

Probability Theory

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1 Review of Measure Theory

Modern probability theory is based on measure theory and we will in this section recall some notions from measure theory.

Definition 1.1 (Algebra). Given a set Ω , a set of subsets \mathcal{A} of Ω is an algebra if $\Omega \in \mathcal{A}$ and \mathcal{A} is closed under finite union and complements.

It follows straight away that an algebra is also closed under finite intersections.

Definition 1.2 (Finitely Additive Measure). A function $\mu : \mathcal{A} \rightarrow [0, \infty]$ where \mathcal{A} is an algebra, is a finitely additive measure if for any disjoint sets $A, B \in \mathcal{A}$,

$$\mu(A \cup B) = \mu(A) + \mu(B).$$

Definition 1.3 (σ -Algebra). A σ -algebra \mathcal{F} is an algebra that is closed under countable unions.

Similarly, it follows that \mathcal{F} is closed under countable intersections.

Definition 1.4 (Measure). A function $\mu : \mathcal{F} \rightarrow [0, \infty]$ where \mathcal{F} is a σ -algebra, is a σ -additive measure (or simply measure) if given a sequence of pairwise disjoint sets A_1, A_2, \dots of \mathcal{F} , we have

$$\mu\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} \mu(A_i).$$

We call a measure a probability measure if $\mu(\Omega) = 1$.

Definition 1.5 (σ -Finite Measure). A measure μ is said to be σ -finite if there exists a sequence of pairwise disjoint sets A_1, A_2, \dots of \mathcal{F} , such that $\bigcup_{i=1}^{\infty} A_i = \Omega$ and for all i , $\mu(A_i) < \infty$.

Definition 1.6 (Probability Space). A probability space is the triple $(\Omega, \mathcal{F}, \mathbb{P})$ consisting of a set Ω , a σ -algebra \mathcal{F} on Ω and \mathbb{P} a probability measure on \mathcal{F} .

We call elements of \mathcal{F} (i.e. a \mathcal{F} -measurable set) an event.

Proposition 1.1 (Continuity of Measures). Let $(A_n)_{n \in \mathbb{N}} \subseteq \mathcal{F}$, then

- (continuity from below) if (A_n) is increasing, then

$$\mathbb{P}\left(\bigcup_{n=1}^{\infty} A_n\right) = \lim_{n \rightarrow \infty} \mathbb{P}(A_n).$$

- (continuity from above) if (A_n) is decreasing, then

$$\mathbb{P}\left(\bigcap_{n=1}^{\infty} A_n\right) = \lim_{n \rightarrow \infty} \mathbb{P}(A_n).$$

We recall the the finiteness of the measure is vital for continuity from below while continuity from above is also valid for general measures.

Proof. Exercise. □

Proposition 1.2. A finitely additive probability measure on the σ -algebra \mathcal{F} is a probability measure if and only if it is continuous at 0.

Proof. The forward direction follows from above so we will prove the reverse. Suppose μ is finitely additive and for any decreasing $(A_n) \subseteq \mathcal{F}$ with $\bigcap A_n = \emptyset$, we have $\lim_{n \rightarrow \infty} \mu(A_n) = 0$. Then, μ is continuous from below, and so for any sequence of disjoint sets (B_n) , we have $(C_n) := (\bigcup_{i=1}^n B_i)$ is a sequence of increasing sets and thus,

$$\mathbb{P}\left(\bigcup_{i=1}^{\infty} B_i\right) = \mathbb{P}\left(\bigcup_{i=1}^{\infty} C_i\right) = \lim_{n \rightarrow \infty} \mathbb{P}(C_n) = \lim_{n \rightarrow \infty} \mathbb{P}\left(\bigcup_{i=1}^n B_i\right) = \lim_{n \rightarrow \infty} \sum_{i=1}^n \mathbb{P}(B_i)$$

implying μ is σ -additive and so, μ is a measure. □

Proposition 1.3. Given a collection $\{\mathcal{F}_i\}_{i \in I}$ σ -algebras of Ω , $\bigcap_{i \in I} \mathcal{F}_i$ is also a σ -algebra on Ω .

Definition 1.7 (σ -Algebra Generated By Sets). Given a collection of subsets S of Ω , the σ -algebra generated by S is

$$\sigma(S) := \bigcap \{\mathcal{F} \text{ a } \sigma\text{-algebra} \mid S \subseteq \mathcal{F}\}.$$

Definition 1.8 (Borel σ -Algebra). Given a topological space (X, \mathcal{T}) , the Borel σ -algebra on X is $\mathcal{B}(X) := \sigma(\mathcal{T})$.

Definition 1.9 (Product σ -Algebra). Given measurable spaces $(\Omega_1, \mathcal{F}_1), (\Omega_2, \mathcal{F}_2)$, the product σ -algebra on $\Omega_1 \times \Omega_2$ is

$$\mathcal{F}_1 \otimes \mathcal{F}_2 := \sigma(\mathcal{F}_1 \times \mathcal{F}_2) = \sigma(\{A_1 \times A_2 \mid A_1 \in \mathcal{F}_1, A_2 \in \mathcal{F}_2\}).$$

Definition 1.10 (Cylindrical σ -Algebra). A set $C \subseteq \mathbb{R}^\infty$ is said to be cylindrical if is of the form

$$C = \{x \in \mathbb{R}^\infty \mid (x_1, \dots, x_n) \in C_n\}$$

where $C_n \in \mathcal{B}(\mathbb{R}^n)$. The set of cylindrical sets $\mathcal{B}(\mathbb{R}^\infty)$ form a σ -algebra on \mathbb{R}^∞ and is called the cylindrical σ -algebra.

Definition 1.11 (Consistent). The sequence of measures \mathbb{P}_n on $(\mathbb{R}^n, \mathcal{B}(\mathbb{R}^n))$ is said to be consistent if for all $n \in \mathbb{N}$, $\mathbb{P}_{n+1}(B_n \times \mathbb{R}) = \mathbb{P}_n(B_n)$ for all $B_n \in \mathcal{B}(\mathbb{R}^n)$.

Theorem 1 (Kolmogorov). Given any consistent sequence of measures \mathbb{P}_n , there exists a unique probability measure \mathbb{P} on $(\mathbb{R}^\infty, \mathcal{B}(\mathbb{R}^\infty))$ such that,

$$\mathbb{P}(\{x \in \mathbb{R}^\infty \mid (x_1, \dots, x_n) \in B_n\}) = \mathbb{P}_n(B_n)$$

for all $n \geq 1$, $B_n \in \mathcal{B}(\mathbb{R}^n)$.

Proof. Simply define the inner measure on the generating sets as claimed and use the Caratheodory extension (which provides both existence and uniqueness). □

Recall that a nondecreasing function g on \mathbb{R} is continuous up to possibly countably many discontinuities of the first kind. Furthermore, the derivative g' exists λ -a.e. (where λ is the Lebesgue measure on \mathbb{R}).

Proposition 1.4. Let $(\mathbb{R}, \mathcal{B}(\mathbb{R}), \mathbb{P})$ be a probability space. Defining $F(x) := \mathbb{P}(-\infty, x]$, we have

- F is nondecreasing;
- $\lim_{x \rightarrow -\infty} F(x) = 0$ and $\lim_{x \rightarrow \infty} F(x) = 1$;
- F is continuous on the right.

Proof. Clear by the monotonicity, continuity of measures (from above). \square

Definition 1.12 (Distribution Function). Any function $F : \mathbb{R} \rightarrow [0, 1]$ satisfying the above three properties is said to be a distribution function on \mathbb{R} .

It is clear that any probability measure induces a distribution. On the other hand the converse is also true.

Proposition 1.5. Given a distribution function F , there exists a unique probability measure \mathbb{P} on $(\mathbb{R}, \mathcal{B}(\mathbb{R}))$ such that $F(x) = \mathbb{P}(-\infty, x]$ for all $x \in \mathbb{R}$.

Proof. Use Caratheodory extension theorem on the algebra $\{(-\infty, x] \mid x \in \mathbb{R}\}$ mapping $(-\infty, x] \mapsto F(x)$. The uniqueness of the probability measure follows by the uniqueness of the Caratheodory extension. \square

Definition 1.13 (Null-set). Given a measure μ , a set $S \subseteq \Omega$ is a null-set if there exists some measurable set $N \subseteq \Omega$ with measure 0 such that $S \subseteq N$.

Definition 1.14 (Complete Measure). A measure μ is complete if every μ -null set is measurable.

If a measure on the σ -algebra Σ is not complete, we may complete the σ -algebra by extending Σ to

$$\overline{\Sigma} := \sigma(\Sigma \cup \{N \mid N \text{ is a null-set}\}).$$

Clearly, the null-sets will have measure 0.

We note that the probability space $(\mathbb{R}, \mathcal{B}(\mathbb{R}), \mathbb{P})$ is not complete as there exists subsets of a Borel null-set which are not Borel. With this in mind, we denote the completion of $\mathcal{B}(\mathbb{R})$ by $\mathcal{M}(\mathbb{R})$ and we say \mathbb{P} is the Lebesgue-Stieltjes measure.

Recall, that in elementary probability theory, we considered three types of distributions, namely discrete, absolutely continuous and singular continuous. Let us now consider them again in a formal measure theoretic setting.

- (Discrete) A random variable $X : \Omega \rightarrow \mathbb{R}$ is said to have discrete distribution if there exists some countable (including finite) set $A \subseteq \mathbb{R}$, such that for all $E \in \mathcal{B}(\mathbb{R})$, the push-forward measure satisfies

$$X_*\mathbb{P}(E) = \sum_{x \in A} p(x)\delta_x(E)$$

where $p(x) = X_*\mathbb{P}(\{x\})$ and δ_x is the Dirac measure at x .

We note that a distribution function F corresponds to a discrete random variable if and only if for all $x_0 \in \mathbb{R}$,

$$F(x_0) = \sum_{x \in A \cap \{\leq x_0\}} p(x).$$

It is clear that $\sum_A p(x) = 1$ since $\sum_A p(x) = X_*\mathbb{P}(\mathbb{R}) = 1$.

- (Absolutely continuous) A random variable $X : \Omega \rightarrow \mathbb{R}$ is said to be absolutely continuous if $X_*\mathbb{P} = f\lambda$ for some Lebesgue integrable function f and λ denotes Lebesgue measure. Thus, the distribution function corresponding to X satisfies

$$F(x) = \int_{(-\infty, x]} f d\lambda.$$

In particular, recalling the Radon-Nikodym theorem, we have X is absolutely continuous if and only if $X_*\mathbb{P} \ll \lambda$ (hence the name “absolutely continuous”).

Before introducing the last type of distribution, let us consider the following definition.

Definition 1.15 (Concentrated). A measure μ on the measurable space X is said to be concentrated on a measurable set A if $\mu(E) = 0$ for all $E \subseteq X \setminus A$.

- (Singular continuous) A random variable $X : \Omega \rightarrow \mathbb{R}$ is singular continuous if its distribution function F is continuous and $X_*\mathbb{P}$ is concentrated on a set A of Lebesgue measure 0 for which $F'(x) = 0$ for all $x \in A$ almost everywhere.

We note that, since $\lambda(A) = 0$, $X_*\mathbb{P} \perp \lambda$ by the set A (hence the name “singular”). Moreover, by continuity, $X_*\mathbb{P}(\{x\}) = 0$ for all $x \in \mathbb{R}$ in contrast to the discrete measure.

Analogous to the Lebesgue decomposition of measures, we may decompose any distribution function into a discrete, absolutely continuous and singular continuous distribution.

Theorem 2 (Hahn Decomposition for Distributions). Given a distribution function F , there exists $a_1 + a_2 + a_3 = 1$ and $F_{\text{disc}}, F_{\text{ac}}, F_{\text{sc}}$ discrete, absolutely continuous and singular continuous distribution functions respectively, such that

$$F = a_1 F_{\text{disc}} + a_2 F_{\text{ac}} + a_3 F_{\text{sc}}.$$

Proof. Recalling the refinement of the Lebesgue decomposition where we may decompose a measure μ with

$$\mu = \mu_d + \mu_a + \mu_s$$

where μ_d is a discrete measure, $\mu_a \ll \lambda$ and μ_s is singular continuous (i.e. mutually singular with respect to the Lebesgue measure and $\mu_s\{x\}$ for all x). Thus, by simply taking the decomposition of the measure corresponding to F (i.e. $X_*\mathbb{P}$), we obtain the required decomposition after normalization. \square

2 Random Variables

We will continue to let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space.

Definition 2.1 (Random Variable). A function $\xi : \Omega \rightarrow \mathbb{R}$ is said to be a random variable if it is \mathcal{F} -measurable (i.e. for any $B \in \mathcal{B}(\mathbb{R})$, we have $\xi^{-1}(B) \in \mathcal{F}$).

While we have already introduced the notion of a distribution within the previous section, we will present it here again for organization.

Definition 2.2 (Distribution of a Random Variable). Given a random variable ξ , the distribution of ξ is the push-forward measure of \mathbb{P} along ξ . Furthermore, the distribution function corresponding to ξ is

$$F(x) := X_*\mathbb{P}(-\infty, x].$$

Definition 2.3. Given a random variable ξ , we define $\mathcal{F}_\xi \subseteq \mathcal{F}$ to be the σ -algebra

$$\mathcal{F}_\xi := \{\xi^{-1}(B) \mid B \in \mathcal{B}(\mathbb{R})\}.$$

This is the least σ -algebra for which ξ is measurable.

We will recall some standard results about measurable functions. All proofs are left as exercises and can be found in the second year measure theory notes.

Lemma 2.1. If $\mathcal{B}(\mathbb{R}) = \sigma(\mathcal{D})$ for a collection of sets \mathcal{D} , ξ is a random variable if $\xi^{-1}(D) \in \mathcal{F}$ for all $D \in \mathcal{D}$.

Lemma 2.2. Given random variables f, g and $c \in \mathbb{R}$, $f + g, f - g, c \cdot f, |f|, fg, \max(f, g)$, and $\min(f, g)$ are all random variables. Furthermore, if $g(x) \neq 0$ for all x , then f/g is also a random variable.

Lemma 2.3. If (f_n) is a sequence of random variables, then

$$\sup_n f_n, \inf_n f_n, \lim_n f_n$$

are random variables if they exist.

Lemma 2.4. If ξ is a random variable and $f : \mathbb{R} \rightarrow \mathbb{R}$ is continuous, then $f(\xi)$ is a random variable.

Definition 2.4 (Simple Function). A random variable ξ is simple if there exists a partition of Ω , D_1, \dots, D_n such that

$$\xi(\omega) = \sum_{i=1}^n x_i 1_{D_i}(\omega)$$

for some x_1, \dots, x_n for all $\omega \in \Omega$.

Lemma 2.5. For any non-negative random variable ξ , there exists a sequence of nondecreasing simple random variables (ξ_n) such that for all $\omega \in \Omega$,

$$\xi_n(\omega) \uparrow \xi(\omega).$$

Definition 2.5 (Random Vector). A function $\xi : \Omega \rightarrow \mathbb{R}^n$ is a random vector if it is measurable. Again, we define its distribution to be its push-forward measure.

Lemma 2.6. $\xi : \Omega \rightarrow \mathbb{R}^n$ is a random vector if and only if $\xi_i := \text{pr}_i \circ \xi$ is a random variable for all $i = 1, \dots, n$ (where $\text{pr}_i : \mathbb{R}^n \rightarrow \mathbb{R}$ is the i -th projection function).

Definition 2.6 (Independent Random Variables). Two random variables $\xi, \eta : \Omega \rightarrow \mathbb{R}$ are said to be independent if

$$(\xi, \eta)_* \mathbb{P} = \xi_* \mathbb{P} \otimes \eta_* \mathbb{P}.$$

Since, to check that two measures are equal, it suffices to check equality on generating sets, ξ, η are independent if for all $A, B \in \mathcal{B}(\mathbb{R})$,

$$\mathbb{P}(\xi \in A, \eta \in B) = \mathbb{P}(\xi \in A) \mathbb{P}(\eta \in B).$$

Let us quickly recall the construction of the Lebesgue integral.

1. Define the Lebesgue integral for simple functions.
2. Define the Lebesgue integral for non-negative functions by taking the limit of the Lebesgue integral of the monotone sequence of simple functions which converge to the said function.
3. Define the Lebesgue integral for general real-valued functions f by taking $\int f = \int f^+ - \int f^-$ if $\int |f| < \infty$.

Definition 2.7 (Expectation). Given a random variable $\xi : \Omega \rightarrow \mathbb{R}$, the expectation of ξ is simply

$$\mathbb{E}(\xi) := \int \xi d\mathbb{P}$$

if it exists. Furthermore, we say ξ is integrable if $\mathbb{E}(|\xi|) < \infty$.

Proposition 2.1. Let ξ, η be integrable random variables and let $c \in \mathbb{R}$, then

- $\mathbb{E}(c) = c$;
- $\mathbb{E}(\xi + \eta) = \mathbb{E}(\xi) + \mathbb{E}(\eta)$;
- $\xi \leq \eta$ a.e. implies $\mathbb{E}(\xi) \leq \mathbb{E}(\eta)$;
- $\xi = \eta$ a.e. implies $\mathbb{E}(\xi) = \mathbb{E}(\eta)$;
- $\xi \geq 0$ a.e. and $\mathbb{E}(\xi) = 0$ implies $\xi = 0$ a.e.

Proof. Follows directly from the properties of the Lebesgue integral. □

Let us recall some convergence theorems for the Lebesgue integral.

Theorem 3 (Dominated Convergence Theorem). Let (ξ_n) be a sequence of random variables such that $\xi_n \rightarrow \xi$ almost everywhere. If there exists some integrable η such that $|\xi_n| \leq \eta$ for all n , then, ξ is integrable and

$$\lim_{n \rightarrow \infty} \mathbb{E}(\xi_n) = \mathbb{E}(\xi).$$

Theorem 4 (Monotone Convergence Theorem). Let (ξ_n) be a sequence of non-negative increasing random variables. Then,

$$\lim_{n \rightarrow \infty} \mathbb{E}(\xi_n) = \mathbb{E} \lim_{n \rightarrow \infty} \xi_n.$$

We note that the right hand side limit always exists since for all $\omega \in \Omega$, $\xi_n(\omega)$ is increasing any bounded by ∞ .

We remark that the monotone convergence theorem applies if there exists some random variable η such that $\mathbb{E}(\eta) > -\infty$ such that $\eta \leq \xi_n$ for all n by considering $\xi_n - \eta$.

Corollary 4.1. If (η_n) is a sequence of non-negative random variables, then

$$\sum_{i=1}^{\infty} \mathbb{E}(\eta_i) = \mathbb{E} \left(\sum_{i=1}^{\infty} \eta_i \right).$$

Corollary 4.2 (Fatou's lemma). Let ξ_n be a sequence of non-negative random variables. Then,

$$\mathbb{E}(\liminf_n \xi_n) \leq \liminf_n \mathbb{E}\xi_n.$$

Proof. Apply the monotone convergence theorem to $\lambda_n := \inf_{k>n} \xi_k$. □

Again, the non-negative condition can be replaced by the existence of some random variable η such that $\mathbb{E}(\eta) > -\infty$ and $\eta \leq \xi_n$ for all n . On the other hand, if $\mathbb{E}(\eta) < \infty$ and $\xi_n \leq \eta$, the theorem holds with limit supremum instead.

We note that in all above theorems, the statement still holds by replacing Ω with any measurable set by restricting the measure onto that set.

Theorem 5 (Change of Variables). Given a random variable ξ , a measurable function $g : \mathbb{R} \rightarrow \mathbb{R}$ and a measurable set A , we have

$$\int_A g d\xi_* \mathbb{P} = \int_{\xi^{-1}(A)} g \circ \xi d\mathbb{P},$$

where both integrals either exist or not exist simultaneously.

Proof. Apply usual method where one first prove the statement for indicator functions. Then, it follows that it holds for simple functions by the linearity of the integral. Finally, for any non-negative measurable function, we take a sequence of monotonically increasing simple functions, and apply the monotone convergence theorem. For arbitrary functions, the result follows by taking $f = f^+ - f^-$. □

Corollary 5.1 (Law of the Unconscious Statistician). Given a random variable ξ and a measurable function $g : \mathbb{R} \rightarrow \mathbb{R}$,

$$\mathbb{E}(g(\xi)) = \int_{\mathbb{R}} g d\xi_* \mathbb{P}.$$

Corollary 5.2. Let $g : \mathbb{R} \rightarrow \mathbb{R}$ be measurable, then, if ξ be a discrete random variable,

$$\mathbb{E}(g(\xi)) = \sum_{x \in A} g(x)p(x).$$

On the other hand, if ξ is absolutely continuous, i.e. there exists some f such that $f\lambda = \xi_* \mathbb{P}$, then

$$\mathbb{E}(g(\xi)) = \int_{\mathbb{R}} g(x)f(x)\lambda(dx).$$

Proof. In the discrete case, we have

$$\mathbb{E}(g(\xi)) = \int g d \left(\sum_{x \in A} p(x) \delta_x \right) = \sum_{x \in A} p(x) \int g d \delta_x.$$

By considering $\int g d \delta_x = \int_{\{x\}} g d \delta_x + \int_{\mathbb{R} \setminus \{x\}} g d \delta_x = \delta_x(\{x\})g(x) + 0 = g(x)$. We have

$$\mathbb{E}(g(\xi)) = \sum_{x \in A} g(x)p(x),$$

as required.

On the other hand, if $f\lambda = \xi_*\mathbb{P}$, we have

$$\mathbb{E}(g(\xi)) = \int_{\mathbb{R}} g d(f\lambda) = \int_{\mathbb{R}} g(x)f(x)\lambda(dx)$$

as required. \square

Theorem 6 (Fubini's Theorem). Let $(E_1, \Sigma_1, \mu_1), (E_2, \Sigma_2, \mu_2)$ be σ -finite measure spaces. Then, for any $\Sigma_1 \otimes \Sigma_2$ -measurable functions $g : E_1 \times E_2 \rightarrow \mathbb{R}$, $g(\cdot, y_0)$ is Σ_1 -measurable for all $y_0 \in E_2$, and $g(x_0, \cdot)$ is Σ_2 -measurable for all $x_0 \in E_1$. Furthermore, $\int_{E_1} g d\mu_1, \int_{E_2} g d\mu_2$ are Σ_2 and Σ_1 -measurable respectively. Finally, if $\int |g| d\mu_1 \otimes \mu_2 < \infty$, then,

$$\int g d\mu_1 \otimes \mu_2 = \int \left(\int g(x, y) \mu_2(dy) \right) \mu_1(dx) = \int \left(\int g(x, y) \mu_1(dx) \right) \mu_2(dy).$$

2.1 Inequalities

Lemma 2.7 (Jensen's Inequality). Let ξ be an integrable random variable and let $g : \mathbb{R} \rightarrow \mathbb{R}$ be a measurable, convex function, then,

$$g(\mathbb{E}\xi) \leq \mathbb{E}g(\xi).$$

Proof. Recall that the function g is convex if for all $x_0 \in \mathbb{R}$, there exists some λ such that $g(x) \geq g(x_0) + (x - x_0)\lambda$ (graphically, λ is the slope (more accurately, a subderivative) of g at x_0 and so, the inequality is saying that the graph lies above the tangent line).

Setting $x = \xi$ and $x_0 = \mathbb{E}\xi$. Then, the above inequality becomes

$$g(\xi) \geq g(\mathbb{E}\xi) - (\xi - \mathbb{E}\xi)\lambda.$$

Thus, applying the expectation to both sides results in the required inequality by the linearity of the integral. \square

Corollary 6.1 (Lyapunov's Inequality). Let ξ be a random variable and let $0 < s < t$ be real numbers, then

$$\mathbb{E}(|\xi|^s)^{1/s} \leq \mathbb{E}(|\xi|^t)^{1/t}.$$

Proof. Use Jensen's inequality with $g(x) = |x|^{t/s}$.

Alternatively, setting $\eta = |\xi|^s$, by Hölder's inequality, we have

$$\|\xi^s\|_1 = \|\eta\|_1 \leq \|\eta\|_{t/s} = \|\xi\|_t^s.$$

Thus, taking both sides to the power of $1/s$, we obtain $\|\xi\|_s = (\|\xi^s\|_1)^{1/s} \leq (\|\xi\|_t^s)^{1/s} = \|\xi\|_t$ as required. \square

Proposition 2.2 (Markov Inequality). Let $\xi \geq 0$ be an integrable random variable and let $c > 0$. Then

$$\mathbb{P}(\xi \geq c) \leq \frac{\mathbb{E}(\xi)}{c}.$$

Proof. $\mathbb{E}(\xi) \geq \mathbb{E}(\xi \mathbf{1}_{\xi \geq c}) \geq c \mathbb{P}(\xi \geq c) = c \mathbb{P}(\xi \geq c)$. \square

Definition 2.8 (Variance). The variance (or dispersion) of a random variable ξ is defined to be

$$V_\xi := \mathbb{E}[(\xi - \mathbb{E}\xi)^2]$$

and we define $\sigma := \sqrt{V_\xi}$ the standard deviation of V_ξ .

By expanding the definition, we find $V_\xi = \mathbb{E}\xi^2 - (\mathbb{E}\xi)^2$.

Definition 2.9 (Covariance). The covariance of random variables ξ and η is defined to be

$$\text{cov}(\xi, \eta) := \mathbb{E}[(\xi - \mathbb{E}\xi)(\eta - \mathbb{E}\eta)].$$

Proposition 2.3. For random variables ξ, η , we have

- $V_{\xi+\eta} = V_\xi + V_\eta + 2\text{cov}(\xi, \eta)$;
- if $\text{cov}(\xi, \eta) = 0$, then $V_{\xi+\eta} = V_\xi + V_\eta$.

Proof. Clear. \square

Proposition 2.4 (Chebyshev's Inequality). Let ξ be a integrable random variable. Then, for all $\epsilon > 0$,

$$\mathbb{P}(|\xi - \mathbb{E}\xi| \geq \epsilon) \leq \frac{V_\xi}{\epsilon^2}.$$

Proof. By Markov inequality, for $\xi \geq 0$, we have

$$\mathbb{P}(\xi \geq \epsilon) = \mathbb{P}(\xi^2 \geq \epsilon^2) \leq \frac{\mathbb{E}\xi^2}{\epsilon^2}.$$

Thus, by replacing ξ by $|\xi - \mathbb{E}\xi|$, we have the required inequality. \square

Proposition 2.5 (Exponential Chebyshev's Inequality). Let $\xi \geq 0$ be a random variable and let $\epsilon, t > 0$ such that $\xi, e^{t\xi}$ are integrable. Then,

$$\mathbb{P}(\xi \geq \epsilon) \leq e^{-t\epsilon} \mathbb{E}(e^{t\xi}).$$

Proof. We observe, by Markov inequality

$$\mathbb{P}(\xi \geq \epsilon) = \mathbb{P}(e^{t\xi} \geq e^{t\epsilon}) \leq \frac{E(e^{t\xi})}{e^{t\epsilon}}.$$

□

Proposition 2.6 (Tail Probability). Let $\xi \geq 0$ be an integrable random variable. Then,

$$\mathbb{E}(\xi) = \int_{(0,\infty)} \mathbb{P}(\xi \geq x) \lambda(dx).$$

Proof. By change of variable, we have,

$$\mathbb{E}\xi = \int_{\Omega} \xi d\mathbb{P} = \int_{(0,\infty)} x(\xi_*\mathbb{P})(dx) = \int_{(0,\infty)} \int_{[0,x]} \lambda(dt)(\xi_*\mathbb{P})(dx).$$

Then, by Fubini's theorem to the function $g : (t, x) \mapsto \mathbf{1}_{[0,x]}(t)$, we have

$$\int_{(0,\infty)} \int_{[0,x]} \lambda(dt)(\xi_*\mathbb{P})(dx) = \int_{(0,\infty)^2} g(t, x) \lambda(dt)(\xi_*\mathbb{P})(dx) = \int_{(0,\infty)} \mathbb{P}(\xi \geq x) \lambda(dx)$$

as required. □

Definition 2.10 (Normal Random Variable). A random variable ξ is said to be norm if $\xi_*\mathbb{P} = f\lambda$ where

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-m)^2}{2\sigma^2}}$$

for some $m \in \mathbb{R}, \sigma > 0$. We denote this by $\xi \sim \mathcal{N}(m, \sigma^2)$.

Proposition 2.7. Let $\xi \sim \mathcal{N}(m, \sigma^2)$. Then, $\mathbb{E}\xi = m$ and $V_{\xi} = \sigma^2$, and so, a normal random variable is fully determined by its mean and variance.

Proof. Exercise. □

Definition 2.11 (Moment). Given a random variable ξ , we define the k -th moment of ξ to be $\mathbb{E}(\xi^k)$.

2.2 Transformation of Random Variables

Let $F_{\xi}(x)$ be a distribution function of a random variable ξ . Then, if ϕ is a real valued continuous function, we would like to consider the distribution of η where $\eta = \phi(\xi)$. An easy observation is that

$$F_{\eta}(y) = \mathbb{P}(\eta \leq y) = \mathbb{P}(\eta \in \phi^{-1}(-\infty, y]) = \int_{\phi^{-1}(-\infty, y]} d(\xi_*\mathbb{P}).$$

As one might expect, elementary methods from first year probability are sufficient for most cases we encounter (consider the case ϕ is linear or quadratic).

Suppose now that ξ is absolutely continuous (and so, has a density by Radon-Nikodym), we would like to find the density of $\eta := \phi(\xi)$. Assume first that $\xi(\Omega) \in I$ where I is a finite or infinite open interval and let ϕ be continuously differentiable and strictly increasing on I . Denoting $h(y) = \phi^{-1}(\{y\})$ which is well-defined and differentiable, for all $y \in \phi(I)$,

$$F_\eta(y) = \mathbb{P}(\eta \leq y) = \mathbb{P}(\xi \leq \phi^{-1}(\{y\})) = \int_{(-\infty, h(y)]} f_\xi d\lambda = \int_{(-\infty, y]} f_\xi(h(z))h'(z)\lambda(dz)$$

where f_ξ is the density of ξ . Hence, the density of η is $f_\eta(h(y))h'(y) = f_\xi(h(y))|h'(y)|$. Similarly, if ϕ is strictly decreasing, η remain to have the density $f_\xi(h(y))|h'(y)|$. With this in mind, we may obtain the density for a large class of transformations by de composing the density into a strictly increasing and decreasing parts.

In the case that (ξ, η) is a random vector with joint distribution F and let ϕ be a continuous function. Then, $\phi(\xi, \eta)$ has distribution

$$F_{\phi(\xi, \eta)}(z) = \int_{\phi^{-1}(-\infty, z]} d((\xi, \eta)_* \mathbb{P}).$$

2.3 Independence

We recall that two random variables ξ, η are said to be independent if $(\xi, \eta)_* \mathbb{P} = \xi_* \mathbb{P} \otimes \eta_* \mathbb{P}$. Thus, if F_ξ, F_η are distributions of ξ and η respectively, then, $F_{(\xi, \eta)}(x, y) = F_\xi(x)F_\eta(y)$ for all $x, y \in \mathbb{R}$ where $F_{(\xi, \eta)}$ is a distribution of the random vector (ξ, η) .

Proposition 2.8. If ξ, η are independent random variables, then the distribution of $\xi + \eta$ is

$$F_{\xi+\eta}(z) = \int_{\mathbb{R}} F_\eta(z-x) \xi_* \mathbb{P}(dx) = \int_{\mathbb{R}} F_\xi(z-y) \eta_* \mathbb{P}(dy).$$

Proof. By taking $\phi : \mathbb{R}^2 \rightarrow \mathbb{R} : (x, y) \mapsto x + y$, we have

$$F_{\xi+\eta}(z) = \int_{\phi^{-1}(-\infty, z]} d((\xi, \eta)_* \mathbb{P}) = \int_{\phi^{-1}(-\infty, z]} d(\xi_* \mathbb{P} \otimes \eta_* \mathbb{P})$$

by independence. By considering that $(x, y) \in \phi^{-1}(-\infty, z]$ if and only if $x + y \leq z$, we have by Fubini's theorem,

$$\begin{aligned} \int_{\phi^{-1}(-\infty, z]} d(\xi_* \mathbb{P} \otimes \eta_* \mathbb{P}) &= \int_{\mathbb{R}^2} \mathbf{1}_{x+y \leq z} d(\xi_* \mathbb{P} \otimes \eta_* \mathbb{P}) \\ &= \int_{\mathbb{R}} \left(\int_{\mathbb{R}} \mathbf{1}_{x+y \leq z} \eta_* \mathbb{P}(dy) \right) \xi_* \mathbb{P}(dx) \\ &= \int_{\mathbb{R}} F_\eta(z-x) \xi_* \mathbb{P}(dx) = \int_{\mathbb{R}} F_\xi(z-y) \eta_* \mathbb{P}(dy). \end{aligned}$$

□

Recalling the definition of the convolution of a function, we may reformulate the above as the following corollaries.

Definition 2.12 (Convolution). Given two real-valued functions $f, g : \Omega \rightarrow \mathbb{R}$, we define the convolution of f with g by

$$f * g := t \mapsto \int_{\mathbb{R}} f(x)g(t-x)\mu(dx).$$

Corollary 6.2. The distribution function of the sum of two independent random variables is the convolution of their distribution functions.

Corollary 6.3. If ξ, η are independent absolutely continuous random variables, then, the density of $\xi + \eta$ is the convolution of their densities.

Proposition 2.9. Let ξ, η be independent integrable random variables. Then, $\xi \cdot \eta$ is integrable and $\mathbb{E}(\xi \cdot \eta) = \mathbb{E}(\xi)\mathbb{E}(\eta)$.

Proof. It is clearly true for indicator functions and so, we may extend to simple function by the linearity of expectation. Hence, by monotone convergence, the statement is true for non-negative random variables and hence true for arbitrary random variables by taking $\xi = \xi^+ - \xi^-$ and $\eta = \eta^+ - \eta^-$. \square

Definition 2.13. Random variables ξ, η are said to be uncorrelated if $\text{cov}(\xi, \eta) = 0$.

Proposition 2.10. Independent random variables are uncorrelated.

Proof. Clear since $\text{cov}(\xi, \eta) = \mathbb{E}(\xi \cdot \eta) - \mathbb{E}\xi\mathbb{E}\eta$. \square

We note that the converse is not true. Namely, uncorrelated does not imply independence.

2.4 Weak Law of Large Numbers

Consider $(\Omega_n, \mathcal{A}, \mathbb{P}_n)$ as a (finite) probability space such that

$$\Omega_n := \{\omega \mid \omega = (a_1, \dots, a_n), a_i \in \{0, 1\}\}, \mathcal{A} = \mathcal{P}(\Omega_n)$$

and

$$\mathbb{P}_n(\{\omega\}) = p(\omega) = p^{\sum_{i=1}^n a_i} q^{n - \sum_{i=1}^n a_i}$$

for some $0 < p < 1$ and $q = 1 - p$. Let $\xi_1, \dots, \xi_n : \Omega_n \rightarrow \{0, 1\}$ be random variables such that $\xi_i(\omega) = a_i$. It is easy to check that ξ_i are independent and identically distributed (iid.). Indeed,

$$(\xi_i)_* \mathbb{P}_n(\{1\}) = \mathbb{P}_n(\{\omega \mid a_j(\omega) = 1\}) = p \sum_{k=0}^{n-1} \binom{n-1}{k} p^k q^{n-k} = p$$

and $(\xi_i)_* \mathbb{P}_n(\{0\}) = (\xi_i)_* \mathbb{P}_n(\{1\}^c) = 1 - p = q$.

Now defining $S_n := \sum_{i=1}^n \xi_i$, we observe that

$$\mathbb{E}S_n = \sum_{i=1}^n \mathbb{E}\xi_i = \sum_{i=1}^n p = np.$$

Thus, the expectation of $\frac{1}{n}S_n$ is simply p . We now ask what is $|\frac{1}{n}S_n(\omega) - p|$. Immediately, we observe that $|\frac{1}{n}S_n(\omega) - p|$ cannot tends to 0 point-wise since $S_n(0) = 0$ for all n . Nonetheless, we observe that $\mathbb{P}_n(S_n = 0) = q^n \rightarrow 0$ as $n \rightarrow \infty$.

Recalling that the Kolmogorov extension theorem, as $\{\mathbb{P}_n\}$ is consistent, there exists a unique probability measure \mathbb{P} on $(\mathbb{R}^\infty, \mathcal{B}(\mathbb{R}^\infty))$ such that $\mathbb{P}(\xi \in \mathbb{R}^\infty \mid (\xi_1, \dots, \xi_n) \in B_n) = \mathbb{P}_n((\xi_1, \dots, \xi_n) \in B_n)$. With this measure in mind, using Chebyshev's inequality, and the fact that they are independent and hence, uncorrelated, we obtain

$$\mathbb{P}\left(\left|\frac{1}{n}S_n - p\right| \geq \epsilon\right) \leq \frac{V(\frac{1}{n}S_n)}{\epsilon^2} = \frac{1}{\epsilon^2} \sum_{j=1}^n V\left(\frac{1}{n}\xi_j\right) = \frac{1}{n^2\epsilon^2} \sum_{i=1}^n V_{\xi_i} = \frac{npq}{n^2\epsilon^2} \rightarrow 0$$

as $n \rightarrow \infty$. Thus, $\frac{1}{n}S_n$ converges to p in measure (probability).

This fact is known as Bernoulli's law of large numbers. By noting that we only used the uncorrelated fact, we obtain a more general theorem.

In general (with arbitrary ξ_i), we call the quantity $\frac{1}{n}S_n$ the time average and the (weak) law of large numbers tells us the time average converges in measure to the space average $\mathbb{E}\xi_i$.

Theorem 7 (Weak Law of Large Numbers). let ξ_1, ξ_2, \dots be integrable random variables. Defining $S_n^{(c)} = \sum_{i=1}^n (\xi_i - \mathbb{E}\xi_i)$ (we note that $\mathbb{E}S_n^{(c)} = 0$). Then, if ξ_1, ξ_2, \dots are uncorrelated and for all i , $V_{\xi_i} \leq C$ for some $C > 0$, for all $\epsilon > 0$,

$$\lim_{n \rightarrow \infty} \mathbb{P}\left(\left|\frac{1}{n}S_n^{(c)}\right| \geq \epsilon\right) = 0$$

i.e. $\frac{1}{n}S_n^{(c)} \rightarrow 0$ in measure.

Proof. By Chebyshev's inequality, as ξ_i are uncorrelated,

$$\mathbb{P}\left(\left|\frac{1}{n}S_n^{(c)}\right| \geq \epsilon\right) \leq \frac{V(\frac{S_n^{(c)}}{n})}{\epsilon^2} \leq \frac{c}{n\epsilon^2} \rightarrow 0$$

as $n \rightarrow \infty$. □

Corollary 7.1. If ξ_1, ξ_2, \dots integrable iid. random variables such that $V_{\xi_i} < \infty$, then, $\frac{1}{n} \sum_{i=1}^n \xi_i$ converges to $\mathbb{E}\xi_i$ in measure.

We would also like to consider the limiting behaviour of the distribution. Let us first recall the big and small-O notation.

Given sequences of functions $(f_n), (g_n)$, we denote $g_n = O_{n \rightarrow \infty}(|f_n|)$ if $|g_n/f_n|$ is eventually bounded, i.e. there exists some c, N such that for all $n \geq N$, $|g_n| \leq c|f_n|$. On the other hand, we denote $g_n = o_{n \rightarrow \infty}(|f_n|)$ if $\lim_{n \rightarrow \infty} |g_n/f_n| = 0$.

Returning to our example of the Bernoulli random variables, we have the following result.

Theorem 8 (Local Limit Theorem). For any $0 < p < 1$,

$$\max_{0 \leq k \leq n} \left| \mathbb{P}(S_n = k) - \frac{1}{\sqrt{2\pi np(1-p)}} e^{-\frac{x^2}{2p(1-p)}} \right| = o\left(\frac{1}{\sqrt{n}}\right),$$

where $x = x_{k,n} := \frac{k-np}{\sqrt{n}}$.

Proof. Let $A_n > 0$ such that $A_n = o(n)$ (e.g. $A_n = n^\epsilon$ for some $0 < \epsilon < 1$). Then, let $k \in \mathbb{N}$ such that $|x_{k,n}| \leq A_n/\sqrt{n}$ and so,

$$np - A_n \leq k \leq np + A_n.$$

Then, $k, n - k \rightarrow \infty$ as $n \rightarrow \infty$ and using Stirling's formula,

$$\begin{aligned} \mathbb{P}(S_n = k) &= \frac{n!}{k!(n-k)!} p^k (1-p)^{n-k} \\ &= \frac{\sqrt{2\pi n}}{\sqrt{2\pi k} \sqrt{2\pi(n-k)}} e^{n \log n - k \log k - (n-k) \log(n-k)} p^k (1-p)^{n-k} \left(1 + O\left(\frac{1}{n}\right)\right). \end{aligned}$$

Then, by noting that $np + x\sqrt{n} = k = np(1 + O(A_n/n))$ and $n(1-p) - x\sqrt{n} = n - k = n(1-p)(1 + O(A_n/n))$, we have

$$\begin{aligned} \mathbb{P}(S_n = k) &= \sqrt{\frac{n}{np2\pi n(1-p)}} \left(1 + O\left(\frac{A_n}{n}\right)\right) e^{n \log n - (np+x\sqrt{n})(\log(np) + \frac{x}{p\sqrt{n}} - \frac{x^2}{2p^2\sqrt{n}} + O(\frac{x}{\sqrt{n}})^3)} \\ &\quad e^{-(n(1-p)-x\sqrt{n})(\log(n(1-p) - \frac{x}{(1-p)\sqrt{n}} - \frac{x^2}{2(1-p)^2n} + O(\frac{x}{\sqrt{n}})^3))} e^{k \log p + (n-k) \log(1-p)} \\ &= \sqrt{\frac{1}{2\pi np(1-p)}} \left(1 + O\left(\frac{A_n}{n}\right)\right) e^{-(np+x\sqrt{n})(\frac{x}{p\sqrt{n}} - \frac{x^2}{2p^2n} + O(\frac{x}{\sqrt{n}})^3)} \\ &\quad e^{(n(1-p)-x\sqrt{n})(\frac{x}{(1-p)\sqrt{n}} - \frac{x^2}{2(1-p)^2n} + O(\frac{x}{\sqrt{n}})^3)} \\ &= \sqrt{\frac{1}{2\pi np(1-p)}} \left(1 + O\left(\frac{A_n}{n}\right)\right) \\ &\quad e^{-\sqrt{n}x + \frac{x^2}{2p} - \frac{x^2}{p} + \frac{x^3}{2p^2\sqrt{n}}} e^{O(\frac{x}{\sqrt{n}})^3 + \sqrt{n}x + \frac{x^2}{2(1-p)} - \frac{x^2}{1-p} - \frac{x^3}{2(1-p)^2\sqrt{n}}} \\ &= \frac{1 + O\left(\frac{A_n}{n}\right)}{\sqrt{2\pi np(1-p)}} e^{-\frac{x^2}{2}(\frac{1}{p} + \frac{1}{1-p}) + O(\frac{A_n^3}{\sqrt{n}^3\sqrt{n}})} \\ &= \frac{1}{\sqrt{2\pi p(1-p)n}} e^{-\frac{x^2}{2p(1-p)}} \left(1 + O\left(\frac{A_n^3}{n^2}\right) + O\left(\frac{A_n}{n}\right)\right). \end{aligned}$$

Hence, choosing $A_n = n^{7/12}$, the result follows for $np - A_n \leq k \leq np + A_n$. Finally, by monotonicity of $P(S_n = k)$, we have the result follows for all $k \leq n$. \square

Theorem 9 (de Moivre-Laplace Central Limit Theorem). For any $0 < p < 1$ and $x \in \mathbb{R}$,

$$\lim_{n \rightarrow \infty} \mathbb{P}\left(\frac{S_n - np}{\sqrt{np(1-p)}} \leq x\right) = \Phi(x),$$

where $\Phi(x)$ is the distribution of the normal random variable with parameters $m = 0, \sigma^2 = 1$.

3 Convergence and Characteristic Functions

3.1 Different Modes of Convergence

Let us recall some different notions of convergence in probability theory.

Definition 3.1. A sequence of random variables $(\xi_n)_n$ is said to converge

- almost surely if $\xi_n(\omega) \rightarrow \xi(\omega)$ almost everywhere on Ω ;
- in probability if for all $\epsilon > 0$,

$$\lim_{n \rightarrow \infty} \mathbb{P}(|\xi_n - \xi| \geq \epsilon) = 0;$$

- in L^p for $1 \leq p < \infty$ (also known as convergence in mean of order p) if

$$\lim_{n \rightarrow \infty} \mathbb{E}(|\xi_n - \xi|^p) = 0;$$

- in distribution (or weakly) if for any bounded continuous function f ,

$$\mathbb{E}(f(\xi_n)) \rightarrow \mathbb{E}(f(\xi)).$$

Namely, this condition requires laws $(\xi_n)_* \mathbb{P}$ converges weakly to $\xi_* \mathbb{P}$. Note that this does not require (ξ_n) to be defined on the same probability space.

Proposition 3.1. We have the following relations between the different notions of convergence.

- Convergence almost surely implies convergence in probability.
- Convergence in probability implies the existence of a convergent almost surely subsequence.
- Convergence in L^p implies convergence in probability.

Proof. Exercise. □

Theorem 10. Let $\xi_n \geq 0$ be a sequence of integrable random variable which converges almost surely to ξ . Then, if $\mathbb{E}\xi_n \rightarrow \mathbb{E}\xi$, we have $\xi_n \rightarrow \xi$ in L^1 .

Proof. We see

$$\mathbb{E}|\xi_n - \xi| = \mathbb{E}(\xi - \xi_n)\mathbf{1}_{\xi \geq \xi_n} + \mathbb{E}(\xi_n - \xi)\mathbf{1}_{\xi_n > \xi} = 2\mathbb{E}(\xi - \xi_n)\mathbf{1}_{\xi \geq \xi_n} + \mathbb{E}(\xi_n - \xi).$$

Thus, the result follows by considering $0 \leq (\xi - \xi_n)\mathbf{1}_{\xi \geq \xi_n} \leq \xi$, and applying dominated convergence. □

Proposition 3.2. Convergence in probability implies convergence in distribution for integrable random variables.

Proof. Let (ξ_n) be a sequence of integrable random variables which converges in probability to ξ . Fix $\epsilon > 0$ and let f be a continuous bounded function such that $|f(x)| \leq C$ for all x . By Markov's inequality, there exists some N such that

$$\mathbb{P}(|\xi| > N) \leq \frac{\epsilon}{4c}.$$

Now, as f is continuous, there exists some $\delta > 0$ such that $|f(x) - f(y)| < \epsilon/2$ for all $|x| \leq N$, $|x - y| \leq \delta$. Then,

$$\begin{aligned} \mathbb{E}(|f(\xi_n) - f(\xi)|) &= \mathbb{E}(|f(\xi) - f(\xi_n)|) \mathbf{1}_{|\xi_n - \xi| \leq \delta, |\xi| \leq n} \\ &\quad + \mathbb{E}(|f(\xi) - f(\xi_n)|) \mathbf{1}_{|\xi_n - \xi| \leq \delta, |\xi| > n} + \mathbb{E}(|f(\xi) - f(\xi_n)|) \mathbf{1}_{|\xi_n - \xi| > \delta} \\ &\leq \frac{\epsilon}{2} + 2c \frac{\epsilon}{4c} + 2c \mathbb{P}(|\xi_n - \xi| > \delta) \\ &\leq \frac{\epsilon}{2} + \frac{\epsilon}{2} + 2c \mathbb{P}(|\xi_n - \xi| > \delta) \end{aligned}$$

where the last equality follows as ξ_n converges to ξ in probability. Thus, as the ϵ is arbitrary,

$$\lim_{n \rightarrow \infty} \mathbb{E}(|f(\xi_n) - f(\xi)|) = 0.$$

□

3.2 Weak Convergence of Measures

Theorem 11 (Portmanteau Theorem). Let (ξ_n) be a sequence of integrable random variables. Then, the following are equivalent:

1. ξ_n converges in distribution to ξ .
2. $\limsup_n \mathbb{P}(\xi_n \in B) \leq \mathbb{P}(\xi \in B)$ for all closed B .
3. $\liminf_n \mathbb{P}(\xi_n \in A) \geq \mathbb{P}(\xi \in A)$ for all open A .
4. $\lim_n \mathbb{P}(\xi_n \in C) = \mathbb{P}(\xi \in C)$ for all borel C with $\mathbb{P}(\xi \in \partial C) = 0$.
5. The distribution functions F_{ξ_n} converges to F_ξ at all points of continuity of F_ξ .

Proof. (1 \implies 2). Let B be closed and let $f := \mathbf{1}_B$ and let $f_\epsilon(x) := g(\epsilon^{-1} \text{dist}(x, B))$ where

$$g(t) := \begin{cases} 1, & t \leq 0, \\ 1 - t, & 0 \leq t \leq 1, \\ 0, & t > 1. \end{cases}$$

We observe that $f_\epsilon(x)$ is 0 if x is not in a ϵ -neighbourhood of B . In particular, defining $B_\epsilon := \{x \mid \text{dist}(x, B) \leq \epsilon\}$, $f_\epsilon(B_\epsilon^c) = \{0\}$. We note that $B_\epsilon \downarrow B$ as $\epsilon \downarrow 0$ and so $f_\epsilon \rightarrow f$. Now, by considering

$$\mathbb{P}(\xi_n \in B) = \int f \, d(\xi_n)_* \mathbb{P} \leq \int f_\epsilon \, d(\xi_n)_* \mathbb{P},$$

we have

$$\limsup_n \mathbb{P}(\xi_n \in B) \leq \limsup_n \int f_\epsilon \, d(\xi_n)_* \mathbb{P} = \int f_\epsilon \, d\xi_* \mathbb{P} \leq \mathbb{P}(\xi \in B_\epsilon)$$

where the equality follows by convergence in distribution and the second inequality follows as $f_\epsilon(B_\epsilon^c) \leq 1$. Thus, by the continuity of measure, $\mathbb{P}(\xi \in B_\epsilon) \rightarrow \mathbb{P}(\xi \in B)$ as $\epsilon \downarrow 0$ and hence, $\limsup_n \mathbb{P}(\xi_n \in B) \leq \mathbb{P}(\xi \in B)$ as required.

(2 \iff 3). Follows taking the complement.

(2, 3 \implies 4). Taking any borel set C , we recall that $\overline{C} = C \cup \partial C$ and $C^\circ = C \setminus \partial C$. Then, as $\mathbb{P}(\xi \in \partial C) = 0$, we have

$$\limsup_n \mathbb{P}(\xi_n \in C) \leq \limsup_n \mathbb{P}(\xi_n \in \overline{C}) \leq \mathbb{P}(\xi \in \overline{C}) = \mathbb{P}(\xi \in C),$$

and

$$\liminf_n \mathbb{P}(\xi_n \in C) \geq \liminf_n \mathbb{P}(\xi_n \in C^\circ) \geq \mathbb{P}(\xi \in C^\circ) = \mathbb{P}(\xi \in C).$$

Hence, $\lim_n \mathbb{P}(\xi_n \in C) = \mathbb{P}(\xi \in C)$ as required.

(4 \implies 5). Obvious from 5.

(5 \implies 1). This implication is a bit more complicated. The general idea uses the fact that if 5 holds, then there exists a probability space $(\Omega', \mathcal{F}', \mathbb{P}')$ and a sequence of random variables (η_n) on this probability space such that $(\xi_n)_* \mathbb{P}_n = (\eta_n)_* \mathbb{P}'$ (I'm writing \mathbb{P}_n to especially indicate ξ_n are on possibly different measurable spaces) $\xi_n \rightarrow \eta$ almost surely.

Assuming this theorem (which will prove below), let f be a continuous bounded function. As f is continuous, $f(\eta_n) \rightarrow f(\eta)$ almost surely. Now, as they are bounded, we may apply dominated convergence such that

$$\mathbb{E}(f(\xi_n)) = \mathbb{E}(f(\eta_n)) \rightarrow \mathbb{E}(f(\eta)) = \mathbb{E}(f(\xi))$$

where the equalities follows by change of variables. \square

Lemma 3.1. Let ξ be a random variable with distribution function F_ξ and let U be a uniformly distributed random variable on the interval $[0, 1]$ where we denote

$$F^{-1}(u) := \sup\{y \mid F_\xi(y) < u\}.$$

Then, $\xi_* \mathbb{P} = (F^{-1}(U))_* \lambda$.

Proof. This follows as $u \leq F_\xi(x) \iff F_\xi^{-1}(u) \leq x$ since F_ξ is nondecreasing and continuous from the right. Indeed, if $u \leq F_\xi(x)$ and suppose for contradiction that $F_\xi^{-1}(u) > x$, then, by definition, $\sup\{y \mid F_\xi(y) < u\} > x$ and so, $\sup\{y \mid F_\xi(y) < u \leq F_\xi(x)\} > x$. However, this is not possible as F_ξ is nondecreasing, thus a contradiction. On the other hand, if $x \leq F_\xi^{-1}(u) = \sup\{y \mid F_\xi(y) < u\}$, as F_ξ is nondecreasing and continuous from the right, $F_\xi(x) \geq F_\xi(\sup\{y \mid F_\xi(y) < u\}) \geq u$. \square

Theorem 12 (Single Probability Space Theorem). Let (ξ_n) be a sequence of integrable random variables such that the distribution functions F_{ξ_n} converges to F_ξ at all points of continuity of F_ξ . Then, there exists a probability space $(\Omega', \mathcal{F}', \mathbb{P}')$ and a sequence of random variables (η_n) on this probability space such that $(\xi_n)_* \mathbb{P}_n = (\eta_n)_* \mathbb{P}'$, $\xi_* \mathbb{P} = \eta_* \mathbb{P}'$ and $\eta_n \rightarrow \eta$ almost surely.

Proof. Let $(\Omega', \mathcal{F}', \mathbb{P}') := ([0, 1], \mathcal{B}([0, 1]), \lambda)$. Then, it suffices to prove that $F_{\xi_n}^{-1}(u) \rightarrow F_{\xi}^{-1}(u)$ for all u such that $|F_{\xi}^{-1}(u)| < \infty$ and the measure of all other points is zero as they form a countable set (nondecreasing function can have at most countably many discontinuities). Indeed, defining $\eta_n := F_{\xi_n}^{-1}(U)$ converges to $\eta := F_{\xi}^{-1}(U)$ almost surely. Then, by the above lemma, the result follows.

Suppose u is a point such that $|F_{\xi}^{-1}(u)| < \infty$. Then, for any $x < F_{\xi}^{-1}(u)$, $F(x) < u$ and if x is a point of continuity of F_{ξ} , we have $F_{\xi_n}(x) \rightarrow F_{\xi}(x)$ and so, $F_{\xi_n} < u$ for sufficiently large n . Thus, $x \leq F_{\xi_n}^{-1}(u)$ implying $x \leq \liminf_n F_{\xi_n}^{-1}(u)$. Now, by taking a sequence (x_n) of points of continuity of F_{ξ} such that $x_n \uparrow F^{-1}(u)$, we obtain $F^{-1}(u) \leq \liminf_n F_{\xi_n}^{-1}(u)$.

On the other hand, if $x > F_{\xi}^{-1}(u)$, we have $F(x) \geq u$. If $F_{\xi}(x) = u$, then we see that $|F_{\xi}^{-1}(u)| = \infty$. So, $F_{\xi}(x) > u$ and from which by the same argument as the less than case, we obtain $F^{-1}(u) \geq \limsup_n F_{\xi_n}^{-1}(u)$ implying convergence as required. \square

Definition 3.2 (Relatively Compact). A family of probability measures $\mathcal{P} := \{\mathbb{P}_{\alpha} \mid \alpha \in A\}$ and the corresponding set of distribution functions F_{α} is called relatively compact if every sequence of measures of \mathcal{P} has a subsequence which converges weakly (not necessary in \mathcal{P} hence the name “relatively”).

We note that the weak convergence of measures is metrizable with the corresponding metric known as the Levy-Prokhorov metric. In this sense, relatively compactness of \mathcal{P} is equivalent to $\overline{\mathcal{P}}$ is compact in this metric space.

Denote \mathcal{G} the set of functions $F : \mathbb{R} \rightarrow [0, 1]$ which is non-decreasing and continuous from the right (and we call functions of this form generalised distribution functions), we have the following result.

Theorem 13 (Helly’s Theorem). The set \mathcal{G} is sequentially compact, i.e. any sequence $(F_n) \subseteq \mathcal{G}$ has a subsequence (F_{n_k}) such that $F_{n_k}(x) \rightarrow F(x)$ for all $x \in C_F$.

Proof. Let us enumerate \mathbb{Q} with $(q_i)_{i=1}^{\infty}$. Then, as $\{F_n(q_1) \mid n\}$ is bounded, it has a convergent subsequence $F_{n_k^{(1)}}(q_1) \rightarrow f(q_1)$. Similarly, as $\{F_{n_k^{(1)}}(q_2) \mid n\}$ is bounded, we can find a subsequence $F_{n_k^{(1)}}, F_{n_k^{(2)}}$ which converges for both q_1, q_2 . Then, iterating this process, for all i , we obtain a subsequence of $F_{n_k^{(i-1)}}$ which converges at q_1, \dots, q_i . Thus, by defining $F_{n_k} := F_{n_k^{(k)}}$, i.e. the diagonal, we have found a subsequence of F_n which converges at all rational numbers. Let us denote the limit of $F_{n_k}(q_i)$ as $f(q_i)$.

Now, defining

$$F(x) := \inf\{f(q) \mid q \in \mathbb{Q}_{>x}\},$$

I claim that $F \in \mathcal{G}$, $F(q) \geq f(q)$ for all $q \in \mathbb{Q}$ and $F(x) \leq f(q)$ for all $x < q$, $q \in \mathbb{Q}$. Indeed, the last inequality is clear by construction while the second inequality follows as f is non-decreasing. Thus, by the definition of \inf , F is continuous from the right and so, $F \in \mathcal{G}$.

Finally, it remains to check that $F_{n_k}(x) \rightarrow F(x)$ at all points of continuity. Let x be a point of continuity and fix $\epsilon > 0$. As x is a point of continuity, there exists some $y < p < x < q$, such that

$$F(x) - \epsilon < F(y) \leq F(p) \leq F(x) \leq F(q) < F(x) + \epsilon.$$

Then, as $F(y) \leq f(p)$ for all $y < p$, $F(x) - \epsilon < f(p)$. On the other hand, $F(p) \geq f(p)$, so

$$F(x) - \epsilon < f(p) \leq F(p) \leq F(x) \leq f(q) < F(x) + \epsilon.$$

Thus, for sufficiently large k ,

$$F(x) - \epsilon < F_{n_k}(p) \leq F_{n_k}(x) \leq F_{n_k}(q) < F(x) + \epsilon,$$

implying $|F_{n_k}(x) - F(x)| < \epsilon$ and hence, $F_{n_k}(x) \rightarrow F(x)$ as required. \square

Definition 3.3 (Tight). A family of probability measures $\mathcal{P} := \{\mathbb{P}_\alpha \mid \alpha \in A\}$ on $(\mathbb{R}, \mathcal{B}(\mathbb{R}))$ is called tight if for all $\epsilon > 0$, there exists a compact set $K \subseteq \mathbb{R}$ such that

$$\sup_{\alpha \in A} \mathbb{P}_\alpha(K^c) \leq \epsilon.$$

Theorem 14 (Prokhorov's Theorem). A family of probability measures $\mathcal{P} := \{\mathbb{P}_\alpha \mid \alpha \in A\}$ is tight if and only if it is relatively compact.

Proof. Suppose first \mathcal{P} is relatively compact and suppose for contradiction it is not tight. That is, there exists some $\epsilon > 0$, for all compact $K \subseteq \mathbb{R}$,

$$\sup_{\alpha \in A} \mathbb{P}_\alpha(K^c) > \epsilon.$$

In particular, we may choose $K_n := [-n, n]$ for any n . Thus, for any n , there exists some $\mathbb{P}_{\alpha_n} \in \mathcal{P}$ such that $\mathbb{P}_{\alpha_n}(K_n^c) > \epsilon$. Now, since \mathcal{P} is relatively compact, it has a weakly convergent subsequence $\mathbb{P}_{\alpha_{n_k}} \rightarrow \mathbb{P}$ for some probability measure \mathbb{P} . As \mathbb{P}_{α_n} are probability measures, we have for all $m > k$,

$$\mathbb{P}_{\alpha_m}([-k, k]) \leq \mathbb{P}_{\alpha_m}([-m, m]) < 1 - \epsilon.$$

So, by weak convergence,

$$\mathbb{P}([-k, k]) = \int \mathbf{1}_{[-k, k]} d\mathbb{P} = \lim_{m \rightarrow \infty} \int \mathbf{1}_{[-k, k]} d\mathbb{P}_{\alpha_{n_m}} < 1 - \epsilon.$$

On the other hand, as \mathbb{P} is a probability measure,

$$1 = \mathbb{P}(\mathbb{R}) = \mathbb{P}\left(\bigcup_{k \in \mathbb{N}} [-k, k]\right) = \lim_{k \rightarrow \infty} \mathbb{P}([-k, k]) \leq 1 - \epsilon < 1,$$

which is a contradiction. Hence, \mathcal{P} is tight.

Suppose now \mathcal{P} is tight and let $(\mathbb{P}_{n_k}) \subseteq \mathcal{P}$ be a sequence of probability measures with corresponding distribution functions (F_{n_k}) . By Helly's theorem, there exists a subsequence (F_{n_k}) which converges to some $F \in \mathcal{G}$ at all points of continuity of F . It remains to show $\lim_{x \rightarrow -\infty} F(x) = 0$ and $\lim_{x \rightarrow \infty} F(x) = 1$.

Fix $\epsilon > 0$. Then, by tightness, there exists some compact K such that $\sup \mathbb{P}_{n_k}(K^c) \leq \epsilon$. Now, since $K \subseteq \mathbb{R}$ it is bounded by some M (choose M such that $\pm M$ is a point of continuity of F) and so, $K \subseteq [-M, M]$ and

$$\sup F_{n_k}(-M) = \sup \mathbb{P}_{n_k}((-\infty, -M]) \leq \sup \mathbb{P}_{n_k}([-M, M]^c) \leq \sup \mathbb{P}_{n_k}(K^c) \leq \epsilon.$$

Thus, as $F(M) \leq \epsilon$, and as it is non-decreasing, $\lim_{x \rightarrow -\infty} F(x) = 0$. Similarly, we have

$$\inf F_{n_k}(M) = \inf \mathbb{P}_{n_k}((-\infty, M]) \geq \inf \mathbb{P}_{n_k}([-M, M]) \geq \inf \mathbb{P}_{n_k}(K) \geq 1 - \epsilon.$$

Thus, $\lim_{x \rightarrow \infty} F(x) = 1$. Hence, \mathbb{P}_{n_k} converges to the probability measure corresponding to \mathbb{P} weakly by Portmanteau's theorem as required. \square

Corollary 14.1. If F_{ξ_n} is a sequence of distribution functions which tends to the distribution function F at all points of continuity of F , then $\{\xi_n\}$ is tight.

To check F is indeed a distribution function, by recalling Helly's theorem, it suffices to check $\lim_{x \rightarrow -\infty} F(x) = 0$ and $\lim_{x \rightarrow \infty} F(x) = 1$.

Prokhorov's theorem remains true for in a much general context. Namely, Prokhorov's theorem holds for measures on $\mathbb{R}^n, \mathbb{R}^\infty$ and measures on any complete separable metric space with a Borel σ -algebra.

4 Characteristic Functions

Definition 4.1 (Characteristic Function). The characteristic function of a random variable ξ is

$$\phi_\xi(t) := \mathbb{E}e^{it\xi} = \int e^{itx} dF_\xi(x),$$

which is the Fourier-Stieltjes transform of the distribution F_ξ .

If $\xi = (\xi_1, \dots, \xi_n)$ is a random vector, then the characteristic function of ξ is

$$\phi_\xi(t_1, \dots, t_n) := \mathbb{E}e^{i\sum_{k=1}^n t_k \xi_k}.$$

By the very definition of the characteristic function, we observe the following properties of the characteristic function.

- If ξ is a random variable and a, b are constants, then defining $\eta := a\xi + b$, we have

$$\phi_\eta(t) = e^{itb} \mathbb{E}e^{iat\xi}.$$

- $|\phi(t)| \leq \phi(0) = 1$.
- If ξ is a random variable, then ϕ_ξ is uniformly continuous on \mathbb{R} . Indeed, we observe

$$|\phi(t+h) - \phi(t)| = |\mathbb{E}e^{it\xi}(e^{ih\xi} - 1)| \leq \mathbb{E}|e^{ih\xi} - 1|$$

where the last term is independent of t and tends to 0 as $h \rightarrow 0$ by dominated convergence.

- If ξ_1, \dots, ξ_n are independent random variables, and we define $S := \xi_1 + \dots + \xi_n$, then

$$\phi_S(t) = \prod_{k=1}^n \phi_{\xi_k}(t).$$

Proposition 4.1. If $\xi \sim \mathcal{N}(m, \sigma^2)$, then $\phi_\xi(t) = e^{itm - \frac{t^2 \sigma^2}{2}}$.

Proof. Let $\eta = (\xi - m)/\sigma$ so that $\eta \sim \mathcal{N}(0, 1)$. Then, $\xi = \sigma\eta + m$ and hence, by the previous result,

$$\phi_\xi(t) = e^{itm} \mathbb{E}e^{i\sigma t\eta}.$$

Thus, it is sufficient to show $\phi_\eta(t) = e^{-t^2/2}$. Indeed, we compute

$$\begin{aligned} \phi_\eta(t) &= \mathbb{E}e^{it\eta} = \frac{1}{\sqrt{2\pi}} \int e^{itx} f_\eta(x) \lambda(dx) = \frac{1}{\sqrt{2\pi}} \int e^{itx - \frac{x^2}{2}} \lambda(dx) \\ &= e^{-\frac{t^2}{2}} \frac{1}{\sqrt{2\pi}} \int e^{-\frac{1}{2}(x-it)^2} \lambda(dx) = e^{-\frac{t^2}{2}} \frac{1}{\sqrt{2\pi}} \int_{-\infty-it}^{\infty+it} e^{-\frac{z^2}{2}} \lambda(dz). \end{aligned}$$

Then, by contour integration, we find $\int_{-\infty-it}^{\infty+it} e^{-\frac{z^2}{2}} \lambda(dz) = \sqrt{2\pi}$ and so, $\phi_\eta(t) = e^{-t^2/2}$ as required. \square

4.1 Moments and Inversion

Theorem 15. Let ξ be a random variable with characteristic function ϕ_ξ and distribution F_ξ . Then, if

- $\mathbb{E}|\xi|^n < \infty$ for some $n \geq 1$, then $\phi_\xi^{(r)}$ exists for any $0 \leq r \leq n$ and is given by

$$\phi_\xi^{(r)}(t) = \int_{\mathbb{R}} (ix)^r e^{ixt} dF_\xi(x).$$

Furthermore, $\mathbb{E}\xi^r = i^{-r} \phi_\xi^{(r)}(0)$. By Taylor's theorem,

$$\phi(t) = \sum_{r=0}^{n-1} \frac{(it)^r}{r!} \mathbb{E}\xi^r + \frac{it^n}{n!} \epsilon_n(t)$$

where $|\epsilon_n(t)| \leq 3\mathbb{E}|\xi|^n$ and $\epsilon_n(t) \rightarrow 0$ as $t \rightarrow 0$.

- $\mathbb{E}|\xi|^n < \infty$ for all $n \geq 1$ and

$$\limsup_{n \rightarrow \infty} \frac{(\mathbb{E}|\xi|^n)^{1/n}}{n} = \frac{1}{eT} < \infty$$

for some T (note that $(\mathbb{E}|\xi|^n)^{1/n} = \|\xi\|_n$), then

$$\phi(t) = \sum_{n=0}^{\infty} \frac{(it)^n}{n!} \mathbb{E}\xi^n$$

for all $|t| < T$.

Proof. By noting that $\mathbb{E}|\xi|^n = \|\xi\|_n^n$, and $L_p(\Omega) \subseteq L_q(\Omega)$ for $q \leq p$ and Ω a finite measure space, $\mathbb{E}|\xi|^n < \infty$ implies $\xi \in L_p$ for all $p \leq n$ and thus, $\mathbb{E}|\xi|^p < \infty$ for all $p \leq n$. By noting that

$$\left| \frac{e^{ihx} - 1}{h} \right| \leq x$$

by dominated convergence and l'Hopital's rule,

$$\begin{aligned} \phi'_\xi(t) &= \lim_{h \rightarrow 0} \frac{\phi(t+h) - \phi(t)}{h} = \lim_{h \rightarrow 0} \mathbb{E} \left(e^{it\xi} \frac{e^{ih\xi} - 1}{h} \right) = i\mathbb{E}\xi e^{it\xi} \\ &= i \int_{\mathbb{R}} x e^{itx} dF_\xi(x). \end{aligned}$$

With this, we obtain the r -th derivative by induction.

By Taylor's theorem, $e^{it\xi} = \sum_{k=0}^{n-1} \frac{(it\xi)^k}{k!} + \frac{(it\xi)^n}{n!} (\cos t\xi\theta_1(\omega) + i \sin t\xi\theta_2(\omega))$. Thus, as the sum is finite, by the linearity of expectation,

$$\mathbb{E}e^{it\xi} = \sum_{k=0}^{n-1} \frac{(it)^k}{k!} \mathbb{E}\xi^k + \frac{(it)^n}{n!} (\mathbb{E}\xi^n + \epsilon_n(t))$$

where $\epsilon_n(t) = \mathbb{E}(\xi^n (\cos t\xi\theta_1 + i \sin t\xi\theta_2 - 1))$. Thus, $|\epsilon_n(t)| \leq \mathbb{E}|3\xi^n| = 3\mathbb{E}|\xi|^n$ and $\epsilon_n(t) \rightarrow 0$ as $t \rightarrow 0$ by dominated convergence.

Let $0 < t_0 < T$. Then,

$$\limsup_{n \rightarrow \infty} \frac{(\mathbb{E}|\xi|^n)^{1/n}}{n} < \frac{1}{et_0}$$

implying

$$\limsup_{n \rightarrow \infty} \left(\frac{\mathbb{E}|\xi|^n}{n^n} t_0^n e^n \right)^{1/n} < 1.$$

We note that this formula is very similar to Stirling's formula and

$$\limsup_{n \rightarrow \infty} \left(\frac{\mathbb{E}|\xi|^n t_0^n}{n!} \right)^{1/n} < L < 1$$

and so $\sum_{n=0}^{\infty} \frac{\mathbb{E}|\xi|^n}{n!} t_0^n$ converges by the root test. \square

With this theorem in mind, we see that if $\mathbb{E}|\xi|^n < \infty$ exists for all $n \geq 1$, the characteristic function ϕ_ξ is uniquely determined on $[-T, T]$ by the moments of ξ . Furthermore, taking $s = T/2$, by the same proof, we obtain

$$\phi_\xi(t) = \sum_{k=0}^{\infty} i^k \frac{(t-s)^k}{k!} \phi_\xi^{(k)}(s)$$

where $\phi_\xi^{(k)}(s) = \mathbb{E}\xi^k e^{is\xi}$ for $-T/2 < s < T/2$. Thus, ϕ_ξ is uniquely determined by moments on $3/2T$. Continuing this process, we find $\phi_\xi(t)$ is uniquely determined by moments for all t .

A stronger sufficient condition for the unique determination of ϕ_ξ is that

$$\sum_{n=0}^{\infty} \frac{1}{(\mathbb{E}\xi^{2n})^{1/2n}} = \infty.$$

This is called Carleman's test. We will omit the proof. In some sense, if $\mathbb{E}\xi^n$ grows too fast, there may be multiple $\phi(t)$ with these moments.

Theorem 16 (Inversion). Let ξ be a random variable with characteristic function ϕ_ξ and distribution F_ξ . Then, if $a < b$ are points of continuity of F_ξ ,

$$F_\xi(b) - F_\xi(a) = \lim_{T \rightarrow \infty} \frac{1}{2\pi} \int_{[-T, T]} \frac{e^{-ita} - e^{-itb}}{it} \phi_\xi(t) \lambda(dt).$$

Furthermore, if $\int |\phi| d\lambda < \infty$, then F_ξ is absolutely continuous with density f_ξ such that

$$f_\xi(x) = \frac{1}{2\pi} \int_{\mathbb{R}} e^{-itx} \phi_\xi(t) \lambda(dt).$$

By recalling that the characteristic function is the Fourier-Stieltjes transform of F_ξ , this is simply the Fourier-Stieltjes inverse formula.

Proof. If F_ξ is absolutely continuous, the Lebesgue-Stieltjes measure λ_{F_ξ} is absolutely continuous with respect to the Lebesgue measure λ . Thus, by Radon-Nikodym, there exists some f such that $\lambda_{F_\xi} = f\lambda$. Thus,

$$\phi_\xi(t) = \int_{\mathbb{R}} e^{itx} f(x) \lambda(dx)$$

implying $\int |\phi_\xi| d\lambda = \int |f| d\lambda = 1 < \infty$ and so ϕ_ξ is integrable. Then, by the inverse Fourier transform, provided f satisfy some regularity conditions (Dini's condition, c.f. Fourier analysis course notes)

$$f(x) = \frac{1}{2\pi} \int_{\mathbb{R}} e^{-itx} \phi_\xi(t) \lambda(dt).$$

Hence, integrating this and applying Fubini's theorem,

$$F(b) - F(a) = \int_{(a,b]} f d\lambda = \frac{1}{2\pi} \int_{(a,b] \times \mathbb{R}} e^{-itx} \phi_\xi(t) \lambda(dt) \lambda(dx) = \frac{1}{2\pi} \int_{\mathbb{R}} \frac{e^{-ita} - e^{-itb}}{it} \phi_\xi(t) \lambda(dt).$$

It remains to show that f satisfy Dini's condition which we shall omit.

Now, let us consider the general case. Define

$$\phi_T := \frac{1}{2\pi} \int_{[-T,T]} \frac{e^{-ita} - e^{-itb}}{it} \phi(t) \lambda(dt) = \int \psi_T(x) dF_\xi(dx)$$

where $\psi_T(x) := \frac{1}{2\pi} \int_{[-T,T]} \frac{e^{-ita} - e^{-itb}}{it} e^{itx} \lambda(dt)$, where the last equality follows by Fubini's theorem. Now, by noting

$$\begin{aligned} \psi_T(x) &= \frac{1}{2\pi} \int_{[-T,T]} \frac{\sin(t(x-a)) - \sin(t(x-b))}{t} \lambda(dt) \\ &= \frac{1}{2\pi} \left(\int_{-T(x-a)}^{T(x-a)} - \int_{-T(x-b)}^{T(x-b)} \right) \frac{\sin u}{u} \lambda(du), \end{aligned}$$

and $\int_{[-R,R]} \frac{\sin u}{u} \lambda(du) \rightarrow \pi$ as $R \rightarrow \infty$, there exists some $c > 0$, such that $|\psi_T(x)| < c$ for all x . Furthermore, by observing that

$$\psi_T(x) \rightarrow \psi(x) = \begin{cases} 0, & x \notin [a, b], \\ 1/2, & x \in \{a, b\}, \\ 1, & x \in (a, b), \end{cases}$$

by dominated convergence, as $T \rightarrow \infty$,

$$\phi_T = \int \psi_T dF \rightarrow \int \psi dF = F(b-) - F(a) + \frac{1}{2}(F(a) - F(a-) + F(b) - F(b-)).$$

Finally, as a, b are points of continuity, $\phi_T \rightarrow F(b) - F(a)$ as claimed.

Now, if $\int |\phi| d\lambda < \infty$, the function

$$f(x) := \frac{1}{2\pi} \int e^{-itx} \phi(t) \lambda(dt).$$

exists and by dominated convergence, is continuous, differentiable (and is integrable on $[a, b]$). Then, by Fubini's theorem,

$$\int_{[a,b]} f d\lambda = \int \frac{1}{2\pi} \phi(t) \frac{e^{-ita} - e^{-itb}}{it} \lambda(dt) = F(b) - F(a),$$

where $F(x) = \int_{(-\infty, x]} f d\lambda$. Thus, F is absolutely continuous and by the arguments above, we conclude the proof. \square

Corollary 16.1. Probability distributions on $(\mathbb{R}, \mathcal{B}(\mathbb{R}))$ and characteristic functions are in 1-to-1 correspondence.

Proof. Since points of continuity of a distribution function is dense, by the continuity of measure to define $F(b) - F(a)$ uniquely for any a, b from any characteristic functions. \square

We will now provide some useful theorems without proof.

Theorem 17 (Bochner). Let ϕ be continuous on \mathbb{R} such that $\phi(0) = 1$. Then, ϕ is a characteristic function if and only if ϕ is positive semi-definite, i.e.

$$\sum_{k,j=1}^n \phi(t_j - t_k) \alpha_j \bar{\alpha}_k$$

for all $t_1, \dots, t_n \in \mathbb{R}$ and $\alpha_1, \dots, \alpha_n \in \mathbb{C}$.

Theorem 18 (Polya). Let $\phi \geq 0$ be continuous, even, $\phi(0) = 1, \lim_{x \rightarrow \infty} \phi(x) = 0$ and is convex on $0 \leq t < \infty$. Then ϕ is a characteristic function.

An example of the above is $e^{-|t|}$.

Theorem 19 (Marcinkiewicz). If a characteristic function is of the form $e^{p(t)}$ where $p(t)$ is a polynomial, then p has degree at most 2.

Definition 4.2 (Cumulant). If there exists an expansion such that

$$\log \phi_\xi(t) = \sum_{k=0}^n \frac{(it)^k}{k!} S_k + o_{t \rightarrow 0}(|t|^n),$$

then the coefficients S_k are called cumulants of ξ .

As an exercise, one can show $\mathbb{E}\xi = S_1$ and $V_\xi = S_2$.

In general, by the Marcinkiewicz theorem, if there ξ is a random variable such that $S_k = 0$ for all $k \geq n$ for some n , then, $S_k = 0$ for all $k \geq 3$. One may show that if ξ is normal, then this condition is satisfied and thus, as this determines the characteristic function which in turn determines the distribution, if there exists some n such that $S_k = 0$ for all $k \geq n$, then $\xi \sim \mathcal{N}(S_1, S_2)$.

Theorem 20. Let ϕ be a characteristic function of ξ such that $|\phi(t_0)| = 1$ for some $t_0 \neq 0$. Then, ξ is a pure point random variable.

Proof. As $|\phi(t_0)| = 1$, there exists some $a \in \mathbb{R}$ such that $\phi(t_0) = e^{iat_0}$. Then, by the definition of characteristic functions,

$$\int e^{it_0 x} dF_\xi(x) = e^{iat_0}.$$

Thus, by comparing the real parts, we obtain

$$1 = \int \cos(t_0(x - a)) dF_\xi(x)$$

and so $\int 1 - \cos(t_0(x - a)) dF_\xi(x) = 0$. But $1 - \cos(t_0(x - a)) \geq 0$ implying $\cos(t_0(x - a)) = 1$ almost everywhere with respect to λ_{F_ξ} . Thus, λ_{F_ξ} is concentrated at points of the form $a + \frac{2\pi}{t_0}n$ for $n \in \mathbb{Z}$. \square

4.2 Central Limit Theorem

We had previously prove the central limit theorem for Bernoulli random variables. However, as we know, the central limit theorem applies to a much wider class of random variables. We shall in this section prove the general central limit theorem using the characteristic functions.

Theorem 21 (Levy's Continuity Theorem). Let ϕ_n be a sequence of characteristic functions of distribution functions F_n . Then,

- if $F_n \rightarrow F$ at all points of continuity (i.e. the corresponding random variables converges weakly) where F is a distribution function, then $\phi_n \rightarrow \phi$ point-wise where ϕ is the characteristic function of F , i.e.

$$\phi(t) = \int e^{itx} dF(x).$$

- if $\lim_{n \rightarrow \infty} \phi_n(t) = \phi(t)$ exists and ϕ is continuous at 0, then ϕ is a characteristic function of some distribution F and $F_n \rightarrow F$ at all points of continuity.
- if ϕ_n corresponds to F_n and ϕ corresponds to F , then $\phi_n \rightarrow \phi$ point-wise if and only if $F_n \rightarrow F$ at all points of continuity.

Proof. The first statement is obvious by the very definition of weak convergence. While the third statement is implied by the first two. We omit the second. \square

Theorem 22 (Central Limit Theorem for i.i.d. Random Variables). Let (ξ_n) be independent and identically distributed random variables such that $\mathbb{E}\xi_1^2 < \infty$ and nondegenerate (i.e. $V_\xi \neq 0$). Then, defining $S_n = \sum_{i=1}^n \xi_n$,

$$\mathbb{P} \left(\frac{S_n - \mathbb{E}S_n}{\sqrt{V_{S_n}}} \leq x \right) \rightarrow \frac{1}{\sqrt{2\pi}} \int_{(-\infty, x]} e^{-y^2/2} \lambda(dy)$$

as $n \rightarrow \infty$ for all $x \in \mathbb{R}$.

Namely, the central limit theorem states that $\frac{S_n - \mathbb{E}S_n}{\sqrt{V_{S_n}}}$ converges in distribution to $\mathcal{N}(0, 1)$.

Proof. Suppose $m = \mathbb{E}\xi$ and $\sigma^2 = V_\xi$. Then, setting $\phi(t) = \mathbb{E}e^{it(\xi-m)}$ and $\phi_n(t) = \mathbb{E}e^{it\frac{S_n - \mathbb{E}S_n}{\sqrt{V_{S_n}}}}$. Then, by independence,

$$\phi_n(t) = \left(\phi \left(\frac{t}{\sigma\sqrt{n}} \right) \right)^n.$$

As $\mathbb{E}\xi^2 < \infty$, $\phi(t) = 1 - \frac{\sigma^2 t^2}{2} + o(t^2)$. So, $\phi_n(t) = \left(1 - \frac{t^2}{2n} + o\left(\frac{1}{2n}\right)\right)^n$ which tends to $e^{-t^2/2}$ as $n \rightarrow \infty$. Thus, as this is the characteristic function of the standard normal distribution, by the continuity theorem, we obtain the result. \square

Theorem 23 (Central Limit Theorem for Independent Random Variables). Let (ξ_n) be independent with $\mathbb{E}\xi_i^2 < \infty$ and distribution function F_i for all i . Then, denoting $m_i = \mathbb{E}\xi_i$, $\sigma_i^2 = V_{\xi_i}$, $S_n = \sum_{i=1}^n \xi_n$ and $D_n^2 = \sum_{i=1}^n \sigma_i^2$. Then, if the Lindenberg condition holds, namely for all $\epsilon > 0$, $\frac{1}{D_n^2} \sum_{k=1}^n \int_{\{|x-m_k| \geq \epsilon D_n\}} (x - m_k)^2 dF_k(x)$ which tends to 0 as $n \rightarrow \infty$. Then,

$$\frac{S_n - \mathbb{E}S_n}{\sqrt{V_{S_n}}} \rightarrow \mathcal{N}(0, 1)$$

in distribution.

Let us note some conditions which implies the Lindenberg condition.

- Lyapunov condition: there exists some $\delta > 0$ such that

$$\frac{1}{D_n^{2+\delta}} \sum_{k=1}^n (\mathbb{E}|\xi_k - m_k|^{2+\delta}) \rightarrow 0$$

as $n \rightarrow \infty$.

Proof. Indeed, for all $\epsilon > 0$,

$$\begin{aligned} \mathbb{E}|\xi_k - m_k|^{2+\delta} &= \int |x - m_k|^{2+\delta} dF_k(x) \\ &\geq (\epsilon D_n)^\delta \int_{\{|x-m_k| \geq \epsilon D_n\}} |x - m_k|^2 dF_k(x). \end{aligned}$$

Thus,

$$\frac{1}{D_n^2} \sum_{k=1}^n \int_{\{|x-m_k| \geq \epsilon D_n\}} (x - m_k)^2 dF_k(x) \leq \frac{1}{\epsilon^\delta D_n^{2+\delta}} \sum_{k=1}^n \mathbb{E}|\xi_k - m_k|^{2+\delta} \rightarrow 0$$

allowing us to conclude. \square

- There exists some $K > 0$ such that $|\xi_k| \leq K$ for all k and $D_n \rightarrow \infty$ as $n \rightarrow \infty$.

Proof. Exercise. \square

We will now consider the rate of convergence.

Theorem 24 (Berry-Esseen Inequality). Let (ξ_k) be i.i.d. random variables with $\mathbb{E}|\xi_1|^3 < \infty$. Then

$$\sup_x \left| \mathbb{P} \left(\frac{S_n - \mathbb{E}S_n}{\sqrt{V_{S_n}}} \leq x \right) - \Phi(x) \right| \leq \frac{c \mathbb{E}|\xi_1 - \mathbb{E}\xi_1|^3}{\sigma^3 \sqrt{n}}$$

where Φ is the distribution function of a standard normal distribution and c is a constant such that $\frac{1}{\sqrt{2\pi}} \leq c \leq 1/2$.

Proof. Omitted. □

We remark that the decay $o(1/\sqrt{n})$ is the optimal rate of convergence by considering a sequence of Bernoulli random variables taking values in ± 1 .

We note that we required the random variables to have finite second moment in the statement of CLT. This condition is necessary. Indeed, if (ξ_n) is a sequence of i.i.d. Cauchy random variables with density

$$f(x) = \frac{\theta}{\pi(x^2 + \theta^2)}, \theta > 0.$$

It is easy to show that $\mathbb{E}\xi^2 = \infty$ and

$$\phi_\xi(t) = \frac{\theta}{\pi} \int \frac{e^{itx}}{x^2 + \theta^2} \lambda(dx) = e^{-\theta|t|}.$$

Thus, by independence,

$$\phi_{S_n/n}(t) = (e^{-\theta|t|/n})^n = e^{-\theta|t|} = \phi_\xi(t).$$

Hence, S_n/n also has the Cauchy distribution!

5 Results in Probability Theory

We will in this section consider some important results in probability theory. In particular, we will study the Borel-Cantelli lemmas, Kolmogorov's inequality and the strong law of large numbers.

5.1 Borel-Cantelli Lemmas

Theorem 25 (First Borel-Cantelli Lemma). Let (A_n) be a sequence of measurable sets. Then, if $\sum_{n=1}^{\infty} \mathbb{P}(A_n) < \infty$, we have

$$\mathbb{P}(\{A_n \text{ i.o.}\}) = \mathbb{P}\left(\limsup_{n \rightarrow \infty} A_n\right) = \mathbb{P}\left(\bigcap_{n=0}^{\infty} \bigcup_{k \geq n} A_k\right) = 0.$$

Proof. Define $B_n := \bigcup_{k \geq n} A_k$ so that B_n is decreasing. Then, by continuity,

$$\mathbb{P}\left(\bigcap_{n=0}^{\infty} B_n\right) = \lim_{n \rightarrow \infty} \mathbb{P}(B_n) = \lim_{n \rightarrow \infty} \mathbb{P}\left(\bigcup_{k \geq n} A_k\right) \leq \lim_{n \rightarrow \infty} \sum_{k \geq n} \mathbb{P}(A_k) = 0.$$

Hence, as $\bigcap_{n=0}^{\infty} B_n = \{A_n \text{ i.o.}\}$, we have the required result. \square

Theorem 26 (Second Borel-Cantelli Lemma). Let (A_n) be a sequence of **independent** measurable sets. Then, if $\sum_{n=1}^{\infty} \mathbb{P}(A_n) = \infty$, we have

$$\mathbb{P}(\{A_n \text{ i.o.}\}) = \mathbb{P}\left(\limsup_{n \rightarrow \infty} A_n\right) = \mathbb{P}\left(\bigcap_{n=0}^{\infty} \bigcup_{k \geq n} A_k\right) = 1.$$

Proof. Consider $\{A_n \text{ i.o.}\}^c = \bigcup_{n=1}^{\infty} \bigcap_{k \geq n} A_k^c$, so by continuity and independence,

$$1 - \mathbb{P}(\{A_n \text{ i.o.}\}) = \lim_{n \rightarrow \infty} \mathbb{P}\left(\bigcap_{k \geq n} A_k^c\right) = \lim_{n \rightarrow \infty} \prod_{k \geq n} \mathbb{P}(A_k^c).$$

Now, by noting $\log(1-x) \leq -x$ for all $x \in [0, 1)$, we have

$$\log \mathbb{P}\left(\bigcap_{k \geq n} A_k^c\right) = \log \prod_{k \geq n} (1 - \mathbb{P}(A_k)) \leq - \sum_{k \geq n} \mathbb{P}(A_k)$$

where $\sum_{k \geq n} \mathbb{P}(A_k) = \infty$ for all n .

Hence, $\log \mathbb{P}\left(\bigcap_{k \geq n} A_k^c\right) = -\infty$ and thus, $1 - \mathbb{P}(\{A_n \text{ i.o.}\}) = 0$ as required. \square

We will provide a new proof for the existence of a subsequence which converges almost everywhere from a sequence which converges in probability.

Corollary 26.1. If $\xi_n \rightarrow \xi$ in probability then there exists a subsequence $(\xi_{n_k}) \subseteq (\xi_n)$ which converges almost everywhere to ξ .

Proof. Since $\lim_{n \rightarrow \infty} \mathbb{P}(|\xi_n - \xi| > 1/k) = 0$ for all k , there exists a subsequence (ξ_{n_k}) of (ξ_n) such that

$$\mathbb{P}(|\xi_{n_k} - \xi| > 1/k) \leq 2^{-k}.$$

Hence, as $\sum_{k=1}^{\infty} \mathbb{P}(|\xi_{n_k} - \xi| > 1/k) \leq \sum_{k=1}^{\infty} 2^{-k} = 1 < \infty$, the first Borel-Cantelli lemma implies $\mathbb{P}(\{|\xi_{n_k} - \xi| > 1/k \text{ i.o.}\}) = 0$ which is exactly the set of elements which do not converge. \square

Corollary 26.2. If (ξ_n) is a sequence of random variables such that $\xi_1 \geq \xi_2 \geq \dots \geq 0$ and $\xi_n \rightarrow 0$ in probability. Then, ξ_n converges almost everywhere to 0.

Proof. As $\xi_n \rightarrow 0$ in probability, it has a subsequence which converges to 0 almost everywhere. Then, as ξ_n is decreasing, a element ω satisfying $\xi_{n_k}(\omega) \rightarrow 0$ also satisfies $\xi_n(\omega) \rightarrow 0$. Thus, the set of non-convergent elements has measure 0 as required. \square

5.2 Strong Law of Large Numbers

Definition 5.1. A sequence of random variables (ξ_n) is said to satisfy the weak/strong law of large numbers (LLN) if ξ_n are integrable and $\frac{S_n - \mathbb{E}S_n}{n}$ converges in probability/almost everywhere to 0 where $S_n := \sum_{i=1}^n \xi_i$.

We had previously proved the weak law of large numbers for sequences of i.i.d. random variables with finite variance. We will now consider strong LLN.

Lemma 5.1 (Cantelli's Strong Law of Large Numbers). Let (ξ_n) be a sequence of i.i.d random variables such that $\mathbb{E}\xi_1^4 < \infty$. Then, (ξ_n) satisfies the strong LLN.

Proof. WLOG. we may assume $\mathbb{E}\xi_1 = 0$ and so $\mathbb{E}S_n = 0$. Let $\epsilon > 0$, and define

$$A_n := \{\omega \mid |S_n(\omega)/n| > \epsilon\}.$$

Then, it suffices to show that $\mathbb{P}(\limsup |S_n/n| > \epsilon) = \mathbb{P}(\{A_n \text{ i.o.}\}) = 0$. Indeed, this follows by Borel-Cantelli where

$$\mathbb{P}(A_n) < \frac{1}{\epsilon^4} \mathbb{E} \left(\left| \frac{S_n}{n} \right|^4 \right)$$

by Chebyshev's inequality and so, by considering

$$\begin{aligned} \mathbb{E}S_n^4 &= \sum_{i=1}^n \mathbb{E}\xi_i^4 + \sum_{i < j} \binom{4}{2} \mathbb{E}\xi_i^2 \mathbb{E}\xi_j^2 + \mathbb{E} \dots = n\mathbb{E}\xi_1^4 + 3n(n-1)(\mathbb{E}\xi_1^2)^2 \\ &\leq n\mathbb{E}\xi_1^4 + 3n(n-1)\mathbb{E}\xi_1^4 \leq n^2(4\mathbb{E}\xi_1^4) \end{aligned}$$

where $\mathbb{E} \dots = 0$ as they represent terms with a singular ξ_i , and hence, by independence $\mathbb{E} \dots = \mathbb{E}\xi_i \mathbb{E} \dots = 0$. Thus, $\sum_n \mathbb{E}(|S_n/n|^4) \leq \sum_n 1/n^2 < \infty$, implying $\sum_n \mathbb{P}(A_n) < \infty$ which is the condition needed for Borel-Cantelli. \square

Theorem 27 (Kolmogorov's Inequality). Let (ξ_n) be a sequence of independent random variables with finite variance (we note that this is implied by finite second (or higher) moment by Lyapunov inequality). Then, for all $n \geq 1$, $\epsilon > 0$,

$$\mathbb{P} \left(\max_{1 \leq k \leq n} |S_k - \mathbb{E}S_k| \geq \epsilon \right) \leq \frac{V_{S_n}}{\epsilon^2}.$$

Proof. WLOG. we may assume $\mathbb{E}\xi_i = 0$. Defining $A := \{\max_{1 \leq k \leq n} |S_k| \geq \epsilon\}$, and

$$A_k := \{|S_j| < \epsilon, j = 1, \dots, k-1, |S_k| \geq \epsilon\}.$$

It is clear that A_k are mutually disjoint, $A = \bigcup_{i=1}^n A_i$ and, $\mathbb{E}S_n^2 \geq \mathbb{E}S_n^2 \mathbf{1}_A = \sum_{i=1}^n \mathbb{E}S_n^2 \mathbf{1}_{A_i}$. Considering

$$\begin{aligned} \mathbb{E}S_n^2 \mathbf{1}_{A_i} &= \mathbb{E}(S_i + \xi_{i+1} + \dots + \xi_n)^2 \mathbf{1}_{A_i} \\ &= \mathbb{E}S_i^2 \mathbf{1}_{A_i} + 2\mathbb{E}S_i(\xi_{i+1} + \dots + \xi_n) \mathbf{1}_{A_i} + \mathbb{E}(\xi_{i+1} + \dots + \xi_n)^2 \mathbf{1}_{A_i} \end{aligned}$$

where by independence,

$$\mathbb{E}S_i(\xi_1 + \dots + \xi_n) \mathbf{1}_{A_i} = \mathbb{E}S_i \mathbf{1}_{A_i} \mathbb{E}(\xi_1 + \dots + \xi_n) = 0,$$

and so,

$$\mathbb{E}S_n^2 \mathbf{1}_{A_i} = \mathbb{E}S_i^2 \mathbf{1}_{A_i} + \mathbb{E}(\xi_{i+1} + \dots + \xi_n)^2 \mathbf{1}_{A_i} \geq \mathbb{E}S_i^2 \mathbf{1}_{A_i}.$$

Hence, combining the inequalities, we obtain $\mathbb{E}S_n^2 \geq \mathbb{E}S_i^2 \mathbf{1}_A \geq \epsilon^2 \mathbb{P}(A)$ which is exactly the required inequality after dividing both sides by ϵ^2 . \square

Theorem 28 (Two-series Theorem). Let (ξ_n) be a sequence of independent random variable. If $\sum_{n=1}^{\infty} \mathbb{E}\xi_n$ and $\sum_{n=1}^{\infty} V_{\xi_n}$ both converge, then the series $\sum_{n=1}^{\infty} \xi_n$ converges almost everywhere.

Proof. In the case that $\mathbb{E}\xi_i = 0$, we will show (S_n) is Cauchy almost everywhere, i.e.

$$\lim_{k \rightarrow \infty} \sup_{m, n \geq k} |S_m - S_n| = 0$$

almost everywhere. However, defining $T_k := \sup_{m, n \geq k} |S_m - S_n|$, we have a sequence of decreasing, non-negative random variables, and hence, it suffice to show T_n converges to 0 in probability. We note that

$$0 \leq \sup_{n, m \geq k} |S_n - S_k| \leq \sup_{n, m \geq k} (|S_n - S_k| + |S_k - S_m|) = 2 \sup_{n \geq k} |S_n - S_k| =: 2\sigma_k.$$

Hence, it suffice to show $\sigma_k \rightarrow 0$ in probability. Fix $\epsilon > 0$, then, by Kolmogorov's inequality,

$$\mathbb{P}(\sigma_k \geq \epsilon) = \lim_{m \rightarrow \infty} \mathbb{P}(\max_{k \leq n \leq m} |S_n - S_k| \geq \epsilon) \leq \frac{1}{\epsilon^2} \lim_{m \rightarrow \infty} \sum_{n=k+1}^m V_{\xi_n} = \frac{1}{\epsilon^2} \sum_{n=k+1}^{\infty} V_{\xi_n}.$$

Now, as $\sum V_{\xi_n}$ converges, the right hand side tends to 0 as $k \rightarrow \infty$ and thus, $\sigma_k \rightarrow 0$ in probability.

In the case that $\mathbb{E}\xi_j \neq 0$, one may simply consider

$$\sum_{n=1}^{\infty} \xi_n = \sum_{n=1}^{\infty} (\xi_n - \mathbb{E}\xi_n) + \sum_{n=1}^{\infty} \mathbb{E}\xi_n$$

where $\sum_{n=1}^{\infty} \mathbb{E}\xi_n$ converges by assumption and $\sum_{n=1}^{\infty} (\xi_n - \mathbb{E}\xi_n)$ converges almost everywhere by the first case. \square

Lemma 5.2 (Toeplitz). Let (a_n) be a non-negative real sequence and $b_n := \sum_{k=1}^n a_k$ such that $b_n \rightarrow \infty$. Furthermore, let $x_n \rightarrow x$. Then,

$$\lim_{n \rightarrow \infty} \frac{1}{b_n} \sum_{k=1}^n a_k x_k = x.$$

Proof. Elementary analysis. \square

Lemma 5.3 (Kronecker). Let (a_n) be a non-negative real sequence and $b_n := \sum_{k=1}^n a_k$ such that $b_n \rightarrow \infty$. Furthermore, let (x_n) be a sequence such that $\sum_{k=1}^{\infty} x_n$ converge. Then,

$$\lim_{n \rightarrow \infty} \frac{1}{b_n} \sum_{k=1}^{\infty} b_k x_k = 0.$$

Proof. Follows by summation by parts. \square

Theorem 29. Let (ξ_n) be a sequence of independent random variables with finite variance and let (b_n) be a sequence of non-negative real numbers such that $b_n \rightarrow \infty$, then, if $\sum_{k=1}^{\infty} \frac{V_{\xi_k}}{b_k^2} < \infty$, we have

$$\frac{S_n - \mathbb{E}S_n}{b_n} \rightarrow 0$$

almost everywhere.

We note that if we simply choose $b_n = n$, we obtain a condition for strong LLN.

Proof. We observe that

$$\frac{S_n - \mathbb{E}S_n}{b_n} = \frac{1}{b_n} \sum_{k=1}^n b_k \frac{\xi_k - \mathbb{E}\xi_k}{b_k}$$

and $V_{\sum_{k=1}^n \frac{\xi_k - \mathbb{E}\xi_k}{b_k}} = \sum_{k=1}^n \frac{V_{\xi_k}}{b_k^2}$ which converges by assumption. Then, by the two-series theorem, $\sum_{k=1}^n \frac{\xi_k - \mathbb{E}\xi_k}{b_k}$ converges almost surely. Thus, applying Kronecker's lemma at all points of convergence, we obtain the required result. \square

Lemma 5.4. Let $\xi \geq 0$ be a non-negative random variable. Then,

$$\sum_{n=1}^{\infty} \mathbb{P}(\xi \geq n) \leq \mathbb{E}\xi \leq 1 + \sum_{n=1}^{\infty} \mathbb{P}(\xi \geq n).$$

Proof. This follows by

$$\begin{aligned} \sum_{n=1}^{\infty} \mathbb{P}(\xi \geq n) &= \sum_{n=1}^{\infty} \sum_{k \geq n} \mathbb{P}(k \leq \xi \leq k+1) = \sum_{k=1}^{\infty} k \mathbb{P}(k \leq \xi < k+1) \\ &= \sum_{k=1}^{\infty} \mathbb{E}(k \mathbf{1}_{[k, k+1)}(\xi)) \leq \sum_{k=1}^{\infty} \mathbb{E}(\xi \mathbf{1}_{[k, k+1)}(\xi)) = \mathbb{E}\xi \\ &\leq \sum_{k=1}^{\infty} \mathbb{E}((k+1) \mathbf{1}_{[k, k+1)}(\xi)) = 1 + \sum_{n=1}^{\infty} \mathbb{P}(\xi \geq n). \end{aligned}$$

□

Theorem 30 (Kolmogorov's Strong Law of Large Numbers). Let $(\xi_n)_{n=1}^\infty$ be a sequence of i.i.d. random variables with finite expectation so $\mathbb{E}\xi_1 = m$, then

$$\frac{1}{n}S_n \rightarrow m$$

almost everywhere.

We note that $S_n/n \rightarrow m$ almost everywhere is equivalent to $\sum_{k=1}^n (\xi_k - \mathbb{E}\xi_k) = o(n)$ as $n \rightarrow \infty$ almost everywhere.

Proof. WLOG. we may assume $\mathbb{E}\xi_1 = 0$, and it suffices to show $S_n/n \rightarrow 0$ almost everywhere. As $\mathbb{E}|\xi_1| < \infty$, by the above lemma we have $\sum_{n=1}^\infty \mathbb{P}(|\xi_n| \geq n) \leq \mathbb{E}|\xi_1| < \infty$. Hence, by Borel-Cantelli, $\mathbb{P}(|\xi_n| \geq n \text{ i.o.}) = 0$ and so $\mathbb{P}(|\xi_n| < n \text{ e.v.}) = 1$.

Now, defining

$$\tilde{\xi}_n = \begin{cases} \xi_n, & |\xi_n| < n, \\ 0, & |\xi_n| \geq n, \end{cases}$$

the above implies $S_n/n \rightarrow 0$ almost everywhere if and only if $\sum_{k=1}^n \tilde{\xi}_k/n \rightarrow 0$ almost everywhere. Noting

$$\mathbb{E}\tilde{\xi}_n = \mathbb{E}\xi_n \mathbf{1}_{|\xi_n| < n} = \mathbb{E}\xi_1 \mathbf{1}_{|\xi_1| < n}$$

tends to $\mathbb{E}\xi_1$ as $n \rightarrow \infty$ almost everywhere. Now, by Toeplitz lemma with $x_n := \mathbb{E}\tilde{\xi}_n$, we have $\frac{1}{n} \sum_{k=1}^n \mathbb{E}\tilde{\xi}_k \rightarrow 0$ almost everywhere as $n \rightarrow \infty$. Thus, $S_n/n \rightarrow 0$ if and only if $\frac{1}{n}(\sum_k (\tilde{\xi}_k - \mathbb{E}\tilde{\xi}_k)) \rightarrow 0$ almost everywhere. Then, applying Kronecker's lemma with $b_n = n$, $x_n = (\tilde{\xi}_n - \mathbb{E}\tilde{\xi}_n)/n$, we obtain $\frac{1}{n}(\sum_k (\tilde{\xi}_k - \mathbb{E}\tilde{\xi}_k)) \rightarrow 0$ if $\sum_{n=1}^\infty \frac{\tilde{\xi}_n - \mathbb{E}\tilde{\xi}_n}{n}$ converges. Then, applying the two-series theorem, it suffices to show

$$\sum_{n=1}^\infty V\left(\frac{\tilde{\xi}_n - \mathbb{E}\tilde{\xi}_n}{n}\right) = \sum_{n=1}^\infty \frac{V(\tilde{\xi}_n - \mathbb{E}\tilde{\xi}_n)}{n^2} < \infty.$$

But this follows by considering

$$\begin{aligned} \sum_{n=1}^\infty \frac{V(\tilde{\xi}_n - \mathbb{E}\tilde{\xi}_n)}{n^2} &\leq \sum_{n=1}^\infty \frac{\mathbb{E}\tilde{\xi}_n^2}{n^2} = \sum_{n=1}^\infty \frac{1}{n^2} \mathbb{E}(\xi_1^2 \mathbf{1}_{|\xi_1| < n}) \\ &= \sum_{n=1}^\infty \frac{1}{n^2} \sum_{k=1}^n \mathbb{E}(\xi_1^2 \mathbf{1}_{k-1 \leq |\xi_1| < k}) \\ &= \sum_{k=1}^\infty \mathbb{E}(\xi_1^2 \mathbf{1}_{k-1 \leq |\xi_1| < k}) \sum_{n \geq k} \frac{1}{n^2} \\ &\leq 2 \sum_{k=1}^\infty \mathbb{E}(|\xi_1| \mathbf{1}_{k-1 \leq |\xi_1| < k}) = 2\mathbb{E}|\xi_1| < \infty. \end{aligned}$$

□