# Scalar on Function Regression

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Presentation Day

#### Introduction

- Near-infrared (NIR) spectroscopy enables fast diagnostics by using the NIR region of the electromagnetic spectrum
- Suited for field-monitoring / on-line analysis and diagnostics of e.g: prediction of octane ratings!
- Spectroscopy results in high-dimensional dataset.
- This set of measurements along a continuum can be viewed as set of smooth spectral curves
- Regression to determine relationship between octane rating and spectral curves



# Theory

A simple functional dataset is given by

$$\{x_i(t_{j,i}) \in \mathbb{R} \mid i = 1, 2, ..., N, j = 1, 2, ..., J_i, t_{j,i} \in [T_1, T_2]\}$$

- Continuous underlying process, where  $x_i(t)$  exists  $\forall t \in [T_1, T_2]$
- Only observed at  $x_i(t_{j,i})$
- Growth curves, financial data, human perception (pitch), ...
- To abstract information from the curves, they must be interpretable!



# Theory

#### Jonghun

- Random Functions (name square integrable functions)
- Motivate continuous stochastic processes (growth curves/electricity consumption/yield curves/stonks)
- Use curves to predict a scalar response (show typical dgp)

# Theory

### Jonghun

- Basis expansions (b-splines and fourier)
- Talk about purposes
- Plots and show bias variance tradeoff

#### Random Function

Random element is a function  $X:\Omega\to\mathcal{S}$  which is defined on a common probability space  $(\Omega,\mathcal{F},\mathbb{P})$  where  $\Omega$  is a probability space with a  $\sigma$ -algebra  $\mathcal{F}$  and a probability measure  $\mathbb{P}$ 

- lacksquare If  $\mathcal{S}=\mathbb{R} o \mathsf{X}$  is a random variable
- If  $\mathcal{S} = \mathbb{R}^n o \mathsf{X}$  is a random vector
- If S is a space of functions, X is called a random function

#### Random Function

Let E be the index set and this can be described as

$${X(t,\omega): t \in E, \omega \in \Omega},$$

where  $X(t,\cdot)$  is  ${\mathcal F}$ -measurable function on the sample space  $\Omega$ 

- lacksquare It can be shortened to X(t) by omitting  $\omega$
- The function is realized when the X(t) have been observed for every  $t \in E$

# Square Integrable Function

If a function *f* satisfies:

$$\int_0^1 \{f(t)\}^2 dt < \infty$$

the function f is called square integrable function

- Without loss of generality, the interval is defined in [0,1]
- lacksquare L<sup>2</sup> is the set of all square integrable functions
- We focus on  $\mathbb{L}^2[0,1]$  since the dimension of the domain of our function is 1

# Stochastic Process Perspective

Functional data are the sample curves observed from continuous time stochastic precess

- From this perspective, X(t) is a random variable (?)
- **\mathbf{X}(\cdot)** is a collection of random variables by each time index(?)
- Realizations of a random function belong in large collection of functions

### **Plots**

Two plots of stock and human growth



### **Plots**

point plots to link basis expansion

# Basis Expansion

Basis expansion is a linear combination of functions defining a function as described:

$$X_i(t) pprox \sum_{k=1}^K c_{ik} \phi_k(t), \quad 1 \leq i \leq n, \forall t \in E$$

, where  $\phi_k(t)$  is the  $k^{th}$  basis function of the expansion and  $c_{ik}$  is the corresponding coefficient

- To make the function smoother
- To replace the original scalar data  $X_n(t_{jn})$  by a smaller collection of  $c_{nm}$



# Two Typical Types of Basis Function

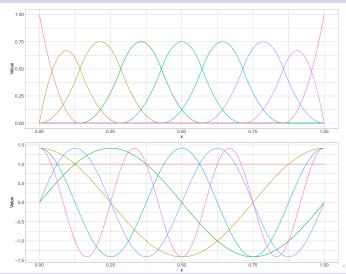
Fourier basis function is written as

$$f(x) = a_0 + \sum_{n=1}^{\infty} a_n cos(2\pi nx) + b_n sin(2\pi nx)$$

B-spline basis function is a flexible curve defined by degree and knots Each B-spline basis function, i-th B-spline basis function of degree p,  $N_{i,p}(u)$  is defined on Cox-de Boor recursion formula(Do you think I need to put the formula?)



### Plots of Basis Functions



#### Trade Off between Bias and Variance

How do we choose the number K of basis functions?

- The larger K, the better fit to the data with also fitting noise
- If K is too small, it would miss some significant information that we want to estimate

### Trade Off between Bias and Variance

■ Bias[
$$\hat{x}(t)$$
] =  $x(t) - E[\hat{x}(t)]$ 

• 
$$Var[\hat{x}(t)] = E[\{\hat{x}(t) - E[x(t)]\}^2]$$

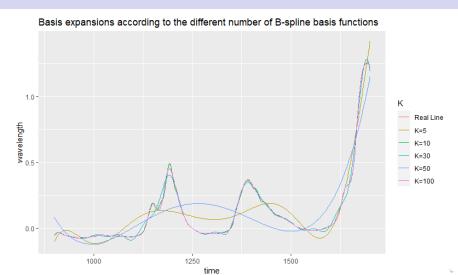
■ 
$$MSE[\hat{x}(t)] = E[\{\hat{x}(t) - x(t)\}]$$

■ 
$$MSE[\hat{x}(t)] = Bias^2[\hat{x}(t)] + Var[\hat{x}(t)]$$

In the sense of that, we need to concentrate on decreasing **MSE**.



### Trade Off between Bias and Variance



Assume the following data generating process

$$Y(\omega) = \alpha + \int_0^1 \beta(s) F(\omega)(s) ds + \epsilon(\omega)$$

lacksquare  $Y(\omega)$  and  $\epsilon(\omega)$  realize in  $\mathbb R$  and  $F(\omega)$  realizes in  $\mathbb L^2[0,1]$ 

Assume that we have a data set containing observations each of which is made up of:

- $y_i$ : a scalar realization of  $Y(\omega)$
- $f_i(t)$ : a realization of  $F(\omega)$

Let  $\{b_i(t) \mid i=1,\ldots,\infty\}$  be a basis of  $\mathbb{L}^2[0,1]$ 

Then we have the following representation of  $\beta(t)$ 

$$eta(t) = \sum_{j=1}^{\infty} \psi_j b_j(t) = \sum_{j=1}^{L} \psi_j b_j(t) + \delta(t) pprox \sum_{j=1}^{L} \psi_j b_j(t)$$

and we can transform the data generating process into:

$$Y(\omega) = \alpha + \int_0^1 \left[ \left( \sum_{j=1}^\infty \psi_j b_j(s) \right) F(\omega)(s) \right] ds + \epsilon(\omega)$$
$$= \alpha + \sum_{j=1}^\infty \left[ \psi_j \int_0^1 F(\omega)(s) b_j(s) ds \right] + \epsilon(\omega)$$

$$Z_j(\omega) = \int_0^1 F(\omega)(s)b_j(s)\mathrm{d}s$$

This is a **scalar random variable** leading to the following transformation

$$Y(\omega) = \alpha + \sum_{j=1}^{\infty} \psi_j Z_j(\omega) + \epsilon(\omega)$$

Each combination of observation  $f_i(t)$  and deterministic basis function  $b_j(t)$  effectively gives us a realization of this random variable.

$$Z_{i,j} = \int_0^1 f_i(s)b_j(s)\mathrm{d}s$$



This allows us to write each observation in the data set as

- $y_i$ : a scalar realization of Y
- $(Z_{i,j})_{j \in \mathbb{N}}$ : a countably infinite sequence of scalars

Truncating the functional basis allows us to approximate the data set in the usual multivariate form.

- y<sub>i</sub>: a scalar realization of Y
- $(Z_{i,1} \ldots Z_{i,L})'$ : a vector of scalar regressors

Coefficients can then be estimated using theory from **multivariate** regression leading to an estimated coefficient vector  $\hat{\psi}_L \in \mathbb{R}^L$ .



This can be translated into an estimated coefficient function  $\hat{\beta}(t)$ :

$$\hat{\beta}_L(t) = \sum_{j=1}^L \hat{\psi}_{L,j} b_j(t)$$

This is dependent on...

- The functional basis  $(b_j(t))_{j\in\mathcal{I}}$  for the estimation of  $\beta(t)$
- The truncation parameter *L*
- (The functional basis used to approximate the observations)
- (The truncation parameter in the approximation of the observations)



# Spectral Representation of Random Vectors

Let  $X(\omega)$  be a random vector realizing in  $\mathbb{R}^p$ .

- Let  $\mu_X = \mathbb{E}(X)$  and  $\Sigma_X = Cov(X)$
- Let  $\{\gamma_i \mid i=1,\ldots,p\}$  be the orthonormal **Eigenvectors** of  $\Sigma_X$
- Let  $\{\lambda_i \mid i=1,\ldots,p\}$  be the corresponding **Eigenvalues** of  $\Sigma_X$

Then X can also be represented as

$$X(\omega) = \mu_{\mathsf{x}} + \sum_{i=1}^{p} \xi_{i}(\omega) \gamma_{i}$$

where the  $\xi_i(\omega)$  have the following properties

$$\mathbb{E}[\xi_i(\omega)] = 0$$

2 
$$Var(\xi_i(\omega)) = \lambda_i$$

3 
$$Cov(\xi_i(\omega), \xi_j(\omega)) = 0$$
 for  $i \neq i$ 



# Karhunen-Loéve Expansion

Mean Function:

$$\mu(t) = \mathbb{E}\left[F(\omega)(t)\right]$$

**Autocovariance Function:** 

$$c(t,s) = \mathbb{E}\big[\left(F(\omega)(t) - \mu(t)\right)\left(F(\omega)(s) - \mu(s)\right)\big]$$

The **Eigenvalues** and **Eigenfunctions**:  $\{(\lambda_i, \nu_i) \mid i \in \mathcal{I}\}$  are solutions of the following equation:

$$\int_0^1 c(t,s)\nu(s)\mathrm{d}s = \lambda\nu(t)$$



# Karhunen-Loéve Expansion

A random function  $F(\omega)$  can be expressed in terms of its mean function and its Eigenfunctions:

$$F(\omega)(t) = \mu(t) + \sum_{j=1}^{\infty} \xi_j(\omega)\nu_j(t)$$

Where the  $\xi_j$  are scalar-valued random variables with the following properties.

$$\mathbb{E}[\xi_i(\omega)] = 0$$

$$2 Var(\xi_i(\omega)) = \lambda_i$$

3 
$$Cov(\xi_i(\omega), \xi_j(\omega)) = 0$$
 for  $i \neq j$ 

This representation is called the **Karhunen-Loéve Expansion** of the random function F and the Eigenfunctions can serve as a basis to represent the function.

# Principal Component Analysis

A related concept is **Principal Component Analysis** (PCA).

### $\Sigma_X$ unknown o sample analogues

- Let  $\mathbf{X} \in \mathbb{R}^{n \times p}$  contain the standardized regressors
- Let  $\hat{\Sigma}_X = \frac{\mathbf{X}'\mathbf{X}}{n}$
- Let  $\{\hat{\gamma}_i \mid i=1,\ldots,p\}$  be the orthonormal **Eigenvectors** of  $\hat{\Sigma}_X$
- Let  $\{\hat{\lambda}_i \,|\, i=1,\ldots,p\}$  be the corresponding **Eigenvalues** of  $\hat{\Sigma}_X$

Then  $Z_i(\omega) = \hat{\gamma}_i' X(\omega)$  is called the i'th principal component and

3 
$$Cov(Z_i(\omega), Z_j(\omega)) = 0$$
 for  $i \neq j$ 



# Functional Principal Component Analysis

This idea can be extended to functional regressors in the form of **Functional Principal Component Analysis** (FPCA).

**Empirical Mean Function:** 

$$\hat{\mu}(t) = \frac{1}{n} \sum_{j=1}^{n} f_j(t)$$

**Empirical Autocovariance Function:** 

$$\hat{c}(t,s) = \frac{1}{n} \sum_{j=1}^{n} (f_j(t) - \hat{\mu}(t)) (f_j(s) - \hat{\mu}(s))$$

# Functional Principal Component Analysis

The **Eigenvalues** and **Eigenfunctions**:  $\{(\hat{\lambda}_i, \hat{\nu}_i) \mid i \in \mathcal{I}\}$  are solutions of the following equation:

$$\int_0^1 \hat{c}(t,s)\hat{\nu}(s)\mathrm{d}s = \hat{\lambda}\hat{\nu}(t)$$

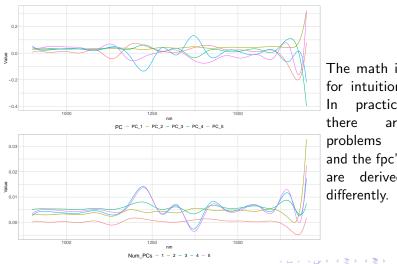
The  $\{\hat{\nu}_i(s) \mid i \in \mathcal{I}\}$  are called **Functional Principal Components** and can serve as a basis for representing the original curves.

The corresponding scores  $\hat{\xi}_i$  can be derived as

$$\hat{\xi}_j(\omega) = \int_0^1 (F(\omega)(s) - \hat{\mu}(s)) \hat{\nu}_j(s) \mathrm{d}s$$



### FPCA - Plots



The math is for intuition. practice there are problems and the fpc's are derived differently.

# Simulation Setup & Application

- Use the **gasoline dataset** (NIR-spectroscopy,  $60 \times 401$ ) to predict octane ratings.
- Generate similar curves from gasoline dataset:

$$ilde{\mathcal{F}}(\omega)(t) = \hat{\mu}(t) + \sum_{j=1}^{\infty} ilde{\xi_j}(\omega)\hat{
u}_j(t)$$

- lacksquare  $ilde{\xi}_j \sim \mathcal{N}(0,\hat{\lambda}_j)$  and  $ilde{\xi}_j \perp \!\!\! \perp \!\!\! \tilde{\xi}_k$  for j 
  eq k
- Simplification: the  $\xi_j$  do not follow a normal
- $\tilde{F}(\omega)(t)$ ,  $\hat{\mu}(t)$  and  $\hat{\nu}_i(t)$  are approximated as vectors in  $\mathbb{R}^{401}$ .

# Simulation Setup & Application cont.

Following **Reiss and Ogden (2007)**, let  $f_1(t)$  and  $f_2(t)$  be two coefficient functions:

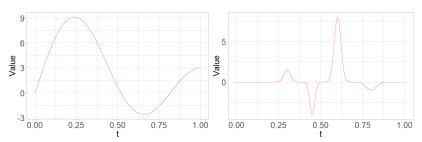


Figure:  $f_1(t)$ , smooth function Figure:  $f_2(t)$ , bumpy function

# Simulation Setup & Application cont.

Let

$$Y_{1,f} = \langle NIR, f \rangle + Z \frac{var(\langle NIR, f \rangle)}{0.9} - var(\langle NIR, f \rangle)$$
  
 $Y_{2,f} = \langle NIR, f \rangle + Z \frac{var(\langle NIR, f \rangle)}{0.9} - var(\langle NIR, f \rangle)$ 

where  $Z \sim \mathcal{N}(0,1)$  be two responses for  $f \in \{f_1(t), f_2(t)\}.$ 

- Four combinations with different number of cubic basis-function  $n_{basis} \in (5, 6, ..., 25)$  to perform regression using basis expansion and the FPCR approach.
- Compare results via criteria (CV, Mallows CP,...)
- add results here!



# Simulation Setup & Application cont.

- Use insights from the simulation study to uncover dependence.
- Similar setup, but using only bsplie basis expansion and initial 60 spectral curves.
- Validation set approach: Scores of testdata needs to be estimated by the trainingdata. explain in detail?
- Report results by MSE scaled by variance.

# Summary

Jona

Just summarize what we have done...

# further reading

Put footnotes here!

