PRINCIPAL COMPONENT ANALYSIS

A DIFFERENT PERSPECTIVE & PRACTICAL CONSIDERATIONS

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MOTIVATION



- In Machine Learning, we often deal with **high-dimensional data** (features)
- e.g., an image $X \in \mathbb{R}^{1024 \times 1024}$ (\approx 1 million dimensions)

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Dimensionality Reduction

- · Many of these features might be redundant
- · We wish to find a lower-dimensional (compressed) representation of our data
- · Useful for ...
 - Feature extraction
 - Visualization
 - · Reducing computational load
 - · Compression per se (smaller file size)

PRINCIPAL COMPONENT ANALYSIS (PCA)



Idea 🤔

- Variance in the data amounts to information
 - $\boldsymbol{\cdot}\,$ e.g., a feature that is constant for all data points is \boldsymbol{not} informative

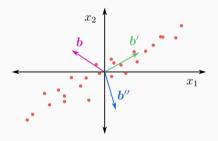
PRINCIPAL COMPONENT ANALYSIS (PCA)



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$$\mathcal{D} = \{\mathbf{x}^{(1)}, \dots, \mathbf{x}^{(N)}\}$$
 with $\mathbf{x}^{(i)} \in \mathbb{R}^2$



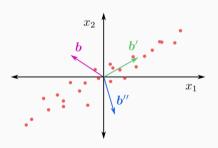
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• A given **b** projects $\mathbf{x}^{(i)}$ to $z^{(i)} \in \mathbb{R}$

$$z^{(i)} = \mathbf{b}^{\mathsf{T}} \mathbf{x}^{(i)}$$

- In the original coordinate system, the projection is then $\hat{\textbf{x}}=z\textbf{b}$
- Find b (unit length) such that the variance in the projections z is maximized, i.e.,

$$b^* = \underset{b^\intercal b = 1}{\mathsf{argmax}} \ \mathsf{Var}\left(b^\intercal x^{(1)}, \dots, b^\intercal x^{(N)}\right)$$

PCA IN ARBITRARY DIMENSIONS



- Let's generalize this to $\mathcal{D} = \{\mathbf{x}^{(1)}, \dots, \mathbf{x}^{(N)}\}$ with $\mathbf{x}^{(i)} \in \mathbb{R}^{\mathcal{D}}$

Orthonormal Basis

For each subspace $U \subseteq \mathbb{R}^D$ there exists a set of **orthonormal basis vectors** that span U.

- Let $\{\mathbf{b}_1,\ldots,\mathbf{b}_D\}$ be an orthonormal basis of \mathbb{R}^D , collected in $B=(\mathbf{b}_1\ \mathbf{b}_2\ \ldots\ \mathbf{b}_D)$
- B is orthonormal, i.e., $B^TB = I$ and thus, $B^{-1} = B^T$
 - · Columns have unit norm and are pairwise orthogonal

Change of Basis

Recall that any $\mathbf{x} \in \mathbb{R}^D$ can be expressed as coordinates w.r.t. B (called \mathbf{z}):

$$\mathbf{x} = B\mathbf{z} \iff B^{-1}\mathbf{x} = \mathbf{z} \iff B^{\mathsf{T}}\mathbf{x} = \mathbf{z}$$

PCA (CONT.)



- $\mathbf{x} = B\mathbf{z}$ and $B^T\mathbf{x} = \mathbf{z}$
- We can transform z back into the original coordinate system: $x = \underline{\mathcal{B}}\underline{\mathcal{B}}^Tx$
- · So far we have just played with coordinate systems, no compression vet
 - · Since we have D basis vectors
- Let's compress $\mathbf{x} \in \mathbb{R}^D$ into a representation $\mathbf{z} \in \mathbb{R}^M$ with M < D

PROJECTING INTO SUBSPACE



- Let $B_M = (\mathbf{b}_1 \dots \mathbf{b}_M)$ and $B_R = (\mathbf{b}_{M+1} \dots \mathbf{b}_D)$
- Assume we are given B_M . How do we project \mathbf{x} into the subspace spanned by B_M ?

Optimal Projection

Given $\mathbf{x} \in \mathbb{R}^D$ and B_M , find $\mathbf{z}^* \in \mathbb{R}^M$ such that

$$\mathbf{z}^* = \underset{\mathbf{z}}{\operatorname{argmin}} \|B_{\mathsf{M}}\mathbf{z} - \mathbf{x}\|_2^2.$$

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Solution

- Looks just like finding parameters for linear regression
- · Closed form solution recovers **orthogonal projection**:

$$\mathbf{z}^* = (B_{\mathsf{M}}{}^{\mathsf{T}}B_{\mathsf{M}})^{-1}B_{\mathsf{M}}{}^{\mathsf{T}}\mathbf{x} = B_{\mathsf{M}}{}^{\mathsf{T}}\mathbf{x}.$$

• Transform back into original coordinates: $\hat{\mathbf{x}} = B_{\mathsf{M}} B_{\mathsf{M}}^{\mathsf{T}} \mathbf{x}$

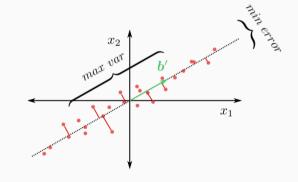


- $\cdot \hat{\mathbf{x}} = B_{\mathsf{M}} B_{\mathsf{M}}^{\mathsf{T}} \mathbf{x}$
- Note that $B_M B_M^T$ is not identity anymore !
 - The inverse of B_M does not exist
- How to find a good B_M ?



We want to find B_M s.t. projections have minimal average squared projection error!

$$\mathcal{L} = \frac{1}{N} \sum_{i=1}^{N} \|\mathbf{x}^{(i)} - \underbrace{B_{M}B_{M}^{T}\mathbf{x}^{(i)}}_{\hat{\mathbf{x}}^{(i)}}\|_{2}^{2}$$



REWRITING THE LOSS



- $\cdot B = (\mathbf{b}_1 \ \mathbf{b}_2 \ \dots \ \mathbf{b}_D), B_M = (\mathbf{b}_1 \ \dots \ \mathbf{b}_M) \text{ and } B_R = (\mathbf{b}_{M+1} \ \dots \ \mathbf{b}_D)$
- Note that

$$BB^{T} = \sum_{j=1}^{N} \mathbf{b}_{j} \mathbf{b}_{j}^{T} = \sum_{j=1}^{M} \mathbf{b}_{j} \mathbf{b}_{j}^{T} + \sum_{j=M+1}^{D} \mathbf{b}_{j} \mathbf{b}_{j}^{T} = B_{M} B_{M}^{T} + B_{R} B_{R}^{T}$$

- Since $\mathbf{x} = BB^T\mathbf{x}$, we have $\mathbf{x} = \left(B_MB_M^T + B_RB_R^T\right)\mathbf{x}$
- · We rewrite the residual

$$\mathbf{x} - \hat{\mathbf{x}} = \underbrace{\left(B_{M}B_{M}^{T} + B_{R}B_{R}^{T}\right)\mathbf{x}}_{\mathbf{x}} - \underbrace{B_{M}B_{M}^{T}\mathbf{x}}_{\hat{\mathbf{x}}} = B_{R}B_{R}^{T}\mathbf{x}$$

- The error is the projection on the orthogonal complement of the principal subspace
- · We thus minimize

$$\mathcal{L} = \frac{1}{N} \sum_{i=1}^{N} \| \mathbf{B}_{R} \mathbf{B}_{R}^{\mathsf{T}} \mathbf{x}^{(i)} \|_{2}^{2}$$



$$\mathcal{L} = \frac{1}{N} \sum_{i=1}^{N} \|B_{R}B_{R}^{T} \mathbf{x}^{(i)}\|_{2}^{2} = \frac{1}{N} \sum_{i=1}^{N} \left(B_{R}B_{R}^{T} \mathbf{x}^{(i)}\right)^{T} \left(B_{R}B_{R}^{T} \mathbf{x}^{(i)}\right)$$

$$= \frac{1}{N} \sum_{i=1}^{N} \mathbf{x}^{(i)T} B_{R} \underbrace{B_{R}^{T} B_{R}}_{I} B_{R}^{T} \mathbf{x}^{(i)}$$

$$= \frac{1}{N} \sum_{i=1}^{N} \mathbf{x}^{(i)T} \underbrace{B_{R}B_{R}^{T} \mathbf{x}^{(i)}}_{j=M+1} \mathbf{b}_{j} \mathbf{b}_{j}^{T} \mathbf{x}^{(i)}$$

$$= \frac{1}{N} \sum_{i=1}^{N} \sum_{i=M+1}^{N} \mathbf{x}^{(i)T} \mathbf{b}_{j} \mathbf{b}_{j}^{T} \mathbf{x}^{(i)} = \frac{1}{N} \sum_{i=1}^{N} \sum_{j=M+1}^{D} \mathbf{b}_{j}^{T} \mathbf{x}^{(i)T} \mathbf{b}_{j}$$



$$\mathcal{L} = \frac{1}{N} \sum_{i=1}^{N} \sum_{j=M+1}^{D} \mathbf{b}_{j}^{T} \mathbf{x}^{(i)} \mathbf{x}^{(i)T} \mathbf{b}_{j}$$
$$= \sum_{j=M+1}^{D} \mathbf{b}_{j}^{T} \underbrace{\left(\frac{1}{N} \sum_{i=1}^{N} \mathbf{x}^{(i)} \mathbf{x}^{(i)T}\right)}_{C} \mathbf{b}_{j}$$

Data Covariance Matrix

C is the **empirical covariance matrix** of X:

$$C = \frac{1}{N} X^{T} X \qquad X = \begin{pmatrix} \mathbf{x}^{(1)^{T}} \\ \mathbf{x}^{(2)^{T}} \\ \vdots \\ \mathbf{x}^{(N)^{T}} \end{pmatrix}$$

How to pick $B_R = (\mathbf{b}_{M+1} \dots \mathbf{b}_D)$ such that $\sum_{i=M+1}^D \mathbf{b}_i^T C \mathbf{b}_i$ is minimized?



$$\min_{\mathbf{b}_{M+1},\dots,\mathbf{b}_{D}}\mathbf{b}_{M+1}^{T}C\mathbf{b}_{M+1}+\dots+\mathbf{b}_{D}^{T}C\mathbf{b}_{D}$$

• Consider $min_b b^T C b$ subject to $b^T b = 1$



$$\min_{\mathbf{b}_{M+1},\dots,\mathbf{b}_{D}} \mathbf{b}_{M+1}^{T} C \mathbf{b}_{M+1} + \dots + \mathbf{b}_{D}^{T} C \mathbf{b}_{D}$$

- Consider $min_b b^T Cb$ subject to $b^T b = 1$
- $\boldsymbol{\cdot} \; \textbf{Constrained minimization} \to \mathsf{setup} \; \textbf{Lagrangian}$

·
$$L(\mathbf{b}, \lambda) = \mathbf{b}^T C \mathbf{b} - \lambda \underbrace{(\mathbf{b}^T \mathbf{b} - 1)}_{\text{Constraint}}$$

- $\nabla_{\lambda} L(\mathbf{b}) = 1 \mathbf{b}^{\mathsf{T}} \mathbf{b} = 0 \Leftrightarrow \mathbf{b}^{\mathsf{T}} \mathbf{b} = 1$ (just the constraint)
- $\nabla_{\mathbf{b}} L(\mathbf{b}, \lambda) = 2C\mathbf{b} 2\lambda \mathbf{b} = 0 \Leftrightarrow C\mathbf{b} = \lambda \mathbf{b}$
- The optimal **b** is an **eigenvector** of *C*
- Since $\mathbf{b}^T C \mathbf{b} = \mathbf{b}^T (\lambda \mathbf{b}) = \lambda \mathbf{b}^T \mathbf{b} = \lambda$, pick \mathbf{b} to be the eigenvector associated with the smallest eigenvalue



- Since *C* is symmetric, its eigenvectors (with different eigenvalues) are **orthogonal** to each other (Spectral Theorem)
- Pick $B_R = (\mathbf{b}_{M+1} \dots \mathbf{b}_D)$ to be the orthonormal eigenvectors of C with the smallest eigenvalues
- B_R is the orthogonal complement of $B_M = (\mathbf{b}_1 \dots \mathbf{b}_M)$
- \cdot Thus, we construct B_M with the **remaining** eigenvectors (with the **largest eigenvalues**)



TL;DR

To minimize projection error, pick the subspace spanned by the orthonormal eigenvectors of *C* that have the largest eigenvalues.

Duality

 $\label{eq:minimizing Projection Error} \Leftrightarrow \text{Maximizing Variance in the Projections}$

PRACTICAL CONSIDERATIONS



- We need to find the M eigenvectors with the largest eigenvalues of X^TX
- · How do we actually do this? 😲
- · np.linalg.eigh(X.T ∂ X) works in theory, but is wasteful
 - $\cdot X^T X \in \mathbb{R}^{D \times D}$
 - · We compute D eigenvalues, although we only need the top-M (M \ll D) $\stackrel{ ext{(a)}}{=}$
 - Use **algorithms that only compute** the *D* largest eigenvalues (and their eigenvectors)



What if D is large (e.g. images) and we have $N \ll D$ data points?



· We want to find an eigenvector **b** of $C = \frac{1}{N} X^T X \in \mathbb{R}^{D \times D}$ (with eigenvalue λ)

$$\frac{1}{N}X^{T}Xb = \lambda b$$

$$\frac{1}{N}XX^{T}\underbrace{Xb}_{a} = \lambda \underbrace{Xb}_{a}$$

$$\frac{1}{N}XX^{T}a = \lambda a$$



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$$\frac{1}{N}XX^{T}a = \lambda a$$

- a is an eigenvector of $\frac{1}{N}XX^T \in \mathbb{R}^{N \times N}$
- · If we find a, we can convert it into an eigenvector of C:

$$\frac{1}{N}XX^{T}\mathbf{a} = \lambda\mathbf{a}$$

$$\frac{1}{N}X^{T}X(X^{T}\mathbf{a}) = \lambda(X^{T}\mathbf{a})$$

$$C(X^{T}\mathbf{a}) = \lambda(X^{T}\mathbf{a})$$

PCA DEMO WITH scikit-learn