

5/17/2021 GMAC analysis and pvalue plots

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1. Overview

Both MRPC-LOND and MRPC-ADDIS techniques inferred a large number of trans mediated trios. The trans-mediation model has been previously identified, but is not the commonly acknowledged mode of mediation. Since this result is surprising relative to the existing literature, we sought to apply another method for inferring mediation on a subset of GTEx trios analyzed herein by MRPC. The Genomic Mediation analysis with Adaptive Confounding (GMAC) algorithm allows for a unique selection of a subset of potential confounders, \mathbf{X}_{ij} from a larger covariate pool, \mathbf{H} , for each trio. By taking advantage of the Principle of Mendelian Randomization, the authors filter \mathbf{H} by removing common child and intermediate confounding variables (e.g variables associated with the eQTL as well as the cis/trans genes). Post-filtering, GMAC performs a mediation test on the edge between the cis gene and trans gene via the regression of the trans-gene T_j on the cis-eQTL L_i , cis-gene C_i , and the set of adaptively selected confounders \mathbf{X}_{ij} :

$$T_j = \beta_0 + \beta_1 C_i + \beta_2 L_i + \mathbf{F}\mathbf{X}_{ij} + \epsilon \quad (1)$$

The mediation statistic is the observed t -value of the cis-gene coefficient β_1 . A null distribution for no-mediation is constructed by iteratively permuting the values of the cis-transcript within each genotype and repeating the above regression. The authors argue that the permutation of the cis-transcript within the genotypes of the cis-eQTL removes the association between the cis and trans gene transcripts while preserving the higher order associations with the cis-eQTL. The resulting permutation test for mediation compares the observed relationship between the trans and cis gene to a null distribution constructed from a model with no association and assuming that possible confounding has been well adjusted via the selected covariates.

It is important to note that the above mediation test describes only the association between cis-gene and trans-gene transcripts ($C_i \leftrightarrow T_j$) and does not consider possible effects between the cis-eQTL and the cis-gene transcript ($L_i \rightarrow C_i$), or the cis-eQTL and trans-gene transcript ($L_i \rightarrow T_j$).

2. Methods

2.1 Applying GMAC to GTEx Trios

To compare the GMAC and MRPC algorithms, we applied the GMAC algorithm to the top five GTEx tissues by sample size. Following with the creators of GMAC, we used the full set of principle components retained from the PCA of the expression matrix as the covariate pool, and three additional known confounders: the PCR used, the platform used, and sex of the individual in each sample [yang2017identifying].

Consistent with yang2017identifying, the analysis was performed using a common child and intermediate variable filtering FDR of 10% and a confounder selection FDR of 5% for each trio. Each trio supplied to

GMAC consisted of the cis-QTL and the PEER normalized cis and trans gene transcripts with the highest association to the eQTL. To mitigate missing values in the eQTL matrix, multiple imputation of the matrix of unique cis-eQTLs was performed via multiple correspondence analysis (MCA) prior to its use in GMAC [JOSSE2016MISSMDA]. The analysis was performed twice on each trio, first with the cis gene as the mediator and second with the trans gene as the mediator. This allowed for GMAC inferred trios to be decomposed into the three groupings used under MRPC: 1) Cis-gene mediation, 2) Trans-gene mediation, 3) both (undirected).

2.2 Comparison of GMAC and MRPC Results

After applying GMAC to each tissue, the false discovery rate among the retained mediation p-values was controlled at the more liberal rate of 10% [YANG2017IDENTIFYING]. Each trio determined to have significant mediation after FDR filtering was compared with the regulatory network type inferred by MRPC-ADDIS. MRPC-ADDIS can infer three types of regulatory networks that contain an edge between the cis and trans gene (M1, M2, or M4). Since GMAC considers only the presence of the edge and not its direction, trios inferred to be one of M1, M2, or M4 under ADDIS, that were also significant under GMAC, were considered consistent (e.g. $C_i \rightarrow T_j$; $T_j \rightarrow C_i$; $C_i \leftrightarrow T_j$ are synonymous under GMAC).

2.3 Simulations Using Real Trios

- The Small True Model Simulation (STM)

To further understand the conflict in edge determination between MRPC-ADDIS and GMAC, we simulated the mediation test - the test for the β_1 coefficient in the presence of all adaptively selected confounders, \mathbf{X}_{ij} , as described by the regression in **eq (1)** - under two different scenarios. (i) To observe the predictive power of the mediation test when the trans gene comes from a set of explanatory variables smaller than those in \mathbf{X}_{ij} , We simulated the trans gene of each trio using the linear relationship:

$$T_j^* = \hat{\beta}_0 + \hat{\beta}_1 C_i + \hat{\beta}_2 L_i + \hat{\mathbf{\Gamma}}_W \mathbf{W}_{ij} + \epsilon \quad (2)$$

where T_j^* is the simulated trans gene, the coefficients are replaced by their estimates from the regression in **eq (1)**, the errors are $\epsilon \sim N(0, \hat{\sigma})$, and \mathbf{W}_{ij} is a subset of confounders in \mathbf{X}_{ij} representing the “highly” significant confounders from **eq (1)** ($p < 0.001$). Note that if the GMAC inferred mediation type was trans gene mediation only then the cis gene was simulated and the mediation test was performed on the β_1 coefficient from regression of the cis gene: $C_i = \beta_0 + \beta_1 T_j + \beta_2 L_i + \mathbf{\Gamma} \mathbf{X}_{ij} + \epsilon$.

We refer to the trans-gene generating function in **(2)** as the Small True Model (STM) as the simulated trans gene comes from a model that is a subset of the explanatory variables in the analysis model described in **(1)**. The simulated mediation test then replaces T_j in **(1)** with the simulated trans gene T_j^* . Therefore, the simulated mediation test under the STM can be decomposed as the test on the β_1 coefficient obtained from the regression:

$$T_j^* = \beta_0 + \beta_1 C_i + \beta_2 L_i + \mathbf{\Gamma}_{1,W} \mathbf{W}_{ij} + \mathbf{\Gamma}_{2,M} \mathbf{M}_{ij} + \epsilon \quad (3)$$

where \mathbf{M}_{ij} represents the additional confounders in \mathbf{X}_{ij} that are not included in \mathbf{W}_{ij} see **Figure 1**.

- The Large True Model Simulation (LTM)

Conversely, a second simulation model was implemented to observe the power of the mediation test when the generating model for the trans gene is larger than the model used to infer the mediation relationship. In this scenario, the trans gene is simulated via:

$$T_j^* = \hat{\beta}_0 + \hat{\beta}_1 C_i + \hat{\beta}_2 L_i + \hat{\Gamma}_V \mathbf{V}_{ij} + \epsilon \quad (4)$$

which we refer to as the Large True Model (LTM). Note that the coefficient estimates in (4) come from the regression of the original data on a larger set than in eq. (1):

$$T_j = \beta_0 + \beta_1 C_i + \beta_2 L_i + \Gamma \mathbf{X}_{ij} + \Gamma_G \mathbf{G}_{ij} + \epsilon \quad (5)$$

where \mathbf{G}_{ij} represents the additional explanatory variables in \mathbf{V}_{ij} not in \mathbf{X}_{ij} (see **Figure 1**). The additional variables represented by \mathbf{G}_{ij} were randomly selected from the confounder pool \mathbf{H} with equal probability. The four sets of p -values (GMAC mediation p -value, GMAC permutation p -value, the STM mediation p -value, and the LTM mediation p -value) were visually inspected to determine the effect of model mis-specification on the power of the mediation and permutation test(s).

- True GMAC Model Simulation (TGM)

We performed a third simulation of the trans gene using the analysis model in GMAC given by (1) as the trans gene generating process. We then applied the MRPC model

$$T_j = \beta_0 + \beta_1 C_i + \beta_2 L_i + \Gamma_z \mathbf{Z}_{ij} \quad (6)$$

where the column dimension \mathbf{Z}_{ij} is a less than or equal to the column dimension of \mathbf{W}_{ij} and represents only the confounders selected under the MRPC PC-selection methodology. The goal was to analyze the power of MRPC in detecting the mediation edge when the data is generated from the much larger GMAC model.

- GMAC Self Simulation (GSS)

We created a fourth simulation scenario to test the importance of permutation on the analysis result. Each of the above simulations seeks to understanding the inferential power of MRPC/GMAC under model mis-specification. In this scenario we take a different approach and use the model suggested by GMAC for each trio given by eqn. (1) to generate the trans gene such that:

$$T_j^* = \hat{\beta}_0 + \hat{\beta}_1 C_i + \hat{\beta}_2 L_i + \hat{\Gamma} \mathbf{X}_{ij} + \epsilon$$

where $\epsilon \sim N(0, \hat{\sigma})$. We then apply the correct model (1) to each simulated trio under this scenario to obtain the parametric and permutation inferences when we know the analysis model is correct.

2.4 Simulations Using Synthetic Trios

- **Purpose:** In the simulations above, it becomes increasingly difficult to determine the robustness of the GMAC performance compared with other algorithms because we do not know the true model of the trio. We decided to perform a second set of simulations where each trio is generated under a specific model type. Therefore, with the truth known prior to inference, we are better able to compare the inferential differences between MRGN, GMAC, and MRPC-ADDIS. Additionally, we simulate each trio with “known” confounders, u “unknown” confounders, w intermediate variables, and z common child variables as defined by Yang et. al., 2017. In doing so, we are also able to compare the confounder selection procedures of GMAC and MRGN/MRPC to better discern the appropriate procedure for selecting the confounder set.

- **Method:** We simulated 10,000 trios of all model types (M0 - M4) with u “unknown” confounding variables (PCs) represented by the matrix \mathbf{U} and 3 simulated “known” confounders represented by the matrix \mathbf{K} , w intermediate variables represented by the matrix \mathbf{W} , and z common child variables represented by the matrix

Z . we simulate with $u, w, z \sim N(\mu = 10, \sigma = 2)$ such that $u, w, z \in [1, \infty)$. We include \mathbf{K} to satisfy the input requirement of GMAC. The possible PCs involved in each trio are selected randomly from the whole genome PC matrix for the GTEx tissue Whole Blood. Additionally, the known confounders for each trio are the “pcr”, “platform”, and “sex” covariates from Whole Blood. We simulate all confounder effects uniformly where $\gamma_i \sim \text{unif}(a = 0.15, b = 0.5)$ and select half at random to be negative. Further, we simulate the SNP effects as $\beta_{11}, \beta_{12} \sim \text{unif}(a = 0.5, b = 1.5)$, and the mediation effects as $\beta_{21}, \beta_{22} \sim \text{unif}(a = 0.5, b = 1.0)$ following Yang et. al., 2017. We simulated the effects of the known confounders such that $\phi_i \sim \text{unif}(0.01, 0.1)$.

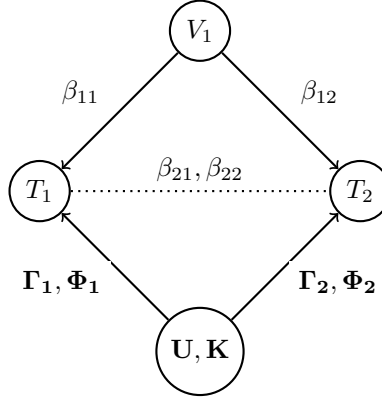
(1.) - We simulate the variant $V_1 = [v_1, v_2, \dots, v_n]$ from a multinomial distribution under Hardy-Weinberg equilibrium:

$$Pr(v_i = 0) = (1 - \theta)^2; Pr(v_i = 1) = 2\theta(1 - \theta); Pr(v_i = 2) = \theta^2$$

(2.) - We denote the j^{th} molecular phenotype as T_j and its set of parents by $\mathbf{Parents}(T_j)$. The set of parent nodes $\mathbf{Parents}(T_j)$ may be empty, contain V -nodes, T -nodes, or a combination of both (Badsha and Fu, 2019). We model T_j as following a normal distribution where

$$T_j \sim N \left(\beta_0 + \sum_{i \in \mathbf{Parents}(T_j)} \beta_i V_i + \left(\sum_{k \in \mathbf{Parents}(T_j)} \beta_k T_k \right) + \mathbf{\Gamma} \mathbf{U} + \mathbf{\Phi} \mathbf{K}, \sigma_j \right) \quad (7)$$

Figure 1: The confounder model



$$T_j \sim N \left(\beta_0 + \sum_{i \in \mathbf{Parents}(T_j)} \beta_i V_i + \left(\sum_{k \in \mathbf{Parents}(T_j)} \beta_k T_k \right) + \mathbf{\Gamma} \mathbf{U} + \mathbf{\Phi} \mathbf{K}, \sigma_j \right) \quad (8)$$

$$T_j \sim N \left(\beta_0 + \sum_{i \in \mathbf{Parents}(T_j)} \beta_i V_i + \left(\sum_{k \in \mathbf{Parents}(T_j)} \beta_k T_k \right) + \mathbf{\Gamma} \mathbf{U} + \mathbf{\Phi} \mathbf{K} + \mathbf{\Omega}_2 \mathbf{W}, \sigma_j \right) \quad (9)$$

$$W_j \sim N \left(\beta_0 + \sum_{i=1}^w \omega_{1,i} T_i, \sigma_j \right) \quad (10)$$

Figure 2: The common child model

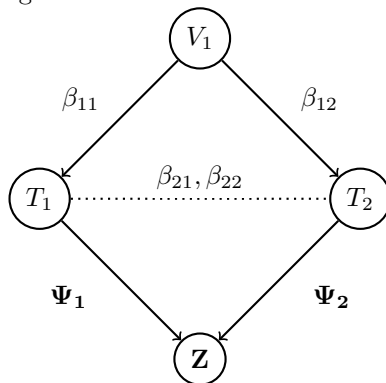
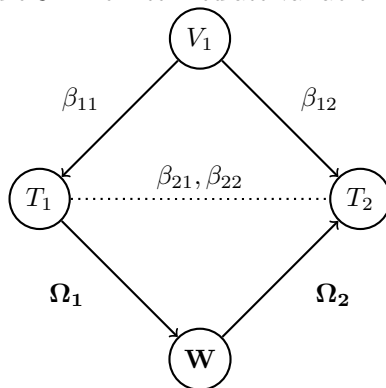


Figure 3: The intermediate variable model



• **Assessing Performance** We apply GMAC twice to each simulated trio, once with T_1 as the mediator and second with T_2 as the mediator. This allows for a determination of the direction of the mediation edge. That is, combining the two inferences allows for the determination of Cis mediation ($T_1 \rightarrow T_2$), Trans mediation ($T_2 \rightarrow T_1$) or both (bi-directional $T_1 - T_2$). To compare the three algorithms across the simulated trios, we define an accuracy metric based on the mediation edge (i.e the edge between nodes T_1 and T_2) for GMAC. If the true model adjacency matrix \mathbf{A} contains an edge between the genes $T_1 - T_2$; $T_1 \rightarrow T_2$; $T_2 \rightarrow T_1$ and GMAC correctly predicts that edge (and its direction), we apply an accuracy of 100%. In the case of the inference performed by GMAC we define:

$$Acc = \begin{cases} 1, & \text{if Inferred} = \text{Cis Mediation} | \text{Truth} = M1.1 \text{ or } M2.2 \\ 1, & \text{if Inferred} = \text{Trans Mediation} | \text{Truth} = M1.2 \text{ or } M2.1 \\ 1, & \text{if Inferred} = \text{Both} | \text{Truth} = M4 \\ 1, & \text{if Inferred} = \text{No Med} | \text{Truth} = M0.1, M0.2 \text{ or } M3 \\ 0, & \text{else} \end{cases}$$

For each simulated trio we perform the GMAC, MRPC-ADDIS, and MRGN algorithms

3. Results

3.1 Comparing GMAC and MRPC

In light of the surprising number of trans-gene mediation trios inferred by MRPC, we sought to compare our results with GMAC by applying the GMAC method to the top five GTEx tissues by sample size. It is important to note that the test for mediation used by GMAC describes only the association between cis-gene and trans-gene transcripts ($C_i \leftrightarrow T_j$) and does not consider the possible effects between the cis-eQTL and the cis-gene ($L_i \rightarrow C_i$), or the cis-eQTL and trans-gene ($L_i \rightarrow T_j$). Therefore, since GMAC considers only the presence of the mediation edge, trios inferred to be one of M1, M2, or M4 under ADDIS, that were also significant under GMAC, were considered consistent (e.g $C_i \rightarrow T_j$; $T_j \rightarrow C_i$; $C_i \leftrightarrow T_j$ are synonymous under GMAC).

At the 10% false discovery rate, GMAC identified 2,160 trios with an edge between the cis and trans genes out of 55,446 total trios tested across the five selected tissues: Adipose subcutaneous, Tibial artery, Muscle skeletal, Sun exposed skin, and Whole blood. Of the trios with mediation edges, 653 were identified as the cis gene mediating the trans gene, 245 as trans gene mediating the cis gene and 1,345 as both (29.1%, 10.9%, and 60% respectively). As can be seen from **Table 2**, the consistency in inferred mediation edges between the two methods varied between 39% and 50% of the trios across tissues.

3.2 Simulation Results

- - We noticed that the mediation parametric p-value can vary quite a bit. we ran each simulation scenario 1000 times on each trio and obtained the median p-value from each of the 4 scenarios. We used this as a robust estimate of the parametric p-value
- - The nominal p-value and mediation p-value give more or less the same value except for rare variant trios
- - We noticed that Badsha implemented the PC selection code for MRPC in a way that led to the qvalues of each correlation test being applied after the p-values had already been bonferroni adjusted. This lead to extremely few PCs being selected for each trio. After implementing the correct PC selection process, MRPC and GMAC select equivalent numbers of PCs
- - Figure 1 panel D supports that the nominal p-value appears to be better at determining the presence of a mediation edge for rare variants (points on the nomial p axis are all significant where as some trios/rare

variants are insignificant on the parametric axis). This supports the use of the permuted regression test when the instrumental variable/ cis-eQTL contains a rare allele

- - Figures 4-5 and Tables 4-9 represent two cases where the marginal tests performed by MRPC falsely remove an edge in the graph that is supported by higher order tests. These cases justify the need to bypass the marginal tests and proceed directly to conditional tests which have increased power in these scenarios.

1 Tables and Figures

Table 1: Descriptive statistics for the distribution of missing values across the eQTL’s for each tissue used in GMAC

	Adipose Subcutaneous	Artery Tibial	Muscle Skeletal	Skin Sun Exposed	Whole Blood
Min.	0.000000	0.000000	0.000000	0.000000	0.000000
1st Qu.	0.000000	0.000000	0.000000	0.000000	0.000000
Median	0.000000	0.000000	0.000000	0.000000	0.000000
Mean	0.006365	0.006625	0.006560	0.006701	0.006103
3rd Qu.	0.003442	0.003425	0.002833	0.003306	0.002985
Max.	0.156627	0.159247	0.158640	0.160331	0.155224

Table 2: The breakdown of unique trios with inferred significant cis or trans mediation under GMAC across their respective ADDIS inferred regulatory networks. The column “Percentage In Common” is the proportion of significant trios that also contained a mediation edge in the regulatory network inferred under ADDIS

Tissue	M0	M1	M2	M3	M4	Other	Total GMAC Inferred	Percentage In Common
AdiposeSubcutaneous	107	47	12	102	134	2	404	0.4777
ArteryTibial	88	37	10	108	112	0	355	0.4479
MuscleSkeletal	126	37	8	145	132	2	450	0.3933
SkinSunExposed	118	42	12	132	139	0	443	0.4357
WholeBlood	107	55	29	142	171	4	508	0.5020

Table 3: Breakdown of trios with inferred mediation under GMAC across both cis and trans mediation types. The column “Unique Both” represents the intersection of the columns “Total Cis Mediated” and “Total Trans Mediated”

Tissue	Sample Size	Tested Trios	Cis-Gene Mediation Trios	Trans-Gene Mediation Trios	Unique Cis	Unique Trans	Unique Both	Unique Total
AdiposeSub	581	11850	374	282	122	30	252	404
ArteryTibial	584	11471	334	226	129	21	205	355
MuscleSkeletal	706	10257	401	314	136	49	265	450
SunExp.Skin	605	13045	404	333	110	39	294	443
WholeBlood	670	8823	451	398	110	57	341	508

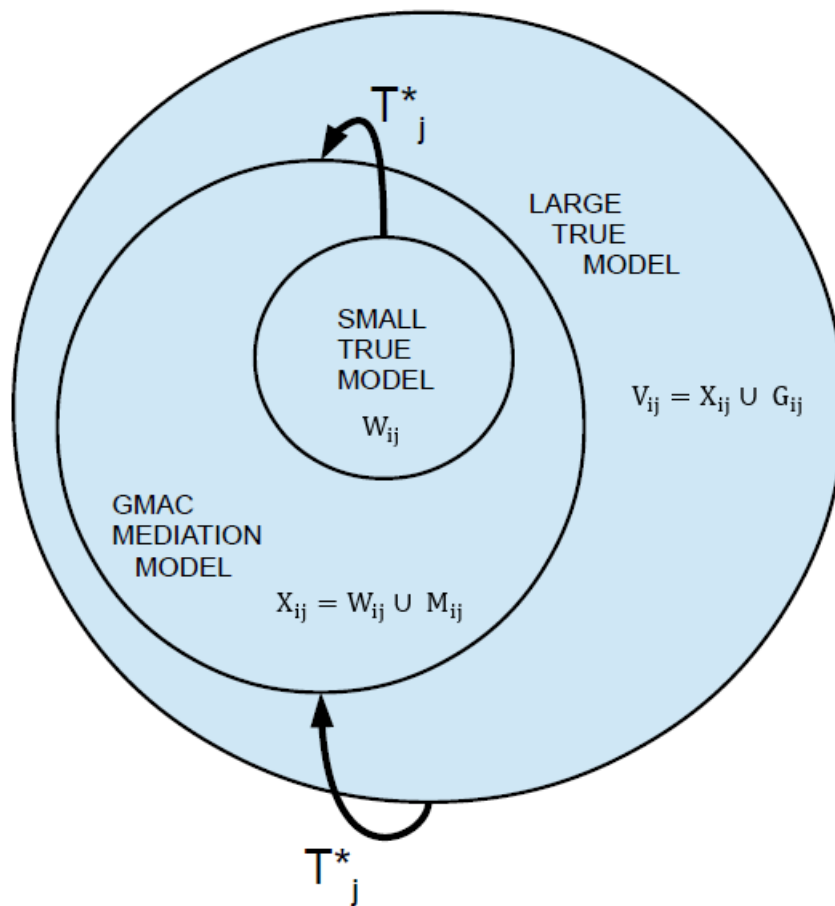


Figure 4: A visual depiction of the relationship between the STM and LTM simulation models for the trans gene and the regression used for the mediation test

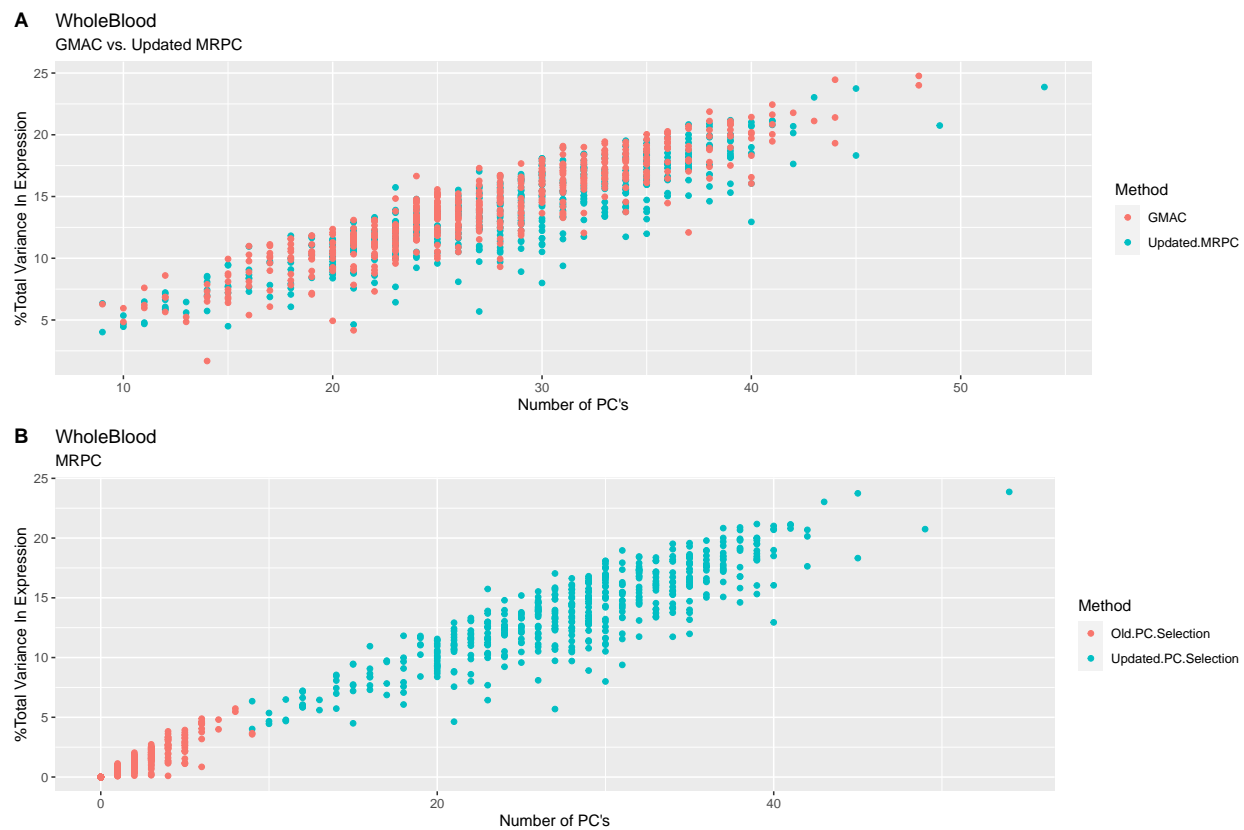


Figure 5: A) The number of PC's included for each trio in whole blood after using the corrected PC selection for MRPC with the number of PC's selected by GMAC. B) compares the number of PC's selected for each trio under Badsha's original code vs the corrected version of PC selection

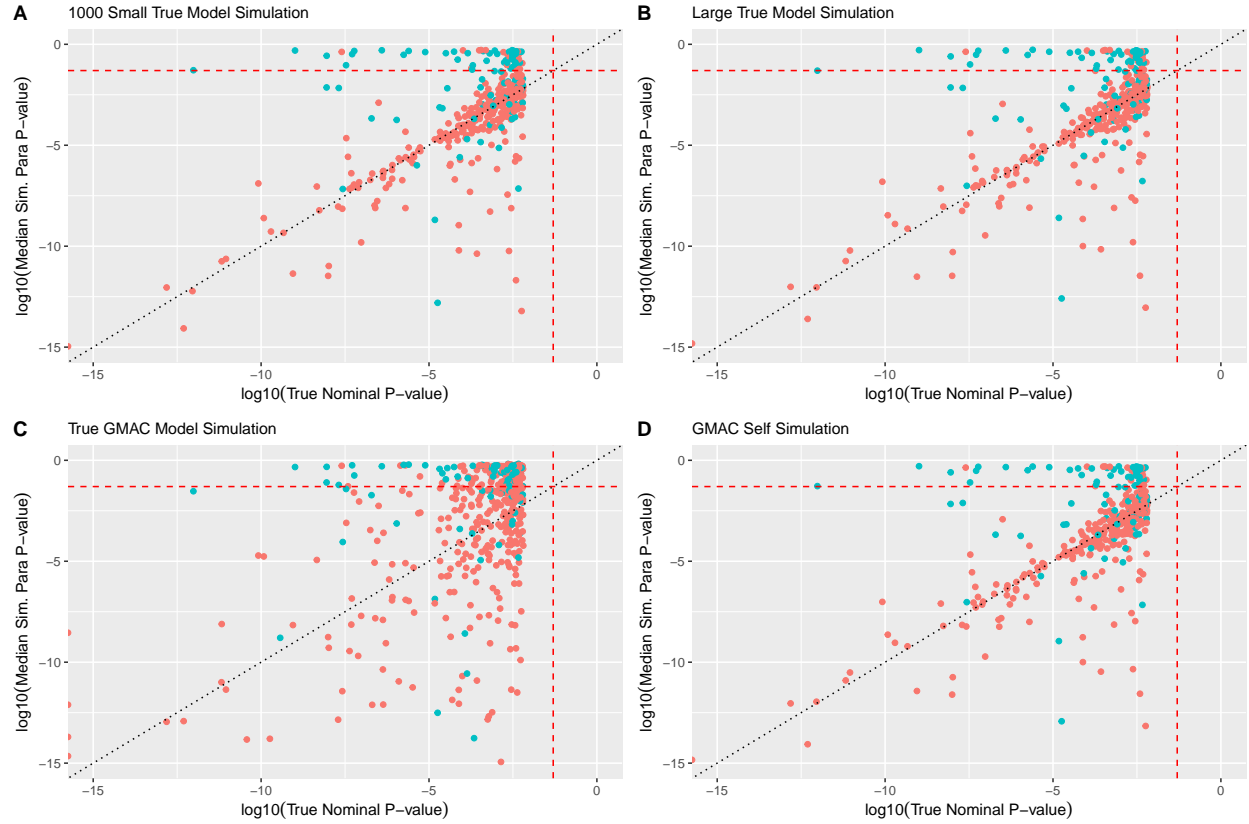
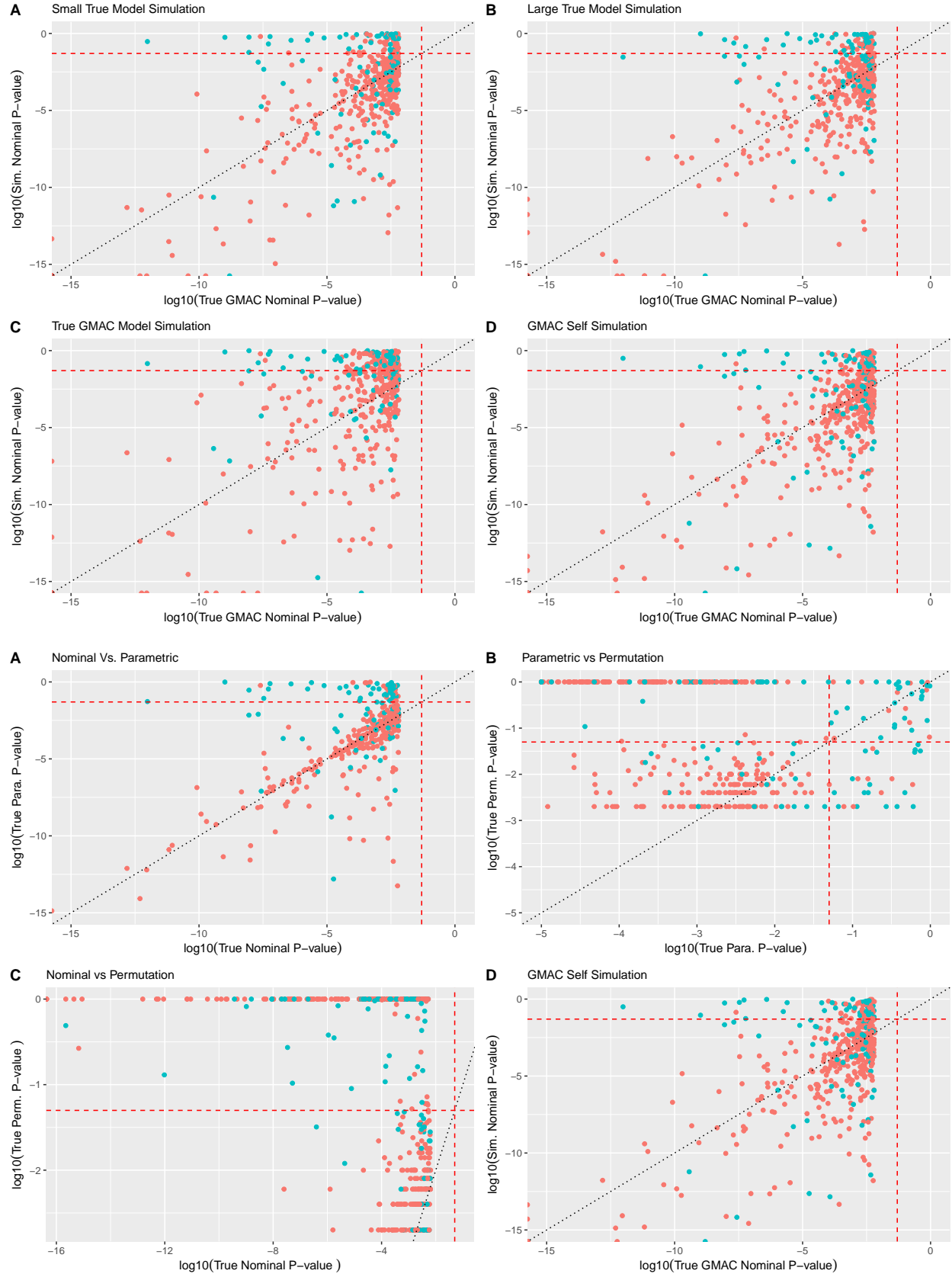


Figure 6: All panels: comparison of the median p -value from the test for the mediation edge from 1000 iterations of each simulation scenario for each trio (y-axis) with the nominal p -value for the edge from the true model (x-axis). Each panel corresponds to a specific simulation scenario. These graphs support the need to use the permutation test when the genotype has a low frequency allele.



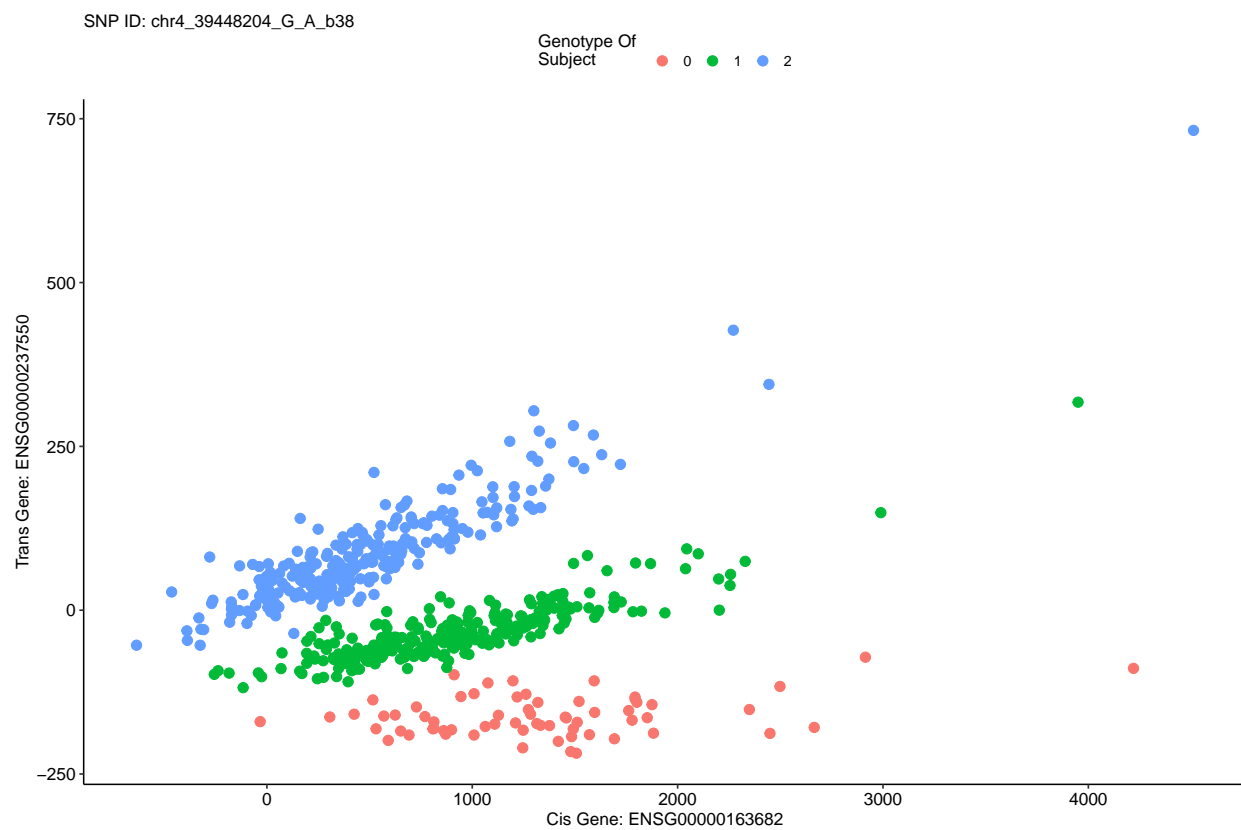


Figure 7: Trio number 9 from the GTEx tissue Adipose Subcutaneous

Table 4: Results showing the terms of interest for the Regression:
 $\text{cis.gene} \sim \text{SNP} + \text{trans.gene} + \text{PCs}$ for trio 9 from the GTEx tissue
 Adipose Subcutaneous

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	2070.6858	75.9773	27.2540	0
trans.gene	5.0385	0.2468	20.4178	0
SNP	-1050.0066	35.6044	-29.4909	0

Table 5: Results showing the terms of interest for the Regression:
 $\text{trans.gene} \sim \text{SNP} + \text{cis.gene} + \text{PCs}$ for trio 9 from the GTEx tissue
 Adipose Subcutaneous

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-260.4547	10.3493	-25.1664	0
cis.gene	0.0856	0.0042	20.4178	0
SNP	159.5456	3.0487	52.3325	0

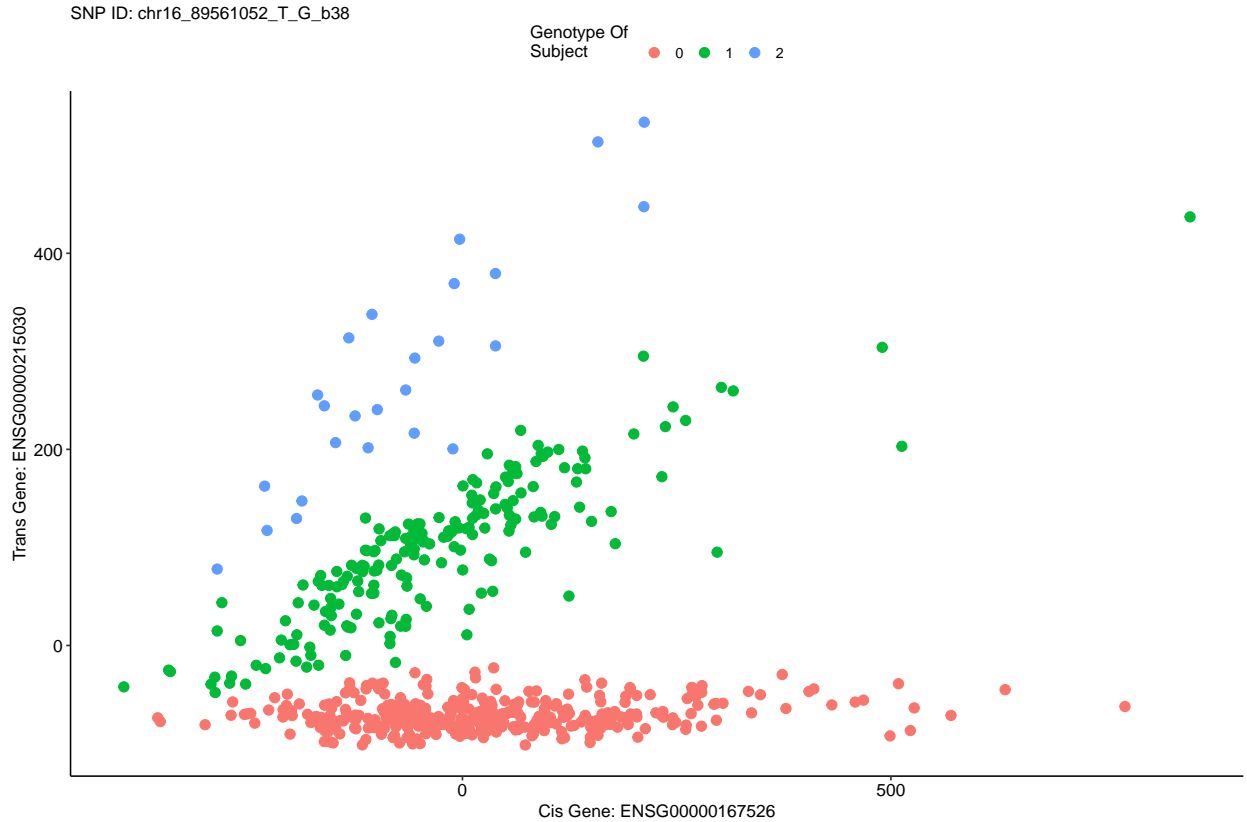


Figure 8: Trio number 1922 from the GTEx tissue Artery Tibial

Table 6: Results showing the terms of interest for the Regression: $\text{cis.gene} \sim \text{SNP} + \text{trans.gene} + \text{PCs}$ for trio 1922 from the GTEx tissue Artery Tibial

	Estimate	Std. Error	t value	$\text{Pr}(> t)$
(Intercept)	41.2423	27.1619	1.5184	0.1295
trans.gene	1.0308	0.1027	10.0398	0.0000
SNP	-222.7941	19.8243	-11.2384	0.0000

Table 7: Results showing the terms of interest for the Regression: $\text{trans.gene} \sim \text{SNP} + \text{cis.gene} + \text{PCs}$ for trio 1922 from the GTEx tissue Artery Tibial

	Estimate	Std. Error	t value	$\text{Pr}(> t)$
(Intercept)	-90.2013	9.5820	-9.4136	0
cis.gene	0.1480	0.0147	10.0398	0
SNP	180.7298	3.2883	54.9611	0

Table 8: Correlations between the cis and trans genes trio 9 from the GTEx tissue Adipose Subcutaneous and trio 1922 from Artery Tibial. Notice that the correlation between the cis and trans gene is quite weak despite an obvious within-genotype correlation between the two molecular phenotypes. Thus MRPC removes this edge in the marginal tests (given by the pvalues) despite being clearly supported in the higher order conditional tests.

	cor	df	t	p-value
Trio 9 AS	0.0681	579	1.6417	0.1012
Trio 1922 ArT	0.0561	582	1.3546	0.1761

	cor	df	t	p-value
Trio 1922 ArT	0.0560624	582	1.354617	0.1760653

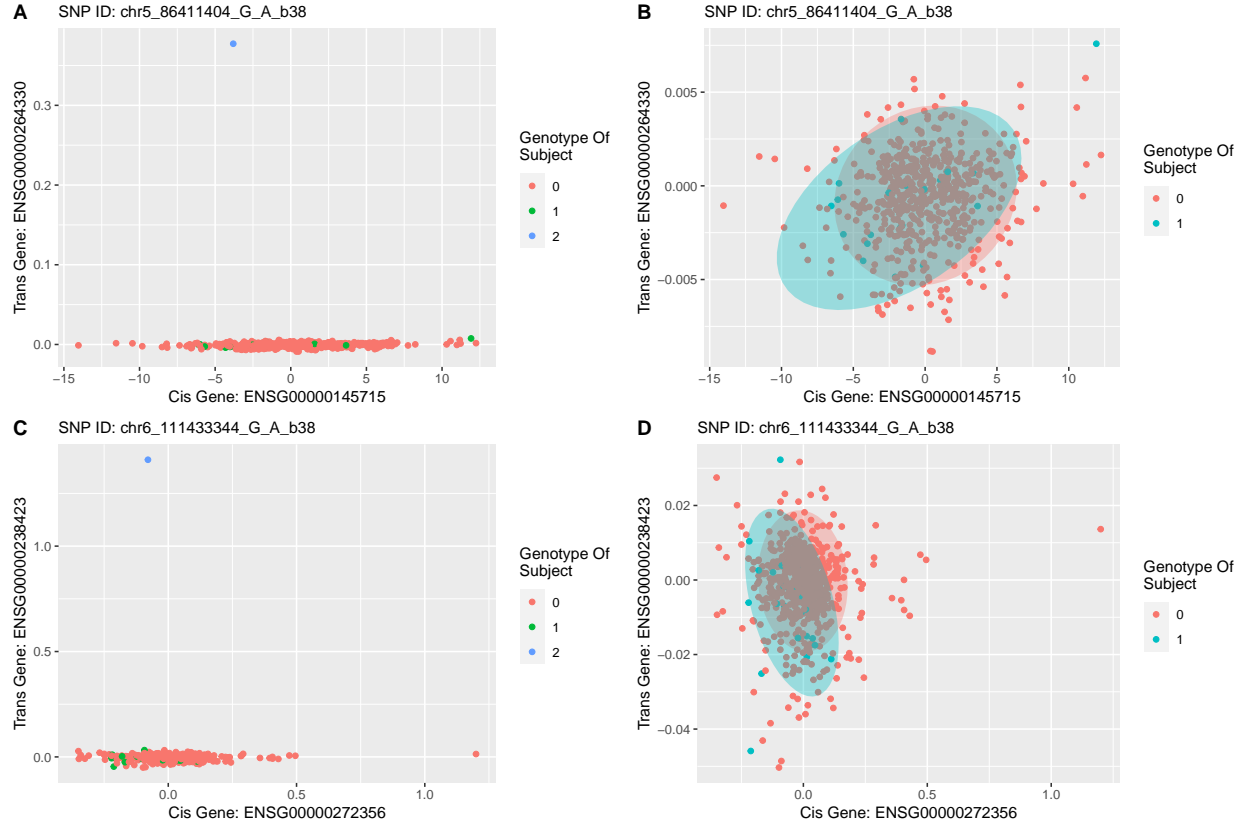


Figure 9: **A** and **C**: Example scatter plots of trios from subcutaneous adipose tissue and whole blood (respectively) with a rare allele present in the sample: Note that 0 indicates individuals homozygous for the reference allele, 1 indicates heterozygous individuals and 2 indicates individuals who are homozygous for the alternative (rare) allele. The apparent outlier represents a single individual in the sample who was homozygous for the rare allele, and **B** and **D** are the scatter plots with the point(s) for the homozygous-alternative individuals removed and a confidence ellipse calculated over the remaining homozygous-reference and heterozygous individuals.