Nested Sampling and the Evaluation of the 'Evidence' for Bayesian Model Selection

Paul D. Baines, Nicholas Ulle University of California, Davis

February 13, 2014

1 Introduction

- Explanation of Nested Sampling
- Intuitive explanation of the algorithm
- Basics of computing the evidence and Bayesian model selection

Example I

Let:

$$Y_i \sim N(\mu, \sigma^2), \qquad i = 1, \dots, n,$$

with prior $p(\mu) \propto N(\mu_0, \tau_0^2)$ and σ^2 known. Letting $C = (2\pi)^{-(n+1)/2} (\tau_0^2)^{-1/2} (\sigma^2)^{-n/2}$, the evidence, or marginal likelihood, is:

$$\begin{split} p(y) &= \int p(y_1, \dots, y_n | \mu) p(\mu) d\mu = \int p(\mu) \prod_{i=1}^n p(y_i | \mu) d\mu \\ &= \int C \times \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2 - \frac{1}{2\tau_0^2} (\mu - \mu_0)^2\right\} d\mu \\ &= C \times \int \exp\left\{-\frac{1}{2} \left(\frac{n}{\sigma^2} + \frac{1}{\tau_0^2}\right) \mu^2 + \frac{1}{2} \mu \left(\frac{\mu_0}{\tau_0^2} + \frac{\sum_{i=1}^n y_i}{\sigma^2}\right) - \frac{1}{2} \left(\frac{\mu_0^2}{\tau_0^2} + \frac{\sum_{i=1}^n y_i^2}{\sigma^2}\right)\right\} d\mu \\ &= C \times \exp\left\{-\frac{1}{2} \left(\frac{\mu_0^2}{\tau_0^2} + \frac{\sum_{i=1}^n y_i^2}{\sigma^2}\right) + \frac{1}{2} \left(\frac{n}{\sigma^2} + \frac{1}{\tau_0^2}\right)^{-1} \left[\frac{\mu_0}{\tau_0^2} + \frac{\sum_{i=1}^n y_i}{\sigma^2}\right]^2\right\} \\ &\times \int \exp\left\{-\frac{1}{2} \left(\frac{n}{\sigma^2} + \frac{1}{\tau_0^2}\right) \left(\mu - \left(\frac{n}{\sigma^2} + \frac{1}{\tau_0^2}\right)^{-1} \left[\frac{\mu_0}{\tau_0^2} + \frac{\sum_{i=1}^n y_i}{\sigma^2}\right]\right)^2\right\} d\mu \\ &= C \times \exp\left\{-\frac{1}{2} \left(\frac{\mu_0^2}{\tau_0^2} + \frac{\sum_{i=1}^n y_i^2}{\sigma^2}\right) + \frac{1}{2} \left(\frac{n}{\sigma^2} + \frac{1}{\tau_0^2}\right)^{-1} \left[\frac{\mu_0}{\tau_0^2} + \frac{\sum_{i=1}^n y_i}{\sigma^2}\right]^2\right\} \times (2\pi)^{1/2} \left(\frac{n}{\sigma^2} + \frac{1}{\tau_0^2}\right)^{-1/2} \\ &= C \times (2\pi)^{1/2} \left(\frac{n}{\sigma^2} + \frac{1}{\tau_0^2}\right)^{-1/2} \exp\left\{-\frac{1}{2} \left(\frac{\mu_0^2}{\tau_0^2} + \frac{\sum_{i=1}^n y_i^2}{\sigma^2}\right) + \frac{1}{2} \left(\frac{n}{\sigma^2} + \frac{1}{\tau_0^2}\right)^{-1} \left[\frac{\mu_0}{\tau_0^2} + \frac{\sum_{i=1}^n y_i}{\sigma^2}\right]^2\right\}. \end{split}$$

This will allow us to verify the results obtained using nested sampling. In the simple case where $\mu_0 = 0$, $\tau_0^2 = 1$, n = 1, $\sigma^2 = 1$ we obtain:

$$Z = (2\pi)^{-1/2} (2)^{-1/2} \exp\left\{-\frac{1}{2}y^2 + \frac{1}{2} (2)^{-1} y^2\right\} = \frac{1}{2\sqrt{\pi}} \exp\left\{-\frac{y^2}{4}\right\}.$$

2 Example II

Here we take a look at the classic mixture of normals:

$$Y_i = \sum_{j=1}^{K} I_{ij} Z_{ij}, \qquad i = 1, \dots, n,$$

where:

$$I_{i} = (I_{i1}, \dots, I_{iK}) \sim \text{Multinomial}(1, p),$$

$$Z_{ij} \stackrel{iid}{\sim} N(\mu_{j}, 1).$$

The parameters in the model are the mixture proportions $p = (p_1, \ldots, p_K)$ (with $\sum_j p_j = 1$) and the mixture locations $\mu = (\mu_1, \ldots, \mu_k)$. The number of mixture components K will be fixed for a given model, and we will use the evidence to motivate a model selection procedure to select the appropriate K. For convenience we choose conditionally conjugate priors for μ and p:

$$\mu \sim N\left(\mu_0, \tau_0^2\right), \qquad p \sim \text{Multinomial}\left(\alpha\right),$$

where μ_0, τ_0^2 and α are fixed hyperparameters chosen by the analyst.

2.1 Posterior Distributions

TODO

2.2 Evaluating the Evidence: Analytically

TODO

2.3 Evaluating the Evidence: Nested Sampling

TODO

2.4 Evaluating the Evidence: Other Methods

TODO