

# Probability distributions EBP038A05

## Assignments

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## INTRODUCTION

Here we just provide the exercises of the assignments. For information with respect to grading we refer to the course manual.

Each assignment contains code in Python and R that shows how to implement an example or the solution of an exercises of BH. You have to run the code, read the output, and explain how the code works.

We include Python and R code, and leave the choice to you what to use. In the exam we will also include both languages in the same problem, so you can stay with the language you like. You should know, however, that many of you will need to learn multiple languages later in life. For instance, when you have to access databases to obtain data about customers, patients, clients, suppliers, inventory, demand, lifetimes (whatever), you often have to use sql. Once you have the raw data, you process it with R or Python to do statistics or make plots.<sup>1</sup> For your interest, based on the statistics [here](#) or [here](#), Python scores (much) higher than R in popularity; if you opt for a business career, the probability you have to use Python is simply higher than to have to use R.

You should become familiar with looking up documentation on coding on the web, no matter your programming language of choice. Invest time in understanding the, at times, rather technical and terse, explanations. Once you are used to it, the core documentation is faster to read, i.e., less clutter. In the long run, it pays off.

The rules:

1. For each assignment, you have to turn in a pdf document typeset in  $\text{\LaTeX}$ . Include a title, group number, student names and ids, and date.
2. We expect brief answers, just a few sentences, or a number plus some short explanation. The idea of the assignment is to help you study and improve your coding skills by showing good code, not to turn you in a writer.
3. When you have to turn in a graph, provide decent labels and a legend, ensure the axes have labels too.

---

<sup>1</sup> (While I (= NvF) worked at a bank, I used Fortran for numerical work, awk for string parsing and making tables, excel, SAS to access the database, and matlab for other numerical work, all next to each other. I got extremely tired of this, so I went to using Python as this can do all of this stuff, but within one language.

## 1 ASSIGNMENT 1

1.1 *Why is the Exponential Distribution so important?*

At the Paris metro, a train arrives every 3 minutes on a platform. Suppose that 50 people arrive between the departure of a train and an arrival. It seems entirely reasonable to me to model the arrival times of the individual people as distributed on the interval  $[0, 3]$ . What is the distribution of the inter-arrival times of these people? It turns out to be exponential!

You might want to compare your final result to Figure BH.13.1 (It is not forbidden to read the book beyond what you have to do for this course!). In this exercise, we use simulation to see that clustering of arrival times.

## Python Code

```

1 import numpy as np
2
3 np.random.seed(3)
4
5
6 num = 5 # small sample at first, for checking.
7 start, end = 0, 3
8 labda = num / (end - start) # per minute
9 print(1 / labda)
10
11 A = np.sort(np.random.uniform(start, end, size=num))
12 print(A) #this
13 print(A[1:])
14 print(A[:-1])
15 X = A[1:] - A[:-1]
16 print(X)
17
18 print(X.mean(), X.std())

```

## R Code

```

1 set.seed(3)
2
3
4 num = 5
5 start = 0
6 end = 3
7 labda = num / (end - start)
8 print(1 / labda)
9
10 A = sort(runif(num, min = start, max = end))
11 print(A) #this
12 print(A[-1])

```

```

13 print(A[-length(A)])
14 X = A[-1] - A[-length(A)]
15 print(X)
16
17 print(mean(X))
18 print(sd(X))

```

---

**Ex 1.1.** Explain the result of line P. 12 (R.12) marked as 'this'.

**Ex 1.2.** Compare the result of line P.13 and P.14 (R.12, R.13); explain what is  $A[1:]$  ( $A[-1]$ )

**Ex 1.3.** Compare the result of line P.12 and P.14 (R.11 and R.13); explain what is  $A[: -1]$  ( $A[-length(A)]$ ).

**Ex 1.4.** Explain what is  $X$  in P.15 (R.14).

**Ex 1.5.** Why do we compare  $1/\lambda$  and  $X.mean()$ ?

**Ex 1.6.** Recall that  $E[X] = \sigma(X)$  when  $X \sim \text{Exp}(\lambda)$ . Hence, what do you expect to see for  $X.std()$ ?

**Ex 1.7.** Run the code for a larger sample, e.g. 50, and discuss (very briefly) your results.

## 1.2 BH.5.6.5

Read this example of BH first. We chop up the exercise in many small exercises.

For the Python code below, run it for a small number of samples; here, I choose `samples=2`. Read the print statements, and use that to answer the questions below.

Python Code

---

```

1 import numpy as np
2 from scipy.stats import expon
3
4 np.random.seed(10)
5
6 labda = 6
7 num = 3
8 samples = 2
9
10 X = expon(scale=labda).rvs((samples, num))
11 print(X) #this
12 T = np.sort(X, axis=1)
13 print(T)
14 print(T.mean(axis=0))
15
16 expected = np.array([labda / (num - j) for j in range(num)])
17 print(expected)
18 print(expected.cumsum())

```

---

## R Code

```

1 set.seed(10)
2
3 labda = 6
4 num = 3
5 samples = 2
6
7 X = matrix(rexp(samples * num, rate = 1 / labda), nrow = samples, ncol = num)
8 print(X) #this
9 bigT = X
10 for (i in 1:samples) {
11     bigT[i,] = sort(bigT[i,])
12 }
13 print(bigT)
14 print(colMeans(bigT))
15
16 expected = rep(0, num)
17 for (j in 1:num) {
18     expected[j] = labda / (num - (j - 1))
19 }
20 print(expected)
21 print(cumsum(expected))

```

**Ex 1.8.** In line P.11 (R.8),<sup>2</sup> also marked as 'this', we print the value of X in line P.10 (R.7), respectively. What is the meaning of X?

**Ex 1.9.** What is the meaning of T in line P.12 (R.11)?

**Ex 1.10.** What do we print in line P.14 (R.14)?

**Ex 1.11.** What is meaning of the variable expected?

**Ex 1.12.** What is the cumsum of expected?

**Ex 1.13.** Now that you understand what is going on, rerun the simulation for a larger number of samples, e.g., 1000, and discuss the results briefly.

### 1.3 BH.7.1

First read and solve BH.7.1. Answers and hints are in the study guide.

**Ex 1.14.** Sketch, on paper, the 2D area that corresponds to the event that Alice and Bob meet. Just make a photo with your mobile phone and include it. (Like this you learn how to include jpeg or png files in  $\text{\LaTeX}$ .)

<sup>2</sup> Line P.x refers to line x of the Python code. Line R.x refers to line x of the R code.

**Ex 1.15.** Write  $A$  and  $B$  for the rvs corresponding to the arrival times of Alice and Bob. Explain you have to compute the expectation of  $I_{A < B + 0.25} \cdot I_{B - 0.25 < A}$  to compute the probability that Alice and Bob meet. Then, explain that is equal to

$$\int_0^1 \int_0^1 I_{a < b + 0.25} \cdot I_{b - 0.25 < a} da db. \quad (1.1)$$

**Ex 1.16.** What is the probability that they meet (i.e., solve the integral by hand)?

**Ex 1.17.** Explain this code to compute numerically the probability of overlap. Check the web on what is an anonymous function in Python or R, and explain why we use that here. (Hint: anonymous functions are useful "when it's not worth the effort to give it a name". Do we use elsewhere the anonymous function we define here?)

Python Code

```
1 from scipy.integrate import dblquad
2
3 area = dblquad(lambda a, b: (b - 0.25 < a) * (a < b + 0.25), 0, 1, 0, 1)
4 print(area)
```

R Code

```
1 library(cubature)
2
3 area = adaptIntegrate(function(x){ (x[2] - 0.25 < x[1]) * (x[1] < x[2] + 0.25)},
4                             lowerLimit = c(0, 0), upperLimit = c(1, 1))
5 print(area$integral)
6 print(area$error)
```

**Ex 1.18.** Explain the output, i.e., the results of the print statements, of this piece of code.

Python Code

```
1 import numpy as np
2 from scipy.stats import uniform
3
4 np.random.seed(3)
5
6 n = 5
7 alice = uniform(0, 1).rvs(n)
8 bob = uniform(0, 1).rvs(n)
9 print(alice)
10 print(bob)
11 print((alice < bob))
```

R Code

```
1 set.seed(3)
2
```

```

3 n = 5
4 alice = runif(n,0,1)
5 bob = runif(n,0,1)
6
7 print(alice)
8 print(bob)
9 print(alice<bob)

```

---

I tend to set the seed of the random number generator so that I always get the same results. This helps a lot when debugging, because the numbers don't change time and again.

**Ex 1.19.** Run this piece of code, and then check that you get different output every time.

---

Python Code

---

```

1 import numpy as np
2 from scipy.stats import uniform
3
4 n = 5
5 alice = uniform(0, 1).rvs(n)
6 bob = uniform(0, 1).rvs(n)
7 print(alice)
8 print(bob)
9 overlap = (alice < bob) # this
10 print(overlap)

```

---



---

R Code

---

```

1 n = 5
2 alice = runif(n,0,1)
3 bob = runif(n,0,1)
4 print(alice)
5 print(bob)
6 overlap = alice<bob #this
7 print(overlap)

```

---

Explain that the line marked as 'this' implements an indicator function.

**Ex 1.20.** Run this piece of code, and explain the output.

---

Python Code

---

```

1 import numpy as np
2 from scipy.stats import uniform
3
4 np.random.seed(3)
5
6 n = 5

```



```

7  alice = uniform(0, 1).rvs(n)
8  bob = uniform(0, 1).rvs(n)
9  overlap = (bob - 0.1 < alice) * (alice < bob + 0.1)
10 print(alice)
11 print(bob)
12 print(overlap) # this

```

---

R Code

---

```

1  set.seed(3)
2
3  n = 5
4  alice = runif(n,0,1)
5  bob = runif(n,0,1)
6  overlap = (bob - 0.1 < alice) * (alice < bob + 0.1)
7  print(alice)
8  print(bob)
9  print(overlap) #this

```

**Ex 1.21.** To find the probability of overlap, we have to modify the code above the line marked as ‘this’ to

---

Python Code

---

```

1  print(overlap.mean())

```

---

R Code

---

```

1  print(mean(overlap))

```

Change this in your code and explain what it does.

**Ex 1.22.** Modify the code of the previous exercise so that you sample 1,000 times an appointment between Alice and Bob, and such that they only meet when 15 minutes apart. Include your code and show the output.

**Ex 1.23.** Contrary to [1.14] suppose Bob is not prepared to wait longer than 5 minutes, while Alice is still prepared to wait for 15 minutes. Write down the integral that corresponds to the probability that they meet. Solve it by hand. Modify the above code such that you can evaluate it numerically. Include your code in the answers.

**Ex 1.24.** Contrary to [1.14] suppose Bob is prepared to wait for 20 minutes when he arrives before 1230, but just 10 when he arrives after 1230, while Alice is still prepared to wait for 15 minutes. Write down the integral that corresponds to the probability that they meet. Explain that this is the code to compute this.

## Python Code

```

1 from scipy.integrate import dblquad
2 area1 = dblquad(lambda a, b: (b - 1 / 3.0 < a) * (a < b + 0.25), 0, 1, 0, 0.5)
3 area2 = dblquad(lambda a, b: (b - 1 / 6.0 < a) * (a < b + 0.25), 0, 1, 0.5, 1)
4 print(area1[0] + area2[0])

```

## R Code

```

1 library(cubature)
2
3 area1 = adaptIntegrate(function(x){(x[2] - 1 / 3.0 < x[1]) * (x[1] < x[2] + 0.25)},
4                           lowerLimit = c(0, 0), upperLimit = c(1, 0.5))
5 area2 = adaptIntegrate(function(x){(x[2] - 1 / 6.0 < x[1]) * (x[1] < x[2] + 0.25)},
6                           lowerLimit = c(0, 0.5), upperLimit = c(1, 1))
7 print(area1$integral+ area2$integral)

```

**Ex 1.25.** Contrary to [1.14] suppose Bob's arrival time is  $\text{Exp}(\lambda)$  with  $\lambda = 1/2$ , while Alice's arrival time is still uniform. Explain the code below.

## Python Code

```

1 import numpy as np
2 from scipy.integrate import dblquad
3 from scipy.stats import uniform, expon
4
5 np.random.seed(3)
6
7 labda = 1.0 / 2
8 end = 5
9
10 value = dblquad(
11     lambda a, b: (b - 0.25 < a) * (a < b + 0.25) * np.exp(-labda * b), 0, 1, 0, end
12 )
13 print(value)
14
15 area = value[0] * labda
16 print(area)
17
18 n = 100000
19 alice = uniform(0, 1).rvs(n)
20 bob = expon(scale=1 / labda).rvs(n)
21 overlap = (bob - 0.25 < alice) * (alice < bob + 0.25)
22 print(overlap.mean())

```

## R Code

```

1 library(cubature)
2 set.seed(3)

```

```

3
4  labda = 1.0/2
5  end = 5
6
7  value = adaptIntegrate(function(x){(x[2] - 0.25 < x[1]) * (x[1] < x[2] + 0.25) *
8                      exp(-labda * x[2])}, lowerLimit = c(0, 0), upperLimit = c(1, end))
9  print(value)
10
11 area = value$integral * labda
12 print(area)
13
14 n = 100000
15 alice = runif(n, 0, 1)
16 bob = rexp(n, labda)
17 overlap = (bob - 0.25 < alice) * (alice < bob + 0.25)
18 print(mean(overlap))
19
20 print(mean(bob))

```

---

**Ex 1.26.** Changing the value of the variable `end` to 100 in the code above, and run it again. Then change it to 50, and finally to 1,000. When I do this, I get a very strange answers for 50 and 1,000. Try to explain what goes wrong.

Hopefully you got by now the point: modeling is an interesting, but quite difficult, activity. Once you have the model, the computer can help solve the problem.

You should memorize that integration in multiple dimensions quickly becomes unmanageable. In fact, most higher D integrals are approximated with simulation, like we do here in the simulation part.

You should also memorize that computers are extremely dumb, and that they can live with any answer that comes out of a computation. Here, since we compare a simulation with a direct integration, we see a difference, and that makes us suspicious about the correctness of both. Mind, we humans can do pretty smart things, and we should typically distrust the output of computers. (That is why I often include tests.)

#### 1.4 BH.7.9

First read and solve BH.7.9. Here we provide a simulation.

**Ex 1.27.** Run this code, and use the documentation of `scipy.stats.geom` in Python ( `rgeom()` function in R) to explain why it fails. Perhaps you can also use the wikipedia page on the geometric distribution. (Hint: is this the first success distribution a geometric distribution?)

---

Python Code

---

```

1  import numpy as np
2  from scipy.stats import randint, geom

```

```

3
4 np.random.seed(3)
5
6 p, num = 0.1, 10000
7 q = 1 - p
8 rv = geom(p) # This line
9 X = rv.rvs(num)
10
11 if X.min() != 0:
12     raise ValueError("The minimal value of X is not 0")
13 if not np.isclose(q / p, X.mean(), 0.1):
14     raise ValueError("The mean of X is not ok")

```

---

R Code

---

```

1 library(dplyr)
2 set.seed(3)
3
4 p = 0.1
5 num = 100000
6 q = 1 - p
7 X = rgeom(num, p) + 1 #This line
8
9 if (min(X) != 0){
10     stop("The minimal value of X is not 0")
11 }
12 if (!(near(q/p, mean(X), 0.1))){
13     stop("The mean of X is not ok")
14 }

```

**Ex 1.28.** The code of the previous exercise failed, on purpose. It will work if we change the line marked as “this line”, into this:

---

Python Code

---

```

1 rv = geom(p, loc=-1)

```

---

R Code

---

```

1 X = rgeom(num, p)

```

Explain the difference, and why it now works. Relate this location-scale transformation to one of the distributions you know from BH.

Memorize from this exercise that with such checks you can build code that fails when it gets unexpected input.

**Ex 1.29.** Run this code and explain in detail the output. Make, and include, some simple example code to show that you are correct.

---

Python Code

---

```
1 import numpy as np
2
3 N = np.array([3, 4, 4, 8, 4])
4 X = np.array([8, 7, 6, 5, 3])
5 print(np.argwhere(N == 4))
6 Xn = X[np.argwhere(N == 4)]
7 print(Xn)
```

---

---

R Code

---

```
1 N = c(3, 4, 4, 8, 4)
2 X = c(8, 7, 6, 5, 3)
3 print(which(N == 4))
4 Xn = X[which(N == 4)]
5 print(Xn)
```

---

**Ex 1.30.** Explain the lines marked with ‘this in the code of Listing 1 or Listing 2. You already explained the other parts so you don’t have to do that again.

**Ex 1.31.** Explain the results. Do they match with the theoretical results?

## Python Code

```

1 import numpy as np
2 from scipy.stats import randint, geom
3 import matplotlib.pyplot as plt
4
5 np.random.seed(3)
6
7 p, num = 0.1, 100000
8 q = 1 - p
9 rv = geom(p, loc=-1)
10 X = rv.rvs(num)
11
12 Y = rv.rvs(num)
13 N = X + Y
14
15 fig, axes = plt.subplots(3, 3, sharey="all", figsize=(6, 3)) # this
16 n = 7 # gives nice results, after some experimentation
17 for ax in axes.flatten(): # this
18     Xn = X[np.argwhere(N == n)]
19     y, bins = np.histogram(Xn, bins=np.linspace(0, n + 1, n + 2)) # this
20     ax.plot(range(n + 1), y, 'bo', ms=3, label='N=8') # this
21     ax.vlines(range(n + 1), 0, y, colors='b', lw=2, alpha=0.5) # this
22     ax.set_title(f'{n=} ({len(Xn)})')
23     ax.set_ylim(0) # this
24     n += 1

```

Listing 1: BH.7.9, Python code.

## R Code

```

1 library(ggplot2)
2 library(gridExtra)
3
4 p = 0.1
5 num = 100000
6 q = 1 - p
7 X = rgeom(num, p)
8
9 Y = rgeom(num, p)
10 N = X + Y
11
12 pdf(file="bh-7-9.pdf", width = 8, height = 7,
13     bg = "white", colormodel = "cmyk", paper = "A4")
14
15 plots <- list()
16 par(oma= c(0,0,0,0))
17
18 i = 1
19 n = 7 # gives nice results, after some experimentation
20 while (i <= 9){ #this
21   Xn = X[which(N == n)]
22   df = data.frame(Xn)
23   plots[[i]] = ggplot()+
24     geom_histogram(df, mapping = (aes(x = Xn)),
25                   bins = n + 1, color = "blue")+ # this
26   ylim(0,600)+ # this
27   labs(title=paste0("n=", n, ", (", length(Xn) ,")")+
28   theme(axis.title.x=element_blank(),
29         axis.title.y = element_blank())
30   n = n + 1
31   i = i + 1
32 }
33
34 do.call("grid.arrange", c(plots, ncol=3))
35 dev.off()

```

Listing 2: BH.7.9, R code.

## 2 ASSIGNMENT 2

## 2.1 BH.7.48

Read this exercise first and solve it. Then consider the code below.

## Python Code

```

1 import numpy as np
2
3 np.random.seed(3)
4
5
6 def find_number_of_maxima(X):
7     num_max = 0
8     M = -np.infty
9     for x in X:
10         if x > M:
11             num_max += 1
12             M = x
13     return num_max
14
15
16 num = 10
17 X = np.random.uniform(size=num)
18 print(X)
19
20 print(find_number_of_maxima(X))
21
22 samples = 100
23 Y = np.zeros(samples)
24 for i in range(samples):
25     X = np.random.uniform(size=num)
26     Y[i] = find_number_of_maxima(X)
27
28 print(Y.mean(), Y.var(), Y.std())

```

## R Code

```

1 set.seed(3)
2
3 find_number_of_maxima = function(X) {
4     num_max = 0
5     M = -Inf
6     for (x in X) {
7         if(x > M) {
8             num_max = num_max + 1

```



```

9         M = x
10     }
11 }
12 return(num_max)
13 }
14
15
16 num = 10
17 X = runif(num, min = 0, max = 1)
18 print(X)
19
20 print(find_number_of_maxima(X))
21
22 samples = 100
23 Y = rep(0, samples)
24 for (i in 1:samples) {
25     X = runif(num, min = 0, max = 1)
26     Y[i] = find_number_of_maxima(X)
27 }
28
29 print(mean(Y))
30 print(var(Y))
31 print(sd(Y))

```

---

**Ex 2.1.** Explain how the small function in lines P6 to P13 (R.4-R.12) works. (You should know that `x += 1` is an extremely useful abbreviation of the code `x = x + 1`. Unfortunately, R does not have such increment operators.)

**Ex 2.2.** Explain the code in lines P25 and P26 (R.25, R.26).

## 2.2 BH.7.53

**Ex 2.3.** Read and solve BH.7.53.

**Ex 2.4.** The code in Listing 3 or Listing 4 simulates this exercise. Run the code. (To improve your understanding, just change some parameters here and there, and check the output.) Explain the lines marked with `#*`. (In particular, why do we do a runs number of runs to estimate the covariance?) Finally, explain the output.

**Ex 2.5.** Modify the code such that the drunkard makes steps of size 2 when moving left or right; the stepsizes up or down remain the same. What happens with  $E[R_n^2]$ ?

**Ex 2.6.** Optional: Add a third dimension.

## Python Code

```

1 import numpy as np
2 from scipy.stats import randint
3
4 np.random.seed(3)
5
6
7 num = 4
8 steps = randint(0, 4).rvs(num) # *
9 print(steps)
10 X = (2 * steps - 1) * (steps < 2) # *
11 Y = (2 * steps - 5) * (steps >= 2) # *
12 print(X)
13 print(Y)
14 R2 = X * X + Y * Y
15 print(f"{num=}, {R2.sum()}=")
16
17
18 def cov(X, Y):
19     return (X * Y).mean() - X.mean() * Y.mean() # *
20
21
22 num = 10
23 runs = 300
24 Ss = np.zeros(runs) # *
25 Ts = np.zeros(runs)
26 for i in range(runs): # *
27     steps = randint(0, 4).rvs(num)
28     X = (2 * steps - 1) * (steps < 2)
29     Y = (2 * steps - 5) * (steps >= 2)
30     Ss[i] = X.sum() # *
31     Ts[i] = Y.sum() # *
32
33 print(f"{cov(Ss,Ts)=}") # *

```

Listing 3: BH.7.53, python code.

## 2.3 BH 7.86

In line with Exercise BH.8.33, we are now going to analyze the effect on  $P\{D|T\}$  when the sensitivity is not known exactly. So, why is this interesting? In Example 2.3.9 the sensitivity is given, but in fact, in ‘real’ experiments, this is not always known as accurately as assumed in this example. For example, in this paper: [False-positive COVID-19 results: hidden problems and costs](#) it is claimed that ‘The current rate of operational false-positive swab tests in the UK is

## R Code

```

1  set.seed(3)
2
3  num = 4
4  steps = sample(0:4, num, replace = TRUE)
5  print(steps)
6  X = (2 * steps - 1) * (steps < 2) # *
7  Y = (2 * steps - 5) * (steps >= 2) # *
8  print(X)
9  print(Y)
10 R2 = X * X + Y * Y
11 print(paste0("num= ", num, " R2.sum() = ", sum(R2) ))
12
13 cov = function (X, Y){
14   return (mean(X * Y) - mean(X) * mean(Y)) # *
15 }
16
17 num = 10
18 runs = 300
19 Ss = replicate(runs, 0) # *
20 Ts = replicate(runs, 0)
21 for (i in runs){
22   steps = sample(0:4, num, replace = TRUE)
23   X = (2 * steps - 1) * (steps < 2)
24   Y = (2 * steps - 5) * (steps >= 2)
25   Ss[i] = sum(X) # *
26   Ts[i] = sum(Y) # *
27 }
28
29 print(paste0("cov(Ss, Ts) = ", cov(Ss, Ts)))

```

Listing 4: BH.7.53, R code.

unknown; preliminary estimates show it could be somewhere between 0.8% and 4.0%.'. Hence, even though it is claimed that PCR tests 'have analytical sensitivity and specificity of greater than 95%', it may be 4% lower. Simply put, the specificity and sensitivity are not precisely known, hence this must affect  $P\{D|T\}$ .

To help you, we show how to make one graph. Then we ask you to make a few on your own, and comment on them.

#### 2.4 BH.2.3.9

I write  $p_{D_gT}$  for  $P\{D|T\}$ . Here is how this can be implemented in python.

## Python Code

```

1 sensitivity = 0.95
2 specificity = 0.95
3 p_D = 0.01
4
5 p_T = sensitivity * p_D + (1-specificity)*(1-p_D)
6 p_D_g_T = sensitivity * p_D/p_T
7 p_D_g_T

```

## R Code

```

1 sensitivity = 0.95
2 specificity = 0.95
3 p_D = 0.01
4
5 p_T = sensitivity * p_D + (1-specificity)*(1-p_D)
6 p_D_g_T = sensitivity * p_D/p_T
7 p_D_g_T

```

1. Make a plot of  $P\{D|T\}$  in which you vary the sensitivity from 0.9 to 0.99. Explain what you see.
2. Make a plot of  $P\{D|T\}$  in which you vary the specificity from 0.9 to 0.99.
3. Make a plot of  $P\{D|T\}$  in which you vary  $P\{D\}$  from 0.01 to 0.5. Explain what you see.

## 3 ASSIGNMENT 3

## 3.1 BH.8.4 Figure

The challenge is to make Figure BH.8.4 for a specific case.

We throw a coin a number of times. We take some beta distribution as prior for the probability  $p$  that Heads appears. Then we update the PDF of  $p$  according to outcomes of the throws.

**Ex 3.1.** Run the follow code and explain the output.

We define the variables H and T as these outcomes are easier to memorize than 1 and 0. Like this, there is less possibility for confusion, hence less bugs.

## Python Code

```

1 H, T = 1, 0
2 outcomes = [H, T]
3
4 h, t = 1, 1 # this
5 h += outcomes[0] == H
6 t += outcomes[0] == T
7 print(f"{h=}, {t=}")
8
9 h += outcomes[1] == H
10 t += outcomes[1] == T
11 print(f"{h=}, {t=}")

```

## R Code

```

1 H = 1
2 T = 0
3 outcomes = c(H,T)
4
5 h = 1 #this
6 t = 1 #this
7 h = h + as.numeric(outcomes[1] == H)
8 t = t + as.numeric(outcomes[1]== T)
9 print(paste0("h=", h, ", t=", t))
10
11 h = h + as.numeric(outcomes[2] == H)
12 t = t + as.numeric(outcomes[2]== T)
13 print(paste0("h=", h, ", t=", t))

```

**Ex 3.2.** Explain the code below. Why does the line marked as ‘this’ relate to the prior distribution? What is that prior distribution?

## Python Code

```

1 import numpy as np
2 from scipy.stats import beta
3 import matplotlib.pyplot as plt
4
5 support = np.arange(0, 1, 0.03)
6
7 H, T = 1, 0
8 outcomes = [H, H, T, H, H, H, T, H, H]
9 h, t = 1, 1 # this
10
11 fig, axes = plt.subplots(3, 3, sharey="all", figsize=(6, 3)) # this
12 for i, ax in enumerate(axes.flatten()): # this
13     h += outcomes[i] == H
14     t += outcomes[i] == T
15     X = beta(h, t).pdf(support)
16     ax.plot(support, X)
17     ax.set_title(f'h={h-1}, t={t-1}')
18
19 plt.tight_layout()
20 plt.savefig("beta.pdf")

```

## R Code

```

1 library(ggplot2)
2 library(gridExtra)
3
4 support = seq(0, 1, 0.03)
5
6 H = 1
7 T = 0
8 outcomes = c(H, H, T, H, H, H, T, H, H)
9 h = 1 #this
10 t = 1 #this
11
12 pdf(file="beta.pdf", width = 8, height = 7,
13     bg = "white", colormodel = "cmyk", paper = "A4")
14
15 plots <- list()
16 par(oma= c(0,0,0,0))
17
18 i = 1
19 while (i <= length(outcomes)){ #this
20     h = h + (outcomes[i] == H)
21     t = t + (outcomes[i] == T)
22     X = dbeta(support, h, t)

```

```

23 df = data.frame(support,X)
24 plots[[i]] = ggplot(data=df, aes(x=support, y=X, group=1)) +
25             geom_line()+
26             ylim(0,3.5)+
27             labs(title=paste0("h=", h-1, ", t=", t-1),x=" ", y = " ")
28 i = i+1
29 }
30
31 do.call("grid.arrange", c(plots, ncol=3))
32
33 dev.off()

```

---

**Ex 3.3.** What do the PDFs actually model? Explain why the PDFs behave as they do. For instance, why does the location of the peak move from right to left to right?

**Ex 3.4.** Which conjugacy relation between prior and data do we use here to update the prior to the posterior?

**Ex 3.5.** Adapt the code above such that the prior is  $a = 3.5, b = 1.5$ , the outcomes are T, H, T, H, T, H, and the output is a figure with 3 x 2 subfigures. Besides the explanation, include the figure.

### 3.2 BH-8-4-3

Let us investigate BH.8.4.3. The rvs  $X$  and  $Y$  are independent,  $X \sim \text{Exp}(\lambda)$  and  $Y \sim \text{Exp}(\mu)$ . We want to compare  $X + Y$  to  $\Gamma(2, \mu)$  and  $\Gamma(2, \lambda)$ , just to see what happens.

**Ex 3.6.** Run the code below. Explain the lines marked as 'this'. In particular, why do we simulate outcomes for  $X$  and  $Y$ ? Then, why are the shape and scale variables as they are? Then explain the three graphs.

Python Code

---

```

1 import numpy as np
2 from scipy.stats import expon, gamma
3 import matplotlib.pyplot as plt
4
5 labda, mu = 3, 5
6
7 n = 1000
8 X = expon(scale=1 / labda)
9 Y = expon(scale=1 / mu)
10 Z = X.rvs(n) + Y.rvs(n)
11 G1 = gamma(2, scale=1 / labda)
12 G2 = gamma(2, scale=1 / mu)
13

```

```

14 support = np.arange(0, 5, 0.03)
15
16 plt.hist(Z, label="Z", density=True)
17 plt.plot(support, G1.pdf(support), label="G1")
18 plt.plot(support, G2.pdf(support), label="G2")
19 plt.legend()
20 plt.tight_layout()
21 plt.savefig("figures/gamma1.pdf")

```

---

R Code

---

```

1 library(ggplot2)
2
3 labda = 3
4 mu = 5
5
6 n = 1000
7 X = rexp(n, labda)
8 Y = rexp(n, mu)
9 Z = X + Y
10
11 support = seq(0, 5, 0.03)
12
13 G1 = dgamma(support, 2, scale=1 / labda)
14 G2 = dgamma(support, 2, scale=1 / mu)
15
16 df1 = data.frame(Z)
17 df2 = data.frame(support, G1, G2)
18
19 pdf(file="figures/gamma1.pdf", width = 8, height = 7,
20     bg = "white", colormodel = "cmyk", paper = "A4")
21
22 ggplot()+
23   geom_histogram(df1, mapping = (aes(x = Z, y = ..density.., fill = "Z")), bins = 20)+
24   geom_line(df2, mapping = (aes(x = support, y = G1, color="G1")))+
25   geom_line(df2, mapping = (aes(x = support, y = G2, color = "G2")))+
26   scale_fill_manual(name = " ", values = c("Z" = "blue")) +
27   scale_color_manual(name = " ", values = c( "G1" = "orange", "G2" = "green"))+
28   theme(axis.title.x=element_blank(),axis.title.y = element_blank())
29
30 dev.off()

```

---

**Ex 3.7.** Modify the above code such that  $X$  and  $Y$  are both  $\text{Exp}(\lambda)$ . Then make a histogram of a simulation of  $X + Y$  and compare it to the PDF of the correct gamma distribution. Discuss your result, and include the figure of course.



## 3.3 BH.8.1

**Ex 3.8.** Here is the code to solve BH.8.1. Explain the relevant steps, i.e., the lines marked as ‘this’. Is the result according to your expectation?

## Python Code

```

1 import numpy as np
2 from scipy.stats import expon
3 import matplotlib.pyplot as plt
4
5 X = expon(scale=1) # this
6 Y = np.exp(-X.rvs(1000)) # this
7 plt.hist(Y, density=True, label="sim")
8
9 x = np.linspace(0, 1, 30)
10 dxdy = np.exp(x) # this
11 pdf_y = X.pdf(x) * dxdy # this
12 plt.plot(x, pdf_y, label="f_Y")
13 plt.legend()
14 plt.tight_layout()
15 plt.savefig("figures/bh-8-1-fig.pdf")

```

## R Code

```

1 library(ggplot2)
2
3 X = rexp(1000, 1) #this
4 Y = exp(-X) #this
5 df1 = data.frame(Y)
6
7 x = seq(0,1,length.out = 30)
8 dxdy = exp(x) #this
9 pdf_y = dexp(x,1) * dxdy #this
10 df2 = data.frame(x,pdf_y)
11
12 pdf(file="figures/bh-8-1-fig.pdf", width = 8, height = 7,
13     bg = "white", colormodel = "cmyk", paper = "A4")
14
15 ggplot()+
16   geom_histogram(df1, mapping = (aes(x = Y, y = ..density.., fill = "sim")), bins = 20)+
17   geom_line(df2, mapping = (aes(x=x,y=pdf_y, color = "f_y")))+
18   scale_fill_manual(name = " ", values = c("sim" = "blue")) +
19   scale_color_manual(name = " ", values = c( "f_y" = "orange"))+
20   theme(axis.title.x=element_blank(),axis.title.y = element_blank())
21
22 dev.off()

```

**Ex 3.9.** What error would we make if we would use this code: `dx dy = np.exp(-x) (dx dy = exp(-x))`?

### 3.4 BH.8.18

How many ping pong balls fit into an Airbus Beluga? One way to answer this is as follows. According to this [wiki-page](#) the cargo volume  $V$  of this airplane is  $1500\text{ m}^3$ . But this number is based on the physical dimensions that is available to store containers, tanks, and so on. So, I estimate the volume as about twice that amount, i.e.,  $V = 2500\text{ m}^3$ . The volume of a ping pong ball is  $v = 4\pi r^3/3 = 33.49333333333333\text{ cm}^3$  with  $r = 2\text{ cm}$ . A plain division gives 74.6268656716418 ping pong balls. Note, I left out the  $10^6$  conversion from meters to cm, and I do not take into account the sphere packing factor. Besides that, I hope you agree with me that providing a result with the precision as given here is plain ridiculous.<sup>3</sup>

In fact, I know that the volumes of an airplane and a ping pong ball is an estimate, rather than a precise number as assumed above. It seems to be better to approximate  $V$  and  $v$  as rvs. Let's assume that

$$V \sim N(2500, 500^2), v \sim N(33.5, 0.5^2),$$

where the variances express my trust in my guess work. What is now the mean of  $N = V/v$  and its std? Can we get the CDF? From BH.8.18 you know that finding the closed form expression for the distribution of  $N$  is not entirely simple. However, with simulation it's easy to get an estimates for the mean and the standard deviation.

**Ex 3.10.** Explain lines marked as 'this'.

---

Python Code

---

```

1 import numpy as np
2 from scipy.stats import norm
3
4 num = 500
5
6 np.random.seed(3)
7
8 V = norm(loc=2500, scale=500) # this
9 v = norm(loc=33.5, scale=0.5) # this
10
11 print(V.mean(), V.std()) # just a check
12
13 N = V.rvs(num) / v.rvs(num) # this
14 print(N.mean(), N.std())
15
16 print(2500/33.5)
17 print(np.sqrt(500*0.5))

```

---

<sup>3</sup> For reasons incomprehensible to me, even professional econometricians sometimes report results with 10 digits or more, without questioning the precision.

## R Code

```

1 num <- 500
2
3 set.seed(3)
4
5 V = rnorm(num, 2500, 500) # this
6 v = rnorm(num, 33.5, 0.5) # this
7
8 N = V / v # this
9 print(paste(mean(N), sd(N)))

```

**Ex 3.11.** Contrary to BH.7.1.25 if you run the code above, you'll see that  $E[N] < \infty$ , and is very near to the deterministic answer. But isn't this strange: we divide two normal random variables, just like BH.7.1.25, but there the expectation is infinite. Explain why it works for the Beluga case, but not for the case discussed in BH.7.1.25.

**Ex 3.12.** Include code to make a histogram of the PDF of  $N$ .

**Remark 3.13.** Here is a final, and optional, question. The numerical results suggest the interesting guess  $V[N] \approx V[V] * V[v]$ , but is this true more generally?

## 4 ASSIGNMENT 4

## 4.1 BH.8.27

We start from BH.8.27 (which you have to read now). We are interested in the difference between the distribution of  $X + Y + Z$  and the normal distribution. But why the normal distribution? As it turns out, the central limit law, see BH.10, states that the distribution of sums of r.v.s converge to the normal distribution (in a specific sense).

Here is some code to simulate.

## Python Code

```

1 import numpy as np
2 from scipy.stats import norm
3
4 import matplotlib.pyplot as plt
5 import seaborn as sns
6
7 sns.set()
8
9 np.random.seed(3)
10
11 k = 3
12 Zexact = norm(loc=k / 2, scale=np.sqrt(k / 12))
13 X = np.arange(0, 3, 0.1)
14
15 XYZ = np.random.uniform(size=(4000, k))
16 # print(XYZ) # if you want to see it.
17 Z = XYZ.sum(axis=1)
18 sns.distplot(Z)
19 plt.plot(X, Zexact.pdf(X))
20 plt.show()

```

## R Code

```

1 set.seed(3)
2
3 k = 3
4 X <- seq(0, 3, by = 0.1)
5 Zexact <- dnorm(X, mean = k / 2, sd = sqrt(k / 12))
6
7
8 XYZ <- matrix(NA, 4000, k)
9 # print(XYZ) # if you want to see it.
10 for (i in 1:k) {
11   XYZ[,i] <- runif(4000, min = 0, max = 1)
12 }

```

```

13 Z <- rowSums(XYZ)
14
15 par()
16 hist(Z, prob = TRUE, breaks = 31)
17 lines(X, Zexact, type = "l", col = "orange")
18 lines(density(Z), col = "blue")

```

---

**Ex 4.1.** What is the shape of XYZ in the code above, i.e., how many rows and columns does it have? If you don't know, run the code, and print it.

**Ex 4.2.** What is the shape (rows and columns) of Z?

**Ex 4.3.** Explain the values for `loc` and `shape` in `Zexact` ( and corresponding mean and sd in R). (Read the documentation of `scipy.stats.norm` (or `norm()` for R) on the web if necessary.) To which definition in BH does this loc-scale transformation relate?

**Ex 4.4.** Change the seed to your student id, or any other number you like, run the code, and include the graph produced by your simulation. Explain what you see.

Now we do an exact computation.

Python Code

```

1 import numpy as np
2 from scipy.stats import norm
3
4 import matplotlib.pyplot as plt
5 import seaborn as sns
6
7 sns.set()
8
9 N = 200
10 x = np.linspace(0, 2, 2 * N)
11 fx = np.ones(N) / N
12 f2 = np.convolve(fx, fx)
13 f3 = np.convolve(f2, fx)
14
15 k = 3
16
17 x = np.linspace(0, k, len(f3))
18 Zexact = norm(loc=k / 2, scale=np.sqrt(k / 12))
19
20
21 plt.plot(x, N * f3, label="conv")
22 plt.plot(x, Zexact.pdf(x))
23 plt.legend()
24 plt.show()

```

---

## R Code

```

1 N = 200
2 x = seq(0, 2, length.out = 2 * N)
3 fx = rep(1, N) / N
4 f2 = convolve(fx, fx, type = "open")
5 f3 = convolve(f2, fx, type = "open")
6
7 k = 3
8
9 x = seq(0, k, length.out = length(f3))
10 Zexact = dnorm(x, mean = k/2, sd = sqrt(k / 12))
11
12 par()
13 plot(x, N * f3, col = "blue", type = "l", ylim = c(0, 0.8))
14 lines(x, Zexact, type = "l", col = "orange")
15 legend("topright", legend = "conv", bty = "n",
16        lwd = 2, cex = 1.2, col = "blue", lty = 1)

```

**Ex 4.5.** Read the documentation of `np.convolve` (or `convolve()` in R). Why is it called like this?

**Ex 4.6.** In the code, what is `f2`?

**Ex 4.7.** What is `f3`?

**Ex 4.8.** Why do we set `k=3`?

**Ex 4.9.** A bit harder, why do we plot `N*f3`, i.e., why do we have to multiply with `N`? Relate this to the meaning of  $\int f(x) dx$ , where  $f$  the density of some random variable. (Think hard about this problem; the ideas are important to understand how numerical integration works. ) Suppose we chop up the area under some arbitrary function  $g$  in blocks of height  $g(x)$  and length  $\Delta x$ . Then the area of such a block is  $g(x)\Delta x$ .

In our case, we chop up the interval in parts with length  $\Delta x = 1/N$ . The elements of `f3` are such that  $f3[i] = f_3(x_i)\Delta x$ , where  $x_i$  lies in the  $i$ th interval and  $f_3$  is the density of the sum of the three r.v.s. But then,  $f_3(x_i) = f3[i]/\Delta x = Nf3[i]$ .

So, why should we scale with  $\Delta x = 1/N$ ? (Forgetting this step is a common error when dealing with densities.)

To memorize,  $\int f(x) dx$  is the area of a block of height  $f(x)$  and length  $dx$ .

**Ex 4.10.** Yet a tiny bit harder, consider `f4 = np.convolve(f3, fx)` (`f4 = convolve(f3, fx, type = "open")` in R) and `g4 = np.convolve(f2, f2)` (`g4 = convolve(f2, f2, type = "open")` in R). Why are they, numerically speaking, equal?

**Ex 4.11.** When you would compute the maximum of `np.abs(f4 - g4)` (`abs(f4 - g4)` in R) you would see that this is about  $10^{-10}$ , or so. Hence, a small number. This is not equal to 0, but we know that this is due to rounding effects.

How can we use the function `np.isclose()` to get around this problem? (You should memorize from this question that you should take care when testing on whether floating point numbers are the same or not.)

## 4.2 BH.8.36

We simulate the post office part of BH.8.36. Read it now, i.e., before reading the text below. Then read the code below. In the questions, we ask you to explain what the code does. There are lots of print statements that have been commented out, but we left them in for you to include while experimenting with the code to see how the code works. (I often use print statements of intermediate results when writing a program, just to see whether I am still on track. Once I checked, I remove them, because they clutter the looks of the code.)

## Python Code

```

1 import numpy as np
2 from scipy.stats import expon
3
4 np.random.seed(3)
5
6 labdas = np.array([3, 4])
7
8 N = 10
9
10 T1 = expon(scale=1 / labdas[0]).rvs(N)
11 # print(T1.mean())
12 T2 = expon(scale=1 / labdas[1]).rvs(N)
13 T = np.zeros([2, N])
14 T[0, :] = T1
15 T[1, :] = T2
16 # print(T)
17 server = np.argmin(T, axis=0)
18 # print(server)
19
20 # BH.8.36.b
21 print((server == 1).mean())
22
23 # BH.8.36.c
24 # print(labdas[server])
25 S = expon(scale=1).rvs(N) / labdas[server] # this
26 # print(S)
27
28 D = np.min(T, axis=0) + S
29 # print(D)
30
31 print(D.mean(), D.std())

```

## R Code

```

1 set.seed(3)
2

```

```

3  labdas = 3:4
4
5  N = 10
6
7  T1 = rexp(N, rate = labdas[1])
8  T2 = rexp(N, rate = labdas[2])
9  #print(mean(T1))
10
11 bigT = matrix(0, nrow = 2, ncol = N)
12 bigT[1,] = T1
13 bigT[2,] = T2
14 #print(T)
15
16 server = rep(0, ncol(bigT))
17
18 for (i in 1:ncol(bigT)) {
19   server[i] = which.min(bigT[,i])
20
21   #print(server)
22
23   #BH.8.36.b
24   print(mean((server == 1)))
25
26   #BH.8.36.c
27   #print(labdas[server])
28   S = rexp(N, rate = 1) / labdas[server] #this
29   #print(S)
30
31   D = apply(bigT, MARGIN = 2, FUN = min) + S
32   #print(D)
33
34   print(mean(D))
35   print(sd(D))

```

---

**Ex 4.12.** Explain why T1 corresponds to a number of service times of the first server.

**Ex 4.13.** To what do the rows of T (bigT) correspond?

**Ex 4.14.** What is the content of server? Why do we compute this?

**Ex 4.15.** Explain how we use the fundamental bridge in line P21 (R.24) to answer BH.8.36.b.

**Ex 4.16.** Alice is taken into service by the server that finishes first. We need to simulate the service time that Alice needs at that server. Explain how we do this in line P25 (R.28). Hint, reread BH.5.5 on the exponential. BTW, this is a good time to reread BH.5.3.



**Ex 4.17.** Why is `np.min(T, axis=0)` in the Python code (`apply(bigT, MARGIN=2, FUN=min)` in R) the time Alice spends waiting in queue, i.e., the time Alice spends in the post office before her service starts?

**Ex 4.18.** Why is `D` the departure time of Alice, i.e., the time Alice spends in the post office?

**Ex 4.19.** Set `N` to 1000 or so, or any other large number to your liking, but not so large that your computer will keep simulating for a month... Compare the values of the simulation to the theoretical result that you have to compute in BH.8.36.c.

**Ex 4.20.** Run the code for  $\lambda_1 = 1$  fixed, and take  $\lambda_2$  equal to 0.5, 1, 1.5 and 2 successively. Compute the mean time waiting time and mean sojourn time of Alice, and put your results in a table. Compare the results of the simulation to the theoretical values.

There is an important lesson to learn here. With simulation it is, often, reasonably simple to get numerical answers, but it requires many simulations to see a pattern in the numbers. For patterns, we can better use theory, as theory gives us formulas that show how the output of some model depends on the input.

### 4.3 BH.8.54

Read and solve BH.8.54. Perhaps you should read the hints in the study guide on this problem too. First we consider the case with  $r = 1$ , and we need to tackle some technical details.

**Ex 4.21.** Just as in the problem, let  $Y = pX/q$ . Why is  $E[Y] = 1$ ?

**Ex 4.22.** If  $Z \sim \text{Exp}(1)$ , then explain that  $M_Z(s) = 1/(1 - s)$  for  $s < 1$ .

**Ex 4.23.** Use the mathematical solution of the problem to explain that  $M_Y(s) \rightarrow M_Z(s)$  as  $p \rightarrow 0$ .

**Ex 4.24.** Here is code to compare  $M_Y(s)$  to  $M_Z(s)$ . Explain what the function  $M(s)$  does. Then explain why in the line marked as ‘this’ we multiply with  $1 - s_i$ . Explain why we limit  $s$  to the interval  $[0, 0.8]$ . Finally, explain why the graph goes through the point  $x = 0, y = 1$ .

#### Python Code

```

1 import numpy as np
2 from scipy.stats import geom, gamma
3 import matplotlib.pyplot as plt
4
5 np.random.seed(3)
6
7 p = 0.1
8 q = 1 - p
9 n = 10000
10
11
```

```

12 def M(s):
13     X = geom(p).rvs(n)
14     Y = p * X / q
15     return np.exp(s * Y).mean()
16
17
18 num = 30
19 m = np.zeros(num)
20 s = np.linspace(0, 0.8, num)
21 for i in range(num):
22     m[i] = M(s[i]) * (1 - s[i]) # this
23
24 plt.plot(s, m, label="MY")
25 plt.tight_layout()
26 plt.savefig("figures/moments.pdf")

```

---

R Code

```

1 set.seed(3)
2
3 p = 0.1
4 q = 1 - p
5 n = 100000
6
7 M = function(s){
8     X = rgeom(n,p) - 1
9     Y = p * X / q
10    return (mean((exp(s * Y))))
11 }
12
13
14 num = 30
15 m = rep(0,num)
16 s = seq(0,0.8,length.out = num)
17 for (i in 1:num){
18     m[i] = M(s[i]) * (1 - s[i]) # this
19 }
20
21 pdf(file="figures/moments.pdf",
22     width = 8, height = 7, bg = "white",
23     colormodel = "cmyk", paper = "A4")
24
25 plot(s,m, type = "l", ann = FALSE, col = "blue")
26
27 dev.off()

```

---

**Ex 4.25.** The match is not terrric when  $p = 0.1$ . Rerun the code, but now with  $p = 1/1000$ . Which line to you have to change for this? Include your graph. Is the match better now?

**Ex 4.26.** The problem asks for the general  $r$ . What lines of code have to change to deal with general  $r$ ? Modify the code appropriately, and explain why your code is correct.

#### 4.4 BH.9.1

Here is the code. (BTW, for this assignment, you don't have to read all of BH.9, reading the problem suffices. Hence, it is not a typo that this assignment appears in week 4, not week 5.)

##### Python Code

```

1 import numpy as np
2 from scipy.stats import randint, norm
3
4 mus = np.array([2, 5, 8])
5 stds = np.array([0.1, 0.5, 1])
6
7
8 m2 = stds ** 2 + mus ** 2 # this
9 v = m2.mean() - (mus.mean()) ** 2 # this
10 print(mus.mean(), v)
11
12 routes = []
13 for i in range(3):
14     routes.append(norm(mus[i], stds[i])) # this
15
16
17 num = 200
18 options = randint(0, 3).rvs(num) # this
19
20 times = np.zeros(num) # this
21 for i in range(num):
22     times[i] = routes[options[i]].rvs() # this
23
24 print(times.mean(), times.var())

```

##### R Code

```

1 mus = c(2, 5, 8)
2 stds = c(0.1, 0.5, 1)
3
4 m2 = stds ** 2 + mus ** 2 #this
5 v = mean(m2) - (mean(mus)) ** 2 #this
6 print(paste(mean(mus), v))
7

```

```

8  num = 200
9  options = sample(1:3, num, replace = TRUE) #this
10
11  times = rep(0, num)
12  for(i in 1:num){
13    times[i] = rnorm(1, mus[options[i]],stds[options[i]]) # this
14  }
15
16  print(paste(mean(times),var(times)))

```

---

**Ex 4.27.** Explain the lines marked as ‘this’, i.e., how does the program work?

**Ex 4.28.** Run the code and explain the results. Are the simulated results in line with the theory?

**Ex 4.29.** Modify the code so that you can plot the histogram of the `times` vector. Include your code and a figure. Set also a seed equal to, e.g., the day of the month in which you do the assignment.

**Ex 4.30.** Set  $\mu_3$  (the third element in `mus`) to 80 instead of 8. Rerun the code and explain what happens to the mean and the variance. Then change to 800 and include a histogram of the PDF of the times.

**Ex 4.31.** Change  $\sigma_1$  from 0.1 to 10. Run the code and describe what you get. What’s wrong with such a large standard deviation? Are you sure that the route lengths remain positive?

## 5 ASSIGNMENT 5

## 5.1 BH.9.1.7

The aim is to use a simulator to analyze the expected profit for the game of BH.9.1.7 but for a different prior than the uniform. First, however, we build BH.9.1.7 with the uniform prior.

First read and solve BH.9.1.7.

**Ex 5.1.** Run the code below. What are the values of accepted? What does it represent? How is the gain computed? Why do we take the mean?

## Python Code

```

1 import numpy as np
2 from scipy.stats import uniform
3
4 np.random.seed(3)
5
6 alpha = 2 / 3
7 n = 100
8 V = uniform(0, 1).rvs(n)
9
10 bid = 0.5
11 accepted = bid > alpha * V
12 gain = (V - bid) * accepted
13 print(gain.mean())

```

## R Code

```

1 set.seed(3)
2
3 alpha = 2 / 3
4 n = 100
5 V = runif(n, 0, 1)
6
7 bid = 0.5
8 accepted = bid > alpha * V
9 gain = (V - bid) * accepted
10 print(mean(gain))

```

**Ex 5.2.** Here is some code to compute the gain for multiple bids. Explain how the function  $f$  works. In other words, how does it handles the formatting of floating point numbers? The rest of the code must be clear to you, so you don't have to explain that.

Printing pretty floating points is something you really have to learn. If you unquote the line marked as 'unquote this line' and rerun the program, you'll see why people like to see formatted numbers.

Hint, search the web (Python or R documentation) on formatting floating point numbers.

## Python Code

```

1 import numpy as np
2 from scipy.stats import uniform
3
4
5 def f(x): # format floating point numbers
6     # return f"{x}" # unquote this line
7     return f"{x:3.3f}"
8
9
10 np.random.seed(3)
11
12
13 alpha = 2 / 3
14 n = 100
15 V = uniform(0, 1).rvs(n)
16
17 for b in np.arange(0.1, 1, 0.1):
18     accepted = b > alpha * V
19     gain = (V - b) * accepted
20     print(f"{f(b)}, {f(gain.mean())}")

```

## R Code

```

1 f = function(x) # format floating point numbers
2     # return(x) # unquote this line
3     return(signif(x, digits = 3))
4
5 set.seed(3)
6
7 alpha = 2 / 3
8 n = 100
9 V = runif(n, 0, 1)
10
11 for (b in seq(0.1, 0.9, 0.1)){
12     accepted = b > alpha * V
13     gain = (V - b) * accepted
14     print(paste0(f(b), ", ", f(mean(gain))))
15 }

```

**Ex 5.3.** Use the code of the previous exercise as a starting point to make a graph of the gain as a function of the bid. Compute the bids for 0.05, 0.1, 0.15, etc, i.e., in steps of 0.05. Include your code, a figure, and comment on the graph. Use your student id as seed.

**Ex 5.4.** It is very easy to compute the expected gain by means of numerical integration. Run the code below and explain how the res variable is computed. Is the result similar to the

simulation? Once again (so that you really remember what an anonymous function is!), explain very briefly why we use an anonymous function here.

---

Python Code

---

```

1 import numpy as np
2 from scipy.integrate import quad
3
4 alpha = 2 / 3
5 n = 100
6 b = 0.5
7
8 res = quad(lambda v: (v - b) * (b > alpha * v), 0, 1)
9 print(res)

```

---



---

R Code

---

```

1 library(cubature)
2
3 alpha = 2 / 3
4 n = 100
5 b = 0.5
6
7 res = adaptIntegrate(function(v){(v - b) * (b > alpha * v)},
8     lowerLimit = c(0), upperLimit = c(1))
9 print(paste(res$integral, res$error))

```

---

**Ex 5.5.** Modify the numerical integrator such that the prior on  $V$  is Beta(8,2). Include your code and give the result for a bid of  $1/2$ .

## 5.2 BH.9.1.9

This is a nice variation of one-step analysis of the coin throwing example of BH.9.1.9.

We have a mouse that sits on one corner of a cube whose edges are made of wire, and diagonally opposite the mouse there is some cheese. The mouse chooses at random (uniform) any edge and moves to the next corner. How many steps does it take, on average, for the mouse to reach the cheese (assuming the cheese does not move)?

For ease, we label all corners with the minimal number of edges that need to be traversed from the corner to reach the cheese into states. Thus, the starting corner will be called state 3, then we have a state 2, and 1, and finally state 0 is the corner with the cheese.

**Ex 5.6.** Explain in the code below how the while loop simulates the moves of the mouse. To understand the line marked as ‘this’, read what is a, so-called, ternary operator.

---

Python Code

---

```

1 import numpy as np
2 from scipy.stats import uniform

```

```

3 import matplotlib.pyplot as plt
4
5 move = uniform()
6
7
8 def do_run():
9     state = 3
10    n_steps = 0
11    while state != 0:
12        n_steps += 1
13        p = move.rvs()
14        if state == 3:
15            state = 2
16        elif state == 2:
17            state = 1 if p < 2 / 3 else 3 # this
18        elif state == 1:
19            state = 2 if p < 2 / 3 else 0
20    return n_steps
21
22
23 n = 200
24 steps = np.zeros(n)
25 for i in range(n):
26     steps[i] = do_run()
27
28 print(steps.mean(), steps.std())
29 plt.hist(steps, density=True)
30 plt.tight_layout()
31 plt.savefig("figures/mouse.pdf")

```

---

R Code

---

```

1 library(ggplot2)
2
3 do_run = function(){
4     state = 3
5     n_steps = 0
6     while (state != 0){
7         n_steps = n_steps + 1
8         p = runif(1, 0, 1)
9         if (state == 3){
10             state = 2
11         }
12         else if (state == 2){
13             state = ifelse(p < 2 / 3, 1, 3) # this
14         }

```



```

15     else if (state == 1){
16         state = ifelse(p < 2 / 3, 3, 0) # this
17     }
18 }
19 return (n_steps)
20 }
21
22 n = 200
23 steps = rep(0, n)
24
25 for (i in 1:n){
26     steps[i] = do_run()
27 }
28
29 print(paste(mean(steps),sd(steps)))
30
31 pdf(file="figures/mouse.pdf",
32     width = 8, height = 7, bg = "white",
33     colormodel = "cmyk", paper = "A4")
34
35 df1 = data.frame(steps)
36 ggplot()+
37 geom_histogram(df1, mapping = (aes(x = steps, y = ..density.., fill = "steps")))+
38 scale_fill_manual(name = " ", values = c("steps" = "blue")) +
39 theme(axis.title.x=element_blank(),axis.title.y = element_blank())
40
41 dev.off()

```

---

### 5.3 BH.9.6.1

Read BH.9.6.1 first. We will redo the computations and check the results with simulation. We will also cover many other topics as we proceed.

**Ex 5.7.** We need to model the distribution of an individual demand  $X$ . As an example, we can use the beta distribution. In the example,  $\mu$  and  $\sigma$  are assumed as given, but for the beta distribution we need an  $a$  and a  $b$ . One way is to solve by hand for  $a$  and  $b$  for given  $\mu$  and  $\sigma$ . However, we can also be lazy and let the computer do the work. Here is the code to do that.

Explain how this code works. (You can use the relevant documentation on the web on how the function `fsolve` works.) Why do I use  $\mu = 1/2$  and  $\sigma^2 = 1/12$  as test case? Are the  $a$  and  $b$  that comes out of the solver correct? What does the final check do?

---

Python Code

---

```

1 from scipy.stats import beta
2 from scipy.optimize import fsolve
3

```

```

4 mu, var = 1 / 2, 1 / 12 # why this test case?
5
6
7 def func(x):
8     a, b = x[0], x[1]
9     m = a / (a + b)
10    v = m * (1 - m) / (a + b + 1)
11    return [m - mu, v - var]
12
13
14 a, b = fsolve(func, [2, 2])
15 print(a, b)
16 X = beta(a, b)
17 print(X.stats(moments='mv')) # final check.

```

---

R Code

---

```

1 library(pracma)
2
3 mu = 1 / 2
4 var = 1 / 12 # why this test case?
5
6 func = function(x){
7     a = x[1]
8     b = x[2]
9     m = a / (a + b)
10    v = m * (1 - m) / (a + b + 1)
11    return (c(m - mu, v - var))
12 }
13
14 output = fsolve(func, c(2,2))$x
15 a = output[1]
16 b = output[2]
17 print(paste(a, b))
18
19 X = rbeta(1000,a, b)
20 print(paste(mean(X), var(X)))

```

**Ex 5.8.** What do you expect to get when you would choose  $\mu = 2$ ? If it fails, why is that?

**Ex 5.9.** What do you expect to get when you would choose  $\sigma > 1/2$ ? If it fails, why is that? Hint, when an rv has a support on  $[0, 1]$ , why is the variance at most  $1/4$ ?

**Remark 5.10.** When setting up simulations and choosing variables, make sure that the variables you choose actually make any sense.

Now we know how to find  $a$  and  $b$  for given  $\mu$  and  $\sigma$ , we henceforth assume that  $a = 1$  and  $b = 2$ , as this is easier to work with.

**Ex 5.11.** Now that we know how to simulate individual demands, we can compute the mean and the variance with the formulas of the book and compare this to simulation.

Explain the lines marked as ‘this’ of the code. (The rest must be clear for you. If not, then you have to be seriously concerned about your understanding of programming.)

---

Python Code

---

```

1 import numpy as np
2 from scipy.stats import poisson, beta
3 from scipy.optimize import fsolve
4
5 np.random.seed(3)
6
7 labda = 3
8 N = poisson(labda)
9 # print(N.rvs(100).mean()) # why this line?
10
11 X = beta(a=1, b=2)
12 mu, var = X.stats(moments='mv')
13
14 days = 1000
15 demands = np.zeros(days)
16 for i in range(days):
17     demands[i] = X.rvs(N.rvs()).sum() # this
18
19 print(demands.mean(), demands.var())
20 print(labda * mu, var * labda + mu * mu * labda) # this

```

---



---

R Code

---

```

1 set.seed(3)
2
3 labda = 3
4 # print(mean(rpois(100,labda))) # why this line?
5
6 days = 1000
7 demands = rep(0, days)
8
9 X = rbeta(1000, 1, 2)
10 mu = mean(X)
11 var = var(X)
12
13 for (i in 1:days){
14     demands[i] = sum(rbeta(rpois(1,labda), 1, 2)) # this
15 }

```

```

16
17 print(paste(mean(demands), var(demands)))
18 print(paste(labda * mu, var * labda + mu * mu * labda)) # this

```

---

**Ex 5.12.** Here is the code to compute the mean and variance with Adam and Eve's laws. Explain how it works. In particular:

1. explain the lines marked as 'this';
2. what are the contents of the demand matrix;
3. what is the  $P$ ;
4. how are  $EX$ ,  $EV$  and  $VE$  computed;
5. Note in particular the difference between  $EXn$  and  $EXN$ , and similarly for the variance. Use BH.9.2.1 to explain why I use a small  $n$  and a large  $N$  in the naming of the variables.

Remark: I realize that this is a hard piece of code, but your understanding of Eve's law will increase by a lot. Some subtle points about the code that you *don't* have to explain:

1. Suppose the largest amount of customers (jobs) that we saw during the simulation is 7. Then on any day we can have  $0, \dots, 7$  demands, i.e, 8 possible outcomes. Thus, per row in the demands matrix we need to have 8 columns.
2. The numpy function (mean() function in R) that computes the mean gives a warning when it is applied to an empty array. (Just try it to see how it works.) We prevent this by computing the histogram (i.e., the empirical PDF) for  $n = 0$  separately. The same reasoning applies to the case when  $idx$  has zero length.

---

Python Code

---

```

1 import numpy as np
2 from scipy.stats import poisson, beta, rv_discrete
3
4 np.random.seed(3)
5
6 labda = 3
7 N = poisson(labda)
8
9 X = beta(a=1, b=2)
10 mu, var = X.stats(moments='mv')
11
12 days = 1000
13 n_jobs = N.rvs(days)
14 n_max = n_jobs.max() + 1 # see the comments
15 demands = np.zeros((days, n_max))
16 for day in range(days):

```

```

17     n = n_jobs[day]
18     demands[day, :n] = X.rvs(n) # this
19
20
21     P = np.zeros(n_max)
22     EXn = np.zeros(n_max)
23     VXn = np.zeros(n_max)
24     P[0] = np.argwhere(n_jobs == 0).size / days
25     for n in range(1, n_max):
26         idx = np.argwhere(n_jobs == n)
27         if idx.size == 0:
28             continue
29         P[n] = idx.size / days # this
30         EXn[n] = demands[idx, :n].mean() * n # this
31         VXn[n] = demands[idx, :n].var() * n # this
32
33     EXN = rv_discrete(values=(EXn, P)) # this
34
35     print("mean:", labda * mu, EXN.mean())
36
37     VXN = rv_discrete(values=(VXn, P)) # this
38     print("EV: ", var * labda, VXN.mean())
39
40     VE = EXN.var()
41     print("VE: ", mu * mu * labda, VE)

```

---

R Code

---

```

1  library(discreteRV)
2  set.seed(3)
3
4  labda = 3
5
6  X = rbeta(1000, 1, 2)
7  mu = mean(X)
8  var = var(X)
9
10 days = 20
11 n_jobs = rpois(days, labda)
12 n_max = max(n_jobs) + 1 # see the comments
13 demands = matrix(0, days, n_max)
14 for (day in 1:days){
15     n = n_jobs[day]
16     if(n == 0) next
17     demands[day, 1:n] = rbeta(n, 1, 2)
18 }

```

```

19
20 P = rep(0, n_max)
21 EXn = rep(0, n_max)
22 VXn = rep(0, n_max)
23 P[1] = length(n_jobs[which(n_jobs == 0)])/days
24 for(n in 1:n_max){
25     idx = which(n_jobs == n)
26     if (length(idx) == 0){
27         next
28     }
29     P[n] = length(idx) / days # this
30     m =
31     EXn[n] = mean(demands[idx, c(1:n)]) * n #this
32     VXn[n] = sd(demands[idx, c(1:n)]) * sd(demands[idx, c(1:n)]) * n #this
33 }
34
35 EXN = RV(outcomes = EXn, probs = P) # this
36 print(paste("mean:", labda * mu, mean(EXN) ))
37
38 VXN = RV(outcomes = VXn, probs = P) # this
39 print(paste("EV: ", var * labda, VXN.mean()))
40
41 VE = var(EXN)
42 print(paste("VE: ", mu * mu * labda, VE))

```

---

#### 5.4 BH.9.7

We use simulation to solve BH.9.7. Read it now, i.e., before reading the text below, then read the code below. Note how short this code is; amazing, isn't it?

Python Code

```

1 import numpy as np
2 from scipy.stats import uniform
3 import matplotlib.pyplot as plt
4
5 np.random.seed(3)
6
7
8 N = 1000
9 a, b = 0, 10000000
10 V = uniform(a, b).rvs(N)
11
12 x_range = np.linspace(b / 5, b / 2, num=50)
13 y = np.zeros_like(x_range)
14

```

```

15 for i, b in enumerate(x_range):
16     payoff = (V - b) * (b >= V / 4)
17     y[i] = payoff.mean()
18
19
20 plt.plot(x_range, y)
21 plt.show()

```

---

R Code

---

```

1 set.seed(3)
2
3 N = 1000
4 a = 0
5 b = 1000000
6 V = runif(N, min = a, max = b)
7
8 x_range = seq(b / 5, b / 2, length.out = 50)
9 y = rep(0, length(x_range))
10
11
12 i = 1
13 for (b in x_range) {
14     payoff = (V - b) * (b >= V / 4)
15     y[i] = mean(payoff)
16     i = i + 1
17 }
18
19 plot(x_range, y, type = "l", col = "blue")

```

**Ex 5.13.** For the Python code, use the scipy documentation to explain why  $V \sim \text{Unif}([0, 10^6])$ . For R, explain the function `runif()`.

**Ex 5.14.** What are the smallest and the largest value of `x_range`?

**Ex 5.15.** Run the code above and make a graph. Include the graph in your report, and explain what you see in the graph. For instance, is there a maximum? If so, can you explain where the maximum occurs? Can you explain how the maximum should be?

**Ex 5.16.** Suppose after seeing the graph of the payoffs, and this graph would only increase, or decrease, how would you change `x_range`? Do you expect to see a maximum?

**Ex 5.17.** For small N, e.g. N=10, you can get quite strange values. Why is that?

**Ex 5.18.** Change the acceptance threshold from  $V/4$  to  $V/5$  (or  $V/6$ , or some other value you like), and make a graph of the payoffs. Include the graph in your report.

**Ex 5.19.** Change the payoff function to e.g.  $\sqrt{V-b}$ , or some weird function that you like particularly such as  $\sin|V-b|$  (any non-trivial function goes). Make a graph of the mean and the standard deviation of the payoff. Can you explain your graph?

## 6 ASSIGNMENT 6

## 6.1 BH.9.25

This simulation exercise is based on BH.9.25. Please read the exercise first, and then the code below.

**Ex 6.1.** How does this code work? Use the print statements to find out what  $S$  is, and explain why we multiply by 2 and subtract 1 in one of the lines marked as ‘this’.

## Python Code

```

1 import numpy as np
2 from scipy.stats import bernoulli
3
4 np.random.seed(3)
5
6 n = 5
7 num = 10
8
9 p = 0.5
10 S = bernoulli(p).rvs([num, n]) * 2 - 1 # this
11 # print(S)
12
13 x = np.zeros([num, n])
14 x[:, 0] = 100 # this
15 # print(x)
16 f = 0.25
17
18 for i in range(1, n):
19     x[:, i] = x[:, i - 1] * (1 + f * S[:, i]) # this
20
21 print(x.mean(axis=0), x.std(axis=0))

```

## R Code

```

1 set.seed(3)
2
3 n = 5
4 num = 10
5
6 p = 0.5
7 S = matrix(rbinom(n * num, size = 1, prob = p), num, n) * 2 - 1 # this
8 # print(S)
9
10 x = matrix(0, num, n)
11 x[, 1] = 100 # this

```



```

12 #print(x)
13 f = 0.25
14
15 for (i in 2:n) {
16   x[, i] = x[, i - 1] * (1 + f * S[, i]) # this
17 }
18
19 print(colMeans(x))
20 print(apply(x, MARGIN = 2, sd))

```

---

**Ex 6.2.** 1. Run the code below.

2. Check the values of  $x$ . What do you notice?
3. Then compare both figures. You should see that in `Kelly-no-jitter.pdf` many of the marks (the  $x$  below the  $y$ -axis) fall on top of each other. To repair for this, we add some noise (jitter) to the  $x$  and  $y$  values of the marks. Explain how randomness is used here.
4. Jitter plots, like this one, are meant to give some more insight into the data to check whether we use the right number of bins. Change the number of bins to 2, include the plot with the jitters, and explain how the marks help to see that two bins is much too small.

---

Python Code

---

```

1 import numpy as np
2 from scipy.stats import bernoulli, uniform
3 import matplotlib.pyplot as plt
4
5
6 np.random.seed(3)
7
8 n = 10
9 n_experiments = 50
10
11 p = 0.4
12 S = bernoulli(p).rvs([n, n_experiments]) * 2 - 1 # this
13 # print(S)
14
15 x = np.zeros([n, n_experiments])
16 x[0, :] = 100 # this
17 f = 0.25
18
19 for i in range(1, n):
20     x[i, :] = x[i - 1, :] * (1 + f * S[i, :]) # this
21
22 print(x[-1, :].mean(), x[-1, :].std())

```

```

23 print(x[-1, :])
24
25 plt.hist(x[-1, :], bins=20, density=True)
26 plt.plot(x[-1, :], np.full_like(x[-1, :], -0.01), "xk", markeredgewidth=1)
27 plt.savefig("figures/Kelly-no-jitter.pdf")
28
29
30 x_jitter = x[-1, :] + uniform(loc=-3, scale=6).rvs(size=n_experiments)
31 y_jitter = np.full_like(x[-1, :], -0.01) + uniform(loc=-0.01, scale=0.02).rvs(
32     size=n_experiments
33 )
34 plt.plot(x_jitter, y_jitter, "xk", markeredgewidth=1)
35 plt.savefig("figures/Kelly-jitter.pdf")

```

---

R Code

---

```

1 library(ggplot2)
2
3 set.seed(3)
4
5 n = 10
6 n_experiments = 50
7
8 p = 0.4
9 S = matrix(0, n, n_experiments)
10 for(i in 1:n){
11     S[i,1:n_experiments] = rbinom(n_experiments, 1, p) * 2 -1 # this
12 }
13 #print(S)
14
15 x = matrix(0, n, n_experiments)
16 x[1,] = 100
17 f = 0.25
18
19 for(i in 2:n){
20     x[i,] = x[i-1,] * (1 + f * S[i,]) # this
21 }
22
23 pdf(file="figures/Kelly-no-jitter.pdf",
24     width = 8, height = 7, bg = "white",
25     colormodel = "cmyk", paper = "A4")
26
27 df1 = data.frame(x[nrow(x),])
28 colnames(df1) = c("X")
29 df1$Y = -0.01
30

```

```

31 p = ggplot()+
32   geom_histogram(df1, mapping = (aes(x = X, y = ..density..)), bins = 20)+
33   geom_point(df1, mapping =(aes(x=X, y = Y)))+
34   theme(axis.title.x=element_blank(),axis.title.y = element_blank())
35 p
36
37 dev.off()
38
39 pdf(file="figures/Kelly-jitter.pdf",
40     width = 8, height = 7, bg = "white",
41     colormodel = "cmyk", paper = "A4")
42
43 x_jitter = x[nrow(x),] + runif(n_experiments,-3, 3)
44 df2 = data.frame(x_jitter)
45 colnames(df2) = c("x_jitter")
46 df2$y_jitter = -0.01 + runif(n_experiments,-0.01, 0.01)
47
48 p + geom_histogram(df2, mapping = (aes(x = x_jitter, y = ..density..)), bins = 20)+
49   geom_point(df2, mapping =(aes(x=x_jitter, y = y_jitter)))+
50   theme(axis.title.x=element_blank(),axis.title.y = element_blank())
51 p
52
53 dev.off()

```

---

## 6.2 BH.9.37

With this exercise we illustrate why (and how) to use bootstrapping. At first it appears a bit strange: we have a set of observations (samples)  $X = \{X_1, X_2, \dots, X_n\}$ . For this we can compute the mean, standard deviation, and the empirical distribution with standard procedures, so why sample (bootstrap) from  $X$ ?

One type of question that seems hard to answer by classical means is to characterize the median, in particular the standard deviation and distribution of the median, of the population (that is, the population from which we took the sample  $X$ ).

Bootstrapping is nowadays used a lot in data science.

**Ex 6.3.** Run the code below, and then explain what is sample, bootstrap and medians. Why do we use `axis=1`?

**Ex 6.4.** Change some numbers (such as the range of the random numbers from which we obtained the initial sample). Include your code, make the plot, and include it in your assignment.

---

Python Code

---

```

1 import numpy as np
2 import matplotlib.pyplot as plt
3

```

```

4 np.random.seed(3)
5
6 sample = np.random.randint(200, size=100)
7 n_boot = 5000
8 bootstrap = np.random.choice(sample, replace=True, size=(n_boot, len(sample)))
9
10 medians = np.median(bootstrap, axis=1)
11 # print(medians.mean(), medians.std())
12 # print(np.percentile(medians, [2.5, 97.5]))
13
14 plt.hist(medians, bins=100, density=True)
15 plt.axvline(np.percentile(medians, 2.5), c='r', lw=2)
16 plt.axvline(medians.mean(), c='r', lw=2)
17 plt.axvline(np.percentile(medians, 97.5), c='r', lw=2)
18 plt.savefig("figures/bootstrap.pdf")

```

---

## R Code

```

1 library(ggplot2)
2 library(matrixStats)
3
4 np.random.seed(3)
5
6 sample = sample(1:200,100, replace=TRUE)
7 n_boot = 5000
8 bootstrap = matrix(0,n_boot,length(sample))
9 for(i in 1:nrow(bootstrap)){
10   bootstrap[i,] = sample(sample, length(sample), replace = TRUE)
11 }
12
13 medians = rowMedians(bootstrap)
14 #print(paste(mean(medians), sd(medians)))
15 #print(quantile(medians, c(.025, 0.975)))
16
17 pdf(file="figures/bootstrap.pdf",
18     width = 8, height = 7, bg = "white",
19     colormodel = "cmyk", paper = "A4")
20
21 df1 = data.frame(medians)
22 ggplot()+
23   geom_histogram(df1, mapping = (aes(x = medians, y = ..density..)), bins = 100)+
24   geom_vline(xintercept = quantile(medians, c(0.025)), color = "red", size=2) +
25   geom_vline(xintercept = quantile(medians, c(0.9725)), color = "red", size=2)+
26   geom_vline(xintercept = mean(medians), color = "red", size=2)
27
28 dev.off()

```

---

**Ex 6.5.** Finally, let us compare the information we obtained from bootstrap to the medians we would obtain when we would sample from the real population.

1. What is here the real population?
2. In a real life situation, what is less costly: real sampling or bootstrapping?
3. Explain the ideas of the code below.
4. Compare this figure to the one obtained by bootstrapping. Comment on major differences.

---

Python Code

---

```

1 import numpy as np
2 import matplotlib.pyplot as plt
3
4 np.random.seed(3)
5
6 samples = np.random.randint(200, size=(500, 100))
7 medians = np.median(samples, axis=1)
8 plt.hist(medians, bins=100, density=True)
9 plt.savefig("figures/real_medians.pdf")

```

---



---

R Code

---

```

1 library(ggplot2)
2 library(matrixStats)
3
4 set.seed(3)
5
6 sample = matrix(0, 500, 100)
7 for(i in 1:nrow(sample)){
8   sample[i,] = sample(1:200, 100, replace = TRUE)
9 }
10 medians = rowMedians(bootstrap)
11
12 pdf(file="figures/real_medians.pdf",
13     width = 8, height = 7, bg = "white",
14     colormodel = "cmyk", paper = "A4")
15
16 df1 = data.frame(medians)
17 ggplot()+
18 geom_histogram(df1, mapping = (aes(x = medians, y = ..density..)), bins = 100)
19
20 dev.off()

```

---

## 6.3 BH.9.50

This simulation exercise is based on BH.9.50. Please read the exercise first. Do the exercise while thinking about the code.

**Ex 6.6.** Here is some code to simulate  $N$ . Explain the non-trivial steps. (By now you should have an idea about boiler-plate code and what code is specific for the problem.)

## Python Code

```

1 import numpy as np
2 from scipy.stats import poisson, expon
3
4 np.random.seed(3)
5
6 num = 100
7 N = np.zeros(num)
8 Labda = expon(scale=1)
9 labdas = Labda.rvs(num)
10 for i in range(num):
11     labda = labdas[i]
12     N[i] = poisson(labda).rvs()
13
14 print(N.mean(), N.var())
15 adam = Labda.mean()
16 eve = Labda.mean() + Labda.var()
17 print(adam, eve)

```

## R Code

```

1 library(ggplot2)
2 library(matrixStats)
3
4 set.seed(3)
5
6 num = 100
7 N = rep(0, num)
8 labdas = rexp(num, 1)
9
10 for(i in 1:num){
11     labda = labdas[i]
12     N[i] = rpois(1, labda)
13 }
14
15 Labda = rexp(10000, 1)
16 print(paste(mean(N), var(N)))
17 adam = mean(Labda)

```

```

18 eve = mean(Labda) + var(Labda)
19 print(paste(adam, eve))

```

---

**Ex 6.7.** Run the above code and compare the result to the theoretical value. What do you observe? Rerun the code with `num = 10000`. Does this result in a better estimate? What do you learn from this?

**Ex 6.8.** The next step is to estimate the mean and variance of the total claim size  $S$ . Here is the code, run it and explain the relevant parts.

---

Python Code

---

```

1 import numpy as np
2 from scipy.stats import poisson, expon, lognorm
3
4 np.random.seed(3)
5
6 num = 100
7 Labda = expon(scale=1)
8 labdas = Labda.rvs(num)
9 mu, sigma = 3, 1
10 X = lognorm(loc=mu, s=sigma)
11 S = np.zeros(num)
12 for i in range(num):
13     N = poisson(labdas[i]).rvs()
14     S[i] = X.rvs(N).sum()
15
16 print(f"{S.mean()}=")
17 adam = X.mean() * Labda.mean()
18 print("mean_theoretical: ", adam)
19 print(f"{S.var()}=")
20 eve = Labda.mean() * X.var() + X.mean() ** 2 * (Labda.mean() + Labda.var())
21 print("var_theoretical: ", eve)

```

---



---

R Code

---

```

1 set.seed(3)
2
3 num = 100
4 labdas = rexp(num, 1)
5 mu = 3
6 sigma = 1
7 S = rep(0, num)
8
9 for(i in 1:num){
10     N = rpois(1, labdas[i])
11     S[i] = sum(rlnorm(N, mean = 0, sd = sigma)+ mu)

```

```

12 }
13
14 Labda = rexp(100000,1)
15 X = rlnorm(100000, mean = 0, sd= sigma) + mu
16
17 print(paste("S.mean()=", mean(S)))
18 adam = mean(X) * mean(Labda)
19 print(paste("mean_theoretical: ", adam))
20 print(paste("S.var()=", var(S)))
21 eve = mean(Labda) * var(X) + mean(X) ** 2 * (mean(Labda)+var(Labda))
22 print(paste("var_theoretical: ", eve))

```

---

**Ex 6.9.** Here is the code to plot the distribution. Change  $\lambda$  such that it is  $\sim \text{Exp}(1/2)$  and include the figure.

---

Python Code

---

```

1 import numpy as np
2 from scipy.stats import poisson, expon
3 import matplotlib.pyplot as plt
4
5 np.random.seed(3)
6
7 num = 100
8 N = np.zeros(num)
9 Labda = expon(scale=1)
10 labdas = Labda.rvs(num)
11 for i in range(num):
12     labda = labdas[i]
13     N[i] = poisson(labda).rvs()
14
15 plt.hist(N, density=True)
16 plt.savefig("figures/bh-9.50-fig.pdf")

```

---



---

R Code

---

```

1 set.seed(3)
2
3 num = 100
4 N = rep(0,num)
5 labdas = rexp(num, 1)
6 for(i in 1:num){
7     labda = labdas[i]
8     N[i] = rpois(1,labda)
9 }
10
11 pdf(file="figures/bh-9.50-fig.pdf",

```



```

12     width = 8, height = 7, bg = "white",
13     colormodel = "cmyk", paper = "A4")
14
15 df1 = data.frame(N)
16 ggplot()+
17     geom_histogram(df1, mapping = (aes(x = N, y = ..density..)))
18
19 dev.off()

```

---

**Ex 6.10.** Finally, let's do the integration numerically to compute  $P\{N = 3\}$ . Explain the code. Why, in particular, do I compare the result against `geom(1/2).pmf(4)`? (Hint, read the documentation of `scipy.stats.geom(dgeom())` function in R) to understand why.)

---

Python Code

---

```

1 import numpy as np
2 from scipy.integrate import quad
3 from scipy.stats import expon, geom
4
5 Labda = expon(scale=1)
6
7 def integrand(labda):
8     return np.exp(-labda) * labda ** 3 / 6 * Labda.pdf(labda)
9
10 res = quad(lambda labda: integrand(labda), 0, np.infty)
11 print(res)
12 print(geom(1 / 2).pmf(3))
13 print(geom(1 / 2).pmf(4))

```

---



---

R Code

---

```

1 library(cubature)
2
3 integrand = function(labda){
4     return (exp(-labda) * labda ** 3 / 6 * exp(-labda) )
5 }
6
7 res = adaptIntegrate(function(x){integrand(x[1])}, lowerLimit = c(0), upperLimit = c(Inf))
8 print(res)
9 print(dgeom(3 - 1, 1 / 2))
10 print(dgeom(4 - 1, 1 / 2))

```

---

#### 6.4 BH.10.2.3

Let us try to understand the weak law of large numbers by means of simulation. An easy example is to take  $X_i \sim \text{Unif}(0, 1)$ , so that is what we do here.

## Python Code

```

1 import numpy as np
2 from numpy.random import uniform
3
4 np.random.seed(3)
5
6 n = 10
7 N = 50 # num samples
8
9 mu = 1 / 2
10 var = 1 / 12
11 eps = 0.1
12
13 X = uniform(size=[num_samples, n])
14
15 Y = X.mean(axis=1)
16
17 larger = np.abs(Y - mu) > eps
18 count = larger.sum()
19 P = count / N
20 RHS = var / (n * eps * eps)
21 print(P, RHS)

```

## R Code

```

1 set.seed(3)
2
3 n = 10
4 N = 50 # num samples
5
6 mu = 1 / 2
7 var = 1 / 12
8 eps = 0.1
9
10 X = matrix(runif(N * n), N, n)
11
12 Y = rowMeans(X)
13
14 larger = abs(Y - mu) > eps
15 count = sum(larger)
16 P = count / N
17 RHS = var / (n * eps * eps)
18 print(P)
19 print(RHS)

```

**Ex 6.11.** Explain  $\mu$  and  $\text{var}$ .

**Ex 6.12.** What is  $Y$ ? What is the symbol that BH use for this?

**Ex 6.13.** What are the meanings of `larger` and `count`?

**Ex 6.14.** What is `RHS` in the notation of BH?

**Ex 6.15.** What inequality of BH do we check by printing `RHS` and `P`?

**Ex 6.16.** Choose some different values for  $n$  and the sample size  $N$ . Is the inequality always true?

## 7 ASSIGNMENT 7

## 7.1 BH.10.9

**Ex 7.1.** In the code below we compute or simulate the various expressions of BH.10.9. Check the results and compare it to the analytical results. Are both results in line? If not, do you know why that might be the case? What can you do about that?

I don't know how to build  $E[V[Y|X]]$  in any sensible way. Since  $Y$  is independent of  $X$ , the conditioning does not have any influence. Thus, this numerical exercise is skipped.

## Python Code

```

1 import numpy as np
2 from scipy.stats import expon
3
4 np.random.seed(3)
5
6 X = expon(2)
7 Y = expon(2)
8
9 print(np.exp(-X.mean()), X.expect(lambda x: np.exp(-x)))
10
11 n_sample = 1000
12 X_sample = X.rvs(n_sample)
13 Y_sample = Y.rvs(n_sample)
14 p1 = sum(X_sample > Y_sample + 3) / n_sample
15 p2 = sum(Y_sample > X_sample + 3) / n_sample
16 print(p1, p2)
17
18 p3 = sum(X_sample > Y_sample - 3) / n_sample
19 print(p1, p3)
20
21 EXY = X_sample @ Y_sample / n_sample
22 EX4 = X.expect(lambda x: x ** 4)
23 print(EX4, EXY * EXY)
24
25 p = sum(np.abs(X_sample + Y_sample) > 3) / n_sample
26 print(p, X.mean())

```

## R Code

```

1 library(cubature)
2 set.seed(3)
3
4 expectation = adaptIntegrate(function(x){exp(-(x[1]+2)) * exp(-x[1]) },
5                               lowerLimit = c(0), upperLimit = c(Inf))
6 print(paste(exp(-mean(rexp(10000,1)+2)), expectation$integral))

```

```

7
8 n_sample = 1000
9 X_sample = rexp(n_sample, 1) + 2
10 Y_sample = rexp(n_sample, 1) + 2
11 p1 = sum(X_sample > Y_sample + 3) / n_sample
12 p2 = sum(Y_sample > X_sample + 3) / n_sample
13 print(paste(p1, p2))
14
15 p3 = sum(X_sample > Y_sample - 3) / n_sample
16 print(paste(p1, p3))
17
18 EXY = X_sample %*% Y_sample / n_sample
19 expectation2 = adaptIntegrate(function(x){((x[1]+2) ** 4) * exp(-x[1]) },
20                               lowerLimit = c(0), upperLimit = c(Inf))
21
22 print(paste(expectation2$integral, EXY * EXY))
23
24 p = sum(abs(X_sample + Y_sample) > 3) / n_sample
25 print(paste(p, mean(X_sample)))

```

---

**Ex 7.2.** Why do I use simulation to estimate  $E[XY]$ ? Recall an earlier assignment in which we used numerical integration. How successful was that?

**Ex 7.3.** What would you do to increase the credibility of the claims? (Note these claims are based on case checking and simulation? There is no actual proof, however, they can act as counter-examples.)

## 7.2 BH.10.28

**Ex 7.4.** Let us first plot the PMF of the standardized version of  $X_n$  for  $n = 10$ . Explain the code below BEFORE making the plot.

Python Code
-------------

```

1 import numpy as np
2 from scipy.stats import poisson, norm
3 import matplotlib.pyplot as plt
4
5 n = 10
6 X = poisson(n)
7 N = norm()
8
9 pmf = np.zeros(2 * n)
10 for i in range(len(pmf)):
11     pmf[i] = X.pmf(i)
12 supp = (np.linspace(0, len(pmf), num=len(pmf)) - n) / np.sqrt(n)

```

```

13
14
15 plt.plot(supp, pmf) # this
16 plt.plot(supp, norm.pdf(supp))
17 plt.savefig("figures/bh-10-28.pdf")

```

---

R Code

---

```

1 library(ggplot2)
2
3 n = 10
4
5 pmf = rep(0, 2 * n)
6 for(i in 1:length(pmf)){
7   pmf[i] = dpois(i, n)
8 }
9 supp = (seq(0, length(pmf), length.out = length(pmf)) - n) / sqrt(n)
10
11 pdf(file="figures/bh-10-28.pdf",
12     width = 8, height = 7, bg = "white",
13     colormodel = "cmyk", paper = "A4")
14
15 df1 = data.frame(supp)
16 df1$pmf = pmf
17 df1$norm_pdf = dnorm(supp,0,1)
18
19 p = ggplot(df1) +
20   geom_line(mapping = (aes(x=supp, y = pmf)), color = "blue")+ # this
21   geom_line(mapping = (aes(x=supp, y = norm_pdf)), color = "orange")+
22   theme(axis.title.x=element_blank(),axis.title.y = element_blank())
23 p
24
25 dev.off()

```

**Ex 7.5.** Now make the plot and explain what is wrong.

**Ex 7.6.** The line marked as ‘this’ should be replaced with

---

Python Code

---

```

1 plt.plot(supp, pmf * np.sqrt(n))

```

---

R Code

---

```

1 df1$pmf = pmf * sqrt(n)

```

Explain why. (You should memorize that comparing PMFs and PDFs is not straightforward.)

**Ex 7.7.** Make the plot of  $n = 2$  and  $n = 20$ . Include both plots and explain the results.

**Ex 7.8.** The code below computes  $M_Y(s)$  where  $Y$  is the standardized version of  $X_n$ . Then it compares  $M_Y$  to the MGF of the standard normal distribution.

1. Explain how it works.

---

Python Code

---

```

1 import numpy as np
2 from scipy.stats import poisson, norm
3 import matplotlib.pyplot as plt
4
5 n = 10
6 X = poisson(n)
7 num = 50
8 S = np.linspace(-1, 1, num)
9 M = np.zeros(num)
10 pmf = X.pmf(range(100)) # this
11
12 for i in range(num):
13     for j in range(len(pmf)):
14         M[i] += np.exp(S[i] * (j - n) / np.sqrt(n)) * pmf[j]
15
16
17 plt.plot(S, M, label="M X")
18 plt.plot(S, np.exp(S * S / 2), label="M N")
19 plt.xlabel("s")
20 plt.legend()
21 plt.savefig("figures/bh-10-28-mgf1.pdf")

```

---



---

R Code

---

```

1 library(ggplot2)
2
3 n = 10
4 num = 50
5 S = seq(-1, 1, length.out = num)
6 M = rep(0, num)
7 pmf = dpois(1:100,n) #this
8
9 for(i in 1:num){
10     for(j in 1:length(pmf)){
11         M[i] = M[i] + exp(S[i] * (j - n)/sqrt(n)) * pmf[j]
12     }
13 }
14
15 pdf(file="figures/bh-10-28-mgf1.pdf",

```

```

16     width = 8, height = 7, bg = "white",
17     colormodel = "cmyk", paper = "A4")
18
19 df1 = data.frame(S)
20 df1$M_X = M
21 df1$M_N = exp(S * S / 2)
22
23 p = ggplot(df1) +
24     geom_line(mapping = (aes(x=S, y = M_X, color = "M X")))+
25     geom_line(mapping = (aes(x=S, y = M_N, color = "M N")))+
26     scale_color_manual(name = " ", values = c("M X" = "blue", "M N" = "orange"))+
27     theme(axis.title.y = element_blank())
28 p
29
30 dev.off()

```

---

**Ex 7.9.** The line marked as ‘this’ computes the PMF upfront. Why is that a good idea? Why do we stop the computation of the PMF at 100?

**Ex 7.10.** In the code below, one of the for loops is removed.

1. What is the change?
2. Will this change have an effect on the speed of the computation?
3. Which of the two alternatives do you find more readable?

---

Python Code

---

```

1  import numpy as np
2  from scipy.stats import poisson, norm
3  import matplotlib.pyplot as plt
4
5  n = 10
6  X = poisson(n)
7  num = 50
8  S = np.linspace(-1, 1, num)
9  M = np.ones(num)
10
11  for i in range(num):
12      M[i] = X.expect(lambda j: np.exp(S[i] * (j - n) / np.sqrt(n)))
13
14  plt.plot(S, M, label="M X")
15  plt.plot(S, np.exp(S * S / 2), label="M N")
16  plt.xlabel("s")
17  plt.legend()
18  plt.savefig("figures/bh-10-28-mgf2.pdf")

```

---



## R Code

```

1 library(ggplot2)
2
3 n = 10
4 num = 50
5 S = seq(-1, 1, length.out = num)
6 M = rep(1, num)
7
8 for(i in 1:num){
9   j = 1:10000
10  X = exp(S[i] * (j-n) / sqrt(n))
11  M[i] = weighted.mean(X,dpois(j,n))
12 }
13
14 pdf(file="figures/bh-10-28-mgf2.pdf",
15     width = 8, height = 7, bg = "white",
16     colmodel = "cmyk", paper = "A4")
17
18 df1 = data.frame(S)
19 df1$M_X = M
20 df1$M_N = exp(S * S / 2)
21
22 p = ggplot(df1) +
23   geom_line(mapping =(aes(x=S, y = M_X, color = "M X")))+
24   geom_line(mapping =(aes(x=S, y = M_N, color = "M N")))+
25   scale_color_manual(name = " ", values = c("M X" = "blue", "M N" = "orange"))+
26   theme(axis.title.y = element_blank())
27 p
28
29 dev.off()

```

**Ex 7.11.** Change  $n$  to 20 and to 50. Does the quality improve? Choose some value for  $n$  that you like, and include the graph.

### 7.3 BH.10.30

**Ex 7.12.** Explain how the code below implements the solution of parts a and b BH.10.30. (Of course we skip the proof of the convergence to  $g(\alpha)$  here.)

## Python Code

```

1 import numpy as np
2 from scipy.stats import bernoulli
3 import matplotlib.pyplot as plt
4
5 np.random.seed(3)

```

```

6
7 num = 100
8 p, alpha = 0.5, 0.5
9 revenue = 0.5 + 1.2 * bernoulli(p).rvs(num)
10
11 Y = np.ones(num)
12 Y[0] = 1
13 for i in range(1, num):
14     Y[i] = (1 - alpha + alpha * revenue[i]) * Y[i - 1]
15
16 print(np.log(Y[-1]) / num)
17
18 # plt.plot(Y) # this
19
20 def g(alpha):
21     return 0.5 * np.log((1 + 0.7 * alpha) * (1 - 0.5 * alpha))
22
23
24 plt.ylim(-10 * g(alpha), 10 * g(alpha)) # this
25 plt.axhline(g(alpha), c='r', lw=2, label="g")
26 plt.plot(np.log(Y)[1:] / range(1, num), label="(log Yn)/n")
27 plt.legend()
28 plt.savefig("figures/bh-10-30.pdf")

```

---

R Code

---

```

1 library(cubature)
2 set.seed(3)
3
4 num = 100
5 p = 0.5
6 alpha = 0.5
7 revenue = 0.5 + 1.2 * rbinom(num, 1, p)
8
9 Y = rep(1, num)
10 Y[1] = 1
11 for(i in 2:num){
12     Y[i] = (1 - alpha + (alpha * revenue[i])) * Y[i - 1]
13 }
14
15 print(log(Y[length(Y)]) / num)
16
17 # plot(Y, type = "l")
18
19 g = function(alpha){
20     return(0.5*log((1 + 0.7 * alpha)*(1 - 0.5 *alpha)))

```

```

21 }
22
23 pdf(file="figures/bh-10-30.pdf",
24     width = 8, height = 7, bg = "white",
25     colormodel = "cmyk", paper = "A4")
26
27 df = data.frame(log(Y)/(1:num))
28 colnames(df) = c("lnY")
29 p = ggplot()+
30     geom_hline(mapping = aes(yintercept = g(alpha), color = "g"), size = 2)+
31     geom_line(df,mapping=aes(x = 1:num, y =lnY, color = '(log Yn)/n')) +
32     scale_color_manual( name = " ", values = c("g" = "red", "(log Yn)/n" = "blue"))+
33     ylim(-10* g(alpha), 10 * g(alpha)) + #this
34     theme(axis.title.x = element_blank(), axis.title.y = element_blank())
35 p
36
37 dev.off()

```

---

**Ex 7.13.** Uncomment the relevant line to obtain a plot of  $Y_n$  as a function  $n$ . Make the plot and include it in your assignment.

**Ex 7.14.** Take num much larger, e.g., 100 000. What is  $(\log Y_n)/n$  now?

**Ex 7.15.** Print  $(\log Y_n)/n$  for various values of  $\alpha$ . For instance, take  $\alpha = 1/10, 2/10, 3/10, 4/10$ .

**Ex 7.16.** If you like a simple challenge, include a plot of  $(\log Y_n)/n$  as a function of  $\alpha$ . However, skip this if you don't have time, or interest in this extension of the problem. (Perhaps, if you find programming hard, you should do it to improve your skills :- )

#### 7.4 BH.10.39

BH.10.39.a asks about the first time such that some exponential rv exceeds a certain threshold. Part b is about when the sum of a number of r.v.s exceed a threshold. Such problems are called *hitting times*.

#### Python Code

```

1 import numpy as np
2 from scipy.stats import expon
3
4 np.random.seed(3)
5
6 X = expon(scale=1)
7
8 # part a
9

```

```

10 N = 10
11 res = np.zeros(N)
12 for i in range(N):
13     n = 0
14     while X.rvs() < 1:
15         n += 1
16     res[i] = n
17
18 print(res.mean(), res.std())
19
20 # part b
21
22 N = 100
23 res = np.zeros(N)
24 for i in range(N):
25     M = X.rvs()
26     n = 1
27     while M < 10:
28         M += X.rvs()
29         n += 1
30     res[i] = n
31
32 print(res.mean(), res.std())

```

---

R Code

---

```

1 set.seed(3)
2
3 # part a
4
5 N = 10
6 res = rep(0, N)
7 for (i in 1:N) {
8     n = 0
9     while (rexp(1, rate = 1) < 1) {
10         n = n + 1
11         res[i] = n
12     }
13 }
14
15 print(mean(res))
16 print(sd(res))
17
18 # part b
19
20 N = 100

```

```
21 res = rep(0, N)
22 for (i in 1:N) {
23   M = rexp(1, rate = 1)
24   n = 1
25   while(M < 10) {
26     M = M + rexp(1, rate = 1)
27     n = n + 1
28     res[i] = n
29   }
30 }
31
32 print(mean(res))
33 print(sd(res))
```

---

By now you have so much experience with reading code that you must be able to explain all of the code without intermediate steps to guide you.

**Ex 7.17.** Explain how the first part of the code simulates BH.10.39.a.

**Ex 7.18.** Explain how the second part of the code simulates BH.10.39.b.