# COMP 540 Assignment #1

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# 0 Background refresher(30 points)

• Plot the categorical distribution.

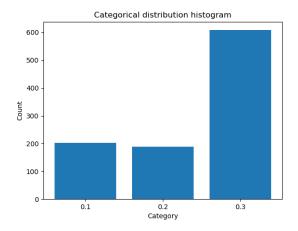


Figure 1: Categorical distribution

• Plot the Univariate normal distribution with mean of and standard deviation of 1.

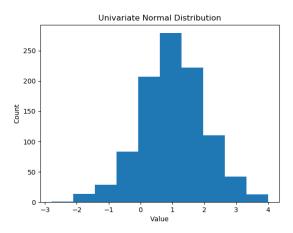


Figure 2: Univariate Normal Distribution

• Produce a scatter plot of the samples for a 2-D Gaussian.

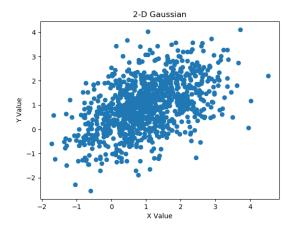


Figure 3: Univariate Normal Distribution

• Test mixture sampling code Code can be seen in sampler.py. Mixture Gaussian plot is shown below

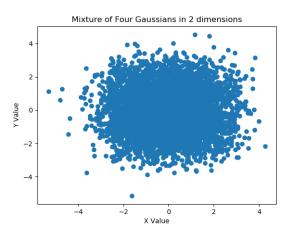


Figure 4: Univariate Normal Distribution

• Prove that the sum of two independent Poisson random variables is also a Poisson random variable. Suppose  $X \sim \mathcal{P}(\lambda)$  and  $Y \sim \mathcal{P}(\mu)$ . Now Prove that  $X + Y \sim \mathcal{P}(\lambda + \mu)$ .

$$\begin{split} P(X+Y=k) &= \sum_{i=0}^{k} P(X+Y=k,X=i) \\ &= \sum_{i=0}^{k} P(Y=k-i,X=i) \\ &= \sum_{i=0}^{k} P(Y=k-i)P(X=i) \\ &= \sum_{i=0}^{k} e^{-\mu} \frac{\mu^{k-i}}{(k-i)!} e^{-\lambda} \frac{\lambda^{i}}{i!} \\ &= e^{-(\mu+\lambda)} \frac{1}{k!} \sum_{i=0}^{k} \frac{k!}{i!(k-i)!} \mu^{k-i} \lambda^{i} \\ &= e^{-(\mu+\lambda)} \frac{1}{k!} \sum_{i=0}^{k} \binom{k}{i} \mu^{k-i} \lambda^{i} \\ &= \frac{(\mu+\lambda)^{k}}{k!} \cdot e^{-(\mu+\lambda)} \end{split}$$

So  $X + Y \sim \mathcal{P}(\lambda + \mu)$ .

• Find  $\alpha, \mu_1$  and  $\sigma_1$ . We have  $X_0$  and  $X_1$  be continuous random variables. If

$$p(X_0 = x_0) = \alpha_0 e^{-\frac{(x_0 - \mu_0)^2}{2\sigma_0^2}}$$
$$P(X_1 = x_1 | X_0 = x_0) = \alpha_1 e^{-\frac{(x_1 - x_0)^2}{2\sigma^2}}$$

$$\begin{split} p(X_1 = x_1) &= \int P(X_1 = x_1 | X_0 = x_0) \cdot p(X_0 = x_0) dx_0 \\ &= \alpha_0 \alpha_1 \int e^{-\frac{\sigma^2 (x_0 - \mu_0)^2 + \sigma_0^2 (x_1 - x_0)^2}{2\sigma_0^2 \sigma^2}} dx_0 \\ &= \alpha_0 \alpha_1 \int e^{-\frac{(\sigma^2 + \sigma_0^2) x_0^2 - 2(\sigma^2 \mu_0 + \sigma_0^2 x_1) x_0 + \sigma^2 \mu_0^2 + \sigma_0^2 x_1^2}{2\sigma_0^2 \sigma^2}} dx_0 \\ &= \alpha_0 \alpha_1 \int e^{-\frac{1}{2\sigma_0^2 \sigma^2} [(\sqrt{\sigma^2 + \sigma_0^2} x_0 - \frac{-\sigma^2 \mu_0 + \sigma_0^2 x_1}{\sqrt{\sigma^2 + \sigma_0^2}})^2 + \sigma^2 \mu_0^2 + \sigma_0^2 x_1^2 - \frac{\sigma^2 \mu_0^2 + \sigma_0^2 x_1^2}{\sigma^2 + \sigma_0^2}]} dx_0 \end{split}$$

Since

$$\int \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx = 1$$

$$p(X_1 = x_1) = \frac{\alpha \alpha_0 \sqrt{2\pi} \sigma \sigma_0}{\sqrt{\sigma^2 + \sigma_0^2}} e^{-\frac{1}{2\sigma_0^2 \sigma^2} (\sigma^2 \mu_0^2 + \sigma_0^2 x_1^2 - \frac{\sigma^2 \mu_0^2 + \sigma_0^2 x_1^2}{\sigma^2 + \sigma_0^2})}$$
$$= \alpha e^{-\frac{(x_1 - \mu_0)^2}{2(\sigma^2 + \sigma_0^2)}}$$

Thus we can solve that:

$$\alpha = \frac{\alpha \alpha_0 \sqrt{2\pi} \sigma \sigma_0}{\sqrt{\sigma^2 + \sigma_0^2}}$$
$$\mu_1 = \mu_0$$
$$\sigma_1 = \sqrt{\sigma^2 + \sigma_0^2}$$

• Show that if P(A|B,C) > P(A|B) then  $P(A|B,C^C) < P(A|B)$ 

$$P(A|B,C) = \frac{P(A \cap B \cap C)}{P(B \cap C)}$$

$$\frac{P(A \cap B \cap C)}{P(B \cap C)} > \frac{P(A \cap B)}{P(B)}$$

$$\frac{P(C|A \cap B)}{P(C|B)} > 1$$

$$P(C|A \cap B) > P(C|B)$$

$$1 - P(C|A \cap B) < 1 - P(C|B)$$

$$P(C^C|A \cap B) < P(C^C|B)$$

$$\frac{P(C^C|A \cap B)}{P(C^C|B)} < 1$$

$$\begin{split} P(A|B,C^C) &= \frac{P(A \bigcap B \bigcap C^C)}{P(B \bigcap C^C)} \\ &= \frac{P(C^C|A \bigcap B)P(A \bigcap B)}{P(C^C|B)P(B)} \\ &= \frac{P(C^C|A \bigcap B)}{P(C^C|B)P(B)}P(A|B) \\ &< P(A|B) \end{split}$$

Thus, if P(A|B,C) > P(A|B) then  $P(A|B,C^C) < P(A|B)$ .

• Consider the vectors  $u = \begin{bmatrix} 1 & 2 \end{bmatrix}^T$  and  $v = \begin{bmatrix} 2 & 3 \end{bmatrix}^T$ . Define the matrix  $M = uv^T$ . Compute the eigenvalues and eigenvectors of M.

$$\begin{aligned} \boldsymbol{M} &= \boldsymbol{u} \boldsymbol{v}^{\mathrm{T}} \\ &= \begin{bmatrix} 1 & 2 \end{bmatrix}^{\mathrm{T}} \cdot \begin{bmatrix} 2 & 3 \end{bmatrix} \\ &= \begin{bmatrix} 2 & 3 \\ 4 & 6 \end{bmatrix} \end{aligned}$$

$$\begin{vmatrix} \lambda \mathbf{I} - \mathbf{M} \end{vmatrix} = \begin{vmatrix} \lambda - 2 & -3 \\ -4 & \lambda - 6 \end{vmatrix} = 0$$

Then we can solve it as

$$(\lambda - 2)(\lambda - 6) - 12 = 0$$
$$\lambda^2 - 8\lambda = 0$$
$$\lambda(\lambda - 8) = 0$$
$$\lambda_1 = 0, \lambda_2 = 8$$

Let  $\lambda = 0$ :

$$(\lambda \mathbf{I} - \mathbf{M}) \cdot \begin{bmatrix} x_1 & x_2 \end{bmatrix}^{\mathrm{T}} = 0$$

$$\begin{bmatrix} -2 & -3 \\ -4 & -6 \end{bmatrix} \cdot \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$$

$$-2x_1 - 3x_2 = 0$$

Let  $x_1 = 3$ , Then  $x_2 = -2$ . So eigenvector is  $\begin{bmatrix} 3 & -2 \end{bmatrix}^T$ . Let  $\lambda = 8$ :

$$(\lambda \mathbf{I} - \mathbf{M}) \cdot \begin{bmatrix} x_1 & x_2 \end{bmatrix}^{\mathrm{T}} = 0$$
$$\begin{bmatrix} 6 & -3 \\ -4 & 2 \end{bmatrix} \cdot \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$$
$$6x_1 - 3x_2 = 0$$
$$-4x_1 + 2x_2 = 0$$

Let  $x_1 = 1$ , Then  $x_2 = 2$ . So eigenvector is  $\begin{bmatrix} 1 & 2 \end{bmatrix}^T$ . Thus, when eigenvalue  $\lambda = 0$ , eigenvector is  $\begin{bmatrix} 3 & -2 \end{bmatrix}^T$ , when eigenvalue  $\lambda = 8$ , eigenvector is  $\begin{bmatrix} 1 & 2 \end{bmatrix}^T$ .

• Show that if A is positive semi-definite, then all eigenvalues of A are non-negative. Definition of an eigenvalue and eigenvector is:

$$A\mathbf{v} = \lambda \mathbf{v}$$

If A is positive semi-definite,  $x^T A x \ge 0$  for all x.

$$\mathbf{v}^T A \mathbf{v} = \mathbf{v}^T \mathbf{v} \lambda$$

Since  $v^T v$  is necessarily a positive number, in order for  $v^T A v$  to be greater than or equal to 0,  $\lambda$  must be  $\lambda \geq 0$ .

• Provide one example for each of the following cases. As for  $(A+B)^2 \neq A^2 + 2AB + B^2$ . Suppose  $A = \begin{bmatrix} 0 & 1 \\ 0 & 0 \end{bmatrix}$  and  $B = \begin{bmatrix} 1 & 0 \\ 0 & 0 \end{bmatrix}$ . Since  $A \cdot B = \begin{bmatrix} 0 & 0 \\ 0 & 0 \end{bmatrix}$  and  $B \cdot A = \begin{bmatrix} 0 & 1 \\ 0 & 0 \end{bmatrix}$ . So  $(A+B)^2 = A^2 + AB + BA + B^2 \neq A^2 + 2AB + B^2$ . As for  $AB = 0, A \neq 0, B \neq 0$ , Suppose  $A = \begin{bmatrix} 0 & 1 \\ 0 & 0 \end{bmatrix}$  and  $B = \begin{bmatrix} 1 & 0 \\ 0 & 0 \end{bmatrix}$ .

• Show that A is orthogonal. Given  $u^{T}u = 1$  and  $A = I - 2uu^{T}$ .

$$\begin{aligned} \boldsymbol{A}^{\mathrm{T}}\boldsymbol{A} &= (\boldsymbol{I} - 2\boldsymbol{u}\boldsymbol{u}^{\mathrm{T}})^{\mathrm{T}}(\boldsymbol{I} - 2\boldsymbol{u}\boldsymbol{u}^{\mathrm{T}}) \\ &= (\boldsymbol{I} - 2\boldsymbol{u}\boldsymbol{u}^{\mathrm{T}})(\boldsymbol{I} - 2\boldsymbol{u}\boldsymbol{u}^{\mathrm{T}}) \\ &= \boldsymbol{I} - 2\boldsymbol{u}\boldsymbol{u}^{\mathrm{T}} - 2\boldsymbol{u}\boldsymbol{u}^{\mathrm{T}} + 4\boldsymbol{u}\boldsymbol{u}^{\mathrm{T}}\boldsymbol{u}\boldsymbol{u}^{\mathrm{T}} \\ &= \boldsymbol{I} - 2\boldsymbol{u}\boldsymbol{u}^{\mathrm{T}} - 2\boldsymbol{u}\boldsymbol{u}^{\mathrm{T}} + 4\boldsymbol{u}\boldsymbol{u}^{\mathrm{T}} \\ &= \boldsymbol{I} \end{aligned}$$

So  $\boldsymbol{A}$  is orthogonal.

• Prove the following assertions. –As for  $f(x) = x^3$  for  $x \ge 0$ .

$$f(x) = x^3$$
$$f''(x) = 6x$$

Since  $x \ge 0$ ,  $f''(x) = 6x \ge 0$ . Thus  $f(x) = x^3$  is convex for  $x \ge 0$ . -As for  $f(x_1, x_2) = max(x_1, x_2)$ . Let  $\mathbf{x} = (x_1, x_2)$  and  $\mathbf{y} = (y_1, y_2)$  and  $\lambda \in [0, 1]$ .

$$f(\lambda \mathbf{x} + (1 - \lambda)\mathbf{y}) = f(\lambda x_1 + (1 - \lambda)y_1, \lambda x_2 + (1 - \lambda)y_2)$$

$$= \max(\lambda x_1 + (1 - \lambda)y_1, \lambda x_2 + (1 - \lambda)y_2)$$

$$\leq \max(\lambda x_1, \lambda x_2) + \max((1 - \lambda)y_1, (1 - \lambda)y_2)$$

$$= \lambda \max(x_1, x_2) + (1 - \lambda)\max(y_1, y_2)$$

$$= \lambda f(\mathbf{x}) + (1 - \lambda)f(\mathbf{y})$$

So  $f(x_1, x_2) = max(x_1, x_2)$  is convex on  $\mathbb{R}^2$ .

-As for function f+g. If univariate functions f and g are convex on S, then  $f''(x) \ge 0$  and  $g''(x) \ge 0$ . So  $(f+g)''(x) = f''(x) + g''(x) \ge 0$ . Thus if univariate functions f and g are convex on S, then f+g is convex on S.

-As for fg. If univariate functions f and g are convex and non-neggetive on S.

$$(fg)''(x) = (f'g + fg')'(x)$$

$$= (f''g + f'g' + f'g' + fg'')(x)$$

$$= (f''g + 2f'g' + fg'')(x)$$

Since f and g have their minimum within S at the same point. Before the minimum point, both f and g are decreasing. After the minimum point, both f and g are increasing. So  $f'g' \geq 0$ ,  $f'' \geq 0$ ,  $g'' \geq 0$ ,  $f \geq 0$  and  $g \geq 0$ . Thus  $(fg)''(x) = (f''g + 2f'g' + fg'')(x) \geq 0$ . Then fg is convex on S.

Find the highest entropy of categorical distribution.
 The entropy of a categorical distribution on K values is defined as

$$H(p) = -\sum_{i=1}^{K} p_i log(p_i)$$

The probability and constrains are defined below:

$$P(X = x_i) = p_i$$
 for  $i = 1, 2, ..., K$ 

s.p. 
$$\begin{cases} \sum_{i=1}^{K} p_i = 1 \\ p_i \ge 0 \quad for \quad i = 1, 2, ..., K \end{cases}$$

Constrain function is

$$\varphi(p_i) = \sum_{i=1}^K p_i - 1 = 0$$

Define Lagrange multipliers:

$$\mathcal{L} = H(p) + \lambda \varphi(p_i)$$

$$= -\sum_{i=1}^{K} p_i log(p_i) + \lambda (\sum_{i=1}^{K} p_i - 1)$$

To find the highest entropy, we should find the point where derivative is 0.

$$\frac{\partial \mathcal{L}}{\partial p_i} = -log(p_i) - 1 + \lambda = 0$$
$$p^* = e^{\lambda - 1}$$

Thus when all  $p_i$ , i = 1, 2, ..., K are equally equal to  $e^{\lambda - 1}$ , the categorical distribution has the highest entropy.

# 1 Locally weighted linear regression(20 points)

• Find an appropriate diagonal matrix W.

$$J(\theta) = (X\theta - y)^T W (X\theta - y)$$

Let W be

$$W = \begin{bmatrix} \frac{1}{2}w^{(1)} & & & & \\ & \frac{1}{2}w^{(2)} & & & \\ & & \ddots & & \\ & & & \frac{1}{2}w^{(i)} \end{bmatrix}$$

X is the  $m \times d$  input matrix and y is a  $m \times 1$  vector.

$$X = \begin{bmatrix} x_1^{(1)} & x_2^{(1)} & x_3^{(1)} & \cdots & x_d^{(1)} \\ x_1^{(2)} & x_2^{(2)} & x_3^{(2)} & \cdots & x_d^{(2)} \\ \vdots & \vdots & \vdots & \cdots & \vdots \\ x_1^{(m)} & x_2^{(m)} & x_3^{(m)} & \cdots & x_d^{(m)} \end{bmatrix}$$

$$y = \begin{bmatrix} y^{(1)} \\ y^{(2)} \\ \vdots \\ y^{(m)} \end{bmatrix}$$

$$\theta = \begin{bmatrix} \theta_{(1)} \\ \theta_{(2)} \\ \vdots \\ \theta_{(d)} \end{bmatrix}$$

$$X\theta - y = \begin{bmatrix} (\theta_1 x_1^{(1)} + \theta_2 x_2^{(1)} + \dots + \theta_d x_d^{(1)}) - y^{(1)} \\ (\theta_1 x_1^{(2)} + \theta_2 x_2^{(2)} + \dots + \theta_d x_d^{(2)}) - y^{(2)} \\ \vdots \\ (\theta_1 x_1^{(m)} + \theta_2 x_2^{(m)} + \dots + \theta_d x_d^{(m)}) - y^{(m)} \end{bmatrix}$$

$$W(X\theta - y) = \begin{bmatrix} \frac{1}{2} w^{(1)} \times ((\theta_1 x_1^{(1)} + \theta_2 x_2^{(1)} + \dots + \theta_d x_d^{(1)}) - y^{(1)}) \\ \frac{1}{2} w^{(2)} \times ((\theta_1 x_1^{(2)} + \theta_2 x_2^{(2)} + \dots + \theta_d x_d^{(2)}) - y^{(2)}) \\ \vdots \\ \frac{1}{2} w^{(m)} \times ((\theta_1 x_1^{(m)} + \theta_2 x_2^{(m)} + \dots + \theta_d x_d^{(m)}) - y^{(m)}) \end{bmatrix}$$

$$(X\theta - y)^{T}W(X\theta - y) = \frac{1}{2}w^{(1)} \times ((\theta_{1}x_{1}^{(1)} + \theta_{2}x_{2}^{(1)} + \dots + \theta_{d}x_{d}^{(1)}) - y^{(1)})^{2}$$
$$+ \frac{1}{2}w^{(2)} \times ((\theta_{1}x_{1}^{(2)} + \theta_{2}x_{2}^{(2)} + \dots + \theta_{d}x_{d}^{(2)}) - y^{(2)})^{2} + \dots$$
$$+ \frac{1}{2}w^{(m)} \times ((\theta_{1}x_{1}^{(m)} + \theta_{2}x_{2}^{(m)} + \dots + \theta_{d}x_{d}^{(m)}) - y^{(m)})^{2}$$

So  $J(\theta) = (X\theta - y)^T W(X\theta - y)$  can be written in the from  $J(\theta) = (X\theta - y)^T W(X\theta - y)$  when choosing W as above.

• Solve for  $\theta$ 

$$J(\theta) = (X\theta - y)^T W (X\theta - y)$$

$$= ((X\theta)^T - y^T) W (X\theta - y)$$

$$= ((X\theta)^T W - y^T W) (X\theta - y)$$

$$= \theta^T X^T W X \theta - 2\theta^T X^T W y + y^T W y$$

$$\nabla_{\theta} J(\theta) = 2X^T W X \theta - 2X^T W y = 0$$
$$\theta = (X^T W X)^{-1} X^T W y$$
$$= (X^T X)^{-1} X^T y$$

Thus 
$$\theta = (X^T X)^{-1} X^T y$$
.

• Write down an algorithm for calculating  $\theta$  by batch gradient descent for locally weighted linear regression. We know that  $\theta = (X^T X)^{-1} X^T y$ .

Algorithm is

**Algorithm 1:** Calculating  $\theta$  by batch gradient descent

$$\begin{split} & \textbf{Input: } x \\ & \textbf{Output: } \theta \\ & \textbf{for } i < 10000 \textbf{ do} \\ & & \middle| \begin{array}{l} \theta_j \leftarrow \theta_j - \alpha \sum_i w^{(i)} (\sum_k (\theta_k x_x^{(i)}) - y^{(i)}), j = 0 \\ & \underline{\theta_j} \leftarrow \theta_j - \alpha \sum_i w^{(i)} (\sum_k (\theta_k x_x^{(i)}) - y^{(i)}) x_j^{(i)}, j \neq 0 \end{array} \end{split}$$

Locally weighted linear regression is a non-parametric method.

## 2 Properties of the linear regression estimator(10 points)

• Show that  $E[\theta] = \theta^*$ 

$$\theta = (X^T X)^{-1} X^T y$$

$$= (X^T X)^{-1} X^T (X \theta^* + \epsilon)$$

$$= (X^T X)^{-1} X^T X \theta^* + (X^T X)^{-1} X^T \epsilon$$

$$E(\theta) = E((X^T X)^{-1} X^T X \theta^* + (X^T X)^{-1} X^T \epsilon)$$

$$= E((X^T X)^{-1} X^T X \theta^*) + E(X^T X)^{-1} X^T \epsilon)$$

$$= \theta^* + (X^T X)^{-1} X^T E(\epsilon)$$

$$= \theta^*$$

• Show the variance of the least squares estimator is  $Var(\theta) = (X^TX)^{-1}\sigma^2$ . Suppose **b** is an estimator of  $\theta$ 

$$Var(\theta) = E[(\boldsymbol{b} - \theta)(\boldsymbol{b} - \theta)^T | X]$$

$$= E[(X^T X)^{-1} X^T \epsilon \epsilon^T X (X^T X)^{-1} | X]$$

$$= (X^T X)^{-1} X^T E[\epsilon \epsilon^T | X] X (X^T X)^{-1}$$

$$= (X^T X)^{-1} \sigma^2$$

# 3 Implementing linear regression and regularized linear regression(90 points)

#### 3.1 Implementing linear regression(45 points)

- A1: Computing the cost function  $J(\theta)$ See the implementation for the loss function in the file linear\_regressor.py.
- A2: Implementing gradient descent See the implementation for the *train* function in the file *linear\_regressor.py*.

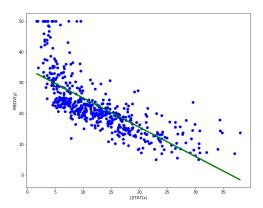


Figure 5: Fitting a linear model

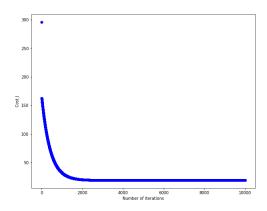


Figure 6: Convergence of gradient descent

- A3: Predicting on unseen data
  For lower status percentage = 5, we predict a median home value of 298034.49
  For lower status percentage = 50, we predict a median home value of -129482.13
- B1: Feature normalization See the implementation for the feature\_normalize function in the file utils.py.
- B2: Loss function and gradient descent
  See the implementation for the train and loss function in the file linear\_regressor\_multi.py.

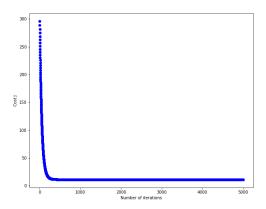


Figure 7: Convergence of gradient descent with multiple variables

- B3: Making predictions on unseen data For average home in Boston suburbs, we predict a median home value of 230467.11
- B4: Normal equations See the implementation for the  $normal_eqn$  function in the file  $linear\_regressor\_multi.py$ . Theta computed by direct solution is: [3.64594884e + 01 1.08011358e 014.64204584e 022.05586264e 022.68673382e + 00 1.77666112e + 013.80986521e + 006.92224640e 04 1.47556685e + 003.06049479e 01 1.23345939e 02 9.52747232e 019.31168327e 03 5.24758378e 01] For average home in Boston suburbs, we predict a median home value of 230406.54.

### • B5: Exploring convergence of gradient descent

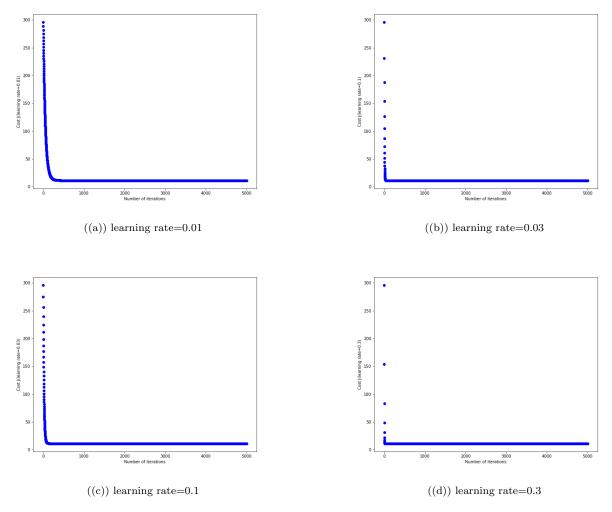


Figure 8: Convergence of gradient descent with multiple variables according to different learning rates According to the above figures, when learning rate = 0.01 is a good choice.

When learning rate = 0.01, chosing different iteration times.

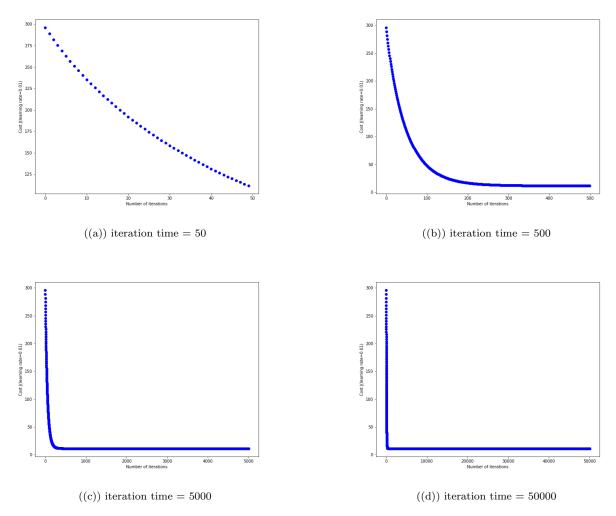


Figure 9: Convergence of gradient descent with multiple variables according to different iteration times According to the above figures, iteration time is 500 is a good choice.

### 3.2 Implementing regularized linear regression(45 points)

- A1: Regularized linear regression cost function
  See the implementation for the loss function in the class Reg\_LinearRegression\_SquaredLoss in the file
  reg\_Linear\_regressor\_multi.py.
- A2: Gradient of the Regularized linear regression cost function
  See the implementation for the grad<sub>l</sub>oss function in the class Reg\_LinearRegression\_SquaredLoss in the file
  reg\_linear\_regressor\_multi.py.
- A3: Learning curves
- A4: Adjusting the regularization parameter

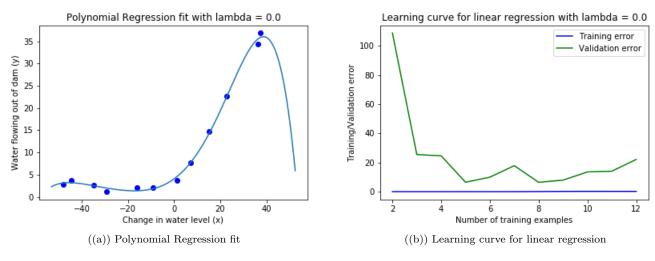


Figure 10:  $\lambda = 0.0$ 

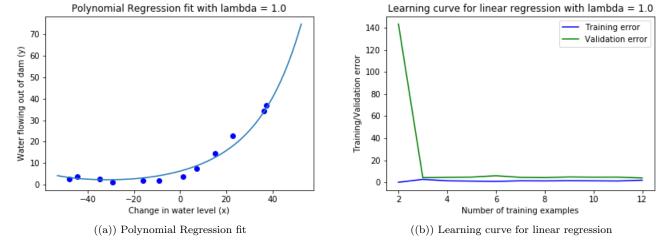


Figure 11:  $\lambda = 1.0$ 

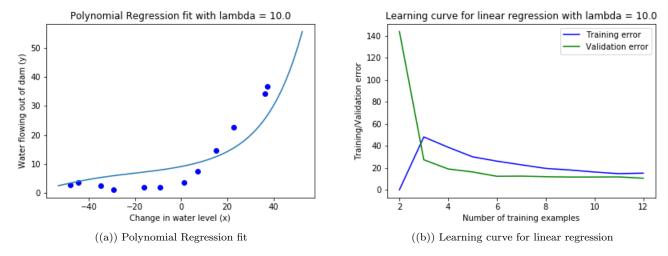


Figure 12:  $\lambda = 10.0$ 

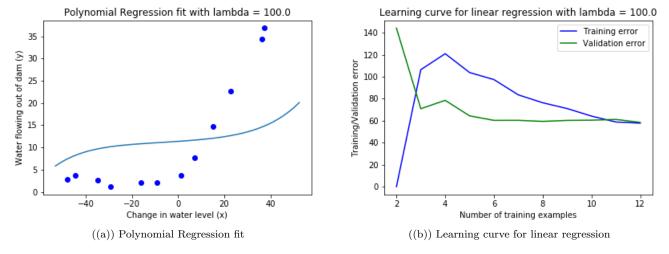


Figure 13:  $\lambda = 100.0$ 

• A5: Selecting  $\lambda$  using a validation set According to the figure above, the best  $\lambda$  is 3.

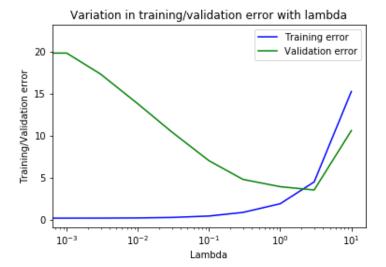


Figure 14: Fitting a linear model

- A6: Computing test set error Test error is 5.944297830865587.
- A7: Plotting learning curves with randomly selected examples
- A8: Comparing ridge regression and lasso regression models Lasso regression converges faster than Ridge regression.

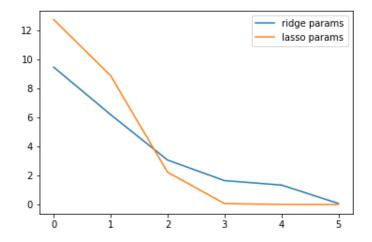


Figure 15: Fitting a linear model

### 4 Extra Credits

• The lowest achievable error on the test set with  $\lambda=0$  Test error is 12.930601133114878 when  $\lambda=0$ .

• Select the best value for  $\lambda$  and report the test set error with the best  $\lambda$ . The best  $\lambda=30.0$ . Test error at  $\lambda=30.0$  is 13.256266000933937.

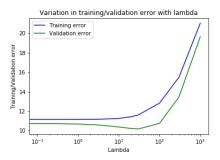


Figure 16: Fitting a linear model

• What is the test set error with quadratic features with the best  $\lambda$  chosen with the validation set? The best  $\lambda = 30.0$ . Test error at  $\lambda = 30.0$  is 4.747514558496794.

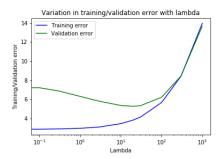


Figure 17: Fitting a linear model

• What is the test set error with cubic features with the best  $\lambda$  chosen with the validation set? The best  $\lambda = 30.0$ . Test error at  $\lambda = 30.0$  is 4.767539493628556.

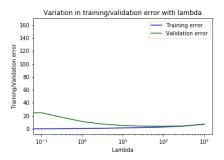


Figure 18: Fitting a linear model