Explanatory or Predictive Models for Pair-Contest Success

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Structured Bradley-Terry Model

- 'Ability', α_i , as a function of predictors
- Logistic probability function
- Generalised Linear Mixed Model

$$\lambda_{i} = \log(\alpha_{i})$$

$$= \sum_{p=1}^{P} \beta_{p} x_{i,p} + b_{i},$$

$$\eta_{i,j} = \lambda_{i} - \lambda_{j}$$

$$\mathbb{P}(i \ beats \ j) = \frac{1}{1 + e^{-\eta_{i,j}}}$$

Penalized Quasi-Likelihood Approach

- $\mathbb{E}[\mathbf{y}|\mathbf{b}] = g^{-1}(\mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{b})$
- $\mathbb{E}(y_i|\mathbf{b}) = \mu_i^b$, $Var(y_i|\mathbf{b}) = \phi a_i v(\mu_i^b)$, where v(.) is a specified variance function, a_i is a known constant and ϕ is a dispersion parameter. Assume that $\mathbf{b} \sim N(\mathbf{0}, \mathbf{D}(\theta))$.
- The integrated quasi-likelihood function is given by

$$e^{ql(\boldsymbol{\beta}, \boldsymbol{\theta})} \propto |\mathbf{D}|^{-1/2} \int \exp \left\{ -\frac{1}{2\phi} \sum_{i=1}^{n} d_i(y_i, \mu_i^b) - \frac{1}{2} \mathbf{b}^T \mathbf{D}^{-1} \mathbf{b} \right\} d\mathbf{b}$$

• We use Laplace's method of integral approximation and maximise the resulting expression to estimate β and θ .

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Laplace Method Bias

- Laplace's method involves a normal approximation.
- Our data is Binomial and sparse.

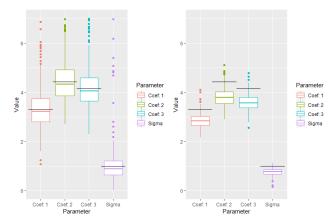


Figure: Summaries of parameters estimated from 500 contests, each with 1000 matches. Left: 10 players. Right: 100 players.

Sequential Reduction Method

Cliques and Maximal cliques

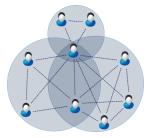


Figure: Source from:

http://c0de-x.com/how-to-find-maximal-cliques/cliques/

Sequential Reduction Method

- Conditional dependence structure of random effects
 - Posterior dependence graph
 - · A vertex for each random effect
 - An edge between two vertices if in the design matrix for the random effects, the two columns contain at least a row that both have non-zero values
- Factorise the integrand of likelihood over maximal cliques

$$g(b_1, \dots, b_m | \beta, D, W) = \prod_{i=1}^n \prod_{j:j>i} f(W_{i,j} | \eta_{i,j}) \prod_{\ell=1}^m \varphi(b_\ell)$$
$$= \prod_{C \in M(\mathcal{G})} g_C(\mathbf{b}_C)$$

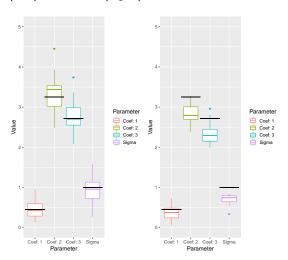
• The Hammersley-Clifford theorem [Besag(1974)]

Sequential Reduction Method

- Factorise the likelihood over maximal cliques
- Por each of the random effects
 - 1 Integrate over the random effect by integrating over the maximal cliques that contains the random effect
 - 2 Factorise the marginal density obtained over the maximal cliques of the new posterior dependence graph

Sequential Reduction vs PQL Results

• 100 players and the simulation is repeated 20 times for sequential reduction (left) and PQL (right).



MCMC Approach with Flat Prior Bayesian Formulation

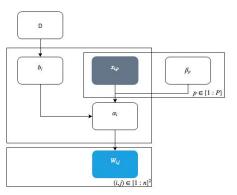


Figure: Plate diagram showing the dependency on observed win 'counts': $\{W_{ij}\}_{i,j}$ based on covariates $\{x_{i,j}\}_{i,j}$ and the unknowns: β_p, α_i, b_i and D

- Flat priors
- Gibbs sampling with nested rejection sampling
- [Zeger and Karim(1991)]

MCMC vs PQL Results

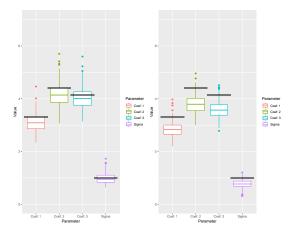


Figure: #players = 100, simulated data with 1000 matches. Left: Boxplot of realised states in a single Markov Chain (of length 500) for Bayesian formulation detailed above. Right: S = 500 simulated PQL estimates. Horizontal black lines indicate the 'true' values from the data generating process.

Conclusion

- The PQL approach suffers from significant bias when faced with sparse, binary data
- Sequential reduction can also be computationally expensive in some scenarios
- The MCMC described approach is fairly complicated and computationally expensive
- Further work is required to identify an improved method

References



Spatial interaction and the statistical analysis of lattice systems. Journal of the Royal Statistical Society. Series B (Methodological), pages 192–236, 1974.

Scott L Zeger and M Rezaul Karim.

Generalized linear models with random effects; A Gibbs sampling approach.

Journal of the American statistical association, 86(413):79–86, 1991.

Likelihood of Binomial GLMM

$$L(\beta, D) = \int \prod_{i=1}^{n} \prod_{j:j>i} f(W_{i,j}|\eta_{i,j}) \varphi_D(b_i) \mathbf{db}$$

$$f(W_{i,j}|\eta_{i,j}) = {W_{i,j} + W_{j,i} \choose W_{i,j}} \left(\frac{1}{1 + e^{-\eta_{i,j}}}\right)^{W_{i,j}} \left(\frac{1}{1 + e^{\eta_{i,j}}}\right)^{W_{j,i}}$$

$$\varphi_D(b_i) = \frac{1}{|D|^{-\frac{1}{2}}} e^{-\frac{1}{2}b_i D^{-1}b_i}$$