 0.49 V_s, & z < 10 \text{ km}, \\ 0.65 V_s, & z > 15 \text{ km}. \end{cases} \quad (4)$$

171 The fourth case employs the USGS model for the Illapel earthquake according to Goldberg et al.
 172 (2022), which incorporates an average rupture velocity (\bar{V}_r) of 1.5 km/s and a depth-dependent
 173 scaling,

174
$$V_r = \begin{cases} 0.28 V_s, & z < 10 \text{ km}, \\ 0.40 V_s, & z > 15 \text{ km}. \end{cases} \quad (5)$$

175 The last case defines our preferred model, as it yields an \bar{V}_r of approximately 2 km/s, consistent
 176 with regional studies (e.g., Riquelme et al., 2020). This model also includes a depth-dependent
 177 scaling following,

178
$$V_r = \begin{cases} 0.30 V_s, & z < 10 \text{ km}, \\ 0.50 V_s, & z > 15 \text{ km}. \end{cases} \quad (6)$$

179 *2.1.4. Rise time*

180 We test four alternative cases for t_0 , which is the duration of the slip. In the first three cases, t_0
 181 follows the formulation of Graves and Pitarka (2010), where the local t_0 of each subfault, τ_{o_i} ,
 182 scales with the square root of slip, $s_i^{1/2}$, and varies with depth according to Kagawa et al.
 183 (2004). For subduction earthquakes, we apply the modification proposed by Melgar et al. (2016),
 184 which adjusts the depth-dependent limits,

185
$$\tau_{o_i} = \begin{cases} 2k s_i^{1/2}, & z < 10 \text{ km}, \\ k s_i^{1/2}, & z > 15 \text{ km}, \end{cases} \quad (7)$$

186 Assuming a linear transition between 10 and 15 km. The proportionality constant k is determined
 187 by the average rise time, t_a , over the entire fault matches a target rise time. This t_a is defined by
 188 empirical relations between t_0 and seismic moment (M_0) in $N \cdot m$. We test three formulations for
 189 t_a . The first is prescribed by Somerville et al. (1999),

190
$$t_a = 4.308 \times 10^{-7} M_0^{1/3}. \quad (8)$$

191 The second relation, from Graves and Pitarka (2015), introduces a geometry-dependent scaling
 192 factor,

193
$$t_a = \alpha_\tau 3.124 \times 10^{-7} M_0^{1/3}, \quad (9)$$

194 where α_τ depends on dip and rake. The third is the formulation from Melgar and Hayes (2017),

195
$$t_a = 4.74 \times 10^{-6} M_0^{0.293}. \quad (10)$$

196 As a fourth case, we test the formulation of Di Toro et al. (2011), assuming a constant slip rate,
 197 \dot{S} , of approximately 1 m/s, and taking directly from this relation the t_0 ,

198
$$\dot{S} = s_i/t_0 \sim 1 \text{ m/s}, \quad (11)$$

199
$$t_0 = s_i/1 \text{ m/s}. \quad (12)$$

200 *2.1.5. Hypocentre location*

201 We test three assumptions for hypocentre locations, enabling us to quantify the sensitivity of
 202 rupture directivity and ground-motion to the assumed nucleation position.

203 In the first case, the nucleation point is fixed at the observed hypocentre of the Illapel
 204 earthquake, located at latitude -31.5952° , longitude -71.6728° and a depth of 29.0 km. This case
 205 has the limitation that it cannot be used in future earthquakes. The second case follows the
 206 recommendation by Mai et al. (2005), locating the hypocentre at latitude -31.847° , longitude -

207 72.4994 and depth of 3.215 km. The hypocentre is located within a high-slip region but not at the
 208 point of maximum slip,

$$209 \quad \frac{1}{3}D_{\max} < D_i < \frac{2}{3}D_{\max}, \quad (13)$$

210 where D_{\max} is the maximum slip of the source model and D_i is the slip of each subfault.

211 The third case is based on Melgar and Hayes (2019). Here, the hypocentre is selected randomly
 212 using an exponential PDF constrained by observations, located at latitude -31.6343°, longitude -
 213 71.2357° and a depth of 31.865 km.

214 *2.1.6. From Low-Resolution to High-Resolution Rupture Modelling*

215 The finite-fault model used in this study comes from a low-frequency inversion. Its spatial
 216 resolution limits the ability to capture high-frequency rupture behaviour (Hartzell, 1989). To
 217 address this, we refine the fault-plane discretisation by subdividing each subfault into point
 218 sources, following the approach of Spudich and Archuleta (1987). The minimum wavelength,
 219 λ_{\min} , is defined as follows,

$$220 \quad \lambda_{\min} = \frac{V_{s,\min}}{F_{\max}}, \quad (14)$$

221 where $V_{s,\min}$ is the minimum shear-wave velocity and F_{\max} is the maximum target frequency
 222 considered in the simulation. The minimum point source spacing, $d_{e,\min}$, which controls the total
 223 number of point sources across the fault plane, is given by,

$$224 \quad d_{e,\min} = \frac{\lambda_{\min}}{5}. \quad (15)$$

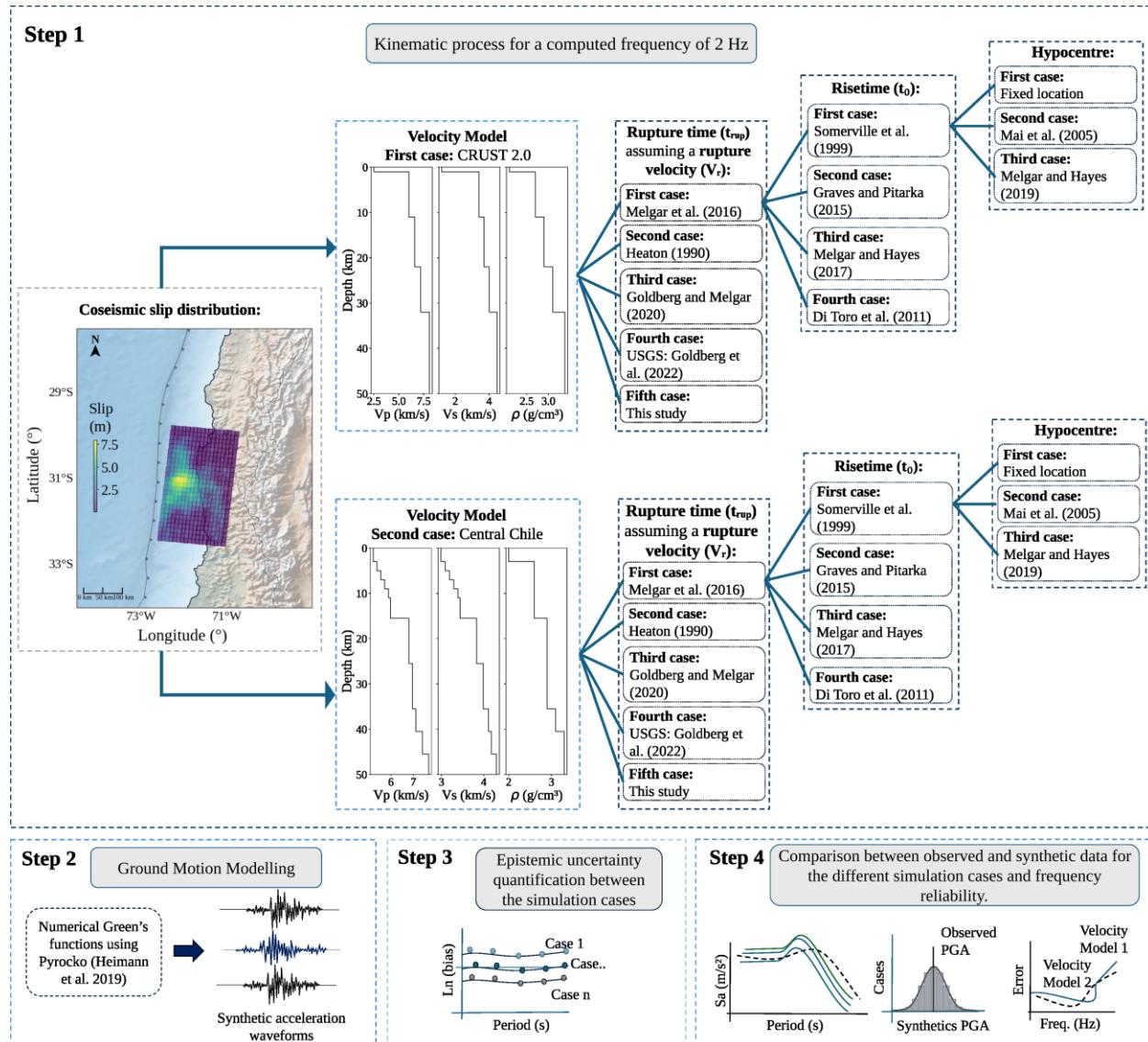
225 Once the grid is refined, the M_0 is interpolated across the rupture area. The t_{rup} at each point
 226 source depends on its distance from the hypocentre, ensuring a consistent spatiotemporal
 227 evolution of the rupture propagation at higher resolution.

228 *2.2. Ground-Motion simulations*

229 We simulate ground-motion at eight near-field stations of the Chilean Seismological Service
230 (CSN: Figure 3), computing numerical Green's functions for the two 1D velocity models,
231 CRUST 2.0 and Central Chile, using the Fomosto module of Pyrocko (Heimann et al., 2019;
232 Figure 2: Step 2). The sampling rate is twice the maximum target frequency, which is 2 Hz,
233 consistent with the discretisation of the subfaults, which avoids aliasing effects and ensures
234 numerical stability and accurate interpolation during the convolution of the sources.
235 Additionally, to represent the local slip-time history, we adopt a triangular source time function.
236 This choice is consistent with the near-triangular moment rate evolution observed in large
237 subduction earthquakes (Meier et al., 2017), supporting a physical description of the rupture
238 process.
239 Once the ground-motions are computed we calculate the geometric-mean of observed and
240 synthetic response spectral accelerations within a 0.1–1 Hz bandpass-filtered frequency range.
241 We then quantify epistemic uncertainty (Figure 2: Step 3) across simulation cases by evaluating
242 the mean $\ln(bias)$, defined as mean of $\ln(obs/syn)$. To characterize the error, we compute the
243 SD of residuals and the *RMSE* defined as,

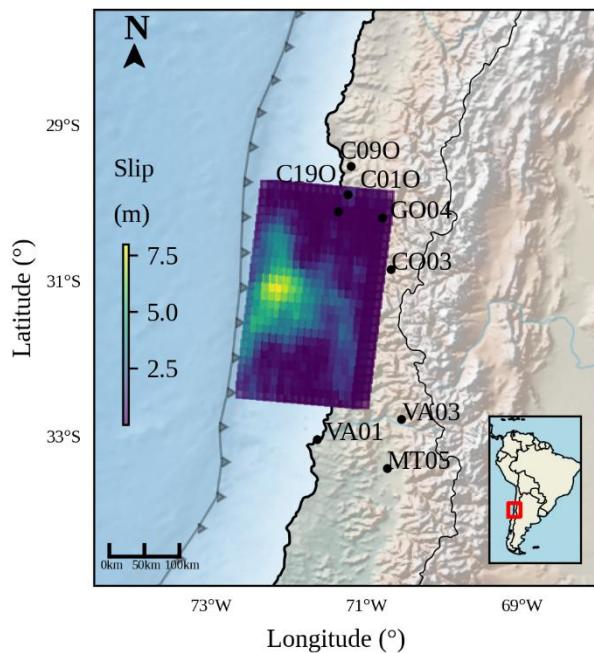
$$244 \quad RMSE = \sqrt{\frac{1}{n} \sum_{i=1}^n (obs_i - syn_i)^2}. \quad (16)$$

245 After evaluating the uncertainties, we compare the synthetic and observed data to access how
246 well the simulation reproduces the ground motions of a well-recorded subduction earthquake and
247 to identify up to which frequency our simulations are reliable (Figure 2, Step 4). The results of
248 the workflow proposed in Figure 2 are presented in the section 3.



250 Figure 2. Workflow to evaluate epistemic uncertainty in kinematic rupture parameters using slip
 251 distributions from subduction earthquakes. Step 1: logic tree, composed of 24 branches,
 252 evaluating different rupture parameter cases. The input of this analysis is the coseismic slip
 253 distribution from the USGS inversion of the 2015 Illapel Mw 8.3 earthquake. Two 1D velocity
 254 models are tested, each simulating a kinematic rupture process up to 2 Hz, while varying rupture
 255 velocity (V_r), to constrain rupture time (t_{rup}), rise time(t_0), and hypocentre location based on
 256 different studies and this work; Step 2: ground-motion modelling using numerical Green's
 257 functions to generate synthetic acceleration waveforms; Step 3: quantification of epistemic

258 uncertainty between the simulation cases, evaluating bias and error. Step 4: comparison between
 259 the geometric-mean of observed and synthetic response spectral accelerations, in a bandpass
 260 filtered frequency range of 0.1 to 1 Hz, across the eight near-field stations, followed by a reliable
 261 frequency range assessment.



262
 263 Figure 3. Spatial distribution of near-field strong-motion stations from the Chilean Seismological
 264 Service (CSN) used in this study, together with the coseismic slip distribution of the 2015 Illapel
 265 Mw 8.3 earthquake.

266 3. Results and Discussion

267 3.1. Sensitivity Analysis of Kinematic Rupture Parameter Cases

268 We evaluate all rupture parameter cases included in the logic tree (Figure 2, Step 1) and detailed
 269 in the methodology section. Figure 4 summarises the ensemble of the 24 logic-tree branches (V_r ,
 270 t_0 , and hypocentre location variations) and how they influence the rupture propagation across the
 271 fault plane. The full branches define the total epistemic uncertainty explored in this study.
 272 Overall, Figure 4 shows considerable variability among the branches, both in the time ranges for

273 the t_{rup} and t_0 , as well as in the propagation of rupture across the different formulations used.

274 However, across all parameter variations, the two 1D velocity models, CRUST 2.0 (Figures 4a-

275 c) and Central Chile (Figures 4d-f), produce similar rupture-propagation behaviour, suggesting

276 that the choice of the velocity model has only a minor influence on the kinematic rupture results.

277 Examining the contribution of each branch of the logic tree, Figures 4a and 4d illustrate the

278 impact of the V_r formulation on rupture evolution. Models proposed by Melgar et al. (2016) and

279 Heaton (1980) yield similarly fast rupture propagation, with a t_{rup} of about 80 s. This similarity

280 arises because, although Melgar et al. (2016) prescribe a depth-dependent V_r following the

281 recommendations of Graves and Pitarka (2010, 2015), both parameterisations effectively

282 converge to a nearly constant $V_r = 0.80 V_s$. From a practical perspective, choosing between these

283 two formulations may be of minor importance for kinematic source modelling. The formulation

284 by Goldberg and Melgar (2020) is also close to Melgar et al. (2016) and Heaton (1980), but it

285 tends to produce slightly slower rupture propagation. By contrast, the USGS formulation and the

286 model introduced in this study prescribe overall lower V_r , leading to systematically longer

287 rupture durations across the fault plane. All these formulations have been tested in subduction

288 earthquake simulations, yet our results show maximum t_{rup} ranging from ~80 s to more than 150

289 s. This variability indicates that V_r is still not well understood and as noted in the Introduction,

290 requires additional physical observations to be better constrained.

291 Figures 4b and 4e show similar spatial patterns in t_0 distributions, which are controlled by the

292 slip, reflecting the common assumption in the four formulations and consistent with the physics

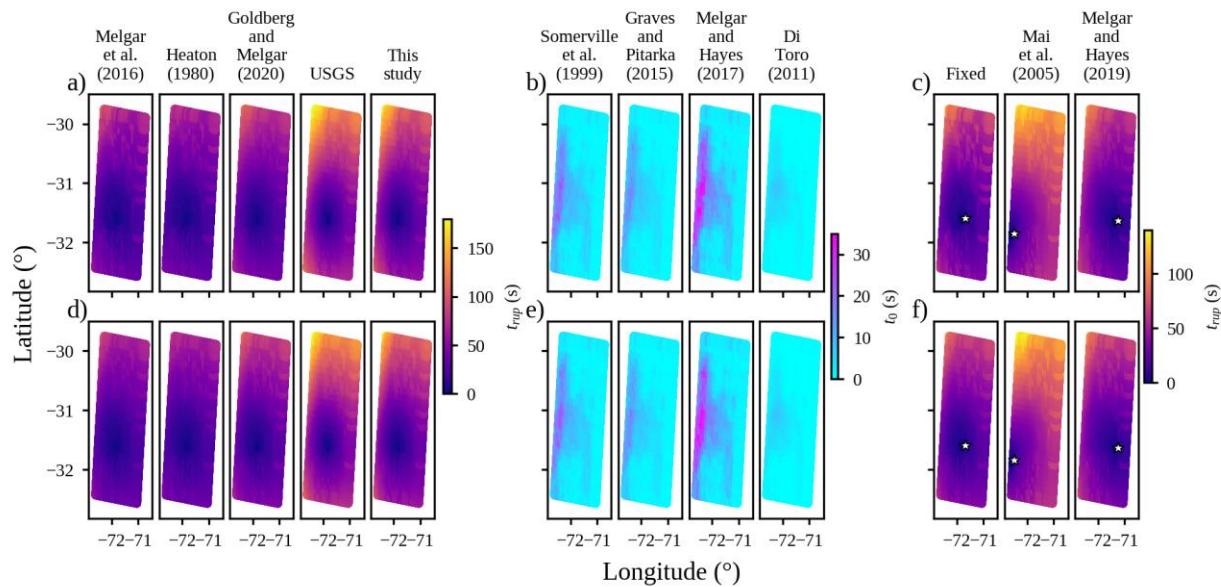
293 of the rupture. Despite using the same fast V_r model (Melgar et al., 2016), the Somerville et al.

294 (1999), Graves and Pitarka (2015), and Melgar and Hayes (2017) formulations produce

295 considerable variability in the absolute t_0 values, showing a stronger sensitivity to the definition

296 of t_a . However, t_a is dependent on the M_0 , this dependence is implemented differently in each
297 formulation: in Somerville et al. (1999), it is related to the distance from the hypocentre to the
298 nearest asperity, in Graves and Pitarka (2015) involves the focal mechanism, and in Melgar and
299 Hayes (2017) it is linked to the mean rise time derived from kinematic inversions of multiple
300 inverse large earthquakes. In contrast, the Di Toro et al. (2011) distribution yields t_0 values that
301 are approximately proportional to slip, because it implicitly assumes a fast \dot{S} on the order of ~ 1
302 m/s. Interestingly, its spatial pattern remains very similar to the other formulations. Overall, such
303 differences in how each formulation parameterizes t_a and the \dot{S} drive the variability in the
304 resulting t_0 distributions.

305 We show the influence of hypocentre location cases in Figures 4c and 4f. The approach of Mai et
306 al. (2005) produces a slower rupture, which is expected to result in a weaker directivity. In
307 contrast, the fixed case, corresponding to the observed hypocentre of the Illapel earthquake and
308 the Melgar and Hayes (2019) case generate a similar fast rupture, despite their different
309 nucleation points. This result confirms that rupture propagation and directivity effects are
310 sensitive to the assumed nucleation point. In both the fixed case, which corresponds to the
311 observed mainshock location, and the Melgar and Hayes (2019) case, the hypocentre is located
312 within a high-slip region but not at the maximum-slip, consistent with the Mai et al. (2005)
313 assumption.

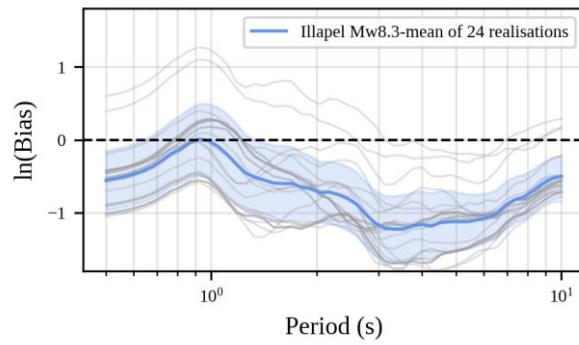


314

315 Figure 4. Kinematic rupture models for the Mw 8.3 Illapel earthquake computed at 2.0 Hz: a), d)
 316 rupture time (t_{rup}) distributions for all rupture velocity (V_r) formulations; b), e) rise time (t_0)
 317 distributions; c), f) t_{rup} distributions across all tested hypocentre locations; a-c) use the CRUST
 318 2.0 1D velocity model (Bassin et al., 2000) and d-f) use the Central Chile 1D velocity model
 319 (Caballero et al., 2023).

320 Using the kinematic rupture models forementioned, we compute synthetic acceleration
 321 waveforms using numerical Green's functions across the eight near-field stations. For all
 322 stations, we calculate the geometric-mean response spectra of the synthetic and observed
 323 horizontal acceleration components in a bandpass filter frequency range of 0.1 to 1 Hz. Using
 324 these response spectra, we compute the mean $\ln(bias)$, to capture the epistemic uncertainty
 325 associated with each branch of the logic tree and quantify whether our simulations tend to
 326 overestimate or underestimate the observed ground-motion. We then summarise the logic-tree
 327 ensemble by computing the ensemble mean $\ln(bias)$ and the ± 1 SD across the 24 realisations,
 328 which captures the variability associated with all rupture parameters and for both 1D velocity
 329 models (Figure 5).

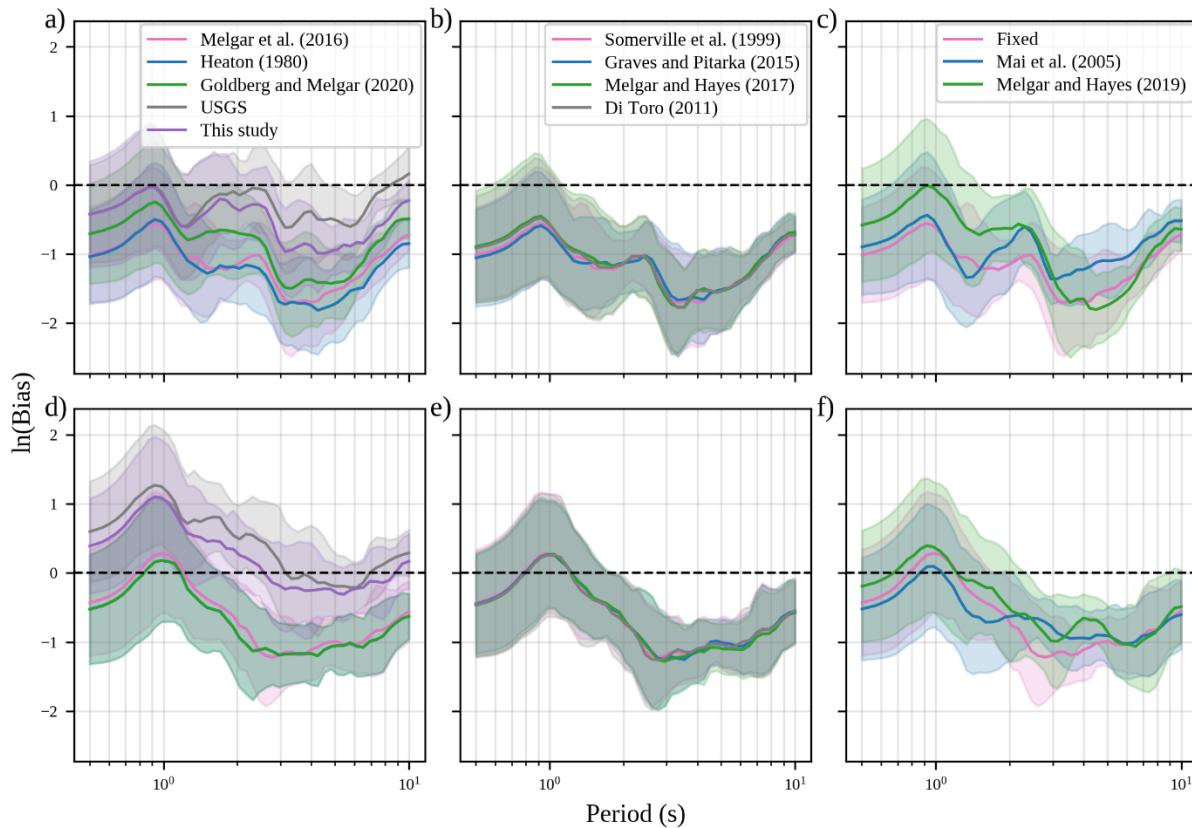
330 In the period range analysed ($\sim 0.5\text{--}10$ s), the ensemble mean shows relatively small bias at short
 331 periods, particularly for periods shorter than ~ 2 s, where $\ln(\text{bias})$ remains close to zero. At
 332 intermediate periods, the ensemble mean becomes increasingly negative, with most realisations
 333 exhibiting a systematic negative $\ln(\text{bias})$ between ~ 4 and 7 s, indicating an overestimation of the
 334 observed spectral amplitudes in this band. Two cases show a pronounced deviation at periods
 335 shorter than ~ 3 s, indicating an underestimation. Overall, the observations are covered by the
 336 range of simulations, as the dispersion of $\ln(\text{bias})$ across the 24 realisations remains distributed
 337 around zero over most of the analysed period range.



338

339 Figure 5. Bias in the geometric-mean horizontal response spectra, expressed as $\ln(\text{bias})$ residuals
 340 between synthetics and observations for the 24 realisations in the logic-tree ensemble in a band-
 341 pass filter range of 0.1–1.0 Hz. Thin grey lines show, for each realisation, the mean $\ln(\text{bias})$
 342 averaged over the eight stations, the thick line shows the ensemble mean, and the shaded
 343 envelope indicates ± 1 Standard Deviation (SD) across the realisations.
 344 Figure 6 shows period-dependent model performance for each logic-tree branch, summarized by
 345 $\ln(\text{bias})$ and the mean residual SD. To isolate the influence of V_r , we vary only this parameter
 346 while keeping the Somerville et al. (1999) t_0 model and a fixed hypocentre location (real case for
 347 Illapel earthquake), as is shown in the logic tree of Figure 2: Step 1. The V_r formulations produce
 348 the largest variability (Figures 6a and 6d), particularly when using the CRUST 2.0 model (Figure

349 6a). The variation is significantly larger for this parameter than for t_0 or hypocentre location.
350 Additionally, we find that the mean $\ln(\text{bias})$ exhibits stronger period-dependent variability.
351 For the t_0 variations, we fix the V_r model of Melgar et al. (2016) and a fixed hypocentre (Figure
352 2, Step 1). In contrast to V_r , the t_0 cases show remarkably low variations (Figures 6b and 6e).
353 The SD of residuals remains almost constant across periods, and the mean log-bias curves cluster
354 tightly around zero for the Central Chile velocity model (Figure 6e).
355 For the hypocentre location cases, we use the Melgar et al. (2016) V_r model and the Somerville
356 et al. (1999) t_0 model (Figure 2, Step 1). The hypocentre formulations (Figures 6c and 6f) show
357 an intermediate sensitivity: the variability is larger than for t_0 but smaller than for V_r . As with the
358 other parameters, the CRUST 2.0 model yields broader scatter than the Central Chile model.
359 Even so, the mean $\ln(\text{bias})$ curves remain relatively consistent across the different hypocentre
360 assumptions, supporting the suitability of the approaches by Mai et al. (2005) and Melgar and
361 Hayes (2019) for future scenarios.
362 Across all parameter variations, the use of the Central Chile velocity model produces slightly
363 lower epistemic uncertainty, the SD of residuals is marginally smaller than in the CRUST 2.0
364 model, and the mean $\ln(\text{bias})$ curves lie closer to zero, particularly in the t_0 and hypocentre cases
365 (Figures 6e and 6f). The improvement is modest but indicates that the regional structure yields
366 more stable and less biased predictions.



367

368 Figure 6. Bias between synthetics and observed geometric-mean horizontal response spectra.

369 Curves show the mean log-bias, and shaded areas are ± 1 Standard Deviation (SD) of residuals.

370 Seismograms are band-pass filtered in the 0.1-1.0 Hz range: a-c) results using the CRUST 2.0

371 velocity model (Bassin et al., 2000) for rupture velocity (V_r), rise time (t_0) and hypocentre

372 location cases, respectively; d-f) corresponding results using the Central Chile velocity model

373 (Caballero et al., 2023). a), d) only V_r is varied while t_0 is fixed to Somerville et al. (1999) and

374 the hypocentre location is fixed to the observed Illapel earthquake; b), e) only t_0 is varied while

375 V_r is fixed to Melgar et al. (2016) and the hypocentre remains fixed; c), f), only the hypocentre

376 location is varied while V_r is fixed to Melgar et al. (2016) and t_0 to Somerville et al. (1999).

377 To quantify the sensitivity of each parameter, we compute three metrics (Table 1): mean $\ln(\text{bias})$,

378 Bias Factor, and RMSE. The mean $\ln(\text{bias})$ describes whether the simulations systematically

379 overestimate (negative values), underestimate (positive values) or unbiased (near zero values) the

380 observed response spectra. The Bias Factor provides similar information, but values close to one
381 represent minimal bias. The RMSE quantifies the overall accuracy of the synthetic spectra.
382 For the CRUST 2.0 velocity model, we observe the largest variability across the V_r formulations,
383 consistent with the strong epistemic uncertainty shown in Figure 6. The USGS case provides the
384 best central tendency, with a mean $\ln(\text{bias})$ of -0.238 and a Bias Factor of 0.788 , both slightly
385 overestimating the observed response spectra. However, the RMSE of the simulations associated
386 to this velocity model is the lowest with 0.29769 , indicating good accuracy.
387 In contrast, the use of the Central Chile velocity model significantly reduces this variability.
388 Across all V_r formulations, our preferred case performs best, with a mean $\ln(\text{bias})$ of 0.088 and a
389 Bias Factor of 1.092 , both of which are very close to unity, while also producing the lowest
390 RMSE of this velocity model with 0.35623 , which is slightly higher than the USGS case in the
391 CRUST2.0 model. Heaton (1980), the commonly used V_r assumption, yields essentially the
392 same metrics as Goldberg and Melgar (2020). This indicates that, when the Central Chile
393 velocity model is used, the variability associated with choosing between these two formulations
394 is limited.
395 The t_0 cases behave similarly across all formulations, matching the low variability seen in Figure
396 6. This result further confirms that the t_0 formulations is not a major source of uncertainty in the
397 ground-motion simulations, particularly when using a regional velocity model.
398 The hypocentre location cases show intermediate sensitivity, as shown in Figure 6. Under
399 CRUST 2.0, the Melgar and Hayes (2019) case produces the lowest RMSE (0.41957), while the
400 Mai et al. (2005) case shows the strongest bias (-0.85705). However, the three cases notably
401 perform similar. Under the Central Chile model, all three formulations are comparable too, with
402 narrow $\ln(\text{bias})$ ranges (0.69 to 0.55), Bias Factors around 0.5 , and only minor variations in

403 RMSE between the Fixed and Melgar and Hayes (2019) cases, principally. This indicates that
 404 hypocentre assumptions become less influential when using a regional velocity model. Although
 405 we evaluate the fixed case, it does not inform decisions for future events because it corresponds
 406 to the true hypocentre of the studied earthquake. Instead, it serves as a reference to determine
 407 which formulation best reproduces the observed ground-motion.

408 Table 1. Bias and error metrics comparison for different rupture velocity (V_r), rise time(t_0) and
 409 hypocentre cases.

| Case | Kinematic parameter | Velocity model | Mean ln(bias) | Bias Factor | RMSE |
|------------------------------|---------------------|----------------|---------------|-------------|---------|
| Melgar et al. (2016) | V_r | CRUST 2.0 | -1.11373 | 0.32833 | 0.52971 |
| Heaton (1980) | V_r | CRUST 2.0 | -1.20819 | 0.29874 | 0.54028 |
| Goldberg and Melgar (2020) | V_r | CRUST 2.0 | -0.85255 | 0.42633 | 0.37547 |
| USGS: Goldberg et al. (2022) | V_r | CRUST 2.0 | -0.23801 | 0.78819 | 0.29769 |
| This study | V_r | CRUST 2.0 | -0.56406 | 0.56890 | 0.31830 |
| Melgar et al. (2016) | V_r | Central Chile | -0.69149 | 0.50083 | 0.39988 |
| Heaton (1980) | V_r | Central Chile | -0.76444 | 0.46560 | 0.39725 |
| Goldberg and Melgar (2020) | V_r | Central Chile | -0.76444 | 0.46560 | 0.39725 |
| USGS: Goldberg et al. (2022) | V_r | Central Chile | 0.28138 | 1.32496 | 0.36655 |
| This study | V_r | Central Chile | 0.08815 | 1.09215 | 0.35623 |
| Sommerville et al. (1999) | t_0 | CRUST 2.0 | -1.11373 | 0.32833 | 0.52971 |
| Graves and Pitarka (2015) | t_0 | CRUST 2.0 | -1.11526 | 0.32783 | 0.50624 |
| Melgar and Hayes (2017) | t_0 | CRUST 2.0 | -1.06476 | 0.34481 | 0.48680 |
| Di Toro (2011) | t_0 | CRUST 2.0 | -1.08481 | 0.33797 | 0.49410 |
| Sommerville et al. (1999) | t_0 | Central Chile | -0.69149 | 0.50083 | 0.39988 |
| Graves and Pitarka (2015) | t_0 | Central Chile | -0.70483 | 0.49419 | 0.39637 |
| Melgar and Hayes (2017) | t_0 | Central Chile | -0.72149 | 0.48603 | 0.39791 |
| Di Toro (2011) | t_0 | Central Chile | -0.70062 | 0.49628 | 0.40925 |
| Fixed | Hypocentre | CRUST 2.0 | -1.11373 | 0.32833 | 0.52971 |
| Mai et al. (2005) | Hypocentre | CRUST 2.0 | -0.85705 | 0.42441 | 0.50832 |
| Melgar and Hayes (2019) | Hypocentre | CRUST 2.0 | -0.92372 | 0.39704 | 0.41957 |
| Fixed | Hypocentre | Central Chile | -0.69149 | 0.50083 | 0.39988 |

| | | | | | | |
|-----|--------------------------------|------------|---------------|----------|---------|---------|
| 410 | Mai et al. (2005) | Hypocentre | Central Chile | -0.67611 | 0.50859 | 0.52971 |
| 411 | Melgar and Hayes (2019) | Hypocentre | Central Chile | -0.55218 | 0.57569 | 0.37298 |

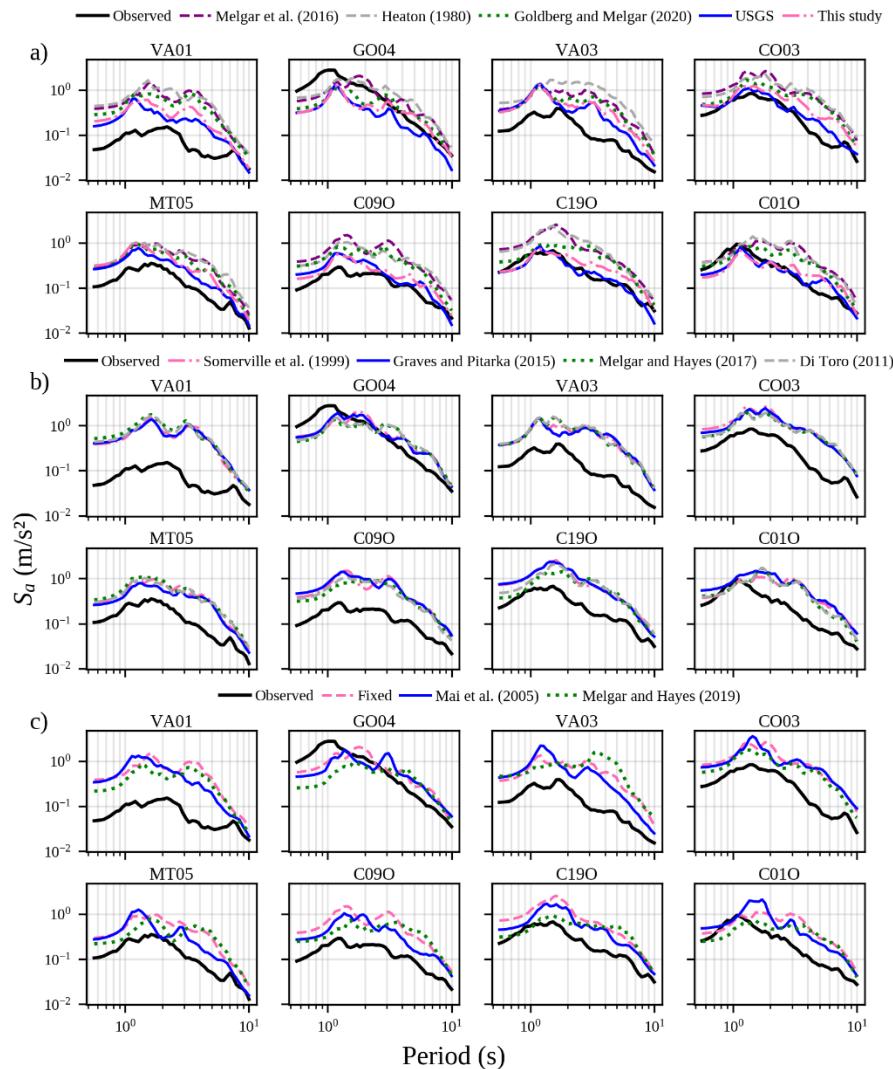
412 *3.2. Comparison between observed and synthetic data*

413 After understanding the epistemic uncertainties involved in the simulation cases, we compare the
 414 geometric-mean response spectra of the synthetic and observed horizontal acceleration
 415 components. This comparison enables us to assess how each rupture formulation affects the
 416 ground-motion across the eight near-field stations. Figures 7 and 8 show such comparisons for
 417 each rupture parameter case and for both 1D velocity models.

418 Figures 7a and 8a show the greatest variability, confirming the high sensitivity and uncertainty
 419 associated with V_r . As seen in Figures 4a, 4d and Table 1, Melgar et al. (2016) and Heaton
 420 (1980) produce nearly identical rupture propagation. This similarity is also reflected in the
 421 response spectra, which closely align, especially when using the Central Chile velocity model
 422 (Figure 8a). In contrast, the other V_r formulations cover a broader amplitude range. The USGS
 423 model and the formulation introduced in this study match the observations more closely in both
 424 velocity models, suggesting that slower V_r values may provide a better representation of the
 425 rupture for this event.

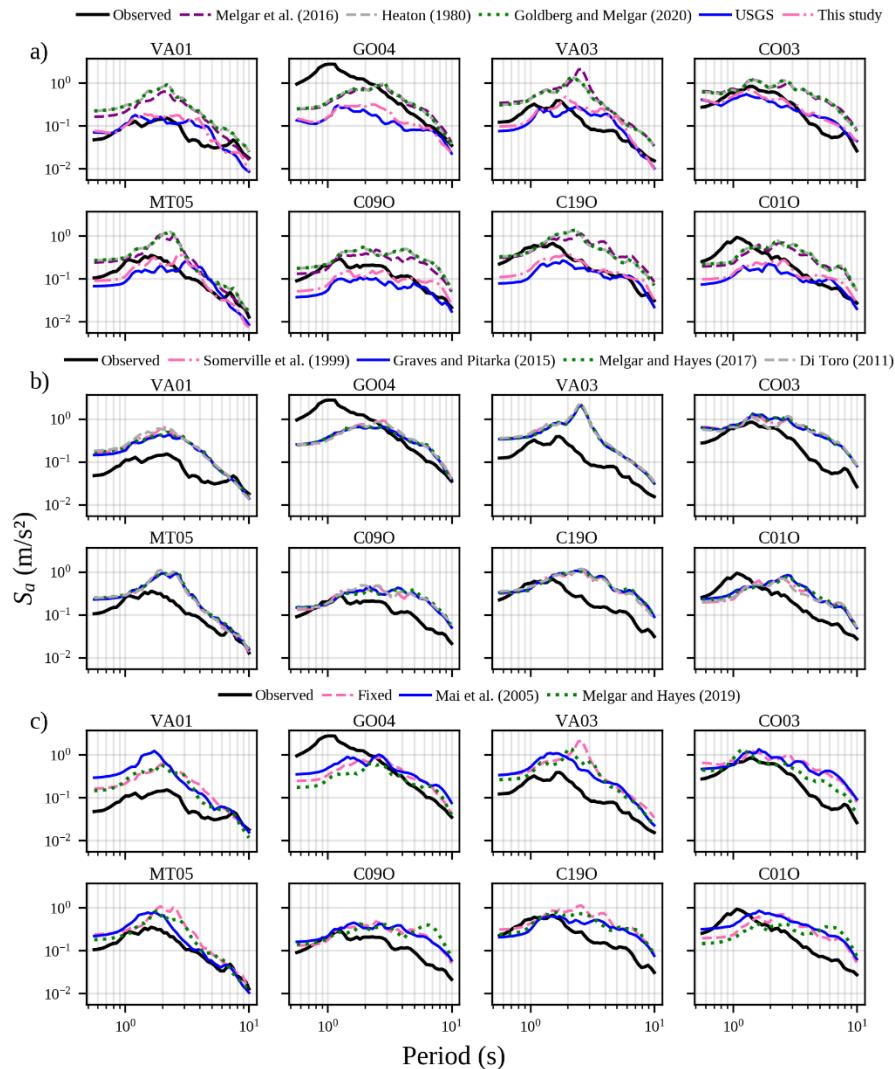
426 Figures 7b and 8b show the t_0 variations where the synthetic spectra do not match the
 427 observations, consistent with the strong influence of the Melgar et al. (2016) model, which is
 428 fixed in these variations. However, the spectra from the different t_0 formulations remain tightly
 429 clustered. This behaviour is unexpected given the large differences in the t_0 distributions (Figures
 430 4b and 4e) and indicates that t_0 exerts only a minor influence on the predicted ground-motion.
 431 Figures 7c and 8c show some variability in the hypocentre location cases, although the
 432 differences are smaller than those in the V_r cases. Moreover, the response spectra obtained from

433 the fixed hypocentre are quite similar to those of the Mai et al. (2005) and Melgar and Hayes
 434 (2019), despite being in clearly different positions and producing different directivity effects.
 435 These results are consistent with the uncertainty metrics in Table 1, which suggest that both
 436 heuristic approaches provide reasonable nucleation points for future scenario modelling.



437

438 Figure 7. Comparison between observed and synthetic geometric-mean response spectral
 439 accelerations for every case: a) rupture velocity (V_r); b) rise time (t_0) and c) hypocentre location
 440 cases using the CRUST 2.0 1D velocity model (Bassin et al., 2000). Results are shown for a
 441 band-pass filtered frequency range of 0.1–1.0 Hz across the eight near-field stations.



442

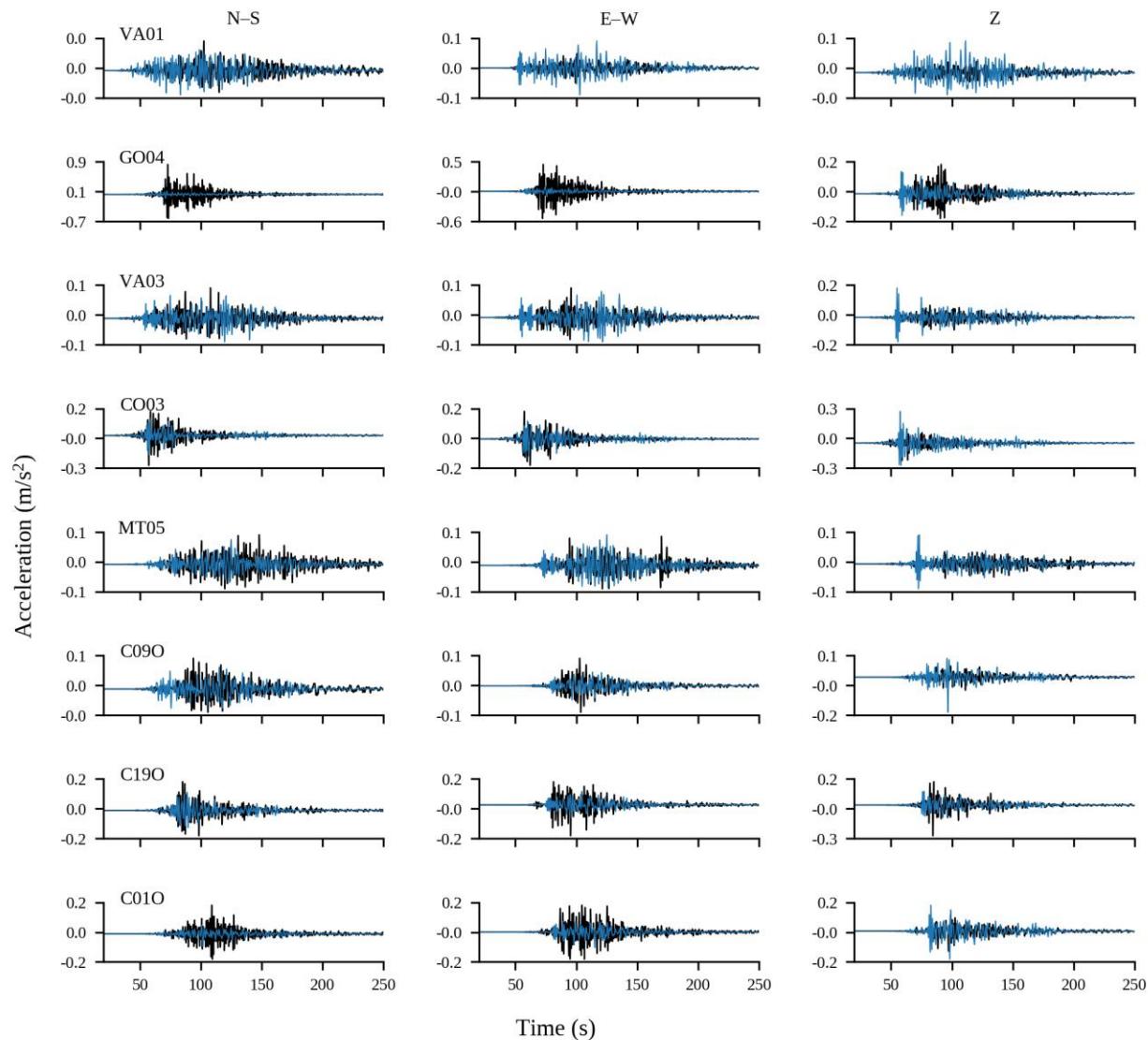
443 Figure 8. Comparison between observed and synthetic geometric-mean response spectral
 444 accelerations for every case: a) rupture velocity (V_r); b) rise time (t_0) and c) hypocentre location
 445 using the Central Chile 1D velocity model (Caballero et al., 2023). Results are shown for a band-
 446 pass filtered frequency range of 0.1–1.0 Hz across the eight near-field stations.
 447 Additionally, we compare the observed acceleration waveforms with the best-fitting logic-tree
 448 branch synthetics. This branch has the lowest bias (mean $\ln(bias)$ closest to zero) and a low
 449 RMSE. In particular, we use the 1-D velocity model of Central Chile, the V_r formulation

450 proposed in this study, the Somerville et al. (1999) as the to model, and the observed hypocentre
451 location of the Illapel earthquake, corresponding to the fixed case.

452 Figure 9 shows the comparison of observed and synthetic acceleration waveforms in a bandpass
453 filter frequency range of 0.1 to 1 Hz for the eight near-field stations used in this analysis. The
454 comparison provides a direct measure of how closely the selected branch reproduces the
455 observed waveform amplitudes and durations across the three components.

456 Overall, the simulations closely reproduce the main features of the recordings, including the
457 waveform similarity, spectral content and amplitude. The best agreement is obtained at the rock
458 stations VA01, VA03 and C19O. On the vertical components at VA01 and VA03, the synthetics
459 tend to exceed the observations, with an overestimation associated with the main pulses. In
460 contrast, GO04 and C01O show a clear underestimation, most notably at GO04 and mainly on
461 the horizontal components. This mismatch may reflect local site effects that are not captured by
462 the 1D velocity structure adopted here, and it could also indicate more complex (potentially
463 nonlinear) site behaviour at GO04. However, MT05, which is also located in soil, shows the
464 opposite tendency; the horizontal components are captured well by the simulations, while the
465 vertical component exhibits a moderate overestimation of the main pulse.

466 Although we rely on a single event to quantify uncertainty, this comparison suggests that the
467 selected best-fitting branch captures the characteristics of the near-field ground-motion time
468 histories for a large megathrust earthquake, such as the Illapel event.



470 Figure 9. Comparison of observed (black) and synthetic (blue) waveforms for the
 471 Mw 8.3 Illapel earthquake at eight near-field stations. Waveforms are band-pass filtered between
 472 0.1 to 1 Hz. Synthetics correspond to the best-fitting logic-tree branch using the Central Chile 1D
 473 velocity model, the V_r formulation proposed in this study, the t_0 according by Somerville et al.
 474 (1999), and the observed (fixed) hypocentre location.

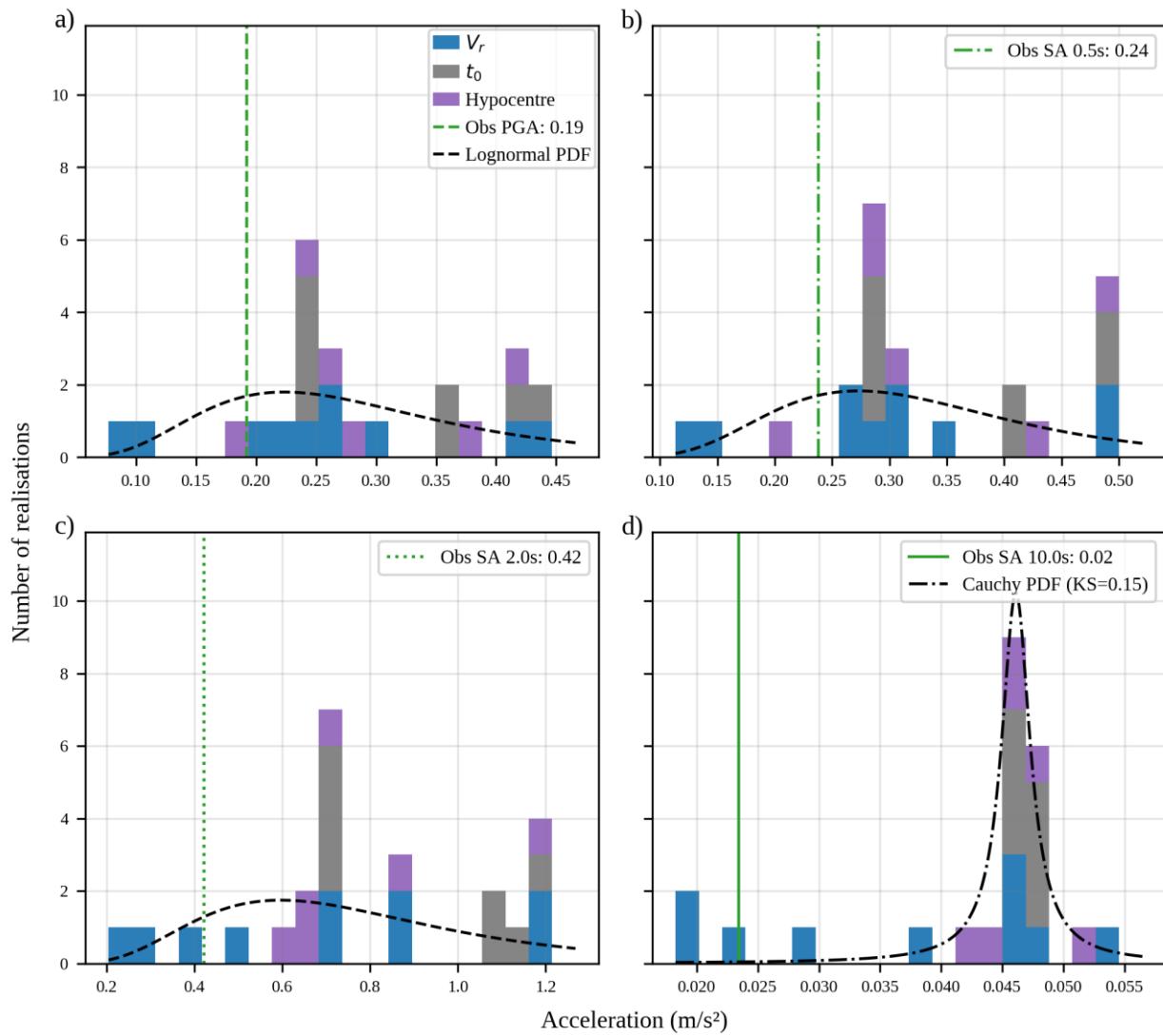
475 In addition to these comparisons, we then test whether a single event, such as the 2015 Illapel
 476 earthquake, can capture the variability needed to constrain future scenarios. Figure 10 shows the
 477 mean synthetic and observed PGA and SA at 0.5, 2.0, and 10.0 s across all parameter variations.

478 PGA and SA at 0.5 and 2.0 s follow lognormal distributions (Figures 10a-c), while SA at 10 s
479 exhibits much heavier tails and is better described by a Cauchy distribution (Figure 10d).
480 The observed PGA (0.19 m/s^2) and SA at 0.5 s (0.24 m/s^2) fall near the central tendency of the
481 synthetic realisations, indicating that one well-recorded event can capture short-period
482 variability. However, the observed SA at 2.0 s (0.42 m/s^2) lies somewhat off the distribution
483 mean (Figure 10c), and the observed SA at 10 s (0.02 m/s^2) falls in the extreme tail of the fitted
484 Cauchy PDF (Figure 10d). These results suggest that a single event is insufficient to represent
485 the spectral acceleration in long-period beyond about 2 s, which limits its use for future long-
486 period scenario development.
487 Extending this analysis to multiple subduction earthquakes would broaden the range of rupture
488 behaviours and help to reduce the epistemic uncertainties. Because this study focuses on
489 uncertainties from the kinematic rupture process, future work should also assess the epistemic
490 variability introduced by stochastic slip distribution generation schemes.

491 *3.3. Evaluation of Frequency Reliability in Ground-Motion Modelling*

492 We complete this study by evaluating the reliability of the computed frequency content, an
493 aspect often overlooked in rupture model generators. We calculate the mean SD of residuals and
494 the bias factor for the lowest-uncertainty formulation within each velocity model (Figure 11). For
495 CRUST 2.0, we adopt the V_r from the USGS model, the t_0 from Melgar and Hayes (2017)
496 formulation and a fixed hypocentre location. For the Central Chile model, we use the V_r inferred
497 in this study, the t_0 from Somerville et al. (1999) and the same fixed hypocentre.
498 Because source resolution depends on the lowest V_s and the target frequency computed, the
499 CRUST 2.0 model supports frequencies for the studied earthquake up to 5 Hz, with a
500 corresponding spatial resolution of 48 m. In contrast, the Central Chile model reaches 10 Hz with

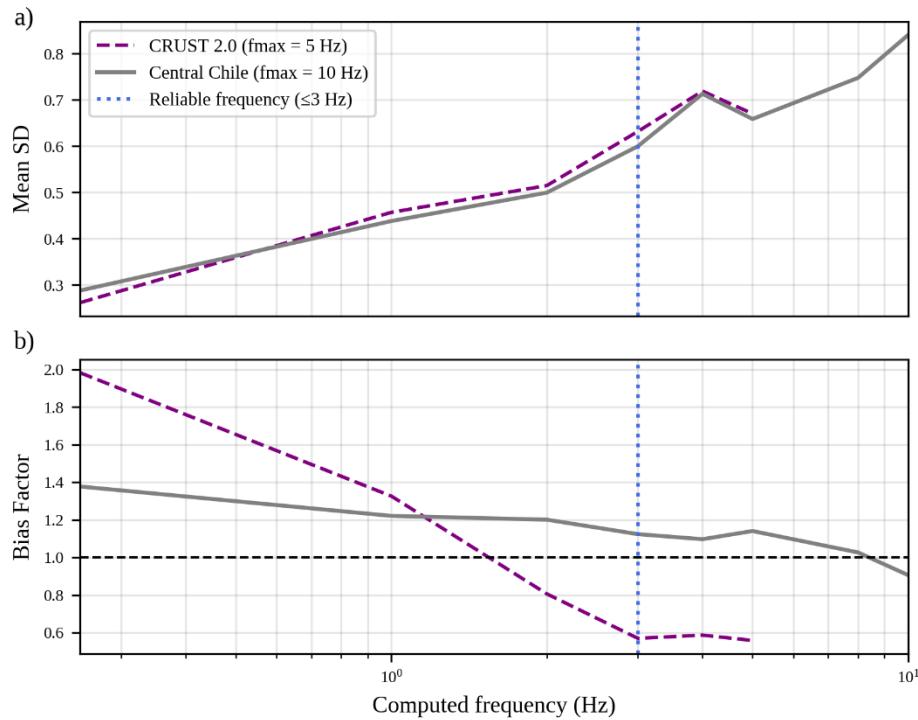
501 a resolution of 60 m. Across the computed frequency range, the mean SD of residuals increases
 502 for both models. The bias factor approaches one, indicating minimal bias, near 2 Hz for CRUST
 503 2.0 and remains stable around 3 Hz for Central Chile before decaying. The behaviour of the SD
 504 of residuals across the computed frequencies aligns with Pilz et al. (2025), confirming that 1D
 505 models remain reliable below 3 Hz, which represents the site-source conditions for this
 506 earthquake.



507

508 Figure 10. Histograms of mean horizontal geometric-mean synthetic, for the 24 realisations of
 509 the logic-tree framework, and observed acceleration values: a) distribution of simulated peak

510 ground acceleration (PGA); b) spectral accelerations (SA) at 0.5 s, c) 2.0 s, and d) 10.0 s. The
 511 bars show the results grouped according to the defined cases of variation in rupture velocity
 512 (V_r), rise time (t_0), and hypocentre location. Vertical green lines correspond to the observed
 513 mean values for each period. The black dashed and dash-dotted lines represent the best-fitting
 514 lognormal and Cauchy probability density functions (PDFs), respectively.



515
 516 Figure 11. Uncertainty quantification of the computed frequency response for the two 1D
 517 velocity models analysed. For CRUST 2.0, we use a rupture velocity (V_r), from USGS (Goldberg
 518 et al., 2022), a rise time (t_0) from Melgar and Hayes (2017), and a fixed hypocentre location. For
 519 the Central Chile model, we use the (V_r), inferred in this study, a t_0 from Somerville et al. (1999),
 520 and a fixed hypocentre: a) mean Standard Deviation (SD) of residuals and b) Bias Factor. The
 521 green dotted line marks the reliable frequency range (≤ 3 Hz), which effectively captures the site
 522 response according to Pilz et al. (2025).

523 **4. Conclusions**

524 We quantify epistemic uncertainties in ground-motion simulations derived from kinematic
525 rupture models based on slip distributions of large subduction earthquakes. Using the 2015
526 Illapel Mw 8.3 event as a case study, we test multiple formulations of rupture velocity (V_r), to
527 constrain rupture time (t_{rup}), rise time (t_0), and hypocentre location across two one-dimensional
528 (1D) velocity models. We evaluate their effects on simulated ground-motions at eight near-field
529 stations by comparing synthetic and observed geometric-mean response spectra. We further
530 assess how well the records of a single event are captured by the range of simulated intensity
531 measures, and finally, we analyse the frequency reliability of the simulations across both 1D
532 velocity models.

533 The main conclusions are:

- 534 • The V_r formulations produce the greatest variability in the ground-motion simulations.
535 The behaviour of the mean $\ln(\text{bias})$, Bias Factor, and mean RMSE confirms that V_r is the
536 dominant source of epistemic uncertainty. The depth-dependent V_r formulation, proposed
537 here, with $0.3V_s$ at depths <10 km and $0.5V_s$ at depths >15 km, minimises the uncertainty
538 metrics for this particular case, especially when combined with the Central Chile velocity
539 model.
- 540 • Despite the strong differences in the spatial distributions of t_0 , the impact on ground-
541 motion remains negligible. All metrics stay stable across formulations, indicating that t_0
542 does not substantially influence the predicted ground-motion.
- 543 • Hypocentre location has a moderate influence on the simulated ground-motion, and this
544 effect becomes much smaller when we use the Central Chile velocity model. This

545 indicates that the influence of hypocentre assumptions depends on the choice of velocity
546 model.

- 547 • Our frequency analysis shows that the simulations are reliable until 3 Hz. This confirms
548 that 1D velocity models are adequate for representing low to intermediate frequencies in
549 ground-motion modelling but not high frequencies, even if the computational resolution
550 makes high-frequency computation feasible.
- 551 • The ground-motions of the well-recorded 2015 Illapel subduction earthquake in Central
552 Chile are better captured by the range of the models at intermediate frequencies (0.5-
553 0.3Hz) rather than at low frequencies (<0.5Hz).
- 554 • Overall, the results show that the regional velocity model (Central Chile) provides the
555 best performance. Further refinement of regional velocity models is therefore a key
556 pathway to reduce epistemic variability in ground-motion simulations
- 557 • Our final goal is not only to evaluate epistemic uncertainty but also to decrease it. The
558 lowest-uncertainty models identified here provide a basis for simulating future synthetic
559 earthquake scenarios to analyse aleatory variability in a way that complements the
560 epistemic uncertainty evaluated in this study. However, because our analysis is based on
561 a single earthquake, additional observations are necessary to more accurately calibrate
562 and represent the rupture velocities for future simulations of large subduction
563 earthquakes. This remains difficult in subduction zones with few well-recorded
564 tsunamigenic earthquakes. By reducing epistemic uncertainty in the kinematic rupture
565 propagation, used to simulate ground-motion, our proposed framework supports a more
566 physics-based seismic source integration for seismic-tsunami hazard assessment in
567 subduction zones.

568 **Data and Resources**

569 The dataset and codes generated for this study are available at:

570 <https://github.com/monsebo/Kinematic-Rupture-From-a-Slip-Distribution-.git>

571 The codes were implemented using Python 3 (Van Rossum and Drake, 2009), Pyrocko and its
572 Fomosto module (Heimann et al., 2019). The slip distribution was used from the USGS inversion
573 in <https://earthquake.usgs.gov/earthquakes/eventpage/us20003k7a/finite-fault> (last accessed on
574 30 November 2025).

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