

Homework 1

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1. Consider the queue model we discussed in class (i.e. a single server model). As we did in class, assume that the interarrival times, T_i are iid as are the service times S_i . Assume further that each T_i is exponentially distributed with rate λ and each S_i is exponentially distributed with rate μ . Let $Q(t)$ be the number of customers waiting in line at time t . Let W_i be the waiting time of the i th individual. Assume that initially the queue is empty, so $Q(0) = 0$ and $W_1 = 0$.

(a) Determine $P(W_2 \geq c)$ for c a positive number. (write down an integral and evaluate it, you won't need a computer). Note this is a little different than what we did in class when we considered $P(W_2 = 0)$.

$$w_2 = \begin{cases} 0, & \text{if } A_2 > D_1 \\ D_1 - A_2, & \text{otherwise} \end{cases} \quad \text{where } A_2 = T_1 + T_2 \text{ and } D_1 = A_1 + S_1 = T_1 + S_1$$

$$\begin{aligned} P(w_2 \geq c) &= P(0 \geq c \mid A_2 - D_1 > 0) + P(D_1 - A_2 \geq c \mid A_2 - D_1 \leq 0) \\ &= 0 + P(D_1 - A_2 \geq c \mid A_2 - D_1 \leq 0) \\ &= P(D_1 - A_2 \geq c \mid A_2 - D_1 \leq 0) \\ &= P(D_1 - A_2 \geq c \mid D_1 - A_2 \geq 0) \\ &= P(T_1 + S_1 - T_1 - T_2 \geq c \mid T_1 + S_1 - T_1 - T_2 \geq 0) \\ &= P(S_1 - T_2 \geq c \mid S_1 - T_2 \geq 0) \end{aligned}$$

Since $T_i \sim \text{Exp}(\lambda)$, $S_i \sim \text{Exp}(\mu)$, and T_i, S_i are independent, we have

$$\begin{aligned} f_T(t) &= \lambda e^{-\lambda t} \text{ for } t \geq 0 \\ f_S(s) &= \mu e^{-\mu s} \text{ for } s \geq 0 \\ \implies f_{T,S}(t,s) &= \lambda e^{-\lambda t} \cdot \mu e^{-\mu s} \text{ for } t, s \geq 0 \end{aligned}$$

Let $Z_2 = S_1 - T_2$, we'll find the CDF $F_{Z_2}(z_2)$. Since $w_2 \geq c$ ($c > 0$) occurs only when the next event is arrival, not departure, we may suppose $z_2 \geq 0$, that is $S_1 \geq T_2$. Take the integral, we have:

$$\begin{aligned}
P(Z_2 \leq z_2) &= P(S - T \leq z_2) = \int_0^\infty \int_0^{t+z_2} \lambda e^{-\lambda t} \cdot \mu e^{-\mu s} ds dt \\
&= \int_0^\infty (-e^{-\mu s}|_0^{t+z_2}) \lambda e^{-\lambda t} dt \\
&= \int_0^\infty (1 - e^{-\mu(t+z_2)}) \lambda e^{-\lambda t} dt \\
&= \int_0^\infty \lambda e^{-\lambda t} - \lambda e^{-\lambda t} \cdot e^{-\mu(t+z_2)} dt \\
&= \int_0^\infty \lambda e^{-\lambda t} - \lambda e^{-\mu z_2} \int_0^\infty e^{-(\lambda+\mu)t} dt \\
&= (-e^{-\lambda t})|_0^\infty + \frac{\lambda}{\lambda+\mu} \cdot e^{-\mu z_2} \cdot (e^{-(\lambda+\mu)t}|_0^\infty) \\
&= 1 + \frac{\lambda}{\lambda+\mu} \cdot e^{-\mu z_2} \cdot (-1) \\
&= 1 - \frac{\lambda}{\lambda+\mu} e^{-\mu z_2}
\end{aligned}$$

Thus, $F_{Z_2}(z_2) = 1 - \frac{\lambda}{\lambda+\mu} e^{-\mu z_2}$, for $z_2 \geq 0$.

Take the differentiate of the CDF $F_{Z_2}(z_2)$, we have the density function $f_{Z_2}(z_2) = \frac{\lambda\mu}{\lambda+\mu} e^{-\mu z_2}$, which would be used in b).

Given $c > 0$, we can use the CDF above to calculate $P(Z_2 \geq c)$, which is equivalent to $P(w_2 \geq c)$ as below:

$$\begin{aligned}
P(w_2 \geq c) &= P(S_1 - T_2 \geq c \mid S_1 - T_2 \geq 0) \\
&= P(Z_2 \geq c \mid Z_2 \geq 0) \\
&= 1 - P(Z_2 \leq c \mid Z_2 \geq 0) \\
&= 1 - (1 - \frac{\lambda}{\lambda+\mu} e^{-\mu c}) \\
&= \frac{\lambda}{\lambda+\mu} e^{-\mu c}
\end{aligned}$$

(b) Write down an integral expression for $P(W_3 \geq c)$ (You don't need to evaluate the integral, unless you want to. Your answer may be the sum of two integrals.)

$$w_3 = \begin{cases} 0, & \text{if } A_3 > D_2 \\ D_2 - A_3, & \text{otherwise} \end{cases} \text{ where } A_3 = T_1 + T_2 + T_3 \text{ and } D_2 = A_2 + w_2 + S_2 = T_1 + T_2 + w_2 + S_2$$

$$\begin{aligned}
P(w_3 \geq c) &= P(0 \geq c \mid A_3 - D_2 > 0) + P(D_2 - A_3 \geq c \mid A_3 - D_2 \leq 0) \\
&= 0 + P(D_2 - A_3 \geq c \mid A_3 - D_2 \leq 0) \\
&= P(D_2 - A_3 \geq c \mid A_3 - D_2 \leq 0) \\
&= P(D_2 - A_3 \geq c \mid D_2 - A_3 \geq 0) \\
&= P(T_1 + T_2 + w_2 + S_2 - T_1 - T_2 - T_3 \geq c \mid T_1 + T_2 + w_2 + S_2 - T_1 - T_2 - T_3 \geq 0) \\
&= P(w_2 + S_2 - T_3 \geq c \mid w_2 + S_2 - T_3 \geq 0)
\end{aligned}$$

We know $S_2 \sim \text{Exp}(\mu)$, $T_3 \sim \text{Exp}(\lambda)$. Follow the same logic in a), let $Z_3 = S_2 - T_3$, easily we have $F_{Z_3}(z_3) = 1 - \frac{\lambda}{\lambda+\mu} e^{-\mu z_3}$, for $z_3 \geq 0$. Take the differentiate of the CDF $F_{Z_3}(z_3)$, we have the density function $f_{Z_3}(z_3) = \frac{\lambda\mu}{\lambda+\mu} e^{-\mu z_3}$. From a), we also have when $A_2 \leq D_1$, that is when the second person arrives before the first person leaves, w_2 has the same distribution as the random variable Z_2 defined in a). Thus, $F_{W_2}(w_2) = 1 - \frac{\lambda}{\lambda+\mu} e^{-\mu w_2}$.

Furthermore,

$$\begin{aligned}
P(w_3 \geq c) &= P(w_2 + S_2 - T_3 \geq c \mid w_2 + S_2 - T_3 \geq 0) \\
&= P(w_2 + Z_3 \geq c \mid w_2 + Z_3 \geq 0) \\
&= P(0 + Z_3 \geq c \mid 0 + Z_3 \geq 0) \cdot P(w_2 = 0) + P(w_2 + Z_3 \geq c \mid w_2 + Z_3 \geq 0) \cdot P(w_2 > 0)
\end{aligned}$$

Since

$$\begin{aligned}
P(w_2 = 0) &= P(Z_2 = 0) = F_{Z_2}(0) = \frac{\mu}{\lambda + \mu} \\
P(w_2 > 0) &= 1 - P(w_2 = 0) = \frac{\lambda}{\lambda + \mu}
\end{aligned}$$

Thus,

$$\begin{aligned}
P(w_3 \geq c) &= P(0 + Z_3 \geq c \mid 0 + Z_3 \geq 0) \cdot P(w_2 = 0) + P(w_2 + Z_3 \geq c \mid w_2 + Z_3 \geq 0) \cdot P(w_2 > 0) \\
&= \frac{\mu}{\lambda + \mu} \cdot P(Z_3 \geq c \mid Z_3 \geq 0) + \frac{\lambda}{\lambda + \mu} \cdot P(Z_2 + Z_3 \geq c \mid Z_2 + Z_3 \geq 0) \\
&= \frac{\mu}{\lambda + \mu} \cdot (1 - P(Z_3 < c \mid Z_3 \geq 0)) + \frac{\lambda}{\lambda + \mu} \cdot P(Z_2 + Z_3 \geq c \mid Z_2 + Z_3 \geq 0) \\
&= \frac{\mu}{\lambda + \mu} \cdot (1 - F_{Z_3}(c)) + \frac{\lambda}{\lambda + \mu} \cdot P(Z_3 \geq c - Z_2 \mid Z_2 + Z_3 \geq 0) \\
&= \frac{\mu}{\lambda + \mu} \cdot (1 - \int_0^c f_{Z_3}(z_3) dz_3) + \frac{\lambda}{\lambda + \mu} \cdot \int_0^\infty \int_{c-z_2}^\infty f_{z_3}(z_3) f_{z_2}(z_2) dz_3 dz_2 \\
&\text{where } f_{Z_3}(z_3) = \frac{\lambda\mu}{\lambda + \mu} e^{-\mu z_3} \text{ for } z_3 \geq 0 \\
&\text{and } f_{Z_2}(z_2) = \frac{\lambda\mu}{\lambda + \mu} e^{-\mu z_2} \text{ for } z_2 \geq 0
\end{aligned}$$

(c) Write a function `WaitingTimes(n, λ, μ)` that samples the waiting times of the first `n` customers. Your function should return a vector of length `n` with the sampled waiting time. Show the output of your function for `n = 10`, `λ = 1`, `μ = 1`.

```

WaitingTimes <- function(n, lambda, mu){
  # Initialize times array
  #inter_arrival.times= array()
  wait.times = array()
  service.times = array()
  arrival.times = array()
  depart.times = array()
  # First customer
  inter_arrival.first = rexp(1,rate = lambda)
  service.first = rexp(1,rate = mu)
  wait.first = 0
  #inter_arrival.times[1] = inter_arrival.first
  arrival.times[1] = inter_arrival.first
  wait.times[1] = wait.first
  service.times[1] = service.first
  depart.times[1] = arrival.times[1] + wait.times[1] + service.times[1]
  # Prepare for the next customer
  arrival.previous = inter_arrival.first
}

```

```

depart.previous = inter_arrival.first + wait.first + service.first

for (i in 2:n){
  # Generate random time for the next customer
  inter_arrival.next = rexp(1,rate = lambda)
  service.next = rexp(1, rate = mu)
  arrival.next = arrival.previous + inter_arrival.next
  wait.next = ifelse(depart.previous <= arrival.next, 0, depart.previous - arrival.next)
  ## Record current "next" customer into arrays to simulate that he has left
  ## Update arrival and depart time accordingly
  # inter_arrival.times[i] = inter_arrival.next
  arrival.times[i] = arrival.next
  wait.times[i] = wait.next
  service.times[i] = service.next
  depart.times[i] = arrival.next + wait.next + service.next
  arrival.previous = arrival.times[i]
  depart.previous = depart.times[i]
}
return(wait.times)
}

set.seed(5)
WaitingTimes(10,1,1)

```

```

## [1] 0.000000000 0.297899250 0.644094200 1.183404403 0.057271227
## [6] 0.915878392 0.000000000 0.750367477 0.000000000 0.003400472

```

(d) Write a function `plotQ(t, λ, μ)` that simulates (in other words samples) the queue and plots $Q(t)$ up to a time t . Show a single simulation for $t = 20$, $\lambda = 1$, $\mu = 1$.

After the first arrival, the next event could be an arrival or a departure. From b), we have

$$P(w_2 > 0) = \frac{\lambda}{\lambda + \mu}$$

This probability is equivalent to the probability that the next event is arrival. To determine whether that event is an arrival or a departure, we generate a random number p between 0 and 1. If $p > \frac{\lambda}{\lambda + \mu}$, then the event is an arrival, otherwise it's a departure.

```

plotQ <- function(t,lambda, mu){
  # Initialization
  t.elapsed = 0
  Q_history = 0
  # Generate first customer
  T1 = rexp(1,rate=lambda)
  Q = 1
  event_times = T1
  t.elapsed = T1
  num_event = 1

  while (t.elapsed < t) {
    num_event = num_event+1
    if(Q>0) {

```

```

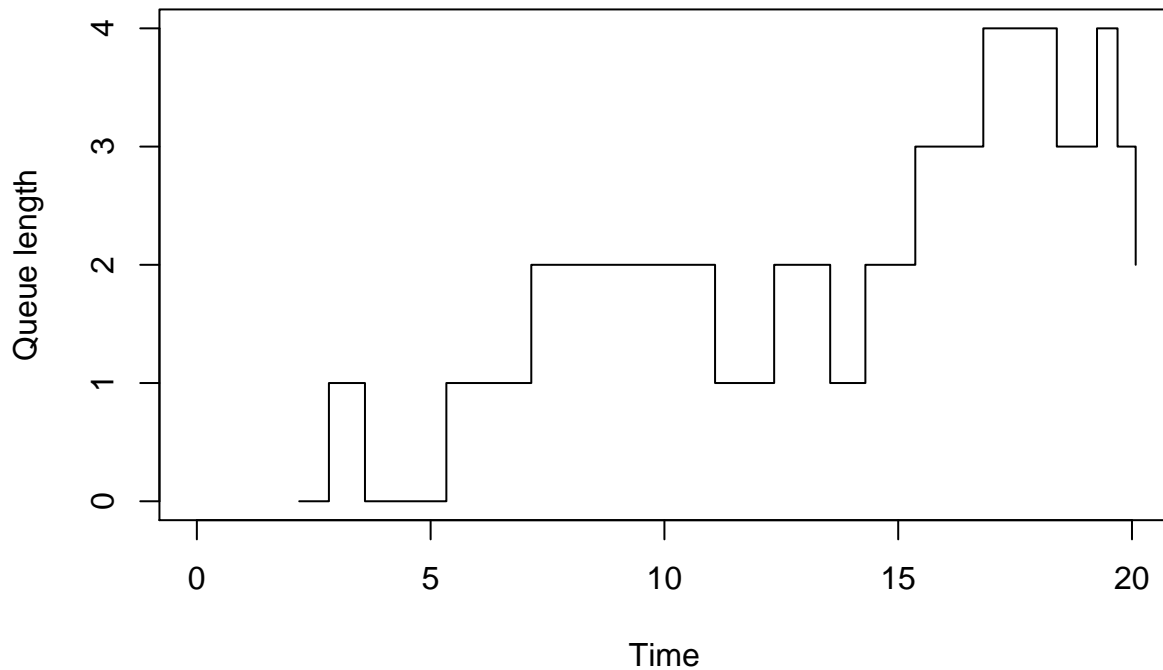
    # check to make sure queue was not empty
    T1 = rexp(1,rate=lambda)
    # roll the dice
    # generate a random number to determine the type of the next event
    p = runif(1,0,1)
    Q_history[num_event] = Q
    # if arrival, Q increase by 1; if departure, Q decreases by 1
    Q = ifelse(p<lambda/(lambda+mu),Q+1,Q-1)
  } else {
    # here, the queue was empty, so only arrivals are possible
    T1 = rexp(1,rate=lambda)
    Q_history[num_event] = Q
    Q = 1
  }
  t.elapsed = t.elapsed+T1
  event_times[num_event] = T1
}

plot(cumsum(event_times),Q_history,type="s",
     xlab="Time",ylab="Queue length",
     xlim = c(0,t),
     main=paste("Queue Simulation (t =",t, ", lambda =",lambda, ", mu =", mu, ")",
               sep = ""))
}

lambda = 1
mu = 1
t = 20
set.seed(12)
plotQ(t,lambda, mu)

```

Queue Simulation (t =20, lambda =1, mu =1)



(e) Using a Monte Carlo approach, estimate $P(W_2 \geq 1)$. Assume $\lambda = 1$, $\mu = 1$. Compare your estimate to the exact answer you derived in part (a). Repeat for $P(W_{100} \geq 1)$, except in this case you won't have the exact answer.

Simulate for $P(W_2 \geq 1)$

```
N <- 20000
lambda <- 1
mu <- 1
w2.simulated.1 = array()
for (i in 1:N){
  set.seed(i)
  w2.simulated.1[i] = WaitingTimes(2,lambda,mu)[2]
}
w2.p1 <- sum(w2.simulated.1 >=1)/N
cat("Simulated P (W2 >= 1) = ", w2.p1)
```

```
## Simulated P (W2 >= 1) = 0.1875
```

Calculated probability in a):

$P(w_2 \geq 1) = \frac{\lambda}{\lambda + \mu} e^{-\mu \cdot 1} = \frac{1}{2 \cdot e} \approx 0.1839$. The simulated probability is close to the theoretical probability.

Repeat for $P(W_{100} \geq 1)$

```
w2.simulated.100 = array()
for (i in 1:N){
  set.seed(i)
  w2.simulated.100[i] = WaitingTimes(100,lambda,mu)[100]
}
w2.p100 <- sum(w2.simulated.100 >=1)/N
cat("Simulated P (W2 >= 100) = ", w2.p100)

## Simulated P (W2 >= 100) = 0.8857
```