

On the Assessment of Monte Carlo Error in Simulation-Based Statistical Analyses

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2017 (updated 2025)

Overview

- (i) Drawing attention on the importance of reporting Monte Carlo Error (MCE) in statistical simulation studies.
- (ii) Define simple and practical tools for estimating MCE.
- (iii) Providing means for determining the number of replications R required to achieve a prespecified desired level of accuracy.

Notes: *The uncertainty as measured by Monte Carlo Error (MCE) or sampling variability - in published literature is mostly not reported. Its variability comes from simulations and number of simulations. This uncertainty comes from the fact that statistical procedures (measures) are repeated on simulated data (finite sample from an infinite population). Thus two different simulation runs will yield different results. According to a survey (of the same authors) published in 2007 in Biometrics, Biometrika and JASA, out of 223 articles where simulations studies were performed, only 8 reported a formal justification of the number of replications used or a MCE (standard deviation).*

Simulation study: experiment 1 (1/4)

Setting: Quantification of association between Y and X , two binary r.v. We assume the following generating process:

$$\text{logit } P(Y = 1 | X) = \beta_0 + \beta_1 X$$

The MCE is defined as:

$$MCE(\hat{\phi}_R) = \sqrt{\text{var}(\hat{\phi}_R)}$$

$N = 100$ (sample size), $R = 100, 500, 1, 000, 2, 500, 5, 000, 10, 000$ (number of replicates), $M = 1, 000$ (number of simulations)

Notes: In a context of logistic regression with known values for the parameters $\beta_0 = -1$ and $\beta_1 = \ln(2)$, we want to have an estimation of three operating characteristics of the Maximum Likelihood Estimator (MLE) for the slope parameter β_1 (log-odds ratio). The MCE is the standard deviation of the estimated quantities of interest (operating characteristics).

Simulation study: experiment 1 (2/4)

The three chosen operating characteristics:

$\hat{\phi}_R^b$, a MC estimate of **percent bias** for MLE of β_X :

$$\hat{\phi}_R^b = \frac{1}{R} \sum_{r=1}^R \frac{\hat{\beta}_X^r - \bar{\beta}_X}{\bar{\beta}_X}$$

$\hat{\phi}_R^c$, a MC estimate of **coverage probability of the 95% CI**:

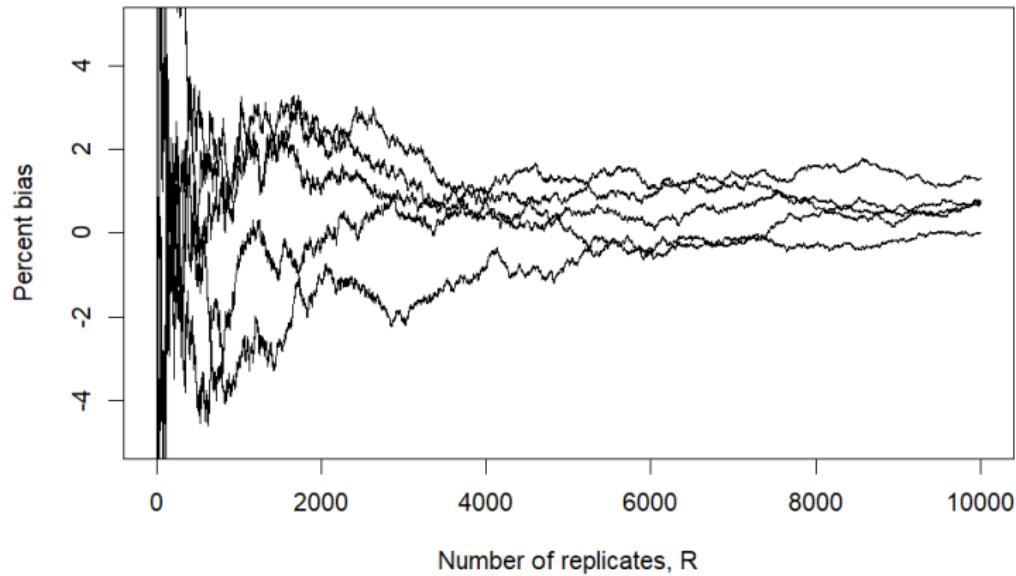
$$\hat{\phi}_R^c = \sum_{r=1}^R 1 \left[\hat{\beta}_X^r - 1.96 \text{se}(\hat{\beta}_X^r) \leq \beta_X \leq \hat{\beta}_X^r + 1.96 \text{se}(\hat{\beta}_X^r) \right]$$

$\hat{\phi}_R^p$, a MC estimate of **power** to detect an association:

$$\hat{\phi}_R^p = \sum_{r=1}^R 1 \left[\frac{\hat{\beta}_X^r}{\text{se}(\hat{\beta}_X^r)} < 1.96 \right]$$

Simulation study: experiment 1 (3/4)

MC estimates of percent bias for the MLE



Notes: Still a surprising amount of variability for the MC estimate of percent bias for the MLE after 10000 replications. The range of percent bias (after 10000 replicates) is around 2 (this range is also a random variable).

Simulation study: experiment 1 (4/4))

| Operating characteristic | R | min. | max. | mean | MCE |
|--------------------------|--------|--------|--------|-------|------|
| Percent bias | 100 | -22.48 | 24.42 | 1.03 | 6.82 |
| | 500 | -10.04 | 11.76 | 0.92 | 3.21 |
| | 1,000 | -6.91 | 7.74 | 0.98 | 2.32 |
| | 2,500 | -3.71 | 4.80 | 0.96 | 1.36 |
| | 5,000 | -2.38 | 3.63 | 0.90 | 0.95 |
| | 10,000 | -1.52 | 3.09 | 0.90 | 0.66 |
| Coverage rate | 100 | 85.00 | 100.00 | 94.65 | 2.23 |
| | 500 | 90.80 | 97.40 | 94.52 | 0.99 |
| | 1,000 | 92.30 | 96.50 | 94.55 | 0.71 |
| | 2,500 | 93.04 | 96.00 | 94.53 | 0.48 |
| | 5,000 | 93.22 | 95.52 | 94.50 | 0.36 |
| | 10,000 | 93.57 | 99.36 | 94.50 | 0.32 |
| Power | 100 | 20.00 | 50.00 | 33.16 | 4.68 |
| | 500 | 26.20 | 39.40 | 32.99 | 2.16 |
| | 1,000 | 28.40 | 37.70 | 33.04 | 1.55 |
| | 2,500 | 29.32 | 36.52 | 33.03 | 0.96 |
| | 5,000 | 30.70 | 35.26 | 32.99 | 0.66 |
| | 10,000 | 31.44 | 34.39 | 32.99 | 0.46 |

Note: I use $M = 1,000$ (instead of 500,000). Still, the results are close. Main differences: larger min and smaller max (i.e. for $R = 100$. The authors get -34 and 35). I have about the same MCE. It suggest that more than $R = 10,000$ to be within one unit -0.1% to 1.9%, 95% on average. Cov. rate means are different (they obtain 95.3). Also, less MCE, indicating that $R = 2,500$ is optimal (within one unit of the true value 95% of the time).

Simulation study: experiment 2 (1/2)

Setting: Quantification of MCE.

By the strong LLN, we have that: $\hat{\phi}_R \rightarrow E[\phi(X)] = \phi$. Also, from the CLT, we know that: $\sqrt{(\hat{\phi}_R - \phi)} \rightarrow N(0, \sigma_\phi^2)$

We will consider two measures:

$$\widehat{MCE}_{clt_{\hat{\phi}_R}} = \frac{\hat{\sigma}_\phi}{R} = \frac{1}{R} \sqrt{\sum_{r=1}^R \left(\phi(X_r) - \hat{\phi}_R \right)^2}$$

$$\widehat{MCE}_{boot_{\hat{\phi}_R}} = \frac{1}{B} \sqrt{\sum_{b=1}^B \left(\hat{\phi}_R(\mathbf{X}_b^*) - \overline{\hat{\phi}_R(\mathbf{X}_b^*)} \right)^2}$$

Note: A MCE bootstrap estimate is the standard deviation across bootstrap statistics. We use $B = 100, 200$ and 500 for $MCE_{boot_{\hat{\phi}_R}}$.

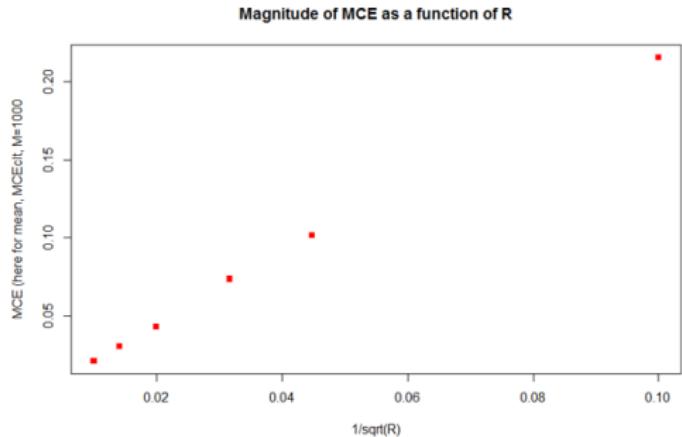
Simulation study: experiment 2 (2/2)

| Characteristic | R | value | $\widehat{MCE}_{clt \hat{\phi}_R}$ | $\widehat{MCE}_{boot \hat{\phi}_R}$ | | |
|----------------|--------|---------|------------------------------------|-------------------------------------|---------|---------|
| | | | | B = 100 | B = 200 | B = 500 |
| Mean | 100 | 1.0254 | 0.2156 | 0.2210 | 0.2268 | 0.2128 |
| | 500 | 0.9248 | 0.1016 | 0.0958 | 0.1027 | 0.1038 |
| | 1,000 | 0.9810 | 0.0734 | 0.0724 | 0.0782 | 0.0727 |
| | 2,500 | 0.95572 | 0.0430 | 0.0416 | 0.0397 | 0.0432 |
| | 5,000 | 0.9045 | 0.0302 | 0.0308 | 0.0307 | 0.0297 |
| | 10,000 | 0.8984 | 0.0208 | 0.0209 | 0.0224 | 0.0209 |
| MCE | 100 | 6.8191 | NA | 0.1601 | 0.1548 | 0.1600 |
| | 500 | 3.2131 | NA | 0.0678 | 0.0660 | 0.0699 |
| | 1,000 | 2.3205 | NA | 0.0482 | 0.0524 | 0.0501 |
| | 2,500 | 1.3603 | NA | 0.0291 | 0.0279 | 0.0293 |
| | 5,000 | 0.9540 | NA | 0.0211 | 0.0191 | 0.0201 |
| | 10,000 | 0.6571 | NA | 0.0146 | 0.0128 | 0.0143 |

Note: Consistent with what the authors obtained. The MCE is reduced in all cases by about 90% when we increase R from 100 to 10,000, whatever the number of simulations M. The authors use M = 500,000 but I use M = 1,000 for computational reasons. We only give point estimation (but could also have reported CI's).

Remarks

- The magnitude of MCE seems to be linear in $1/\sqrt{R}$
- As $1/\sqrt{R}$ tends to 0, MCE tends to 0
- Further work: Use a constrained linear regression to estimate R to achieve a desired level of MCE.



Notes: As $1/\sqrt{R}$ tends to 0, MCE will tend to 0 (useful to assess desired proportion of MCE). The authors proposed an R package called MCE (which has been archived) implementing the BCG plot. The estimated slope helps estimate the number of replications R to achieve an acceptable MCE.

Conclusions

- (i) The MCE can be more substantial than traditionally thought.
- (ii) Even after 500,000 simulations, there is residual uncertainty.
- (iii) The magnitude of the MCE depends on various factors, i.e. parameters, operating characteristics, variability in the data.
- (iv) The MCE can be drastically reduced by increasing R , the number of replicates.
- (iv) In simulations of most published scientific studies, the MCE is rarely reported though documenting uncertainty (usually by standard errors, p-values and CIs) is commonly insisted on.

Note: *In more complex settings, the MCE is more complicated to estimate.*

References

Elizabeth KOEHLER, Elizabeth BROWN, and Sebastien J.-P. A. HANEUSE, 2009. On the Assessment of Monte Carlo Error in Simulation-Based Statistical Analyses.

Christian P. ROBERT, George CASELLA, 2005. Monte Carlo Statistical Methods.

Maria L. RIZZO, 2008. Statistical computing with R.

F. H. C. MARRIOTT, 1979. Barnard's Monte Carlo Tests: How Many Simulations?

Full R code of this reproduced study available upon request.