

ECE QE CS-1 2018

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1 Question 1

(a)

$$\begin{aligned} P(\min(X, Y) = k) &= P(X = k \cap Y > k) + P(Y = k \cap X > k) + P(X = k \cap Y = k) \\ &= \frac{1}{2^k} * \sum_{i=k+1}^{+\infty} \frac{1}{2^i} + \frac{1}{2^k} * \sum_{i=k+1}^{+\infty} \frac{1}{2^i} + \frac{1}{2^k} * \frac{1}{2^k} \\ &= \left(\frac{1}{2}\right)^k * 2\left(\frac{1}{2}\right)^{k+1} + \left(\frac{1}{2}\right)^k * 2\left(\frac{1}{2}\right)^{k+1} + \left(\frac{1}{2}\right)^{2k} \\ &= 3 * \left(\frac{1}{2}\right)^{2k} = 3\left(\frac{1}{4}\right)^k \end{aligned} \tag{1}$$

(b)

$$\begin{aligned} P(X = Y) &= \sum_{i=1}^{+\infty} \frac{1}{2^i} * \frac{1}{2^i} \\ &= \sum_{i=1}^{+\infty} \left(\frac{1}{4}\right)^i = \frac{(1/4) * (1-0)}{1 - (1/4)} \\ &= \frac{1}{3} \end{aligned} \tag{2}$$

(c)

$$\begin{aligned} P(Y > X) &= \sum_{i=1}^{+\infty} \sum_{j=i+1}^{+\infty} \frac{1}{2^i} * \frac{1}{2^j} \\ &= \sum_{i=1}^{+\infty} \left(\frac{1}{2}\right)^i * \frac{(1/2)^{i+1}}{1 - (1/2)} \\ &= \sum_{i=1}^{+\infty} \left(\frac{1}{4}\right)^i \\ &= \frac{1}{3} \end{aligned} \tag{3}$$

(d)

$$\begin{aligned} P(Y = kX) &= \sum_{i=1}^{+\infty} \frac{1}{2^i} * \frac{1}{2^{ki}} \\ &= \frac{1}{1 - \left(\frac{1}{2}\right)^{k+1}} \end{aligned} \tag{4}$$

2 Question 2

(a)

$$\begin{aligned} P(N = n) &= \int_{-\infty}^{+\infty} p_N(n|X = x) f_X(x) dx = \int_0^1 x^n (1-x) dx \\ &= \frac{1}{n+1} x^{n+1} - \frac{1}{n+2} x^{n+2} \Big|_0^1 = \frac{1}{n+1} - \frac{1}{n+2} = \frac{1}{(n+1)(n+2)} \end{aligned} \tag{5}$$

(b) Using Bayes' theorem,

$$f_X(x|N=n) = \frac{p_N(n|X=x)f_X(x)}{p_N(n)} = (n+1)(n+2)x^n(1-x) \quad (6)$$

,when $0 \leq x \leq 1$

(c)

$$\begin{aligned} MMSE = E(X|N=n) &= \int_{-\infty}^{+\infty} x * f_X(x|N=n)dx = \int_0^1 (n+1)(n+2)x^{n+1}(1-x)dx \\ &= (n+1)(n+2) \left(\frac{1}{n+2}x^{n+2} - \frac{1}{n+3}x^{n+3} \right) \Big|_0^1 \\ &= \frac{n+1}{n+3} \end{aligned} \quad (7)$$

3 Question 3

(a) If $X(t), Y(t)$ is WSS, $\mu_{X(t)} = \mu_X$ and $\mu_{Y(t)} = \mu_Y$. Therefore, $\mu_{Z(t)} = \mu_X \mu_Y$.

$$\begin{aligned} r_{ZZ}(t_1, t_2) &= E(X(t_1)Y(t_1)X(t_2)Y(t_2)) \\ &= E(X(t_1)X(t_2)) * E(Y(t_1)Y(t_2)) \\ &= r_{XX}(l) * r_{YY}(l) \end{aligned} \quad (8)$$

This result only depends on the time difference. Therefore, Z is WSS.

(b) Since Y is WSS. We only need to verify X is WSS.

$$\begin{aligned} E(X(t)) &= \int_{-\infty}^{+\infty} \theta \cos(\omega_0 t + \theta) d\theta \\ &= \theta \sin(\omega_0 t + \theta) \Big|_{-\infty}^{+\infty} - \int_{-\infty}^{+\infty} \sin(\omega_0 t + \theta) d\theta \\ &= 0 \end{aligned} \quad (9)$$

$$\begin{aligned} r_{XX}(t_1, t_2) &= E(X(t_1)X(t_2)) = E(\cos(\omega_0 t_1 + \theta) * \cos(\omega_0 t_2 + \theta)) \\ &= \frac{1}{2} E(\cos(\omega_0 t_1 - \omega_0 t_2) + \cos(\omega_0 t_1 + \omega_0 t_2 + 2\theta)) \\ &= \frac{1}{2} (\cos(\omega_0 t_1 - \omega_0 t_2) + \frac{1}{2} E(\cos(\omega_0 t_1 + \omega_0 t_2) \cos(2\theta) - \sin(\omega_0 t_1 + \omega_0 t_2) \sin(2\theta))) \\ &= \frac{1}{2} (\cos(\omega_0 t_1 - \omega_0 t_2) + \int_0^{2\pi} \theta \cos(2\theta) d\theta * \cos(\omega_0 t_1 + \omega_0 t_2) - \int_0^{2\pi} \theta \sin(2\theta) d\theta * \sin(\omega_0 t_1 + \omega_0 t_2)) \end{aligned} \quad (10)$$

Since two integral parts are both 0. The result is $\frac{1}{2} \cos(\omega_0 t_1 - \omega_0 t_2)$. It only depends on the difference of t_1 and t_2 . So Z is also WSS.

Then we will find the power spectral density of Z:

$$\begin{aligned} R_Z &= R_X R_Y = \frac{1}{2} \cos(\omega_0 t) e^{-\alpha|t|} \\ &= \frac{1}{2} \frac{e^{i\omega_0 t} + e^{-i\omega_0 t}}{2} e^{-\alpha|t|} \end{aligned} \quad (11)$$

$$\begin{aligned} F\{e^{-\alpha|t|}\} &= \int_{-\infty}^{+\infty} e^{-\alpha|t|} e^{i\omega t} dt \\ &= \int_{-\infty}^0 e^{\alpha t} e^{i\omega t} dt + \int_0^{\infty} e^{-\alpha t} e^{i\omega t} dt \\ &= \frac{1}{\alpha - i\omega} + \frac{1}{\alpha + i\omega} \\ &= \frac{2\alpha}{\alpha^2 + \omega^2} \end{aligned} \quad (12)$$

$$S_Z(\omega) = F\{R_Z(t)\} = \frac{\alpha}{2(\alpha^2 + (\omega - \omega_0)^2)} + \frac{\alpha}{2(\alpha^2 + (\omega + \omega_0)^2)} \quad (13)$$

4 Question 4

$$\begin{aligned}\Phi_Y(y) &= E(e^{iwy}) = E(e^{iw*(1/n) \sum_{i=1}^n X_i}) = \prod_{i=1}^n E(e^{iw \frac{X_i}{n}}) \\ &= \prod_{i=1}^n e^{\frac{-|w|}{n}} = e^{-|w|}\end{aligned}\tag{14}$$

$$\begin{aligned}f_Y(y) &= \frac{1}{2\pi} \int_{-\infty}^{+\infty} e^{-iwy} e^{-|w|} dw \\ &= \frac{1}{2\pi} \int_{-\infty}^0 e^{w(-iy+1)} dw + \int_0^{+\infty} e^{w(-iy-1)} dw \\ &= \frac{1}{2\pi} \left(\frac{1}{1-iy} + \frac{-1}{-1-iy} \right) \\ &= \frac{1}{\pi} \frac{1}{1+y^2}\end{aligned}\tag{15}$$

Therefore, the pdf shows that $Y_1, Y_2, Y_3 \dots$ converge in distribution. And it is convergent to a Cauchy distribution.