

$$L(\theta; x) = p(x; \theta)$$

data parameter

what is the "likelihood"
of these parameters
given this data

what is the prob. of the data
under these values of
the parameter

Mathematically equivalent, conceptually two different universes

$$\hat{\theta}_{MLE} := \underset{\theta \in \Theta}{\operatorname{argmax}} \{ L(\theta; x) \}$$

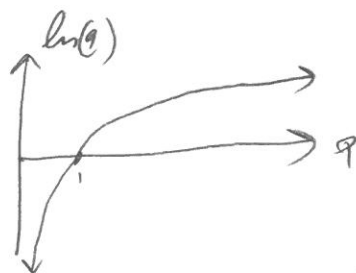
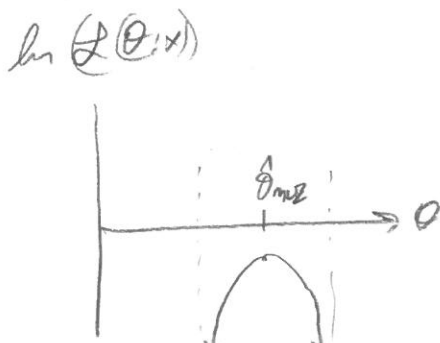
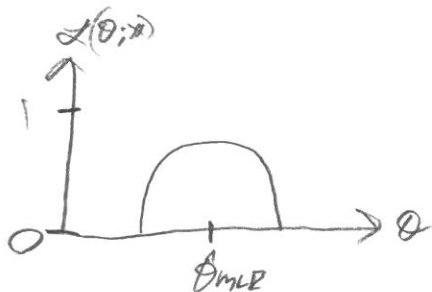
i.e. the most likely parameter value
assuming this data

Note: a positive monotonic transformation g is one that preserves the following:

$$x > y \Rightarrow g(x) > g(y) \quad \text{i.e. a strictly increasing function}$$

If $\hat{\theta}_{MLE}$ is the best because it maximizes $L(\theta; x)$ then
it will also maximize

$$g(L(\theta; x))$$



let

$$l(\theta; x) := \ln(L(\theta; x))$$

$$\hat{\theta}_{MLE} = \underset{\theta \in \Theta}{\operatorname{argmax}} \{ l(\theta; x) \}$$

$$X_1, \dots, X_6 \stackrel{iid}{\sim} \text{bin}(\theta) \quad \text{sum } 0,0,1,0,1,0,7$$

(2)

Let's go back for this comp... what to you think it is?

$$\ell(\theta; x) = \ln(L(\theta; x)) = \ln\left(\prod_{i=1}^6 \theta^{x_i} (1-\theta)^{1-x_i}\right)$$

$$= \sum_{i=1}^6 \ln \theta^{x_i} (1-\theta)^{1-x_i}$$

$$= \sum_{i=1}^6 x_i \ln \theta + (1-x_i) \ln(1-\theta)$$

$$= \ln \theta \sum x_i + (6 - \sum x_i) \ln(1-\theta)$$

$$\ell(\theta; x_1, \dots, x_6) = L(\theta; x_1, \dots, x_6)$$

$$\prod \Rightarrow \sum$$

prod \Rightarrow sum
(HARD) (EASY)

$$\text{Roll } \bar{x} = \frac{1}{n} \sum x_i \Rightarrow \sum x_i = n \bar{x}$$

$$= \ln \theta (6 \bar{x}) + (6 - 6 \bar{x}) \ln(1-\theta)$$

$$= 6 \left(\bar{x} \ln \theta + (1-\bar{x}) \ln(1-\theta) \right)$$

Now... we maximize...

use calculus!

$$0 \stackrel{\text{set}}{=} \frac{d}{d\theta} [\] = 6 \left(\bar{x} \left(\frac{1}{\theta} \right) + (1-\bar{x}) \left(\frac{-1}{1-\theta} \right) \right)$$

$$0 = \bar{x} \frac{1}{\theta} - (1-\bar{x}) \frac{1}{1-\theta}$$

$$0 = \bar{x} \left(\frac{1-\theta}{\theta} \right) - (1-\bar{x})$$

$$0 = \bar{x}(1-\theta) - (1-\bar{x})\theta = \bar{x} - \bar{x}\theta - \theta + \bar{x}\theta \Rightarrow \boxed{\theta = \bar{x} = \hat{\theta}_{MLE}} = \hat{p}$$

$$\text{From case } \bar{x} = \frac{2}{6} = \frac{1}{3}$$

Our calc. above was independent of $n=6$. It works for all n .

$\hat{\theta}_{MLE} = \bar{x}$, $\hat{\theta}_{MLE} = \bar{X}$ "estimator"
 estimate quicker notation!

MLE's give us only estimates but they have really nice properties

① $\hat{\theta}_{MLE} \xrightarrow{p} \theta$

By def: $\forall \epsilon > 0 \quad \lim_{n \rightarrow \infty} P(|\hat{\theta}_{MLE} - \theta| \geq \epsilon) = 0$

$\Rightarrow \hat{\theta}_{MLE}$ becomes arbitrarily close to θ with high n

② Asymptotic Normality

$$\hat{\theta}_{MLE} \xrightarrow{d} N(\theta, SE(\hat{\theta}_{MLE})^2)$$

③ Efficiency $SE(\hat{\theta}_{MLE})$ is the smallest standard error for all consistent estimators (Cramer-Rao bound)

In the Bernoulli case $\hat{\theta}_{MLE} = \bar{X}$

$$SE(\hat{\theta}_{MLE})$$

$$\hat{\theta}_{MLE} = \bar{X} \approx N\left(\theta, \left(\sqrt{\frac{\theta(1-\theta)}{n}}\right)^2\right) \approx N\left(\bar{X}, \left(\sqrt{\frac{\bar{X}(1-\bar{X})}{n}}\right)^2\right)$$

$$X_1, \dots, X_n \stackrel{iid}{\sim} \text{Bern}(\theta) \Rightarrow \hat{\theta}_{MLE} = \bar{X}$$

$$X_1, \dots, X_n \stackrel{iid}{\sim} \text{Geom}(\theta) := (1-\theta)^X \theta$$

different parameterization than in B241

Here, X is the # of failures before the first success.

$$\text{Supp}(X) = \{0, 1, \dots\} (= \mathbb{N}_0)$$

④ = (0.1) param spec

$$f(\theta; x) = P(x; \theta) = \prod_{i=1}^n (1-\theta)^x \theta = \theta^n (1-\theta)^{\sum x_i}$$

$$l(\theta; x) = \ln(\cdot) = n \ln \theta + \sum x_i \ln(1-\theta)$$

$$l'(\theta; x) = \frac{n}{\theta} - \frac{\sum x_i}{1-\theta} \stackrel{\text{set}}{=} 0 \Rightarrow n(1-\theta) - \theta \sum x_i = 0$$

$$\Rightarrow n - n\theta = \theta \sum x_i$$

$$\Rightarrow n = \theta (\sum x_i + n)$$

$$\Rightarrow \hat{\theta}_{MLE} = \frac{n}{\sum x_i + n} = \frac{n}{\bar{x}n + n} = \frac{1}{\bar{x} + 1}$$

What would it be for old parametrization? $\frac{1}{\bar{x}}$ sure \bar{x} always gets one more
trick

$SE[\hat{\theta}_{MLE}] = SE\left[\frac{1}{\bar{x}+1}\right] \dots$ HARDER but totally possible

Inference Based on MLE's (preparations needed)

(1) Pt Estimate $\hat{\theta}_{MLE}$

(2) Conf set $CI_{\theta, 1-\alpha} := \left[\hat{\theta}_{MLE} \pm z_{\frac{\alpha}{2}} SE[\hat{\theta}_{MLE}] \right]$

(3) Hyp testing: $H_0: \theta = \theta_0$

$H_1: \theta \neq \theta_0$

your choice

Rejection Region: $\left[\theta_0 \pm z_{\frac{\alpha}{2}} SE[\hat{\theta}_{MLE}] \right]_{\theta_0}$
if $\hat{\theta}_{MLE} \in \text{Rejection Region} \Rightarrow \text{Fail to reject}$

Observed Data \rightarrow Pick $K \rightarrow$ Do Inference via MLE

What's wrong with this? P P P P