

# Binomial Likelihood

Suppose that a r.v.  $X$  obeys a Binomial distribution  $\mathcal{Bin}(p)$ , characterized by the following Probability Mass Function (PMF)

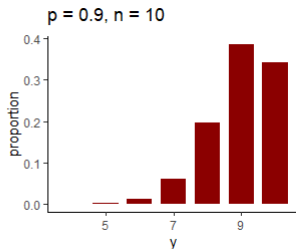
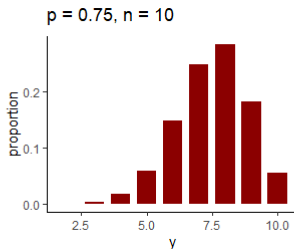
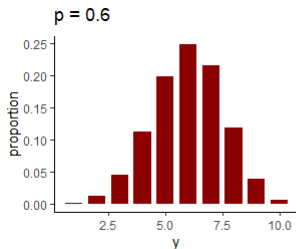
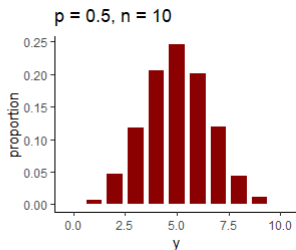
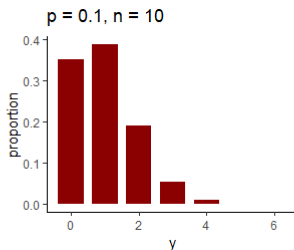
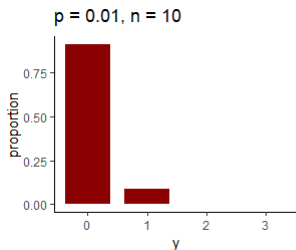
$$f_p(x) = \binom{n}{x} p^x (1-p)^{n-x}$$

If we observe a sample  $x_1, \dots, x_n$ , then the likelihood is just the product of the individual PMF

$$\begin{aligned}\pi(x_1, \dots, x_n \mid p) &= \prod_{i=1}^n f_p(x_i) = \prod_{i=1}^n \binom{n}{x_i} p^{x_i} (1-p)^{n-x_i} \\ &= \prod_{i=1}^n \binom{n}{x_i} p^{\sum_{i=1}^n x_i} (1-p)^{n-\sum_{i=1}^n x_i}\end{aligned}$$

where the model parameter  $p$  estimated by  $\hat{p}$  is the proportion of successes in  $n$  independent Bernoulli trials obtained by  $\hat{p} = \frac{1}{n} \sum_{i=1}^n x_i$ .

# Some examples of Binomial samples



## Beta prior

In Bayesian statistics, to model the uncertainty about  $p$ , a convenient prior would be the Beta distribution  $Beta(\alpha, \beta)$ , since its support is the interval  $[0, 1]$ . The Beta distribution has the following Probability Density Function (PDF)

$$\pi(p) = \frac{1}{B(\alpha, \beta)} p^{\alpha-1} (1-p)^{\beta-1} = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} p^{\alpha-1} (1-p)^{\beta-1}$$

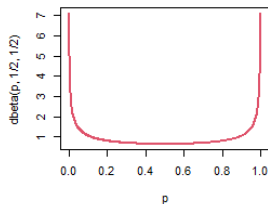
where  $\alpha$  and  $\beta$  are shape parameters,  $B(\alpha, \beta)$  is the Beta function and  $\Gamma(\alpha)$  is the Gamma function, defined as  $\int_0^\infty x^{\alpha-1} e^{-x} dx = (\alpha - 1)!$ .

Let  $p \sim Beta(\alpha, \beta)$ . Then  $E[p] = \frac{\alpha}{\alpha + \beta}$  and  $var(p) = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}$ .

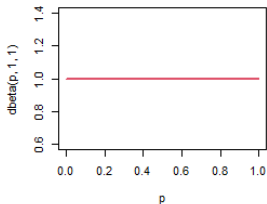
It is a convenient and flexible choice since the Beta distribution can take a wide variety of shapes.

# Some examples of the Beta family

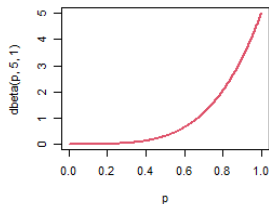
Beta(1/2, 1/2)



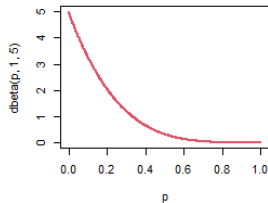
Beta(1, 1)



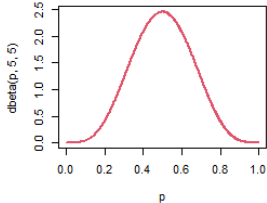
Beta(5, 1)



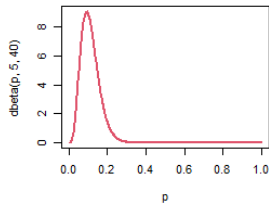
Beta(5, 5)



Beta(5, 5)



Beta(5, 40)



## Beta posterior (1/2)

The posterior distribution for  $p$  if  $X$  is Binomial and a Beta prior is chosen for  $p$  will also have the functional form of a Beta r.v. That is why we say that Beta is the conjugate prior for a Binomial likelihood. Using the Bayes theorem, we have that

$$\pi(p \mid x_1, \dots, x_n) = \frac{\pi(x_1, \dots, x_n \mid p)\pi(p)}{\pi(x_1, \dots, x_n)}$$

and since  $\sum_{i=1}^n x_i = n\bar{x}$ , we have

$$\begin{aligned} \prod_{i=1}^n \binom{n}{x_i} p^{\sum_{i=1}^n x_i} (1-p)^{n-\sum_{i=1}^n x_i} \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} p^{\alpha-1} (1-p)^{\beta-1} \\ = \underbrace{\prod_{i=1}^n \binom{n}{x_i} \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)}}_{\text{do NOT depend on } p} p^{n\bar{x}} (1-p)^{n-n\bar{x}} p^{\alpha-1} (1-p)^{\beta-1} \\ \propto p^{n\bar{x}+\alpha-1} (1-p)^{n-n\bar{x}+\beta-1} \end{aligned}$$

## Beta posterior (2/2)

So we find that  $p \mid x_1, \dots, x_n \propto \text{Beta}(n\bar{x} + \alpha, n - n\bar{x} + \beta)$ . The posterior mean and variance are given by

$$E[p \mid x] = \frac{n\bar{x} + \alpha}{n + \alpha + \beta} \quad \text{var}(p \mid x) = \frac{(n\bar{x} + \alpha)(n - n\bar{x} + \beta)}{(n + \alpha + \beta)^2(n + \alpha + \beta + 1)}$$

where  $n\bar{x} = \sum_{i=1}^n x_i$  is the average number of successes in the sample size. Let's take a few examples and plot the likelihood, a possible prior and the posterior, all at once in R.

# General remark on Bayesian inference

Unlike in the traditional Frequentist framework, the Bayesian approach views parameters as random variables rather than fixed, unknown quantities. Given a Poisson sample  $x_1, \dots, x_n$  and a Poisson parameter  $\lambda$ , from the Bayes theorem, we can write

$$\pi(p \mid x_1, \dots, x_n) = \frac{\pi(x_1, \dots, x_n \mid p) \pi(p)}{\pi(x_1, \dots, x_n)}$$

Adopting the 'proportional' notation, the constant term in the denominator is dropped so that the above expression is rewritten as  $\pi(p \mid x_1, \dots, x_n) \propto \pi(x_1, \dots, x_n \mid p) \pi(p)$

When conjugate models are used (as in the case of a Beta-Binomial model), the posterior distribution can be identified and closed-form quantities of interest like a mean, a variance or quantiles can be computed. Most of the time in practice, the posterior distribution is intractable so that it is necessary to resort to MCMC techniques.

## Working example

Suppose that we know that about 5% of e-mail received are spams. For 20 periods of 20 days each, we recorded the number of spams in our mailbox so that we have the following data at hand:

$$x_i = 1, 0, 0, 1, 0, 0, 1, 1, 0, 1, 2, 0, 0, 5, 2, 0, 0, 2, 0, 1$$

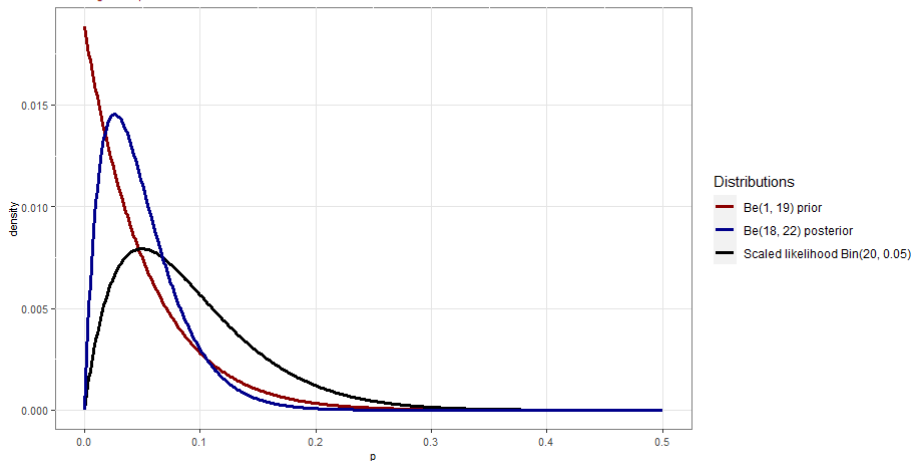
So assuming a Binomial likelihood with parameter  $p = 0.05$ , namely a  $\text{Bin}(20, 0.05)$  likelihood for the data and using a  $\text{Be}(1, 19)$  prior, with mean  $1/(1+19) = 0.05$  and variance  $19/((20)^2(21)) \approx 0.0023$ , what is the posterior mean and the 95% credible interval for the model parameter?



# Working example: posterior

Posterior distribution in blue -  $\text{Be}(18, 22)$

*Working example data*



# Posterior quantities obtained from direct sampling

```
1 # Posterior mean, posterior variance and 95% Credible Interval including the
   sample median
2 set.seed(2023)
3 data1 = rbinom(100000, size = n, prob = p1)
4 alpha_posterior = ((alpha1 + n*mean(data1))) # 19
5 beta_posterior = (n - n*mean(data1) + beta1) # 41
6
7 pmean = alpha_posterior / (alpha_posterior + beta_posterior)
8 pmean
9 # [1] 0.45
10
11 pvariance = (alpha_posterior *beta_posterior) / ((alpha_posterior + beta_
   posterior)^2 + (alpha_posterior + beta_posterior + 1) )
12 pvariance
13 # [1] 0.2413163
14
15 # 95% Credible Interval obtained by direct sampling (simulation)
16 set.seed(2023)
17 round(quantile(rbeta(n = 10^8, alpha_posterior, beta_posterior), probs = c
   (0.025, 0.5, 0.975)),4)
18 #    2.5%    50%   97.5%
19 # 0.3009 0.4492 0.6038
20
21 # Posterior mean obtained from direct sampling
22 set.seed(2023)
23 mean(rbeta(n = 10^8, alpha_posterior, beta_posterior))
24 # [1] 0.4500023
```

## Working example: in conclusion

So the theoretical posterior mean is given by

$$E[p] = \frac{n\bar{x} + \alpha}{n + \alpha + \beta} = \frac{20 * 0.85 + 1}{20 + 1 + 19} = 18/40 = 0.45$$

By direct sampling, using  $10^8$  number of simulations, the posterior sample mean is 0.4500023

By direct sampling, a 95% Credible Interval is given by

$$[0.3009, 0.6038]$$

So, combining modeling and simulations, we are now able to generalize and infer to the whole population of spams in the mailbox those values from a sample of size 20.

## Further reading and code

The R Project for Statistical Computing:

<https://www.r-project.org/>

Accessing the R code:

<https://github.com/JRigh/Beta-Binomial-Example-in-R>