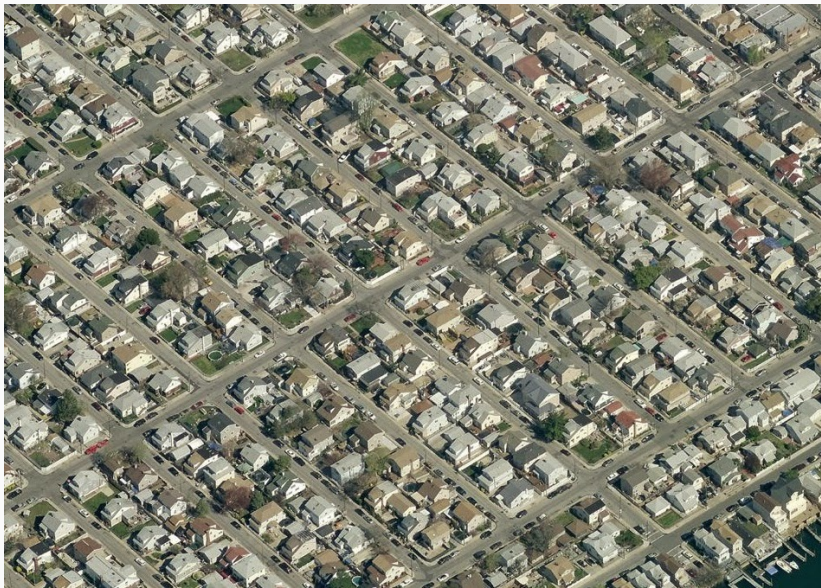


Lot Sizes, Welfare and Urban Form: A View from the United States

James Macek

August 15, 2022



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- Lot size regulation appears to be mediating the burden of high housing prices on low income families, causing considerable residential sorting on income and race (Kulka, 2019) (Song, 2021).
- No previous work in measuring the welfare impacts in general equilibrium.

1. Income sorting has important implications for the *aggregate* welfare consequences of regulation, building on [Hsieh and Moretti \(2019\)](#):
 - Deregulation causes high income, productive households to sort out of productive cities
 - Exacerbated when local amenities are endogenous
 - Ensuing externalities, e.g. [Hamilton \(1976\)](#).

In this paper

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 - Ensuing externalities, e.g. [Hamilton \(1976\)](#).
2. Evidence that *heterogeneous* regulation has altered *urban form* in expensive cities. Crucial for assessing welfare consequences for two reasons:
 - Households can move *within* cities to avoid large lots
 - Within-city income sorting is an important margin bolstering welfare inequality

How?

- Build and calibrate a structural GE model with heterogeneous households, cities and neighborhoods across the contiguous United States
 - (Heterogeneous) regulations cause income sorting, and income sorting shapes the pattern of neighborhood amenities
 - ... and also shape patterns of urban density in a way that matches observation

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 - ...and also shape patterns of urban density in a way that matches observation
- Counterfactual: What if we decreased the lot size restrictions we observe in the suburbs of expensive cities?
- Housing regulation is hard to measure, but recent advances...
 - [Song \(2021\)](#) shows that a structural break detection algorithm works well for minimum lot sizes
 - I hope to leverage assessment data with nation-wide coverage

Literature

- **Macroeconomics of housing regulation** (Hsieh and Moretti, 2019) (Duranton and Puga, 2019) (Parkhomenko, 2020) (Bunten, 2017) (Herkenhoff et al., 2018) (Ganong and Shoag, 2017) (Gyourko et al., 2013)
- **Lot Size/Unit density regulation** (Kulka, 2019) (Song, 2021) (Kulka et al., 2022) (Zabel and Dalton, 2011) (Gyourko et al., 2021) (Grieson and White, 1981) (Gyourko and Voith, 1997) (Davidoff et al., 2022)
- **Housing Regulation + Supply + Affordability** (Baum-Snow and Han, 2021) (Saiz, 2010) (Asquith et al., 2021) (Mast, 2020) (Albouy et al., 2016) (Bertaud and Brueckner, 2005) (Mills, 2005) (Brueckner and Singh, 2020) (Brueckner et al., 2017) (Acosta, 2021) (Martynov, 2022) (Turner et al., 2014) (Gyourko and Molloy, 2015)
- **Urban spatial sorting** (Diamond, 2016) (Couture and Handbury, 2020) (Su, 2022) (Baum-Snow and Hartley, 2020) (Couture et al., 2019) (Amalgro and Dominguez-lino, 2021) (Brueckner et al., 1999) (Glaeser et al., 2008) (Brueckner and Rosenthal, 2009) Lee and Lin (2017)
- **Inequality in cities** (Baum-Snow and Pavan, 2013) (Eeckhout et al., 2014) (Fogli and Guerrieri, 2019)
- **Exclusionary Zoning** (Hamilton, 1975) (Calabrese et al., 2007) (Fernandez and Rogerson, 1997) (Calabrese et al., 2011) (Barseghyan and Coate, 2016) (Brueckner, 2021)

Motivating evidence

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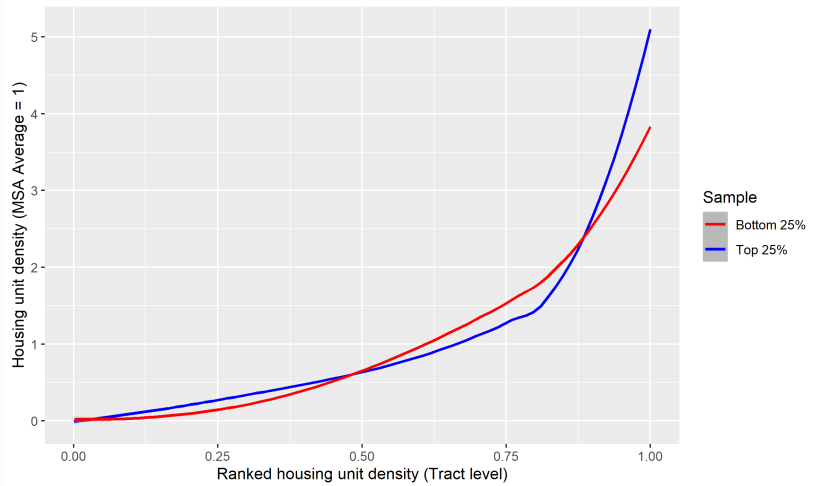
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- Superstars are MSAs in the top quartile of (unadjusted) housing prices, Non-superstars in the bottom quartile
- Housing unit density is **normalized to be on average 1** across tracts within each MSA. Controls for MSA fixed effects.

Fact 1: Density Gradients Differ in Expensive Cities



Fact 1: Why might lot size regulation be generating these patterns?

1. Decompose Housing Unit Density into a *Single Family* margin and an *Other Structures* margin [Go!](#)
2. Further decompose Single Family margin into *Land* and *Density* margins, using assessment data [Go!](#)
3. Fact 1 does *not* replicate in contemporary French cities [Go!](#)
4. TBD: Does Fact 1 replicate in the early 20th century US?

Fact 2: Strong Income Sorting on Density in Expensive Cities



Model and Calibration

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- Each tract i has land $T(i)$ available for residential development.

Developer's Problem

- Start with a standard Cobb-Douglas production function for housing, yielding housing supply per unit of land

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$$A^*(i) = \left[\frac{P(i)}{\lambda(i)} \right]^{\epsilon(i)} l(i) \quad (2)$$

- Note: $P(i)A^*(i)$ is increasing in $P(i)$.

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$$\max_{A, g} b(i, z) A^{\beta} g^{1-\beta} \quad (3)$$

subject to $A \geq A^*(i)$ and $P(i)A + g \leq w(i)z$.

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- Let $V[P(i), w(i), z, A^*(i)] := V(i, z)$ be the indirect utility associated with this problem.

Neighborhood Choice

- Frechet shocks over neighborhoods, allowing for within-city correlation ρ . Labour supply for type z households are

$$L(i, z) = \left[\frac{W(C(i), z)}{\mathbf{W}(z)} \right]^\theta \left[\frac{V(i, z)}{W(C(i), z)} \right]^{\frac{\theta}{1-\rho}} \bar{L}(z) \quad (4)$$

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and

$$\mathbf{W}(z) = \left[\sum_{c \in C} W(c, z)^\theta \right]^{\frac{1}{\theta}}$$

Endogenous Amenities

- Let $\tau(i)$ be the commuting time in i . I allow amenity values $b(i, z)$ to respond to neighborhood income per capita via the equation:

$$\log [b(i, z)] = -\kappa\tau(i) + \Omega \log \left[\frac{\sum_{z' \in Z} w(i)z' L(i, z')}{\sum_{z' \in Z} L(i, z')} \right] + \epsilon(i, z) \quad (5)$$

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- Two microfoundations
 1. Local, non-traded Dixit-Stiglitz market ([Couture et al., 2019](#)) ([Amalgro and Dominguez-lino, 2021](#)) + local disutility of density ([Kulka et al., 2022](#))

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 2. Local, congested public good financed through property taxes ([Calabrese et al., 2007](#)) ([Calabrese et al., 2011](#))

- Numeraire good is produced using only units of labour under *external increasing returns*.
- Inverse labour demand given by

$$w(c) = \sigma(c) \left[\sum_{z \in Z} \sum_{i \in N(c)} zL(i, z') \right]^\alpha \quad (6)$$

for $\alpha > 0$.

- Past literature estimates α as a population elasticity, but with heterogenous households this interpretation is lost.
- Apparently, hard to separately identify agglomeration effects on output and population, See [Combes and Gobillon \(2015\)](#) Section 3.1.

Calibration and Identification

Minimum Lot Sizes

- Follows a structural break algorithm in [Song \(2021\)](#).
 - Idea: Zoning variances are present, but rare
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- I generalize this procedure to multifamily homes (duplexes, triplexes and fourplexes)
- Problem: Lot size policy may vary within census tracts.
 - If multiple minimum lot sizes overlap on a tract, consider the smallest one
 - Assign the measured minimum lot size to a tract if the observed number of housing units does not exceed the *implied* maximum under regulation.

Regional Production Functions

- Follows [Baum-Snow and Han \(2021\)](#) closely, with some differences
 1. Use assessment data to construct high quality measures of housing prices $P(i)$ via a hedonic regression
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- Given stock measures $A(i)$, prices $P(i)$ and elasticities $\epsilon(i)$ from [Baum-Snow and Han \(2021\)](#), supply shifters $\lambda(i)$ are uniquely identified ...
- and so is the minimum housing stock per unit $A^*(i)$

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- For now: choose a value of θ and ρ from the literature
 1. [Hornbeck and Moretti \(2018\)](#) measure $\theta \approx \frac{10}{3}$
 2. [Herzog \(2022\)](#) estimates $\frac{\theta}{1-\rho} = 7.21$ in a closed city model of commuting, implying $\rho \approx 0.5377$

- Recall the estimating equation:

$$\log [b(i, z)] = -\kappa_{\mathcal{T}}(i) + \Omega \log \left[\frac{\sum_{z' \in \mathcal{Z}} w(i) z' L(i, z')}{\sum_{z' \in \mathcal{Z}} L(i, z')} \right] + \epsilon(i, z)$$

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- Two sources preventing identification:
 1. Simultaneity bias, analogous to [Couture et al. \(2019\)](#) and [Diamond \(2016\)](#).
 2. Exogenous, unobserved features that cause amenities and income sorting, **but do not respond to changes in local incomes.**

Estimation of Ω

- I propose an instrument – tract level terrain slopes. There are two competing interpretations of its validity:
 - [Saiz \(2010\)](#): Slopes primarily affect housing prices through *supply constraints*. **Validates the instrument.**
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- To address the amenities interpretation, I control for a large set of observed local amenities from [Lee and Lin \(2017\)](#) and the National Land Cover Database + MSA Fixed Effects
- For most amenities variables v_i and average slope s_i , I construct an additional control $v_i s_i$
 - Accounts for supermodularities between slopes and natural features
 - Example: can have an ocean view from farther away if terrain is sloped

Estimation of Ω : First Stage

VARIABLES	(1) Log(Average Income)	(2) Log(Average Income)	(3) Log(Average Income)
Average Slope	0.022*** (0.002)	0.024*** (0.004)	0.010 (0.012)
Observations	58,498	58,350	58,350
R-squared	0.038	0.007	0.000
KP F-Stat	100.397	35.884	.588
MSA FE	Yes	Yes	Yes
MSA Clustering	Yes	Yes	Yes
Amenity Controls	No	Yes	Yes incl. slope x max july temperature

Robust standard errors in parentheses

*** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$

Figure 1: Amenity Controls: A set of approximately 240 variables, merging data from [Lee and Lin \(2017\)](#) and measures of forest and perennial snow cover from the NLCD. [Lee and Lin \(2017\)](#) controls include distance to various bodies of water, temperature, precipitation, different types of community names (i.e. "gulch", "shores"). Includes most interactions of these variables with average slopes.

Conclusion

- I provide observational evidence that *heterogeneous* lot size regulation has altered urban form – both the shape of cities and how households of various incomes sort into them.
- Income sorting in the face of lot size regulation has important implications for aggregate welfare.
 - A channel that has been (sadly) ignored.
- This motivates a structural model that can simultaneously capture all these forces.
- Thank you for all the help this year!

Appendix

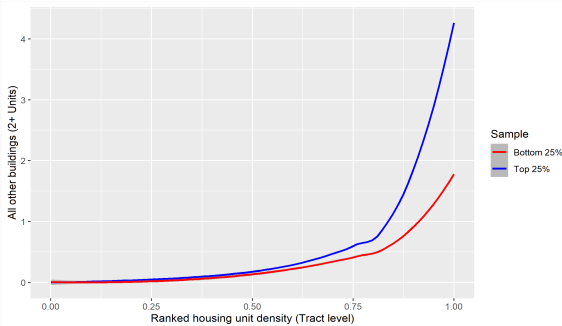
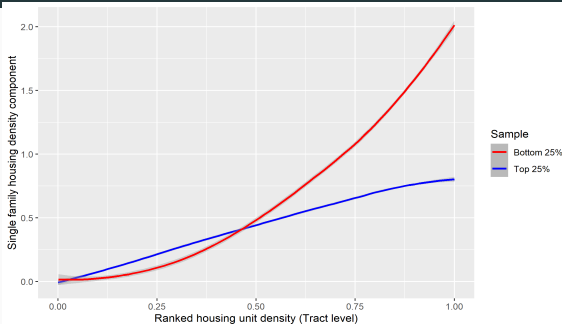
Appendix: Single Family Margin vs Other Structures

- The shape of the distribution *may not* be driven the presence of structures that tend to be heavily regulated.
- Linearly decompose housing unit density $D_{H,im}$ into two margins:

$$D_{H,im} = D_{S,im} + D_{L,im} \quad (7)$$

- where $D_{S,im}$ is the number of single family homes in MSA m and tract i divided by the total land mass of the tract.
- $D_{L,im}$ are defined analogously for all other structures.
- Single family home shares are from the 2008-2012 ACS.
- **Repeat** the regression for each component separately. Justified because the conditional expectation is additive.

Appendix: Single Family and Other Structures margins

[Return!](#)

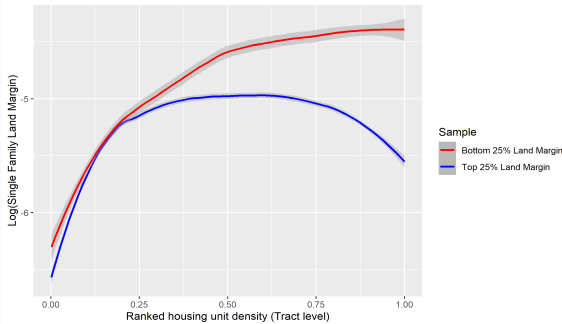
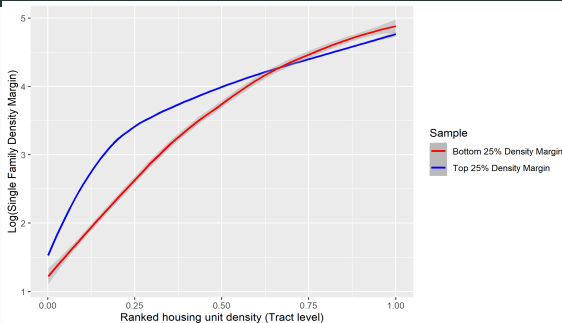
Appendix: Single Family Density and Land margins

- Log-linearly decompose the single family component into *density* and *land* margins, respectively:

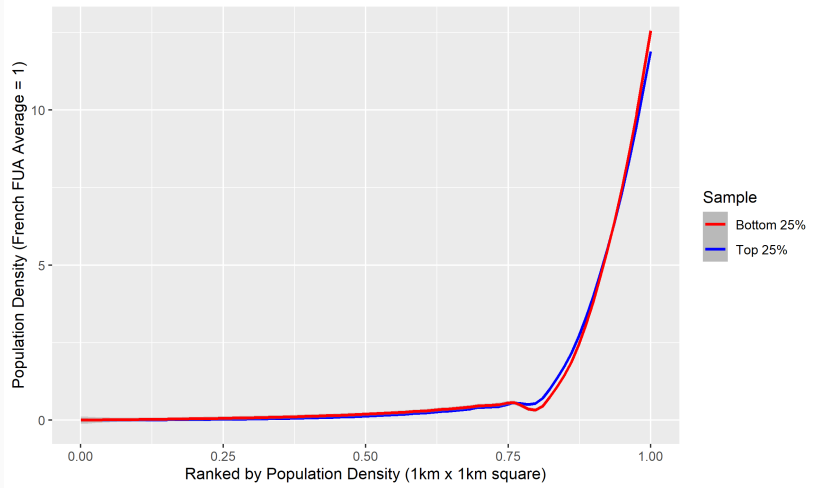
$$\log(D_{S,im}) = \log(\tilde{D}_{S,im}) + \log(L_{S,im}) \quad (8)$$

- where $\tilde{D}_{S,im}$ is the number of single family homes divided by the total land used for single family housing, and $L_{S,im}$ is the fraction of tract land used for single family housing.
- Repeat the regression for each component separately.
- Use 2015 assessment data on single family homes to measure average land use per housing unit, with unit counts from the ACS/Census.

Appendix: Single Family Density and Land margins

[Return!](#)

Appendix: Fact 1 replication with French Data [Return!](#)



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