# CSE 546 Homework 2B

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## 1 Convexity and Norms

B1.

(a) To show that f(x) = ||x|| is a convex function we will use the definition of convexity. Let  $\lambda \in [0, 1]$  and apply the triangle inequality and absolute scalability for any  $x, y \in \mathbb{R}^n$ :

$$\|\lambda x + (1 - \lambda)y\| \le \|\lambda x\| + \|(1 - \lambda)y\| \le \lambda \|x\| + (1 - \lambda)\|y\|.$$

Therefore we have shown f(x) = ||x|| is a convex function.

(b) We want to show that the set  $D := \{x \in \mathbb{R}^n : ||x|| \le 1\}$  is a convex set. Let  $\lambda \in [0,1]$  then from the previous problem in B1-a we have that

$$\|\lambda x\| + \|(1 - \lambda)y\| \le \lambda \|x\| + (1 - \lambda)\|y\|.$$

From the defined set we say that  $||x|| \le 1$  and  $||y|| \le 1$  for any  $x, y \in D$ . Then we can further write

$$\begin{split} \|\lambda x\| + \|(1-\lambda)y\| &\leq \lambda \|x\| + (1-\lambda)\|y\| \leq \lambda + 1 - \lambda = 1. \\ &\implies \lambda x + (1-\lambda)y \in D. \end{split}$$

Thus we have shown that any line segment between the two points lies in D hence it is a convex set.

(c) The set  $D := \{(x,y) : g(x,y) \le 4\}$  where  $g(x,y) = (|x|^{1/2} + |y|^{1/2})^2$  is below. This set is not convex, it is clear to see we can take a point at one of of the "spikey points" of the set and try to draw a line to the other point and it would leave the set, thus breaking the definition of convexity.

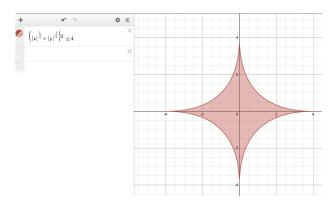


Figure 1: The set D defined by g(x,y) in the problem statement.

## 2 Bounding the Estimate

**B2**.

(a) By definition we can write the least squares estimator as follows:

$$\widehat{\beta} = (X^T X)^{-1} X^T Y$$

$$= (X^T X)^{-1} X^T (X \beta^* + \epsilon)$$

$$= \beta^* + (X^T X)^{-1} X^T \epsilon$$

$$\Longrightarrow \widehat{\beta} - \beta^* = (X^T X)^{-1} X^T \epsilon$$

$$\widehat{\beta} = \beta^* + (X^T X)^{-1} X^T \epsilon$$

Since  $\widehat{\beta}$  is normally distributed and  $\epsilon \sim \mathcal{N}(0, I)$  and assuming X is fixed:

$$\begin{split} \widehat{\beta} &\sim \mathcal{N}(0*(X^TX)^{-1}X^T\epsilon + \beta^*, (X^TX)^{-1}X^TI((X^TX)^{-1}X^T)^T) \\ &= \mathcal{N}(\beta^*, (X^TX)^{-1}X^T((X^TX)^{-1}X^T)^T) \\ &= \mathcal{N}(\beta^*, (X^TX)^{-1}X^TX(X^{-1}X^T)^T) \\ &= \mathcal{N}(\beta^*, X^{-1}X^{-T}) \\ &= \mathcal{N}(\beta^*, (X^TX)^{-1}) \end{split}$$

where  $(X^TX)^{-1}$  is the covariance matrix of  $\widehat{\beta}$ . For each of the elements  $\widehat{\beta}_j$  the variance will be the j-th diagonal element of the covariance matrix. And, since  $\widehat{\beta}_j$  is normally distributed, all of its elements will be as well> Thus we write:

$$\widehat{\beta}_j = \mathcal{N}(B_j^*(X^T X)_{i,j}^{-1})$$

(b) Let's fix  $\delta \in (0,1)$  and let  $j \in [d]$ . From the previous problem we can derive

$$\widehat{\beta}_{j} = \mathcal{N}(B_{j}^{*}(X^{T}X)_{i,j}^{-1})$$

$$\implies \frac{\widehat{\beta}_{j}}{(X^{T}X)_{i,j}^{-1}} \sim \mathcal{N}(0,1)$$

Then by proposition 2, we have that with a probability of at least  $1 - \delta$ ,  $\widehat{\beta}_j$  will be contained within a confidence interval:

$$\begin{bmatrix} \beta^* - \sqrt{2(X^T X)_{j,j}^{-1} \log(2/\delta)}, & \beta^* + \sqrt{2(X^T X)_{j,j}^{-1} \log(2/\delta)} \end{bmatrix}$$
$$= \begin{bmatrix} -\sqrt{2(X^T X)_{j,j}^{-1} \log(2/\delta)}, \sqrt{2(X^T X)_{j,j}^{-1} \log(2/\delta)} \end{bmatrix}$$

Therefore,

$$\mathbb{P}\left(-\sqrt{2(X^TX)_{j,j}^{-1}\log(2/\delta)} \le \widehat{\beta}_j \le \sqrt{2(X^TX)_{j,j}^{-1}\log(2/\delta)}\right) \ge 1 - \delta$$

$$\implies \mathcal{P}\left(\left|\widehat{\beta}_j\right| \le \sqrt{2(X^TX)_{j,j}^{-1}\log(2/\delta)}\right) \ge 1 - \delta.$$

We <u>cannot</u> conclude that with probability at least  $1 - \delta$ ,  $\left| \widehat{\beta}_j \right| \leq \sqrt{2(X^T X)_{j,j}^{-1} \log(2/\delta)}$  for all  $j \in [d]$  simultaneously because each of the events have a probability of at least  $1 - \delta$  but this does not guarantee that all of the events with simultaneously be at least probability of  $1 - \delta$ .

(c) If we run the code below we see that we have 63 out of 10,000  $\hat{\beta}'_{j}s$  outside of the interval.

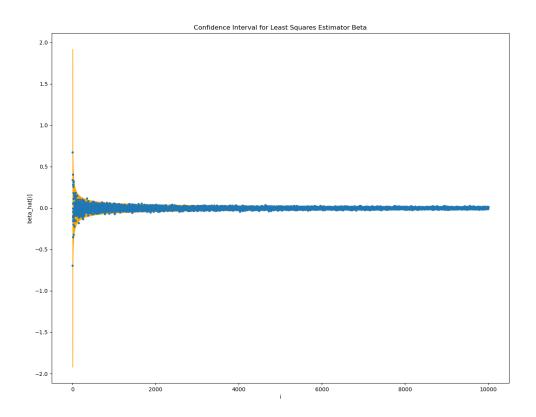


Figure 2: Confidence Interval

```
import numpy as np
from matplotlib import pyplot as plt
import random
random.seed(10)
delta = 0.05
n = 20000
d = 10000
beta = np.zeros(d)
```

```
X = np.empty((n,d))
inv_XTX = np.zeros((d,d))
y = np.random.normal(0, 1, n)
for i in range(n):
    index = (i\%d)+1
    e = np.zeros(d)
    e[index-1] = 1.0
    X[i,:] = np.sqrt(index)*e
XTX = X.T @ X
ci = np.zeros(d)
for i in range(d):
    inv_XTX[i,i] = 1/(XTX[i,i])
    ci[i] =np.sqrt(2*inv_XTX[i,i]*np.log(2/delta))
beta_hat = inv_XTX @ X.T @ y
count = 0
for i in range(d):
    if np.abs(beta_hat[i])> ci[i]:
        count += 1
x = np.linspace(1,d,d)
plt.figure(figsize=(20,20))
plt.plot(x,beta_hat, '.')
plt.fill_between(x, (beta-ci), (beta+ci), color='orange')
plt.xlabel("i")
plt.ylabel("beta_hat[i]")
plt.title("Confidence Interval for Least Squares Estimator Beta")
plt.show()
print(count)
```