

# Regression

COMP9417 Machine Learning and Data Mining

Term 2, 2019

# Acknowledgements

Material derived from slides for the book  
"Elements of Statistical Learning (2nd Ed.)" by T. Hastie,  
R. Tibshirani & J. Friedman. Springer (2009)  
<http://statweb.stanford.edu/~tibs/ElemStatLearn/>

Material derived from slides for the book  
"Machine Learning: A Probabilistic Perspective" by P. Murphy  
MIT Press (2012)  
<http://www.cs.ubc.ca/~murphyk/MLbook>

Material derived from slides for the book  
"Machine Learning" by P. Flach  
Cambridge University Press (2012)  
<http://cs.bris.ac.uk/~flach/mlbook>

Material derived from slides for the book  
"Bayesian Reasoning and Machine Learning" by D. Barber  
Cambridge University Press (2012)  
<http://www.cs.ucl.ac.uk/staff/d.barber/brml>

Material derived from slides for the book  
"Machine Learning" by T. Mitchell  
McGraw-Hill (1997)  
<http://www-2.cs.cmu.edu/~tom/mlbook.html>

Material derived from slides for the course  
"Machine Learning" by A. Srinivasan  
BITS Pilani, Goa, India (2016)

# Aims

After a brief introduction to this course and the topics in it, this lecture will introduce you to machine learning approaches to the problem of numerical prediction. Following it you should be able to reproduce theoretical results, outline algorithmic techniques and describe practical applications for the topics:

- the supervised learning task of numeric prediction
- how linear regression solves the problem of numeric prediction
- fitting linear regression by least squares error criterion
- non-linear regression via linear-in-the-parameters models
- parameter estimation for regression
- local (nearest-neighbour) regression

**Note: slides with titles marked \* are for background only.**

# A Brief Course Introduction

# Overview

This course will introduce you to machine learning, covering some of the core ideas, methods and theory currently used and understood by practitioners, including, but not limited to:

- categories of learning (supervised learning, unsupervised learning, etc.)
- widely-used machine learning techniques and algorithms
- batch vs. online settings
- parametric vs. non-parametric approaches
- generalisation in machine learning
- training, validation and testing phases in applications
- limits on learning

# What we will cover

- core algorithms and model types in machine learning
- foundational concepts regarding learning from data
- relevant theory to inform and generalise understanding
- practical applications

# What we will NOT cover

- lots of probability and statistics
- lots of neural nets and deep learning
- “big data”
- commercial and business aspects of “analytics”
- ethical aspects of AI and ML

although all of these are interesting and important topics!

## Some history

*One can imagine that after the machine had been in operation for some time, the instructions would have been altered out of recognition, but nevertheless still be such that one would have to admit that the machine was still doing very worthwhile calculations. Possibly it might still be getting results of the type desired when the machine was first set up, but in a much more efficient manner. In such a case one would have to admit that the progress of the machine had not been foreseen when its original instructions were put in. It would be like a pupil who had learnt much from his master, but had added much more by his own work.*

From A. M. Turing's lecture to the London Mathematical Society. (1947)



## Some history

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From A. M. Turing's lecture to the London Mathematical Society. (1947)

# Some definitions

*The field of machine learning is concerned with the question of how to construct computer programs that automatically improve from experience.*

“Machine Learning”. T. Mitchell (1997)

*Machine learning, then, is about making computers modify or adapt their actions (whether these actions are making predictions, or controlling a robot) so that these actions get more accurate, where accuracy is measured by how well the chosen actions reflect the correct ones.*

“Machine Learning”. S. Marsland (2015)

## Some definitions

*Machine learning is the systematic study of algorithms and systems that improve their knowledge or performance with experience.*

Machine Learning". P. Flach (2012)

*The term machine learning refers to the automated detection of meaningful patterns in data.*

"Understanding Machine Learning". S. Shalev-Shwartz and S. Ben-David (2014)

*Data mining is the extraction of implicit, previously unknown, and potentially useful information from data.*

"Data Mining". I. Witten *et al.* (2016)

# Machine Learning is ...

*Trying to get programs to work in a reasonable way to predict stuff.*

R. Kohn (2015)

# How is Machine Learning different from ...

Machine learning comes originally from Artificial Intelligence (AI), where the motivation is to build intelligent agents, capable of acting autonomously. Learning is a characteristic of intelligence, so to be successful an agent must ultimately be able to learn, *apply*, *understand* and *communicate* what it has learned.

These are not requirements in:

- statistics — the results are typically mathematical models for humans
- data mining — the results are typically models of “insight” for humans

These criteria are often also necessary, but not always sufficient, for machine learning.

# Supervised and unsupervised learning

The most widely used categories of machine learning algorithms are:

- *Supervised learning* – output class (or label) is given
- *Unsupervised learning* – no output class is given

There are also hybrids, such as semi-supervised learning, and alternative strategies to acquire data, such as reinforcement learning and active learning.

Note: output class can be real-valued or discrete, scalar, vector, or other structure ...

## Supervised and unsupervised learning

Supervised learning tends to dominate in applications.

Why ?

Generally, because it is much easier to define the problem and develop an error measure (loss function) to evaluate different algorithms, parameter settings, data transformations, etc. for supervised learning than for unsupervised learning.

## Supervised and unsupervised learning

Unfortunately ...

In the real world it is often difficult to obtain good labelled data in sufficient quantities

So in such cases unsupervised learning is really what you want ...

but currently, finding good unsupervised learning algorithms for complex machine learning tasks remains a research challenge.



# Machine learning models

Machine learning models can be distinguished according to their main intuition, for example:

- *Geometric* models use intuitions from geometry such as separating (hyper-)planes, linear transformations and distance metrics.
- *Probabilistic* models view learning as a process of reducing uncertainty, modelled by means of probability distributions.
- *Logical* models are defined in terms of easily interpretable logical expressions.

## Machine learning models

Alternatively, can be characterised by *algorithmic properties*:

- *Regression models* predict a numeric output
- *Classification models* predict a discrete class value
- *Neural networks* learn based on a biological analogy
- *Local models* predict in the local region of a query instance
- *Tree-based models* partition the data to make predictions
- *Ensembles* learn multiple models and combine their predictions

# Introduction to Regression

# Introduction to Regression

The “most typical” machine learning approach is to apply supervised learning methods for *classification*, where the task is to learn a model to predict a *discrete* value for data instances ...

... however, we often find tasks where the most natural representation is that of *prediction of numeric values*

## Introduction to Regression

Example – task is to learn a model to predict CPU performance from a dataset of example of 209 different computer configurations:

	Cycle time (ns)	Main memory (Kb)		Cache (Kb)	Channels		Performance
	MYCT	MMIN	MMAX	CACH	CHMIN	CHMAX	PRP
1	125	256	6000	256	16	128	198
2	29	8000	32000	32	8	32	269
...							
208	480	512	8000	32	0	0	67
209	480	1000	4000	0	0	0	45

## Introduction to Regression

Result: a linear regression equation fitted to the CPU dataset.

$$\begin{aligned} \text{PRP} = & \\ & - 56.1 \\ & + 0.049 \text{ MYCT} \\ & + 0.015 \text{ MMIN} \\ & + 0.006 \text{ MMAX} \\ & + 0.630 \text{ CACH} \\ & - 0.270 \text{ CHMIN} \\ & + 1.46 \text{ CHMAX} \end{aligned}$$

# Introduction to Regression

For the class of *symbolic* representations, machine learning is viewed as:

searching a space of **hypotheses** ...

represented in a formal hypothesis language (trees, rules, graphs ...).

# Introduction to Regression

For the class of *numeric* representations, machine learning is viewed as:

“searching” a space of **functions** ...

represented as mathematical models (linear equations, neural nets, ...).

Note: in both settings, the models may be probabilistic ...



# Introduction to Regression

Methods to predict a numeric output from statistics and machine learning:

- linear regression (statistics) determining the “line of best fit” using the least squares criterion
- linear models (machine learning) learning a predictive model from data under the assumption of a linear relationship between predictor and target variables

Very widely-used, many applications

Ideas that are generalised in Artificial Neural Networks

# Introduction to Regression

Regression as a term occurs in many areas of machine learning:

- non-linear regression    by adding non-linear basis functions
- multi-layer neural networks (machine learning)    learning non-linear predictors via hidden nodes between input and output
- regression trees (statistics / machine learning)    tree where each leaf predicts a numeric quantity
- local (nearest-neighbour) regression

# Learning Linear Regression Models

# Regression

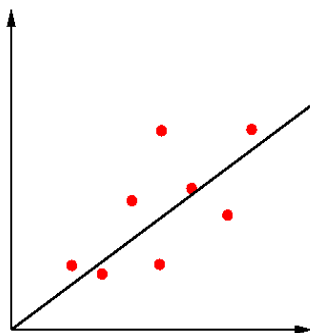
We will look at the simplest model for numerical prediction:  
*a regression equation*

Outcome will be a linear sum of feature values with appropriate weights.

Note: the term *regression* is overloaded – it can refer to:

- the process of determining the weights for the regression equation, or
- the regression equation itself.

# Linear Regression



inputs	outputs
$x1 = 1$	$y1 = 1$
$x2 = 3$	$y2 = 2.2$
$x3 = 2$	$y3 = 2$
$x4 = 1.5$	$y4 = 1.9$
$x5 = 4$	$y5 = 3.1$

Assumes: expected value of the output given an input,  $E[y|x]$ , is linear.

Simplest case:  $\text{Out}(x) = bx$  for some unknown  $b$ .

Learning problem: given the data, estimate  $b$  (i.e.,  $\hat{b}$ ).

# Linear Models

- Numeric attributes and numeric prediction, i.e., regression
- Linear models, i.e. outcome is *linear* combination of attributes

$$y = b_0 + b_1x_1 + b_2x_2 + \dots + b_nx_n$$

- Weights are calculated from the training data
- **Predicted** value for first training instance  $\mathbf{x}^{(1)}$  is:

$$b_0x_0^{(1)} + b_1x_1^{(1)} + b_2x_2^{(1)} + \dots + b_nx_n^{(1)} = \sum_{i=0}^n b_ix_i^{(1)}$$

# Minimizing Squared Error

Difference between *predicted* and *actual* values is the error !

$n + 1$  coefficients are chosen so that sum of squared error on all instances in training data is minimized

Squared error:

$$\sum_{j=1}^m \left( y^{(j)} - \sum_{i=0}^n b_i x_i^{(j)} \right)^2$$

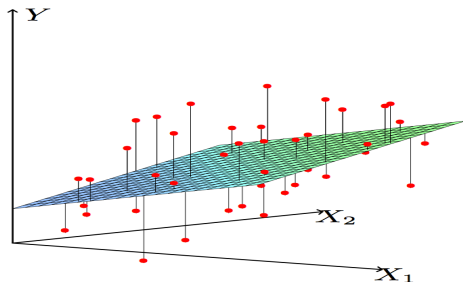
Coefficients can be derived using standard matrix operations

Can be done if there are more instances than attributes (roughly speaking).

Known as “Ordinary Least Squares” (OLS) regression – minimizing the sum of squared distances of data points to the estimated regression line.

# Multiple Regression

Given 2 real-valued variables  $X_1$ ,  $X_2$ , labelled with a real-valued variable  $Y$ , find “plane of best fit” that captures the dependency of  $Y$  on  $X_1$ ,  $X_2$ .



Learning here is by minimizing MSE, i.e., average of squared vertical distances of actual values of  $Y$  from the learned function  $\hat{Y} = \hat{f}(\mathbf{X})$ .



## Step back: Statistical Techniques for Data Analysis

# Probability vs Statistics: The Difference

- **Probability** versus **Statistics**
- Probability: reasons from populations to samples
  - This is deductive reasoning, and is usually *sound* (in the logical sense of the word)
- Statistics: reasons from samples to populations
  - This is inductive reasoning, and is usually *unsound* (in the logical sense of the word)

# Statistical Analyses

- Statistical analyses usually involve one of 3 things:
  - ① The study of populations;
  - ② The study of variation; and
  - ③ Techniques for data abstraction and data reduction
- Statistical analysis is more than statistical computation:
  - ① What is the question to be answered?
  - ② Can it be quantitative (i.e., can we make measurements about it)?
  - ③ How do we collect data?
  - ④ What can the data tell us?

# Sampling

# Where do the Data come from? (Sampling)

- For groups (populations) that are fairly homogeneous, we do not need to collect a lot of data. (We do not need to sip a cup of tea several times to decide that it is too hot.)
- For populations which have irregularities, we will need to either take measurements of the entire group, or find some way of get a good idea of the population without having to do so
- *Sampling* is a way to draw conclusions about the population without having to measure all of the population. The conclusions need not be completely accurate
- All this is possible if the sample closely resembles the population about which we are trying to draw some conclusions

# What We Want From a Sampling Method

- No systematic bias, or at least no bias that we cannot account for in our calculations
- The chance of obtaining an unrepresentative sample can be calculated. (So, if this chance is high, we can choose not to draw any conclusions.)
- The chance of obtaining an unrepresentative sample decreases with the size of the sample

# Estimation

# Estimation from a Sample

- Estimating some aspect of the population using a sample is a common task. Along with the estimate, we also want to have some idea of the accuracy of the estimate (usually expressed in terms of *confidence limits*)
- Some measures calculated from the sample are very good estimates of corresponding population values. For example, the sample mean  $m$  is a very good estimate of the population mean  $\mu$ . But this is not always the case. For example, the range of a sample usually under-estimates the range of the population
- We will have to clarify what is meant by a “good estimate”. One meaning is that an estimator is correct on average. For example, on average, the mean of a sample is a good estimator of the mean of the population



## Estimation from a Sample

- For example, when a number of samples are drawn and the mean of each is found, then average of these means is equal to the population mean
- Such an estimator is said to be *statistically unbiased*

# Sample Estimates of the Mean and the Spread I

**Mean.** This is calculated as follows.

- Find the total  $T$  of  $N$  observations. Estimate the (arithmetic) mean by  $m = T/N$ .
- This works very well when the data follow a symmetric bell-shaped frequency distribution (of the kind modelled by “normal” distribution)
- A simple mathematical expression of this is  $m = \frac{1}{N} \sum_i x_i$ , where the observations are  $x_1, x_2 \dots x_n$
- If we can group the data so that the observation  $x_1$  occurs  $f_1$  times,  $x_2$  occurs  $f_2$  times and so on, then the mean is calculated even easier as  $m = \frac{1}{N} \sum_i x_i f_i$

# Sample Estimates of the Mean and the Spread II

- If, instead of frequencies, you had relative frequencies (i.e. instead of  $f_i$  you had  $p_i = f_i/N$ ), then the mean is simply the observations weighted by relative frequency. That is,  $m = \sum_i x_i p_i$
- We want to connect this up to computing the mean value of observations modelled by some theoretical probability distribution function. That is, we want to a similar counting method for calculating the mean of random variables modelled using some known distribution

# Sample Estimates of the Mean and the Spread III

- Correctly, this is the mean value of the *values of the random variable function*. But this is a bit cumbersome, so we will just say the “mean value of the r.v.” For discrete r.v.’s this is:

$$E(X) = \sum_i x_i p(X = x_i)$$

**Variance.** This is calculated as follows:

- Calculate the total  $T$  and the sum of squares of  $N$  observations. The estimate of the standard deviation is  $s = \sqrt{\frac{1}{N-1} \sum_i (x_i - m)^2}$
- Again, this is a very good estimate when the data are modelled by a normal distribution

# Sample Estimates of the Mean and the Spread IV

- For grouped data, this is modified to
$$s = \sqrt{\frac{1}{N-1} \sum_i (x_i - m)^2 f_i}$$
- Again, we have a similar formula in terms of expected values, for the scatter (spread) of values of a r.v.  $X$  around a mean value  $E(X)$ :

$$\begin{aligned} \text{Var}(X) &= E((X - E(X))^2) \\ &= E(X^2) - [E(X)]^2 \end{aligned}$$

- You can remember this as “the mean of the squares minus the square of the mean”

## Covariance and Correlation

# Correlation I

- The *correlation coefficient* is a number between -1 and +1 that indicates whether a pair of variables  $x$  and  $y$  are associated or not, and whether the scatter in the association is high or low
  - High values of  $x$  are associated with high values of  $y$  and low values of  $x$  are associated with low values of  $y$ , and scatter is low
  - A value near 0 indicates that there is no particular association and that there is a large scatter associated with the values
  - A value close to -1 suggests an inverse association between  $x$  and  $y$
- Only appropriate when  $x$  and  $y$  are roughly linearly associated (doesn't work well when the association is curved)

# Correlation II

- The formula for computing correlation between  $x$  and  $y$  is:

$$r = \frac{\text{cov}(x, y)}{\sqrt{\text{var}(x)}\sqrt{\text{var}(y)}}$$

This is sometimes also called *Pearson's correlation coefficient*

- The terms in the denominator are simply the standard deviations of  $x$  and  $y$ . But the numerator is different. This is the *covariance*, calculated as the average of the product of deviations from the mean:

$$\text{cov}(x, y) = \frac{\sum_i (x_i - \bar{x})(y_i - \bar{y})}{n - 1}$$

- What does “covariance” actually mean ? Consider
  - Case 1:  $x_i > \bar{x}$ ,  $y_i > \bar{y}$
  - Case 2:  $x_i < \bar{x}$ ,  $y_i < \bar{y}$



# Correlation III

3 Case 3:  $x_i < \bar{x}$ ,  $y_i > \bar{y}$

4 Case 4:  $x_i > \bar{x}$ ,  $y_i < \bar{y}$

In the first two cases,  $x_i$  and  $y_i$  vary together, both being high or low relative to their means. In the other two cases, they vary in different directions

- If the positive products dominate in the calculation of  $\text{cov}(x, y)$ , then the value of  $r$  will be positive. If the negative products dominate, then  $r$  will be negative. If 0 products dominate, then  $r$  will be close to 0.
- You should be able to show that:

$$r = \frac{\sum_i (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum (x_i - \bar{x})^2} \sqrt{\sum (y_i - \bar{y})^2}}$$

# Correlation IV

- Computers generally use a short-cut formula:

$$r = \frac{\sum_i x_i y_i - n \overline{xy}}{n - 1}$$

- The same kinds of calculations can be done if the data were not actual values but ranks instead (i.e. ranks for the  $x$ 's and the  $y$ 's).
  - This is called *Spearman's rank correlation*, but we won't do these calculations here.

# What Does Correlation Mean? I

- $r$  is a quick way of checking whether there is some linear association between  $x$  and  $y$
- The sign of the value tells you the direction of the association
- All that the numerical value tells you is about the scatter in the data
- The correlation coefficient does not model any relationship. That is, given a particular  $x$  you cannot use the  $r$  value to calculate a  $y$  value
  - It is possible for two datasets to have the same correlation, but different relationships
  - It is possible for two datasets to have different correlations but the same relationship
- MORAL: Do not use correlations to compare datasets. All you can derive is whether there is a positive or negative relationship between  $x$  and  $y$
- ANOTHER MORAL: Do not use correlation to imply  $x$  causes  $y$  or the other way around

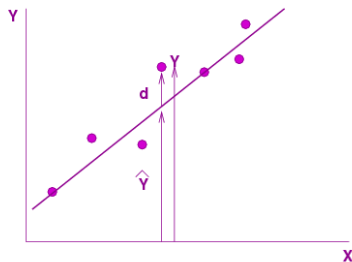
# Regression

# Regression

- Given a set of data points  $x_i, y_i$ , what is the relationship between them? (We can generalise this to the “multivariate” case later)
- One kind of question is to ask: are these linearly related in some manner? That is, can we draw a straight line that describes reasonably well the relationship between  $X$  and  $Y$
- Remember, the correlation coefficient can tell us if there is a case for such a relationship
- In real life, even if such a relationship held, it will be unreasonable to expect all pairs  $x_i, y_i$  to lie precisely on a straight line. Instead, we can probably draw some reasonably well-fitting line. But which one?

# Univariate linear regression

# Linear Relationship Between 2 Variables I



- GOAL: fit a line whose equation is of the form  $\hat{Y} = a + bX$
- HOW: minimise  $\sum_i d_i^2 = \sum_i (Y_i - \hat{Y}_i)^2$  (the “least squares estimator”)

# Linear Relationship Between 2 Variables II

- The calculation for  $b$  is given by:

$$b = \frac{\text{cov}(x, y)}{\text{var}(x)}$$

where  $\text{cov}(x, y)$  is the covariance of  $x$  and  $y$ , given by  $\sum_i (x_i - \bar{x})(y_i - \bar{y})$  as before

- This can be simplified to:

$$b = \sum (xy) / \sum x^2$$

where  $x = (X_i - \bar{X})$  and  $y = (Y_i - \bar{Y})$

- $a = \bar{Y} - b\bar{X}$

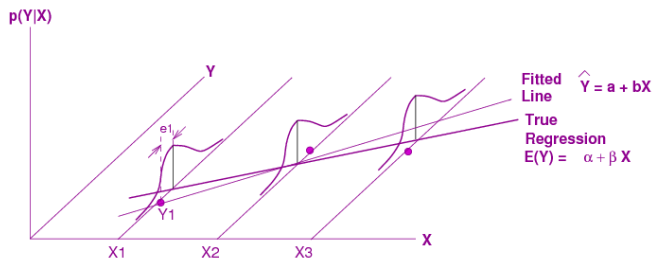


# Meaning of the Coefficients $a$ and $b$

- $b$ : change in  $Y$  that accompanies a unit change in  $X$
- If the values of  $X$  were assigned at random, then  $b$  estimates the unit change in  $Y$  *caused* by a unit change in  $X$
- If the values of  $X$  were not assigned at random (for example, they were data somebody observed), then the change in  $Y$  will include the change in  $X$  and any other confounding variables that may have changed as a result of changing  $X$  by 1 unit. So, you cannot say for example, that a change of  $X$  by 1 unit causes  $b$  units of change in  $Y$
- $b = 0$  means there is no linear relationship between  $X$  and  $Y$ , and then best we can do is simply say is  $\hat{Y} = a = \bar{Y}$ . Estimating the sample mean is therefore a special case of the MSE criterion

# The Regression Model I

- The least-square estimator fits a line using sample data
- To draw inferences about the population requires us to have a (statistical) model about what this line means
- What is being assumed is actually this:



# The Regression Model II

- That is: Obtain  $Y$  values for many instances of  $X_1$ . This will result in a distribution of  $Y$  values  $P(Y|X_1)$ ; and so on for  $P(Y|X_2), P(Y|X_3), \text{etc.}$ . The regression model makes the following assumptions:
  - All the  $Y$  distributions are the same, and have the same spread
  - For each  $P(Y|X_i)$  distribution, the true mean value  $\mu_i$  lies on a straight line (this is the “true regression line”)
  - The  $Y_i$  are independent
- In standard terminology, the  $Y_i$  are identically distributed independent (i.i.d.) random variables with mean  $\mu_i = \alpha + \beta X_i$  and variance  $\sigma^2$
- Or:  $Y_i = \alpha + \beta X_i + e_i$  where the  $e_i$  are independent errors with mean 0 and variance  $\sigma^2$

## Finding the parameters for univariate linear regression

# Univariate linear regression

## Example:

Suppose we want to investigate the relationship between people's height and weight. We collect  $n$  height and weight measurements  $(h_i, w_i), 1 \leq i \leq n$ .

Univariate linear regression assumes a linear equation  $w = a + bh$ , with parameters  $a$  and  $b$  chosen such that the sum of squared residuals  $\sum_{i=1}^n (w_i - (a + bh_i))^2$  is minimised.

## Univariate linear regression

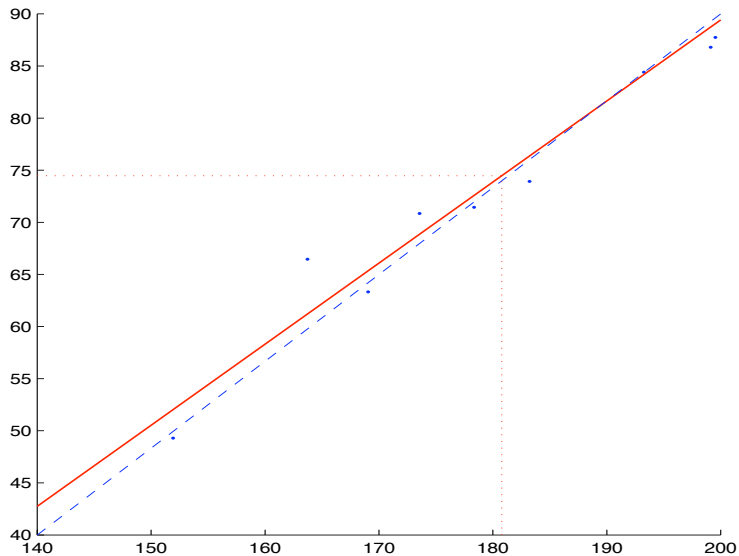
In order to find the parameters we take partial derivatives, set the partial derivatives to 0 and solve for  $a$  and  $b$ :

$$\begin{aligned}\frac{\partial}{\partial a} \sum_{i=1}^n (w_i - (a + bh_i))^2 &= -2 \sum_{i=1}^n (w_i - (a + bh_i)) = 0 \\ \Rightarrow \hat{a} &= \bar{w} - \hat{b}\bar{h}\end{aligned}$$

$$\begin{aligned}\frac{\partial}{\partial b} \sum_{i=1}^n (w_i - (a + bh_i))^2 &= -2 \sum_{i=1}^n (w_i - (a + bh_i))h_i = 0 \\ \Rightarrow \hat{b} &= \frac{\sum_{i=1}^n (h_i - \bar{h})(w_i - \bar{w})}{\sum_{i=1}^n (h_i - \bar{h})^2}\end{aligned}$$

So the solution found by linear regression is  $w = \hat{a} + \hat{b}h = \bar{w} + \hat{b}(h - \bar{h})$ .

# Univariate linear regression



## Univariate linear regression

### Shown on previous slide:

The red solid line indicates the result of applying linear regression to 10 measurements of body weight (on the  $y$ -axis, in kilograms) against body height (on the  $x$ -axis, in centimetres). The orange dotted lines indicate the average height  $\bar{h} = 181$  and the average weight  $\bar{w} = 74.5$ ; the regression coefficient  $\hat{b} = 0.78$ . The measurements were simulated by adding normally distributed noise with mean 0 and variance 5 to the true model indicated by the blue dashed line ( $b = 0.83$ ).



# Linear regression: intuitions

For a feature  $X$  and a target variable  $Y$ , the regression coefficient is the covariance between  $X$  and  $Y$  in proportion to the variance of  $X$ :

$$\hat{b} = \frac{\sigma_{XY}}{\sigma_{XX}}$$

(Here we use  $\sigma_{XX}$  as an alternative notation for  $\sigma_X^2$ ).

This can be understood by noting that the covariance is measured in units of  $X$  times units of  $y$  (e.g., metres times kilograms above) and the variance in units of  $X$  squared (e.g., metres squared), so their quotient is measured in units of  $y$  per unit of  $X$  (e.g., kilograms per metre).

## Linear regression: intuitions

The *intercept*  $\hat{a}$  is such that the regression line goes through  $(\bar{X}, \bar{Y})$ .

Adding a constant to all  $X$ -values (a translation) will affect only the intercept but not the regression coefficient (since it is defined in terms of deviations from the mean, which are unaffected by a translation).

So we could *zero-centre* the  $X$ -values by subtracting  $\bar{X}$ , in which case the intercept is equal to  $\bar{Y}$ .

We could even subtract  $\bar{Y}$  from all  $Y$ -values to achieve a zero intercept, without changing the problem in an essential way.

## Linear regression: intuitions

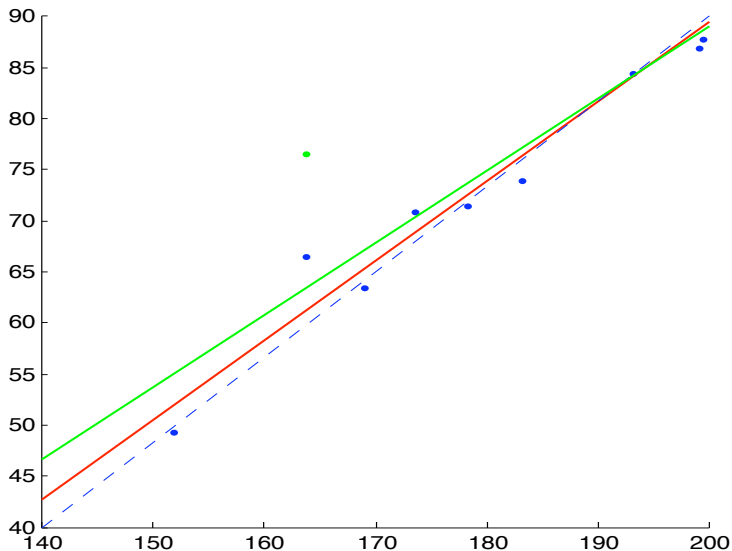
Another important point to note is that the sum of the residuals of the least-squares solution is zero:

$$\sum_{i=1}^n (y_i - (\hat{a} + \hat{b}x_i)) = n(\bar{y} - \hat{a} - \hat{b}\bar{x}) = 0$$

The result follows because  $\hat{a} = \bar{Y} - \hat{b}\bar{X}$ , as derived above.

While this property is intuitively appealing, it is worth keeping in mind that it also makes linear regression susceptible to *outliers*: points that are far removed from the regression line, often because of measurement errors.

# The effect of outliers



## The effect of outliers

### Shown on previous slide:

Suppose that, as the result of a transcription error, one of the weight values from the previous example of univariate regression is increased by 10 kg. The diagram shows that this has a considerable effect on the least-squares regression line.

Specifically, we see that one of the **blue points** got moved up 10 units to the **green point**, changing the **red regression line** to the **green line**.

# Least-Squares as Cost Minimization I

- Finding the least-squares solution is in effect finding the value of  $a$  and  $b$  that minimizes  $\sum_i d_i^2 = \sum_i (Y_i - \hat{Y}_i)^2$ , where  $\hat{Y}_i = a + bX_i$
- This minimum value was obtained analytically by the usual process of differentiating and equating to 0,
- A numerical alternative to the analytical approach is to take (small) steps that decreases the value of the function to be minimised, and stopping when we reach a minimum
- Recall that at a point the gradient vector points in the direction of greatest increase of a function. So, the opposite direction to the gradient vector gives the direction of greatest decrease
  - $b_{i+1} = b_i - \eta \times g_b$
  - $a_{i+1} = a_i - \eta \times g_a$
  - Stop when  $b_{i+1} \approx b_i$  and  $a_{i+1} \approx a_i$
- More on this in a later lecture

## Multivariate linear regression

# Many variables

- Often, we are interesting in modelling the relationship of  $Y$  to several other variables
- In observational studies, the value of  $Y$  may be affected by the values of several variables. For example, carcinogenicity may be gender-specific. A regression model that ignores gender may find that carcinogenicity to be related to some surrogate variable (height, for example)
- Including more variables can give a narrower confidence interval on the prediction being made



# Multivariate linear model

- The  $Y_i$  are identically distributed independent variables with mean  $\mu = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \cdots + \beta_n X_n$  and variance  $\sigma^2$
- Or:  $Y_i = \beta_0 + \beta_1 X_1 + \cdots + \beta_n X_n + e_i$  where the  $e_i$  are independent errors with mean 0 and variance  $\sigma^2$
- As before, this linear model is estimated from a sample by the equation  $\hat{Y} = b_0 + b_1 X_1 + b_2 X_2 + \cdots + b_n X_n$
- With many variables, the regression equation and expressions for the  $b_i$  are expressed better using a matrix representation for sets of equations.

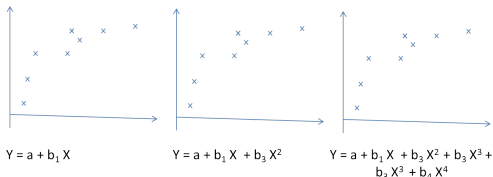
# Regularisation

# Parameter Estimation by Optimization I

*Regularisation* is a general method to avoid overfitting by applying additional constraints to the weight vector. A common approach is to make sure the weights are, on average, small in magnitude: this is referred to as *shrinkage*.

Recall the setting for regression in terms of cost minimization.

- Can add penalty terms to a *cost* function, forcing coefficients to shrink to zero



$$Y = f_{\theta_0, \theta_1, \dots, \theta_n}(X_1, X_2, \dots, X_n) = f_{\theta}(\mathbf{X})$$

# Parameter Estimation by Optimization II

- MSE as a cost function, given data  $(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)$

$$Cost(\theta) = \frac{1}{n} \sum_i (f_{\theta}(\mathbf{x}_i) - y_i)^2$$

and with a penalty function:

$$Cost(\theta) = \frac{1}{n} \sum_i (f_{\theta}(\mathbf{x}_i) - y_i)^2 + \frac{1}{n} \lambda \sum_{i=1}^n \theta_i$$

- Parameter estimation by optimisation will attempt to values for  $\theta_0, \theta_1, \dots, \theta_n$  s.t.  $Cost(\theta)$  is a minimum
- It will be easier to take the  $\frac{1}{n}$  term as  $\frac{1}{2n}$ , which will not affect the minimisation

# Parameter Estimation by Optimization III

- Using gradient descent with the penalty function will do two things:
  - (a) we will move each  $\theta_j$  in a direction that minimises the cost; and
  - (b) each value of  $\theta_j$  will also get “shrunk” on each iteration by multiplying the old value by an amount  $< 1$

$$\theta_j^{(i+1)} = \alpha \theta_j^{(i)} - \eta \nabla_{\theta_j}$$

where  $\alpha < 1$

# Regularised regression

The multivariate least-squares regression problem can be written as an optimisation problem:

$$\mathbf{w}^* = \arg \min_{\mathbf{w}} (\mathbf{y} - \mathbf{X}\mathbf{w})^T (\mathbf{y} - \mathbf{X}\mathbf{w})$$

The regularised version of this optimisation is then as follows:

$$\mathbf{w}^* = \arg \min_{\mathbf{w}} (\mathbf{y} - \mathbf{X}\mathbf{w})^T (\mathbf{y} - \mathbf{X}\mathbf{w}) + \lambda \|\mathbf{w}\|^2$$

where  $\|\mathbf{w}\|^2 = \sum_i w_i^2$  is the squared norm of the vector  $\mathbf{w}$ , or, equivalently, the dot product  $\mathbf{w}^T \mathbf{w}$ ;  $\lambda$  is a scalar determining the amount of regularisation.

## Regularised regression

This regularised problem still has a closed-form solution:

$$\hat{\mathbf{w}} = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y}$$

where  $\mathbf{I}$  denotes the identity matrix. Regularisation amounts to adding  $\lambda$  to the diagonal of  $\mathbf{X}^T \mathbf{X}$ , a well-known trick to improve the numerical stability of matrix inversion. This form of least-squares regression is known as *ridge regression*.

An interesting alternative form of regularised regression is provided by the *lasso*, which stands for 'least absolute shrinkage and selection operator'. It replaces the ridge regularisation term  $\sum_i w_i^2$  with the sum of absolute weights  $\sum_i |w_i|$ . The result is that some weights are shrunk, but others are set to 0, and so the lasso regression favours *sparse solutions*.

## Bias-Variance Decomposition



# The Bias-Variance Tradeoff

- When comparing unbiased estimators, we would like to select the one with minimum variance
- In general, we would be comparing estimators that have some bias and some variance
- We can combine the bias and variance of an estimator by obtaining the *mean square error* of the estimator, or MSE. This is the average value of squared deviations of an estimated value  $V$  from the true value of the parameter  $\theta$ . That is:

$$\text{MSE} = \text{Avg. value of } (V - \theta)^2$$

- Now, it can be shown that:

$$\text{MSE} = (\text{variance}) + (\text{bias})^2$$

- If, as sample size increases, the bias and the variance of an estimator approaches 0, then the estimator is said to be *consistent*.

## The Bias-Variance Tradeoff

- Since

$$\text{MSE} = (\text{variance}) + (\text{bias})^2$$

the lowest possible value of MSE is 0

- In general, we may not be able to get to the ideal MSE of 0. Sampling theory tells us the minimum value of the variance of an estimator. This value is known as the *Cramer-Rao* bound. So, given an estimator with bias  $b$ , we can calculate the minimum value of the variance of the estimator using the CR bound (say,  $v_{min}$ ). Then:

$$\text{MSE} \geq v_{min} + b^2$$

The value of  $v_{min}$  depends on whether the estimator is biased or unbiased (that is  $b = 0$  or  $b \neq 0$ )

- It is not the case that  $v_{min}$  for an unbiased ( $b = 0$ ) estimator is less than  $v_{min}$  for a biased estimator. So, the MSE of a biased estimator can end up being lower than the MSE of an unbiased estimator.

# Decomposition of MSE

$$\text{MSE} = (\text{variance}) + (\text{bias})^2$$

Imagine testing the prediction of our estimator  $\hat{y}$  on many samples of the same size drawn at random from the same distribution. We compute error based on the squared difference between predicted and actual values. Then the MSE can be decomposed like this:

$$\begin{aligned}\text{MSE} &= E[\hat{y} - f(x)]^2 \\ &= E[\hat{y} - E(\hat{y})]^2 + [E(\hat{y}) - f(x)]^2\end{aligned}$$

Note that the first term in the error decomposition (variance) does not refer to the actual value at all, although the second term (bias) does.

## Some theoretical properties

# How Good is the Least-Squares Estimator I

- The line fitted using the least-squares criterion is a sample-based estimate of the true regression line
- To know how good this estimate is, we are really asking questions about the bias and variance of the estimates of  $a$  and  $b$
- It can be shown that under some assumptions, the least-square estimates of  $a$  and  $b$  will be unbiased and that they will have the lowest variance
- The proof of this is called the *Gauss-Markov theorem*. The Gauss-Markov theorem makes the following assumptions:
  - 1 The expected (average) values of residuals is 0 ( $E(e_i) = 0$ )
  - 2 The spread of residuals is constant for all  $X_i$  ( $Var(e_i) = \sigma^2$ )
  - 3 There is no relationship amongst the residuals ( $cov(e_i, e_j) = 0$ )
  - 4 There is no relationship between the residuals and the  $X_i$  ( $cov(X_i, e_i) = 0$ )

# How Good is the Least-Squares Estimator II

- If these assumptions hold, then the Gauss-Markov theorem shows that  $E(a) = \alpha$ ,  $E(b) = \beta$ , and that the variance in these estimates will have the lowest variance
- There is a special case of the assumptions that arises when the residuals are assumed to be distributed according to the Normal distribution, with mean 0
  - In this case, minimising least-squares is equivalent to maximising the probability of the  $Y_i$ , given the  $X_i$  (that is, least-squares is equivalent to *maximum likelihood estimation*)
  - More on this in a later lecture

## Some further issues in learning linear regression models

# What do the Coefficients $b_i$ Mean?

- Consider the two equations:

$$\hat{Y} = a + bX$$

$$\hat{Y} = b_0 + b_1X_1 + b_2X_2$$

- $b$ : change in  $Y$  that accompanies a unit change in  $X$
- $b_1$ : change in  $Y$  that accompanies a unit change in  $X_1$  *provided*  $X_2$  *remains constant*
- More generally,  $b_i$  ( $i > 0$ ) is the change in  $Y$  that accompanies a unit change in  $X_i$  provided all other  $X$ 's are constant
- So: if all relevant variables are included, then we can assess the effect of each one in a controlled manner



# Categoric Variables: $X$ 's I

- “Indicator” variables are those that take on the values 0 or 1
- They are used to include the effects of categoric variables. For example, if  $D$  is a variable that takes the value 1 if a patient takes a drug and 0 if the patient does not. Suppose you want to know the effect of drug  $D$  on blood pressure  $Y$  keeping age ( $X$ ) constant

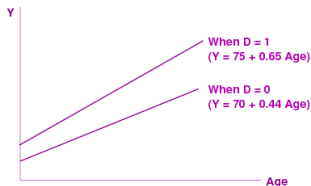
$$\hat{Y} = 70 + 5D + 0.44X$$

So, taking the drug (a unit change in  $D$ ) makes a difference of 5 units, provided age is held constant

## Categoric Variables: $X$ 's II

- How do we capture any interaction effect between age and drug intake? Introduce a new indicator variable  $DX = D \times X$

$$\hat{Y} = 70 + 5D + 0.44X + 0.21DX$$



## Categoric Values: $Y$ values

- Sometimes,  $Y$  values are simply one of two values (let's call them 0 and 1)
- We can't use the regression model as we described earlier, in which the  $Y$ 's can take any real value
- But, we can define a new linear regression model in which predicts not the value of  $Y$ , but what are called the *log odds* of  $Y$ :

$$\log \text{ odds } Y = Odds = b_0 + b_1X_1 + \cdots + b_nX_n$$

- Once *Odds* are estimated, they can be used to calculate the probability of  $Y$ :

$$Pr(Y = 1) = \frac{e^{Odds}}{(1 + e^{Odds})}$$

We can then use the value of  $Pr(Y = 1)$  to decide if  $Y = 1$

- This procedure is called *logistic regression* (we'll see this again)

# Is the Model Appropriate ?

- If there is no systematic pattern to the residuals—that is, there are approximately half of them that are positive and half that are negative, then the line is a good fit
- It should also be the case that there should be no pattern to the residual scatter all along the line. If the average size of the residuals varies along the line (this condition is called *heteroscedasticity*) then the relationship is probably more complex than a straight line
- Residuals from a well-fitting line should show an approximate symmetric, bell-shaped frequency distribution with a mean of 0

# Non-linear Relationships

**A question:** is it possible to do better than the line of best fit?

Maybe. Linear regression assumes that the  $(\mathbf{x}_i, y_i)$  examples in the data are “generated” by the true (but unknown) function  $Y = f(\mathbf{X})$ .

So any training set is a sample from the true distribution  $E(Y) = f(\mathbf{X})$ .

But what if  $f$  is non-linear ?

We may be able to reduce the mean squared error (MSE) value  $\sum_i (y_i - \hat{y})^2$  by trying a different function.

## Non-linear Relationships

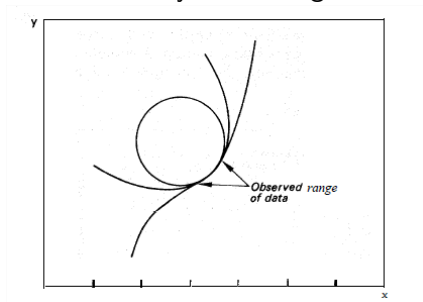
- Some non-linear relationships can be captured in a linear model by a transformation (“trick”). For example, the curved model  $\hat{Y} = b_0 + b_1 X_1 + b_2 X_1^2$  can be transformed by  $X_2 = X_1^2$  into a linear model. This works for polynomial relationships.
- Some other non-linear relationships may require more complicated transformations. For example, the relationship is  $Y = b_0 X_1^{b_1} X_2^{b_2}$  can be transformed into the linear relationship

$$\log(Y) = \log b_0 + b_1 \log X_1 + b_2 \log X_2$$

- Other relationships cannot be transformed quite so easily, and will require full non-linear estimation (in subsequent topics in the ML course we will find out more about these)

## Non-Linear Relationships

- Main difficulty with non-linear relationships is choice of function
  - How to learn ?
  - Can use a form of gradient descent to estimate the parameters
- After a point, almost any sufficiently complex mathematical function will do the job in a sufficiently small range



- Some kind of prior knowledge or theory is the only way to help here.
  - Otherwise, it becomes a process of trial-and-error, in which case, beware of conclusions that can be drawn

# Model Selection

- Suppose there are a lot of variables  $X_i$ , some of which may be representing products, powers, *etc.*
- Taking all the  $X_i$  will lead to an overly complex model. There are 3 ways to reduce complexity:
  - ① Subset-selection, by search over subset lattice. Each subset results in a new model, and the problem is one of model-selection
  - ② Shrinkage, or *regularization* of coefficients to zero, by optimization. There is a single model, and unimportant variables have near-zero coefficients.
  - ③ Dimensionality-reduction, by projecting points into a lower dimensional space (this is different to subset-selection, and we will look at it later)



# Model Selection as Search I

- The subsets of the set of possible variables form a lattice with  $S_1 \cap S_2$  as the g.l.b. or meet and  $S_1 \cup S_2$  as the l.u.b. or join
- Each subset refers to a model, and a pair of subsets are connected if they differ by just 1 element
- A lattice is a graph, and we know how to search a graph
  - $A^*$ , greedy, randomised *etc.*
  - “Cost” of node in the graph: MSE of the model. The parameters (coefficients) of the model can be found
- Historically, model-selection for regression has been done using “forward-selection”, “backward-elimination”, or “stepwise” methods
  - These are greedy search techniques that either: (a) start at the top of the subset lattice, and add variables; (b) start at the bottom of the subset lattice and remove variables; or (c) start at some interior point and proceed by adding or removing single variables (examining nodes connected to the node above or below)

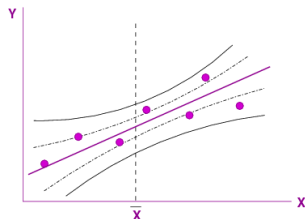
# Model Selection as Search II

- Greedy selection done on the basis of calculating the *coefficient of determination* (often denoted by  $R^2$ ) which denotes the proportion of total variation in the dependent variable  $Y$  that is explained by the model
- Given a model formed with a subset of variables  $X$ , it is possible to compute the observed change in  $R^2$  due to the addition or deletion of some variable  $x$
- This is used to select greedily the next best move in the graph-search

To set other *hyper-parameters*, such as shrinkage parameter  $\lambda$ , can use grid search

# Prediction I

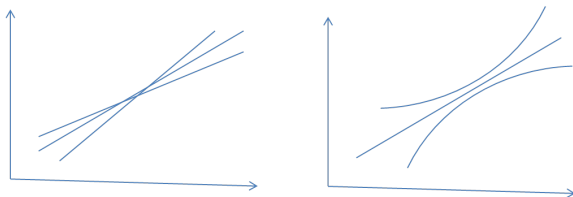
- It is possible to quantify what happens if the regression line is used for prediction:



- The intuition is this:
  - Recall the regression line goes through the mean  $(\bar{X}, \bar{Y})$

# Prediction II

- If the  $X_i$  are slightly different, then the mean is not going to change much. So, the regression line stays somewhat “fixed” at  $(\bar{X}, \bar{Y})$  but with a different slope
- With each different sample of the  $X_i$  we will get a slightly different regression line
- The variation in  $Y$  values is greater further we move from  $(\bar{X}, \bar{Y})$



- MORAL: Be careful, when predicting far away from the centre value
- ANOTHER MORAL: The model only works under the approximately the same conditions that held when collecting the data

## Local (nearest-neighbour) regression

# Local learning

- Related to the simplest form of learning: rote learning, or memorization
- Training instances are searched for instance that **most closely resembles** *query* or test instance
- The *instances* themselves represent the knowledge
- Called: *nearest-neighbour*, *instance-based*, *memory-based* or *case-based* learning; all forms of *local learning*
- The *similarity* or *distance* function defines “learning”, i.e., how to go beyond simple memorization
- Intuition — classify an instance similarly to examples “close by” — neighbours or *exemplars*
- A form of *lazy* learning – don’t need to build a model!

# Nearest neighbour for numeric prediction

Store all training examples  $\langle x_i, f(x_i) \rangle$ .

Nearest neighbour:

- Given query instance  $x_q$ ,
- first locate nearest training example  $x_n$ ,
- then estimate  $\hat{y} = \hat{f}(x_q) = f(x_n)$
- $k$ -Nearest neighbour:
- Given  $x_q$ , take mean of  $f$  values of  $k$  nearest neighbours

$$\hat{y} = \hat{f}(x_q) = \frac{\sum_{i=1}^k f(x_i)}{k}$$

# Distance function

The distance function defines what is “learned”, i.e., predicted.  
Instance  $x_i$  is described by an  $m$ -vector of feature values:

$$\langle x_{i1}, x_{i2}, \dots, x_{im} \rangle$$

where  $x_{ik}$  denotes the value of the  $k$ th feature of  $x_i$ .

Most commonly used distance function is *Euclidean* distance, where the distance between two instances  $x_i$  and  $x_j$  is defined to be:

$$d(x_i, x_j) = \sqrt{\sum_{k=1}^m (x_{ik} - x_{jk})^2}$$



# Local regression

Use  $k$ NN to form a local approximation to  $f$  for each query point  $x_q$  using a linear function of the form

$$\hat{f}(x) = b_0 + b_1x_1 + \dots + b_mx_m$$

where  $x_i$  denotes the value of the  $i$ th feature of instance  $x$ .

Where does this linear regression model come from ?

- fit linear function to  $k$  nearest neighbours
- or quadratic or higher-order polynomial ...
- produces “piecewise approximation” to  $f$

# Summary

# Summary

- Regression gives us a glimpse into many aspects of Machine Learning
    - Terminology.** Training data, test data, resubstitution error, prediction error (later lecture).
    - Conceptual.** Learning as search, learning as optimisation, assumptions underlying a technique
    - Implementation.** Approximate alternatives to analytical methods
    - Application.** Overfitting, problems of prediction
- Each of these aspects will have counterparts in other kinds of machine learning
- Linear models are one way to predict numerical quantities
    - Ordinal regression: predicting ranks (not in the lectures)
    - Neural networks: non-linear regression models (later)
    - Regression trees: piecewise regression models (later)
    - Class-probability trees: predicting probabilities (later)
    - Model trees: piecewise non-linear models (later)