

<Machine Learning HW#2>

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1.32. Consider a vector x of continuous variables with distribution $p(x)$ and corresponding entropy $H(x)$. Suppose that we make a nonsingular linear transformation of x obtain a new variable $y = Ax$. Show that the corresponding entropy is given by $H[y] = H[x] + \ln |A|$ where $|A|$ denotes determinant of A .

Proof)

Vector x 가 distribution $p(x)$ 를 갖는 continuous variables이므로, vector x 의 entropy는 아래와 같이 정의된다.

$$H(x) = - \int p(x) \ln p(x) dx$$

위 식을 토대로, $H[y] = H[x] + \ln |A|$ 를 증명하면, 아래와 같다.

$$H(y) = - \int p(y) \ln p(y) dy \quad (y\text{에 대한 entropy})$$

$p(x) = p(y) \left| \frac{dy_i}{dx_j} \right| = p(y) |A|$ 에 의해, 위 식을 변형하면 아래와 같다.

$$\begin{aligned} (\text{준 식}) &= - \int p(y) \ln p(y) dy = - \int p(x) \ln \{p(x) |A|^{-1}\} dx \\ &= - \int p(x) \{ \ln p(x) + \ln |A|^{-1} \} dx \\ &= - \int p(x) \ln p(x) + p(x) \ln |A|^{-1} dx \\ &= - \int p(x) \ln p(x) dx - \int p(x) \ln |A|^{-1} dx \\ &= - \int p(x) \ln p(x) dx - \ln |A|^{-1} \int p(x) dx \\ &= - \int p(x) \ln p(x) dx + \ln |A| \\ &= H[x] + \ln |A| \end{aligned}$$

따라서 $H[y] = H[x] + \ln |A|$ 는 참이다.

(증명 끝)

1.35. Use the results (1.106) and (1.107) to show that the entropy of the univariate Gaussian (1.109) is given by (1.110).

$$\int_{-\infty}^{\infty} xp(x)dx = \mu \quad (1.106)$$

$$\int_{-\infty}^{\infty} (x - \mu)^2 p(x) dx = \sigma^2 \quad (1.107)$$

$$p(x) = \frac{1}{(2\pi\sigma^2)^{\frac{1}{2}}} \exp\left\{-\frac{(x - \mu)^2}{2\sigma^2}\right\} \quad (1.109)$$

$$H(x) = \frac{1}{2}\{1 + \ln(2\pi\sigma^2)\} \quad (1.110)$$

Proof)

$$H(x) = - \int p(x) \ln p(x) dx$$

$$\begin{aligned} &= - \int p(x) \times \ln \left[\frac{1}{(2\pi\sigma^2)^{\frac{1}{2}}} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\} \right] dx \\ &= - \int p(x) \times \left[\ln \left\{ \frac{1}{(2\pi\sigma^2)^{\frac{1}{2}}} \right\} - \frac{(x - \mu)^2}{2\sigma^2} \right] dx \\ &= - \int p(x) \times \left[-\ln(2\pi\sigma^2)^{\frac{1}{2}} - \frac{(x - \mu)^2}{2\sigma^2} \right] dx \\ &= \int p(x) \times \left[\ln(2\pi\sigma^2)^{\frac{1}{2}} + \frac{(x - \mu)^2}{2\sigma^2} \right] dx \\ &= \int p(x) \times \left[\frac{1}{2} \ln(2\pi\sigma^2) + \frac{(x - \mu)^2}{2\sigma^2} \right] dx \\ &= \int p(x) \times \frac{1}{2} \ln(2\pi\sigma^2) dx + \int p(x) \times \frac{(x - \mu)^2}{2\sigma^2} dx \\ &= \frac{1}{2} \int p(x) \times \ln(2\pi\sigma^2) dx + \int p(x) \times \frac{(x - \mu)^2}{2\sigma^2} dx \\ &= \frac{1}{2} \int p(x) \times \ln(2\pi\sigma^2) dx + \frac{1}{2\sigma^2} \int p(x)(x - \mu)^2 dx \end{aligned}$$

(1.107 식에 의해 $\int p(x)(x - \mu)^2 dx = \sigma^2$ 이므로,)

$$\begin{aligned} &= \frac{1}{2} \int p(x) \times \ln(2\pi\sigma^2) dx + \frac{1}{2\sigma^2} \times \sigma^2 \\ &= \frac{1}{2} \int p(x) \times \ln(2\pi\sigma^2) dx + \frac{1}{2} \\ &= \frac{1}{2} \ln(2\pi\sigma^2) \int p(x) dx + \frac{1}{2} \end{aligned}$$

(확률공리에 의해, $\int p(x) dx = 1$ 이므로,)

$$= \frac{1}{2} \ln(2\pi\sigma^2) + \frac{1}{2}$$

$$= \frac{1}{2} \{ \ln(2\pi\sigma^2) + 1 \}$$

$$\therefore H(x) = \frac{1}{2} \{ 1 + \ln(2\pi\sigma^2) \}$$

(증명 끝)

1.37. Using the definition (1.111) together with the product rule of probability, prove the result (1.112).

$$H[y | x] = - \iint p(y, x) \ln p(y | x) dy dx \quad (1.111)$$

$$H[x, y] = H[y | x] + H[x] \quad (1.112)$$

Proof)

$$H[x, y] = - \iint p(x, y) \ln p(x, y) dx dy$$

Continuous variable의 product rule of probability, $p(x, y) = p(y | x)p(x) = p(x | y)p(y) = p(y, x)$ 이용하면, 아래와 같이 식을 전개할 수 있다.

$$\begin{aligned} (\text{준 식}) &= - \iint p(x, y) \ln p(y | x)p(x) dx dy \\ &= - \iint p(x, y) \ln p(y | x) + p(x, y) \ln p(x) dx dy \\ &= - \iint p(x, y) \ln p(y | x) dx dy - \iint p(x, y) \ln p(x) dx dy \\ &= - \iint p(y, x) \ln p(y | x) dx dy - \iint p(x, y) \ln p(x) dx dy \\ &= H[y | x] - \iint p(x, y) \ln p(x) dx dy \\ &= H[y | x] - \int \ln p(x) \left\{ \int p(x, y) dy \right\} dx \\ &= H[y | x] - \int p(x) \ln p(x) dx = H[y | x] + H[x] \\ \therefore H[x, y] &= H[y | x] + H[x] \end{aligned}$$

(증명 끝)

1.40. By applying Jensen's inequality (1.115) with $f(x) = \ln x$, show that the arithmetic mean of a set of real numbers is never less than their geometrical mean.

$$f\left(\sum_{i=1}^M \lambda_i x_i\right) \leq \sum_{i=1}^M \lambda_i f(x_i) \quad (1.115)$$

$$\frac{1}{M} \sum_{i=1}^M x_i \quad (\text{arithmetic mean})$$

$$\sqrt[M]{\prod_{i=1}^M x_i} \quad (\text{geometric mean})$$

Proof)

명제, "the arithmetic mean of a set of real numbers is never less than their geometrical mean."는 아래와 같은 식으로 나타낼 수 있다.

$$\frac{1}{M} \sum_{i=1}^M x_i \geq \sqrt[M]{\prod_{i=1}^M x_i}$$

위 부등식의 양변에 자연로그(ln)을 취하면 아래와 같다.

$$\ln\left(\frac{1}{M} \sum_{i=1}^M x_i\right) \geq \ln\left(\sqrt[M]{\prod_{i=1}^M x_i}\right)$$

그리고 위 부등식을 변형하면, 아래와 같이 변형할 수 있다.

$$\ln\left(\frac{1}{M} \sum_{i=1}^M x_i\right) \geq \ln\left(\left(\prod_{i=1}^M x_i\right)^{\frac{1}{M}}\right)$$

$$\ln\left(\frac{1}{M} \sum_{i=1}^M x_i\right) \geq \frac{1}{M} \times \ln\left(\prod_{i=1}^M x_i\right)$$

$$\ln\left(\frac{1}{M} \sum_{i=1}^M x_i\right) \geq \frac{1}{M} \times \ln(x_1 x_2 x_3 \cdots x_m)$$

$$\ln\left(\frac{1}{M} \sum_{i=1}^M x_i\right) \geq \frac{1}{M} \times (\ln x_1 + \ln x_2 + \ln x_3 + \cdots + \ln x_M)$$

$$\ln\left(\frac{1}{M} \sum_{i=1}^M x_i\right) \geq \frac{1}{M} \sum_{i=1}^M (\ln x_i)$$

λ_i 의 정의에 의해서 $\sum_i \lambda_i = 1$ 이므로, 위 식을 아래와 같이 변형할 수 있다.

$$\ln \left(\frac{1}{M} \sum_{i=1}^M \lambda_i x_i \right) \geq \frac{1}{M} \sum_{i=1}^M \lambda_i (\ln x_i)$$

$$\ln \left(\sum_{i=1}^M \frac{\lambda_i x_i}{M} \right) \geq \sum_{i=1}^M \frac{\lambda_i (\ln x_i)}{M}$$

$f(x) = \ln x$ 가 concave function이므로, Jensen's inequality에 의해 위 부등식은 참이다.

(증명 끝)

1.41. Using the sum and product rules of probability, show that the mutual information $I(x, y)$ satisfies the relation (1.121).

$$I[x, y] = H[x] - H[x | y] = H[y] - H[y | x] \quad (1.121)$$

Proof)

Mutual information, $I[x, y]$ 는 아래와 같이 정의된다.

$$I[x, y] = - \iint p(x, y) \ln \left(\frac{p(x)p(y)}{p(x, y)} \right) dx dy \quad \dots \quad \text{식 ①}$$

Continuous variable의 product rule of probability, $p(x, y) = p(y | x)p(x) = p(x | y)p(y) = p(y, x)$ 이용하면, 아래와 같이 식을 전개할 수 있다.

$$\begin{aligned} (\text{준 식}) &= - \iint p(x, y) \ln \left(\frac{p(x)p(y)}{p(y | x)p(x)} \right) dx dy \\ &= - \iint p(x, y) \ln \left(\frac{p(y)}{p(y | x)} \right) dx dy \\ &= - \iint p(x, y) \ln p(y) - p(x, y) \ln p(y | x) dx dy \\ &= - \iint p(x, y) \ln p(y) dx dy + \iint p(x, y) \ln p(y | x) dx dy \\ &= - \int \ln p(y) \left\{ \int p(x, y) dx \right\} dy + \iint p(x, y) \ln p(y | x) dx dy \\ &= - \int \ln p(y) \left\{ \int p(y, x) dx \right\} dy + \iint p(x, y) \ln p(y | x) dx dy \end{aligned}$$

Continuous variable의 sum rule of probability, $\int p(y, x) dx = p(y)$ 에 의해 위 식을 변형하면, 아래와 같다.

$$\begin{aligned}
(\text{준 식}) &= - \int p(y) \ln p(y) dy + \iint p(x, y) \ln p(y | x) dx dy \\
&= H[y] + \iint p(x, y) \ln p(y | x) dx dy
\end{aligned}$$

Continuous variable의 product rule of probability, $p(x, y) = p(y | x)p(x) = p(x | y)p(y) = p(y, x)$ 이용하면, 아래와 같이 식을 전개할 수 있다.

$$\begin{aligned}
(\text{준 식}) &= H[y] - \left(- \iint p(y, x) \ln p(y | x) dy dx \right) \\
&= H[y] - H[y | x] \\
\therefore \mathbf{I[x, y]} &= \mathbf{H[y]} - \mathbf{H[y | x]}
\end{aligned}$$

식 ①을 다시 이용하여, $\mathbf{I[x, y]} = \mathbf{H[x]} - \mathbf{H[x | y]}$ 를 증명하면, 아래와 같다.

$$(\text{식 ①}) = - \iint p(x, y) \ln \left(\frac{p(x)p(y)}{p(x, y)} \right) dx dy$$

Continuous variable의 product rule of probability, $p(x, y) = p(y | x)p(x) = p(x | y)p(y) = p(y, x)$ 로부터, $p(x, y) = p(y, x)$ 라는 결과를 이끌어 낼 수 있으므로, 아래와 같이 식을 변형할 수 있다.

$$\begin{aligned}
(\text{준 식}) &= - \iint p(y, x) \ln \left(\frac{p(x)p(y)}{p(y, x)} \right) dx dy \\
&= - \iint p(y, x) \ln \left(\frac{p(x)p(y)}{p(x|y)p(y)} \right) dx dy \\
&= - \iint p(y, x) \ln \left(\frac{p(x)}{p(x | y)} \right) dx dy \\
&= - \iint p(y, x) \ln p(x) - p(y, x)p(x | y) dx dy \\
&= - \iint p(y, x) \ln p(x) dx dy + \iint p(y, x)p(x | y) dx dy \\
&= - \iint p(y, x) \ln p(x) dx dy - \left(- \iint p(y, x)p(x | y) dx dy \right) \\
&= - \iint p(y, x) \ln p(x) dx dy - H[x | y] \\
&= - \int \ln p(x) \left\{ \int p(y, x) dy \right\} dx - H[x | y] \\
&= - \int \ln p(x) \left\{ \int p(x, y) dy \right\} dx - H[x | y]
\end{aligned}$$

Continuous variable의 sum rule of probability, $\int p(x, y) dy = p(x)$ 에 의해 위 식을 변형하면, 아래와 같다.

$$(\text{준 식}) = - \int p(x) \ln p(x) dx - H[x | y]$$

$$= H[x] - H[x | y]$$

$$\therefore I[x, y] = H[x] - H[x | y]$$

결과적으로, $I[x, y] = H[x] - H[x | y]$ 이고, $I[x, y] = H[y] - H[y | x]$ 이므로, 명제 $I[x, y] = H[x] - H[x | y] = H[y] - H[y | x]$ 는 성립한다.

$$\therefore I[x, y] = H[x] - H[x | y] = H[y] - H[y | x]$$

(증명 끝)