

Inverse Gaussian distribution

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In probability theory, the **inverse Gaussian distribution** (also known as the **Wald distribution**) is a two-parameter family of continuous probability distributions with support on $(0,\infty)$.

Its probability density function is given by

$$f(x;\mu,\lambda) = \left[\frac{\lambda}{2\pi x^3}\right]^{1/2} \exp\frac{-\lambda(x-\mu)^2}{2\mu^2x}$$

for $x > 0$, where $\mu > 0$ is the mean and $\lambda > 0$ is the shape parameter.

As λ tends to infinity, the inverse Gaussian distribution becomes more like a normal (Gaussian) distribution. The inverse Gaussian distribution has several properties analogous to a Gaussian distribution. The name can be misleading: it is an "inverse" only in that, while the Gaussian describes a Brownian Motion's level at a fixed time, the inverse Gaussian describes the distribution of the time a Brownian Motion with positive drift takes to reach a fixed positive level.

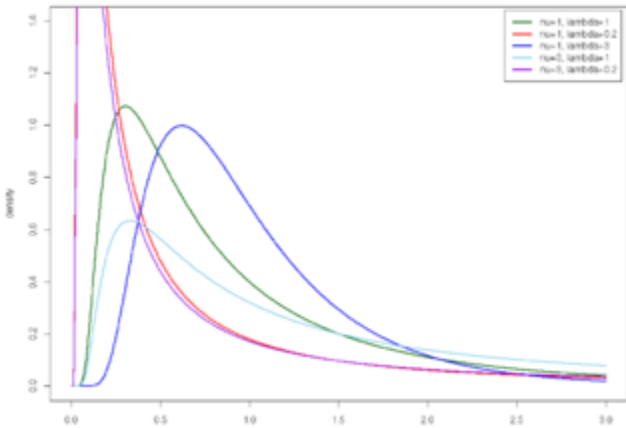
Its cumulant generating function (logarithm of the characteristic function) is the inverse of the cumulant generating function of a Gaussian random variable.

To indicate that a random variable X is inverse Gaussian distributed with mean μ and shape parameter λ we write

$$X \sim IG(\mu,\lambda).$$

Inverse Gaussian

Probability density function



parameters:	$\lambda > 0$ $\mu > 0$
support:	$x \in (0, \infty)$
pdf:	$\left[\frac{\lambda}{2\pi x^3}\right]^{1/2} \exp\frac{-\lambda(x-\mu)^2}{2\mu^2x}$
cdf:	$1 - \Phi\left(\sqrt{\frac{\lambda}{x}}\left(\frac{x}{\mu} - 1\right)\right) + \exp\left(\frac{2\lambda}{\mu}\right) \Phi\left(-\sqrt{\frac{\lambda}{x}}\left(\frac{x}{\mu} + 1\right)\right)$ where $\Phi()$ is the standard normal (standard Gaussian) distribution cdf

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Properties

Summation

If X_i has a $IG(\mu_0 w_i, \lambda_0 w_i^2)$ distribution for $i = 1, 2, \dots, n$ and all X_i are independent, then

$$S = \sum_{i=1}^n X_i \sim IG\left(\mu_0 \sum w_i, \lambda_0 \left(\sum w_i\right)^2\right).$$

Note that

$$\frac{\text{Var}(X_i)}{\text{E}(X_i)} = \frac{\mu_0^2 w_i^2}{\lambda_0 w_i^2} = \frac{\mu_0^2}{\lambda_0}$$

is constant for all i . This is a necessary condition for the summation. Otherwise S would not be inverse Gaussian.

Scaling

For any $t > 0$ it holds that

$$X \sim IG(\mu, \lambda) \quad \Rightarrow \quad tX \sim IG(t\mu, t\lambda).$$

Exponential family

The inverse Gaussian distribution is a two-parameter exponential family with natural parameters $-\lambda/(2\mu^2)$ and $-\lambda/2$, and natural statistics X and $1/X$.

Relationship with Brownian motion

The stochastic process X_t given by

$$\begin{aligned} X_0 &= 0 \\ X_t &= \nu t + \sigma W_t \end{aligned}$$

(where W_t is a standard Brownian motion and $\nu > 0$) is a Brownian motion with drift ν .

Then, the first passage time for a fixed level $\alpha > 0$ by X_t is distributed according to an inverse-gaussian:

$$T_\alpha = \inf\{0 < t \mid X_t = \alpha\} \sim IG\left(\frac{\alpha}{\nu}, \frac{\alpha^2}{\sigma^2}\right).$$

When drift is zero

A common special case of the above arises when the Brownian motion has no drift. In that case, parameter μ tends to infinity, and the first passage time for fixed level α has probability density function

$$f\left(x; \infty, \left(\frac{\alpha}{\sigma}\right)^2\right) = \frac{\alpha}{\sigma\sqrt{2\pi x^3}} \exp\left(-\frac{\alpha^2}{2x\sigma^2}\right).$$

This is a Lévy distribution with parameter $c = \frac{\alpha^2}{\sigma^2}$.

Maximum likelihood

The model where

$$X_i \sim IG(\mu, \lambda w_i), \quad i = 1, 2, \dots, n$$

with all w_i known, (μ, λ) unknown and all X_i independent has the following likelihood function

$$L(\mu, \lambda) = \left(\frac{\lambda}{2\pi}\right)^{\frac{n}{2}} \left(\prod_{i=1}^n \frac{w_i}{X_i^3}\right)^{\frac{1}{2}} \exp\left(\frac{\lambda}{\mu} - \frac{\lambda}{2\mu^2} \sum_{i=1}^n w_i X_i - \frac{\lambda}{2} \sum_{i=1}^n w_i \frac{1}{X_i}\right).$$

Solving the likelihood equation yields the following maximum likelihood estimates

$$\hat{\mu} = \frac{\sum_{i=1}^n w_i X_i}{\sum_{i=1}^n w_i}, \quad \frac{1}{\hat{\lambda}} = \frac{1}{n} \sum_{i=1}^n w_i \left(\frac{1}{X_i} - \frac{1}{\hat{\mu}}\right).$$

$\hat{\mu}$ and $\hat{\lambda}$ are independent and

$$\hat{\mu} \sim IG\left(\mu, \lambda \sum_{i=1}^n w_i\right) \quad \frac{n}{\hat{\lambda}} \sim \frac{1}{\lambda} \chi_{n-1}^2.$$

Generating random variates from an inverse-Gaussian distribution

The following algorithm may be used.^[1]

Generate a random variate from a normal distribution with a mean of 0 and 1 standard deviation

$$\nu = N(0, 1).$$

Square the value

$$y = \nu^2$$

and use this relation

$$x = \mu + \frac{\mu^2 y}{2\lambda} - \frac{\mu}{2\lambda} \sqrt{4\mu\lambda y + \mu^2 y^2}.$$

Generate another random variate, this time sampled from a uniformed distribution between 0 and 1

$$z = U(0, 1).$$

If

$$z \leq \frac{\mu}{\mu + x}$$

then return

$$x$$

else return

$$\frac{\mu^2}{x}.$$

Sample code in Java language:

```
public double inverseGaussian(double mu, double lambda) {
    Random rand = new Random();
    double v = rand.nextGaussian(); // sample from a normal distribution
    double y = v*v;
    double x = mu + (mu*mu*y)/(2*lambda) - (mu/(2*lambda)) * Math.sqrt(1 + 4*lambda*y);
    double test = rand.nextDouble(); // sample from a uniform distribution
    if (test <= (mu)/(mu + x))
        return x;
    else
        return (mu*mu)/x;
}
```

See also

- Generalized inverse Gaussian distribution
- Tweedie distributions

Notes

1. ^ *Generating Random Variates Using Transformations with Multiple Roots* by John R. Michael, William R. Schucany and Roy W. Haas, *American Statistician*, Vol. 30, No. 2 (May, 1976), pp. 88-90

References

- *The inverse gaussian distribution: theory, methodology, and applications* by Raj Chhikara and Leroy Folks, 1989 ISBN 0-8247-7997-5
- *System Reliability Theory* by Marvin Rausand and Arnljot Høyland
- *The Inverse Gaussian Distribution* by Dr. V. Seshadri, Oxford Univ Press, 1993

External links

- Inverse Gaussian Distribution (<http://mathworld.wolfram.com/InverseGaussianDistribution.html>) in Wolfram website.

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