

Question 1

$$\begin{aligned} \text{a) } \int_0^\infty (1 - F_X(x)) dx - \int_{-\infty}^0 F_X(x) dx &= [x(1 - F_X(x))]_0^\infty + \int_0^\infty x f_X(x) dx - [x F_X(x)]_{-\infty}^0 + \int_{-\infty}^0 x f_X(x) dx \\ &= \int_{-\infty}^\infty x f_X(x) dx \\ &= E(X) \end{aligned}$$

$$\text{b) } f_{XY}(x, y) = f_X(x)f_Y(y)$$

$$\begin{aligned} Z &= X + Y & W &= X \\ Y &= Z - W & X &= W \end{aligned}$$

$$J = \begin{vmatrix} 1 & -1 \\ 0 & 1 \end{vmatrix} = 1$$

$$f_{WZ}(w, z) = f_X(w)f_Y(z - w)|1|$$

$$\begin{aligned} f_Z(z) &= \int_{-\infty}^\infty f_X(w)f_Y(z - w) dw \\ &= \int_{-\infty}^\infty f_X(x)f_Y(z - x) dx \end{aligned}$$

Question 2

$$\begin{aligned} \text{a) } M_{X+Y}(t) &= E[e^{(X+Y)t}] \\ &= E(e^{Xt})E(e^{Yt}) \\ &= e^{\lambda(t-1)} \cdot e^{\mu(t-1)} \\ &= e^{(\lambda+\mu)(t-1)} \end{aligned}$$

$$X + Y \sim \text{Poisson}(\lambda + \mu)$$

$$\begin{aligned} \text{b) } P(X = x | X + Y = n) &= \frac{P(X=x, Y=n-x)}{P(X+Y=n)}, x = 0, 1, \dots, n \\ &= \frac{P(X=x)P(Y=n-x)}{P(X+Y=n)} \\ &= \frac{\frac{\lambda^x e^{-\lambda}}{x!} \cdot \frac{\mu^{n-x} e^{-\mu}}{(n-x)!}}{\frac{(\lambda+\mu)^n e^{-(\lambda+\mu)}}{n!}} \\ &= \binom{n}{x} \frac{\lambda^x \mu^{n-x}}{(\lambda+\mu)^n} \cdot \left(\frac{\lambda+\mu}{\lambda+\mu}\right)^x \\ &= \binom{n}{x} \left(\frac{\lambda}{\lambda+\mu}\right)^x \left(\frac{\mu}{\lambda+\mu}\right)^{n-x} \end{aligned}$$

$$P(X = x | Y + X = n) = 0, x \neq 0, 1, \dots, n$$

$$X | X + Y = n \sim \text{Binomial}\left(n, \frac{\lambda}{\lambda+\mu}\right)$$

c) Since $X | X + Y$ does not follow Poisson distribution, they are not independent

Question 3

a) $F_Y(y) = P(Y \leq y)$
 $= P(X^2 \leq y)$
 $= P(-\sqrt{y} \leq X \leq \sqrt{y})$
 $= \int_{-\sqrt{y}}^{\sqrt{y}} f_X(x) dx$
 $= F_X(\sqrt{y}) - F_X(-\sqrt{y})$

$$f_Y(y) = \frac{\partial}{\partial y} [F_X(\sqrt{y}) - F_X(-\sqrt{y})]$$

$$= \frac{f_X(\sqrt{y})}{2\sqrt{y}} + \frac{f_X(-\sqrt{y})}{2\sqrt{y}}$$

$$= \frac{f_X(\sqrt{y}) + f_X(-\sqrt{y})}{2\sqrt{y}}, y \geq 0$$

$$f_Y(y) = 0, y < 0$$

b)

i. $f_{X_1 X_2}(x_1, x_2) = \frac{1}{2\pi\sigma^2} e^{-\frac{(x_1^2 + x_2^2)}{2\sigma^2}}$

$$Y_1 = X_1^2 + X_2^2 \quad Y_2 = \frac{X_1}{\sqrt{X_1^2 + X_2^2}}$$

$$X_1 = Y_2 \sqrt{Y_1} \quad X_2 = \pm \sqrt{Y_1(1 - Y_2^2)}$$

$$J_{X_2 > 0} = \begin{vmatrix} \frac{Y_2}{2\sqrt{Y_1}} & \sqrt{Y_1} \\ \frac{\sqrt{1-Y_2^2}}{2\sqrt{Y_1}} & -\frac{Y_2\sqrt{Y_1}}{\sqrt{1-Y_2^2}} \end{vmatrix}$$

$$= -\frac{1}{2\sqrt{1-Y_2^2}}$$

$$J_{X_2 < 0} = \begin{vmatrix} \frac{Y_2}{2\sqrt{Y_1}} & \sqrt{Y_1} \\ -\frac{\sqrt{1-Y_2^2}}{2\sqrt{Y_1}} & \frac{Y_2\sqrt{Y_1}}{\sqrt{1-Y_2^2}} \end{vmatrix}$$

$$= \frac{1}{2\sqrt{1-Y_2^2}}$$

$$f_{Y_1 Y_2}(y_1, y_2) = \frac{1}{2\pi\sigma^2} e^{-\frac{y_1}{2\sigma^2}} \left| -\frac{1}{2\sqrt{1-y_2^2}} \right| + \frac{1}{2\pi\sigma^2} e^{-\frac{y_1}{2\sigma^2}} \left| \frac{1}{2\sqrt{1-y_2^2}} \right|$$

$$= \frac{1}{2\pi\sigma^2} e^{-\frac{y_1}{2\sigma^2}} \cdot \frac{1}{\sqrt{1-y_2^2}}, y_1 > 0, y_2 \in (-1, 1)$$

ii. Since the support of the joint distribution is the product set of space of Y_1 and space of Y_2 ; and

$$f_{Y_1 Y_2}(y_1, y_2) \text{ can be rewrite as } \left[g(y_1) := \frac{1}{2\sigma^2} e^{-\frac{y_1}{2\sigma^2}} \right] \left[h(y_2) := \frac{1}{\pi\sqrt{1-y_2^2}} \right], \text{ where}$$

$$Y_1 \sim \text{Exponential}(2\sigma^2) \text{ and } Y_2 \sim \frac{1}{\pi\sqrt{1-y_2^2}} \Rightarrow \int_{-1}^1 \frac{1}{\pi\sqrt{1-y_2^2}} dy_2 = \frac{1}{\pi} [\sin^{-1}(y_2)]_{-1}^1 = \frac{\pi}{\pi} = 1, \text{ both are valid pdfs}$$

Y_1 and Y_2 are independent

Question 4

$$\begin{aligned}
 \text{a) } f_Y(y) &= \int_0^\infty f_Y(y|\lambda) f_\Lambda(\lambda) d\lambda \\
 &= \int_0^\infty \frac{\lambda^y e^{-\lambda}}{y!} \cdot \frac{\lambda^{\alpha-1} e^{-\frac{\lambda}{\beta}}}{\Gamma(\alpha)\beta^\alpha} d\lambda \\
 &= \frac{1}{y!\Gamma(\alpha)\beta^\alpha} \int_0^\infty \lambda^{y+\alpha-1} e^{\left(\frac{\beta}{1+\beta}\right)\lambda} \left(\frac{\beta}{1+\beta}\right)^{y+\alpha} d\lambda \\
 &= \frac{1}{y!\Gamma(\alpha)\beta^\alpha} \Gamma(y+\alpha) \left(\frac{\beta}{1+\beta}\right)^{y+\alpha}, y = 0, 1, \dots \\
 &= \binom{y+\alpha-1}{y} \left(\frac{\beta}{1+\beta}\right)^y \left(\frac{1}{1+\beta}\right)^\alpha, \alpha \text{ is a positive integer}
 \end{aligned}$$

$$Y \sim \text{Negative Binomial}\left(\alpha, \frac{1}{1+\beta}\right)$$

$$\begin{aligned}
 E(Y) &= E[E(Y|\Lambda)] \\
 &= E(\Lambda) \\
 &= \alpha\beta
 \end{aligned}$$

$$\begin{aligned}
 \text{Var}(Y) &= \text{Var}[E(Y|\Lambda)] + E[\text{Var}(Y|\Lambda)] \\
 &= \text{Var}(\Lambda) + E(\Lambda) \\
 &= \alpha\beta(\beta + 1)
 \end{aligned}$$

$$\begin{aligned}
 \text{b) } P(Y = y|\lambda) &= \sum_{n=y}^\infty P(Y = y|N = n, \lambda) P(N = n|\lambda) \\
 &= \sum_{n=y}^\infty \binom{n}{y} p^y (1-p)^{n-y} \frac{\lambda^n e^{-\lambda}}{n!} \\
 &= \frac{e^{-\lambda}}{y!} [(1-p)\lambda]^y \sum_{m=0}^\infty \frac{[(1-p)\lambda]^m}{m!} \cdot \frac{e^{-(1-p)\lambda}}{e^{-(1-p)\lambda}}, m = n - y \\
 &= \frac{(p\lambda)^y e^{-p\lambda}}{y!}, y = 0, 1, \dots
 \end{aligned}$$

$$\begin{aligned}
 P(Y = y) &= \int_0^\infty \frac{(p\lambda)^y e^{-p\lambda}}{y!} \cdot \frac{\lambda^{\alpha-1} e^{-\frac{\lambda}{\beta}}}{\Gamma(\alpha)\beta^\alpha} d\lambda \\
 &= \frac{p^y}{y!\Gamma(\alpha)\beta^\alpha} \Gamma(y+\alpha) \left(\frac{\beta}{1+\beta p}\right)^{y+\alpha}, y = 0, 1, \dots \\
 &= \binom{y+\alpha-1}{y} \left(\frac{\beta p}{1+\beta p}\right)^y \left(\frac{1}{1+\beta p}\right)^\alpha, \alpha \text{ is a positive integer}
 \end{aligned}$$

$$Y \sim \text{Negative Binomial}\left(\alpha, \frac{1}{1+\beta p}\right)$$

Question 5

$$\begin{aligned}
 \text{a) } E(Y) &= E[E(Y|X)] \\
 &= E(X) \\
 &= \frac{1}{2}
 \end{aligned}$$

$$\begin{aligned}
 \text{Var}(Y) &= \text{Var}[E(Y|X)] + E[\text{Var}(Y|X)] \\
 &= \text{Var}(X) + E(X^2) \\
 &= \frac{1}{12} + \frac{1}{3} \\
 &= \frac{5}{12}
 \end{aligned}$$

$$\begin{aligned}
 \text{Cov}(X, Y) &= E(XY) - E(X)E(Y) \\
 &= E[E(XY|X)] - \left(\frac{1}{2}\right)\left(\frac{1}{2}\right) \\
 &= E[XE(Y|X)] - \frac{1}{4} \\
 &= E(X^2) - \frac{1}{4} \\
 &= \frac{1}{3} - \frac{1}{4} \\
 &= \frac{1}{12}
 \end{aligned}$$

$$\text{b) Since } Y|X = x \sim N(1, 1), \frac{Y}{X} \text{ and } X \text{ are independent}$$
