Lesson 6

Maximum likelihood & Bayesian Theory

Trees are the natural way of representing relationships among species.

The number of possible trees increases factorially with the number of OTUs.





So far we have seen methods based on generating distances between OTUs

...methods that are based on using distance matrices...



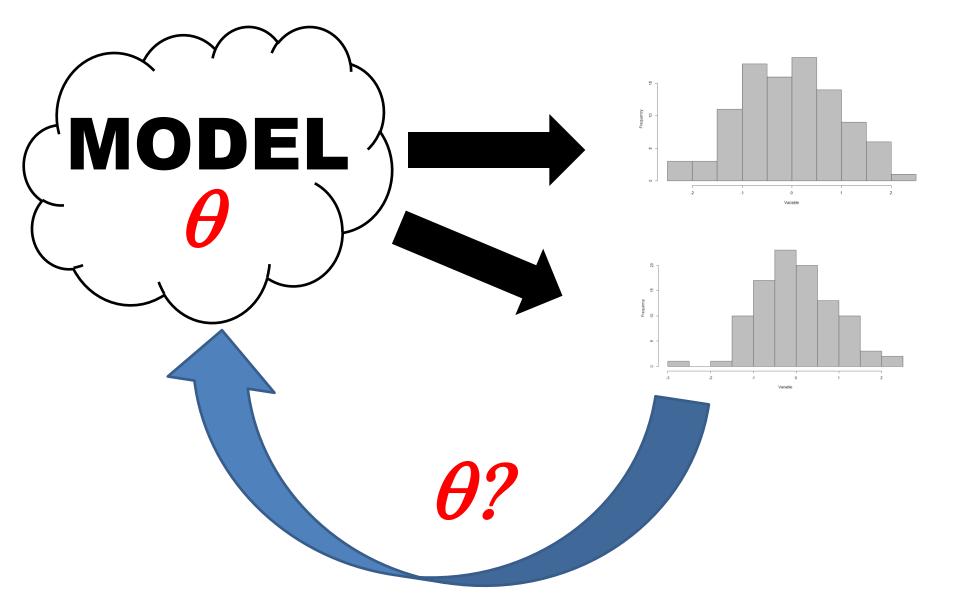


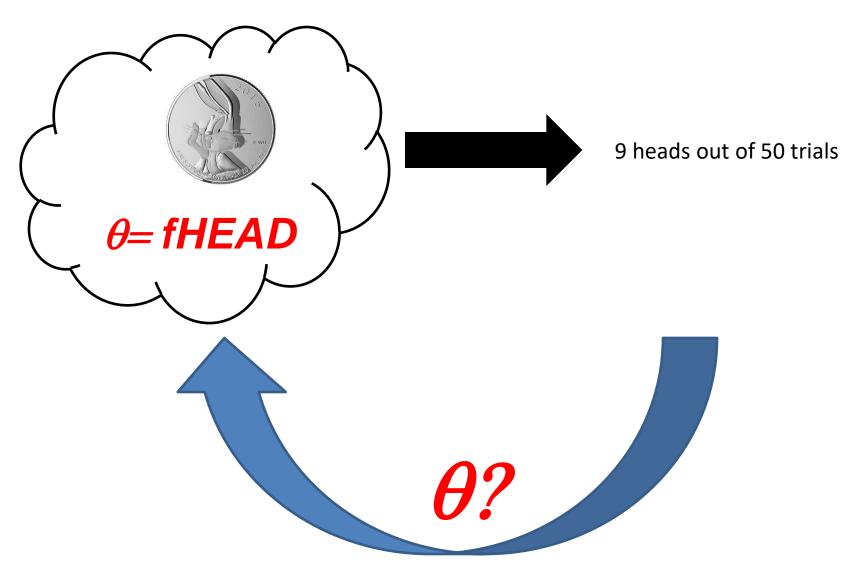
Wait a moment!
LIKELIHOOD!!! That
sounds like statistics!!!!!!

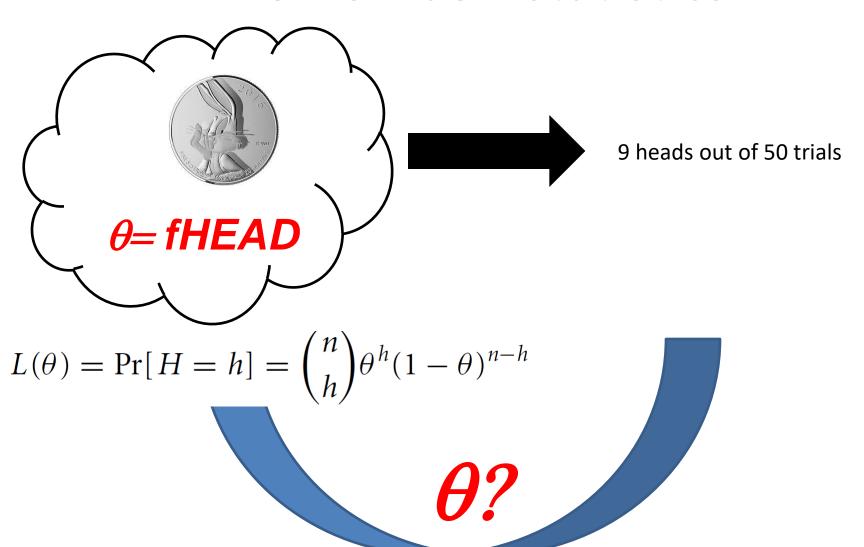
NOOOOOO!!!!! **Statistics NOOOOOOOO**!!!!!!

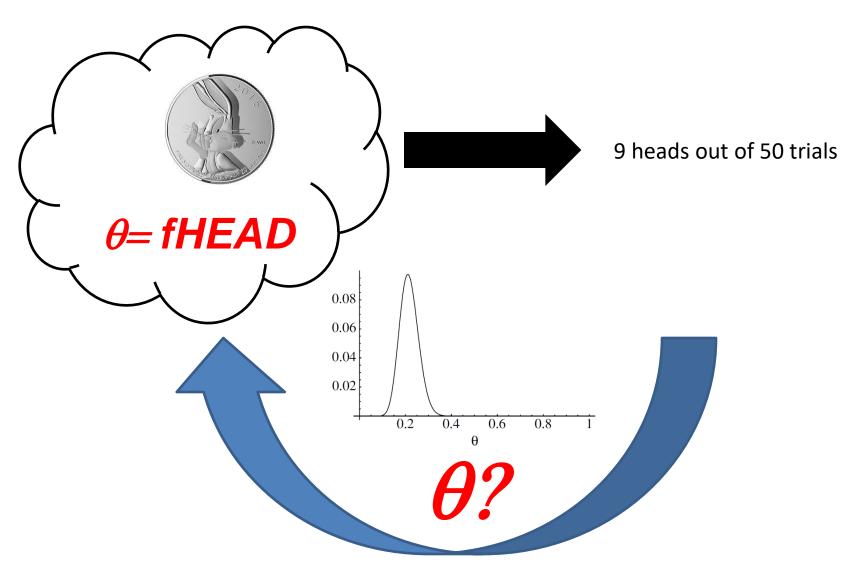


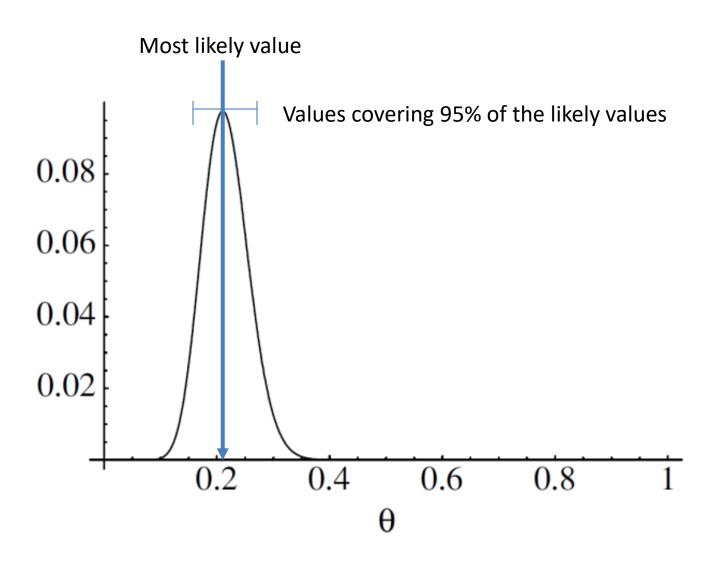












How do we recover the most likely value of the parameter θ that produced the data?

$$L(\theta) = \Pr[H = h] = \binom{n}{h} \theta^h (1 - \theta)^{n-h}$$

$$h = 9$$

$$n = 50$$

$$\log(L(\theta)) = \log(C) + h * \log(\theta) + (n-h)*\log(1 - \theta)$$

$$L(\theta) = \log(C) + h * \log(\theta) + (n-h)*\log(1-\theta)$$

$$\frac{dL}{d\theta} = \frac{h}{\theta} - \frac{n-h}{1-\theta}$$

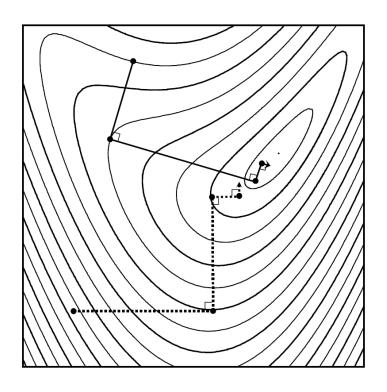
$$\frac{dL}{d\theta} = 0; \theta = \frac{h}{n}$$

How do we recover the most likely value of the parameter θ that produced the data?

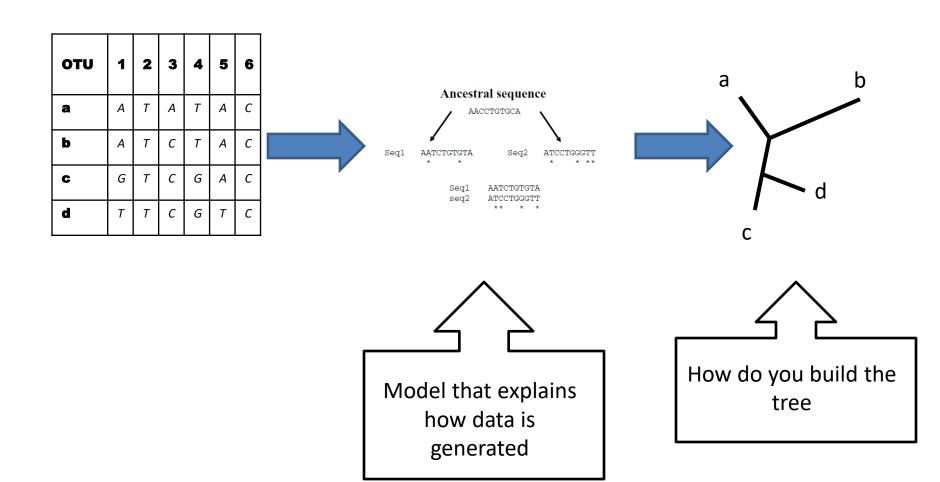
$$\frac{dL}{d\theta} = 0$$
;?

Newton-Raphson method for finding maximum

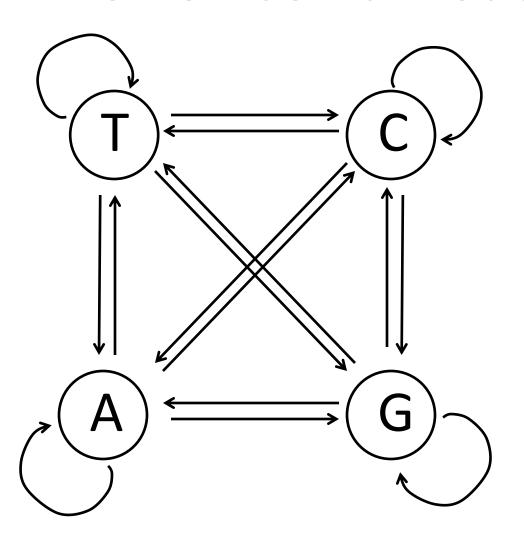
$$\theta_{k+1} = \theta_k - \frac{f'(\theta_k)}{f''(\theta_k)}.$$



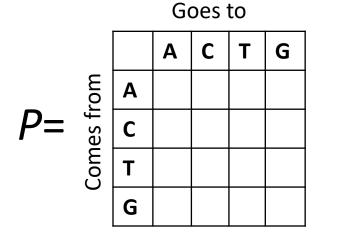
Remember: Character based methods



Remember: a model of evolution



Probability change matrix



Remember: a model of evolution

Table 1.1 Substitution-rate matrices for commonly used Markov models of nucleotide substitution

	From		T	ò			
		T	С	A	G		Т —
JC69 (Jukes and Cantor 1969)	T		λ	λ	λ		11 15
	C	λ		λ	λ	JC69	W/
	Α	λ	λ		λ	JC69	//\
	G	λ	λ	λ			14 % /4
K80 (Kimura 1980)	T		α	β	β		(A)
	C	α		β	β		
	Α	β	β		α		
	G	β	β	α			
F81 (Felsenstein 1981)	T		$\pi_{\mathbf{C}}$	π_{A}	$\pi_{\mathbf{G}}$		т —
	C	π_{T}		π_{A}	$\pi_{\mathbf{G}}$		
	Α	π_{T}	$\pi_{\mathbf{C}}$		$\pi_{\mathbf{G}}$		1 1 1
	G	π_{T}	$\pi_{\mathbf{C}}$	π_{A}		K80	W
HKY85 (Hasegawa et al. 1984, 1985)	T		$\alpha\pi_{\mathbf{C}}$	$\beta\pi_{\rm A}$	$\beta\pi_{\mathbf{G}}$		↓ 4/\}
	C	$\alpha\pi_{\mathrm{T}}$		$\beta\pi_{A}$	$\beta\pi_{\mathbf{G}}$		
	Α	$\beta \pi_{\rm T}$	$\beta\pi_{\rm C}$, ,	$\alpha\pi_{G}$		(A) ~~
	G	$\beta \pi_{\rm T}$	$\beta\pi_{\mathbf{C}}$	$\alpha\pi_{A}$			
F84 (Felsenstein, DNAML program since 1984)	T		$(1 + \kappa/\pi_Y)\beta\pi_C$	$\beta\pi_{\rm A}$	$\beta\pi_{ m G}$		
(,	Č	$(1 + \kappa/\pi_Y)\beta\pi_T$		$\beta\pi_{A}$	$\beta\pi_{\mathbf{G}}$		
	A	$\beta \pi_{\rm T}$	$\beta\pi_{\mathbb{C}}$		$(1 + \kappa/\pi_R)\beta\pi_G$		$T \longrightarrow$
	G	$\beta \pi_{\mathrm{T}}$	$\beta\pi_{\mathbf{C}}$	$(1 + \kappa/\pi_R)\beta\pi_A$			
TN93 (Tamura and Nei 1993)	T		$\alpha_1\pi_{\rm C}$	$\beta \pi_{A}$	$\beta\pi_{\mathbf{G}}$		11 1
Thos (Tallian and Not 1993)	C	$\alpha_1\pi_T$	-10	$\beta\pi_{A}$	$\beta\pi_{\mathbf{G}}$	HKY85	XX
	Α	$\beta \pi_{\rm T}$	$\beta\pi_{\rm C}$		$\alpha_2\pi_{ m G}$		₩ 🐼 📢
	G	$\beta \pi_{\rm T}$	$\beta\pi_{\rm C}$	$\alpha_2\pi_A$			A .
GTR (REV) (Tavaré 1986; Yang 1994b; Zharkikh 1994)	T		$a\pi_{\mathbf{C}}$	$b\pi_{\rm A}$	$c\pi_{\mathbf{G}}$		₩ ←
OTK (KEV) (Tavaic 1900, Tang 19940, Zharkikii 1994)	Ċ	$a\pi_{\mathrm{T}}$		$d\pi_{\rm A}$	$e\pi_{\rm G}$		
	Ā	$b\pi_{\rm T}$	$d\pi_{\mathbb{C}}$	·	$f\pi_{\mathbf{G}}$		
	G	$c\pi_{\mathrm{T}}$	$e\pi_{\mathbf{C}}$	$f\pi_{ m A}$,	T-1-1- 4 1	The 11-
UNREST (Yang 1994b)	Т					Table 4.1	The discrete-ra
OTTEST (Tang 19940)	Ċ		<i>q</i> _{TC} .	q_{TA}	q _{TG}		
	A	q_{CT} q_{AT}		<i>q</i> _{CA} ·	$q_{\rm CG}$ $q_{\rm AG}$	Site class	1 2 3
		4AI	$q_{\rm AC}$		YAG		

 q_{GC}

 q_{GA}

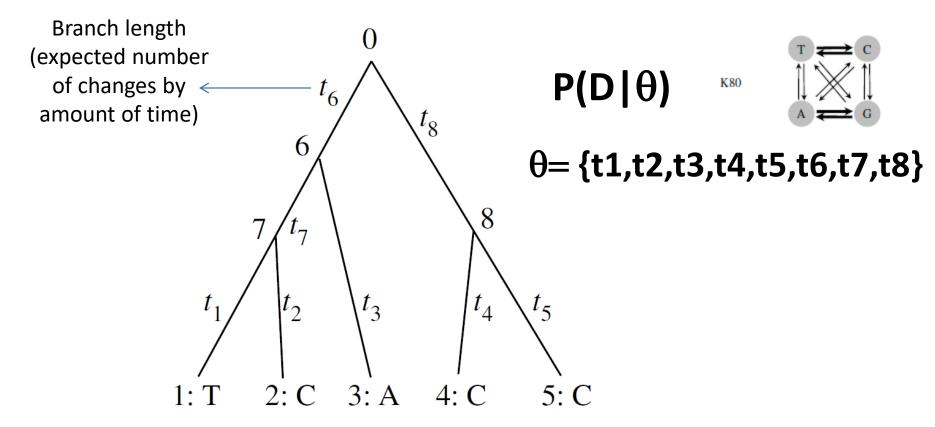
The diagonals of the matrix are determined by the requirement that each row sums to 0. The equilibrium distribution is $\pi = (1/4, 1/4, 1/4, 1/4)$ under JC69 and K80, and $\pi = (\pi_T, \pi_C, \pi_A, \pi_G)$ under F81, F84, HKY85, TN93, and GTR. Under the general unrestricted (UNREST) model, it is given by the equations $\pi Q = 0$ under the constraint $\sum_i \pi_i = 1$.

Table 4.1	The	aiscre	ete-ra	tes mo	oaei
Site class	1	2	3		K
Probability Rate					

Some basic notation

X1 2 T \boldsymbol{C} a b T*T* Α C T

And now... Imagine that we know the topology of the tree



And now... Imagine that we know the topology of the tree

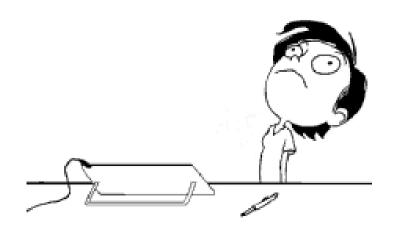
 $P(D|\theta)$

"What is the probability (likelihood) of observing the data given the values of the set θ of parameters"

Which is the likelihood of position *h* given the length of the branches and my model of nucleotide evolution?

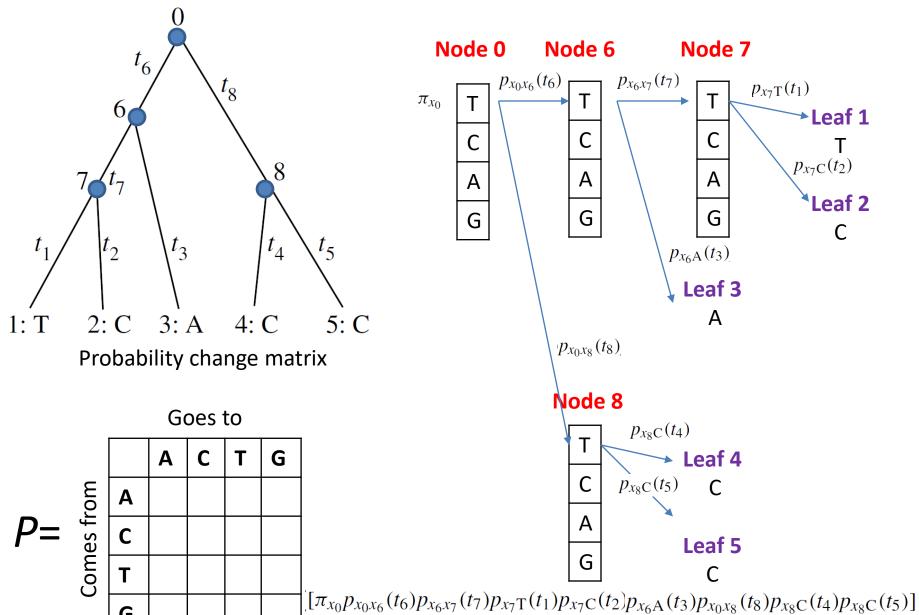
$$f(\mathbf{x}_{h}|\theta) = \sum_{x_{0}} \sum_{x_{6}} \sum_{x_{7}} \sum_{x_{8}} [\pi_{x_{0}} p_{x_{0}x_{6}}(t_{6}) p_{x_{6}x_{7}}(t_{7}) p_{x_{7}T}(t_{1}) p_{x_{7}C}(t_{2})$$

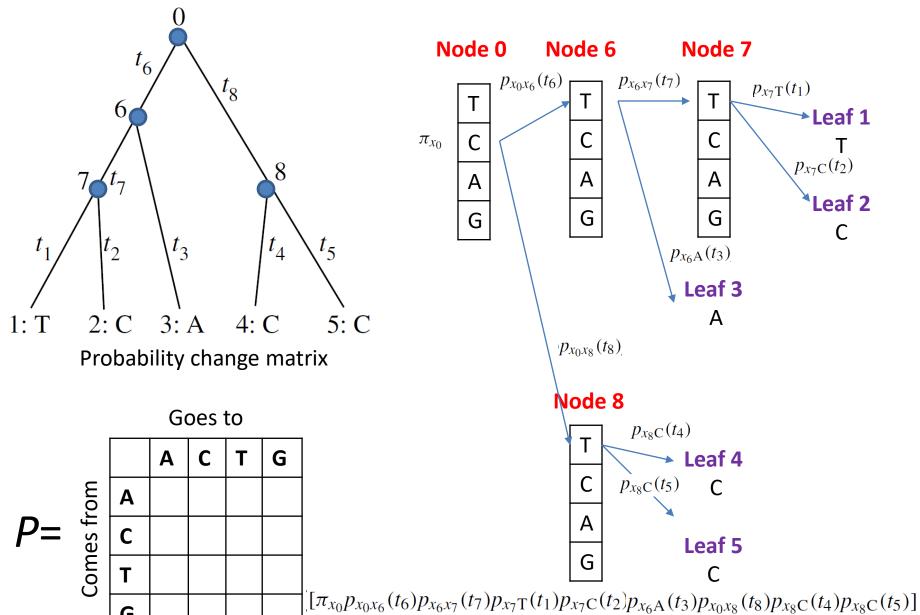
$$\times p_{x_{6}A}(t_{3}) p_{x_{0}x_{8}}(t_{8}) p_{x_{8}C}(t_{4}) p_{x_{8}C}(t_{5})].$$

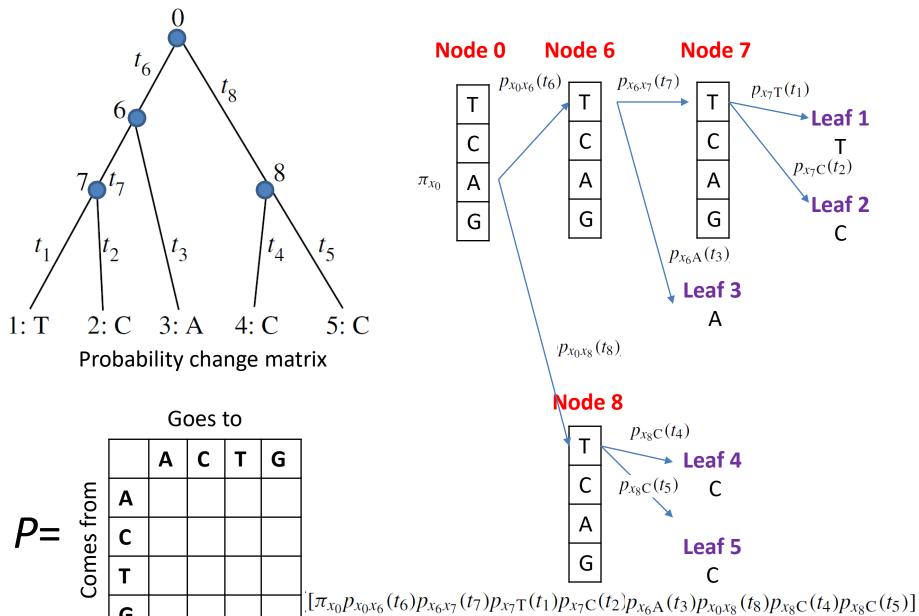


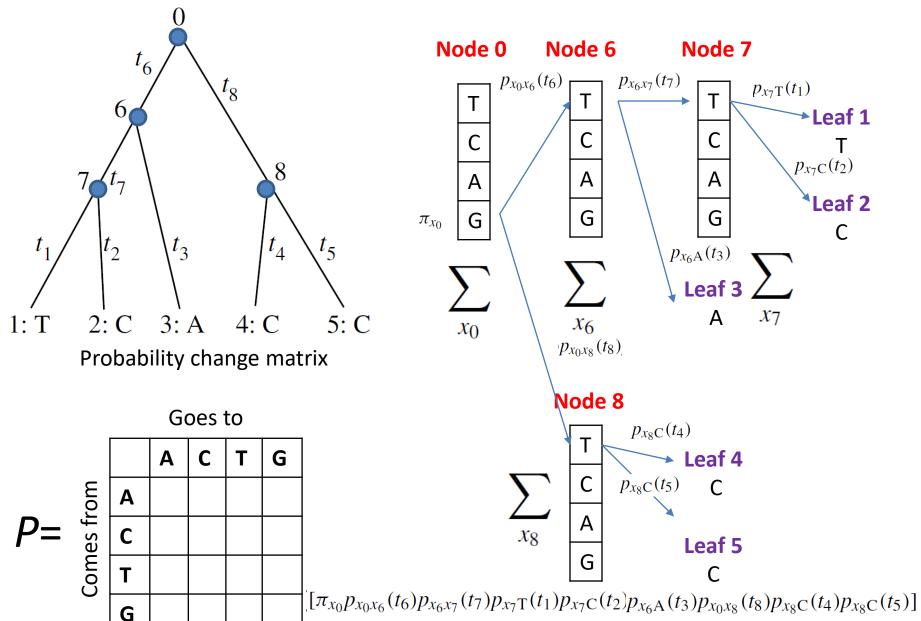
$$f(\mathbf{x}_{h}|\theta) = \sum_{x_{0}} \sum_{x_{6}} \sum_{x_{7}} \sum_{x_{8}} [\pi_{x_{0}} p_{x_{0}x_{6}}(t_{6}) p_{x_{6}x_{7}}(t_{7}) p_{x_{7}T}(t_{1}) p_{x_{7}C}(t_{2})$$

$$\times p_{x_{6}A}(t_{3}) p_{x_{0}x_{8}}(t_{8}) p_{x_{8}C}(t_{4}) p_{x_{8}C}(t_{5})].$$

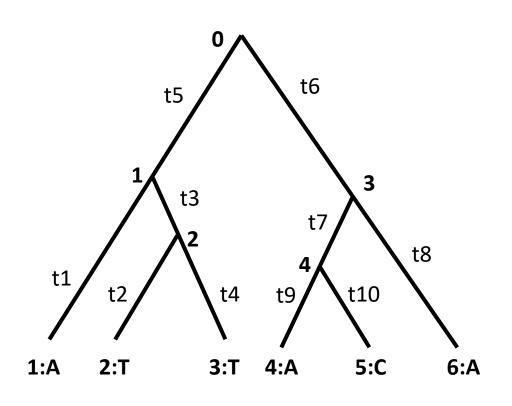








What would be the likelihood of this tree?



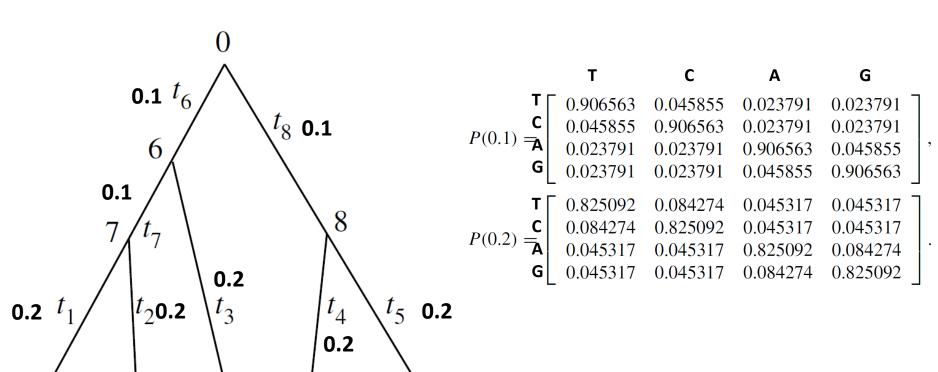


Felselstein's pruning algorithm

$$f(\mathbf{x}_{h}|\theta) = \sum_{x_{0}} \pi_{x_{0}} \left\{ \sum_{x_{6}} p_{x_{0}x_{6}}(t_{6}) \left[\left(\sum_{x_{7}} p_{x_{6}x_{7}}(t_{7})p_{x_{7}T}(t_{1})p_{x_{7}C}(t_{2}) \right) p_{x_{6}A}(t_{3}) \right] \right\}$$

$$\times \left[\sum_{x_{8}} p_{x_{0}x_{8}}(t_{8})p_{x_{8}C}(t_{4})p_{x_{8}C}(t_{5}) \right]$$

$$(4.3)$$

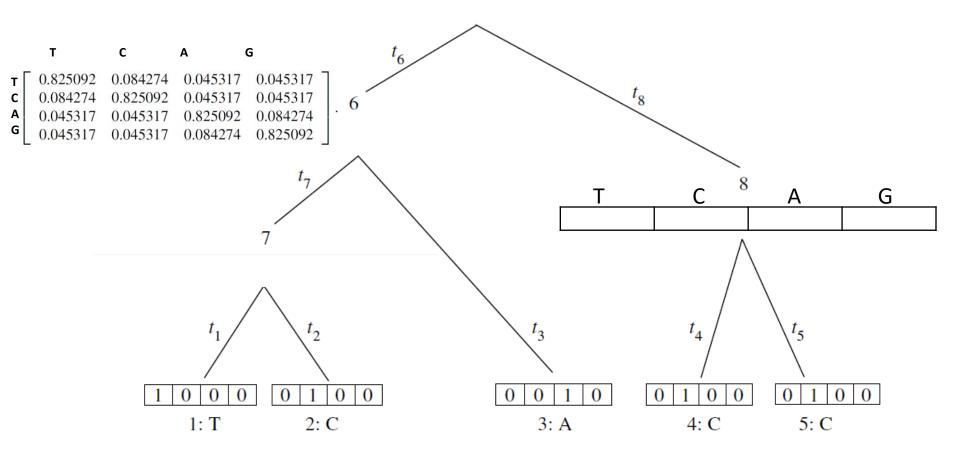


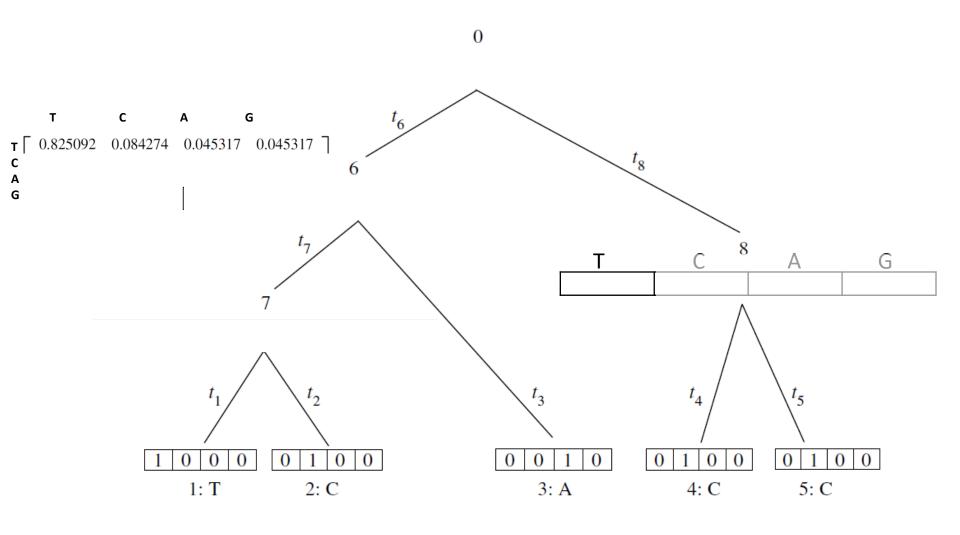
3: A

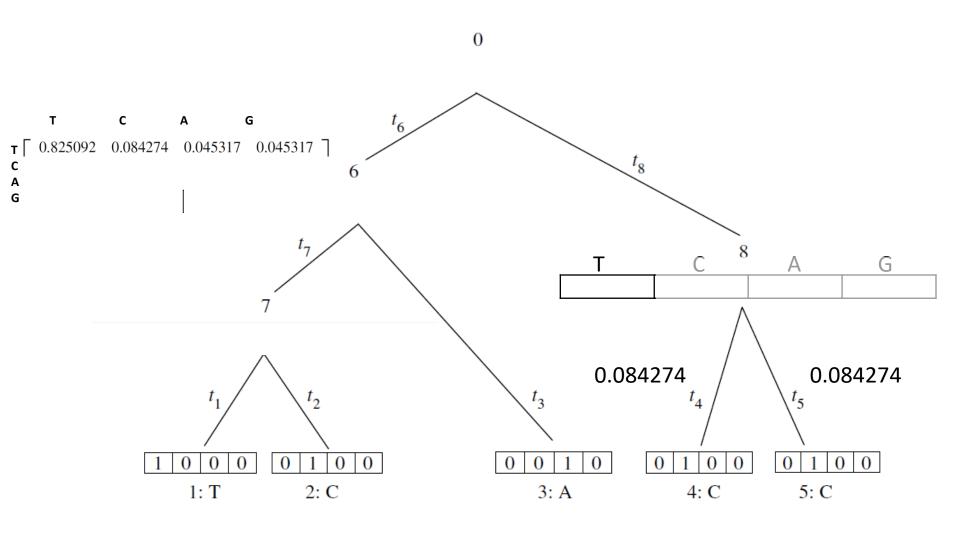
Ma M		Now I am								
ם ס		T	C	A	G					
	Т	0.825092	0.084274	0.045317	0.045317	7				
5	C	0.084274	0.825092	0.045317	0.045317					
Δ	Α	0.045317	0.045317	0.825092	0.084274					
9	G	0.045317	0.045317	0.084274	0.825092					
ב ע		_			_	_				

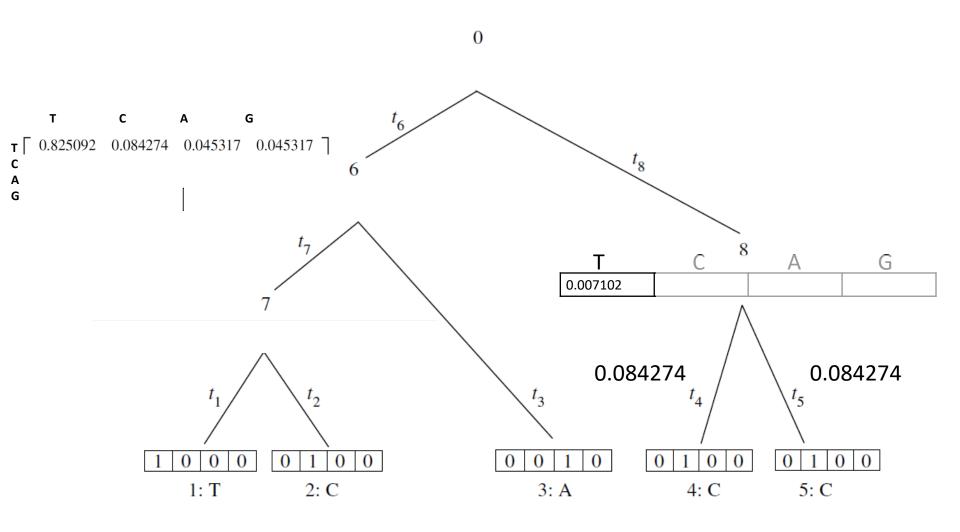
The pruning algorithm for estimating the likelihood

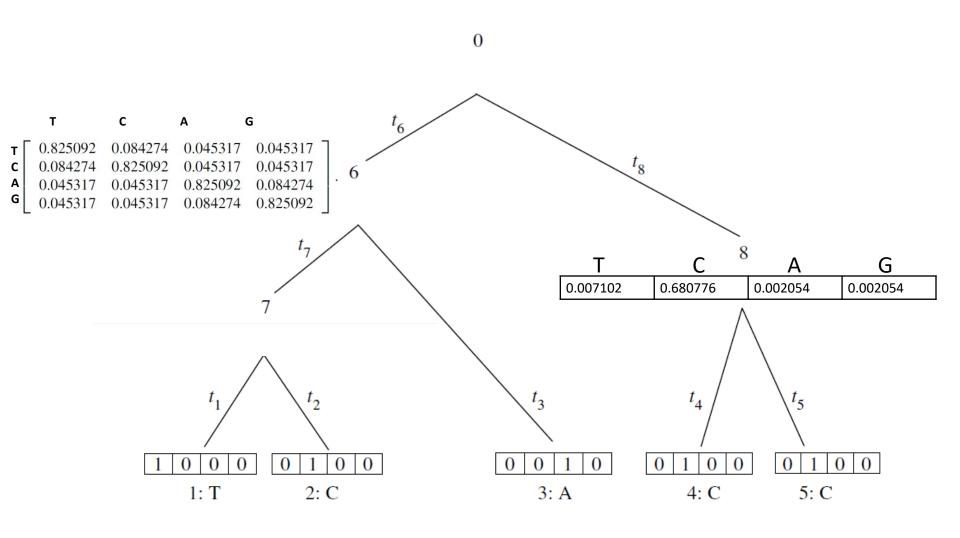
0

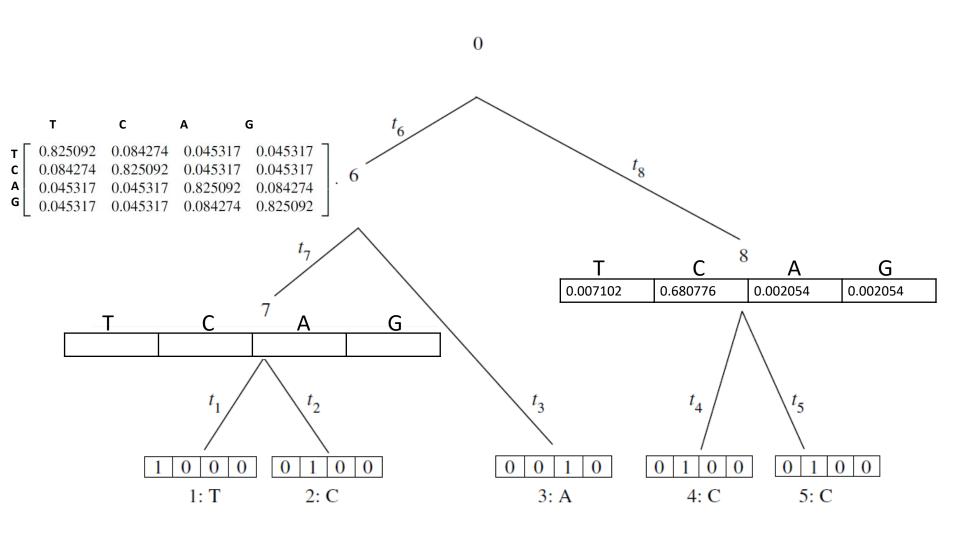


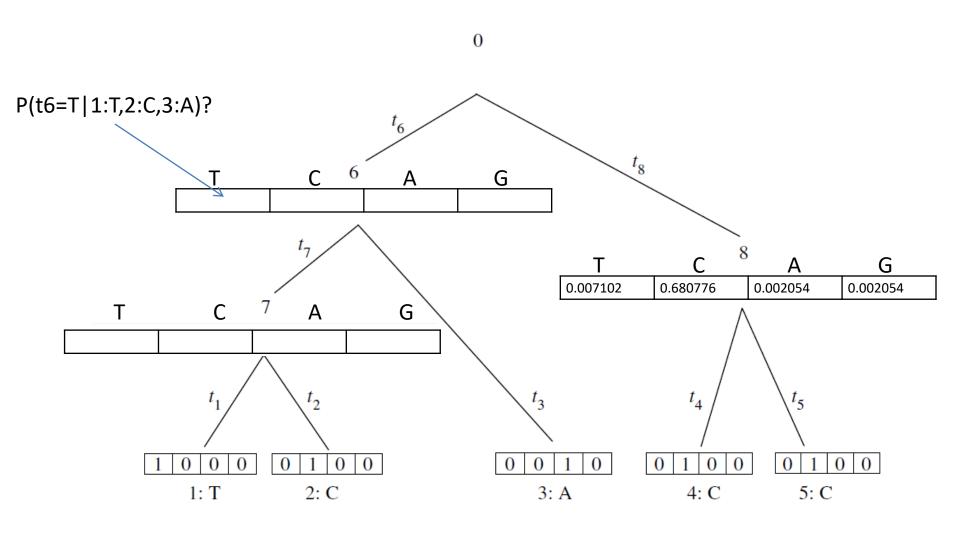


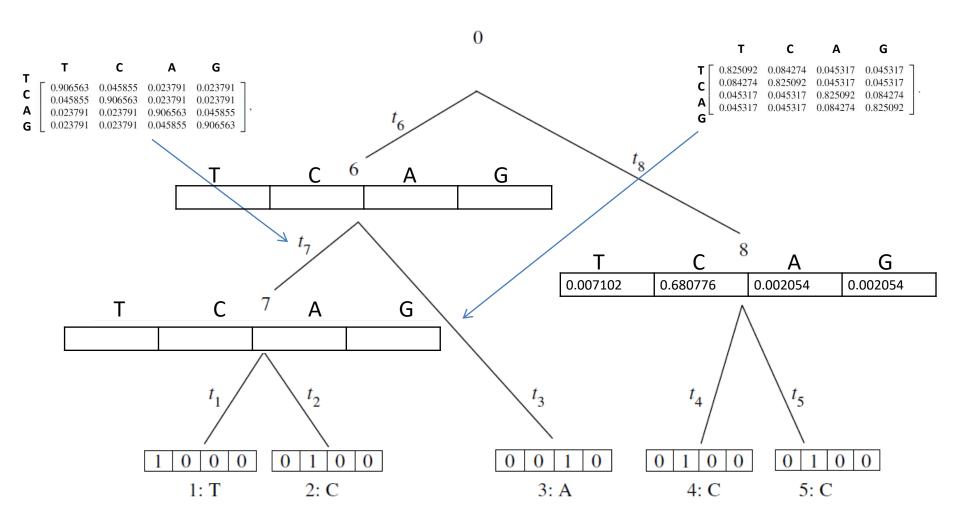


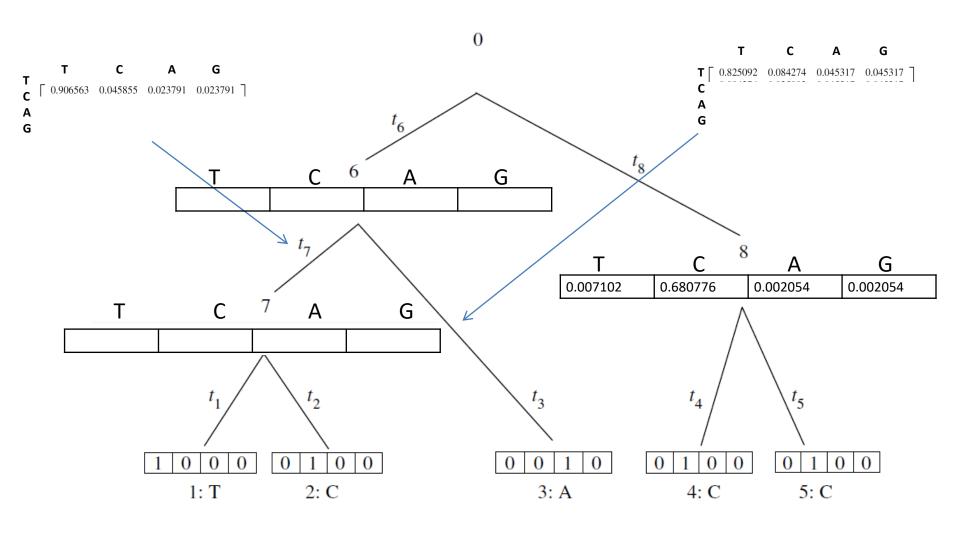


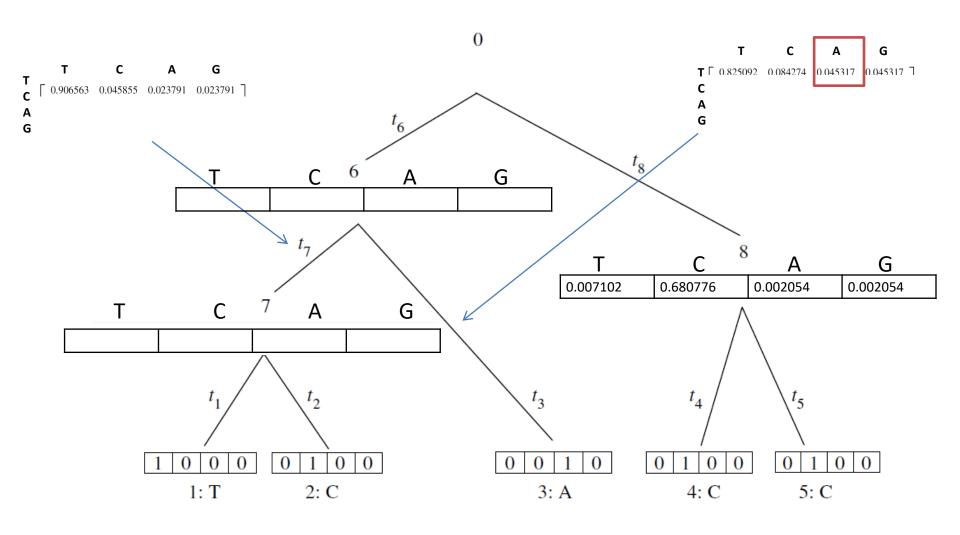


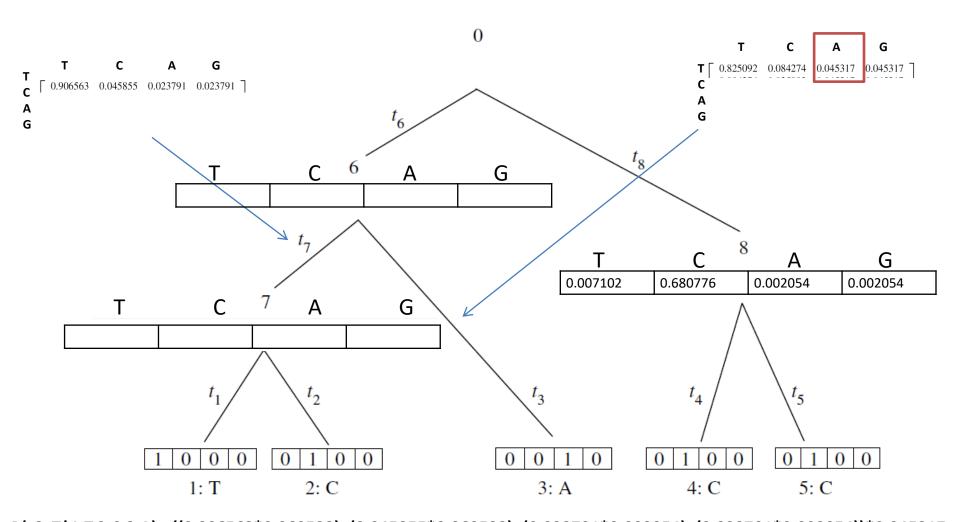




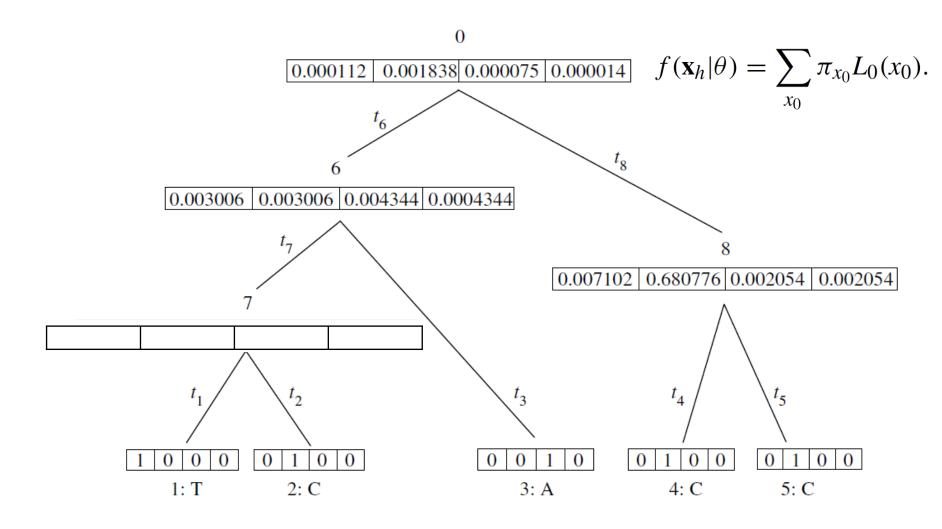




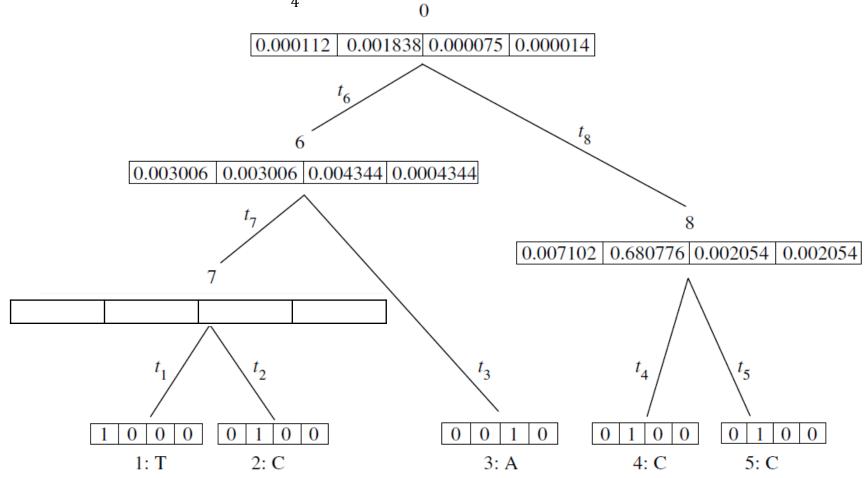




P(t6=T|1:T,2:C,3:A)= ((0.906563*0.069533)+(0.045855*0.069533)+(0.023791*0.002054)+(0.023791*0.002054))*0.045317



 $f(X_h|\theta) = \frac{1}{4}(0.000112 + 0.001838 + 0.000075 + 0.000014)$



And now... Imagine that we know the topology of the tree

 $P(D|\theta)$

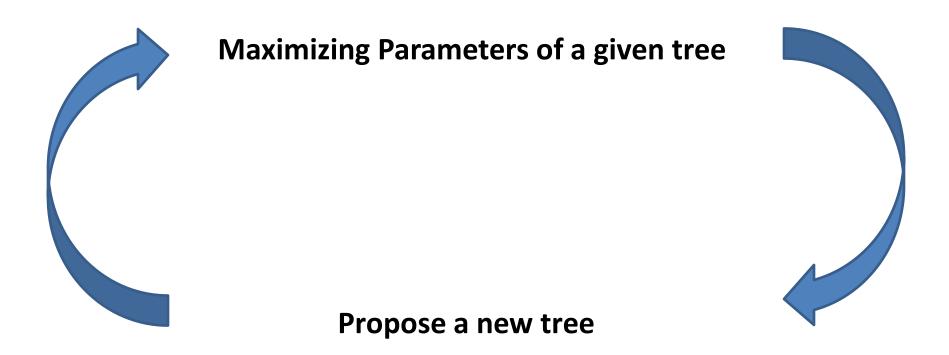
"What is the probability (likelihood) of observing the data given the values of the set of parameters"

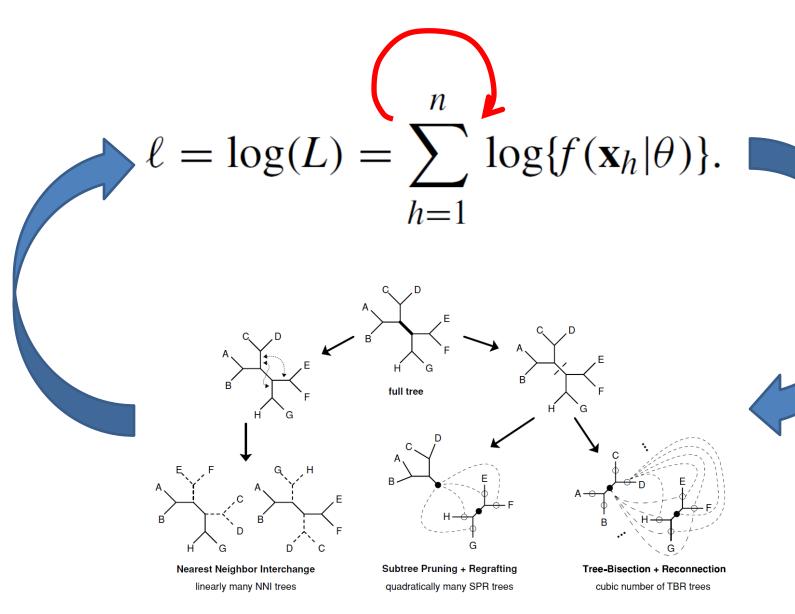
Which is the loglikelihood of all the positions given the length of the branches and my model of nucleotide evolution? n

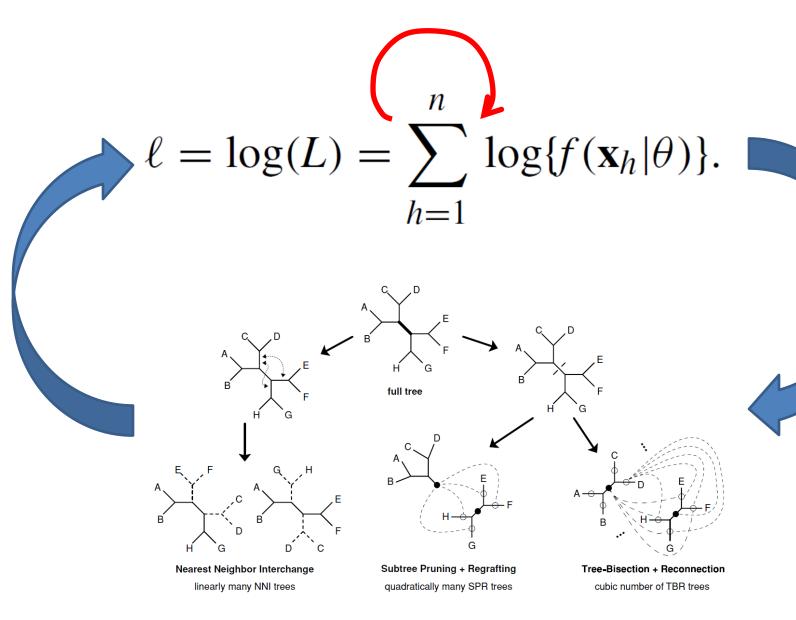
$$\ell = \log(L) = \sum_{h=1}^{\infty} \log\{f(\mathbf{x}_h|\theta)\}.$$

This is the parameter we want to maximize!

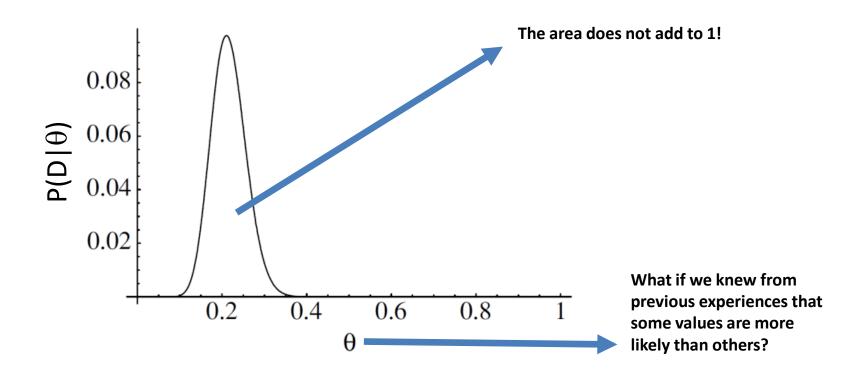
- Parameters
 - Branch length
 - Rates of substitution
- Topology
 - Relationship between OTUs



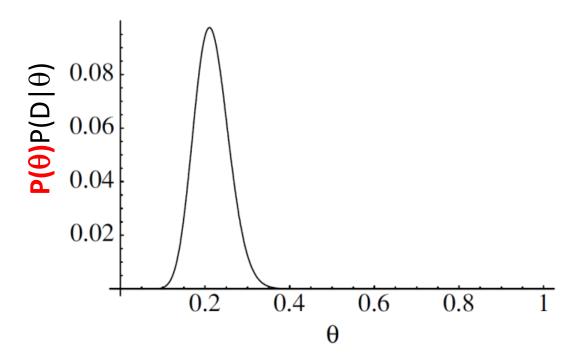




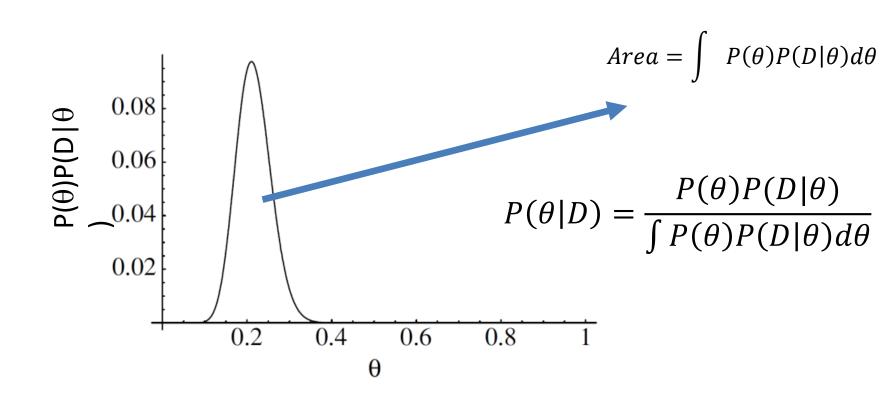
Which are the differences between maximum parsimony and maximum likelihood?



What if we knew from previous experiences that some values are more likely than others?



The area does not add to 1!



A formal way of expressing it...

$$P(D;\theta) = P(\theta)P(D|\theta) = P(D)P(\theta|D)$$
$$P(D) = \int P(\theta)P(D|\theta)d\theta$$

Prior knowledge

$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)}$$

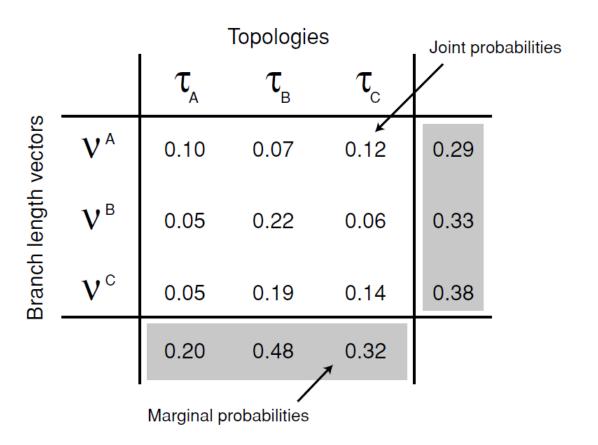
"our current knowledge is a mixture of what we previously knew updated with new data"

Posterior distribution of θ

Normalizing constant

$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{\int P(\theta)P(D|\theta)d\theta}$$
 "integrating over the marginals"

Bayesian methods Some extra definitions



Bayesian methods Remember: Bayesian statistics

$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)} = \frac{P(\theta)P(D|\theta)}{\int P(\theta)P(D|\theta)d\theta}$$

$$\int P(\theta)P(D|\theta)d\theta$$

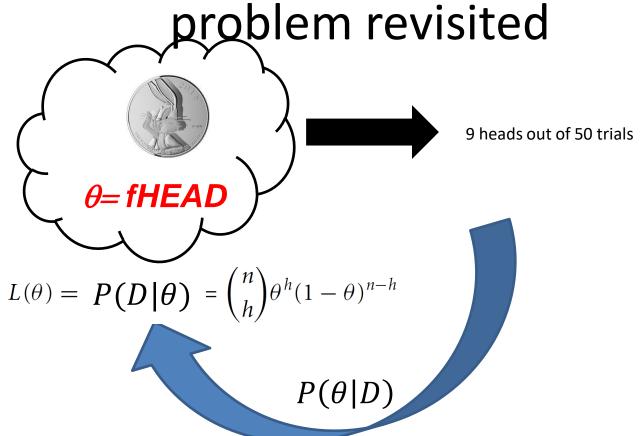
$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)} = \frac{P(\theta)P(D|\theta)}{\int P(\theta)P(D|\theta)d\theta}$$

$$\int P(\theta)P(D|\theta)d\theta$$

What if it does not exist?

Bayesian methods A continuous variable: The head

problem revisited



Bayesian methods How to get the Posterior?

Assume uniform prior knowledge ("the frequency can be any value between 0 and 1 with equal probability")

$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)}$$

$$P(y|\theta)$$

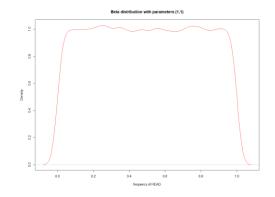
Binomial Distribution

$$P(y|\theta,n) = \frac{n!}{y!(n-y)!}\theta^y(1-\theta)^{n-y}$$

Beta distribution

$$P(\theta | \alpha, \beta) = \frac{\theta^{\alpha - 1} (1 - \theta)^{\beta - 1}}{\left(\frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)}\right)}$$

$$\alpha = 1$$
; $\beta = 1$



Bayesian methods Remember: Bayesian statistics

$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)} = K\theta^{y+1-1}(1-\theta)^{n-y+1-1}$$
Beta distribution
$$P(\theta|\alpha,\beta) = \frac{\theta^{\alpha-1}(1-\theta)^{\beta-1}}{\left(\frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}\right)}$$

"Beta distribution is a conjugate of the Binomial distribution"

Bayesian methods Remember: Bayesian statistics

$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)} = K\theta^{y+1-1}(1-\theta)^{n-y+1-1}$$
Beta distribution
$$P(\theta|\alpha,\beta) = \frac{\theta^{\alpha-1}(1-\theta)^{\beta-1}}{\left(\frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}\right)}$$

 $P(\theta|D) = B(y+1, n-y+1)$ $Mean(B(y+1, n-y+1)) = \frac{y+1}{n+2}$

Remember: Bayesian statistics

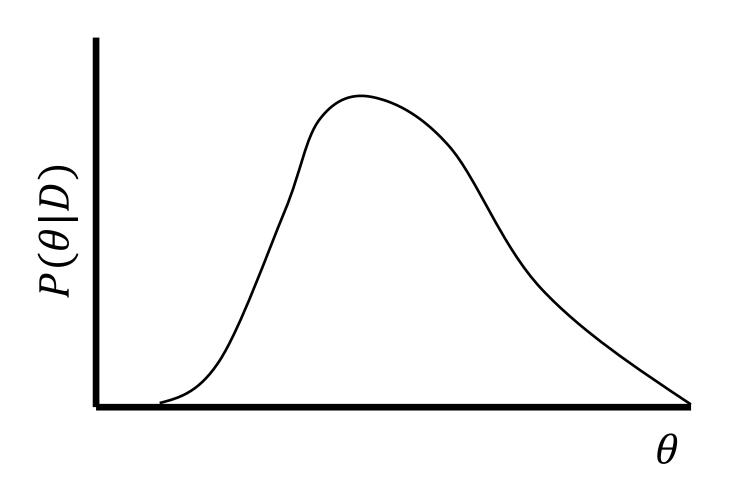
$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)} = \frac{P(\theta)P(D|\theta)}{\int P(\theta)P(D|\theta)d\theta}$$

$$\int P(\theta)P(D|\theta)d\theta$$
What if

there is no

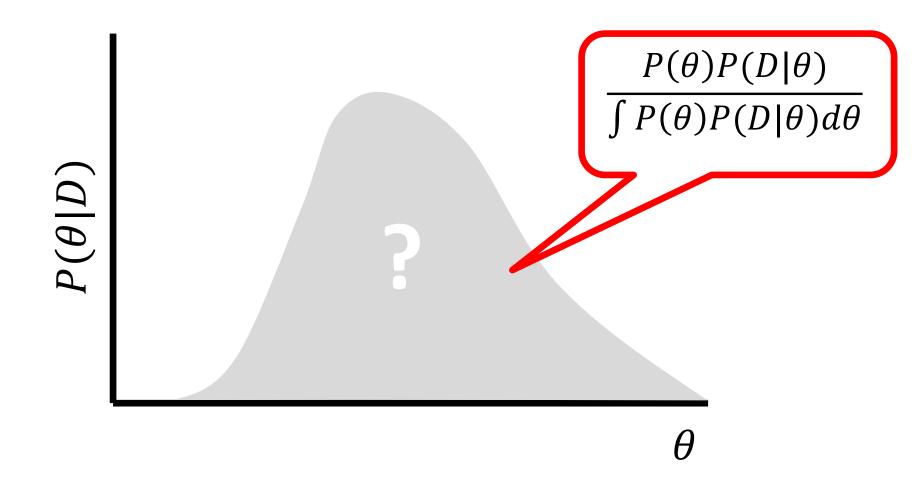
conjugate?

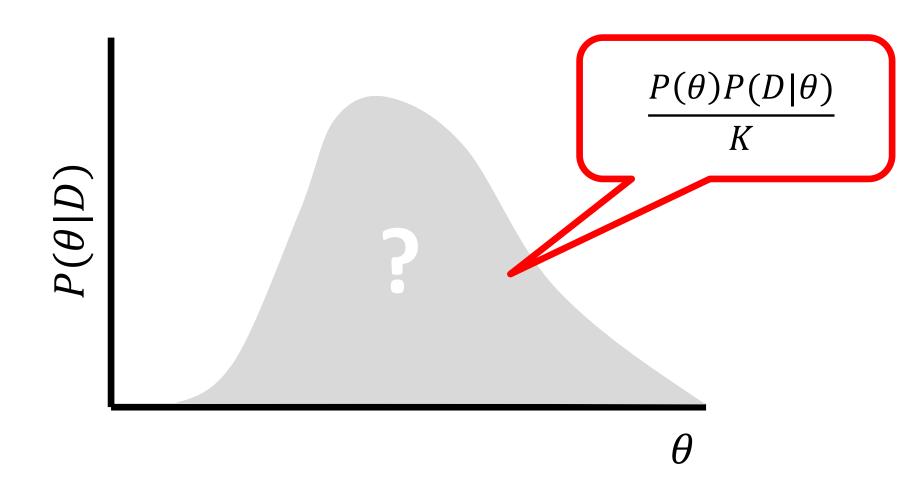
What if it does not exist?



At least, is there any way we can generate samples from this (unknown) posterior distribution so we can broadly regenerate the (unknown) posterior distribution?



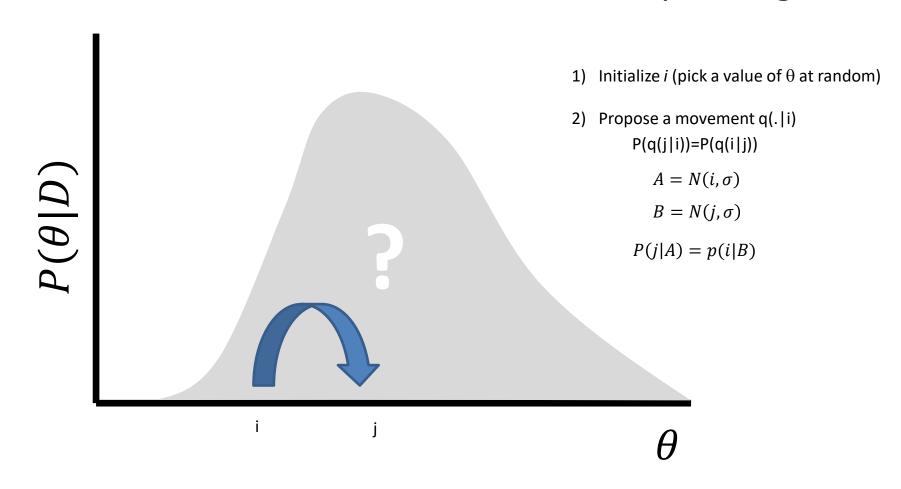


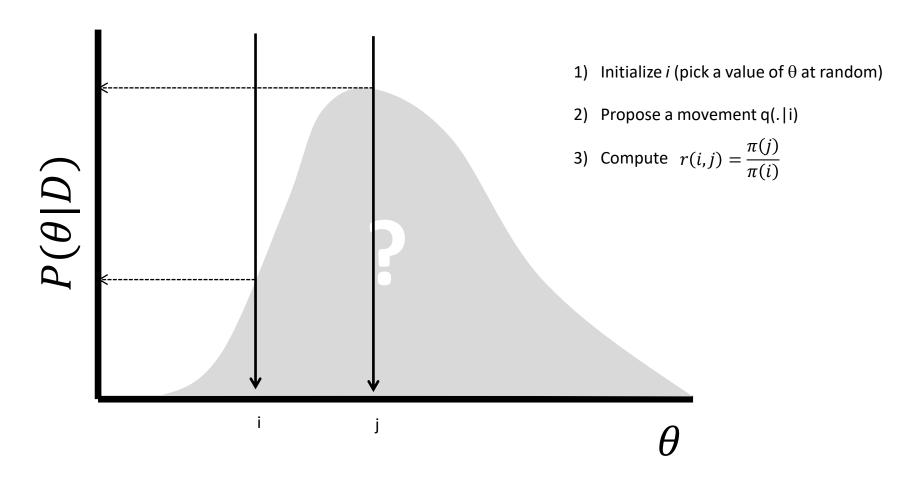


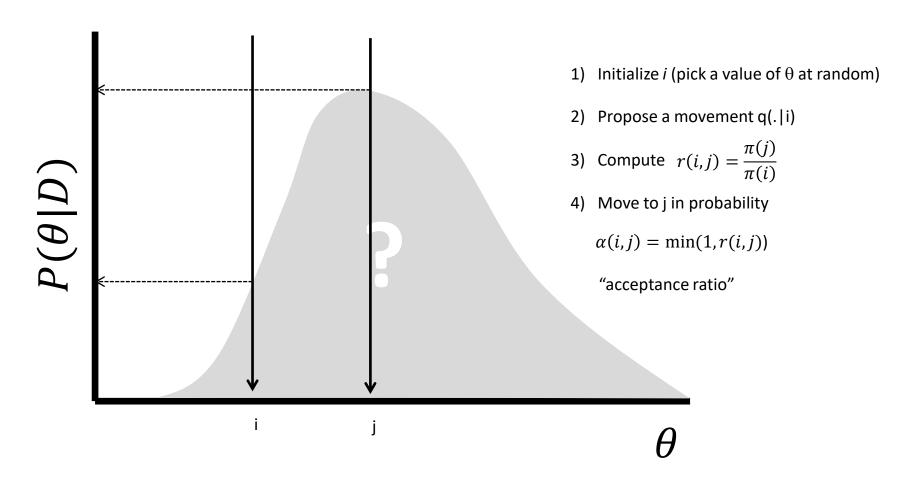
How many times is more likely *j* than *i* in this (unknown) distribution?

$$\mathbf{r}(i,j) = \frac{\frac{P(\theta = i)P(D|\theta = i)}{K}}{\frac{P(\theta = j)P(D|\theta = j)}{K}} = \frac{P(\theta = i)P(D|\theta = i)}{P(\theta = j)P(D|\theta = j)} = \frac{\pi(j)}{\pi(i)}$$

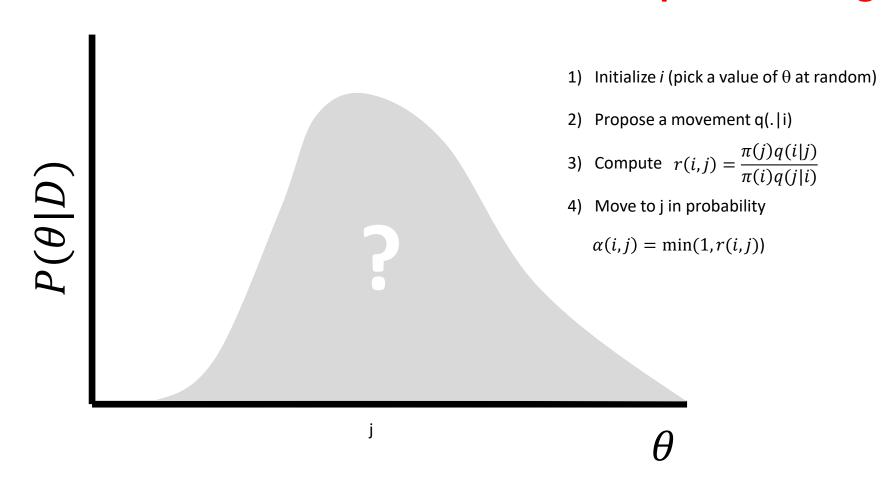




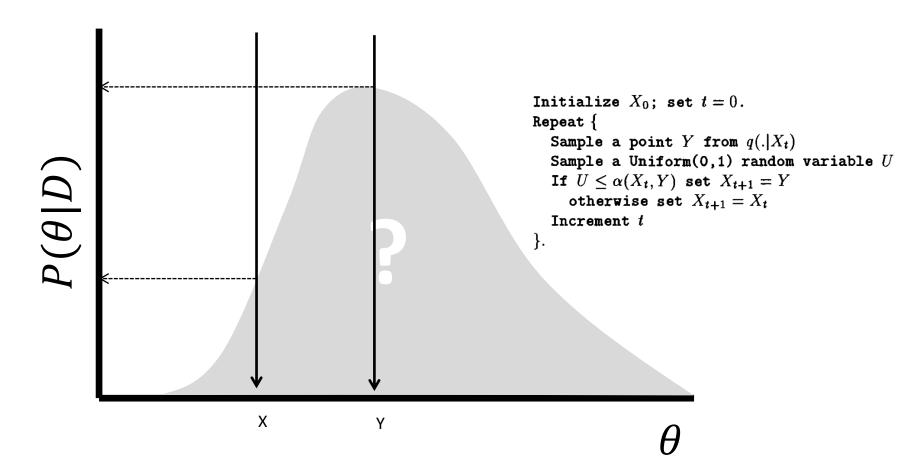




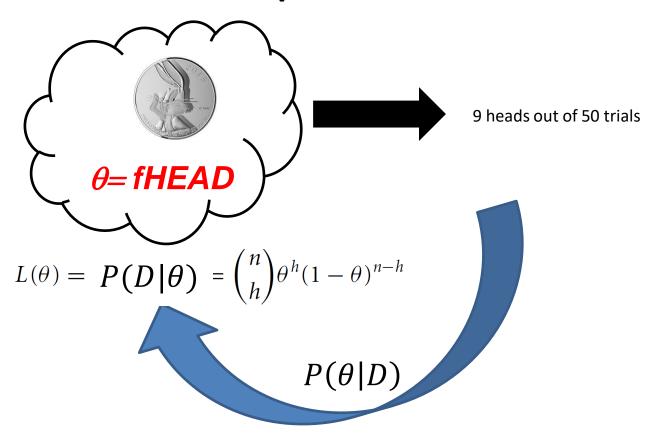
Markov Chain Monte Carlo Methods: Metropolis-Hastings



Markov Chain Monte Carlo Methods: Metropolis-Hastings



Bayesian methods The head problem revisited



$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)} = \frac{U(0,1)Binom(h|\theta,n)}{?}$$

$$Binom(h|\theta,n) = \frac{n!}{h!(n-h)!}\theta^h(1-\theta)^{n-h}$$

$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)} = \frac{U(0,1)Binom(h|\theta,n)}{?}$$

$$Binom(h|\theta,n) = \frac{n!}{h!(n-h)!}\theta^h(1-\theta)^{n-h}$$
 $h = 9;$
 $n = 50$

```
Initialize X_0; set t=0.

Repeat {

q(.|X_t) = N(X_t, 1)

Sample a point Y from q(.|X_t)

Sample a Uniform(0,1) random variable U

If U \leq \alpha(X_t, Y) set X_{t+1} = Y

otherwise set X_{t+1} = X_t

Increment t

}.
```

$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)} = \frac{U(0,1)Binom(h|\theta,n)}{?}$$

$$Binom(h|\theta,n) = \frac{n!}{h!(n-h)!}\theta^h(1-\theta)^{n-h}$$
 $h = 9;$
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```
Initialize X_0; set t = 0.

q(.|X_t) = N(X_t, 1)

Repeat {

Sample a point Y from q(.|X_t)

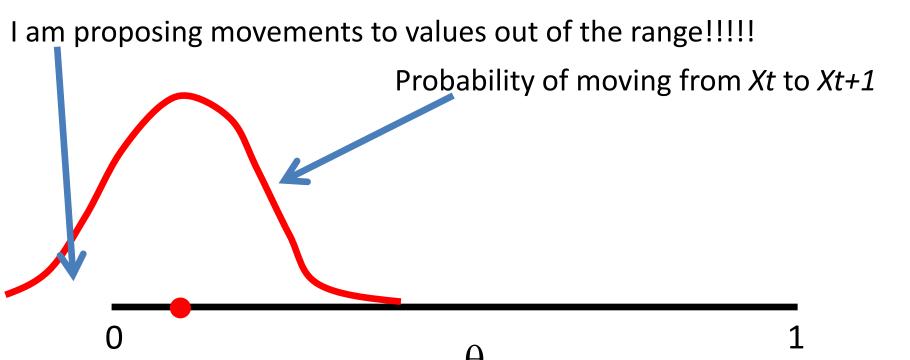
Sample a Uniform(0,1) random variable U

If U \leq \alpha(X_t, Y) set X_{t+1} = Y

otherwise set X_{t+1} = X_t

Increment t
```

Bayesian methods A potential problem with this variable definition...



$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)} = \frac{U(0,1)Binom(h|\theta,n)}{?}$$

$$Binom(h|\theta,n) = \frac{n!}{h! (n-h)!} \theta^{h} (1-\theta)^{n-h} \qquad h = 9;$$

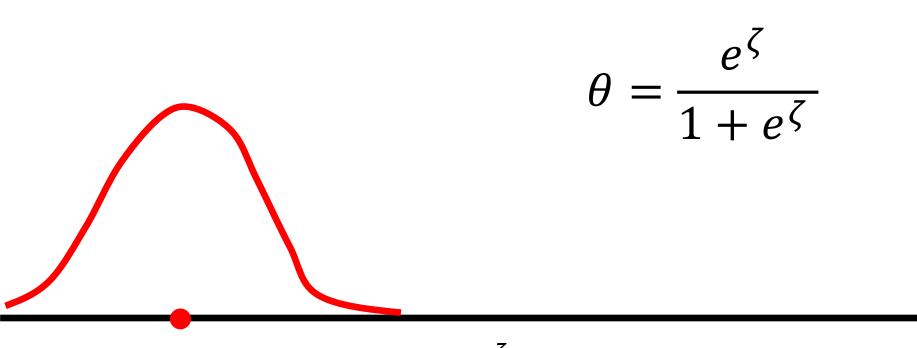
$$n = 50$$

$$q(.|X_{t}) = N(X_{t},1)$$

$$Sample a point Y from q(.|X_{t})$$

$$Sample a Uniform(0,1) random variable U for Use Q(X_{t},Y) set X_{t+1} = Y for the twise set X_{t+1} = X_{t} fo$$

Bayesian methods Now there is no problem anymore!



$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)} = \frac{U(0,1)Binom(h|\theta,n)}{?}$$

$$Binom(h|\theta,n) = \frac{n!}{h!(n-h)!} \theta^h (1-\theta)^{n-h}$$
 $h = 9;$
 $n = 50$

```
Initialize X_0; set t = 0. \zeta_0 = 0; \theta_0 = 0.5; P(9|0.5,50) = 2.2e-06 Repeat {

Sample a point Y from q(.|X_t) Y = -0.46; Y_\theta = 0.38; P(9|0.38,50) = 0.0012 Sample a Uniform(0,1) random variable U If U \le \alpha(X_t, Y) set X_{t+1} = Y_t otherwise set X_{t+1} = X_t Increment t }
```

$$P(\theta|D) = \frac{P(\theta)P(D|\theta)}{P(D)} = \frac{U(0,1)Binom(h|\theta,n)}{?}$$

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```

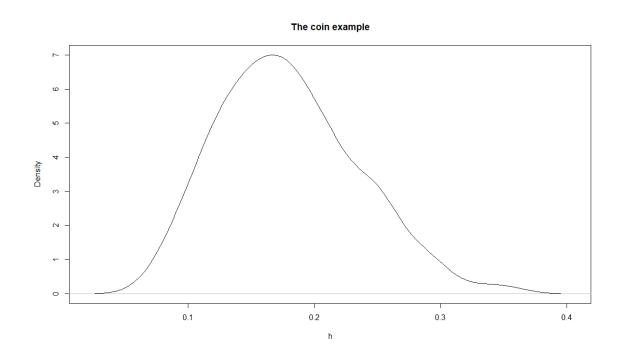
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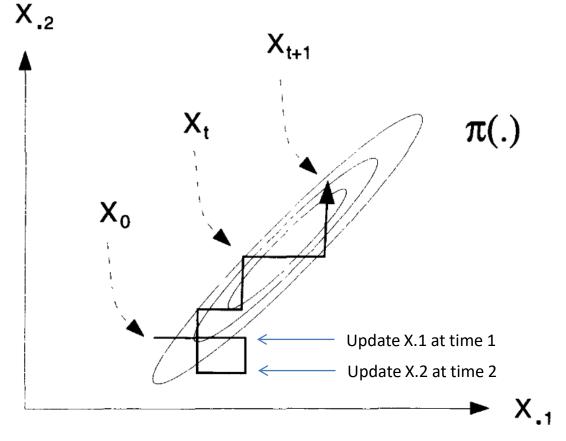
Sample a point Y from q(.|X_t) Y= -0.46; Y_\theta=0.38; P(9|0.38,50)=0.0012 Sample a Uniform(0,1) random variable U U=0.9661937 If U \le \alpha(X_t,Y) set X_{t+1}=Y otherwise set X_{t+1}=X_t \alpha(\theta_0,Y_\theta)=\frac{0.0012}{2.2e-06}=572.28 Increment t \zeta_1=-0.46; \theta_1=0.38; P(9|0.38,50)=0.0012
```

Posterior distribution found by Metropolis

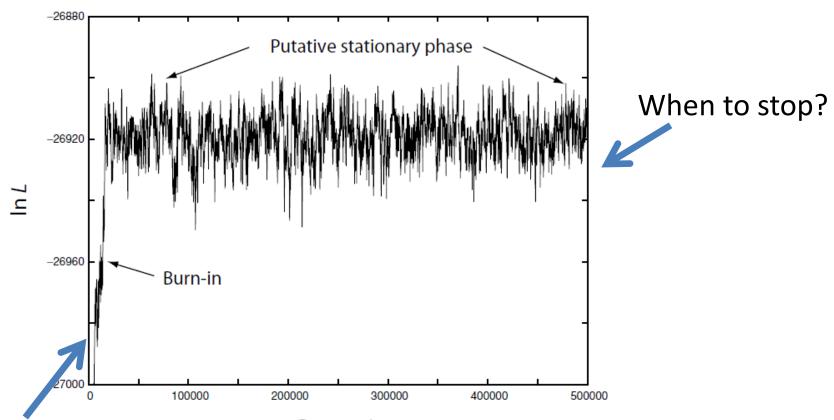


MCMC: Metropolis-Hastings single component update

What happens if we have more than one parameter to estimate?

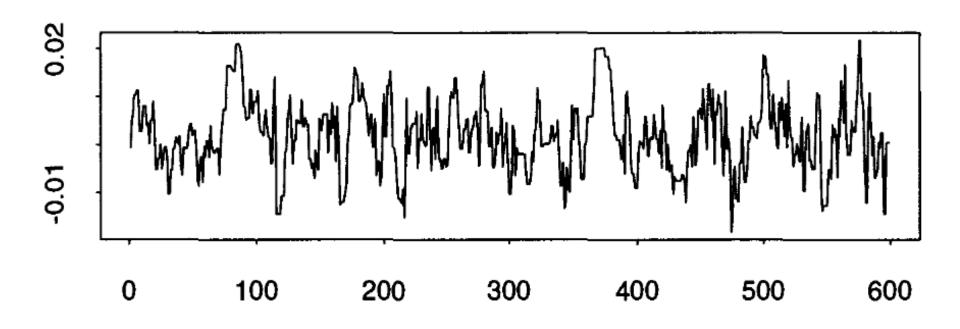


Bayesian methods MCMC in practice

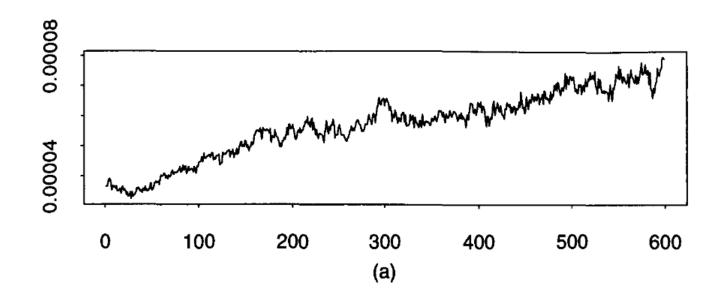


Usually we start out of the most likely values of the posterior

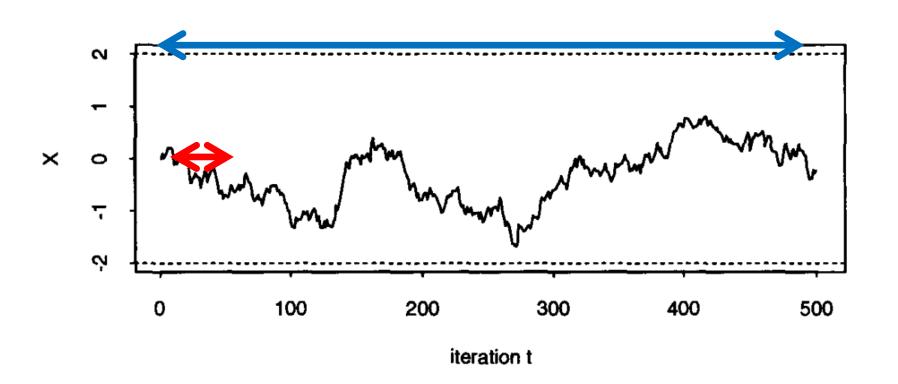
Bayesian methods MCMC in practice: good mixing

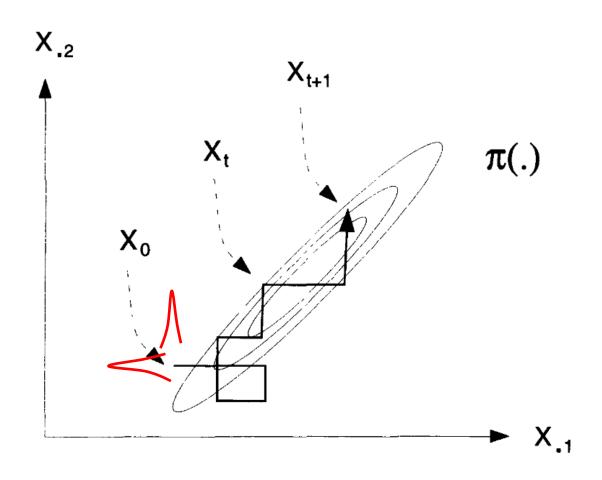


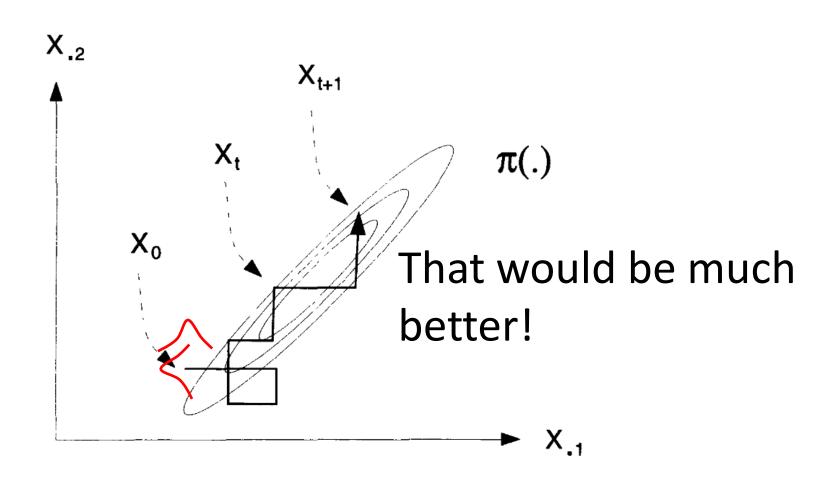
Bayesian methods MCMC in practice: bad mixing



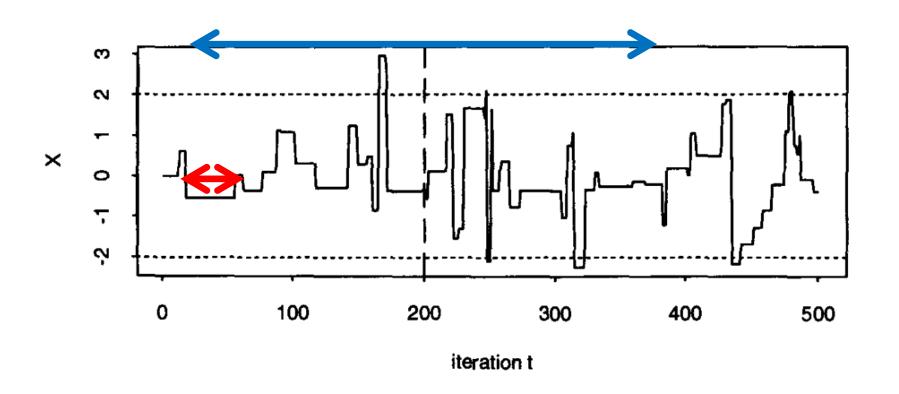
Bayesian methods MCMC in practice: bad mixing

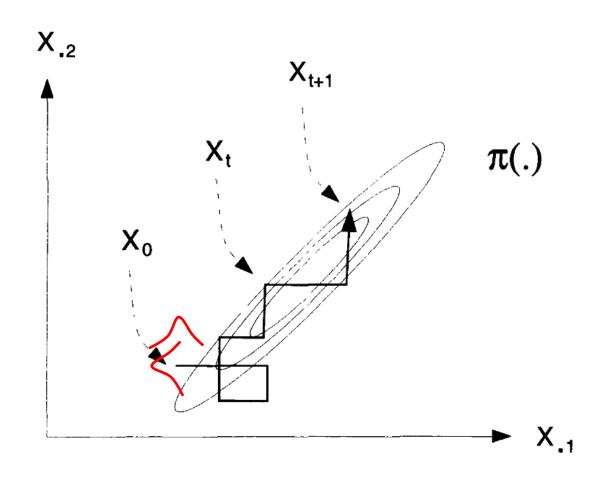


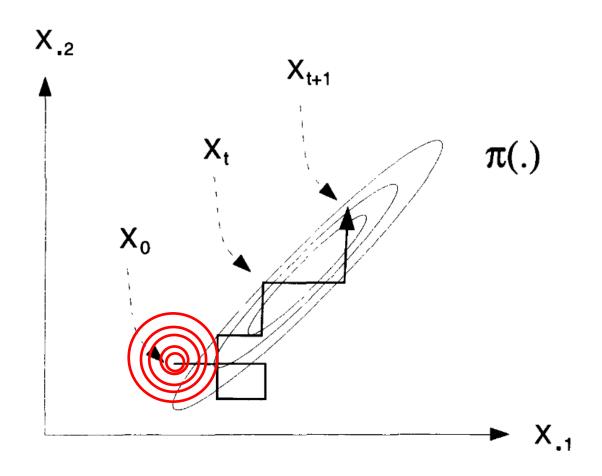


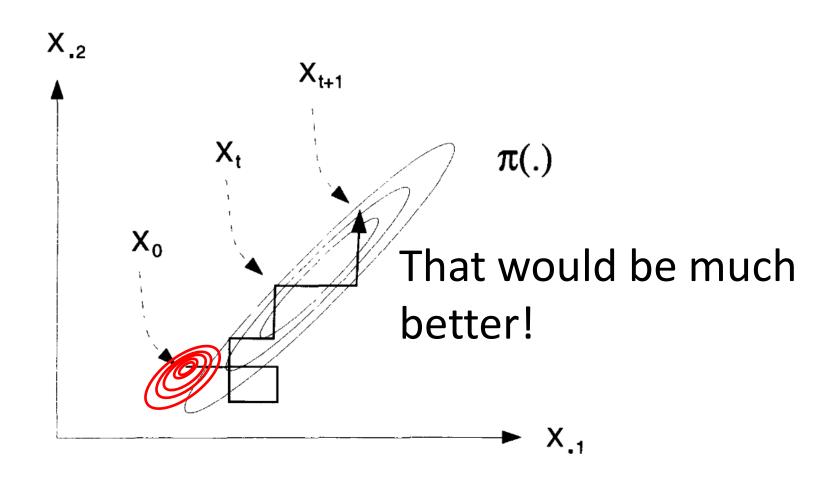


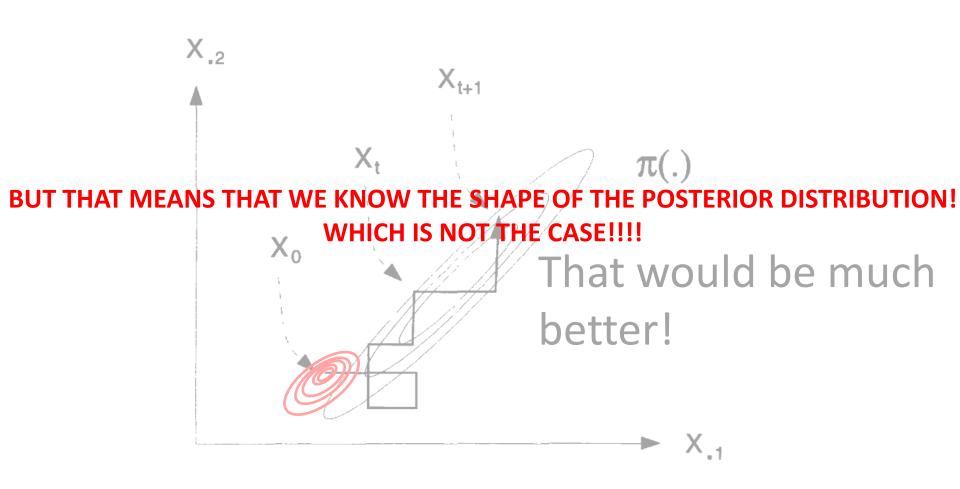
Bayesian methods MCMC in practice: bad mixing





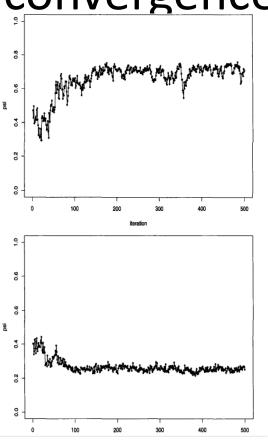




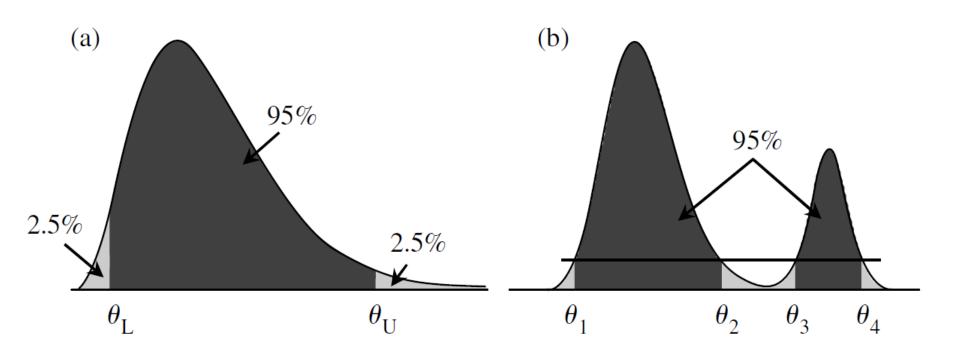


Bayesian methods MCMC in practice: Lack of

convergence



Bayesian methods Yet another issue: multimodality



Bayesian methods MCMC: Be careful



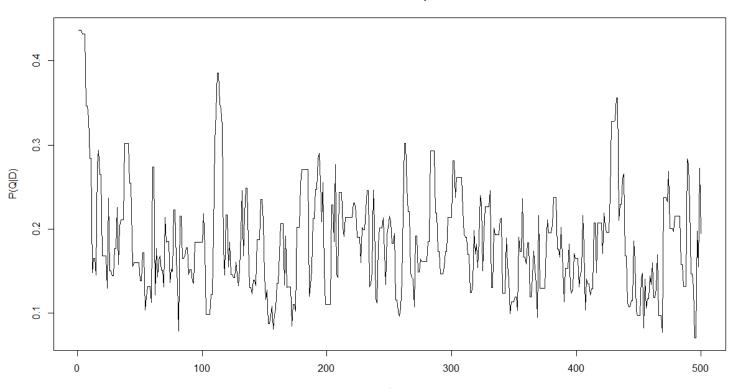
Bayesian methods MCMC in practice

- The final GOAL of MCMC is to retrieve samples from the posterior distribution without having a close form of this distribution
- However, we start at a random point of the parameter space, which probably has a low probability given the data.
- We want independent samples from the posterior distribution. However, we are doing walks that depends from previous sampled values...

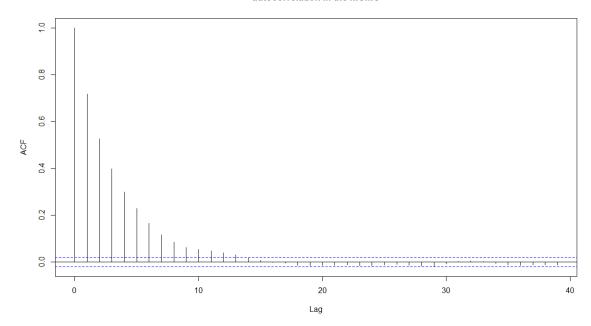
Bayesian methods MCMC: Some hints

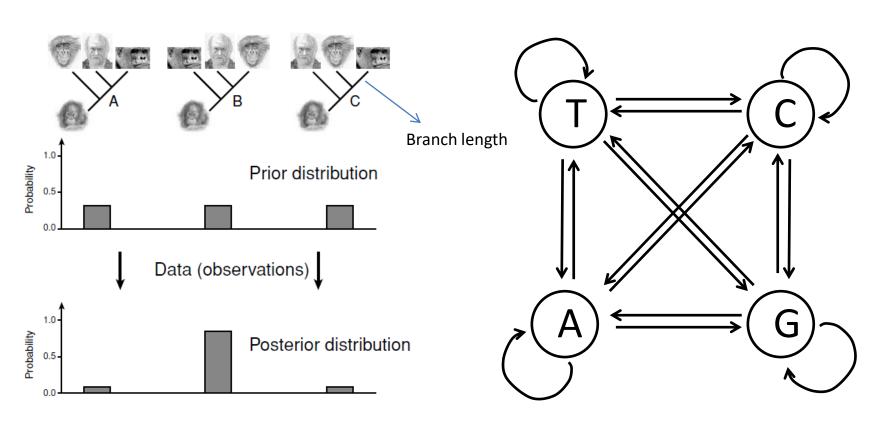
- Define burn-in
- Run more than one MCMC chain and check convergence of the loglikelihood.
- Compute metrics of autocorrelation between t and t+k
 - Is good the mixing?
 - Define lag between accepted runs

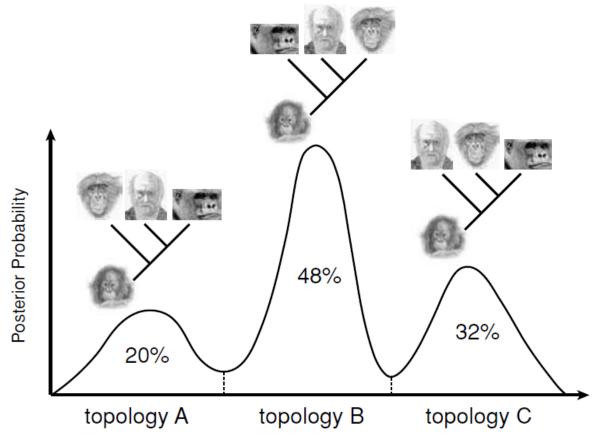
The coin example



autocorrelation in the MCMC

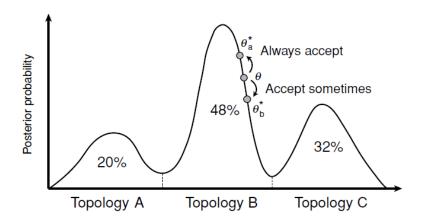






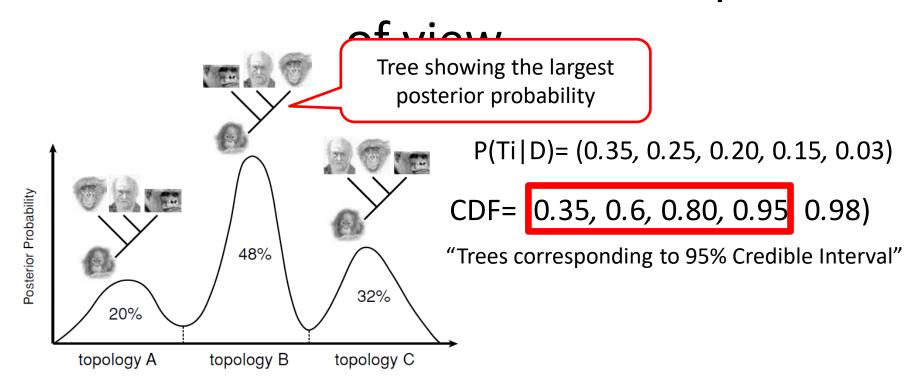
Markov chain Monte Carlo steps

- 1. Start at an arbitrary point (θ)
- 2. Make a small random move (to θ^*)
- 3. Calculate height ratio (r) of new state (to θ^*) to old state (θ)
 - (a) r > 1: new state accepted
 - (b) r < 1: new state accepted with probability rif new state rejected, stay in old state
- 4. Go to step 2

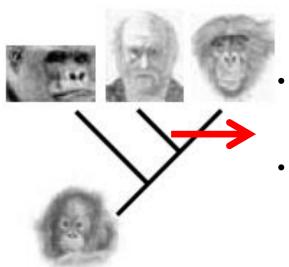


- 1. Start with a random tree τ , with random branch lengths **b**, and random substitution parameters θ .
- 2. In each iteration do the following:
 - a. Propose a change to the tree, by using tree rearrangement algorithms (such as NNI or SPR). This step may change branch lengths **b** as well.
 - b. Propose changes to branch lengths **b**.
 - c. Propose changes to parameters θ .
 - d. Every k iterations, sample the chain: save τ , **b**, θ to disk.
- 3. At the end of the run, summarize the results.

Bayesian phylogenetics: how to summarize results? From a tree point



Bayesian phylogenetics: how to summarize results? From a branch point of view



Branch length

- Mean, median, mode
- 95% Credible interval
- How often this branch is retrieved?