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# Section 1.6 Independent of Events

One has occurred does not affact the prob. of the other's occurance.  $\frac{P(AB)}{P(A)} = P(B|A) = P(B)$  P(AB) = P(B) = P(A)

$$P(AB) = P(A)P(B)$$

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Sometimes an event can affect the next event. We call those Dependent Events, because what happens depends on what happened before.

Consider the following examples:

- (i) Landing on heads after tossing a coin AND rolling a 5 on a single 6-sided die.
- (ii) Choosing a marble from a jar AND landing on heads after tossing a coin.
- (iii) Choosing a 3 from a deck of cards, replacing it, AND then choosing an ace as the second card.
- (iv) Rolling a 4 on a single 6-sided die, AND then rolling a 1 on a second roll of the die.

Independent Events are not affected by previous events.

This is an important idea!

The "operational independence" described indicates that knowledge that one of the events has occurred does not affect the likelihood that the other will occur. For a pair of events A and B, this is the condition

$$P(A|B) = P(A) \text{ or } P(B|A) = P(B).$$

Occurrence of the event A is not "conditioned by" occurrence of the event B. If the occurrence of B does not affect the likelihood of the occurrence of A, we should be inclined to think of the events A and B as being independent in a probability sense.

We take our clue from the condition P(A|B) = P(A) to define the independence of two events.

#### Definition

Two events A and B are independent if

$$P(AB) = P(A)P(B). (1)$$

It is obvious that if P(A) > 0 and P(B) > 0, then independence is equivalent to the statement that the conditional probability of one event given the other is the same as the unconditional probability of the event:

$$P(A|B) = P(A) \Leftrightarrow P(B|A) = P(B) \Leftrightarrow P(AB) = P(A)P(B).$$

The independence of the events in the examples at the beginning of this section are obvious. But things do not always like those.

#### Example

Let three fair coins be tossed.

Let  $A = \{$  all heads or all tails  $\}$ ,  $B = \{$  at least two heads  $\}$ , and  $C = \{$  at most two tails  $\}$ . Of the pairs of events, (A, B), (A, C), and (B, C), which are independent and which are dependent?

**Solution**. The events can be written explicitly:

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A = \{HHH, TTT\}, B = \{HHH, HHT, HTH, THH\}, C = \{HHH, HHT, HTH, THH, HTT, THT, TTH\}.
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Solution. Then

$$P(AB) = 1/8 = (2/8)(4/8) = P(A) \cdot P(B),$$
  
 $P(AC) = 1/8 \neq (2/8)(7/8) = P(A) \cdot P(C),$   
 $P(BC) = 4/8 \neq (4/8)(7/8) = P(B) \cdot P(C).$ 

Thus A and B are independent. A and C are dependent. B and C are dependent.  $\Box$ 

The terms independent and disjoint sound vaguely similar but they are actually very different.

#### Theorem

Suppose that A and B are disjoint events for an experiment, each with positive probability. Then A and B are dependent.

**Proof.** Since A and B are disjoint events,  $P(AB) = P(\emptyset) = 0$ . While P(A)P(B) > 0 because P(A) > 0, P(B) > 0. So  $P(AB) \neq P(A)P(B)$ .

#### Theorem

If A and B are independent events in an experiment, then each of the following pairs of events is independent: (i)  $\overline{A}$  and B; (ii) A and  $\overline{B}$ ; (iii)  $\overline{A}$  and  $\overline{B}$ .

**Proof.** We prove case (i). Suppose that A and B are independent. Then by the difference rule,

$$P(\overline{A}B) = P(B) - P(AB) = P(B) - P(A)P(B)$$
$$= [1 - P(A)]P(B) = P(\overline{A})P(B).$$

Hence  $\overline{A}$  and B are independent.

(ii) and (iii) can be obtained by the similar way.

A,Bindep. => A, Bindep. Aim: P(AB) = P(A) P(B)  $P.H.S=P(\overline{A})P(B)=(I-P/AI)P(B)$ = P(B) - P(A) P(B)= P(B) - P(AB)

= P(AB) = L.H.S.

#### Example

Let A and B be independent events with P(A) = 1/4 and  $P(A \cup B) = 2P(B) - P(A)$ . Find the values of P(B), P(A|B) and  $P(\overline{B}|A)$ .

**Solution**. Since A and B are independent,

$$P(A \cup B) = P(A) + P(B) - P(AB) = P(A) + P(B) - P(A)P(B).$$

Thus,

$$1/4 + P(B) - (1/4)P(B) = 2P(B) - 1/4.$$

This implies that P(B) = 2/5.

Since A and B are independent,

$$P(A|B) = P(A) = 1/4$$
 and  $P(\overline{B}|A) = P(\overline{B}) = \frac{3}{5}$ .

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How to define the independence of three events A, B, C? How to express the occurrence of any event of the three events has nothing to do with the occurrence of the other two events?

#### Definition

Three events A, B and C in the sample space S of a random experiment are said to be **mutually independent** (or independent) if

$$\begin{cases} P(AB) = P(A)P(B), \\ P(AC) = P(A)P(C), \\ P(BC) = P(B)P(C), \\ P(ABC) = P(A)P(B)P(C). \end{cases}$$

Events A, B and C are said to be **pairwise independent** if the first three equations hold.

# (HW)

#### Example

Suppose that a balanced coin is independently tossed two times. Define the following events:

A: Head appears on the first toss;

B: Head appears on the second toss;

C: Both tosses yield the same outcome.

Are events A, B and C mutually independent?

**Solution**. The sample space  $\Omega$  consists of the following outcomes:  $\Omega = \{HH, HT, TH, TT\}$  and the events listed above are  $A = \{HH, HT\}, B = \{TH, HH\}, C = \{HH, TT\}.$  Since the coin is balanced, that is, all outcomes are assigned the same probability, namely 1/4, we get

Solution.

$$P(A) = P(B) = P(C) = 2/4 = 1/2,$$
  
 $P(AB) = P(AC) = P(BC) = \frac{1}{4},$   
 $P(ABC) = P(\{HH\}) = 1/4.$ 

Therefore,

$$P(AB) = P(A)P(B), \quad P(AC) = P(A)P(C),$$
 
$$P(BC) = P(B)P(C).$$

However,

$$P(ABC) \neq P(A)P(B)P(C)$$
.

So A, B and C are pairwise independent.

#### Example

Toss two different standard dice, white and black. Define the following events:

 $A_1 = \{ \text{first die} = 1, 2 \text{ or } 3 \};$ 

 $A_2 = \{ \text{first die} = 3, 4 \text{ or } 5 \};$ 

 $A_3 = \{\text{sum of faces is 9}\}.$ 

Are events  $A_1$ ,  $A_2$  and  $A_3$  mutually independent?

**Solution**. The sample space  $\Omega$  of the outcomes consists of all ordered pairs  $(i, j), i, j = 1, \dots, 6$ , i.e.,

 $\Omega = \{(1,1), (1,2), ..., (6,6)\}$ . Since the two dice are standard, each point in S has probability 1/36. So the probability of the events are

$$P(A_1) = 1/2, P(A_2) = 1/2, P(A_3) = 1/9.$$

**Solution**. Obviously,

$$A_1 A_2 = \{(3,1), (3,2), (3,3), (3,4), (3,5), (3,6)\},$$
  
 $A_1 A_3 = \{(3,6)\}, A_2 A_3 = \{(3,6), (4,5), (5,4)\},$   
 $A_1 A_2 A_3 = \{(3,6)\}.$ 

It follows that

$$P(A_1A_2A_3) = 1/36 = (1/2)(1/2)(1/2) = P(A_1)P(A_2)P(A_3).$$

However,

$$P(A_1A_2) = 1/6 \neq 1/4 = P(A_1)P(A_2),$$
  
 $P(A_1A_3) = 1/36 \neq 1/18 = P(A_1)P(A_3),$   
 $P(A_2A_3) = 1/12 \neq 1/18 = P(A_2)P(A_3).$ 

So events  $A_1$ ,  $A_2$  and  $A_3$  are not mutually independent.

# $C_n^2$

#### Definition

Events  $A_1, A_2, \dots, A_n$  are **pairwise independent** if every possible pair of these events are independent, i.e.,  $P(A_jA_k) = P(A_j) \cdot P(A_k)$  for all  $j, k(j \neq k)$ .

#### Definition

The *n* events  $A_1, A_2, \dots, A_n$  are **independent** if, for every subset  $A_{i_1}, A_{i_2}, \dots, A_{i_k}$  of these events  $(k = 2, 3, \dots, n)$ ,

$$P(A_{i_1}A_{i_2}\cdots A_{i_k}) = P(A_{i_1})P(A_{i_2})\cdots P(A_{i_k}).$$
 (2)

$$C_n^2 + C_n^3 + ... + C_n^n = 2^n - 1 - n$$

$$\left( \left| + 1 \right| \right)^n = C_0 + \dots + C_n$$

$$(a+b)^n = C_n^0 a^n b^0 + ... + C_n^n a^0 b^n$$

#### Theorem

The *n* events  $A_1, A_2, \dots, A_n$  are **independent** if and only if for any subset  $A_{i_1}, A_{i_2}, \dots, A_{i_k}$ ,

$$P(A_{i_1}^* A_{i_2}^* \cdots A_{i_k}^*) = P(A_{i_1}^*) P(A_{i_2}^*) \cdots P(A_{i_k}^*), \quad (k = 2, 3, \dots, n)$$
(3)

where  $A_i^*$  denotes either  $A_i$  or  $\overline{A_i}$  the same on both sides of the equation).

#### Example

Suppose that A, B and C are three independent events such that P(A) = 1/4, P(B) = 1/3 and P(C) = 1/2. (a) Determine the probability that none of these three events will occur; (b) Determine the probability that exact one of these three events will occur.

**Solution**. (a) Since A, B and C are three independent events,

$$P(\overline{A}\ \overline{B}\ \overline{C}) = P(\overline{A})P(\overline{B})P(\overline{C}) = (1-1/4)(1-1/3)(1-1/2) = 1/4.$$

(b) By additivity of probability and independency,

$$P(A\overline{B}\ \overline{C} \cup \overline{A}B\overline{C} \cup \overline{A}\ \overline{B}C)$$

$$= P(A\overline{B}\ \overline{C}) + P(\overline{A}B\overline{C}) + P(\overline{A}\ \overline{B}C)$$

$$= P(A)P(\overline{B})P(\overline{C}) + P(\overline{A})P(B)P(\overline{C}) + P(\overline{A})P(B)P(\overline{C})$$

**Solution**. (b) By additivity of probability and independency,

$$P(A\overline{B} \ \overline{C} \cup \overline{A}B\overline{C} \cup \overline{A} \ \overline{B}C)$$

$$= P(A\overline{B} \ \overline{C}) + P(\overline{A}B\overline{C}) + P(\overline{A} \ \overline{B}C)$$

$$= P(A)P(\overline{B})P(\overline{C}) + P(\overline{A})P(B)P(\overline{C}) + P(\overline{A})P(\overline{B})P(C)$$

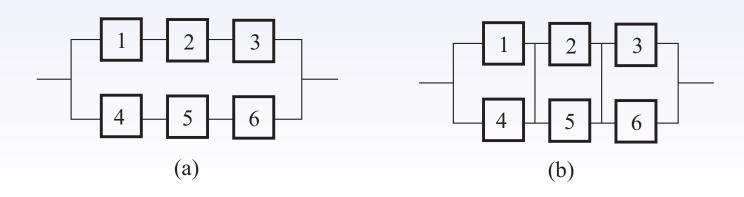
$$= (1/4)(1 - 1/3)(1 - 1/2) + (1 - 1/4)(1/3)(1 - 1/2)$$

$$+ (1 - 1/4)(1 - 1/3)(1/2)$$

$$= 1/12 + 1/8 + 1/4 = 11/24.$$

#### Example

(Reliability analysis) We have two systems illustrated in following Figure. Given a particular lifetime value  $t_0$ , let  $A_i$  denote the event that the lifetime of cell i exceeds  $t_0 (i = 1, 2, \dots, 6)$ . Assume that the  $A_i s$  are independent events (whether any particular cell lasts more than  $t_0$  hours has no bearing on whether or not any other cell does) and that  $P(A_i) = 0.9$  for every i since the cells are identical. Determine the probabilities that the systems' lifetime exceeds  $t_0$ . Which system is more reliable?



$$P(Ai) = \Upsilon$$

$$P(A_{1}...A_{n})U(A_{n+1}...A_{2n})) = 2r^{n} - r^{2n}$$

$$= r^{n}(2-r^{n})$$

$$P((A_{1}UA_{n+1})) (A_{n}UA_{2n})$$

$$= (2r - r^{2})^{n} = r^{n}(2-r)^{n}$$

**Solution**. For the first system,

$$P(\text{system lifetime exceeds } t_0)$$
=  $P((A_1 \cap A_2 \cap A_3) \cup (A_4 \cap A_5 \cap A_6))$   
=  $P(A_1 \cap A_2 \cap A_3) + P(A_4 \cap A_5 \cap A_6) - P((A_1 \cap A_2 \cap A_3) \cap (A_4 \cap A_5 \cap A_6))$   
=  $P(A_1)P(A_2)P(A_3) + P(A_4)P(A_5)P(A_6) - P(A_1)P(A_2)P(A_3)P(A_4)P(A_5)P(A_6)$   
=  $0.9^3 + 0.9^3 - 0.9^6 \approx 0.927$ .

Alternatively,

$$P(\text{system lifetime exceeds } t_0)$$
  
=  $1 - P(\text{both subsystem lives are } \leq t_0)$   
=  $1 - [P(\text{subsystem lives are } \leq t_0)]^2$   
=  $1 - [1 - P(\text{subsystem lives are } > t_0)]^2$   
=  $1 - [1 - (0.9)^3]^2 = 0.927$ .

**Solution**. For the second configuration,

$$P(\text{system lifetime exceeds } t_0)$$
=  $P((A_1 \cup A_4) \cap (A_2 \cup A_5) \cap (A_3 \cup A_6))$   
=  $P(A_1 \cup A_4) \cdot P(A_1 \cup A_4) \cdot P(A_1 \cup A_4)$   
=  $[P(A_1) + P(A_4) - P(A_1A_4)]^3$   
=  $(0.9 + 0.9 - 0.9 \cdot 0.9)^3 \approx 0.97$ 

In reliability theory, we call the probability  $P(\text{system lifetime exceeds } t_0)$  the reliability of the system. Since the reliability of the second system is larger than the first, the second system is more reliable.

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#### Consider the following experiments:

- Flipping a coin and observing which side will appears.
- Rolling a die and observing whether the number six appears.
- In conducting a political opinion poll, choosing a voter at random to ascertain whether that voter will vote "yes" in an upcoming referendum.

All above examples have the same character: the experiment E that has *only two outcomes*. A **Bernoulli experiment** E is such kind of random experiment, the outcome of which can be classified in but one of two mutually exclusive and exhaustive ways, mainly, *success* (denoted by E) or *failure* (denoted by E) (e.g., female or male, life or death, nondefective or defective).





A sequence of Bernoulli trials  $E_n$  occurs when a Bernoulli experiment is performed several independent times so that the probability of success, say, p, remains the same from trial to trial. That is, in such a sequence we let p denote the probability of success on each trial. In addition, frequently q = 1 - p denote the probability of failure; that is, we shall use q and 1 - p interchangeably.

#### Theorem

The probability that the outcome of an experiment that consists of n Bernoulli trials has k successes and n - k failures is given by:

$$P_n(k) := P_n(\#S = k) = \binom{n}{k} p^k q^{n-k}.$$

Here #S means the number of successes.

**Proof.** Now, suppose that, when we performed the n trials, we first obtained k consecutive successes S and then n-k consecutive failures F. By independence,

$$P(\underbrace{SS\cdots S}_{k \text{ times}} \underbrace{FF\cdots F}_{(n-k) \text{ times}}) = P(S)^k P(F)^{n-k} = p^k (1-p)^{n-k}.$$

Hence, given that we can place the k successes among the n trials in  $\binom{n}{k}$  different ways, we deduce that

$$P_n(\#S = k) = \binom{n}{k} p^k q^{n-k}.$$

$$X = \# Success \quad \text{within } \{E_n\}$$

$$P(X=k) = \binom{k}{n} p^k (k-p)^{n-k}$$

$$\sum_{k=0}^{n} C_{n}^{k} a^{k} b^{n-k} = (a+b)^{n}$$

$$1 = \sum_{k=0}^{n} C_{k} p^{k} (1-p)^{n-k}$$

Bernoulli distribution.

Binomial distr. 二项分布

# Summary

- A and B are independent if and only if P(AB) = P(A)P(B).
- The *n* events  $A_1, A_2, \dots, A_n$  are **independent** if and only if for any subset  $A_{i_1}, A_{i_2}, \dots, A_{i_k}$ ,

$$P(A_{i_1}^* A_{i_2}^* \cdots A_{i_k}^*) = P(A_{i_1}^*) P(A_{i_2}^*) \cdots P(A_{i_k}^*), \quad (k = 2, 3, \dots, n)$$

where  $A_i^*$  denotes either  $A_i$  or  $\overline{A}_i$  (the same on both sides of the equation).

Ex. 
$$\{E_n\}$$
...

 $X = \# \text{ times till } 1^{St} \text{ success.}$ 
 $P(X = k) = (1-p)^{k-1} p$ ,

 $Y = \# \text{ times till } 1^{St} \text{ success.}$ 
 $Y = \# \text{ times till } 1^{St} \text{ success.}$ 
 $Y = \# \text{ times till } 1^{St} \text{ success.}$ 

# Thank you for your patience!