

## Joint and conditional Distributions

Definition :  $F : \mathbb{R}^2 \rightarrow [0, 1]$  is a joint distribution function of  $X$  &  $Y$  if

$$F(x, y) = P(X \leq x, Y \leq y)$$

If  $X, Y$  are discrete, the joint mass function  $p : \mathbb{R}^2 \rightarrow [0, 1]$  is given by

$$p(x, y) = P(X = x, Y = y),$$

$$A_x = \{ \omega : X(\omega) = x \}$$

$$B_y = \{ \omega : Y(\omega) = y \}$$

$$p(x, y) = P(A_x \cap B_y)$$

$$P(X = x) = P\left(\bigcup_y (\{X = x\} \cap \{Y = y\})\right)$$

$$= \sum_y P(X = x, Y = y)$$

$$= \sum_y p(x, y)$$

The marginal mass function of  $X$  :

$$p_X(x) = \sum_y p(x, y)$$

Similarly,

$$p_Y(y) = \sum_x p(x, y)$$

Definition :  $E[XY] = \sum_x \sum_y xy p(x, y)$

Theorem :  $E[g(X, Y)] = \sum_x \sum_y g(x, y) p(x, y)$ .

Theorem : If  $X, Y$  are independent,  
then  $E(XY) = E(X) \cdot E(Y)$ .

Proof :  $A_x = \{X = x\}$   
 $B_y = \{Y = y\}$

$$XY = \sum_x \sum_y xy I_{A_x \cap B_y}$$

$$E(XY) = \sum_x \sum_y xy E(I_{A_x \cap B_y})$$

$$= \sum_x \sum_y xy P(A_x \cap B_y)$$

$$= \sum_x \sum_y xy P(A_x) \cdot P(B_y)$$

$$= \sum_x \sum_y xy \cdot p(x, y)$$

□

Remark : ① If  $g, h$  are functions &  
 $X, Y$  are independent, then

$$E[g(X)h(Y)] = \sum_x \sum_y g(x)h(y) p(x, y)$$

② If  $E(XY) = E(X) \cdot E(Y)$ , then  $X$  &  $Y$   
are called uncorrelated.

Ex: Let  $X \sim \text{Geo}(\alpha)$ ,  $Y \sim \text{Geo}(\beta)$  be independent. Find the pmf of  $Z = \min\{X, Y\}$ .

Solution:

$$\begin{aligned} P(Z > k) &= P(X > k, Y > k) \\ &= P(X > k) P(Y > k) \\ &= (1-\alpha)^k \cdot (1-\beta)^k = ((1-\alpha)(1-\beta))^k \end{aligned}$$

Let  $k \in \mathbb{N}$

$$\begin{aligned} P(Z = k) &= P(Z > k-1) - P(Z > k) \\ &= ((1-\alpha)(1-\beta))^{k-1} - ((1-\alpha)(1-\beta))^k \\ &= ((1-\alpha)(1-\beta))^{k-1} (1 - (1-\alpha)(1-\beta)) \end{aligned}$$

$$\Rightarrow Z \sim \text{Geo}((1-\alpha)(1-\beta))$$

Definition: The covariance of  $X$  &  $Y$  is

$$\text{cov}(X, Y) = E[(X - E(X))(Y - E(Y))] = E(XY) - E(X)E(Y)$$

Theorem:  $V(X+Y) = V(X) + V(Y) + 2\text{cov}(X, Y)$

$$E[(X+Y)^2] - (E(X+Y))^2$$

$$\begin{aligned} &= E[X^2 + 2XY + Y^2] - (E(X))^2 - 2E(X)E(Y) - (E(Y))^2 \\ &= E(X^2) - (E(X))^2 + E(Y^2) - (E(Y))^2 + 2[E(XY) - E(X)E(Y)] \\ &= V(X) + V(Y) + 2\text{cov}(X, Y) \end{aligned}$$

Theorem: If  $X$  &  $Y$  are independent, then  $\text{COV}(X, Y) = 0$ .

Remark: If  $X$  &  $Y$  are independent, then  $V(X+Y) = V(X) + V(Y)$ .

Theorem: Cauchy - Schwartz inequality.

If  $E(X^2), E(Y^2) < \infty$ , then

$$(E(XY))^2 \leq E(X^2) E(Y^2)$$

Theorem:  $E(aX + bY)^2 \geq 0$ , for any  $a, b \in \mathbb{R}$ .

$0 \leq E(aX + bY)^2 = a^2 E(X^2) + 2ab E(XY) + b^2 E(Y^2)$   
is a quadratic function of  $a$ .

Since  $E[(aX + bY)^2] \geq 0$ , this can have at most one real solution. Therefore,

$$\begin{aligned} \text{Discriminant} &= (2b E(XY))^2 - 4(E(X^2) b^2 E(Y^2)) \leq 0 \\ \Rightarrow E(XY)^2 &\leq E(X^2) E(Y^2) \quad \square \end{aligned}$$

## Sums of random variables

Theorem : If  $X$  &  $Y$  are independent discrete rvs, then

$$P(X+Y=z) = \sum_x P_{X,Y}(x, z-x)$$

$$\text{Proof: } \{X+Y=z\} = \bigcup \left\{ \{X=x\} \cap \{Y=z-x\} \right\}$$

$$\Rightarrow P(X+Y=z) = \sum_x P(X=x, Y=z-x)$$

Remark : If  $X$  &  $Y$  are independent, then

$$P(X+Y=z) = P_{X+Y}(z) = \sum_x P_X(x) P_Y(z-x)$$

$$\text{or} = \sum_y P_X(z-y) P_Y(y)$$

This sum is called the convolution of  $X$  &  $Y$

$$\text{Notation : } P_{X+Y} = P_X * P_Y$$

Ex: Sums of independent binomial.

$$X_1 \sim B(n_1, p), \quad X_2 \sim B(n_2, p)$$

$$P(X_1 + X_2 = k) = \sum_{m=0}^k P(X_1 = m) \cdot P(X_2 = k - m)$$

$$= \sum_{m=0}^k \binom{n_1}{m} p^m q^{n_1-m} \cdot \binom{n_2}{k-m} p^{k-m} q^{n_2-(k-m)}$$

$$= p^k q^{(n_1+n_2)-k} \cdot \underbrace{\sum_{m=0}^k \binom{n_1}{m} \binom{n_2}{k-m}}_{\binom{n_1+n_2}{k}} = \binom{n_1+n_2}{k} p^k q^{(n_1+n_2)-k}$$

$\Rightarrow X_1 + X_2 \sim B(n_1 + n_2, p)$