



ECON2271

Business Econometrics

(Introductory Econometrics)

Week 3:Topic 2 (ii)

Topic 2(ii): Univariate Regression and OLS *Agenda and learning outcomes*



- Topic 2 (ii): Univariate Regression: $Y = b_0 + b_1 X + e$
 - a) Introduce the classic univariate linear regression model
 - b) Using univariate simple regression with single dummy to test for differences in group means [Continuous Y, single binary X]
 - c) Using OLS to estimating intercept and slope for a linear function [Continuous Y, single continuous X]
 - d) The classical assumptions behind OLS:
 - ➤ Understand meaning of BLUE estimates;
 - > Understand the assumptions about Y and its relationship with X
 - > Understand key assumptions about disturbances (e or u)
 - e) Working with nonlinear relationships:
 - > Recognise some common nonlinear functional forms, and understand how to specify a function of Y which is linear in the model parameters.
 - > Students able to interpret model estimates for simple linear functions, when Y is linear in X, and also some linearized non-linear functions, where Y is not linear in X but is linear in a function of X (i.e. linear in the parameters).
 - Students able to apply techniques required to identify correct functional form and linearized model specification.



OLS and BLUE estimates:

First, a recap: We want to estimate the linear model:

$$Y_i = b_0 + b_1 X_i + e_i$$

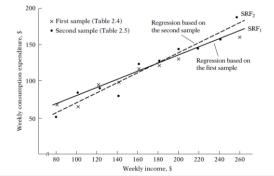
- Y_i: The value of Y for unit (e.g. individual) i
- X_i : The value of X for unit (e.g. individual) i
- e_i : The part of Y_i that is not explained by $(b_0 + b_1 X_i)$, i.e. the error, disturbance, noise... $(=u_i)$
 - \triangleright We use data for X and Y to obtain estimates of the parameters b_0 and b_1
 - The estimator is Ordinary Least Squares (OLS): identifies the straight line of best fit; the set of parameters which minimizes the sum of squared deviations (e's) away from the regression line.
 - OLS will provide the Best Linear Unbiased and Efficient (BLUE) model estimates, though this is only guaranteed under certain favourable conditions... These conditions are identified by a set of Classical Assumptions of OLS.
- Despite the fact that the Classical Assumptions are quite often violated, strictly speaking, OLS tends to be hard to beat. However, that doesn't mean we shouldn't be careful, and any analysis where there is reason to believe that there may be a problem should make attempts to evaluate this and look for solutions, and/or provide some robustness tests.

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Topic 2: Univariate Regression and OLS *d) The Classical Assumptions of OLS*



- A quick reminder about randomness and sampling in the context of regression:
 - > We estimate the parameters of a model using data drawn from a sample
 - Hence, we are using the sample data to make inferences (i.e. best possible guesses) about the true value of the model parameters
 - > If we estimate the model based on a different sample, we are unlikely to arrive at exactly the same model estimates, but hopefully they should be close.
 - Clearly, if we increased our sample, or combined samples, we'd get a better estimate..





- Best Linear Unbiased and Efficient (BLUE) estimates:
 - > What do we mean by unbiasedness?
 - The estimated sample mean should be close to the population mean
 - Darts analogy: your aim has good direction
 - > What do we mean by efficiency? Precision..
 - An efficient estimate is precise, i.e. it has a narrow confidence interval
 - Darts analogy: your throws are precise



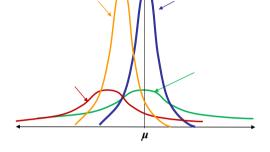






unbiased, imprecise





Topic 2: Univariate Regression and OLS d) The Classical Assumptions of OLS

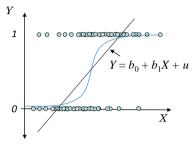


- For OLS to yield estimates that are BLUE, the errors must be random and evenly distributed without any systematic patterns:
 - Errors should be normally distributed around a mean of zero:
 - Errors should have constant variance, which doesn't change with X. That is, errors should be homoscedastic, rather than heteroscedastic.
 - > Errors should be uncorrelated with one another: That is, errors should not be autocorrelated or serially correlated.
 - Errors should be uncorrelated with X: That is, X needs to be exogenous, rather than endogenous.
 - In practice, we avoid problems with the errors by sticking to these rules, as closely
 - 1. The dependent variable (Y) is (at least approximately) unlimited, continuous and
 - 2. We are estimating a model which is linear in its parameters.
 - 3. We are aware of the key symptoms, consequences and solutions for some common problems with errors which tend to arise in specific settings.



• What if the dependent variable (Y) isn't unlimited, continuous and cardinal?

- Consider what happens if Y is binary:
- E.g: Y = indicator for whether a labour force participant is employed (=1) or unemployed (=0)
 X = Years of schooling



- Clearly, the Y values are lumped in two groups.
- These data are not well spaced around the linear regression line which OLS will produce, and the errors will be huge!
- · Consequently, OLS will yield a model with:
 - > Large SEs and low R-squared
 - Predicted Y's are continuous will be <0 and >1 for some many values of X.
 - Nonetheless, OLS will probably tell you that there is a positive relationship between X and Y.
- As you can see, OLS is not well suited to regressions where Y is binary.
- The solution is to estimate a probability model, which is quite different.
- We will cover this later in the unit. For now, just be aware of this limitation for our classical linear regression model.

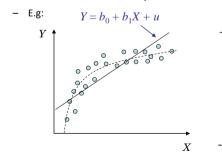
- Can you think why an ordinal Y would also cause problems? Or censored or truncated Y?

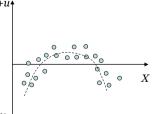
Topic 2: Univariate Regression and OLS d) The Classical Assumptions of OLS



· What if the relationship between X and Y isn't linear?

 If there are nonlinear patterns in the data, but we impose a linear functional form, then our model estimates will clearly not be BLUE...

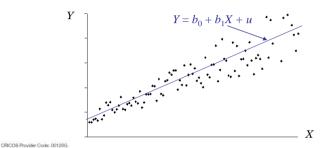




— In order for OLS to yield BLUE estimates, we require Y to be linear in the model parameters. Hence, if Y does not move linearly with Y we may be able to linearise the relationship by transforming X by some function. In this example it looks like the association between X and Y might be logarithmic. That means Y will be nonlinear in X, but it will be linear in InX. More on this later in this lecture...



- "Nonspherical" errors: What if errors are heteroskedastic?
 - Errors must be homoskedastic, i.e. have constant variance for all X
 - For example, let's say that we are examining how consumption (Y) varies with income (X). We observe a general positive association here, but we might also observe some heteroscedasticity in the data: People with higher incomes may exhibit much more variation in their consumption. This means our ability to predict E(Y|X) will diminish with higher X.
 - There are various tests you can perform to detect heteroscedasticity, but nothing beats having a look: then you can see exactly where the variances are blowing out too, which is useful.

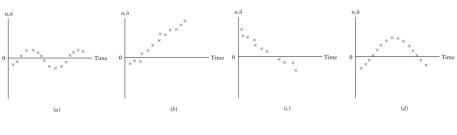




d) The Classical Assumptions of OLS

Topic 2: Univariate Regression and OLS

- · "Nonspherical errors: What if the errors are correlated with one another?
 - Means there are patterns in the error terms: cov $(u_i, u_i) \neq 0$
 - Often referred to as autocorrelation or serial correlation.
 - These problems arise most often in time-series data, which we haven't covered yet, but:
 - In a cross-sectional setting we might find errors "spilling over" to other affected units: if I'm less happy than I should be, given my other observable characteristics, chances are the rest of my household is too. So our errors are correlated.



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- Consequences and solutions for nonspherical errors:
 - > Happily, nonspherical errors do NOT cause OLS estimates to be biased.
 - However, variances are no longer reliable, which means hypothesis testing won't be reliable either.
 - But, happily, there are ways to address that problem, and modern statistical packages make this easy: you can usually select to estimate "robust standard errors", and all is well
 - > Further, OLS estimates won't necessarily be the most efficient, producing the parameters with the lowest variance.
 - ➢ But, happily, there are ways to address this problem too: There is another estimator which will be BLUE, called Generalised Least Squares (GLS) estimator. Instead of minimising the sum of squared residuals, GLS minimises an appropriately weighted sum of squared residuals. (And you can select the appropriate weighting method to suit your specific problem.)

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Topic 2: Univariate Regression and OLS d) The Classical Assumptions of OLS



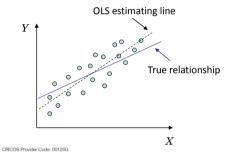
- Consequences and solutions for nonspherical errors: A caveat!
 - ➤ Later in the unit we will consider multivariate regression models, where there are more than one explanatory (right-hand-side) variable. Then we will be able to specify models where Y can be explained by a combination of variables. This is very exciting, but with all fun things comes responsibility: the problem of correct model specification becomes more complex. Not only do we have to worry about linearity issues, we also have to worry about selecting the correct combination of variables.
 - So, before we try to fix a problem with errors, Kennedy (in his text, chapter 5) reminds us to stop and think:

"Funny-looking errors should at first be interpreted as signalling a specification error, not a nonspherical error".

- > As always is the case: the best way to fix a problem is to find the root cause of it, and fix it there. Don't just fix the symptoms, unless it's your last option.
- Just because your software provides lots of fancy options and correction methods in your estimation procedure doesn't mean you should take out the full artillery. You should make sure you understand what it does and why, and you should remember that by trying to fix one problem you can create another one which is even bigger...



- . Endogeneity: What if the errors are correlated with X?
 - The <u>consequence</u>s of endogeneity is <u>biased estimates</u>, and this is much harder to circumvent and fix than problems associated with "nonspherical" errors.
 - It is, essentially, a mostly invisible problem which does not reveal itself by funny-looking errors or any other easily observable feature. Instead, we have to determine whether it is likely to be a problem, in any given setting, using logic, theory, intuition and evidence. If we can't make a good case for why we don't have a problem, then we have to assume that we have one, and try to deal with it as best we can.



There are 3 main root-causes of endogeneity:

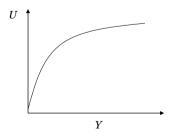
- 1. Omitted variables
- 2. Simultaneity and reverse causality
- 3. Measurement error

The problem of endogenous regressors is complex and difficult to explain in simple terms, so we will return to this later

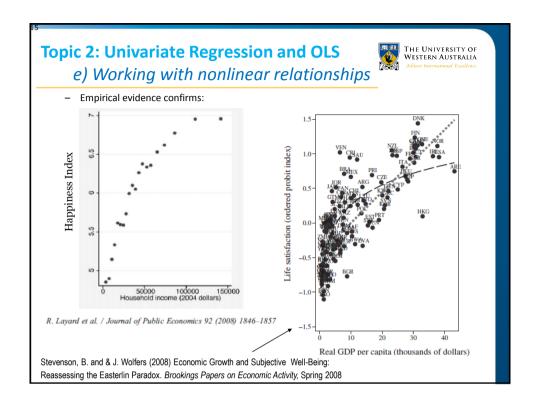
Topic 2: Univariate Regression and OLS *e) Working with nonlinear relationships*

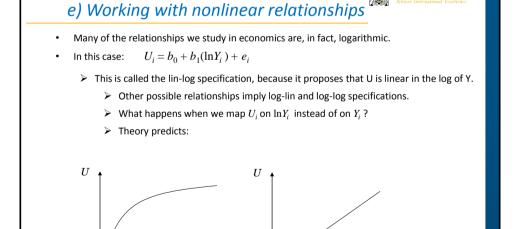


- . What do we do if we are working with nonlinear functions?
 - Many relationships we work with in Economics are nonlinear...
 - E.g. utility functions
 - U = some measure of utility
 - Y = some measure of income
 - Theory predicts:



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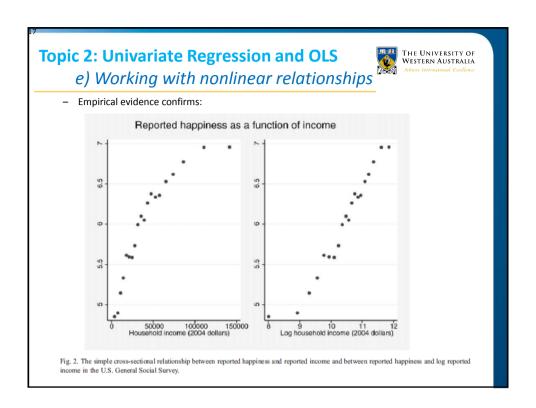


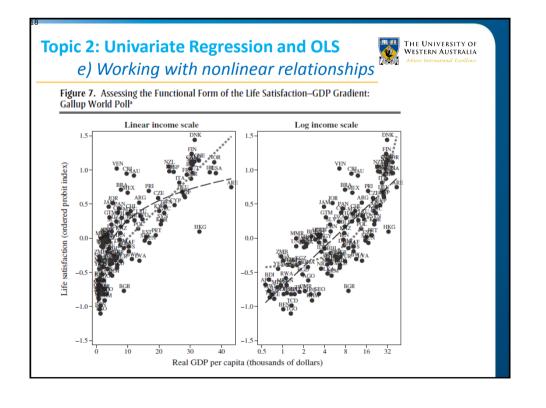
lnY

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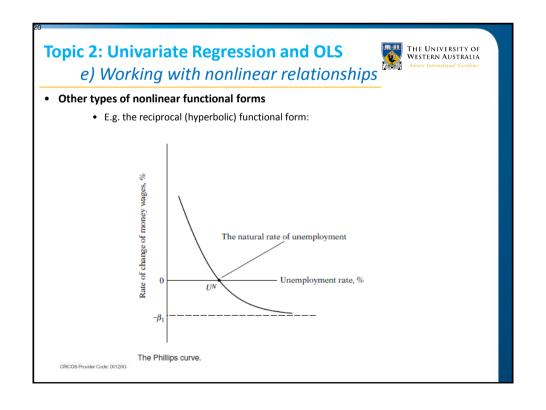
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Y



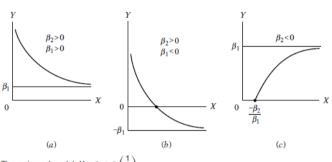


Topic 2: Univariate Regression and OLS THE UNIVERSITY OF WESTERN AUSTRALIA e) Working with nonlinear relationships • Other types of nonlinear functional forms We find other types of nonlinear relationships too... • E.g. The log-log functional form: ln YLog of quantity demanded Quantity demanded $Y = \beta_1 X_i^{-\beta_2}$ $\ln Y = \ln \beta_1 - \beta_2 \ln X_i$ ln XPrice Log of price (a) (b) CRICOS Provider Code: 00126G





- Other types of nonlinear functional forms
 - Reciprocal (hyperbolic) functional forms:



The reciprocal model: $Y = \beta_1 + \beta_2 \left(\frac{1}{X}\right)$

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Topic 2: Univariate Regression and OLS *e) Working with nonlinear relationships*



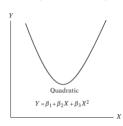
- Other types of nonlinear functional forms
 - An overview of variants of logarithmic and reciprocal (hyperbolic) functional forms:

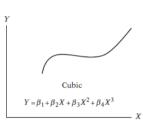
Model	Equation	Slope $\left(=\frac{dY}{dX}\right)$	Elasticity $\left(=\frac{dY}{dX}\frac{X}{Y}\right)$
Linear	$Y = \beta_1 + \beta_2 X$	β2	$\beta_2 \left(\frac{X}{Y}\right)^*$
Log-linear	$\ln Y = \beta_1 + \beta_2 \ln X$	$\beta_2\left(\frac{Y}{X}\right)$	β_2
Log-lin	$\ln Y = \beta_1 + \beta_2 X$	$\beta_2(Y)$	$\beta_2(X)^*$
Lin-log	$Y = \beta_1 + \beta_2 \ln X$	$\beta_2\left(\frac{1}{X}\right)$	$\beta_2 \left(\frac{1}{Y}\right)^*$
Reciprocal	$Y = \beta_1 + \beta_2 \left(\frac{1}{X}\right)$	$-\beta_2\left(\frac{1}{X^2}\right)$	$-\beta_2 \left(\frac{1}{XY}\right)^*$
Log reciprocal	$\ln Y = \beta_1 - \beta_2 \left(\frac{1}{X}\right)$	$\beta_2 \left(\frac{Y}{X^2} \right)$	$\beta_2 \left(\frac{1}{X}\right)^*$

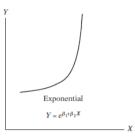
Note: * indicates that the elasticity is variable, depending on the value taken by X or Y or both. When no X and Y values are specified, in practice, very often these elasticities are measured at the mean values of these variables, namely, X and Y.



- Other types of nonlinear functional forms
 - Polynomial and exponential functional forms:







$$\frac{dY}{dX} = \beta_2 + 2\beta_3 X;$$

$$\frac{dY}{dX} = \beta_2 + 2\beta_3 X + 3\beta_3 X^2 \qquad \frac{dY}{dX} = \beta_2 e^{\beta_1 + \beta_2 X}$$

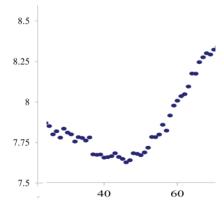
$$\frac{dY}{dX} = \beta_2 e^{\beta_1 + \beta_2 X}$$

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Topic 2: Univariate Regression and OLS e) Working with nonlinear relationships

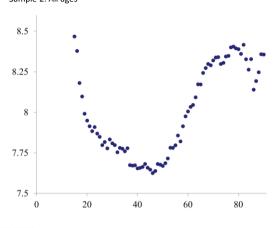


- Other types of nonlinear functional forms
 - E.g. What is the relationship between age and life satisfaction?
 - Sample 1: People of "working age"





- Other types of nonlinear functional forms
 - E.g. What is the relationship between age and life satisfaction?
 - Sample 2: All ages



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Topic 2: Univariate Regression and OLS *e) Working with nonlinear patterns*



- The general approach to working with nonlinear patterns
 - ➤ The important condition for OLS is that Y is linear *in the model parameters*:
 - E.g. for the lin-log model, Y is NOT linear in X, but it is linear in lnX
 - E.g. for the quadratic model, Y is NOT linear in X, but it is linear in the X and X²
 - As long as Y is linear in the parameters, we can estimate the model using OLS
 - > Therefore it is important to establish the correct functional form
 - ➤ How do you know?
 - Look to the theory: e.g. the Phillips Curve; the Utility Curve
 - ➤ Look at the data: Can you see any nonlinear patterns?
 - Check for departures from linearity by adding a squared term (X²) to a standard linear model.



· The general approach to working with nonlinear patterns

- Once you have discovered a departure from linearity in your data, you need to figure out what type of nonlinearity
 - ➤ Just because the coefficient for X² is significant doesn't mean the quadratic specification is best..
 - ➤ A quadratic specification is great for picking up different types of nonlinearity, which can be:
 - Positive monotonic, i.e. increasing for all feasible values of X, but not necessarily at the same rate;
 - Negative monotonic, i.e. decreasing for all feasible values of X, but not necessarily at the same rate;
 - · Concave up, or U-shaped, i.e. it has a minimum turning-point; or
 - · Concave down, or inverted-U-shaped, i.e. it has a maximum turning-point.
 - ➤ However, there might be further "moments" in your data: throw in an X³ too, and see what happens.
 - You might come up with some different alternative functional forms. Estimate these, and check the goodness of fit.

Never underestimate the powerful eye-ball test: Plot the data and look at it!

Topic 2: Univariate Regression and OLS *e) Working with nonlinear patterns*



• Example 1: Financial satisfaction and income

- I have some data on Financial Satisfaction and Income, from wave 14 of the HILDA survey.
- Financial Satisfaction (FS) is measured on a scale between 0 (totally dissatisfied) and 10 (totally satisfied).
- I measure income (INC) as annual equivalised disposable household income (why do you think I use household income instead of personal income?)
- I want to estimate the association between FS and INC, but I am not sure of the correct functional form. I try lin-lin, lin-log, and quadratic:

(1)
$$FS_i = b_0 + b_1 INC_i + u_i$$

(2)
$$FS_i = b_0 + b_1 INC + b_2 (INC_i)^2 + u_i$$

(3)
$$FS_i = b_0 + b_1(\ln INC_i) + u_i$$



- Example 1: Financial satisfaction and income
 - (1) $FS_i = b_0 + b_1 INC_i + u_i$

. regress FS 1	INC						
Source	SS	df	MS		r of ob		17,488
				F(1,	17486)	=	780.17
Model	3735.78915	1	3735.78915	Prob	> F	=	0.0000
Residual	83730.1486	17,486	4.78841065	R-squ	ared	=	0.0427
				- Adj R	-square	d =	0.0427
Total	87465.9377	17,487	5.00176918	Root	MSE	=	2.1882
	I						
FS	Coef.	Std. Err.	t	P> t	[95%	Conf.	Interval]
INC	.0000102	3.65e-07	27.93	0.000	9.48e	-06	.0000109
_cons	5.939023	.0259705	228.68	0.000	5.888	118	5.989927

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Topic 2: Univariate Regression and OLS *e) Working with nonlinear patterns*



- Example 1: Financial satisfaction and income
 - (1) $FS_i = b_0 + b_1 INC_i + u_i$

(room for extra notes)



• Example 1: Financial satisfaction and income

(2)
$$FS_i = b_0 + b_1 INC + b_2 (INC_i)^2 + u_i$$

. generate INC	CSQ=INC^2						
. regress FS	INC INCSQ						
Source	ss	df	MS		er of obs		17,488
					17485)	=	576.04
Model	5406.83368	2	2703.4168	4 Prob	> F	=	0.0000
Residual	82059.1041	17,485	4.6931143	3 R-sq	uared	=	0.0618
				- Adi	R-squared	=	0.0617
Total	87465.9377	17,487	5.0017691	8 Root	MSE	=	2.1664
FS	Coef.	Std. Err.	t	P> t	[95% Co	onf.	Interval]
INC	.0000202	6.44e-07	31.45	0.000	.0000	19	.0000215
INCSQ	-2.63e-11	1.39e-12	-18.87	0.000	-2.90e-	11	-2.36e-11
_cons	5.520841	.0339438	162.65	0.000	5.45430	8 (5.587374

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Topic 2: Univariate Regression and OLS *e) Working with nonlinear patterns*

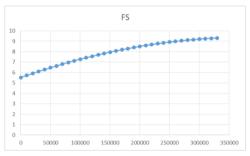


• Example 1: Financial satisfaction and income

(2) $FS_i = b_0 + b_1 INC + b_2 (INC_i)^2 + u_i$ (room for extra notes)



- Example 1: Financial satisfaction and income $FS_i = b_0 + b_1 INC + b_2 (INC_i)^2 + u_i$
 - You can find the theoretical maximum by using calculus (or the short-cut formula for quadratic functions), and then see whether this maximum lies within a feasible range:
 - ightharpoonup X-value for turning point = $-\frac{b_1}{2b_2}$ =
 - > How many people will have incomes above this value? (52, as it turns out... about 0.3% of the sample)
 - > Even better: Plot the model in Excel, using a range of feasible INC values.



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Topic 2: Univariate Regression and OLS *e) Working with nonlinear patterns*



• Example 1: Financial satisfaction and income:

generate LN: regress FS 1	INC=1n(INC+1)	$(3) F_{i}$	$S_i = b_0 +$	$b_1(\ln I)$	$NC_i) + \iota$	ι_i	
Source	ss	df	MS		r of obs	=	17,488
Model	3346.47095	1	3346.47095			_	0.000
Residual	84119.4668	17,486	4.81067521	l R-squ	ared	=	0.038
				- Adj R	squared	=	0.038
Total	87465.9377	17,487	5.00176918	Root I	MSE	=	2.193
FS	Coef.	Std. Err.	t	P> t	[95% Con	f. I	nterval
LNINC	.5235245	.0198494	26.37	0.000	.4846178		.562431
_cons	.90387	.2127524	4.25	0.000	.4868541		1.32088

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• Example 1: Financial satisfaction and income:

(3)
$$FS_i = b_0 + b_1(\ln INC_i) + u_i$$

(Room for extra notes)

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Topic 2: Univariate Regression and OLS *e) Working with nonlinear patterns*



- Example 1: Financial satisfaction and income
 - You can actually also check whether there are departures from linearity in the lin-log model:
 Throw in a squared log term and see if its coefficient is statistically significant!

. generate LN	INCSQ=LNINC^2	(4) E	S = b	h (ln II	V(C) + I	. (ln	INC \2 + u
. regress FS 1	LNINC LNINCSQ	(4) F	$S_i - \nu_0 +$	$\nu_1(mn)$	$(C_i) + \iota$	/ ₁ (III.	$INC_i)^2 + u_i$
Source	SS	df	MS		er of obs	-	17,488
Model	5344.08821	2	2672.0441	, ,	17485)	=	568.92 0.0000
Residual	82121.8495	17,485	4.6967028	6 R-sq	uared	=	0.0611
				- Adj	R-squared	1 =	0.0610
Total	87465.9377	17,487	5.0017691	8 Root	MSE	=	2.1672
	2 6	0.1.7		75.11.1			
FS	Coef.	Std. Err.	t	P> t	[95% (oni.	Interval]
LNINC	7656264	.0655139	-11.69	0.000	89404	102	6372127
LNINCSQ	.0769897	.0037331	20.62	0.000	.06967	724	.084307
_cons	5.834583	.3183587	18.33	0.000	5.2105	68	6.458598



• Example 1: Financial satisfaction and income

(4)
$$FS_i = b_0 + b_1(\ln INC_i) + b_1(\ln INC_i)^2 + u_i$$

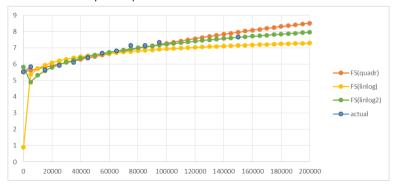
(Room for extra notes)

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Topic 2: Univariate Regression and OLS *e) Working with nonlinear patterns*



- Example 1: Financial satisfaction and income
 - Let's see how they all compare:



- > They all perform well across the mid-range of incomes, except for the lin-log model at low incomes!!! Terrible....
- The lin-log model implies a slipe that is too flat across the mid-range. The quadratic specification is better, but over-shoots the mark for higher income levels.
- > The squared log model performs best is most flexible.



- Example 2: Life satisfaction and age
 - I have some data on Life Satisfaction and age, from wave 14 of the HILDA survey.
 - Life Satisfaction (LS) is measured on a scale between 0 (totally dissatisfied) and 10 (totally satisfied).
 - I want to estimate the association between LS and age. The literature tells me that happiness is U-shaped across the life span. If so, the correct specification is quadratic:

$$LS_i = b_0 + b_1 AGE_i + b_2 AGE_i^2 + u_i$$

. regress ls a	age agesq						
Source	SS	df	MS		r of obs	=	17,503
				- F(2,	17500)	=	185.50
Model	771.455339	2	385.7276	7 Prob	> F	=	0.0000
Residual	36389.6511	17,500	2.0794086	3 R-squ	ared	=	0.0208
				- Adj R	-squared	=	0.0206
Total	37161.1064	17,502	2.1232491	4 Root	MSE	=	1.442
ls	Coef.	Std. Err.	t	P> t	[95% Co	onf.	Interval]
age	0496758	.0028503	-17.43	0.000	055262	27	044089
agesq	.0005481	.0000293	18.72	0.000	.000490	7	.0006055
_cons	8.835998	.0622126	142.03	0.000	8.7140	55	8.957941

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Topic 2: Univariate Regression and OLS *e) Working with nonlinear patterns*



- Example 2: Life satisfaction and age
 - But is there additional nonlinearity in this relationship? Throw in a cubed term...

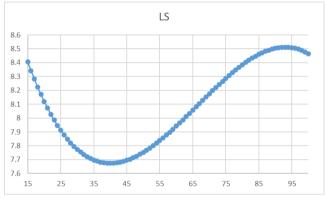
$$LS_i = b_0 + b_1 AGE_i + b_2 AGE_i^2 + b_3 AGE_i^3 + u_i$$

Source	SS	df	MS	Numb	er of obs	=	17,503
				- F(3,	17499)	=	144.04
Model	895.56878	3	298.522927	7 Prob	> F	=	0.0000
Residual	36265.5377	17,499	2.07243486	6 R-sq	uared	=	0.0241
				- Adj	R-squared	=	0.0239
Total	37161.1064	17,502	2.12324914	Root	MSE	=	1.4396
ls	Coef.	Std. Err.	t	P> t	[95% C	onf.	Interval
ls age	Coef.	Std. Err.	-12.30	P> t	[95% C		Interval]
						48	
age	1254487	.0101965	-12.30	0.000	14543	48	105462



• Example 2: Life satisfaction and age

$$LS_i = b_0 + b_1 AGE_i + b_2 AGE_i^2 + b_3 AGE_i^3 + u_i$$



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Topic 2(ii): Univariate Regression and OLS *Suggested guide to further reading*



- Topic 2 (ii): Univariate Regression: $Y = b_0 + b_1 X + e$
 - a) The classical assumptions behind OLS:
 - > Wooldridge Ch 4
 - > Gujarati Ch 3.2-3.4 (and Ch 4.1-4.3)
 - > Kennedy Ch 2.5-2.8
 - e) Working with nonlinear relationships:
 - > Gujarati Ch 2.3; 6.4-6.8
 - ➤ Wooldridge Ch 4.5