

Inverse Weighting Methods

March 11, 2011

IPW Estimator

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$$\widehat{ATE}_{IPW} = \frac{1}{n} \sum_{i=1}^n \left\{ \frac{T_i Y_i}{\hat{\pi}(X_i)} - \frac{(1 - T_i) Y_i}{1 - \hat{\pi}(X_i)} \right\}$$

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- Common alternative:

$$\left\{ \sum_{i=1}^n \frac{T_i}{\hat{\pi}(X_i)} \right\}^{-1} \sum_{i=1}^n \frac{T_i Y_i}{\hat{\pi}(X_i)} - \left\{ \sum_{i=1}^n \frac{1 - T_i}{1 - \hat{\pi}(X_i)} \right\}^{-1} \frac{(1 - T_i) Y_i}{1 - \hat{\pi}(X_i)}$$

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- IPW estimator “upweights” treated units with a low probability of treatment and “down-weights” controls that have a high probability of being in treatment.

Unbiasedness of the IPW Estimator

$$E[\widehat{ATE}_{IPW}] = \frac{1}{n} \sum_{i=1}^n \left\{ E \left[\frac{T_i Y_i}{\pi(X_i)} \right] - E \left[\frac{(1 - T_i) Y_i}{1 - \pi(X_i)} \right] \right\}$$

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- Disadvantages: boundary problems (propensity scores close to 0 or 1): “be a hero, don’t divide by zero”
- Still possible to check balance.

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- If one of the models is correctly specified, then the estimator is consistent.
- If both models are correct, then the AIPW achieves the semiparametric efficiency bound.
- If weights are highly variable, then AIPW is not robust.
- Many estimators have this double robustness property, including regression on a matched dataset.

AIPW Estimator

$$\begin{aligned}\widehat{ATE} &= \frac{1}{n} \sum_{i=1}^n \left[\frac{T_i Y_i}{\hat{\pi}(X_i)} - \frac{(1 - T_i) Y_i}{1 - \hat{\pi}(X_i)} \right] - \frac{T_i - \hat{\pi}(X_i)}{\hat{\pi}(X_i)(1 - \hat{\pi}(X_i))} \\ &\times [(1 - \hat{\pi}(X_i))E(Y_i | T_i = 1, X_i) + \hat{\pi}(X_i)E(Y_i | T_i = 0, X_i)]\end{aligned}$$

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- First line corresponds to basic IPW estimator.
- Second line adjusts this estimator by a weighted average of the two regression estimators.
- The adjustment set X_i need not be the same ones used in the propensity score.

AIPW Estimator

- AIPW estimator is more robust to small weights than the IPW estimator. To see why, set $T_i = 1$. The estimator becomes:

$$\begin{aligned} & \frac{Y_i}{\hat{\pi}(X_i)} - \frac{1}{\hat{\pi}(X_i)} \\ \times & [(1 - \hat{\pi}(X_i))E(Y_i|T_i = 1, X_i) + \hat{\pi}(X_i)E(Y_i|T_i = 0, X_i)] \\ = & \left[\frac{Y_i}{\hat{\pi}(X_i)} - \frac{[1 - \hat{\pi}(X_i)]E(Y|T=1, X)}{\hat{\pi}(X_i)} \right] - E(Y|T = 0, X) \end{aligned}$$

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- When $\hat{\pi}(X_i)$ is close to 0, then $\frac{Y_i}{\hat{\pi}(X_i)}$ will get large in absolute value. This is a problem.
- However, it is mitigated by the fact that $\frac{[1 - \hat{\pi}(X_i)]E(Y|T = 1, X)}{\hat{\pi}(X_i)}$ gets large at the same rate as $\frac{Y_i}{\hat{\pi}(X_i)}$

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