



# ECON2271

## Business Econometrics

### (Introductory Econometrics)

#### Week 2: Topic 2 (i)

## Topic 2: Univariate Regression and OLS

### *Agenda and learning outcomes*



- **Topic 2(i): Univariate Regression:**  $Y = b_0 + b_1X + u$ 
  - a) Introduce the classic univariate linear regression model
    - Able to distinguish between model variables and parameters;
    - Interpret meaning of model parameters and error term
  - b) Using univariate simple regression with single dummy to test for differences in group means [Continuous Y, single binary X]
    - Students able to test simple hypotheses comparing means using dummy regression
  - c) Using OLS to estimating intercept and slope for a linear function [Continuous Y, single continuous X]
    - Understand the basic premise of OLS.
    - Understand how we evaluate goodness of fit (R-squared and F-stat)
    - Students able to estimate simple univariate regression model and interpret key output correctly.
  - d) The classical assumptions behind OLS
  - e) Working with nonlinear relationships

## Topic 2: Univariate Regression and OLS



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### a) The Univariate Model

- **The general form linear univariate function:**  $Y = b_0 + b_1X$ 
  - $Y$ : Dependent variable, the value of which depends on  $X$
  - $X$ : Independent or explanatory variable = the value of which determines the value of  $Y$
  - $b_0$ : The constant or intercept term = the value which  $Y$  will take when  $X = 0$
  - $b_1$ : Captures the change\* in  $Y$  when  $X$  changes by one unit (= *slope* if  $X$  is a continuous variable; *discrete change* if  $X$  is a discrete variable)
    - \* Note on terminology: the term *change* here refers to the *difference* we observe in  $Y$  when we change the value of  $X$  by 1. We must be careful not to wrongfully imply causation where we really mean association. More on this later.
- We can use this as a theoretical model for an imagined closed system, where all relevant information is known, there is no room for error:  $Y$  only depends on  $X$  and nothing else can be going on.
- This works well in controlled environments (“in the lab”) but many environments are messy and complex (e.g. weather, economics). Then, we must try to generalize from a messy and complicated world, about which we know only some information. We dig around to look for patterns in this mess. In economics, we use econometric models for this, which means we need to account for error (denoted  $e$  or  $u$ )

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## Topic 2: Univariate Regression and OLS



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### a) The Univariate Model

- **The econometric (statistical) version:**

$$Y_i = b_0 + b_1X_i + u_i$$

- $Y_i$ : The observed value of  $Y$  for unit (e.g. individual)  $i$
  - $X_i$ : The value of  $X$  for unit (e.g. individual)  $i$
  - $b_0$  and  $b_1$ : Parameters to be estimated; their true values are not known!
  - $u_i$ : The part of  $Y_i$  that is not explained by  $(b_0 + b_1X_i)$ , i.e. the error, disturbance, noise...
- We obtain some data for  $Y$  and  $X$ , and estimate the model parameters:

$$\hat{Y}_i = \hat{b}_0 + \hat{b}_1X_i \Rightarrow Y_i = \hat{b}_0 + \hat{b}_1X_i + \hat{u}_i$$

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## Topic 2: Univariate Regression and OLS

### a) The Univariate Model



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- We can evaluate the “goodness of fit” of our model by the size of the residuals ( $\hat{u}_i$ 's): How much of the  $Y_i$ 's are explained by  $\hat{Y}_i = \hat{b}_0 + \hat{b}_1 X_i$  ?
  - 100%: our model is perfect – we can explain all the variation in  $Y$ !!
  - 0%: our model can't explain any of the variation in  $Y$ ...
  - BUT: we're not always concerned about explaining as much of  $Y$  as possible; sometimes we're more interested in finding out what we can and cannot assume about the true value of  $b_0$  and  $b_1$ , given a desired level of confidence (i.e. probability).
- In this topic we restrict ourselves to situations where  $Y$  is approximately continuous (and unlimited and cardinal...)
  1. First, we look at a model where  $X$  is binary,
  2. Second, we look at a model where  $X$  is (approximately) continuous and unlimited

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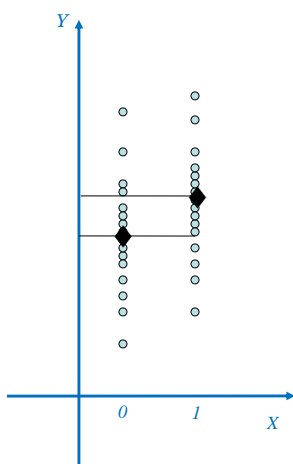
## Topic 2: Univariate Regression and OLS

### b) The Single Dummy Variable Model



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- **Movements in conditional means:** Continuous  $Y$ , binary  $X$



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$$Y_i = b_0 + b_1 X_i + u_i$$

$Y_i$  = A continuous variable (e.g. wages)

$X_i$  = Binary variable (e.g. male = 1, female = 0)

$b_0$  = Conditional mean (expected)  $Y$  when  $X = 0$ :

$b_1$  = Shift in the conditional mean (expected)  $Y$  as we move from  $X = 0$  to  $X = 1$  (e.g. from female to male)

$u_i$  = The difference between the expected  $Y$  for unit (e.g. individual)  $i$ , given the value of  $X$ , and the actual value of  $Y$  for unit  $i$

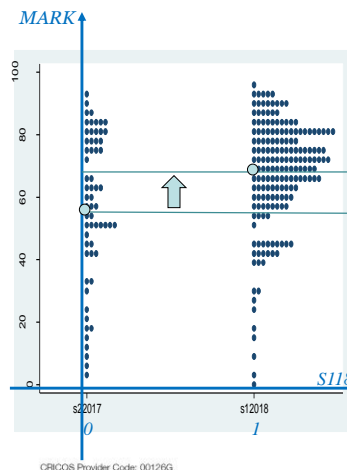
## Topic 2: Univariate Regression and OLS

### b) The Single Dummy Variable Model



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- **Example 1:** Revisit Example 2 from Lecture 1 (ECON1111 marks across two semesters)
  - I combine these sets of marks into one variable (MARK) and create a binary (0/1) dummy variable to indicate whether the mark is from semester 1 2018 (S118).



$$MARK_i = b_0 + b_1(S118_i) + u_i$$

$MARK_i$  = ECON1111 mark for student  $i$ .

$S118_i$  = Binary variable where 1 indicates the student  $i$  did ECON1111 in sem 1 2018, and 0 otherwise (i.e. sem 2 1017).

$b_0$  = Conditional mean (expected) mark when  $S118 = 0$ ;

$b_1$  = Shift in the conditional mean (expected) mark as we move from  $S118 = 0$  to  $S118 = 1$

$u_i$  = The difference between the expected mark for student  $i$ , given the value of  $S118$ , and the actual mark for student  $i$

## Topic 2: Univariate Regression and OLS

### b) The Single Dummy Variable Model



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$$MARK_i = b_0 + b_1(S118_i) + u_i$$

We can estimate this model and generate:  $MARK_i = \hat{b}_0 + \hat{b}_1(S118) + u_i$

$\hat{b}_0$  = Our estimate for  $b_0$  = the conditional mean ( $\overline{MARK} \mid S118 = 0$ )

$\hat{b}_1$  = Our estimate for  $b_1$  = the difference in the conditional means when we move from  $S118 = 0$  to  $S118 = 1$ :  $\overline{MARK} \mid S118 = 1 - \overline{MARK} \mid S118 = 0$

$\widehat{MARK}_i$  = Our estimate of student  $i$ 's mark, given the semester  $s$ /he studied in and the mean marks for each semester.

$$\widehat{MARK}_i = \hat{b}_0 + \hat{b}_1(S118_i) \Rightarrow u_i = MARK_i - \widehat{MARK}_i$$

Recall:  $H_0$ : these means are not different;  $b_1 = 0$

$H_1$ : these means are different;  $b_1 \neq 0$

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## Topic 2: Univariate Regression and OLS

### b) The Single Dummy Variable Model

#### ➤ Model estimate from Stata:

. regress MARK S118						
Source	SS	df	MS	Number of obs	=	240
Model	4781.46497	1	4781.46497	F(1, 238)	=	12.57
Residual	90534.5184	238	380.397136	Prob > F	=	0.0005
Total	95315.9833	239	398.811646	R-squared	=	0.0502
				Adj R-squared	=	0.0462
				Root MSE	=	19.504

MARK	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
S118	10.25167	2.891564	3.55	0.000	4.555344 15.948
_cons	57.2623	2.497202	22.93	0.000	52.34285 62.18174

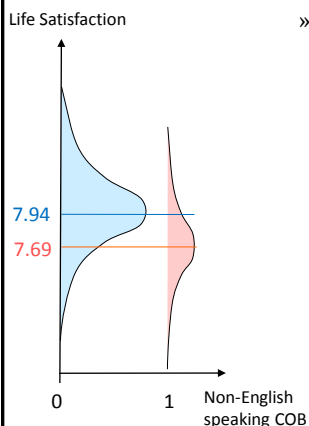
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## Topic 2: Univariate Regression and OLS

### b) The Single Dummy Variable Model

#### • Example 2: Revisit the last exercise from Lecture 1

- A sample of individuals for whom we have data on life satisfaction and information about whether they are born in an English-speaking country or not.
- We found that these two groups are different, and that we can be 95% confident that this difference is not explained by randomness.



$$» \quad LS_i = b_0 + b_1(NES_i) + u_i$$

$LS_i$  = Life satisfaction of individual  $i$ .

$NES_i$  = Binary variable where 1 indicates Non-English-Speaking COB, and 0 otherwise (i.e. English-speaking COB), for individual  $i$ .

$$b_0 = \overline{LS} \mid NES = 0 \quad (= \text{mean LS when } NES = 0)$$

$$b_1 = (\overline{LS} \mid NES = 1) - (\overline{LS} \mid NES = 0) \quad (= \text{shift in conditional mean when moving from } NES=0 \text{ to } NES=1)$$

$u_i$  = The difference between the expected LS for individual  $i$ , given the value of NES, and the actual LS of individual  $i$

## Topic 2: Univariate Regression and OLS

### b) The Single Dummy Variable Model



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```
. reg LS NES
```

Source	SS	df	MS	Number of obs	=	17,498
Model	112.002229	1	112.002229	F(1, 17496)	=	52.93
Residual	37024.5106	17,496	2.11617001	Prob > F	=	0.0000
				R-squared	=	0.0030
				Adj R-squared	=	0.0030
Total	37136.5128	17,497	2.12245029	Root MSE	=	1.4547

LS	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
NES	-.2460357	.0338189	-7.28	0.000	-.3123242 - .1797472
_cons	7.939331	.0117243	677.17	0.000	7.91635 7.962312

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### b) The Single Dummy Variable Model



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#### ➤ A note on model diagnostics: Marks example

```
. regress MARK S118
```

Source	SS	df	MS	Number of obs	=	240
Model	4781.46497	1	4781.46497	F(1, 238)	=	12.57
Residual	90534.5184	238	380.397136	Prob > F	=	0.0005
				R-squared	=	0.0502
				Adj R-squared	=	0.0462
Total	95315.9833	239	398.811646	Root MSE	=	19.504

MARK	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
S118	10.25167	2.891564	3.55	0.000	4.555344 15.948
_cons	57.2623	2.497202	22.93	0.000	52.34285 62.18174

- R-squared: The proportion of variation in Y which is explained by X
- Model F-statistic: tests whether the model explains a significant proportion of the variation in Y (so it's a test statistic for the R-squared, essentially)
  - Here: R-squared is quite low – only 5% of the variation in marks is explained by which semester you're studying in.
  - But is the objective here to explain as much variation in students marks as possible? No!
  - The objective here is only to see if the conditional means are statistically different (i.e. to test if  $b_1 = 0$ )

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### b) The Single Dummy Variable Model



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➤ A note on model diagnostics: Life Satisfaction example:

. reg LS NES									
Source	SS	df	MS	Number of obs	=	17,498	F(1, 17496)	=	52.93
Model	112.002229	1	112.002229	Prob > F	=	0.0000	R-squared	=	0.0030
Residual	37024.5106	17,496	2.11617001	Adj R-squared	=	0.0030	Root MSE	=	1.4547
Total	37136.5128	17,497	2.12245029						
LS	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]				
NES	-.2460357	.0338189	-7.28	0.000	-.3123242	-.1797472			
_cons	7.939331	.0117243	677.17	0.000	7.91635	7.962312			

- R-squared is pretty low... what does this signify?
  - Whether or not you are born in an English-speaking country matters, statistically, in terms of your life satisfaction, but it doesn't explain much of the variation in life satisfaction across individuals on its own (<1%).
- Does it matter?
  - Not unless the point of the exercise is to try to explain as much variation in LS as possible.
  - Here, the point is rather to see whether there is a significant difference between two groups.

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## Topic 2: Univariate Regression and OLS

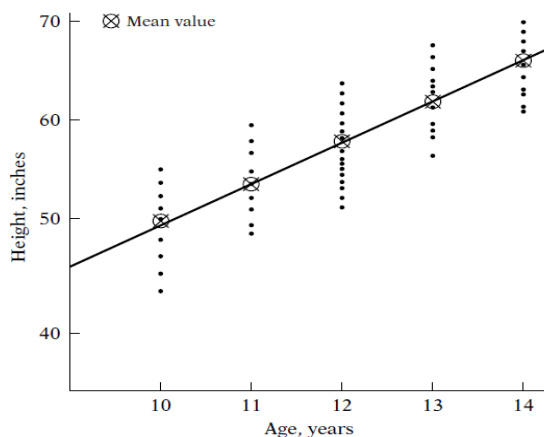
### c) The Classic Linear Regression Model



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- Movements in conditional means: Continuous Y, discrete X

$$Y_i = b_0 + b_1 X_i + u_i$$



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## Topic 2: Univariate Regression and OLS

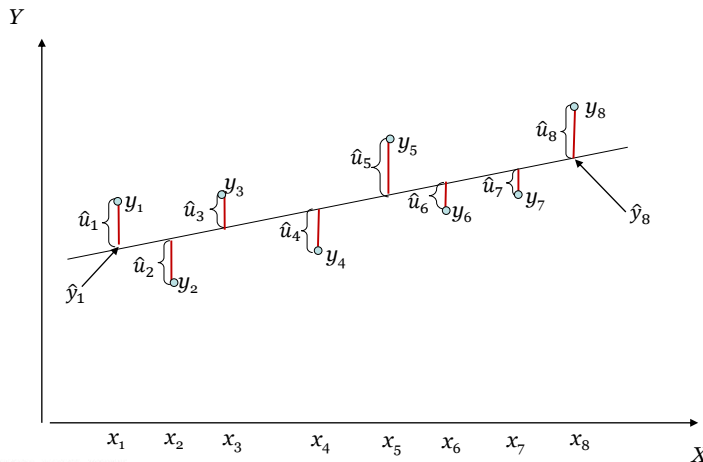
### c) The Classic Linear Regression Model



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- The classical linear regression model: Continuous Y, continuous X

$$Y_i = b_0 + b_1 X_i + u_i$$



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## Topic 2: Univariate Regression and OLS

### c) The Classic Linear Regression Model



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- The classical linear regression by Ordinary Least Squares:

$$Y_i = b_0 + b_1 X_i + u_i$$

- How it works:
  - We have some data for X and Y, which we assume are linearly related.
  - The aim of OLS is to estimate the value of the model parameters (the b's) by identifying the regression line that fits the data the best
  - Any estimate of the b-parameters will yield a predicted value of Y, given the corresponding value of X (or "conditional X"):

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### c) The Classic Linear Regression Model



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- **The Error (disturbance or residual) Term:**

- The error term of a regression model reflects, or captures, the unexplained variation in the dependent variable: the portion of variation in  $Y$  (the dependent variable) that is not explained by  $X$  (the explanatory variable/s).
- Therefore, the observed (total) variation in the dependent variable can be split into the:
  - Explained component,  $\hat{Y}$ ,
  - Unexplained component,  $\hat{u}$ .
- Hence, we can obtain explained, unexplained and total values for each pair of  $X$  and  $Y$ , square these, and thus obtain total sum of squares (SST), explained sum of squares (SSE), and residual sum of squares (SSR).

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## Topic 2: Univariate Regression and OLS

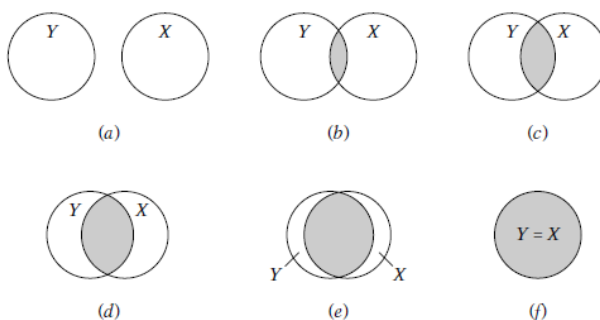
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- **The Coefficient of Determination: R-squared**

- The R-squared ( $R^2$  or  $r^2$ ) statistic of the univariate regression estimate tells you the proportion of variation in  $Y$  which is explained by variation in  $X$ :
  - R-squared =  $SSR/SST$ .



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### c) The Classic Linear Regression Model

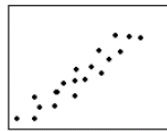


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- Slope and correlation:**

- The slope of the regression line =  $\frac{\Delta Y}{\Delta X}$
- The coefficient of correlation ( $r$ ) between Y and X gives you a measure of how closely they are aligned.
- $r$  is bounded by -1 (=perfectly negatively correlated) and +1 (=perfectly positively correlated).
- $r = \pm\sqrt{r^2}$
- So  $r$  and  $r^2$  are closely connected but not the same!

Degree of Correlation



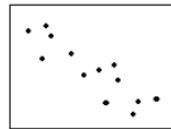
Strong Positive



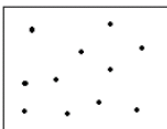
Strong Negative



Weak Positive



Moderate Negative



None



Weak Negative

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## Topic 2: Univariate Regression and OLS

### c) The Classic Linear Regression Model



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- Example 3:** Estimating the health-income gradient

- The health-income gradient refers to the association between income and health
- I obtain some income and health data from a large sample of Australian households (2014):
  - $PH_i$  = The physical health index for individual  $i$ ; PH is a 0-100 index aggregated from responses to a set of survey questions;
  - $INC_i$  = Equivalised annual disposable (after-tax) household income of individual  $i$ ;
- I estimate the health-income gradient by estimating the simple linear regression model:

$$PH_i = b_0 + b_1(INC_i) + u_i$$

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### c) The Classic Linear Regression Model



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- Example 3: Estimating the health-income gradient**

```
. regress ph realeqinc
```

Source	SS	df	MS	Number of obs	=	15,586
Model	185890.094	1	185890.094	F(1, 15584)	=	388.09
Residual	7464497.05	15,584	478.984667	Prob > F	=	0.0000
				R-squared	=	0.0243
				Adj R-squared	=	0.0242
Total	7650387.15	15,585	490.881434	Root MSE	=	21.886

ph	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
realeqinc	.0000751	3.81e-06	19.70	0.000	.0000676 .0000826
_cons	70.67294	.2749388	257.05	0.000	70.13402 71.21185

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### c) The Classic Linear Regression Model



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- Example 3: What happens if we scale the variables?**

- Sometimes we might want to scale a variable to reflect very large or very small values. Here, PH is measured on a 0-100 scale, while INC has a huge range:

```
. summarize realeqinc, detail
```

realeqinc					
Percentiles			Smallest		
1%	7953.2		0		
5%	17694.17		0		
10%	21796.15		0	Obs	23,107
25%	30518.57		0	Sum of Wgt.	23,107
50%	44927.33			Mean	53547.47
			Largest		
75%	64288.8		834071	Std. Dev.	42748.76
90%	91050		834071	Variance	1.83e+09
95%	112484		834071	Skewness	5.994644
99%	197658.3		834071	Kurtosis	71.47739

- See what happens if I generate a new variable I call INC = realeqinc/10,000, and use this instead:

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## Topic 2: Univariate Regression and OLS

### c) The Classic Linear Regression Model



- **Example 3:** What happens if we scale the variables?

```
. generate inc = realeqinc/10000
```

```
. regress ph inc
```

Source	SS	df	MS	Number of obs	=	15,586
Model	185890.095	1	185890.095	F(1, 15584)	=	388.09
Residual	7464497.05	15,584	478.984667	Prob > F	=	0.0000
				R-squared	=	0.0243
				Adj R-squared	=	0.0242
Total	7650387.15	15,585	490.881434	Root MSE	=	21.886

ph	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
inc	.7510782	.0381257	19.70	0.000	.6763474 .825809
_cons	70.67294	.2749388	257.05	0.000	70.13402 71.21185

- Note: Now the “inc” coefficient reflects the change in mean PH as we compare people whose incomes are \$10,000 higher. (SE and CI also change to reflect the new scale)
- Nothing else changes.

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## Topic 2: Univariate Regression and OLS

### c) The Classic Linear Regression Model



- **Example 3:** What happens if we restrict the sample?

- Suppose I want to restrict the sample to only include individuals with an income less than \$100,000. Why might I want to do this?

```
. regress ph inc if inc<10
```

Source	SS	df	MS	Number of obs	=	14,278
Model	436002.383	1	436002.383	F(1, 14276)	=	919.39
Residual	6770086.28	14,276	474.228515	Prob > F	=	0.0000
				R-squared	=	0.0605
				Adj R-squared	=	0.0604
Total	7206088.67	14,277	504.734095	Root MSE	=	21.777

ph	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
inc	2.61757	.0863272	30.32	0.000	2.448357 2.786782
_cons	62.17131	.4391525	141.57	0.000	61.31051 63.03211

- Income coefficient is now 3.5 times larger!!!! Why??
- t-stat for income coefficient = 30.32  $\Rightarrow$  even higher than before
- The R-squared = 0.06  $\Rightarrow$  income explains about 6% of the variation in PH observed across individuals in this sample, which 2.5 times more than with the full sample. Why?

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### c) The Classic Linear Regression Model



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- **Example 3:** What happens if I restrict the sample further?

```
. regress ph inc if inc<5
```

Source	SS	df	MS	Number of obs	=	8,751
Model	277249.223	1	277249.223	F(1, 8749)	=	506.14
Residual	4792436.25	8,749	547.769602	Prob > F	=	0.0000
				R-squared	=	0.0547
				Adj R-squared	=	0.0546
Total	5069685.47	8,750	579.392625	Root MSE	=	23.404

ph	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
inc	5.18068	.230277	22.50	0.000	4.729283 5.632077
_cons	54.09793	.7860433	68.82	0.000	52.5571 55.63876

- The income coefficient almost doubles in size!!
- What does this mean?
- What can I conclude?

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## Topic 2: Univariate Regression and OLS

### c) The Classic Linear Regression Model



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- Some **really, really** important things to bear in mind:



#### Beware of the OLS small print!

- The reliability of OLS estimates rests on a set of assumptions – many of them concerning the error terms. If these are violated, then problems ensue. Different contexts are associated with different common problems. We'll talk more about this next week.



#### The slope parameter does not measure correlation!

- Correlation is a term which refers to the strength or degree of linear association between two variables. If all points lie ON the linear regression line, we have zero error and perfect (100%) correlation. The correlation coefficient is then 1. If X and Y are all over the place, with no systematic linear pattern, then we have 0% correlation (correlation coefficient = 0).
- The slope parameter or coefficient measures  $\frac{\Delta Y}{\Delta X}$ , which can take any value at all; its interpretation depends on the scales on which Y and X are measured.



#### Association is NOT the same as causation!

- The fact that Y is observed to increase along with X does not mean that X causes Y!

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## Topic 2: Univariate Regression and OLS

### *Guide to further reading*



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- **Topic 2(i): Univariate Regression:**  $Y = b_0 + b_1X + e$ 
  - a) Introduce the classic univariate linear regression model
    - Woodridge Ch 4.1; Gujarati Ch 1
  - b) Using univariate simple regression with single dummy to test for differences in group means [Continuous Y, single binary X]
    - Gujarati Ch 2.2 (at a pinch)
  - c) Using OLS to estimating intercept and slope for a linear function [Continuous Y, single continuous X]
    - Gujarati Ch 3.1. Woodridge Ch 4.2 – 4.4