

# Spatial modeling with INLA



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The repo for this workshop is [here](https://tinyurl.com/INLAUP24)  
<https://tinyurl.com/INLAUP24>



## 1 Spatial domains

## 2 Areal modeling

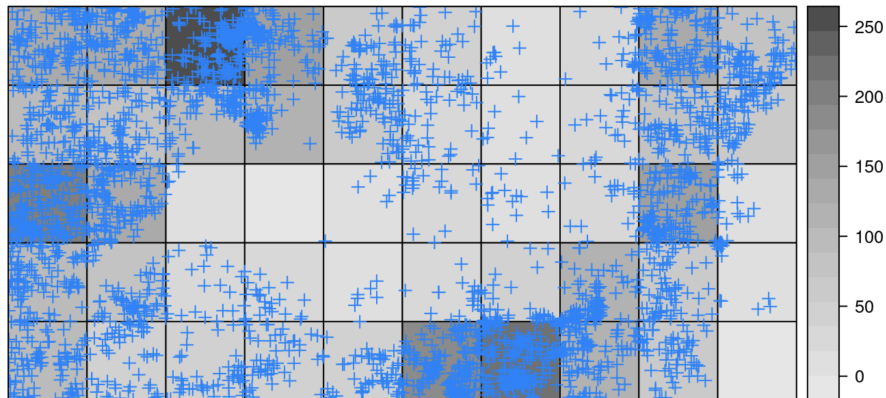
- Besag and BYM
- Non-stationary Besag model
- Spatio-temporal models

## 3 Geostatistics

- Kriging
- Matérn field and extensions
- SPDE approach
- Barrier model - non-stationary Matern field
- Spatially varying coefficient models
- Spatio-temporal models

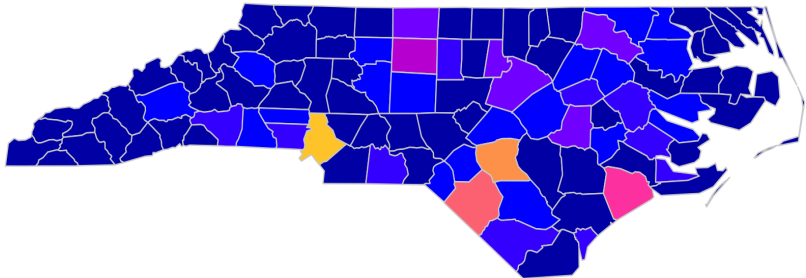


# Lattice type - Day 2





# Irregular lattice - areal data - Day 2



10

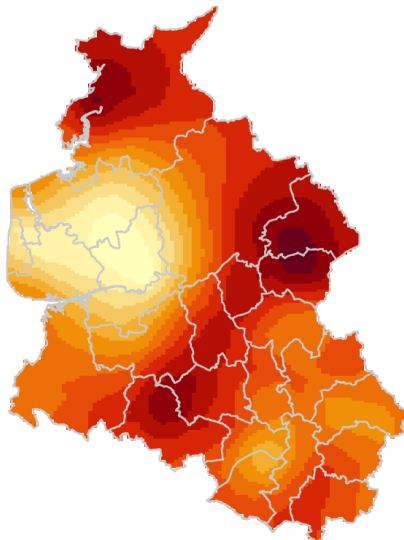
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30

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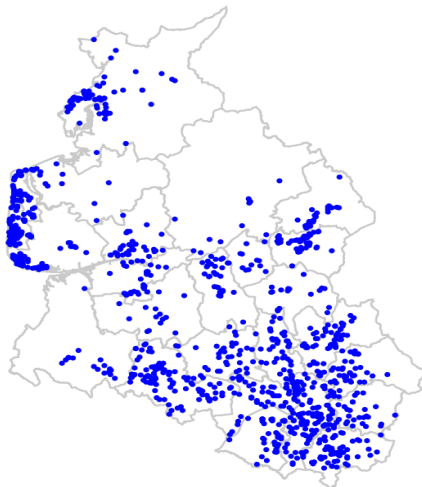


# Continuous domain - geostatistics - Day 2





# Point process - Day 3



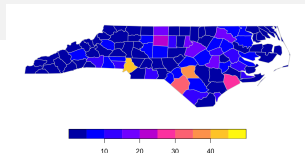


# Universal tools in INLA - also for spatial models

- Model selection metrics - WAIC, DIC
- Cross validation (1 and group) and model-based clustering
- Prediction of unobserved areas or new profiles
- Mean or quantile models
- Joint models
- Multiple imputation
- Coregionalization models
- etc..... ask at <https://groups.google.com/g/r-inla-discussion-group?pli=1> or e-mail [help@r-inla.org](mailto:help@r-inla.org)



# Besag and BYM



Besag model is a "smoother" over space.

$$x_i | \mathbf{x}_{-i} \sim N \left( \frac{1}{n_i} \sum_{j \sim i} x_j, \frac{1}{n_i \tau} \right)$$

BYM (Besag + iid) parameterized for interpretable parameters.

$$x_i = \frac{1}{\sqrt{\tau}} \left( \sqrt{\phi} u_i + \sqrt{1 - \phi} v_i \right).$$





# North Carolina SIDS example

NC sids R markdown file - Example 1 and 2



# Malaria and G6PD example on joint spatial modeling

We can do joint modeling and quantile models as well.<sup>1</sup>

[Malaria and G6PD R markdown file - Example 3 and 4](#)

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<sup>1</sup>Alahmadi, H., Van Niekerk, J., Padellini, T. and Rue, H., 2024. Joint quantile disease mapping with application to malaria and G6PD deficiency. Royal Society Open Science, 11(1), p.230851.



# Flexible Besag model<sup>2</sup>

Instead of one precision for the entire area, we define multiple precision parameters,  $\tau_1, \tau_2, \dots, \tau_P$ , to account for covariance non-stationarity. The conditional density for the spatial effect of area  $i$  is

$$x_i | \mathbf{x}_{-i}, \tau_1, \dots, \tau_P \sim N \left( \frac{1}{2} \sum_{\substack{i \text{ in sub-region } k \\ j \text{ in sub-region } l \\ i \sim j}} (\tau_l + \tau_k) \tau_{x_i}^{-1} x_j, \tau_{x_i}^{-1} \right),$$

and

$$\tau_{x_i} = \frac{1}{2} \left( n_i \tau_k + \sum_l n_{il} \tau_l \right).$$

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<sup>2</sup>Abdul-Fattah, E., Krainski, E., Van Niekerk, J. and Rue, H., 2024. Non-stationary Bayesian spatial model for disease mapping based on sub-regions. Statistical Methods in Medical Research, p.09622802241244613.



# Contraction prior: Non-stationary $\rightarrow$ stationary

The joint PC prior for  $\boldsymbol{\theta} = \log \boldsymbol{\tau}$  can be derived as a convolution of the PC prior for  $\boldsymbol{\tau}$  from the Besag model, as follows

$$\pi(\boldsymbol{\theta}) = 2^{-(P+2)/2} \pi^{-P/2} \lambda \sigma^{-P} \exp \left( -\frac{1}{2} (\boldsymbol{\theta} - \mathbf{1}\bar{\theta})^T \tilde{\boldsymbol{\Sigma}}^{-1} (\boldsymbol{\theta} - \bar{\theta}\mathbf{1}) - \bar{\theta}/2 - \lambda e^{-\bar{\theta}/2} \right),$$

This prior contracts

$$\tau_1, \tau_2, \dots, \tau_P \quad \rightarrow \quad \boldsymbol{\tau}$$



# Dengue risk in Brazil

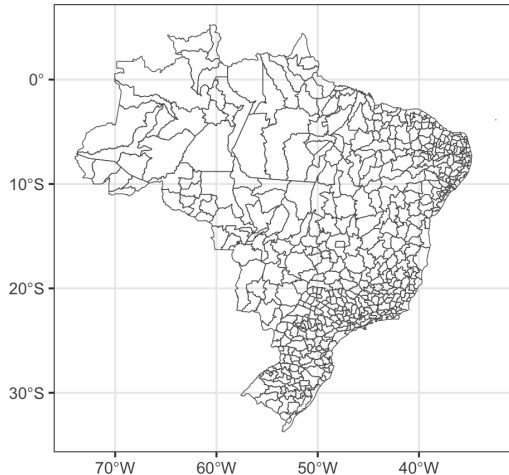
We analyze the effects of hydrometeorological hazards on dengue risk in Brazil. To test the spatial variations in the spread of the virus in different sub-regions of Brazil, we fit dengue counts with a Poisson regression model as follows,

$$\mathbf{y} \sim \text{Poisson}(Ee^{\boldsymbol{\eta}}), \quad \boldsymbol{\eta} = \mathbf{1}^T \mu + \boldsymbol{\alpha}$$

where  $\mathbf{y}$  is the observed counts in November of dengue cases,  $E$  is the expected number of counts,  $\boldsymbol{\eta}$  is the linear predictor,  $\mu$  is the overall intercept, and  $\boldsymbol{\alpha}$  is the Besag or flexible Besag model over space.

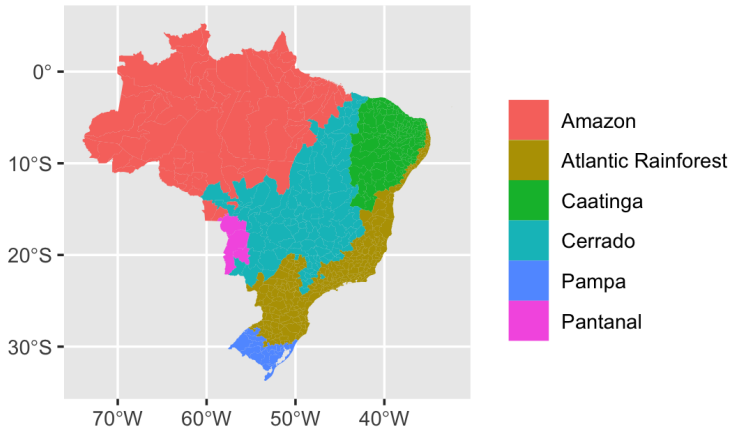


# Dengue risk in Brazil



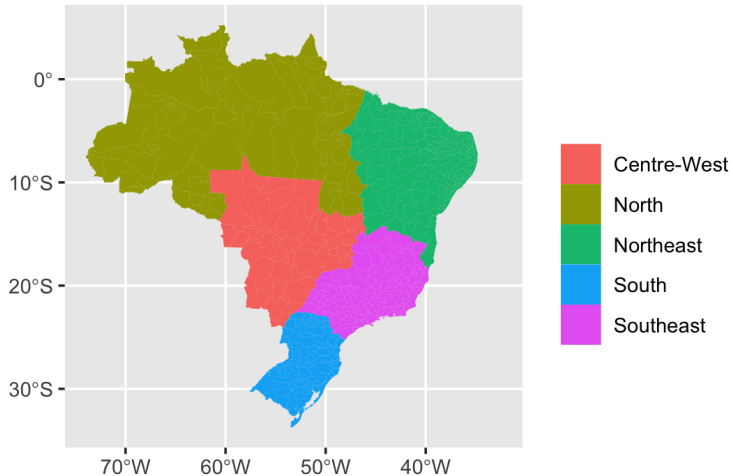


# Dengue risk in Brazil





# Dengue risk in Brazil





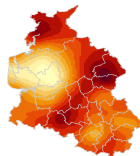


# Dengue risk in Brazil

Dengue risk in Brazil R markdown file - Example 5

Each of these spatial models can be used to form a space-time interaction. In the areal modeling framework this is done with the Kronecker product of the precision matrices of the spatial and temporal components.

NOTE: These models need a lot of constraints to ensure identifiability. Details for current models are here: Knorr-Held models



Kriging provides conditional expectations of the spatial field based on covariance parameters.

With INLA we estimate "covariance" parameters in a Bayesian way and provide the marginal expectation of the spatial field.

So INLA also does "Kriging" and more - Kriging is a model, not a method.



# Matérn field

Consider a set of locations  $\mathbf{s}$ , then the spatial field  $\mathbf{u}$  defined at  $\mathbf{s}$  is multivariate Gaussian with the Matérn covariance function for the elements of  $\Sigma(\theta)$ ,

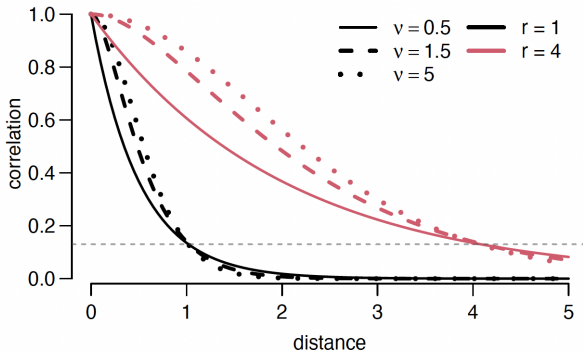
$$\pi(\mathbf{u}|\theta) = (2\pi)^{-n/2} |\Sigma(\theta)|^{-1/2} \exp\left(-\frac{1}{2} \mathbf{u}^T \Sigma(\theta)^{-1} \mathbf{u}\right)$$



# Matérn covariance model

$$\text{Matérn}(1960): \Sigma_{ij} = \frac{\sigma^2(\kappa \|\mathbf{s}_i - \mathbf{s}_j\|)^\nu K_\nu(\kappa \|\mathbf{s}_i - \mathbf{s}_j\|)}{\Gamma(\nu + d/2)(4\pi)^{d/2} \kappa^{2\nu} 2^{\nu-1}}$$

If  $d = 2$  and  $\nu = 1$ : Whittle (1954)



practical range =  $r = \sqrt{8\nu}/\kappa$ ,  $\text{corr}(r) \approx 0.13$



# The Matérn's SPDE

- Whittle (1954), Whittle (1963):
  - Fields with Matérn covariance are solutions to the following Stochastic Partial Differential Equation (SPDE)

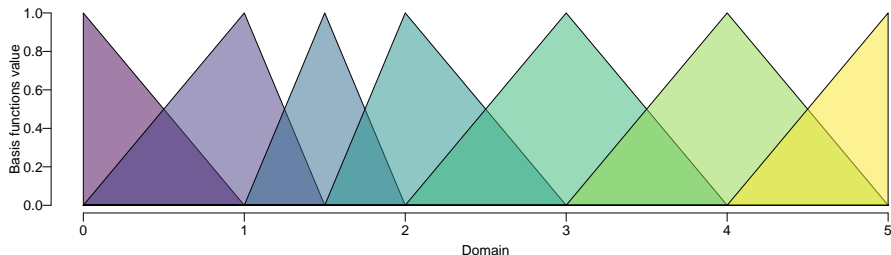
$$\tau(\kappa^2 - \Delta)^{\alpha/2} u(\mathbf{s}) = \mathcal{W}(\mathbf{s})$$

- $\kappa > 0$ : scale parameter
- $\alpha = \nu + d/2$ : smoothness
- $\Delta$  is the Laplacian

$$\Delta = \sum_{i=1}^d \frac{\partial^2}{\partial s_i^2}$$



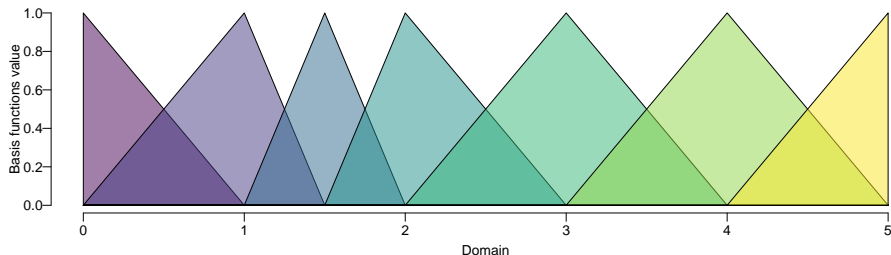
# How to solve the SPDE? FEM



- $u(\mathbf{s}) \approx \sum_{k=1}^m \psi_k(\mathbf{s})u_k = \mathbf{A}(\mathbf{s}, \mathbf{s}_0)u(\mathbf{s}_0),$ 
  - $\psi_k$ : basis functions evaluated at data locations  $\mathbf{s}$
  - $u_k$ : the process at the discretization points  $\mathbf{s}_0$



# How to solve the SPDE? FEM



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# Lindgren, Rue, and Lindström (2011)<sup>3</sup> I

- Discretization

- sparse precision matrix:
- $\mathbf{Q}_\alpha(\tau, \kappa)$ , for  $\alpha \in \{1, 2, \dots\}$ .

- $\alpha$

- $\alpha = 1$ :  $\tau \mathbf{K}_1(\kappa) = \tau(\kappa^2 \mathbf{C} + \mathbf{G})$
- $\alpha = 2$ :  $\tau(\kappa^4 \mathbf{C} + 2\kappa^2 \mathbf{G} + \mathbf{G} \mathbf{C}^{-1} \mathbf{G})$
- $\alpha = 2, 3, 4, \dots$ :  $\tau \mathbf{K}_1(\kappa) \mathbf{C}^{-1} \mathbf{K}_{\alpha-2}(\kappa) \mathbf{C}^{-1} \mathbf{K}_1(\kappa)$

Equivalent models (discretized): Whittle (1954), Besag (1974), Besag (1981), Besag and Kooperberg (1995) and Besag and Mondal (2005).

- $\alpha = 1$ :  $\tau \mathbf{K}_1(\kappa) = \tau(\kappa^2 \mathbf{C} + \mathbf{G})$



# Lindgren, Rue, and Lindström (2011)<sup>4</sup> II



- $d=1, u_1, u_2, \dots, u_n$ , two neighbours

$$\tau^2 \begin{bmatrix} 1 + \kappa^2 & -1 & & & \\ -1 & 2 + \kappa^2 & -1 & & \\ & & \ddots & & \\ & & & -1 & 2 + \kappa^2 & -1 \\ & & & & -1 & 1 + \kappa^2 \end{bmatrix}$$

- $d = 2, \mathbf{C} = \mathbf{I}, \mathbf{G} = \text{Laplacian (4 neighbours)}$

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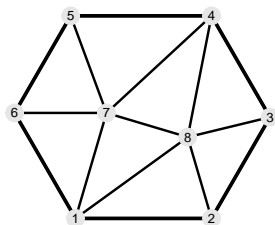
<sup>3</sup>Lindgren, F., Rue, H. and Lindström, J., 2011. An explicit link between Gaussian fields and Gaussian Markov random fields: the stochastic partial differential equation approach. *Journal of the Royal Statistical Society Series B: Statistical Methodology*, 73(4), pp.423-498.

<sup>4</sup>Lindgren, F., Rue, H. and Lindström, J., 2011. An explicit link between Gaussian fields and Gaussian Markov random fields: the stochastic partial differential equation approach. *Journal of the Royal Statistical Society Series B: Statistical Methodology*, 73(4), pp.423-498.  



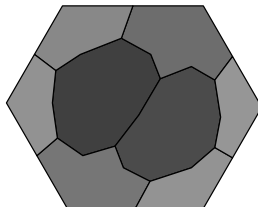
# Piecewise linear basis, FEM matrices

Mesh nodes



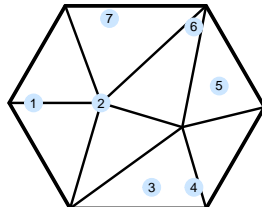
C

Dual mesh

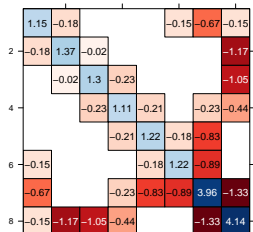
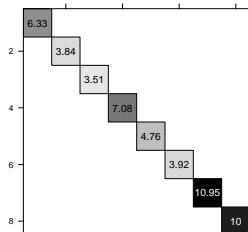


G

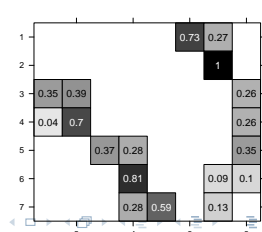
Data locations



A

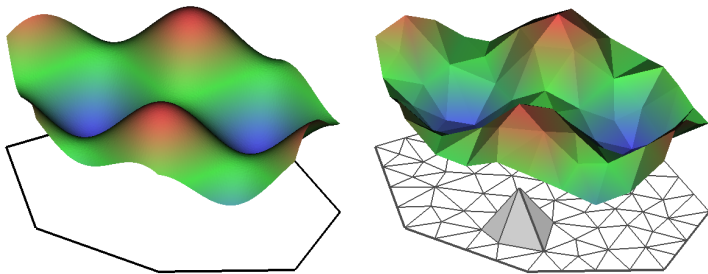


Space - INLA





# FEM in 3D





# Geostatistical survival analysis

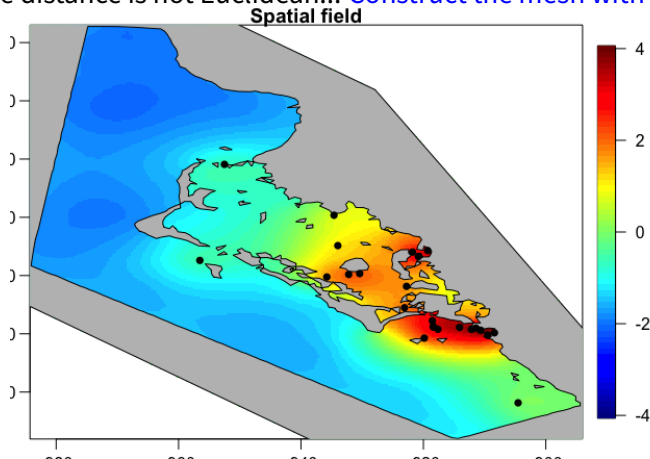
In this example we are studying the spatial distribution of leukemia mortality to inform public health policies, to gain insights for unmeasured covariates.

Leukemia mortality example in R - Example 1



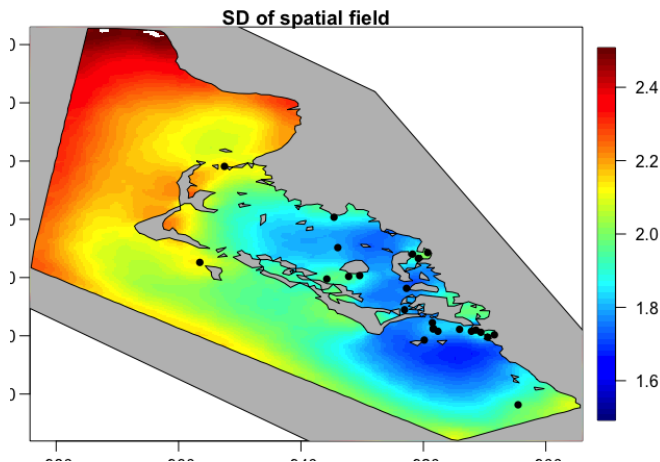
# Non-stationary Matern field based on physical barriers

Now the distance is not Euclidean... Construct the mesh with boundaries





# Non-stationary Matern field based on physical barriers





# Spatially varying coefficient models

We will use inlabru here -SVC using inlabru

$$\eta_{st} = \kappa_s + \alpha_s + \epsilon_s X_{1,st} + \tau_s X_{2,st} \quad (1)$$





# Spatio-temporal models

Rainfall example in Parana state with a hurdle model. In space we have a Matern model and in time we use an AR(1) model. So this is a seperable model. For non-seperable models see [here](#)<sup>5</sup>.

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<sup>5</sup>Lindgren, F., Bakka, H., Bolin, D., Krainski, E. and Rue, H., 2024. A diffusion-based spatio-temporal extension of Gaussian Matérn fields. SORT-Statistics and Operations Research Transactions, pp.3-66.



Gaedke-Merzhäuser, L., van Niekerk, J., Schenk, O., and Rue, H. (2023).  
Parallelized integrated nested Laplace approximations for fast  
Bayesian inference.

*Statistics and Computing*, 33(1):25.



Rue, H., Martino, S., and Chopin, N. (2009).  
Approximate Bayesian inference for latent Gaussian models by using  
integrated nested laplace approximations.

*Journal of the Royal Statistical Society: Series B (Statistical  
Methodology)*, 71(2):319–392.




Van Niekerk, J., Krainski, E., Rustand, D., and Rue, H. (2023).  
A new avenue for Bayesian inference with INLA.

*Computational Statistics & Data Analysis*, 181:107692.



van Niekerk, J. and Rue, H. (2024).  
Low-rank variational Bayes correction to the Laplace method.

*Journal of Machine Learning Research*, 25(62):1–25.



# Thank you • شكرا



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