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Abstract—This manual provides an introduction to linear methods in regression.

1 THE GAUSSIAN DISTRIBUTION

1.1 Generate a Gaussian random number with 0 mean and unit variance.

Solution: Open a text editor and type the following program.

```
#!/usr/bin/env python

#This program generates a Gaussian random
#no with 0 mean and unit variance

#Importing numpy
import numpy as np

print (np.random.normal(0,1))
```

Save the file as gaussian_no.py and run the program.

1.2 The mean of a random variable X is defined as

$$E[X] = \frac{1}{N} \sum_{i=1}^N X_i \quad (1.1)$$

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and its variance as

$$\text{var}[X] = E[X - E[X]]^2 \quad (1.2)$$

Verify that the program in 1.1 actually generates a Gaussian random variable with 0 mean and unit variance.

Solution: Use the header in the previous program, type the following code and execute.

```
#This program generates a Gaussian random
#no with 0 mean and unit variance
```

```
#Importing numpy
import numpy as np
```

```
simlen = int(1e5) #No of samples
```

```
n = np.random.normal(0,1,simlen)#Random
vector
```

```
mean = np.sum(n)/simlen #Mean value
```

```
print (mean)
```

```
var = np.sum(np.square(n - mean*np.ones
((1,simlen))))/simlen
```

```
print (var)
```

1.3 Using the previous program, verify your results for different values of the mean and variance.

2 CDF AND PDF

2.1 A Gaussian random variable X with mean 0 and unit variance can be expressed as $X \sim \mathcal{N}(0, 1)$. Its cumulative distribution function (CDF) is defined as

$$F_X(x) = \Pr(X < x), \quad (2.1)$$

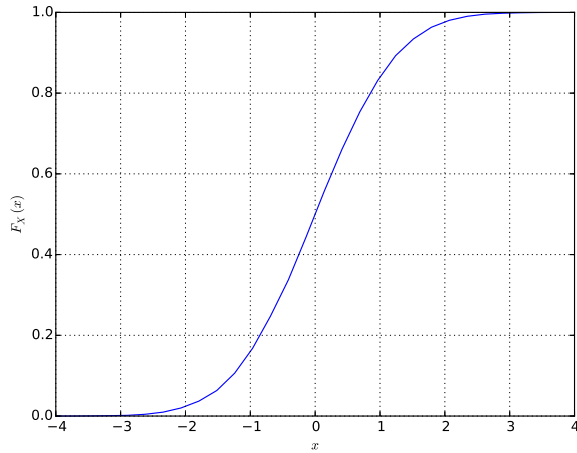


Fig. 2.1: CDF of X

Plot $F_X(x)$.

Solution: The following code yields Fig. 2.1.

```
#Importing numpy, scipy, mpmath and pyplot
import numpy as np
import matplotlib.pyplot as plt

x = np.linspace(-4,4,30)#points on the x axis
simlen = int(1e5) #number of samples
err = [] #declaring probability list
n = np.random.normal(0,1,simlen)

for i in range(0,30):
    err_ind = np.nonzero(n < x[i]) #
        checking probability condition
    err_n = np.size(err_ind) #
        computing the probability
    err.append(err_n/simlen) #storing
        the probability values in a list

plt.plot(x.T,err)#plotting the CDF
plt.grid() #creating the grid
plt.xlabel('$x$')
plt.ylabel('$F_X(x)$')
plt.show() #opening the plot window
```

2.2 List the properties of $F_X(x)$ based on Fig. 2.1.

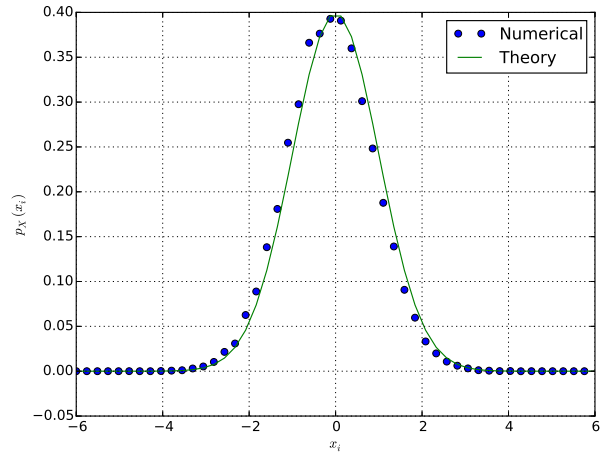


Fig. 2.3: The PDF of X

2.3 Let

$$p_X(x_i) = \frac{F_X(x_i) - F_X(x_{i-1})}{h}, i = 1, 2, \dots, h \quad (2.2)$$

for $x_i = x_{i-1} + h, x_1 = -4$. Plot $p_X(x_i)$. On the same graph, plot

$$p_X(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, -4 < x < 4 \quad (2.3)$$

Solution: The following code yields the graph in Fig. 2.3

https://github.com/gadepall/EE1390/raw/master/manuals/supervised/linear_class/codes/1.4.py

Thus, the PDF is the derivative of the CDF. For $X \sim \mathcal{N}(0, 1)$, the PDF is

$$p_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}, -\infty < x < \infty \quad (2.4)$$

2.4 For $X \sim \mathcal{N}(\mu, \sigma^2)$,

$$p_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, -\infty < x < \infty \quad (2.5)$$

Plot $p_X(x)$ for different values of μ and σ in the same graph. Comment.

3 DETECTION & ESTIMATION

3.1 Use the following code

```
#Importing numpy and pyplot
import numpy as np
import matplotlib.pyplot as plt
```

```
#Function for generating coin toss
def coin(x):
    return 2*np.random.randint(2,size=x)
    -1

simlen = int(1e5)
N = np.random.normal(0,1,simlen)
S = coin(simlen)
A = 4
X = A*S+N
```

to generate X . Obtain a scatterplot of X .

- 3.2 Suppose you wanted to classify X into two groups. How would you do so by looking at the scatterplot?

4 BAYES CLASSIFIER

- 4.1 Let

$$x = A(2s - 1) + n \quad (4.1)$$

where $s \in (0, 1)$, $n \sim \mathcal{N}(0, 1)$.

- 4.2 Show that

$$x|0 \sim \mathcal{N}(-A, 1) \quad (4.2)$$

$$x|1 \sim \mathcal{N}(A, 1) \quad (4.3)$$

- 4.3 Find

$$p_X(x|0) \text{ and } p_X(x|1) \quad (4.4)$$

- 4.4 Find

$$p_X(x|0) \underset{1}{\overset{0}{\geq}} p_X(x|1) \quad (4.5)$$

- 4.5 Show that

$$p_X(0|x) \underset{1}{\overset{0}{\geq}} p_X(1|x) \quad (4.6)$$

$$p_X(x|0) \underset{1}{\overset{0}{\geq}} p_X(x|1) \quad (4.7)$$

if

$$p(0) = p(1) \quad (4.8)$$

5 OPTIMUM CLASSIFIER

- 5.1 Let (\mathbf{X}, \mathbf{G}) be an input/output dataset, whose relation f is unknown. Also

$$\mathbf{g} \in \mathbf{G} = \{\mathbf{g}_k\}_{k=1}^K \quad (5.1)$$

Let

$$C(\mathbf{g}_k, \mathbf{g}_l) = \begin{cases} 1 & k = l \\ 0 & k \neq l \end{cases} \quad (5.2)$$

where \mathbf{g}_i are different classes of output data. Thus C is a *correctness* metric.

- 5.2 Show that

$$\max_{\mathbf{g} \in \mathbf{G}} E[C(\mathbf{G}, f(\mathbf{X}))] = \max_{\mathbf{g} \in \mathbf{G}} p(\mathbf{g}|\mathbf{X} = \mathbf{x}) \quad (5.3)$$

Solution: In the above,

$$\begin{aligned} \max_{\mathbf{g} \in \mathbf{G}} E[C(\mathbf{G}, f(\mathbf{X}))] \\ = \max_{\mathbf{g} \in \mathbf{G}} E_X[E_G\{C(\mathbf{G}, f(\mathbf{x}))\}] \end{aligned} \quad (5.4)$$

$$= \max_{\mathbf{g} \in \mathbf{G}} \sum_{k=1}^K C\{\mathbf{g}_k, \mathbf{g}\} p(\mathbf{g}_k|\mathbf{X} = \mathbf{x}) \quad (5.5)$$

From (4.10), the above expression simplifies to

$$\max_{\mathbf{g} \in \mathbf{G}} E[C(\mathbf{G}, f(\mathbf{X}))] = \max_{\mathbf{g} \in \mathbf{G}} p(\mathbf{g}|\mathbf{X} = \mathbf{x}) \quad (5.6)$$

6 LINEAR DISCRIMINANT ANALYSIS

- 6.1 The multivariate Gaussian distribution is defined as

$$\begin{aligned} p_{\mathbf{x}}(\mathbf{x}) \\ = \frac{1}{\sqrt{(2\pi)^k |\Sigma|}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right\} \end{aligned} \quad (6.1)$$

where $\boldsymbol{\mu}$ is the mean vector, $\Sigma = E[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^T]$ is the covariance matrix and $|\Sigma|$ is the determinant of Σ .

- 6.2 For

$$\mathbf{x} = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}, \quad (6.2)$$

show that

$$\begin{aligned} p_{\mathbf{x}}(\mathbf{x}) &= \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp \left[-\frac{1}{2(1-\rho^2)} \right. \\ &\times \left. \left\{ \frac{(x_1 - \mu_1)^2}{\sigma_1^2} + \frac{(x_2 - \mu_2)^2}{\sigma_2^2} - \frac{2\rho(x_1 - \mu_1)(x_2 - \mu_2)}{\sigma_1\sigma_2} \right\} \right] \end{aligned} \quad (6.3)$$

where

$$\boldsymbol{\mu} = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \Sigma = \begin{pmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{pmatrix} \quad (6.4)$$

6.3 Let

$$\mathbf{s}_0 = \begin{pmatrix} a \\ 0 \end{pmatrix} \quad (6.5)$$

$$\mathbf{s}_1 = \begin{pmatrix} 0 \\ a \end{pmatrix} \quad (6.6)$$

If

$$\mathbf{x} = \mathbf{s} + \mathbf{n} \quad (6.7)$$

where $\mathbf{s} \in \{\mathbf{s}_0, \mathbf{s}_1\}$ and $\mathbf{n} \sim \mathcal{N}(0, \sigma^2 \mathbf{I})$, show that

$$\mathbf{x}|0 = \begin{pmatrix} a + n_1 \\ n_2 \end{pmatrix}, \quad (6.8)$$

and

$$\mathbf{x}|1 = \begin{pmatrix} n_1 \\ a + n_2 \end{pmatrix}, \quad (6.9)$$

6.4 Find

$$p_{\mathbf{x}|\mathbf{s}_0}(\mathbf{x}) \quad (6.10)$$

$$p_{\mathbf{x}|\mathbf{s}_1}(\mathbf{x}) \quad (6.11)$$

6.5 How will you decide between \mathbf{s}_0 and \mathbf{s}_1 if you have \mathbf{x} ?