

Binomial Likelihood

Suppose that a r.v. X obeys a Binomial distribution $\mathcal{Bin}(p)$, characterized by the following Probability Mass Function (PMF)

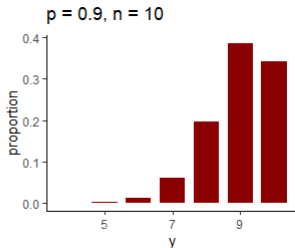
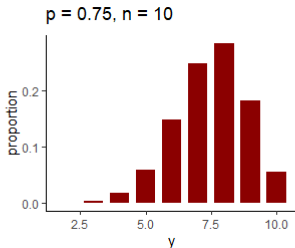
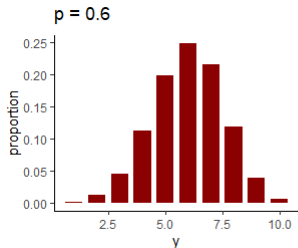
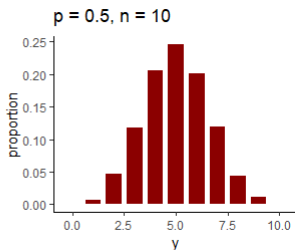
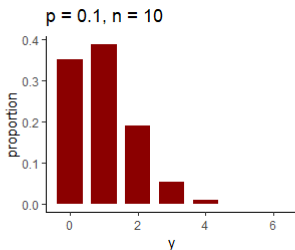
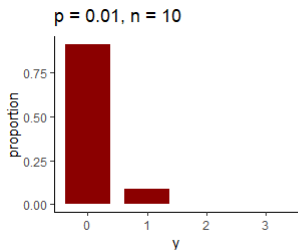
$$f_p(x) = \binom{n}{x} p^x (1-p)^{n-x}$$

If we observe a sample x_1, \dots, x_n , then the likelihood is just the product of the individual PMF

$$\begin{aligned} \pi(x_1, \dots, x_n \mid p) &= \prod_{i=1}^n f_p(x_i) = \prod_{i=1}^n \binom{n}{x_i} p^{x_i} (1-p)^{n-x_i} \\ &= \prod_{i=1}^n \binom{n}{x_i} p^{\sum_{i=1}^n x_i} (1-p)^{n-\sum_{i=1}^n x_i} \end{aligned}$$

where the model parameter p estimated by \hat{p} is the proportion of successes in n independent Bernoulli trials obtained by $\hat{p} = \frac{1}{n} \sum_{i=1}^n x_i$.

Some examples of Binomial samples



Beta prior

In Bayesian statistics, to model the uncertainty about p , a convenient prior would be the Beta distribution $Beta(\alpha, \beta)$, since its support is the interval $[0, 1]$. The Beta distribution has the following Probability Density Function (PDF)

$$\pi(p) = \frac{1}{B(\alpha, \beta)} p^{\alpha-1} (1-p)^{\beta-1} = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} p^{\alpha-1} (1-p)^{\beta-1}$$

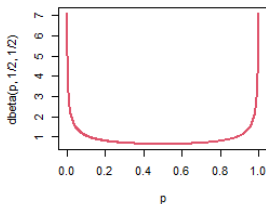
where α and β are shape parameters, $B(\alpha, \beta)$ is the Beta function and $\Gamma(\alpha)$ is the Gamma function, defined as $\int_0^\infty x^{\alpha-1} e^{-x} dx = (\alpha - 1)!$.

Let $p \sim Beta(\alpha, \beta)$. Then $E[p] = \frac{\alpha}{\alpha + \beta}$ and $var(p) = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}$.

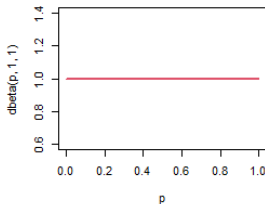
It is a convenient and flexible choice since the Beta distribution can take a wide variety of shapes.

Some examples of the Beta family

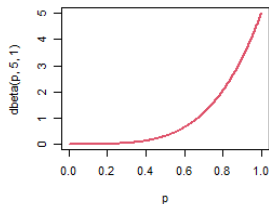
Beta(1/2, 1/2)



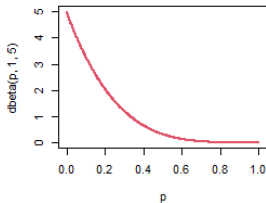
Beta(1, 1)



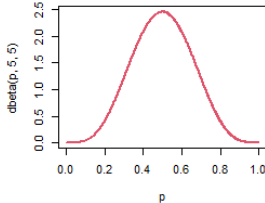
Beta(5, 1)



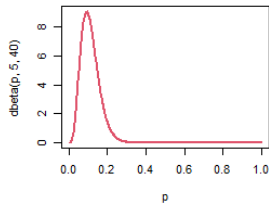
Beta(5, 5)



Beta(5, 5)



Beta(5, 40)



Beta posterior (1/2)

The posterior distribution for p if X is Binomial and a Beta prior is chosen for p will also have the functional form of a Beta r.v. That is why we say that Beta is the conjugate prior for a Binomial likelihood. Using the Bayes theorem, we have that

$$\pi(p \mid x_1, \dots, x_n) = \frac{\pi(x_1, \dots, x_n \mid p)\pi(p)}{\pi(x_1, \dots, x_n)}$$

and since $\sum_{i=1}^n x_i = n\bar{x}$, we have

$$\begin{aligned} \prod_{i=1}^n \binom{n}{x_i} p^{\sum_{i=1}^n x_i} (1-p)^{n-\sum_{i=1}^n x_i} \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} p^{\alpha-1} (1-p)^{\beta-1} \\ = \underbrace{\prod_{i=1}^n \binom{n}{x_i} \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)}}_{\text{do NOT depend on } p} p^{n\bar{x}} (1-p)^{n-n\bar{x}} p^{\alpha-1} (1-p)^{\beta-1} \\ \propto p^{n\bar{x}+\alpha-1} (1-p)^{n-n\bar{x}+\beta-1} \end{aligned}$$

Beta posterior (2/2)

So we find that $p \mid x_1, \dots, x_n \propto \text{Beta}(n\bar{x} + \alpha, n - n\bar{x} + \beta)$. The posterior mean and variance are given by

$$E[p \mid x] = \frac{n\bar{x} + \alpha}{n + \alpha + \beta} \quad \text{var}(p \mid x) = \frac{(n\bar{x} + \alpha)(n - n\bar{x} + \beta)}{(n + \alpha + \beta)^2(n + \alpha + \beta + 1)}$$

where $n\bar{x} = \sum_{i=1}^n x_i$ is the average number of successes in the sample size. Let's take a few examples and plot the likelihood, a possible prior and the posterior, all at once in R.

General remark on Bayesian inference

Unlike in the traditional Frequentist framework, the Bayesian approach views parameters as random variables rather than fixed, unknown quantities. Given a Poisson sample x_1, \dots, x_n and a Poisson parameter λ , from the Bayes theorem, we can write

$$\pi(p \mid x_1, \dots, x_n) = \frac{\pi(x_1, \dots, x_n \mid p) \pi(p)}{\pi(x_1, \dots, x_n)}$$

Adopting the 'proportional' notation, the constant term in the denominator is dropped so that the above expression is rewritten as $\pi(p \mid x_1, \dots, x_n) \propto \pi(x_1, \dots, x_n \mid p) \pi(p)$

When conjugate models are used (as in the case of a Beta-Binomial model), the posterior distribution can be identified and closed-form quantities of interest like a mean, a variance or quantiles can be computed. Most of the time in practice, the posterior distribution is intractable so that it is necessary to resort to MCMC techniques.

Working example

Suppose that we know that about 5% of e-mail received are spams. For 20 periods of 20 days each, we recorded the number of spams in our mailbox so that we have the following data at hand:

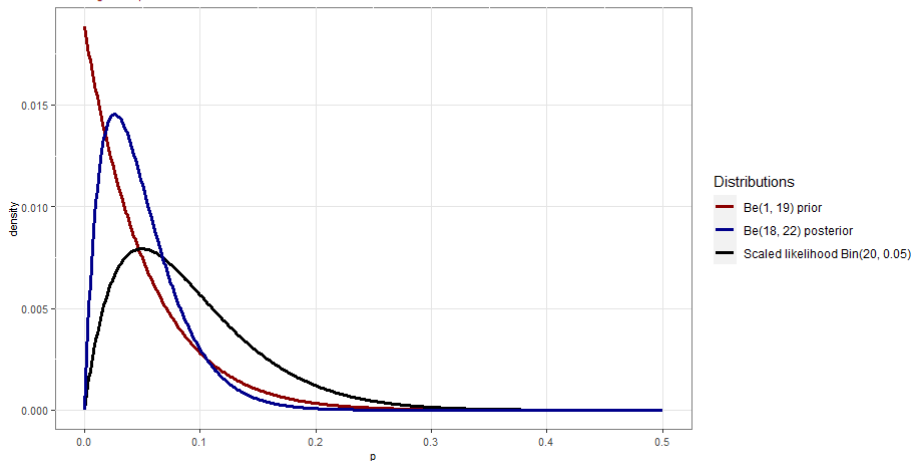
$$x_i = 1, 0, 0, 1, 0, 0, 1, 1, 0, 1, 2, 0, 0, 5, 2, 0, 0, 2, 0, 1$$

So assuming a Binomial likelihood with parameter $p = 0.05$, namely a $\text{Bin}(20, 0.05)$ likelihood for the data and using a $\text{Be}(1, 19)$ prior, with mean $1/(1+19) = 0.05$ and variance $19/((20)^2(21)) \approx 0.0023$, what is the posterior mean and the 95% credible interval for the model parameter?

Working example: posterior

Posterior distribution in blue - $\text{Be}(18, 22)$

Working example data



Posterior quantities obtained from direct sampling

```
1 # Posterior mean, posterior variance and 95% Credible Interval including the
   sample median
2 set.seed(2023)
3 data1 = rbinom(100000, size = n, prob = p1)
4 alpha_posterior = ((alpha1 + n*mean(data1))) # 19
5 beta_posterior = (n - n*mean(data1) + beta1) # 41
6
7 pmean = alpha_posterior / (alpha_posterior + beta_posterior)
8 pmean
9 # [1] 0.45
10
11 pvariance = (alpha_posterior *beta_posterior) / ((alpha_posterior + beta_
   posterior)^2 + (alpha_posterior + beta_posterior + 1) )
12 pvariance
13 # [1] 0.2413163
14
15 # 95% Credible Interval obtained by direct sampling (simulation)
16 set.seed(2023)
17 round(quantile(rbeta(n = 10^8, alpha_posterior, beta_posterior), probs = c
   (0.025, 0.5, 0.975)),4)
18 #    2.5%    50%   97.5%
19 # 0.3009 0.4492 0.6038
20
21 # Posterior mean obtained from direct sampling
22 set.seed(2023)
23 mean(rbeta(n = 10^8, alpha_posterior, beta_posterior))
24 # [1] 0.4500023
```

Working example: in conclusion

So the theoretical posterior mean is given by

$$E[p] = \frac{n\bar{x} + \alpha}{n + \alpha + \beta} = \frac{20 * 0.85 + 1}{20 + 1 + 19} = 18/40 = 0.45$$

By direct sampling, using 10^8 number of simulations, the posterior sample mean is 0.4500023

By direct sampling, a 95% Credible Interval is given by

$$[0.3009, 0.6038]$$

So, combining modeling and simulations, we are now able to generalize and infer to the whole population of spams in the mailbox those values from a sample of size 20.

Further reading and code

The R Project for Statistical Computing:

<https://www.r-project.org/>

Accessing the R code:

[https://github.com/JRigh/Poisson-Gamma-example-in-R/blob/main/
Poisson-Gamma](https://github.com/JRigh/Poisson-Gamma-example-in-R/blob/main/Poisson-Gamma)