

Birth and Death Process — M/G/1 Queue — Pollaczek - Khintchine Formula, Series Queues — Open and Closed Networks

BIRTH AND DEATH PROCESS

U.Q

Let $x(t)$ be the number of the customers at time t [population at time t] in which two types of events occur namely **births** and **deaths**. Then the continuous time discrete random process $\{X(t)\}$ With state space $\{0, 1, 2, \dots\}$ is called a **birth and death process** if the following postulates are satisfied.

If $X(t) = n$, that is the size of the population at time t is n or the system is in state n . Let λ_n , $n = 0, 1, 2, \dots$ be the rate at which births occur in the state n and μ_n , $n = 1, 2, \dots$ be the rate at which deaths occur in state n , such that

- i. $P[1 \text{ birth in } (t, t+h)] = \lambda_n h + O(h)$
- ii. $P[0 \text{ birth in } (t, t+h)] = 1 - \lambda_n h + O(h)$
- iii. $P[2 \text{ or more births in } (t, t+h)] = O(h)$
- iv. Births in $(t, t+h)$ are independent of time since the last birth
- v. $P[1 \text{ death in } (t, t+h)] = \mu_n h + O(h)$
- vi. $P[0 \text{ death in } (t, t+h)] = 1 - \mu_n h + O(h)$
- vii. $P[2 \text{ or more deaths in } (t, t+h)] = O(h)$
- viii. Deaths occurring in $(t, t+h)$ are independent of time since the last death
- ix. Births and deaths occur independently of each other at any time.

Probability Distribution of $X(t)$

U.Q

Let $X(t)$ be a birth and death process.

Let $X(t) = n$ be the event with corresponding probability $P_n(t) = P[X(t) = n]$

Then $P_n(t+h) = P[X(t+h) = n]$ be the probability that the size of the population is n at time $t+h$.

The event $X(t+h) = n$ means that the process will be in state n at time $t+h$ if one of the following mutually exclusive and exhaustive events occurs.

- i. $X(t) = n$ and no birth or death in $(t, t+h)$
- ii. $X(t) = n-1$ and 1 birth and no death in $(t, t+h)$
- iii. $X(t) = n+1$ and no birth and one death in $(t, t+h)$
- iv. $X(t) = n$ and 1 birth and 1 death in $(t, t+h)$

Then, $P[X(t) = n] = P_n(t+h)$

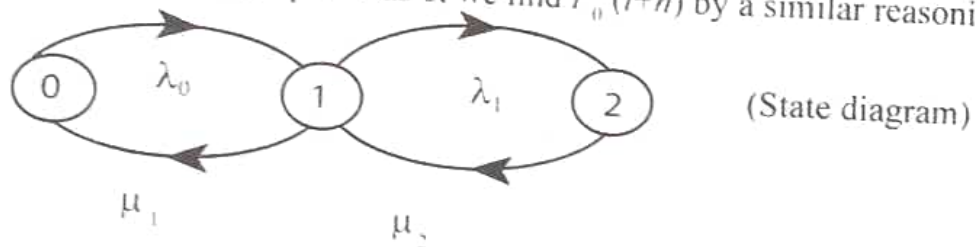
5.2 □ Probability and Queueing Theory

$$\begin{aligned}
 &= P_n(t)[1 - \lambda_n h + O(h)][1 - \lambda_n h + O(h)] + \\
 &\quad P_{n-1}(t)[\lambda_{n-1} h + O(h)][1 - \mu_{n-1} h + O(h)] + \\
 &\quad P_{n+1}(t)[1 - \lambda_{n+1} h + O(h)][\mu_{n+1} h + O(h)] + \\
 &\quad P_n(t)[\lambda_n h + O(h)][\mu_n h + O(h)] \quad [\because P_n(t+h) = P[(i) + (ii) + (iii) + (iv)]] \\
 &= P_n(t)[1 - \lambda_n h - \mu_n h] + O(h) + P_{n-1}(t)[\lambda_{n-1} h] + O(h) + \\
 &\quad P_{n+1}(t)[\mu_{n+1} h] + O(h) + P_n(t)
 \end{aligned}$$

Since $O(h)[1 - \lambda_n h + O(h)] = O(h)$ & $\lambda_n \mu_n h^2 - \lambda_n O(h) + O(h) = O(h)$

$$\begin{aligned}
 &\Rightarrow P_n(t+h) = P_n(t)[1 - (\lambda_n + \mu_n)h] + P_{n-1}(t)\lambda_{n-1}h + P_{n+1}(t)\mu_{n+1}h + O(h) \\
 \therefore P_n(t+h) - P_n(t) &= -P_n(t)(\lambda_n + \mu_n)h + P_{n-1}(t)\lambda_{n-1}h + P_{n+1}(t)\mu_{n+1}h + O(h) \\
 \Rightarrow \frac{P_n(t+h) - P_n(t)}{h} &= -P_n(t)(\lambda_n + \mu_n) + P_{n-1}(t)\lambda_{n-1} + P_{n+1}(t)\mu_{n+1} + \frac{O(h)}{h} \\
 \lim_{h \rightarrow 0} \left[\frac{P_n(t+h) - P_n(t)}{h} \right] &= -(\lambda_n + \mu_n)P_n(t) + \lambda_{n-1}P_{n-1}(t) + P_{n+1}(t)\mu_{n+1} \\
 \text{i.e. } P_n'(t) &= -(\lambda_n + \mu_n)P_n(t) + \lambda_{n-1}P_{n-1}(t) + P_{n+1}(t)\mu_{n+1} \quad \dots(1)
 \end{aligned}$$

This is true of all $n \geq 1$, Since the state space has 0, we find $P_0(t+h)$ by a similar reasoning.



$$\begin{aligned}
 P_0(t+h) &= P_0(t)[1 - \lambda_0 h + O(h)] + P_1(t)[1 - \lambda_0 h + O(h)][\mu_1 h + O(h)] \\
 &= P_0(t) - \lambda_0 h P_0(t) + P_1(t)\mu_1 h + O(h) \\
 \Rightarrow P_0(t+h) - P_0(t) &= -\lambda_0 h P_0(t) + \mu_1 h P_1(t) + O(h) \\
 \therefore \frac{P_0(t+h) - P_0(t)}{h} &= -\lambda_0 P_0(t) + \mu_1 P_1(t) + \frac{O(h)}{h} \\
 \therefore \lim_{h \rightarrow 0} \frac{P_0(t+h) - P_0(t)}{h} &= -\lambda_0 P_0(t) + \mu_1 P_1(t) \\
 \text{i.e. } P_0'(t) &= -\lambda_0 P_0(t) + \mu_1 P_1(t) \quad \dots(2)
 \end{aligned}$$

Solving (1) and (2) we get $P_n(t)$ for $n \geq 0$, which gives the probability distribution of $X(t)$.