

Q 2.1.) log likelihood function of x given λ

$$p(x = k | \lambda) = \frac{\lambda^k e^{-\lambda}}{k!} \quad k \in \{0, 1, 2, \dots\}$$

$$\text{Likelihood} = \prod_{i=1}^n p(x_i = k_i | \lambda)$$

$$= \prod_{i=1}^n \frac{\lambda^{k_i} e^{-\lambda}}{k_i!}$$

- Since all observations are independent

log likelihood:

$$\ln L(\lambda) = \ln \left(\prod_{i=1}^n \frac{\lambda^{k_i} e^{-\lambda}}{k_i!} \right)$$

$$= \sum_{i=1}^n \ln \left(\lambda^{k_i} e^{-\lambda} \right)$$

$$= \sum_{i=1}^n \left(k_i \ln \lambda - \ln k_i! \right)$$

$$= n \ln \lambda + \left(\sum_{i=1}^n k_i \right) \ln \lambda - \sum_{i=1}^n \ln k_i!$$

(Q4) $\hat{\lambda}$ \approx \bar{x}

$$= \sum_{i=1}^n \left[(\ln e^{-\lambda}) - (\ln k_i!) + (\ln \lambda^{k_i}) \right] \quad (S+SS)$$

$$= \sum_{i=1}^n \left[(-\lambda) - (\ln k_i!) + (k_i \ln \lambda) \right]$$

$$(S+P = \lambda + n\lambda - \sum_{i=1}^n (\ln k_i!) + \ln \lambda \sum_{i=1}^n k_i)$$

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Q2.1.2) MLE for λ

$$\hat{\lambda} = \underset{\lambda}{\operatorname{argmax}} \text{ (likelihood)}$$

$$= \underset{\lambda}{\operatorname{argmax}} \text{ (log likelihood)}$$

Differentiating log likelihood wrt λ ;

$$-n + \frac{1}{\lambda} \sum_{i=1}^n k_i = 0$$

$$\therefore \hat{\lambda} = \frac{1}{n} \sum_{i=1}^n k_i$$

Q 2.1.3) MLE using observed (x)

$$L(\lambda) = \left(\frac{1}{n} \sum_{i=1}^n K e(x_i) \right)^n$$

$$\begin{aligned} L(\lambda) &= \left(\frac{1}{7} (4+5+3+5+6+9+3) \right)^7 \\ &= \frac{1}{7} (35) \\ &= \underline{\underline{5}} \end{aligned}$$

(Maximum) Likelihood = 5

$$J(\lambda) = \frac{1}{n} \sum_{i=1}^n \frac{1}{K}$$

$$\frac{1}{n} \sum_{i=1}^n 1 = R$$

$$Q2.1) P(\lambda | \alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda} \quad \lambda > 0$$

Delays resemble Gamma distribution.

$$\lambda \sim \Gamma(\alpha, \beta)$$

$$\text{mean} = \frac{\alpha}{\beta}$$

$$\text{mode} = \frac{\alpha-1}{\beta} \quad \text{for } \alpha > 1$$

prior distribution of $\lambda = \Gamma(\lambda | \alpha, \beta)$

$$P(x = k | \lambda) = \frac{\lambda^k}{k!} e^{-\lambda}$$

$$P(X=k|\lambda) \\ P(\lambda|x \in \Omega)$$

$$P(\lambda|x=k)$$

$$P(X=k|\lambda) \cdot P(\lambda)$$

$$= P(X=k|\lambda) \cdot P(\lambda)$$

$$\sum_{e=0}^{\infty} x^e e^{\alpha - 1} \cdot \lambda^{n+\beta}$$

$$= \lambda^{\alpha} \cdot \lambda^{-\beta} \cdot e^{\alpha - \beta}$$

$$= (\lambda^{\alpha} \cdot e^{\alpha}) \cdot (\lambda^{-\beta} \cdot e^{-\beta})$$

$$= \lambda^{\alpha - \beta} \cdot e^{\alpha - \beta}$$

$$P(X=k|\lambda) = \frac{\lambda^k}{k!} e^{-\lambda}$$

$$\begin{aligned}
 & \frac{1}{\Gamma(\alpha)} \cdot \lambda^{\alpha} \cdot \frac{\lambda^k}{k!} e^{-\lambda} \\
 & = e^{-B\lambda - \lambda} \cdot \lambda^{d-1+k}, \\
 & \left(\frac{\beta^k}{k! \Gamma(k)} \right)
 \end{aligned}$$

prior of λ $P(\lambda)$

\therefore posterior $P(\lambda | x)$

$P(x|\lambda) \propto$

$$P(\lambda | x) = \frac{P(x | \lambda) P(\lambda)}{P(x)}$$

$$\rightarrow (\lambda^n \cdot e^{-\lambda}) \cdot (\lambda^{d-1} \cdot e^{-B\lambda})$$

$$P(\text{Data} | \text{model}) = P(x=k_0 | \lambda) = \frac{\lambda^k}{k!} e^{-\lambda}$$

$$P(\text{model}) = \int p(\lambda) = \frac{\beta^k}{\Gamma(\alpha)} \lambda^{d-1} e^{-B\lambda}$$

$\therefore P(\text{model data}) =$

$$P(\lambda | x=k_0) = \frac{\lambda^k}{k!} e^{-\lambda}$$

$$\int_0^\infty$$

$$R-5 \quad \frac{6}{2^3} = 8 \quad 2^6 = 8 \times 8 \\ 2^3 = 2^3 \cdot 2^3$$

$$x+1 \rightarrow k, R-R_0=8 \rightarrow$$

Q2.3.1

$$n = e^{-2\lambda} (\lambda + \lambda^2 + \dots) \quad \text{as } R \rightarrow \infty$$

$$X \sim \text{Poisson}(\lambda)$$

$$(x) P(X=k|n) = \frac{\lambda^k e^{-\lambda}}{k!}$$

$$\hat{n} = s e^{-2\lambda} \quad (n - s, \lambda)$$

$$n = e^{-2\lambda}$$

$$R-5 \quad \lambda = \cancel{X} = \cancel{(e^{-2\lambda})^2} = (135000 / 10700) \lambda$$

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$$R-5 \quad e^\lambda = \frac{n}{e^{-2}}$$

$$R-5 \quad \lambda = \ln \left(\frac{n}{e^{-2}} \right)$$

$$R-5 \quad \lambda = (57100 / 62)^2$$

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$$\therefore P(X = k_e | n) = \frac{\left(\ln\left(\frac{n}{e^{-2}}\right)\right)^{k_e}}{K!} \cdot \frac{e^{-2}}{n} = M$$

MLE for $n \rightarrow$

Taking $\log; \text{pd} = 0$

likelihood = argmax M
 n

$$\therefore \text{log likelihood} = \log M$$

$$\rightarrow \cancel{n} K_e \log \left(\log \left(\frac{n}{e^{-2}} \right) \right) - \log K! + \log \left(\frac{e^{-2}}{n} \right)$$

.. ignore $(-\log K!)$

$$= \cancel{n} K_e \log \left(\log \frac{n}{e^{-2}} \right) + \log e^{-2} - \log n$$

$$= (\cancel{n} K_e) \log \left[\log n - \log e^{-2} \right] - \log n$$

~~($\cancel{n} K_e$) per example = ignore $\log e^{-2}$~~

$$= K_e \log \log n - \log n$$

$$\approx \cancel{K_e} \cancel{\log \log n}$$

$$\therefore \frac{d}{dn} = \frac{\cancel{K_e}}{n \log n} - \frac{1}{n} = \frac{1}{n} \left(\frac{K_e}{\log n} - 1 \right) = 0$$

Q2.3.1)

$$M = \frac{1}{2} n$$

g.H

$$\therefore (\log n = k\gamma) \Rightarrow (n \approx e^{k\gamma})$$

$$n = e^{-2\gamma}$$

$$\therefore n^{1/2} = e^{-\gamma}$$

$$n^{-1/2} = e^{\gamma}$$

$$\gamma = \log n$$

$$M \text{ median} = \frac{1}{2} \log n$$

$$P(x = k\gamma | n) = \frac{n^{k\gamma}}{k\gamma!} e^{-n}$$

$$\left(\frac{1}{n}\right)^{k\gamma} + \dots + P(x = k\gamma | n) = \left(\frac{-1/2 \log n}{k\gamma!}\right)^{k\gamma} e^{-1/2 \log n}$$

$$\left(\frac{1}{n}\right)^{k\gamma} + \dots + \left(\frac{-1/2 \log n}{k\gamma!}\right)^{k\gamma} n^{-1/2} =$$

$$\text{likelihood} = \arg \max_n \left(\frac{-1/2 \log n}{k\gamma!} \right)^{k\gamma} n^{-1/2}$$

$$\log \text{likelihood} = \arg \max_n \log \left(\frac{-1/2 \log n}{k\gamma!} \right)^{k\gamma} + \log n$$

$$\left(\frac{1}{n} - \frac{1}{2} \right) \frac{1}{n} = \frac{1}{n} - \frac{1}{2} = \frac{1}{n}$$

$$0 =$$

$$= \underset{n}{\operatorname{argmax}} \left[k^e \log \left(\frac{-1/2 \log n}{k^e!} \right) + -\frac{1}{2} \log n \right]$$

$$\rightarrow 1/2 \log n = m$$

$$= \underset{n}{\operatorname{argmax}} \left[k^e \log(-m) - k^e \log k^e! + \frac{1}{2} m \right]$$

$$m = \frac{1}{2} \log n$$

Differentiating wrt n;

$$\frac{dm}{dn} = 1 \pm \frac{1}{n}$$

$$\frac{d}{dn} (k^e \log (-1/2 \log n) + -1/2 \log n)$$

$$= \frac{k^e}{-1/2 \log n} \cdot -\frac{1}{2} \frac{1}{n} \mp \frac{1}{2n}$$

$$= \frac{1}{2n} \left(\cancel{k^e} - 1 - \frac{k^e}{-1/2 \log n} \right)$$

$$\therefore 1/2 \log n = -k^e = x$$

$$\log n = x^2 - 2x$$

$$n = e^{-2x}$$

Q3.1.1)

$$\epsilon_i \sim \mathcal{N}(0, \sigma_i^2)$$

$\epsilon_1 \dots \epsilon_N$ are independent Gaussians.

$$y_i = w^T x_i + \epsilon_i$$

$$\Omega = (g(x_i) - g_p)^2 \leq \epsilon$$

Now, log likelihood calculation:

$$\Omega = \sum_{i=1}^N \log \left[\frac{1}{2\pi\sigma_i^2} \exp \left(\frac{-1}{2\sigma_i^2} (y_i - w^T x_i)^2 \right) \right]$$

$$= \sum_{i=1}^N \left[\log \left(\frac{1}{2\pi\sigma_i^2} \right)^{1/2} - \frac{1}{2\sigma_i^2} (y_i - w^T x_i)^2 \right]$$

$$= \sum_{i=1}^N \left[\text{some } \sigma_i \text{ term} - \frac{1}{2\sigma_i^2} (y_i - w^T x_i)^2 \right]$$

{
constant}

$$\text{minimize} \sum_{i=1}^N \frac{1}{2\sigma_i^2} (y_i - w^T x_i)^2$$

$$se = -\frac{1}{2\sigma_i^2}$$

$$\frac{\partial}{\partial w} \sum_{i=1}^n s_i (y_i - w x_i)^2 =$$

$$2 \sum_{i=1}^n -x_i (y_i - w x_i)$$

$$\Rightarrow \sum_{i=1}^n 2s_i x_i (y_i - w x_i) = 0$$

$$\sum_{i=1}^n 2s_i x_i y_i - \sum_{i=1}^n 2s_i x_i^2 w = 0$$

$$\sum_{i=1}^n x_i y_i = \sum_{i=1}^n w x_i^2$$

$$w = \frac{\sum_{i=1}^n x_i y_i}{\sum_{i=1}^n x_i^2}$$

Q3.2)

$$P(\epsilon_i) = \frac{1}{2b} \exp\left(\frac{-|\epsilon_i|}{b}\right) = \frac{1}{2b} e^{-\frac{|\epsilon_i|}{b}}$$

$$y(x) = w^T x + \epsilon$$

$$\hat{\theta} = \underset{\theta}{\operatorname{argmax}} \log p(D|\theta)$$

⇒ log likelihood

$$\sum_{i=1}^N \log \left[\frac{1}{2b} e^{-\frac{|\epsilon_i|}{b}} \right]$$

$$\Rightarrow \sum_{i=1}^N \log \left[\frac{1}{2b} e^{-\frac{|y_i - w^T x_i|}{b}} \right]$$

$$= \sum_{i=1}^n \left[\log \left(\frac{1}{2b} \right) + -|y_i - w^T x_i| \right]$$

$$= \frac{1}{b} \sum_{i=1}^n |x_i w - y_i| + n \log \frac{1}{2b}$$

constant

maximize log likelihood

$$= \underset{n}{\operatorname{maximize}} \sum_{i=1}^n |x_i w - y_i|$$

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$$Q4.1.1 \quad r_p = y_p - (w^T x_p + b)$$

E.I.F.Q

y_p is +ve if $w^T x_p + b > 0$ else -ve

$\|x\|_0 = \text{no. of non zero } x \text{ entries}$
 $(n) \circ$

$$R = y - w^T x - B$$

$NX \left[y - w^T x - B \right] \leq \epsilon \quad E.I.F.Q$

Time complexity = $O(\|x\|_0)$
 $(n) \circ$

Q4.1.2

$$B_{\text{new}} = \frac{R}{N} + B_{\text{old}}$$

 $(n) \circ$

Time complexity = $O(n) \rightarrow O(N \cdot \|x\|_0)$

where $B \Rightarrow (N \times \|x\|_0)$

Iteration loop: $R_{\text{new}} = R_{\text{old}} - B_{\text{new}} + B_{\text{old}}$
 $(n) \circ$

Q4.1.3

$$R_{\text{new}} \leftarrow R_{\text{old}} - B_{\text{new}} + B_{\text{old}} \quad (n) \circ$$

 $O(n)$

Q4.1.4

$$C_n = 2 [R + w_k x_k] x_k \quad (n) \circ$$

vector x_k
 $(n) \circ$

$(n) \circ + (n) \circ + (n) \circ + (n) \circ$
 $\downarrow k^{\text{th}} \text{ element of } w$

Q4.1.5

Q 4.1.3

$$(d + \alpha x^T w) = \theta_0 + \theta_1 x^T$$

$$R_{\text{new}} = R_{\text{old}} - B_{\text{new}} + B_{\text{old}}$$

$$\theta_0 + \theta_1 x^T - \theta_0 + \theta_1 x^T = \|x\|$$

$O(n)$

$$\theta - x^T w - \theta = \|x\|$$

$$Q 4.1.4 C_k = 2 [R + w_k x_k] x_k$$

$$(\|x\|) O = \text{matrix multiplication}$$

$O(\Sigma k)$

Q 4.1.5

$$R_{\text{new}} = R_{\text{old}} + w_k^T x_k - w_{\text{new}}^T x_k$$

$$\theta_0 + \theta_1 x^T = \text{matrix multiplication}$$

$O(\Sigma k)$

$$(\|x\|) O = \text{matrix multiplication}$$

Q 4.1.6 1. $O(\|x\|)$ $\in O(n)$

2. $O(n)$ $\rightarrow (n \times d)$ matrix considered

3. $O(n)$

4.

4. $(d)(\Sigma k)$

5. $(d)(\Sigma k)(1)$

6. $(d)(\Sigma k)$

$$\therefore \text{Total time} = O(\|x\|) + O(n) + O(d\Sigma k) + O(d\Sigma k)$$

Q4.3 The Δ to precision-recall relation :

As shown in graphs;

Precision increases, as Δ increases

Recall decreases as Δ increases

Increasing σ to 10 \rightarrow

Precision decreased ;
recall remains same

→ To achieve better precision & recall;

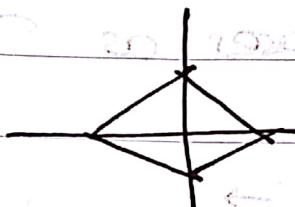
As seen through results ; the Δ should
be increased to increase precision .

There will be a point after which recall
starts falling. Stop increasing Δ at
this point ...

Q 3-3]

Laplace preferred over Gaussian -

- 1) Laplace has thinner tails -
- 2) Laplace solutions intersect the



(curve at

edges; which helps eliminate unnecessary features)

- 3) Sparser solutions --

x_1

$$[0, 1]$$

x_2

$$[0, 1]$$

$$x \rightarrow \max(x_1, 2x_2)$$

$$f_{x_1} = \begin{cases} 1 & 0 < x_1 < 1 \\ 0 & \text{else} \end{cases}$$

$$f_{x_2} = \begin{cases} 1 & 0 < x_2 < 1 \\ 0 & \text{else} \end{cases}$$

$$x_3 = 2x_2$$

$$f_{x_3} = \begin{cases} 1/2 & 0 < x_3 < 2 \\ 0 & \text{else} \end{cases}$$

$$F_{x_1} = \begin{cases} 0 & x_1 < 0 \\ x & 0 < x_1 < 1 \\ 1 & x_1 \geq 1 \end{cases}$$

$$F_{x_3} = \begin{cases} 0 & x_3 < 0 \\ x/2 & 0 < x_3 < 2 \\ 1 & x_3 \geq 2 \end{cases}$$

$$= \left[\frac{x_1}{2} \right] + \left[\frac{x_2}{2} \right] + \left[\frac{x_3}{2} \right]$$

$$f_x = f_{x_1} F_{x_3} + f_{x_3} F_{x_1}$$

$$f_x = \begin{cases} 0 & x < 0 \\ 1/2 & 1 \leq x < 2 \\ 0 & 0 \leq x < 1 \end{cases}$$

$$E(x) = \int_0^1 \frac{3}{4} x^2 dx + \int_1^2 \frac{x^2}{2} dx$$

$$= \left[\frac{x^3}{3} \right]_0^1 + \left[\frac{x^2}{4} \right]_1^2$$

$$= \frac{1}{3} + \frac{3}{4} = \frac{13}{12}$$

$$= \frac{13}{12}$$

$$(13/12)^2$$

$$\text{var}(x) = E(x^2) - (E(x))^2$$

$$= \int_0^1 x^3 dx + \int_1^2 \frac{x^2}{2} dx$$

$$= \left[\frac{x^4}{4} \right]_0^1 + \left[\frac{x^3}{6} \right]_1^2$$

$$= \frac{1}{4} + \frac{8}{6} - \frac{1}{6} = \frac{1}{4} + \frac{7}{6}$$

$$= \frac{6 + 28}{24}$$

$$= \frac{34}{24}$$

$$= \frac{17}{12}$$

$$\bullet \text{ OY}(x, y) = e(x, y) - e(x)e(y)$$

$$\rightarrow \text{OY}(x, x_1) = \cancel{e(x, x_1)} - \cancel{e(x)e(x_1)}$$

$\cancel{13/2} \quad \cancel{1/2}$

$$e(x, x_1) = \cancel{\cancel{xx_1}}$$

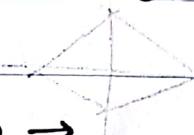
Q4.3 The γ to precision-recall relation: [8-8]

As shown in paragraph; (a)

Precision increases as γ increases

Recall decreases as γ increases

To / from



(changing σ to 10 \rightarrow)

Precision decreased; (b)

recall remains same

→ ADDITIONAL POINTS (c)

→ To achieve better precision & recall;

As seen through results; the γ should be increased to increase precision.

There will be a point after which recall starts falling. Stop increasing γ at this point...

→ Smaller γ takes more iterations to converge -

- Priors added to PDF end of