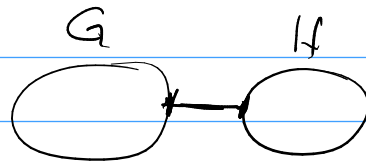


dendrograma

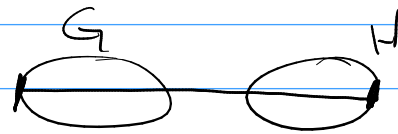
Single linkage

$$d_{SL}(G, H) = \min_{\substack{i \in G \\ i' \in H}} d_{i, i'}$$



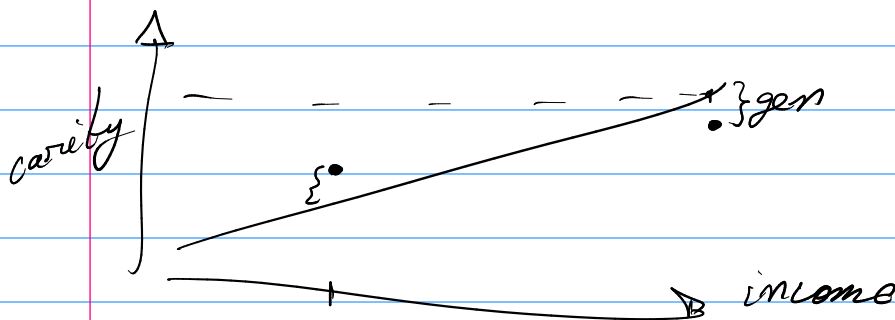
Complete Linkage

$$d_{CL}(G, H) = \max_{\substack{i \in G \\ i' \in H}} d_{i, i'}$$



Group Average

$$d_{GA}(G, H) = \frac{1}{N_G N_H} \sum_{i \in G} \sum_{i' \in H} d_{i, i'}$$



Clasificación

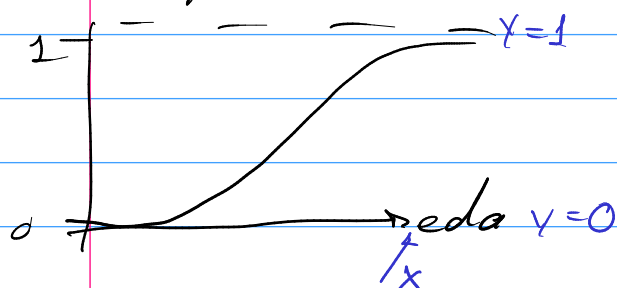
LDA (Linear Discriminant Analysis)

Análisis Discriminante Lineal

Clasificador Óptimo de Bayes:

$$g^* \in \operatorname{argmin}_{g: X \rightarrow Y} P(g(X) \neq Y | X)$$

i.e., que minimiza la probabilidad posterior de equivocarse



Edad \rightarrow Coronarios
 $\leftarrow P$

Supongamos que $Y=k \Rightarrow$

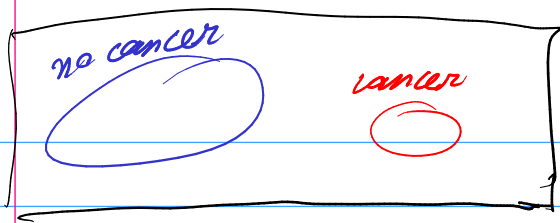
$$g^* \in \operatorname{argmax}_{g: X \rightarrow Y} P(g(x)=k | X=x)$$

teorema de Bayes

$$P(g(x)=k | X=x) = \frac{P(X=x | Y=k) P(Y=k)}{P(X=x)}$$

i.e. $\underbrace{P(g(x)=k | X=x)}_{\text{posterior}} \propto \underbrace{P(X=x | Y=k)}_{\text{verosimilitud}} \underbrace{P(Y=k)}_{\text{previa}} \dots (*)$

De aquí que $P(Y=k | X=x)$ se le llama la proba. posterior y a g^* por eso se le llama clasificador bayesiano óptimo



Regresión logística modela directamente la proba. posterior, en particular si usamos la liga logit:

$$\log \frac{P(Y=k | X=x)}{P(Y=\underline{k} | X=x)} = x^T \vec{\beta}_k$$

$$\log \frac{P(Y=1 | X=x)}{1 - P(Y=1 | X=x)} = x^T \vec{\beta}$$

Pero podríamos modelar el lado derecho

π_k : proba. previa de que $Y \in \text{Pop. } k$

$P(X=x | Y=k)$: Verosimilitud
 $= f_k(x)$

$$\left[\phi_k(x; \mu_k, \Sigma_k) \right]^{\text{mezcla gaussian}}$$

$$P(Y=k | X=x) = \frac{P(X=x | Y=k) P(Y=k)}{P(X=x)}$$

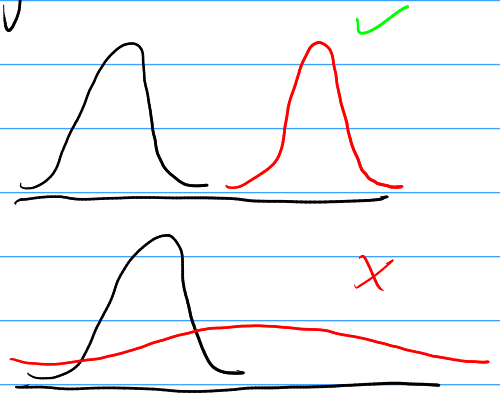
$$= \frac{P(X=x | Y=k) P(Y=k)}{\sum_{l=1}^K P(X=x | Y=l) P(Y=l)}$$

$$= \frac{f_k^*(x) \pi_k}{\sum_{l=1}^K f_l^*(x) \pi_l}$$

also, $\frac{\mathbb{P}(Y=k | X=x)}{\mathbb{P}(Y=j | X=x)} = \frac{f_k^*(x)}{f_j^*(x)} \frac{\pi_k}{\pi_j} \dots (**)$

$$f_i^*(x) = \mathcal{N}_p(\mu_i, \Sigma) \quad \forall i$$

$$x \in \mathbb{R}^p$$



De (**), Vamos a asignar $\forall \in \text{Prob. } k$ si

$$\frac{\mathbb{P}(Y=k | X=x)}{\mathbb{P}(Y=j | X=x)} \geq 1 \quad \forall j \neq k$$

i.e. $f_k^*(x) \pi_k \geq f_j^*(x) \pi_j$

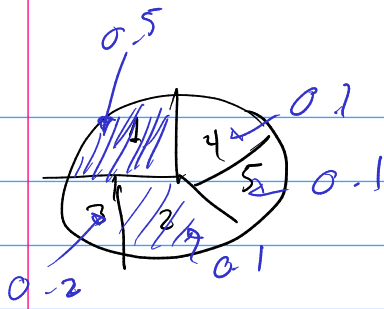
$$\log f_k^*(x) + \log \pi_k \geq \log f_j^*(x) + \log \pi_j$$

$$\Rightarrow -\frac{1}{2} \log |\Sigma| - \frac{1}{2} (x - \mu_k)^T \Sigma^{-1} (x - \mu_k) + \log \pi_k$$

$$\geq -\frac{1}{2} \log |\Sigma| - \frac{1}{2} (x - \mu_j)^T \Sigma^{-1} (x - \mu_j) + \log \pi_j$$

$$\Rightarrow \log \pi_k - \frac{1}{2} x^T \Sigma^{-1} x - \frac{1}{2} \mu_k^T \Sigma^{-1} \mu_k + x^T \Sigma^{-1} \mu_k$$

$$\geq \log \pi_j - \frac{1}{2} x^T \Sigma^{-1} x - \frac{1}{2} \mu_j^T \Sigma^{-1} \mu_j + x^T \Sigma^{-1} \mu_j$$



Supuesto $X=x \mid Y=j \sim \mathcal{N}_p(\mu_j, \Sigma)$

$$f_j(x) = \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}(x-\mu_j)^T \Sigma^{-1}(x-\mu_j)\right\}$$

$\underbrace{\quad}_{\mathbb{R}^p}$

$$\propto |\Sigma|^{-1/2} \exp\left\{-\frac{1}{2}(x-\mu_j)^T \Sigma^{-1}(x-\mu_j)\right\}$$

$$\log f_j(x) \propto -\frac{1}{2} \log |\Sigma| - \frac{1}{2} (x-\mu_j)^T \Sigma^{-1}(x-\mu_j)$$

i.e.

Asignamos a la población k si:

$$(\log \pi_k - \frac{1}{2} \mu_k^T \Sigma^{-1} \mu_k) + \underbrace{x^T (\Sigma^{-1} \mu_k)}_{(\mu_k^T \Sigma^{-1})x} \geq (\log \pi_j - \frac{1}{2} \mu_j^T \Sigma^{-1} \mu_j) + x^T \Sigma^{-1} \mu_j$$

$\forall j \neq k$

Sea $\left\{ \begin{array}{l} w_{0i} = \log \pi_i - \frac{1}{2} \mu_i^T \Sigma^{-1} \mu_i \\ y \quad w_i^T = \mu_i^T \Sigma^{-1} \end{array} \right. \forall i=1, \dots, K$

Asignamos a la Pop. k si

$$\boxed{w_{0k} + w_k^T x \geq w_{0j} + w_j^T x \quad \forall j \neq k}$$

LDA

QDA (Quadratic Discriminant Analysis)

$$X=x | Y=i \sim \mathcal{N}_p(\mu_i, \Sigma_i)$$

Asignamos a la población k si

$$-\frac{1}{2} \log |\Sigma_k| - \frac{1}{2} (x - \mu_k)^T \Sigma_k^{-1} (x - \mu_k) + \log \pi_k$$

$$\gg -\frac{1}{2} \log |\Sigma_j| - \frac{1}{2} (x - \mu_j)^T \Sigma_j^{-1} (x - \mu_j) + \log \pi_j$$

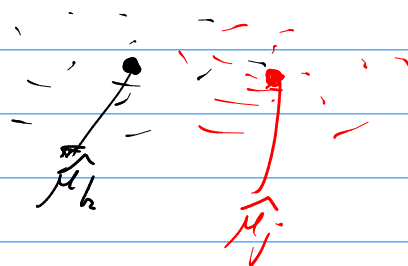
Reglas basadas en datos

$$\mathcal{D}_n = ((X_1, Y_1), \dots, (X_n, Y_n))$$

$$X_i \in \mathbb{R}^p, \quad i=1, \dots, n$$

$$n_k = \sum_{i=1}^n \mathbb{1}_{\{Y_i=k\}}$$

$$\Rightarrow \hat{\mu}_k = \frac{1}{n_k} \sum_{i=1}^n X_i \mathbb{1}_{\{Y_i=k\}}$$



QDA:

$$\hat{\Sigma}_k \equiv S_k = \frac{1}{n_k - 1} \sum_{i=1}^n (X_i - \hat{\mu}_k) (X_i - \hat{\mu}_k)^T \mathbb{1}_{\{Y_i=k\}}$$

LDA:

$$\hat{\Sigma} = \frac{1}{n-K} \left((n_1-1)S_1 + \dots + (n_K-1)S_K \right)$$

