Probability distributions EBP038A05 Assignments

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INTRODUCTION

Here we just provide the exercises of the assignments. For information with respect to grading we refer to the course manual.

Each assignment contains code in Python and R that shows how to implement an example or the solution of an exercises of BH. You have to run the code, read the output, and explain how the code works.

We include Python and R code, and leave the choice to you what to use. In the exam we will also include both languages in the same problem, so you can stay with the language you like. You should know, however, that many of you will need to learn multiple languages later in life. For instance, when you have to access databases to obtain data about customers, patients, clients, suppliers, inventory, demand, lifetimes (whatever), you often have to use sql. Once you have the raw data, you process it with R or Python to do statistics or make plots. For your interest, based on the statistics here or here, Python scores (much) higher than R in popularity; if you opt for a business career, the probability you have to use Python is simply higher than to have to use R.

You should become familiar with looking up documentation on coding on the web, no matter your programming language of choice. Invest time in understanding the, at times, rather technical and terse, explanations. Once you are used to it, the core documentation is faster to read, i.e., less clutter. In the long run, it pays off.

The rules:

- 1. For each assignment, you have to turn in a pdf document typeset in Lag. Include a title, group number, student names and ids, and date.
- 2. We expect brief answers, just a few sentences, or a number plus some short explanation. The idea of the assignment is to help you study and improve your coding skills by showing good code, not to turn you in a writer.
- 3. When you have to turn in a graph, provide decent labels and a legend, ensure the axes have labels too.

¹ (While I (= NvF) worked at a bank, I used Fortran for numerical work, awk for string parsing and making tables, excel, SAS to access the database, and matlab for other numerical work, all next to each other. I got extremely tired of this, so I went to using Python as this can do all of this stuff, but within one language.

1 ASSIGNMENT 1

print(A[-1])

1.1 Why is the Exponential Distribution so important?

At the Paris metro, a train arrives every 3 minutes on a platform. Suppose that 50 people arrive between the departure of a train and an arrival. It seems entirely reasonable to me to model the arrival times of the individual people as distributed on the interval [0,3]. What is the distribution of the inter-arrival times of these people? Amazingly, this turns out to be nearly exponential. To check this fact, we use simulation in this question.

You might want to compare your final result to Figure BH.13.1 (It is not forbidden to read the book beyond what you have to do for this course!).

```
Python Code
   import numpy as np
   np.random.seed(3)
   num = 5 # small sample at first, for checking.
   start, end = 0, 3
   labda = num / (end - start) # per minute
   print(1 / labda)
   A = np.sort(np.random.uniform(start, end, size=num))
   print(A) #this
   print(A[1:])
13
   print(A[:-1])
   X = A[1:] - A[:-1]
15
   print(X)
16
17
   print(X.mean(), X.std())
18
                                         R Code
   set.seed(3)
   num = 5
   start = 0
   end = 3
   labda = num / (end - start)
   print(1 / labda)
  A = sort(runif(num, min = start, max = end))
10
   print(A) #this
11
```

```
print(A[-length(A)])
    X = A[-1] - A[-length(A)]
    print(X)

print(mean(X))
    print(sd(X))
```

- **Ex 1.1.** Explain the result of line P. 12 (R.11) marked as 'this'.²
- Ex 1.2. Compare the result of line P.13 and P.14 (R.12, R.13); explain what is A[1:] (A[-1])
- **Ex 1.3.** Compare the result of line P.12 and P.14 (R.11 and R.13); explain what is A[:-1](A[-length(A)]).
- **Ex 1.4.** Explain what is X in P.15 (R.14).
- **Ex 1.5.** Why do we compare $1/\lambda$ and X.mean()?
- **Ex 1.6.** Recall that $E[X] = \sigma(X)$ when $X \sim Exp(\lambda)$. Hence, what do you expect to see for X.std()?
- Ex 1.7. Run the code for a larger sample, e.g. 50, and discuss (very briefly) your results.

1.2 BH.5.6.5

Read this example of BH first. We chop up the exercise in many small steps.

For the code below, run it first for a small number of samples (Here I choose samples=2). Then read (and if necessary put in more) print statements to see what results you get. Use the output to answer the questions below.

```
import numpy as np
from scipy.stats import expon

np.random.seed(10)

labda = 6
num = 3
samples = 2

X = expon(scale=labda).rvs((samples, num))
print(X) #this
T = np.sort(X, axis=1)
print(T)
```

² Line P.x refers to line x of the Python code. Line R.x refers to line x of the R code.

```
print(T.mean(axis=0))
14
15
   expected = np.array([labda / (num - j) for j in range(num)])
16
   print(expected)
  print(expected.cumsum())
                                         R Code
   set.seed(10)
   labda = 6
   num = 3
   samples = 2
  X = matrix(rexp(samples * num, rate = 1 / labda), nrow = samples, ncol = num)
  print(X) #this
  bigT = X
   for (i in 1:samples) {
     bigT[i,] = sort(bigT[i,])
   }
12
  print(bigT)
   print(colMeans(bigT))
  expected = rep(0, num)
   for (j in 1:num) {
     expected[j] = labda / (num - (j - 1))
18
19
   print(expected)
20
   print(cumsum(expected))
```

- **Ex 1.8.** In line P.11 (R.8), also marked as 'this', we print the value of X in line P.10 (R.7), respectively. What is the meaning of X?
- **Ex 1.9.** What is the meaning of T in line P.12 (R.11)?
- **Ex 1.10.** What do we print in line P.14 (R.14)?
- **Ex 1.11.** What is meaning of the variable expected?
- Ex 1.12. What is the cumsum of expected?
- **Ex 1.13.** Now that you understand what is going on, rerun the simulation for a larger number of samples, e.g., 1000, and discuss the results briefly.

 $[\]overline{}^3$ Line P.x refers to line x of the Python code. Line R.x refers to line x of the R code.

1.3 BH.7.1

First read and solve BH.7.1. Answers and hints are in the study guide.

- **Ex 1.14.** Sketch, on paper, the 2D area that corresponds to the event that Alice and Bob meet. Just make a photo with your mobile phone and include it. (Like this you learn how to include jpeg or png files in LaTeX.)
- **Ex 1.15.** Write *A* and *B* for the rvs corresponding to the arrival times of Alice and Bob. Explain you have to compute the expectation of $I_{A < B+0.25} \cdot I_{B-0.25 < A}$ to compute the probability that Alice and Bob meet. Then, explain that it is equal to

$$\int_0^1 \int_0^1 I_{a < b + 0.25} \cdot I_{b - 0.25 < a} \, \mathrm{d}a \, \mathrm{d}b. \tag{1.1}$$

- **Ex 1.16.** Solve the integral by hand to compute the probability that they meet.
- **Ex 1.17.** Explain the code below to compute numerically the probability of overlap. Check the web on what is an anonymous function in Python or R, and explain why we use that here. (Hint: anonymous functions are useful "when it's not worth the effort to give it a name". Do we use elsewhere the anonymous function we define here?)

```
Python Code

from scipy.integrate import dblquad

area = dblquad(lambda a, b: (b - 0.25 < a) * (a < b + 0.25), 0, 1, 0, 1)

print(area)

R Code

library(cubature)

area = adaptIntegrate(function(x){ (x[2] - 0.25 < x[1]) * (x[1] < x[2] + 0.25)},

lowerLimit = c(0, 0), upperLimit = c(1, 1))

print(area$integral)

print(area$error)
```

Ex 1.18. Explain the output, i.e., the results of the print statements, of this piece of code.

```
import numpy as np
from scipy.stats import uniform

np.random.seed(3)

n = 5
alice = uniform(0, 1).rvs(n)
```

```
bob = uniform(0, 1).rvs(n)
print(alice)
print(bob)
print((alice < bob))

R Code

set.seed(3)

a n = 5
a alice = runif(n,0,1)
bob = runif(n,0,1)
print(alice)
print(bob)
print(bob)
print(alice<bob)</pre>
```

It is useful to set the seed of the random number generator so that we always get the same results. This helps a lot when debugging, because the numbers don't change time and again. (You can remove the line that sets the seed, and then rerun the program a few times. You should get different numbers for each run.)

Ex 1.19. Run this piece of code, and then check that you get different output every time.

```
Python Code
import numpy as np
from scipy.stats import uniform
n = 5
alice = uniform(0, 1).rvs(n)
bob = uniform(0, 1).rvs(n)
print(alice)
print(bob)
overlap = (alice < bob) # this</pre>
print(overlap)
                                       R Code
n = 5
alice = runif(n,0,1)
bob = runif(n,0,1)
print(alice)
print(bob)
overlap = alice<bob #this</pre>
print(overlap)
```

Explain that the line marked as 'this' implements an indicator function.

Ex 1.20. Run this piece of code, and explain the output.

```
Python Code
   import numpy as np
   from scipy.stats import uniform
   np.random.seed(3)
  n = 5
  alice = uniform(0, 1).rvs(n)
  bob = uniform(0, 1).rvs(n)
   overlap = (bob - 0.1 < alice) * (alice < bob + 0.1)
  print(alice)
10
  print(bob)
  print(overlap) # this
                                        R Code
   set.seed(3)
  n = 5
alice = runif(n,0,1)
  bob = runif(n, 0, 1)
  overlap = (bob - 0.1 < alice) * (alice < bob + 0.1)
   print(alice)
8 print(bob)
   print(overlap) #this
```

Ex 1.21. To find the probability of overlap, we have to modify the code above the line marked as 'this' to

```
Python Code

print(overlap.mean())

R Code

print(mean(overlap))
```

Change this in your code and explain what it does.

- **Ex 1.22.** Modify the code of the previous exercise so that you sample 1,000 times an appointment between Alice and Bob, and such that they only meet when 15 minutes apart. Include your code and show the output.
- **Ex 1.23.** Contrary to [1.14] suppose Bob is not prepared to wait longer than 5 minutes, while Alice is still prepared to wait for 15 minutes. Write down the integral that corresponds to the probability that they meet. Solve it by hand. Modify the above code such that you can evaluate it numerically. Include your code in the answers.

Ex 1.24. Contrary to [1.14] suppose Bob is prepared to wait for 20 minutes when he arrives before 1230, but just 10 when he arrives after 1230, while Alice is still prepared to wait for 15 minutes. Write down the integral that corresponds to the probability that they meet. Explain that this is the code to compute this.

```
Python Code

from scipy.integrate import dblquad

area1 = dblquad(lambda a, b: (b - 1 / 3.0 < a) * (a < b + 0.25), 0, 1, 0, 0.5)

area2 = dblquad(lambda a, b: (b - 1 / 6.0 < a) * (a < b + 0.25), 0, 1, 0.5, 1)

print(area1[0] + area2[0])

R Code

library(cubature)

area1 = adaptIntegrate(function(x){(x[2] - 1 / 3.0 < x[1]) * (x[1] < x[2] + 0.25)},

lowerLimit = c(0, 0), upperLimit = c(1, 0.5))

area2 = adaptIntegrate(function(x){(x[2] - 1 / 6.0 < x[1]) * (x[1] < x[2] + 0.25)},

lowerLimit = c(0, 0.5), upperLimit = c(1, 1))

print(area1$integral+ area2$integral)
```

Ex 1.25. Contrary to [1.14] suppose Bob's arrival time is $\text{Exp}(\lambda)$ with $\lambda = 1/2$, while Alice's arrival time is still uniform. Explain the code below.

```
Python Code
   import numpy as np
   from scipy.integrate import dblquad
   from scipy.stats import uniform, expon
   np.random.seed(3)
   labda = 1.0 / 2
   end = 5
   value = dblquad(
10
       lambda a, b: (b - 0.25 < a) * (a < b + 0.25) * np.exp(-labda * b), 0, 1, 0, end
11
   )
12
   print(value)
13
14
   area = value[0] * labda
   print(area)
   n = 100000
   alice = uniform(0, 1).rvs(n)
   bob = expon(scale=1 / labda).rvs(n)
   overlap = (bob - 0.25 < alice) * (alice < bob + 0.25)
   print(overlap.mean())
```

```
R Code
  library(cubature)
   set.seed(3)
   labda = 1.0/2
   end = 5
   value = adaptIntegrate(function(x){(x[2] - 0.25 < x[1]) * (x[1] < x[2] + 0.25) *
           \exp(-labda * x[2]), lowerLimit = c(0, 0), upperLimit = c(1, end))
   print(value)
9
10
   area = value$integral * labda
11
   print(area)
12
13
   n = 100000
14
   alice = runif(n, 0, 1)
   bob = rexp(n, labda)
   overlap = (bob - 0.25 < alice) * (alice < bob + 0.25)
   print(mean(overlap))
18
   print(mean(bob))
```

Ex 1.26. Changing the value of the variable end to 100 in the code above, and run it again. Then change it to 50, and finally to 1,000. When I do this, I get a very strange answers for 50 and 1,000. Try to explain what goes wrong.

Hopefully you got by now the point: modeling is an interesting, but quite difficult, activity. Once you have the model, the computer can help solve the problem.

You should memorize that integration in multiple dimensions quickly becomes unmanageable. In fact, most higher D integrals are approximated with simulation, like we do here in the simulation part.

You should also memorize that computers are extremely dumb, and that they can live with any answer that comes out of a computation. Here, since we compare a simulation with a direct integration, we see a difference, and that makes us suspicious about the correctness of both. Mind, we humans can do pretty smart things, and we should typically distrust the output of computers. (That is why I often include tests.)

1.4 BH.7.9

First read and solve BH.7.9. Here we provide a simulation.

Ex 1.27. Run this code, and use the documentation of scipy.stats.geom in Python (rgeom() function in R) to explain why it fails. Perhaps you can also use the wikipedia page on the geometric distribution. (Hint: is this the first success distribution or the geometric distribution?)

```
Python Code
   import numpy as np
   from scipy.stats import randint, geom
   np.random.seed(3)
   p, num = 0.1, 10000
   q = 1 - p
   rv = geom(p) # This line
   X = rv.rvs(num)
   if X.min() != 0:
11
       raise ValueError("The minimal value of X is not 0")
12
   if not np.isclose(q / p, X.mean(), 0.1):
       raise ValueError("The mean of X is not ok")
                                         R Code
   library(dplyr)
   set.seed(3)
   p = 0.1
   num = 100000
   q = 1 - p
   X = rgeom(num, p) + 1 #This line
   if (\min(X) != 0){
     stop("The minimal value of X is not 0")
10
   }
11
   if (!(near(q/p,mean(X),0.1))){
     stop("The mean of X is not ok")
   }
```

Ex 1.28. The code of the previous exercise failed, on purpose. It will work if we change the line marked as "this line", into this:

```
rv = geom(p, loc=-1)
R Code

X = rgeom(num, p)
```

Explain the difference, and why it now works. Relate this location-scale transformation to one of the distributions you know from BH.

Memorize from this exercise that with such checks you can build code that fails when it gets unexpected input.

Ex 1.29. Run this code and explain in detail the output. Make, and include, some simple example code to show that you are correct.

```
python Code

import numpy as np

N = np.array([3, 4, 4, 8, 4])
X = np.array([8, 7, 6, 5, 3])
print(np.argwhere(N == 4))
Xn = X[np.argwhere(N == 4)]
print(Xn)

R Code

N = c(3, 4, 4, 8, 4)
X = c(8, 7, 6, 5, 3)
print(which(N == 4))
Xn = X[which(N == 4)]
print(Xn)
```

Ex 1.30. Explain the lines marked with 'this line' in the code of Listing 1 or Listing 2. You already explained the other parts so you don't have to do that again.

Ex 1.31. Explain the results. Do they match with the theoretical results?

```
Python Code
   import numpy as np
   from scipy.stats import randint, geom
   import matplotlib.pyplot as plt
   np.random.seed(3)
   p, num = 0.1, 100000
   q = 1 - p
   rv = geom(p, loc=-1)
   X = rv.rvs(num)
10
   Y = rv.rvs(num)
12
   N = X + Y
13
   fig, axes = plt.subplots(3, 3, sharey="all", figsize=(6, 3)) # this
   n = 7 # gives nice results, after some experimentation
   for ax in axes.flatten(): # this
       Xn = X[np.argwhere(N == n)]
       y, bins = np.histogram(Xn, bins=np.linspace(0, n + 1, n + 2)) # this
19
       ax.plot(range(n + 1), y, 'bo', ms=3, label='N=8') # this
20
       ax.vlines(range(n + 1), 0, y, colors='b', lw=2, alpha=0.5) # this
21
       ax.set_title(f'{n=} ({len(Xn)})')
22
       ax.set_ylim(0) # this
23
       n += 1
24
```

Listing 1: BH.7.9, Python code.

```
R Code
   library(ggplot2)
   library(gridExtra)
   p = 0.1
   num = 100000
   q = 1 - p
   X = rgeom(num, p)
   Y = rgeom(num, p)
   N = X + Y
10
11
   pdf(file="bh-7-9.pdf", width = 8, height = 7,
12
       bg = "white", colormodel = "cmyk", paper = "A4")
13
14
   plots <-list()</pre>
15
   par(oma= c(0,0,0,0))
   i = 1
   n = 7 # gives nice results, after some experimentation
   while (i <= 9){  #this
     Xn = X[which(N == n)]
     df = data.frame(Xn)
     plots[[i]] = ggplot()+
23
                    geom_histogram(df, mapping = (aes(x = Xn)),
24
                    bins = n + 1, color = "blue")+ # this
25
                    ylim(0,600) + # this
26
                    labs(title=paste0("n=", n, ", (", length(Xn) ,")"))+
27
                    theme(axis.title.x=element_blank(),
                    axis.title.y = element_blank())
29
     n = n + 1
30
     i = i + 1
31
   }
   do.call("grid.arrange", c(plots, ncol=3))
   dev.off()
```

Listing 2: BH.7.9, R code.

2 ASSIGNMENT 2

2.1 BH.7.48

Read this exercise first and solve it. Then consider the code below.

```
Python Code
   import numpy as np
   np.random.seed(3)
   def find_number_of_maxima(X):
       num_max = 0
       M = -np.infty
       for x in X:
           if x > M:
10
               num\_max += 1
               M = x
       return num_max
   num = 10
16
   X = np.random.uniform(size=num)
   print(X)
18
19
   print(find_number_of_maxima(X))
20
21
   samples = 100
22
   Y = np.zeros(samples)
   for i in range(samples):
       X = np.random.uniform(size=num)
25
       Y[i] = find_number_of_maxima(X)
26
   print(Y.mean(), Y.var(), Y.std())
                                         R Code
   set.seed(3)
   find_number_of_maxima = function(X) {
     num_max = 0
     M = -Inf
     for (x in X) {
       if(x > M) {
         num_max = num_max + 1
```

```
M = x
       }
10
     }
11
     return(num_max)
12
   }
13
14
15
   num = 10
   X = runif(num, min = 0, max = 1)
   print(X)
   print(find_number_of_maxima(X))
20
21
   samples = 100
22
   Y = rep(0, samples)
   for (i in 1:samples) {
24
     X = runif(num, min = 0, max = 1)
     Y[i] = find_number_of_maxima(X)
26
   }
27
28
   print(mean(Y))
   print(var(Y))
   print(sd(Y))
```

Ex 2.1. Explain how the small function in lines P.6 to P.13 (R.4-R.12) works. (You should know that x += 1 is an extremely useful abbreviation of the code x = x + 1. Unfortunately, R does not have such increment operators.)

Ex 2.2. Explain the code in lines P.25 and P.26 (R.25, R.26).

2.2 BH.7.53

- Ex 2.3. Read and solve BH.7.53.
- **Ex 2.4.** The code in Listing 3 or Listing 4 simulates this exercise. Run the code. (To improve your understanding, just change some parameters here and there, and check the output.) Explain the lines marked with #*. (In particular, why do we do a runs number of runs to estimate the covariance?) Finally, explain the output.
- **Ex 2.5.** Modify the code such that the drunkard makes steps of size 2 when moving left or right; the stepsizes up or down remain the same. What happes with $E[R_n^2]$?
- **Ex 2.6.** Optional: Add a third dimension.

```
Python Code
   import numpy as np
   from scipy.stats import randint
   np.random.seed(3)
   num = 4
   steps = randint(0, 4).rvs(num) \# *
   print(steps)
   X = (2 * steps - 1) * (steps < 2) # *
   Y = (2 * steps - 5) * (steps >= 2) # *
   print(X)
12
   print(Y)
   R2 = X * X + Y * Y
   print(f"{num=}, {R2.sum()=}")
16
17
   def cov(X, Y):
       return (X * Y).mean() - X.mean() * Y.mean() # *
21
   num = 10
22
   runs = 300
   Ss = np.zeros(runs)
   Ts = np.zeros(runs)
25
   for i in range(runs): # *
26
       steps = randint(0, 4).rvs(num)
       X = (2 * steps - 1) * (steps < 2)
28
       Y = (2 * steps - 5) * (steps >= 2)
       Ss[i] = X.sum() # *
30
       Ts[i] = Y.sum() # *
   print(f"{cov(Ss,Ts)=}") # *
```

Listing 3: BH.7.53, python code.

2.3 BH 7.86

In line with Exercise BH.8.33, we are now going to analyze the effect on $P\{D \mid T\}$ when the sensitivity is not known exactly. So, why is this interesting? In Example 2.3.9 the sensitivity is given, but in fact, in 'real' experiments, this is not always known as accurately as assumed in this example. For example, in this paper: False-positive COVID-19 results: hidden problems and costs it is claimed that 'The current rate of operational false-positive swab tests in the UK is

```
R Code
   set.seed(3)
2
   num = 4
   steps = sample(0:4, num, replace = TRUE)
   print(steps)
  X = (2 * steps - 1) * (steps < 2) # *
   Y = (2 * steps - 5) * (steps >= 2) # *
   print(X)
   print(Y)
  R2 = X * X + Y * Y
   print(paste0("num= ", num, " R2.sum() = ", sum(R2) ))
11
12
   cov = function (X, Y){
13
     return (mean(X * Y) - mean(X) - mean(Y)) # *
14
   }
15
16
   num = 10
17
   runs = 300
   Ss = replicate(runs, 0) # *
   Ts = replicate(runs, 0)
   for (i in runs){
     steps = sample(0:4, num, replace = TRUE)
     X = (2 * steps - 1) * (steps < 2)
23
     Y = (2 * steps - 5) * (steps >= 2)
24
     Ss[i] = sum(X) # *
25
     Ts[i] = sum(Y) # *
26
   }
27
28
   print(paste0("cov(Ss, Ts) = ", cov(Ss, Ts)))
```

Listing 4: BH.7.53, R code.

unknown; preliminary estimates show it could be somewhere between 0.8% and 4.0%.'. Hence, even though it is claimed that PCR tests 'have analytical sensitivity and specificity of greater than 95%', it may be 4% lower. Simply put, the specificity and sensitivity are not precisely known, hence this must affect $P\{D \mid T\}$.

To help you, we show how to make one graph. Then we ask you to make a few on your own, and comment on them.

2.4 BH.2.3.9

I write $p_D_g T$ for $P\{D \mid T\}$. Here is how this can be implemented in python.

```
sensitivity = 0.95

specificity = 0.95

p_D = 0.01

p_T = sensitivity * p_D + (1-specificity)*(1-p_D)

p_D_g_T = sensitivity * p_D/p_T

p_D_g_T

sensitivity = 0.95

sensitivity = 0.95

p_D = 0.01

p_T = sensitivity * p_D + (1-specificity)*(1-p_D)

p_D_g_T = sensitivity * p_D + (1-specificity)*(1-p_D)

p_D_g_T = sensitivity * p_D + (1-specificity)*(1-p_D)

p_D_g_T = sensitivity * p_D/p_T

p_D_g_T
```

- 1. Make a plot of $P\{D \mid T\}$ in which you vary the sensitivity from 0.9 to 0.99. Explain what you see.
- 2. Make a plot of $P\{D \mid T\}$ in which you vary the specificity from 0.9 to 0.99.
- 3. Make a plot of $P\{D \mid T\}$ in which you vary $P\{D\}$ from 0.01 to 0.5. Explain what you see.

3 ASSIGNMENT 3

3.1 BH.8.4 Figure

The challenge is to make Figure BH.8.4 for a specific case.

We throw a coin a number of times. We take some beta distribution as prior for the probability p that Heads appears. Then we update the PDF of p according to outcomes of the throws.

Ex 3.1. Run the follow code and explain the output.

We define the variables H and T as these outcomes are easier to memorize than 1 and 0. Like this, there is less possibility for confusion, hence less bugs.

```
Python Code
H, T = 1, 0
outcomes = [H, T]
h, t = 1, 1 # this
h += outcomes[0] == H
t += outcomes[0] == T
print(f"{h=}, {t=}")
h += outcomes[1] == H
t \leftarrow outcomes[1] == T
print(f"{h=}, {t=}")
                                      R Code
H = 1
T = 0
outcomes = c(H,T)
h = 1 #this
t = 1 #this
h = h + as.numeric(outcomes[1] == H)
t = t + as.numeric(outcomes[1]== T)
print(paste0("h=", h, ", t=", t))
h = h + as.numeric(outcomes[2] == H)
t = t + as.numeric(outcomes[2]== T)
print(paste0("h=", h, ", t=", t))
```

Ex 3.2. Explain the code below. Why does the line marked as 'this' relate to the prior distribution? What is that prior distribution?

```
Python Code
   import numpy as np
   from scipy.stats import beta
   import matplotlib.pyplot as plt
   support = np.arange(0, 1, 0.03)
   H, T = 1, 0
   outcomes = [H, H, T, H, H, H, T, H, H]
   h, t = 1, 1 # this
10
   fig, axes = plt.subplots(3, 3, sharey="all", figsize=(6, 3)) # this
11
   for i, ax in enumerate(axes.flatten()): # this
12
       h += outcomes[i] == H
13
       t += outcomes[i] == T
       X = beta(h, t).pdf(support)
       ax.plot(support, X)
       ax.set_title(f'h={h-1}, t={t-1}')
   plt.tight_layout()
   plt.savefig("beta.pdf")
                                         R Code
  library(ggplot2)
   library(gridExtra)
   support = seq(0, 1, 0.03)
   H = 1
   outcomes = c(H, H, T, H, H, H, T, H, H)
   h = 1 #this
   t = 1 #this
10
11
   pdf(file="beta.pdf", width = 8, height = 7,
12
       bg = "white", colormodel = "cmyk", paper = "A4")
13
14
   plots <-list()
   par(oma= c(0,0,0,0))
16
17
   i = 1
   while (i <= length(outcomes)){</pre>
                                      #this
     h = h + (outcomes[i] == H)
     t = t + (outcomes[i] == T)
     X = dbeta(support, h, t)
```

```
df = data.frame(support,X)
23
     plots[[i]] = ggplot(data=df, aes(x=support, y=X, group=1)) +
24
                   geom_line()+
25
                   ylim(0,3.5)+
26
                   labs(title=paste0("h=", h-1, ", t=", t-1),x=" ", y = " ")
     i = i+1
28
   }
29
   do.call("grid.arrange", c(plots, ncol=3))
32
   dev.off()
```

- **Ex 3.3.** What do the PDFs actually model? Explain why the PDFs behave as they do. For instance, why does the location of the peak move from right to left to right?
- **Ex 3.4.** Which conjugacy relation between prior and data do we use here to update the prior to the posterior?
- **Ex 3.5.** Adapt the code above such that the prior is a = 3.5, b = 1.5, the outcomes are T, H, T, H, T, H, and the output is a figure with 3 x 2 subfigures. Besides the explanation, include the figure.

3.2 BH-8-4-3

Let us investigate BH.8.4.3. The rvs X and Y are independent, $X \sim \text{Exp}(\lambda)$ and $Y \sim \text{Exp}(\mu)$. We want to compare X + Y to $\Gamma 2$, μ and $\Gamma 2$, λ , just to see what happens.

Ex 3.6. Run the code below. Explain the lines marked as 'this'. In particular, why do we simulate outcomes for *X* and *Y*? Then, why are the shape and scale variables as they are? Then explain the three graphs.

```
import numpy as np
from scipy.stats import expon, gamma
import matplotlib.pyplot as plt

labda, mu = 3, 5

n = 1000
X = expon(scale=1 / labda)
Y = expon(scale=1 / mu)
Z = X.rvs(n) + Y.rvs(n)
G1 = gamma(2, scale=1 / mu)
G2 = gamma(2, scale=1 / mu)
G2 = gamma(2, scale=1 / mu)
```

```
support = np.arange(0, 5, 0.03)
14
15
   plt.hist(Z, label="Z", density=True)
16
   plt.plot(support, G1.pdf(support), label="G1")
   plt.plot(support, G2.pdf(support), label="G2")
   plt.legend()
   plt.tight_layout()
   plt.savefig("figures/gamma1.pdf")
                                         R Code
   library(ggplot2)
   labda = 3
   mu = 5
   n = 1000
   X = rexp(n, labda)
   Y = rexp(n, mu)
   Z = X + Y
   support = seq(0, 5, 0.03)
   G1 = dgamma(support, 2, scale=1 / labda)
13
   G2 = dgamma(support, 2, scale=1 / mu)
14
15
   df1 = data.frame(Z)
16
   df2 = data.frame(support,G1, G2)
17
18
   pdf(file="figures/gamma1.pdf", width = 8, height = 7,
19
       bg = "white", colormodel = "cmyk", paper = "A4")
20
21
   ggplot()+
22
   geom_histogram(df1, mapping = (aes(x = Z, y = ..density.., fill = "Z")), bins = 20)+
   geom\_line(df2, mapping = (aes(x = support, y = G1, color="G1")))+
   geom\_line(df2, mapping = (aes(x = support, y = G2, color = "G2")))+
   scale_fill_manual(name = " ", values = c("Z" = "blue")) +
   scale_color_manual(name = " ", values = c( "G1" = "orange", "G2" = "green"))+
   theme(axis.title.x=element_blank(),axis.title.y = element_blank())
28
   dev.off()
```

Ex 3.7. Modify the above code such that X and Y are both $Exp(\lambda)$. Then make a histogram of a simulation of X + Y and compare it to the PDF of the correct gamma distribution. Discuss your result, and include the figure of course.

3.3 BH.8.1

Ex 3.8. Here is the code to solve BH.8.1. Explain the relevant steps, i.e., the lines marked as 'this'. Is the result according to your expectation?

```
Python Code
   import numpy as np
   from scipy.stats import expon
   import matplotlib.pyplot as plt
   X = expon(scale=1) # this
   Y = np.exp(-X.rvs(1000)) # this
   plt.hist(Y, density=True, label="sim")
  x = np.linspace(0, 1, 30)
   dxdy = np.exp(x) # this
10
  pdf_y = X.pdf(x) * dxdy # this
11
   plt.plot(x, pdf_y, label="f_Y")
  plt.legend()
13
  plt.tight_layout()
14
  plt.savefig("figures/bh-8-1-fig.pdf")
                                         R Code
   library(ggplot2)
  X = rexp(1000, 1) #this
   Y = \exp(-X) \#this
   df1 = data.frame(Y)
   x = seq(0,1,length.out = 30)
   dxdy = exp(x) #this
   pdf_y = dexp(x,1) * dxdy #this
   df2 = data.frame(x,pdf_y)
10
11
   pdf(file="figures/bh-8-1-fig.pdf", width = 8, height = 7,
12
       bg = "white", colormodel = "cmyk", paper = "A4")
13
14
   ggplot()+
   geom_histogram(df1, mapping = (aes(x = Y, y = ..density..., fill = "sim")), bins = 20)+
   geom_line(df2, mapping =(aes(x=x,y=pdf_y, color = "f_y")))+
   scale_fill_manual(name = " ", values = c("sim" = "blue")) +
   scale_color_manual(name = " ", values = c( "f_y" = "orange"))+
19
   theme(axis.title.x=element_blank(),axis.title.y = element_blank())
20
   dev.off()
```

Ex 3.9. What error would we make if we would use this code: dxdy = np.exp(-x)(dxdy = exp(-x))?

3.4 BH.8.18

In fact, I know that the volumes of an airplane and a ping pong ball is an estimate, rather than a precise number as assumed above. It seems to be better to approximate V and v as rvs. Let's assume that

$$V \sim N(2500, 500^2), \nu \sim N(33.5, 0.5^2),$$

where the variances express my trust in my guess work. What is now the mean of N = V/v and its std? Can we get the CDF? From BH.8.18 you know that finding the closed form expression for the distribution of N is not entirely simple. However, with simulation it's easy to get an estimates for the mean and the standard deviation.

Ex 3.10. Explain lines marked as 'this'.

```
Python Code
   import numpy as np
   from scipy.stats import norm
   num = 500
   np.random.seed(3)
6
   V = norm(loc=2500, scale=500) # this
   v = norm(loc=33.5, scale=0.5) # this
10
   print(V.mean(), V.std()) # just a check
11
12
   N = V.rvs(num) / v.rvs(num) # this
   print(N.mean(), N.std())
   print(2500/33.5)
   print(np.sqrt(500*0.5))
```

⁴ For reasons incomprehensible to me, even professional econometricians sometimes report results with 10 digits or more, without questioning the precision.

```
num <- 500

2
3  set.seed(3)

4
5  V = rnorm(num, 2500, 500) # this
6  v = rnorm(num, 33.5, 0.5) # this
7
8  N = V / v # this
9  print(paste(mean(N), sd(N)))</pre>
```

Ex 3.11. Contrary to BH.7.1.25 if you run the code above, you'll see that $E[N] < \infty$, and is very near to the deterministic answer. But isn't this strange: we divide two normal random variables, just like BH.7.1.25, but there the expectation is infinite. Explain why it works for the Beluga case, but not for the case discussed in BH.7.1.25.

Ex 3.12. Include code to make a histogram of the PDF of N.

Remark 3.13. Here is a final, and optional, question. The numerical results suggest the interesting guess $V[N] \approx V[V] * V[v]$, but is this true more generally?

4 ASSIGNMENT 4

4.1 BH.8.27

We start from BH.8.27 (which you have to read now). We are interested in the difference between the distribution of X + Y + Z and the normal distribution. But why the normal distribution? As it turns out, the central limit law, see BH.10, states that the distribution of sums of r.v.s converge to the normal distribution (in a specific sense).

Here is some code to simulate.

```
Python Code
   import numpy as np
   from scipy.stats import norm
   import matplotlib.pylab as plt
   import seaborn as sns
   sns.set()
   np.random.seed(3)
10
11
   Zexact = norm(loc=k / 2, scale=np.sqrt(k / 12))
12
  X = np.arange(0, 3, 0.1)
13
14
  XYZ = np.random.uniform(size=(4000, k))
   # print(XYZ) # if you want to see it.
   Z = XYZ.sum(axis=1)
   sns.distplot(Z)
  plt.plot(X, Zexact.pdf(X))
   plt.show()
```

```
R Code
set.seed(3)

k = 3
    X <- seq(0, 3, by = 0.1)
    Zexact <- dnorm(X, mean = k / 2, sd = sqrt(k / 12))

XYZ <- matrix(NA, 4000, k)
    # print(XYZ) # if you want to see it.
for (i in 1:k) {
    XYZ[,i] <- runif(4000, min = 0, max = 1)
}</pre>
```

```
13 Z <- rowSums(XYZ)
14
15 par()
16 hist(Z, prob = TRUE, breaks = 31)
17 lines(X, Zexact, type = "l", col = "orange")
18 lines(density(Z), col = "blue")</pre>
```

- **Ex 4.1.** What is the shape of XYZ in the code above, i.e., how many rows and columns does it have? If you don't know, run the code, and print it.
- **Ex 4.2.** What is the shape (rows and columns) of Z?
- **Ex 4.3.** Explain the values for loc and shape in Zexact (and corresponding mean and sd in R). (Read the documentation of scipy.stats.norm (or norm() for R) on the web if necessary.) To which definition in BH does this loc-scale transformation relate?
- **Ex 4.4.** Change the seed to your student id, or any other number you like, run the code, and include the graph produced by your simulation. Explain what you see.

Now we do an exact computation.

```
Python Code
   import numpy as np
   from scipy.stats import norm
   import matplotlib.pylab as plt
   import seaborn as sns
   sns.set()
   N = 200
   x = np.linspace(0, 2, 2 * N)
  fx = np.ones(N) / N
   f2 = np.convolve(fx, fx)
   f3 = np.convolve(f2, fx)
13
14
   k = 3
15
16
   x = np.linspace(0, k, len(f3))
   Zexact = norm(loc=k / 2, scale=np.sqrt(k / 12))
19
   plt.plot(x, N * f3, label="conv")
21
   plt.plot(x, Zexact.pdf(x))
   plt.legend()
   plt.show()
```

```
R Code
   N = 200
   x = seq(0, 2, length.out = 2 * N)
   fx = rep(1, N) / N
   f2 = convolve(fx, fx, type = "open")
   f3 = convolve(f2, fx, type = "open")
   k = 3
   x = seq(0, k, length.out = length(f3))
   Zexact = dnorm(x, mean = k/2, sd = sqrt(k / 12))
10
11
  par()
12
   plot(x, N * f3, col = "blue", type = "l", ylim = c(0, 0.8))
   lines(x, Zexact, type = "l", col = "orange")
   legend("topright", legend = "conv", bty = "n",
         lwd = 2, cex = 1.2, col = "blue", lty = 1)
```

Ex 4.5. Read the documentation of np.convolve (or convolve() in R). Why is it called like this?

Ex 4.6. In the code, what is f2?

Ex 4.7. What is f3?

Ex 4.8. Why do we set k=3?

Ex 4.9. A bit harder, why do we plot N*f3, i.e., why do we have to multiply with N? Relate this to the meaning of f(x) dx, where f the density of some random variable. (Think hard about this problem; the ideas are important to understand how numerical integration works.) Suppose we chop up the area under some arbitrary function g in blocks of height g(x) and length Δx . Then the area of such a block is $g(x)\Delta x$.

In our case, we chop up the interval in parts with length $\Delta x = 1/N$. The elements of f3 are such that $f3[i] = f_3(x_i)\Delta x$, where x_i lies in the ith interval and f_3 is the density of the sum of the three r.v.s. But then, $f_3(x_i) = f3[i]/\Delta x = Nf3[i]$.

So, why should we scale with $\Delta x = 1/N$? (Forgetting this step is a common error when dealing with densities.)

To memorize, f(x) dx is the area of a block of height f(x) and length dx.

Ex 4.10. Yet a tiny bit harder, consider f4 = np.convolve(f3, fx) (f4 = convolve(f3, fx, type = "open") in R) and <math>g4 = np.convolve(f2, f2) (g4 = convolve(f2, f2, type = "open") in R). Why are they, numerically speaking, equal?

Ex 4.11. When you would compute the maximum of np.abs (f4 - g4) (abs(f4 - g4) in R) you would see that this is about 10^{-10} , or so. Hence, a small number. This is not equal to 0, but we know that this is due to rounding effects.

How can we use the function np.isclose() to get around this problem? (You should memorize from this question that you should take care when testing on whether floating point numbers are the same or not.)

4.2 BH.8.36

We simulate the post office part of BH.8.36. Read it now, i.e., before reading the text below. Then read the code below. In the questions, we ask you to explain what the code does. There are lots of print statements that have been commented out, but we left them in for you to include while experimenting with the code to see how the code works. (I often use print statements of intermediate results when writing a program, just to see whether I am still on track. Once I checked, I remove them, because they clutter the looks of the code.)

```
Python Code
   import numpy as np
   from scipy.stats import expon
   np.random.seed(3)
   labdas = np.array([3, 4])
   N = 10
   T1 = expon(scale=1 / labdas[0]).rvs(N)
   # print(T1.mean())
   T2 = expon(scale=1 / labdas[1]).rvs(N)
   T = np.zeros([2, N])
13
   T[0, :] = T1
14
   T[1, :] = T2
15
   # print(T)
16
   server = np.argmin(T, axis=0)
17
   # print(server)
18
19
   # BH.8.36.b
   print((server == 1).mean())
21
   # BH.8.36.c
   # print(labdas[server])
   S = expon(scale=1).rvs(N) / labdas[server] # this
   # print(S)
26
27
   D = np.min(T, axis=0) + S
   # print(D)
29
30
   print(D.mean(), D.std())
31
                                          R Code
   set.seed(3)
```

```
labdas = 3:4
   N = 10
  T1 = rexp(N, rate = labdas[1])
  T2 = rexp(N, rate = labdas[2])
   #print(mean(T1)
   bigT = matrix(0, nrow = 2, ncol = N)
   bigT[1,] = T1
   bigT[2,] = T2
   #print(T)
   server = rep(0, ncol(bigT))
16
17
   for (i in 1:ncol(bigT)) {
18
     server[i] = which.min(bigT[,i])
19
20
   #print(server)
21
   #BH.8.36.b
   print(mean((server == 1)))
25
  #BH.8.36.c
   #print(labdas[server])
   S = rexp(N, rate = 1) / labdas[server] #this
   #print(S)
   D = apply(bigT, MARGIN = 2, FUN = min) + S
31
   #print(D)
32
33
   print(mean(D))
   print(sd(D))
```

- **Ex 4.12.** Explain why T1 corresponds to a number of service times of the first server.
- **Ex 4.13.** To what do the rows of T (bigT) correspond?
- **Ex 4.14.** What is the content of server? Why do we compute this?
- **Ex 4.15.** Explain how we use the fundamental bridge in line P.21 (R.24) to answer BH.8.36.b.
- **Ex 4.16.** Alice is taken into service by the server that finishes first. We need to simulate the service time that Alice needs at that server. Explain how we do this in line P.25 (R.28). Hint, reread BH.5.5 on the exponential. BTW, this is a good time to reread BH.5.3.

Ex 4.17. Why is np.min(T, axis=0) in the Python code (apply(bigT, MARGIN=2, FUN=min) in R) the time Alice spends waiting in queue, i.e., the time Alice spends in the post office before her service starts?

Ex 4.18. Why is D the departure time of Alice, i.e., the time Alice spends in the post office?

Ex 4.19. Set N to 1000 or so, or any other large number to your liking, but not so large that your computer will keep simulating for a month...Compare the values of the simulation to the theoretical result that you have to compute in BH.8.36.c.

Ex 4.20. Run the code for $\lambda_1 = 1$ fixed, and take λ_2 equal to 0.5, 1, 1.5 and 2 successively. Compute the mean time waiting time and mean sojourn time of Alice, and put your results in a table. Compare the results of the simulation to the theoretical values.

There is an important lesson to learn here. With simulation it is, often, reasonably simple to get numerical answers, but it requires many simulations to see a pattern in the numbers. For patterns, we can better use theory, as theory gives us formulas that show how the output of some model depends on the input.

4.3 BH.8.54

Read and solve BH.8.54. Perhaps you should read the hints in the study guide on this problem too. First we consider the case with r = 1, and we need to tackle some technical details.

Ex 4.21. Just as in the problem, let Y = pX/q. Why is E[Y] = 1?

Ex 4.22. If $Z \sim \text{Exp}(1)$, then explain that $M_Z(s) = 1/(1-s)$ for s < 1.

Ex 4.23. Use the mathematical solution of the problem to explain that $M_Y(s) \to M_Z(s)$ as $p \to 0$.

Ex 4.24. Here is code to compare $M_Y(s)$ to $M_Z(s)$. Explain what the function M(s) does. Then explain why in the line marked as 'this' we multiply with $1 - s_i$. Explain why we limit s to the interval [0,0.8]. Finally, explain why the graph goes through the point x = 0, y = 1.

```
import numpy as np
from scipy.stats import geom, gamma
import matplotlib.pyplot as plt

np.random.seed(3)

p = 0.1
q = 1 - p
n = 100000
```

```
X = geom(p).rvs(n)
13
       Y = p * X / q
14
       return np.exp(s * Y).mean()
15
16
17
   num = 30
18
   m = np.zeros(num)
   s = np.linspace(0, 0.8, num)
   for i in range(num):
       m[i] = M(s[i]) * (1 - s[i]) # this
22
23
   plt.plot(s, m, label="MY")
24
   plt.tight_layout()
25
   plt.savefig("figures/moments.pdf")
                                          R Code
   set.seed(3)
   p = 0.1
   q = 1 - p
   n = 100000
   M = function(s){
       X = rgeom(n,p) - 1
       Y = p * X / q
        return (mean((exp(s * Y))))
10
   }
11
12
13
   num = 30
14
   m = rep(0, num)
15
   s = seq(0,0.8, length.out = num)
   for (i in 1:num){
       m[i] = M(s[i]) * (1 - s[i]) # this
19
   }
20
   pdf(file="figures/moments.pdf",
21
       width = 8, height = 7, bg = "white",
22
       colormodel = "cmyk", paper = "A4")
23
24
   plot(s,m, type = "l", ann = FALSE, col = "blue")
25
26
   dev.off()
27
```

def M(s):

12

Ex 4.25. The match is not terrric when p = 0.1. Rerun the code, but now with p = 1/1000. Which line to you have to change for this? Include your graph. Is the match better now?

Ex 4.26. The problem asks for the general *r*. What lines of code have to change to deal with general *r*? Modify the code appropriately, and explain why your code is correct.

4.4 BH.9.1

Here is the code. (BTW, for this assignment, you don't have to read all of BH.9, reading the problem suffices. Hence, it is not a typo that this assignment appears in week 4, not week 5.)

```
Python Code
   import numpy as np
   from scipy.stats import randint, norm
   mus = np.array([2, 5, 8])
   stds = np.array([0.1, 0.5, 1])
   m2 = stds ** 2 + mus ** 2 # this
   v = m2.mean() - (mus.mean()) ** 2 # this
   print(mus.mean(), v)
10
11
   routes = []
12
   for i in range(3):
13
       routes.append(norm(mus[i], stds[i])) # this
14
15
16
   num = 200
17
   options = randint(0, 3).rvs(num) # this
   times = np.zeros(num) # this
   for i in range(num):
21
       times[i] = routes[options[i]].rvs() # this
22
23
   print(times.mean(), times.var())
                                          R Code
   mus = c(2, 5, 8)
   stds = c(0.1, 0.5, 1)
   m2 = stds ** 2 + mus ** 2 #this
   v = mean(m2) - (mean(mus)) ** 2 #this
   print(paste(mean(mus), v))
```

```
num = 200
poptions = sample(1:3, num, replace = TRUE) #this

times = rep(0, num)
for(i in 1:num){
   times[i] = rnorm(1, mus[options[i]], stds[options[i]]) # this
}

print(paste(mean(times), var(times)))
```

- Ex 4.27. Explain the lines marked as 'this', i.e., how does the program work?
- **Ex 4.28.** Run the code and explain the results. Are the simulated results in line with the theory?
- **Ex 4.29.** Modify the code so that you can plot the histogram of the times vector. Include your code and a figure. Set also a seed equal to, e.g., the day of the month in which you do the assignment.
- **Ex 4.30.** Set μ_3 (the third element in mus) to 80 instead of 8. Rerun the code and explain what happens to the mean and the variance. Then change to 800 and include a histogram of the PDF of the times.
- **Ex 4.31.** Change σ_1 from 0.1 to 10. Run the code and describe what you get. What's wrong with such a large standard deviation? Are you sure that the route lengths remain positive?

5 ASSIGNMENT 5

5.1 BH.9.1.7

The aim is to use a simulator to analyze the expected profit for the game of BH.9.1.7 but for a different prior than the uniform. First, however, we build BH.9.1.7 with the uniform prior.

First read and solve BH.9.1.7.

Ex 5.1. Run the code below. What are the values of accepted? What does it represent? How is the gain computed? Why do we take the mean?

```
Python Code
   import numpy as np
   from scipy.stats import uniform
   np.random.seed(3)
   alpha = 2 / 3
   n = 100
   V = uniform(0, 1).rvs(n)
  bid = 0.5
10
  accepted = bid > alpha * V
   gain = (V - bid) * accepted
   print(gain.mean())
                                        R Code
   set.seed(3)
2
   alpha = 2 / 3
   n = 100
   V = runif(n, 0, 1)
  bid = 0.5
  accepted = bid > alpha * V
   gain = (V - bid) * accepted
   print(mean(gain))
```

Ex 5.2. Here is some code to compute the gain for multiple bids. Explain how the function f works. In other words, how does it handles the formatting of floating point numbers? The rest of the code must be clear to you, so you don't have to explain that.

Printing pretty floating points is something you really have to learn. If you unquote the line marked as 'unquote this line' and rerun the program, you'll see why people like to see formatted numbers.

Hint, search the web (Python or R documentation) on formatting floating point numbers.

```
Python Code
   import numpy as np
   from scipy.stats import uniform
   def f(x): # format floating point numbers
       # return f"{x}" # unquote this line
       return f"{x:3.3f}"
   np.random.seed(3)
10
11
12
   alpha = 2 / 3
13
   n = 100
   V = uniform(0, 1).rvs(n)
   for b in np.arange(0.1, 1, 0.1):
       accepted = b > alpha * V
18
       gain = (V - b) * accepted
19
       print(f"{f(b)}, {f(gain.mean())}")
20
                                         R Code
   f = function(x) # format floating point numbers
       # return(x) # unquote this line
       return(signif(x, digits = 3))
3
   set.seed(3)
   alpha = 2 / 3
   n = 100
   V = runif(n, 0, 1)
   for (b in seq(0.1, 0.9, 0.1)){}
11
       accepted = b > alpha * V
12
       gain = (V - b) * accepted
13
       print(paste0(f(b), ", " ,f(mean(gain))))
14
   }
15
```

- **Ex 5.3.** Use the code of the previous exercise as a starting point to make a graph of the gain as a function of the bid. Compute the bids for 0.05, 0.1, 0.15, etc, i.e., in steps of 0.05. Include your code, a figure, and comment on the graph. Use your student id as seed.
- **Ex 5.4.** It is very easy to compute the expected gain by means of numerical integration. Run the code below and explain how the res variable is computed. Is the result similar to the

simulation? Once again (so that you really remember what an anonymous function is!), explain very briefly why we use an anonymous function here.

```
Python Code
import numpy as np
from scipy.integrate import quad
alpha = 2 / 3
n = 100
b = 0.5
res = quad(lambda v: (v - b) * (b > alpha * v), 0, 1)
print(res)
                                     R Code
library(cubature)
alpha = 2 / 3
n = 100
b = 0.5
res = adaptIntegrate(function(v){(v - b) * (b > alpha * v)},
      lowerLimit = c(0), upperLimit = c(1))
print(paste(res$integral, res$error))
```

Ex 5.5. Modify the numerical integrator such that the prior on V is Beta(8,2). Include your code and give the result for a bid of 1/2.

5.2 BH.9.1.9

This is a nice variation of one-step analysis of the coin throwing example of BH.9.1.9.

We have a mouse that sits on one corner of a cube whose edges are made of wire, and diagonally opposite the mouse there is some cheese. The mouse chooses at random (uniform) any edge and moves to the next corner. How many steps does it take, on average, for the mouse to reach the cheese (assuming the cheese does not move)?

For ease, we label all corners with the minimal number of edges that need to be traversed from the corner to reach the cheese into states. Thus, the starting corner will be called state 3, then we have a state 2, and 1, and finally state 0 is the corner with the cheese.

Ex 5.6. Explain in the code below how the while loop simulates the moves of the mouse. To understand the line marked as 'this', read what is a, so-called, ternary operator.

```
Python Code ______
import numpy as np
from scipy.stats import uniform
```

```
import matplotlib.pyplot as plt
   move = uniform()
   def do_run():
       state = 3
       n_steps = 0
10
       while state != 0:
           n_steps += 1
           p = move.rvs()
           if state == 3:
                state = 2
           elif state == 2:
16
                state = 1 if p < 2 / 3 else 3 # this
17
            elif state == 1:
18
                state = 2 if p < 2 / 3 else 0
19
       return n_steps
20
21
   n = 200
23
   steps = np.zeros(n)
   for i in range(n):
       steps[i] = do_run()
   print(steps.mean(), steps.std())
28
   plt.hist(steps, density=True)
   plt.tight_layout()
   plt.savefig("figures/mouse.pdf")
```

R Code

do_run = function(){
 state = 3
 n_steps = 0

state = ifelse(p < 2 / 3, 1, 3) # this

```
while (state != 0) {
    n_steps = n_steps + 1
    p = runif(1, 0, 1)
    if (state == 3) {
        state = 2
    }
    else if (state == 2) {
```

library(ggplot2)

13

14

}

```
else if (state == 1){
15
         state = ifelse(p < 2 / 3, 3, 0) # this
16
       }
17
18
     return (n_steps)
19
   }
20
21
   n = 200
   steps = rep(0, n)
24
   for (i in 1:n){
     steps[i] = do_run()
   }
27
28
   print(paste(mean(steps),sd(steps)))
29
30
   pdf(file="figures/mouse.pdf",
31
       width = 8, height = 7, bg = "white",
32
       colormodel = "cmyk", paper = "A4")
33
34
   df1 = data.frame(steps)
35
   ggplot()+
   geom_histogram(df1, mapping = (aes(x = steps, y = ..density.., fill = "steps")))+
   scale_fill_manual(name = " ", values = c("steps" = "blue"))
   theme(axis.title.x=element_blank(),axis.title.y = element_blank())
   dev.off()
```

5.3 BH.9.6.1

Read BH.9.6.1 first. We will redo the computations and check the results with simulation. We will also cover many other topics as we proceed.

Ex 5.7. We need to model the distribution of an individual demand X. As an example, we can use the beta distribution. In the example, μ and σ are assumed as given, but for the beta distribution we need an a and a b. One way is to solve by hand for a and b for given μ and σ . However, we can also be lazy and let the computer do the work. Here is the code to do that.

Explain how this code works. (You can use the relevant documentation on the web on how the function fsolve works.) Why do I use $\mu = 1/2$ and $\sigma^2 = 1/12$ as test case? Are the a and b that comes out of the solver correct? What does the final check do?

```
from scipy.stats import beta
from scipy.optimize import fsolve
```

```
mu, var = 1 / 2, 1 / 12 # why this test case?
5
   def func(x):
       a, b = x[0], x[1]
       m = a / (a + b)
       v = m * (1 - m) / (a + b + 1)
10
       return [m - mu, v - var]
11
   a, b = fsolve(func, [2, 2])
   print(a, b)
   X = beta(a, b)
   print(X.stats(moments='mv')) # final check.
                                         R Code
   library(pracma)
   mu = 1 / 2
   var = 1 / 12 # why this test case?
   func = function(x){
     a = x[1]
     b = x[2]
     m = a / (a + b)
     v = m * (1 - m) / (a + b + 1)
10
     return (c(m - mu, v - var))
11
   }
12
13
   output = fsolve(func, c(2,2))$x
14
   a = output[1]
15
   b = output[2]
   print(paste(a, b))
17
  X_{mean} = a / (a + b)
  X_{var} = (X_{mean} * (1-X_{mean})) / (a + b + 1)
   print(paste(X_mean, X_var)) # final check
```

Ex 5.8. What do you expect to get when you would choose $\mu = 2$? If it fails, why is that?

Ex 5.9. What do you expect to get when you would choose $\sigma > 1/2$? If it fails, why is that? Hint, when an rv has a support on [0,1], why is the variance at most 1/4?

Remark 5.10. When setting up simulations and choosing variables, make sure that the variables you choose actually make any sense.

Now we know how to find a and b for given μ and σ , we henceforth assume that a=1 and b=2, as this is easier to work with.

Ex 5.11. Now that we know how to simulate individual demands, we can compute the mean and the variance with the formulas of the book and compare this to simulation.

Explain the lines marked as 'this' of the code. (The rest must be clear for you. If not, then you have to be seriously concerned about your understanding of programming.)

```
Python Code
   import numpy as np
   from scipy.stats import poisson, beta
   from scipy.optimize import fsolve
   np.random.seed(3)
   labda = 3
   N = poisson(labda)
   # print(N.rvs(100).mean()) # why this line?
   X = beta(a=1, b=2)
   mu, var = X.stats(moments='mv')
12
13
   days = 1000
14
   demands = np.zeros(days)
15
   for i in range(days):
16
       demands[i] = X.rvs(N.rvs()).sum() # this
17
18
   print(demands.mean(), demands.var())
19
   print(labda * mu, var * labda + mu * mu * labda) # this
                                          R Code
   set.seed(3)
   labda = 3
   # print(mean(rpois(100, labda))) # why this line?
   days = 1000
   demands = rep(0, days)
   a = 1
   b = 2
   mu = a / (a + b)
   var = (mu * (1 - mu)) / (a + b + 1)
12
13
   for (i in 1:days){
14
     demands[i] = sum(rbeta(rpois(1,labda), 1, 2)) # this
```

```
16  }
17
18  print(paste(mean(demands), var(demands)))
19  print(paste(labda * mu, var * labda + mu * mu * labda)) # this
```

Ex 5.12. Here is the code to compute the mean and variance with Adam and Eve's laws. Explain how it works. In particular:

- 1. explain the lines marked as 'this';
- 2. what are the contents of the demand matrix;
- 3. what is the P;
- 4. how are EX, EV and VE computed;
- 5. Note in particular the difference between EXn and EXN, and similarly for the variance. Use BH.9.2.1 to explain why I use a small n and a large N in the naming of the variables.

Remark: I realize that this is a hard piece of code, but your understanding of Eve's law will increase by a lot. Some subtle points about the code that you *don't* have to explain:

- 1. Suppose the largest amount of customers (jobs) that we saw during the simulation is 7. Then on any day we can have 0,...,7 demands, i.e, 8 possible outcomes. Thus, per row in the demands matrix we need to have 8 columns.
- 2. The numpy function (mean() function in R) that computes the mean gives a warning when it is applied to an empty array. (Just try it to see how it works.) We prevent this by computing the histogram (i.e., the empirical PDF) for n = 0 separately. The same reasoning applies to the case when idx has zero length.

```
Python Code
   import numpy as np
   from scipy.stats import poisson, beta, rv_discrete
   np.random.seed(3)
   labda = 3
   N = poisson(labda)
   X = beta(a=1, b=2)
   mu, var = X.stats(moments='mv')
11
   days = 1000
12
   n_{jobs} = N.rvs(days)
13
   n_max = n_jobs.max() + 1 # see the comments
14
   demands = np.zeros((days, n_max))
```

```
for day in range(days):
16
       n = n_{jobs}[day]
17
       demands[day, :n] = X.rvs(n) # this
18
19
20
   P = np.zeros(n_max)
21
   EXn = np.zeros(n_max)
   VXn = np.zeros(n_max)
   P[0] = np.argwhere(n_jobs == 0).size / days
   for n in range(1, n_max):
       idx = np.argwhere(n_jobs == n)
       if idx.size == 0:
27
           continue
       P[n] = idx.size / days # this
       EXn[n] = demands[idx, :n].mean() * n # this
30
       VXn[n] = demands[idx, :n].var() * n # this
31
32
   EXN = rv_discrete(values=(EXn, P)) # this
33
34
   print("mean:", labda * mu, EXN.mean())
   VXN = rv_discrete(values=(VXn, P)) # this
   print("EV: ", var * labda, VXN.mean())
   VE = EXN.var()
   print("VE: ", mu * mu * labda, VE)
                                         R Code
   library(discreteRV)
   set.seed(3)
   labda = 3
   a = 1
   b = 2
   mu = a / (a + b)
   var = (mu * (1 - mu)) / (a + b + 1)
10
   days = 1000
11
   n_jobs = rpois(days,labda)
12
   n_max = max(n_jobs) + 1 # see the comments
13
   demands = matrix(0,days, n_max) # this
14
   for (day in 1:days){
15
     n = n_{jobs}[day]
16
     if(n == 0) next
17
```

```
demands[day, 1:n] = rbeta(n, 1, 2) # this
18
   }
19
20
   P = rep(0, n_max)
21
   EXn = rep(0, n_max)
   VXn = rep(0, n_max)
   P[1] = length(n_jobs[which(n_jobs == 0)])/days
   for(n in 1:n_max){
     idx = which(n_jobs == n - 1)
     if (length(idx) == 0){
       next
     }
     P[n] = length(idx) / days # this
31
     EXn[n] = mean(demands[idx, c(1:n)]) * (n - 1) #this
     VXn[n] = sd(demands[idx, c(1:n)]) * sd(demands[idx, c(1:n)]) * (n - 1) #this
33
   }
34
   EXN = RV(outcomes = EXn, probs = P) # this
   print(paste("mean:", labda * mu, E(EXN) ))
   VXN = RV(outcomes = VXn, probs = P) # this
   print(paste("EV: ", var * labda, E(VXN)))
   VE = V(EXN)
   print(paste("VE: ", mu * mu * labda, VE))
```

5.4 BH.9.7

We use simulation to solve BH.9.7. Read it now, i.e., before reading the text below, then read the code below. Note how short this code is; amazing, isn't it?

```
import numpy as np
from scipy.stats import uniform
import matplotlib.pyplot as plt

np.random.seed(3)

N = 1000
a, b = 0, 1000000
V = uniform(a, b).rvs(N)

x_range = np.linspace(b / 5, b / 2, num=50)
```

```
y = np.zeros_like(x_range)
13
14
   for i, b in enumerate(x_range):
15
       payoff = (V - b) * (b >= V / 4)
16
       y[i] = payoff.mean()
17
18
   plt.plot(x_range, y)
   plt.show()
                                           R Code
   set.seed(3)
   N = 1000
   a = 0
   b = 1000000
   V = runif(N, min = a, max = b)
   x_range = seq(b / 5, b / 2, length.out = 50)
   y = rep(0, length(x_range))
10
11
   i = 1
12
   for (b in x_range) {
```

- **Ex 5.13.** For the Python code, use the scipy documentation to explain why $V \sim \text{Unif}([0, 10^6])$. For R, explain the function runif().
- **Ex 5.14.** What are the smallest and the largest value of x_range ?

payoff = (V - b) * (b >= V / 4)

plot(x_range, y, type = "l", col = "blue")

y[i] = mean(payoff)

i = i + 1

16

17 }

- **Ex 5.15.** Run the code above and make a graph. Include the graph in your report, and explain what you see in the graph. For instance, is there a maximum? If so, can you explain where the maximum occurs? Can you explain how the maximum should be?
- **Ex 5.16.** Suppose after seeing the graph of the payoffs, and this graph would only increase, or decrease, how would you change x_range ? Do you expect to see a maximum?
- **Ex 5.17.** For small N, e.g. N=10, you can get quite strange values. Why is that?
- **Ex 5.18.** Change the acceptance threshold from to V/4 to V/5 (or V/6, or some other value you like), and make a graph of the payoffs. Include the graph in your report.

Ex 5.19. Change the payoff function to e.g $\sqrt{V-b}$, or some weird function that you like particularly such as $\sin |V-b|$ (any non-trivial function goes). Make a graph of the mean and the standard deviation of the payoff. Can you explain your graph?

6 ASSIGNMENT 6

x[, 1] = 100 # this

6.1 BH.9.25

This simulation exercise is based on BH.9.25. Please read the exercise first, and then the code below.

Ex 6.1. How does this code work? Use the print statements to find out what *S* is, and explain why we multiply by 2 and subract 1 in one of the lines marked as 'this'.

```
Python Code
   import numpy as np
   from scipy.stats import bernoulli
   np.random.seed(3)
   n = 5
   num = 10
   p = 0.5
   S = bernoulli(p).rvs([num, n]) * 2 - 1 # this
   # print(S)
12
  x = np.zeros([num, n])
   x[:, 0] = 100 # this
   # print(x)
   f = 0.25
   for i in range(1, n):
       x[:, i] = x[:, i - 1] * (1 + f * S[:, i]) # this
19
20
   print(x.mean(axis=0), x.std(axis=0))
                                        R Code
   set.seed(3)
  n = 5
   num = 10
   p = 0.5
   S = matrix(rbinom(n * num, size = 1, prob = p), num, n) * 2 - 1 # this
   # print(S)
  x = matrix(0, num, n)
```

```
#print(x)
f = 0.25

for (i in 2:n) {
    x[, i] = x[, i - 1] * (1 + f * S[, i]) # this
}

print(colMeans(x))
print(apply(x, MARGIN = 2, sd))
```

Ex 6.2. 1. Run the code below.

- 2. Check the values of x. What do you notice?
- 3. Then compare both figures. You should see that in Kelly-no-jitter.pdf many of the markes (the x below the y-axis) fall on top of each other. To repair for this, we add some noise (jitter) to the x and y values of the markes. Explain how randomness is used here.
- 4. Jitter plots, like this one, are meant to give some more insight into the data to check whether we use the right number of bins. Change the number of bins to 2, include the plot with the jitters, and explain how the marks help to see that two bins is much too small.

```
Python Code
   import numpy as np
   from scipy.stats import bernoulli, uniform
   import matplotlib.pyplot as plt
   np.random.seed(3)
   n = 10
   n_{experiments} = 50
10
   S = bernoulli(p).rvs([n, n_experiments]) * 2 - 1 # this
12
   # print(S)
13
   x = np.zeros([n, n_experiments])
   x[0, :] = 100 # this
   f = 0.25
   for i in range(1, n):
19
       x[i, :] = x[i - 1, :] * (1 + f * S[i, :]) # this
20
21
   print(x[-1, :].mean(), x[-1, :].std())
```

```
print(x[-1, :])
23
24
   plt.hist(x[-1, :], bins=20, density=True)
25
   plt.plot(x[-1, :], np.full\_like(x[-1, :], -0.01), "xk", markeredgewidth=1)
   plt.savefig("figures/Kelly-no-jitter.pdf")
28
   x_{i} = x_{i} + uniform(loc=-3, scale=6).rvs(size=n_experiments)
   y_jitter = np.full_like(x[-1, :], -0.01) + uniform(loc=-0.01, scale=0.02).rvs(
       size=n_experiments
32
   )
   plt.plot(x_jitter, y_jitter, "xk", markeredgewidth=1)
   plt.savefig("figures/Kelly-jitter.pdf")
36
                                         R Code
   library(ggplot2)
   set.seed(3)
3
   n = 10
   n_{experiments} = 50
   p = 0.4
   S = matrix(0, n, n_experiments)
   for(i in 1:n){
     S[i,1:n\_experiments] = rbinom(n\_experiments, 1, p) * 2 -1 # this
   }
   #print(S)
13
   x = matrix(0, n, n_experiments)
   x[1,] = 100
16
   f = 0.25
17
18
   for(i in 2:n){
19
     x[i,] = x[i-1,] * (1 + f * S[i,]) # this
20
   }
21
   pdf(file="figures/Kelly-no-jitter.pdf",
       width = 8, height = 7, bg = "white",
       colormodel = "cmyk", paper = "A4")
   df1 = data.frame(x[nrow(x),])
   colnames(df1) = c("X")
   df1\$Y = -0.01
30
```

```
p = qqplot()+
31
     geom\_histogram(df1, mapping = (aes(x = X, y = ..density..)), bins = 20)+
32
     geom_point(dfl, mapping = (aes(x=X, y = Y)))+
33
     theme(axis.title.x=element_blank(),axis.title.y = element_blank())
34
   р
35
36
   dev.off()
   pdf(file="figures/Kelly-jitter.pdf",
       width = 8, height = 7, bg = "white",
       colormodel = "cmyk", paper = "A4")
41
   x_{jitter} = x[nrow(x),] + runif(n_{experiments,-3, 3)
43
   df2 = data.frame(x_jitter)
44
   colnames(df2) = c("x_jitter")
45
   df2\$y_jitter = -0.01 + runif(n_experiments, -0.01, 0.01)
46
47
   p + geom\_histogram(df2, mapping = (aes(x = x_jitter, y = ..density..)), bins = 20)+
48
     geom_point(df2, mapping =(aes(x=x_jitter, y = y_jitter)))+
49
     theme(axis.title.x=element_blank(),axis.title.y = element_blank())
   р
51
   dev.off()
```

6.2 BH.9.37

With this exercise we illustrate why (and how) to use bootstrapping. At first it appears a bit strange: we have a set of observations (samples) $X = \{X_1, X_2, ..., X_n\}$. For this we can compute the mean, standard deviation, and the empirical distribution with standard procedures, so why sample (bootstrap) from X?

One type of question that seems hard to answer by classical means is to characterize the median, in particular the standard deviation and distribution of the median, of the population (that is, the population from which we took the sample X).

Bootstrapping is nowadays used a lot in data science.

Ex 6.3. Run the code below, and then explain what is sample, bootstrap and medians. Why do we use axis=1?

Ex 6.4. Change some numbers (such as the range of the random numbers from which we obtained the initial sample). Include your code, make the plot, and include it in your assignment.

```
import numpy as np
import matplotlib.pyplot as plt
```

```
np.random.seed(3)
   sample = np.random.randint(200, size=100)
   n_boot = 5000
   bootstrap = np.random.choice(sample, replace=True, size=(n_boot, len(sample)))
   medians = np.median(bootstrap, axis=1)
10
   # print(medians.mean(), medians.std())
11
   # print(np.percentile(medians, [2.5, 97.5]))
13
   plt.hist(medians, bins=100, density=True)
   plt.axvline(np.percentile(medians, 2.5), c='r', lw=2)
   plt.axvline(medians.mean(), c='r', lw=2)
   plt.axvline(np.percentile(medians, 97.5), c='r', lw=2)
   plt.savefig("figures/bootstrap.pdf")
                                         R Code -
   library(ggplot2)
   library(matrixStats)
2
3
   np.random.seed(3)
4
   sample = sample(1:200,100, replace=TRUE)
   n_boot = 5000
   bootstrap = matrix(0,n_boot,length(sample))
   for(i in 1:nrow(bootstrap)){
     bootstrap[i,] = sample(sample, length(sample), replace = TRUE)
   }
11
12
   medians = rowMedians(bootstrap)
13
   #print(paste(mean(medians), sd(medians)))
14
   \#print(quantile(medians, c(.025, 0.975)))
15
16
   pdf(file="figures/bootstrap.pdf",
17
       width = 8, height = 7, bg = "white",
18
       colormodel = "cmyk", paper = "A4")
19
   df1 = data.frame(medians)
   ggplot()+
   geom_histogram(df1, mapping = (aes(x = medians, y = ..density..)), bins = 100)+
   geom_vline(xintercept = quantile(medians, c(0.025)), color = "red", size=2) +
   geom_vline(xintercept = quantile(medians, c(0.9725)), color = "red", size=2)+
   geom_vline(xintercept = mean(medians), color = "red", size=2)
26
   dev.off()
28
```

Ex 6.5. Finally, let us compare the information we obtained from bootstrap to the medians we would obtain when we would sample from the real population.

- 1. What is here the real population?
- 2. In a real life situation, what is less costly: real sampling or bootstrapping?
- 3. Explain the ideas of the code below.
- 4. Compare this figure to the one obtained by bootstrapping. Comment on major differences.

```
Python Code
   import numpy as np
   import matplotlib.pyplot as plt
   np.random.seed(3)
   samples = np.random.randint(200, size=(500, 100))
   medians = np.median(samples, axis=1)
  plt.hist(medians, bins=100, density=True)
   plt.savefig("figures/real_medians.pdf")
                                         R Code
   library(ggplot2)
   library(matrixStats)
   set.seed(3)
4
   sample = matrix(0, 500, 100)
   for(i in 1:nrow(sample)){
     sample[i,] = sample(1:200, 100, replace = TRUE)
   }
   medians = rowMedians(bootstrap)
10
11
   pdf(file="figures/real_medians.pdf",
12
       width = 8, height = 7, bg = "white",
13
       colormodel = "cmyk", paper = "A4")
14
15
   df1 = data.frame(medians)
16
   ggplot()+
17
   geom_histogram(df1, mapping = (aes(x = medians, y = ..density..)), bins = 100)
18
   dev.off()
```

6.3 BH.9.50

This simulation exercise is based on BH.9.50. Please read the exercise first. Do the exercise while thinking about the code.

Ex 6.6. Here is some code to simulate N. Explain the non-trivial steps. (By now you should have an idea about boiler-plate code and what code is specific for the problem.)

```
Python Code
   import numpy as np
   from scipy.stats import poisson, expon
   np.random.seed(3)
   num = 100
   N = np.zeros(num)
   Labda = expon(scale=1)
   labdas = Labda.rvs(num)
   for i in range(num):
10
       labda = labdas[i]
11
       N[i] = poisson(labda).rvs()
12
13
   print(N.mean(), N.var())
14
   adam = Labda.mean()
  eve = Labda.mean() + Labda.var()
   print(adam, eve)
                                         R Code
   library(ggplot2)
   library(matrixStats)
   set.seed(3)
   num = 100
   N = rep(0, num)
   labdas = rexp(num, 1)
   for(i in 1:num){
10
     labda = labdas[i]
11
     N[i] = rpois(1, labda)
   }
13
  adam = 1/1
   eve = 1/1 + 1/(1 * 1)
   print(paste(adam, eve))
```

Ex 6.7. Run the above code and compare the result to the theoretical value. What do you observe? Rerun the code with num = 10000. Does this result in a better estimate? What do you learn from this?

Ex 6.8. The next step is to estimate the mean and variance of the total claim size *S*. Here is the code, run it and explain the relevant parts.

```
Python Code
   import numpy as np
   from scipy.stats import poisson, expon, lognorm
   np.random.seed(3)
   num = 100
  Labda = expon(scale=1)
   labdas = Labda.rvs(num)
   mu, sigma = 3, 1
   X = lognorm(loc=mu, s=sigma)
   S = np.zeros(num)
   for i in range(num):
       N = poisson(labdas[i]).rvs()
13
       S[i] = X.rvs(N).sum()
14
   print(f"{S.mean()=}")
   adam = X.mean() * Labda.mean()
   print("mean_theoretical: ", adam)
   print(f"{S.var()=}")
19
   eve = Labda.mean() * X.var() + X.mean() ** 2 * (Labda.mean() + Labda.var())
20
   print("var_theoretical: ", eve)
                                         R Code
   set.seed(3)
  num = 100
   labdas = rexp(num, 1)
   mu = 3
   sigma = 1
   S = rep(0, num)
   for(i in 1:num){
     N = rpois(1,labdas[i])
     S[i] = sum(rlnorm(N, mean = 0, sd = sigma) + mu)
11
12
   }
13
  theta = exp(sigma * sigma / 2)
14
   var = (theta * theta) * (exp(sigma * sigma) - 1)
```

```
print(paste("S.mean()=", mean(S)))
adam = (theta + mu) * (1/1)
print(paste("mean_theoretical: ", adam))
print(paste("S.var()=", var(S)))
eve = (1/1) * var + (theta + mu) ** 2 * (1/1 + 1/1/1)
print(paste("var_theoretical: ", eve))
```

Ex 6.9. Here is the code to plot the distribution. Change λ such that it is $\sim \text{Exp}(1/2)$ and include the figure.

```
Python Code
   import numpy as np
   from scipy.stats import poisson, expon
   import matplotlib.pyplot as plt
   np.random.seed(3)
   num = 100
   N = np.zeros(num)
   Labda = expon(scale=1)
   labdas = Labda.rvs(num)
   for i in range(num):
11
       labda = labdas[i]
12
       N[i] = poisson(labda).rvs()
13
14
   plt.hist(N, density=True)
15
   plt.savefig("figures/bh-9.50-fig.pdf")
                                          R Code
   set.seed(3)
   num = 100
   N = rep(0, num)
   labdas = rexp(num, 1)
   for(i in 1:num){
     labda = labdas[i]
     N[i] = rpois(1, labda)
8
   }
10
   pdf(file="figures/bh-9.50-fig.pdf",
11
       width = 8, height = 7, bg = "white",
12
       colormodel = "cmyk", paper = "A4")
   df1 = data.frame(N)
```

```
16  ggplot()+
17  geom_histogram(df1, mapping = (aes(x = N, y = ..density..)))
18
19  dev.off()
```

Ex 6.10. Finally, let's do the integration numerically to compute $P\{N=3\}$. Explain the code. Why, in particular, do I compare the result against $geom(1/2) \cdot pmf(4)$? (Hint, read the documentation of scipy.stats.geom(dgeom() function in R) to understand why.)

```
Python Code
   import numpy as np
   from scipy.integrate import quad
   from scipy.stats import expon, geom
   Labda = expon(scale=1)
   def integrand(labda):
       return np.exp(-labda) * labda ** 3 / 6 * Labda.pdf(labda)
   res = quad(lambda labda: integrand(labda), 0, np.infty)
10
   print(res)
11
   print(geom(1 / 2).pmf(3))
12
   print(geom(1 / 2).pmf(4))
                                         R Code -
   library(cubature)
   integrand = function(labda){
     return (exp(-labda) * labda ** 3 / 6 * exp(-labda) )
   }
   res = adaptIntegrate(function(x){integrand(x[1])}, lowerLimit = c(\theta), upperLimit = c(Inf))
   print(res)
   print(dgeom(3 - 1, 1 / 2))
   print(dgeom(4 - 1, 1 / 2))
```

6.4 BH.10.2.3

Let us try to understand the weak law of large numbers by means of simulation. An easy example is to take $X_i \sim \text{Unif}(0,1)$, so that is what we do here.

```
Python Code ______
import numpy as np
from numpy.random import uniform
```

```
np.random.seed(3)
4
   n = 10
6
   N = 50 \# num samples
   mu = 1 / 2
   var = 1 / 12
   eps = 0.1
   X = uniform(size=[num_samples, n])
   Y = X.mean(axis=1)
15
16
   larger = np.abs(Y - mu) > eps
17
   count = larger.sum()
18
   P = count / N
19
  RHS = var / (n * eps * eps)
20
   print(P, RHS)
                                         R Code
   set.seed(3)
   n = 10
   N = 50 \# num samples
   mu = 1 / 2
   var = 1 / 12
   eps = 0.1
   X = matrix(runif(N * n), N, n)
10
11
   Y = rowMeans(X)
12
13
  larger = abs(Y - mu) > eps
   count = sum(larger)
  P = count / N
  RHS = var / (n * eps * eps)
   print(P)
   print(RHS)
```

Ex 6.11. Explain mu and var.

Ex 6.12. What is Y? What is the symbol that BH use for this?

- **Ex 6.13.** What are the meanings of larger and count?
- **Ex 6.14.** What is RHS in the notation of BH?
- **Ex 6.15.** What inequality of BH do we check by printing RHS and P?
- **Ex 6.16.** Choose some different values for n and the sample size N. Is the inequality always true?

7 ASSIGNMENT 7

7.1 BH.10.9

Ex 7.1. In the code below we compute or simulate the various expressions of BH.10.9. Check the results and compare it to the analytical results. Are both results in line? If not, do you know why that might be the case? What can you do about that?

I don't know how to build E[V[Y|X]] in any sensible way. Since Y is independent of X, the conditioning does not have any influence. Thus, this numerical exercise is skipped.

```
Python Code
   import numpy as np
   from scipy.stats import expon
   np.random.seed(3)
   X = expon(2)
   Y = expon(2)
   print(np.exp(-X.mean()), X.expect(lambda x: np.exp(-x)))
10
   n_sample = 1000
11
   X_sample = X.rvs(n_sample)
12
   Y_sample = Y.rvs(n_sample)
13
   p1 = sum(X_sample > Y_sample + 3) / n_sample
   p2 = sum(Y_sample > X_sample + 3) / n_sample
   print(p1, p2)
   p3 = sum(X_sample > Y_sample - 3) / n_sample
   print(p1, p3)
   EXY = X_sample @ Y_sample / n_sample
21
   EX4 = X.expect(lambda x: x ** 4)
22
   print(EX4, EXY * EXY)
23
24
   p = sum(np.abs(X_sample + Y_sample) > 3) / n_sample
   print(p, X.mean())
                                         R Code
  library(cubature)
   set.seed(3)
   expectation = adaptIntegrate(function(x)\{exp(-(x[1]+2)) * exp(-x[1]) \},
                 lowerLimit = c(0), upperLimit = c(Inf))
   print(paste(exp(-mean(rexp(10000,1)+2)), expectation$integral))
```

```
n_sample = 1000
  X_{sample} = rexp(n_{sample}, 1) + 2
   Y_sample = rexp(n_sample, 1) + 2
   p1 = sum(X_sample > Y_sample + 3) / n_sample
   p2 = sum(Y_sample > X_sample + 3) / n_sample
   print(paste(p1, p2))
   p3 = sum(X_sample > Y_sample - 3) / n_sample
   print(paste(p1, p3))
   EXY = X_sample %*% Y_sample / n_sample
   expectation2 = adaptIntegrate(function(x)\{((x[1]+2) ** 4) * exp(-x[1]) \},
19
                 lowerLimit = c(0), upperLimit = c(Inf))
20
21
   print(paste(expectation2$integral, EXY * EXY))
22
23
   p = sum(abs(X_sample + Y_sample) > 3) / n_sample
24
   print(paste(p, (2) + 1))
```

- **Ex 7.2.** Why do I use simulation to estimate E[XY]? Recall an earlier assignment in which we used numerical integration. How successful was that?
- **Ex 7.3.** What would you do to increase the credibility of the claims? (Note these claims are based on case checking and simulation? There is no actual proof, however, they can act as counter-examples.)

7.2 BH.10.28

Ex 7.4. Let us first plot the PMF of the standardized version of X_n for n = 10. Explain the code below BEFORE making the plot.

```
import numpy as np
from scipy.stats import poisson, norm
import matplotlib.pyplot as plt

n = 10
X = poisson(n)
N = norm()

pmf = np.zeros(2 * n)
for i in range(len(pmf)):
    pmf[i] = X.pmf(i)
supp = (np.linspace(0, len(pmf), num=len(pmf)) - n) / np.sqrt(n)
```

```
13
14
   plt.plot(supp, pmf) # this
15
   plt.plot(supp, norm.pdf(supp))
   plt.savefig("figures/bh-10-28.pdf")
                                          R Code
   library(ggplot2)
   n = 10
   pmf = rep(0, 2 * n)
   for(i in 1:length(pmf)){
     pmf[i] = dpois(i, n)
8
   supp = (seq(0, length(pmf), length.out = length(pmf)) - n) / sqrt(n)
10
   pdf(file="figures/bh-10-28.pdf",
11
       width = 8, height = 7, bg = "white",
12
       colormodel = "cmyk", paper = "A4")
13
   df1 = data.frame(supp)
   df1$pmf = pmf
   df1$norm_pdf = dnorm(supp,0,1)
17
18
   p = ggplot(df1) +
19
       geom_line(mapping =(aes(x=supp, y = pmf)), color = "blue")+ # this
20
       geom_line(mapping =(aes(x=supp, y = norm_pdf)), color = "orange")+
21
       theme(axis.title.x=element_blank(),axis.title.y = element_blank())
22
   р
23
24
   dev.off()
```

Ex 7.5. Now make the plot and explain what is wrong.

Ex 7.6. The line marked as 'this' should be replaced with

```
plt.plot(supp, pmf * np.sqrt(n))

R Code

df1$pmf = pmf * sqrt(n)
```

Explain why. (You should memorize that comparing PMFs and PDFs is not straightforward.)

- **Ex 7.7.** Make the plot of n = 2 and n = 20. Include both plots and explain the results.
- **Ex 7.8.** The code below computes $M_Y(s)$ where Y is the standardized version of X_n . Then it compares M_Y to the MGF of the standard normal distribution.
 - 1. Explain how it works.

```
Python Code
   import numpy as np
   from scipy.stats import poisson, norm
   import matplotlib.pyplot as plt
   n = 10
   X = poisson(n)
   num = 50
   S = np.linspace(-1, 1, num)
   M = np.zeros(num)
   pmf = X.pmf(range(100)) # this
10
11
   for i in range(num):
12
       for j in range(len(pmf)):
13
           M[i] += np.exp(S[i] * (j - n) / np.sqrt(n)) * pmf[j]
14
15
16
   plt.plot(S, M, label="M X")
17
   plt.plot(S, np.exp(S * S / 2), label="M N")
   plt.xlabel("s")
   plt.legend()
   plt.savefig("figures/bh-10-28-mgf1.pdf")
                                          R Code
   library(ggplot2)
   n = 10
   num = 50
   S = seq(-1, 1, length.out = num)
   M = rep(0, num)
   pmf = dpois(1:100,n) #this
   for(i in 1:num){
     for(j in 1:length(pmf)){
       M[i] = M[i] + exp(S[i] * (j - n)/sqrt(n)) * pmf[j]
11
     }
12
   }
13
14
   pdf(file="figures/bh-10-28-mgf1.pdf",
```

```
width = 8, height = 7, bg = "white",
16
       colormodel = "cmyk", paper = "A4")
17
18
   df1 = data.frame(S)
19
   df1$M_X = M
   df1$M_N = exp(S * S / 2)
   p = ggplot(df1) +
       geom_line(mapping =(aes(x=S, y = M_X, color = "M X")))+
       geom_line(mapping =(aes(x=S, y = M_N, color = M_N))+
25
       scale_color_manual(name = " ", values = c("M X" = "blue", "M N" = "orange"))+
       theme(axis.title.y = element_blank())
   р
28
29
   dev.off()
```

Ex 7.9. The line marked as 'this' computes the PMF upfront. Why is that a good idea? Why do we stop the computation of the PMF at 100?

Ex 7.10. In the code below, one of the for loops is removed.

- 1. What is the change?
- 2. Will this change have an effect on the speed of the computation?
- 3. Which of the two alternatives do you find more readable?

```
Python Code
   import numpy as np
   from scipy.stats import poisson, norm
   import matplotlib.pyplot as plt
   n = 10
   X = poisson(n)
   num = 50
   S = np.linspace(-1, 1, num)
   M = np.ones(num)
10
   for i in range(num):
       M[i] = X.expect(lambda j: np.exp(S[i] * (j - n) / np.sqrt(n)))
12
   plt.plot(S, M, label="M X")
   plt.plot(S, np.exp(S * S / 2), label="M N")
   plt.xlabel("s")
   plt.legend()
   plt.savefig("figures/bh-10-28-mgf2.pdf")
```

```
R Code
   library(ggplot2)
   n = 10
   num = 50
   S = seq(-1, 1, length.out = num)
   M = rep(1, num)
   for(i in 1:num){
     j = 1:10000
     X = \exp(S[i] * (j-n) / \operatorname{sqrt}(n))
10
     M[i] = weighted.mean(X,dpois(j,n))
11
   }
12
13
   pdf(file="figures/bh-10-28-mgf2.pdf",
14
       width = 8, height = 7, bg = "white",
15
       colormodel = "cmyk", paper = "A4")
   df1 = data.frame(S)
18
   df1$M_X = M
   df1$M_N = exp(S * S / 2)
20
21
   p = ggplot(df1) +
22
       geom_line(mapping =(aes(x=S, y = M_X, color = "M X")))+
23
       geom_line(mapping =(aes(x=S, y = M_N, color = M_N)))+
24
       scale_color_manual(name = " ", values = c("M X" = "blue", "M N" = "orange"))+
       theme(axis.title.y = element_blank())
26
   р
   dev.off()
```

Ex 7.11. Change n to 20 and to 50. Does the quality improve? Choose some value for n that you like, and include the graph.

7.3 BH.10.30

Ex 7.12. Explain how the code below implements the solution of parts a and b BH.10.30. (Of course we skip the proof of the convergence to $g(\alpha)$ here.)

```
python Code
import numpy as np
from scipy.stats import bernoulli
import matplotlib.pyplot as plt

np.random.seed(3)
```

```
6
   num = 100
   p, alpha = 0.5, 0.5
   revenue = 0.5 + 1.2 * bernoulli(p).rvs(num)
10
   Y = np.ones(num)
11
   Y[0] = 1
12
   for i in range(1, num):
       Y[i] = (1 - alpha + alpha * revenue[i]) * Y[i - 1]
   print(np.log(Y[-1]) / num)
16
17
   # plt.plot(Y) # this
18
19
   def g(alpha):
20
       return 0.5 * np.log((1 + 0.7 * alpha) * (1 - 0.5 * alpha))
21
22
23
   plt.ylim(-10 * g(alpha), 10 * g(alpha)) # this
24
   plt.axhline(g(alpha), c='r', lw=2, label="g")
   plt.plot(np.log(Y)[1:] / range(1, num), label="(log Yn)/n")
   plt.legend()
   plt.savefig("figures/bh-10-30.pdf")
                                          R Code
   library(cubature)
   set.seed(3)
   num = 100
   p = 0.5
   alpha = 0.5
   revenue = 0.5 + 1.2 * rbinom(num, 1, p)
   Y = rep(1, num)
   Y[1] = 1
   for(i in 2:num){
     Y[i] = (1 - alpha + (alpha * revenue[i])) * Y[i - 1]
12
   }
13
14
   print(log(Y[length(Y)]) / num)
15
16
   # plot(Y, type = "l")
17
18
   g = function(alpha){
19
     return(0.5*log((1 + 0.7 * alpha)*(1 - 0.5 *alpha)))
20
```

```
}
21
22
   pdf(file="figures/bh-10-30.pdf",
23
       width = 8, height = 7, bg = "white",
24
       colormodel = "cmyk", paper = "A4")
25
26
   df = data.frame(log(Y)/(1:num))
   colnames(df) = c("lnY")
   p = ggplot()+
       geom_hline(mapping = aes(yintercept = g(alpha), color = "g"), size = 2)+
       geom\_line(df, mapping=aes(x = 1:num, y = lnY, color = '(log Yn)/n')) +
       scale_color_manual( name = " ", values = c("g" = "red", "(log Yn)/n" = "blue"))+
       ylim(-10* g(alpha), 10* g(alpha)) + #this
33
       theme(axis.title.x = element_blank(), axis.title.y = element_blank())
34
   р
35
36
   dev.off()
```

- **Ex 7.13.** Uncomment the relevant line to obtain a plot of Y_n as a function n. Make the plot and include it in your assignment.
- **Ex 7.14.** Take num much larger, e.g., 100 000. What is $(\log Y_n)/n$ now?
- **Ex 7.15.** Print $(\log Y_n)/n$ for various values of α . For instance, take $\alpha = 1/10, 2/10, 3/10, 4/10$.
- **Ex 7.16.** If you like a simple challenge, include a plot of $(\log Y_n)/n$ as a function of α . However, skip this if you don't have time, or interest in this extension of the problem. (Perhaps, if you find programming hard, you should do it to improve your skills :-))

7.4 BH.10.39

BH.10.39.a asks about the first time such that some exponential rv exceeds a certain threshold. Part b is about when the sum of a number of r.v.s exceed a threshold. Such problems are called *hitting times*.

```
import numpy as np
from scipy.stats import expon

np.random.seed(3)

X = expon(scale=1)

# part a

9
```

```
N = 10
   res = np.zeros(N)
   for i in range(N):
12
        n = 0
13
       while X.rvs() < 1:</pre>
14
            n += 1
15
        res[i] = n
16
   print(res.mean(), res.std())
   # part b
   N = 100
   res = np.zeros(N)
23
   for i in range(N):
24
       M = X.rvs()
25
       n = 1
26
        while M < 10:
27
            M += X.rvs()
28
            n += 1
        res[i] = n
30
   print(res.mean(), res.std())
                                           R Code
   set.seed(3)
2
   # part a
3
   N = 10
   res = rep(0, N)
   for (i in 1:N) {
     n = 0
     while (rexp(1, rate = 1) < 1) {</pre>
       n = n + 1
        res[i] = n
     }
12
   }
13
14
   print(mean(res))
15
   print(sd(res))
16
17
   # part b
18
19
```

N = 100

```
res = rep(0, N)
for (i in 1:N) {
    M = rexp(1, rate = 1)
    n = 1
    while(M < 10) {
    M = M + rexp(1, rate = 1)
    n = n + 1
    res[i] = n
}
res[i] = n
}
rrint(mean(res))
rrint(sd(res))</pre>
```

By now you have so much experience with reading code that you must be able to explain all of the code without intermediate steps to guide you.

- Ex 7.17. Explain how the first part of the code simulates BH.10.39.a.
- **Ex 7.18.** Explain how the second part of the code simulates BH.10.39.b.