

# DATA 624: Project 1

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## Overview

We split the work into three sections for Project 1. Individual team members each took lead on individual problem. Jeremy and Julian focused on Part A, Sang Yoon (Andy) and Vinicio worked on Part B, and Bethany took lead on Part C. Juliann created an overall format for the assignment to be used and all team members collectively worked together on reviewing and merging our finished product.

## Dependencies

The following R libraries were used to complete this assignment:

```
library(easypackages)

libraries('knitr', 'kableExtra', 'default')

# Processing
libraries('readxl', 'tidyverse', 'janitor', 'imputeTS', 'tsoutliers', 'lubridate')
# 'xlsx'

# Timeseries
libraries('psych', 'urca', 'forecast', 'timetk', 'fpp2')

# Graphing
libraries('ggplot2', 'grid', 'gridExtra', 'ggfortify', 'ggpubr', 'scales')
```

## Data

Data was stored within our group repository and imported below using the readxl package. Each individual question was solved within an R script and the data was sourced into our main report. For replication purposes, we also made our R scripts available within our appendix. All forecasts have been exported and saved to a single .xlsx file in our github repository folder named forecasts.

```
# Data Aquisition
atm_data <- read_excel("data/ATM624Data.xlsx")
power_data <- read_excel("data/ResidentialCustomerForecastLoad-624.xlsx")
pipe1_data <- read_excel("data/Waterflow_Pipe1.xlsx")
pipe2_data <- read_excel("data/Waterflow_Pipe2.xlsx")

# Source Code
source('~/.GitHub/CUNY_DATA_624/Project One/scripts/Part-A.R')
source('~/.GitHub/CUNY_DATA_624/Project One/scripts/Part-B.R')
source('~/.GitHub/CUNY_DATA_624/Project One/scripts/Part-C.R')
```

# 1 Part A: ATMs

**Instructions:** In part A, I want you to forecast how much cash is taken out of 4 different ATM machines for May 2010. The data is given in a single file. The variable `Cash` is provided in hundreds of dollars, other than that, it is straight forward. I am being somewhat ambiguous on purpose. I am giving you data, please provide your written report on your findings, visuals, discussion and your R code all within a Word readable document, except the forecast which you will put in an Excel readable file. I must be able to cut and paste your R code and run it in R studio. Your report must be professional - most of all - readable, EASY to follow. Let me know what you are thinking, assumptions you are making! Your forecast is a simple CSV or Excel file that MATCHES the format of the data I provide.

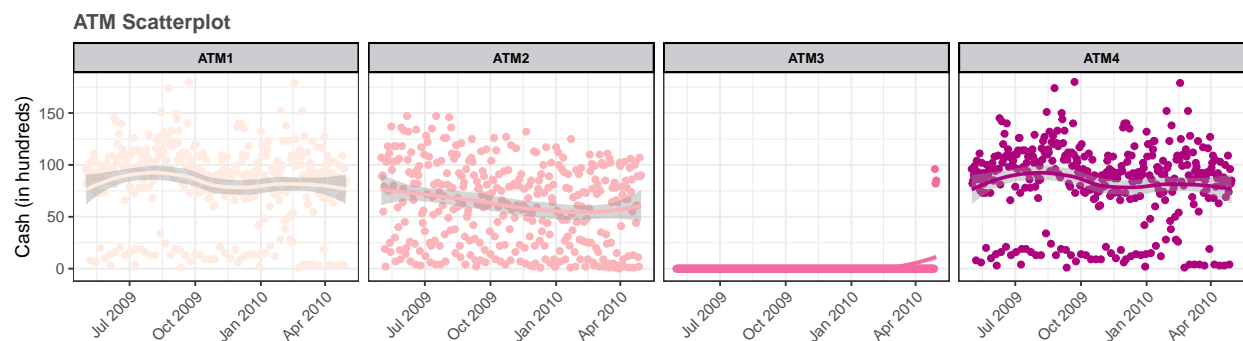
## 1.1 Exploration

The data covers a period of Friday May 1, 2010 through Saturday April 30, 2010. While reviewing the data, we identified that the original data file contained NA values in our `ATM` and `Cash` columns for 14 observations between May 1 and 14, 2010. As these contain no information, we removed these missing values and transformed the dataset into a wide format.

Our initial review also revealed that ATM2 contained one missing value on 2009-10-25 and that ATM4 contained a potential outlier of \$1,123 on 2010-02-09. We replaced both values with the corresponding mean value of each machine.

We examined summary statistics for each ATM time series (a table can be found in the appendix).

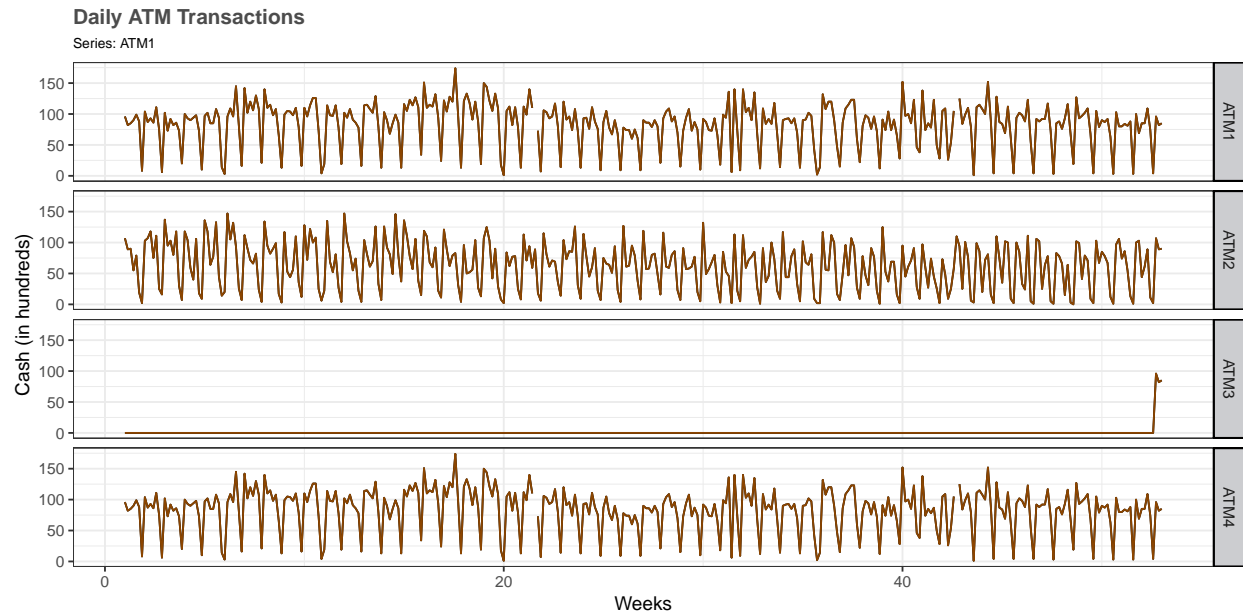
- ATM1 and ATM2 have pretty normal distributions; ATM1's daily mean cash dispensed is \$84, and ATM2's is \$62.
- ATM3 only dispensed cash on the last three days of the time series - as this provides few data points on which to forecast, we'll need to treat it specially.
- ATM4 has a similar mean to ATM1, but skew and kurtosis suggest the impact of an outlier Wednesday, February 10, 2010. If this ATM is located in the Northeastern United States, this may have a relationship to a blizzard which struck on that day.



Last, we used a scatterplot to examine the correlation between cash withdrawals and dates for each machine. We identified similar patterns between ATM1 and ATM4, which show non-linear fluctuations that suggest a potential trend component in these timeseries. ATM2 follows a relatively linear path and decreases overtime. This changes in the last few observations, where withdrawals begin to increase. As mentioned, there are only 3 observed transactions for ATM3 that appear at the end of the captured time period.

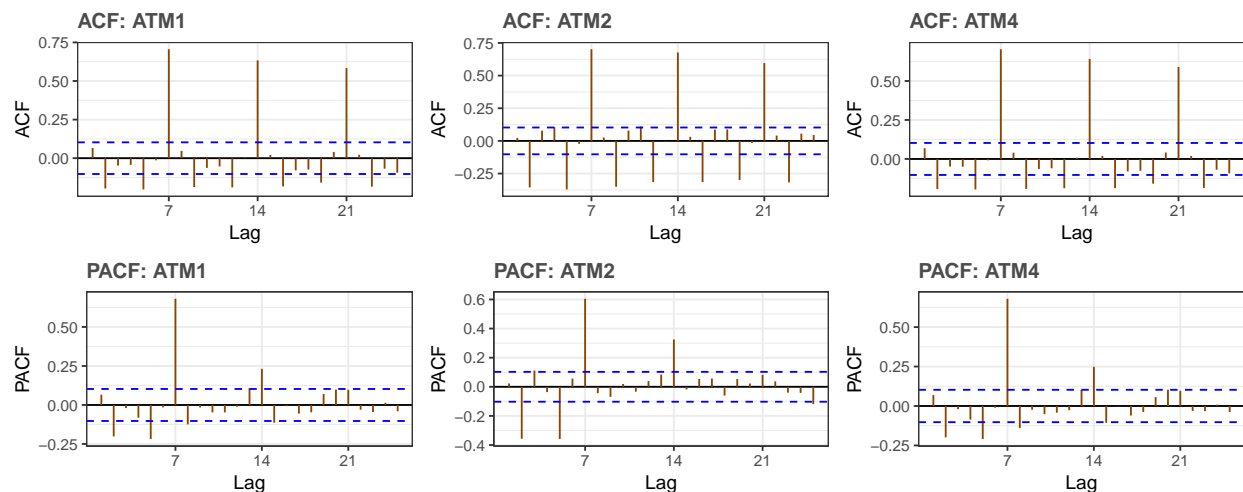
Our cleaned dataframe was then converted into a timeseries format. The time series plots show high weekly variance, for ATM1, ATM2, and ATM4 - consistent with our takeaway from the scatterplots.

These plots also remind us that ATM3 only dispensed cash on 3 days at the end of the timespan, with a daily range between \$82 and \$96. Given the paucity of observations in the training data, the simplest possible approach to forecasting ATM3, averaging, is likely best. Given that ATM3 distributed no cash until April 28, 2010, we'll assume that it was not operating until then and only include the three day window of non-zero observations in the forecast.



## 1.2 Evaluation

We constructed our initial timeseries for ATM1, ATM2, and ATM4 using a weekly frequency. Our ACF plots for each ATM show cases large, decreasing lags starting at 7. This pattern continues in a multiple of seven, which confirms our assumption about seasonality within the observed data. These lags are indicative of a weekly pattern.



Our plots further suggest that the ATM data is non-stationary. We performed a unit root test using the `ur.kpss()` function to confirm this observation. The test results below show that differencing is required on all ATM2 and ATM4 series. ATM1 falls just below the cut-off critical value, but could still benefit from differencing due to the observed seasonal pattern.

Table 1.1: KPSS unit root test

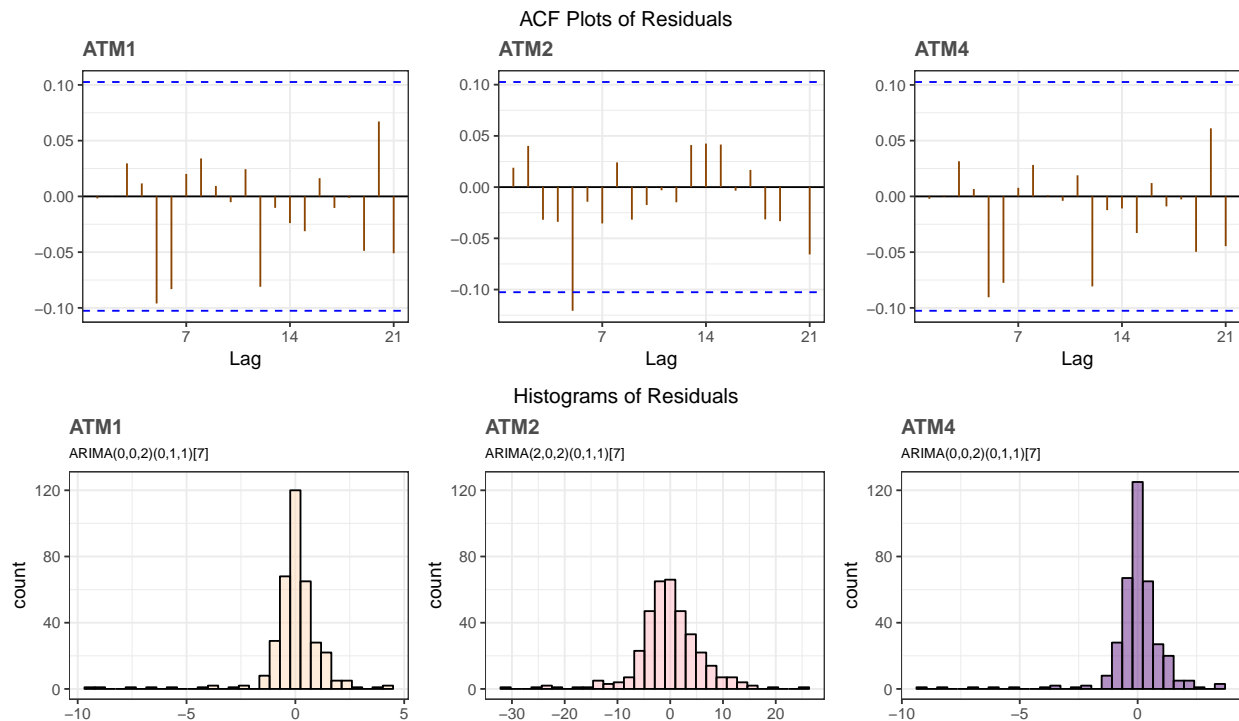
ATM	No-Diff	Diff-1
ATM1	0.4967	0.0219
ATM2	2.0006	0.016
ATM4	0.5182	0.0211

## 1.3 Modeling

We used `auto.arima()` and set  $D=1$  to account for seasonal differencing of our data to select the best ARIMA models for ATM1, ATM2, and ATM4. The full models and accuracy statistics for each series can be viewed in the appendix.

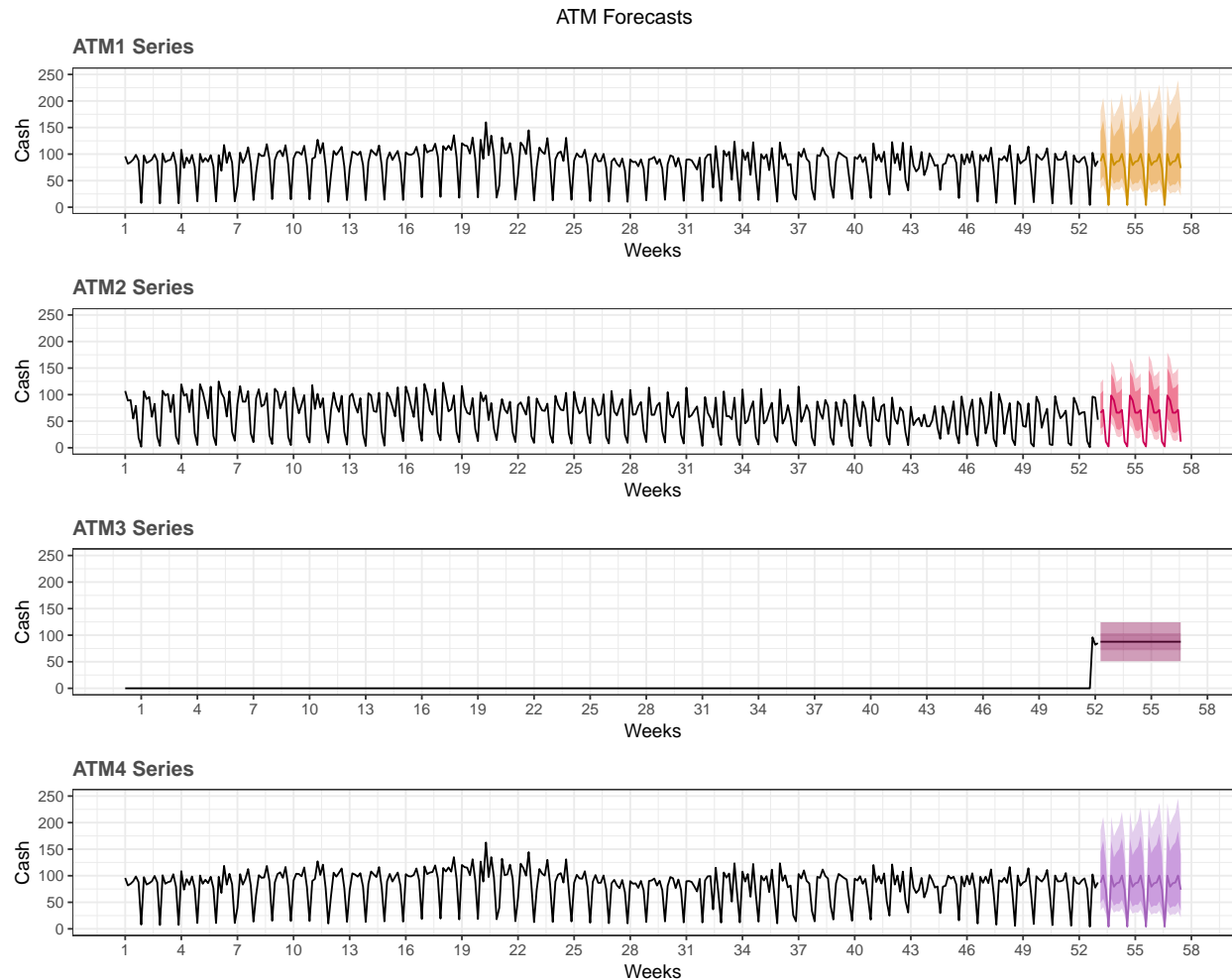
- **ATM1:**  $\text{ARIMA}(0, 0, 2)(0, 1, 1)_7$
- **ATM2:**  $\text{ARIMA}(2, 0, 2)(0, 1, 1)_7$
- **ATM3:** MEAN
- **ATM4:**  $\text{ARIMA}(0, 0, 2)(0, 1, 1)_7$

The residual ACF plots contain no pattern and the lags fall within the critical value, which suggest they are white noise and not autocorrelated. The residual histograms follow a relatively normal distribution that is centered around zero. The p-value from the Ljung-Box test for ATM1, ATM2, and ATM4 all exceeds 0.05, which further supports that residuals happen by chance and the models adequately fit the observed data.



## 1.4 Forecast

A forecast for the month of May will be 31 days in length. We applied a forecast to each series, which spanned across 5 weeks. The numeric forecasts can be viewed in a table output in the appendix section and are also located within our data output folder.



## 1.5 Summary

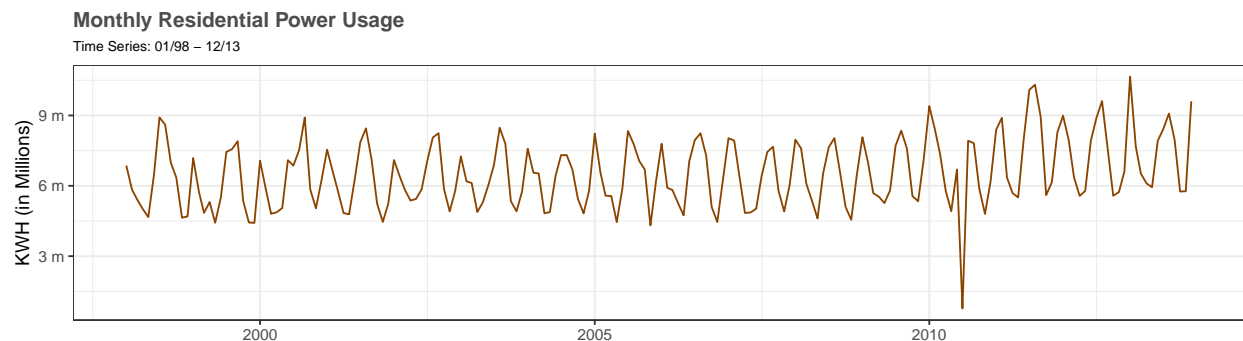
Forecasts for ATM1, ATM2, and ATM4 reprise the clear, persistent weekly pattern found in the historic data, with mid-week troughs and a largely flat trend on a five-week time horizon. ATM1 and ATM4 experience sharper troughs on Wednesdays; ATM2 drops on Tuesdays and bottoms out on Wednesdays. Additionally, ATM2 has a slightly tighter confidence interval than ATM1 and ATM4. The mean forecast for ATM3 based on three data points is a useful estimate insofar as the assumptions it rests on are sound: that the zero observations aren't measurement or data errors, and that the three non-zero observations aren't outliers and in fact convey information about a future pattern.

## 2 Part B: Forecasting Power

**Instructions:** Part B consists of a simple dataset of residential power usage for January 1998 until December 2013. Your assignment is to model these data and a monthly forecast for 2014. The data is given in a single file. The variable 'KWH' is power consumption in Kilowatt hours, the rest is straight forward. Add these to your existing files above - clearly labeled.

### 2.1 Exploration

We observed a missing value in September 2008 and imputed it using `na.interpolation`, which performs a technique in numerical analysis to estimate a value from known data points (in our case, a linear method using first order Taylor polynomials).



Our time series plot reveals annual seasonality; box plots and seasonality demonstrate where power consumption fluctuations occur within each of the cycles. We speculate that this pattern could be due to no major holidays that require power draining decor plus and minimal air conditioning usage during cold months.

### 2.2 Evaluation

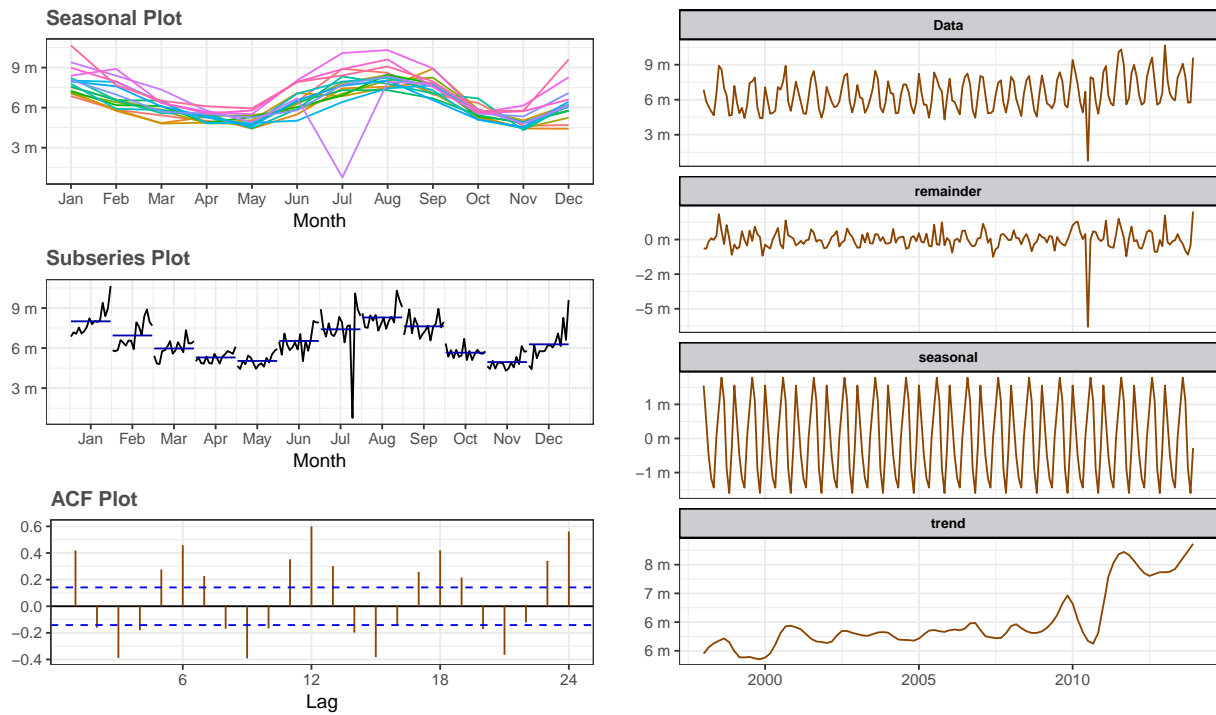
Power consumption increases between the months of June and August, likely in relation to air conditioning usage. It dips from September to November, followed by a small spike in December, which might be due the holidays (perhaps even holiday lights).

Within the overall TS plot a dip in July 2010 is visible; this outlier which may be the result of a power outage during a hot summer month. Using `TSOutliers`, we identify the outlier and replace it using a Box-Cox transformation (by setting the `lambda` argument to automatic).

The ACF plot shows that autocorrelations are well outside the significant space indicating the series is not white noise, non-stationary.



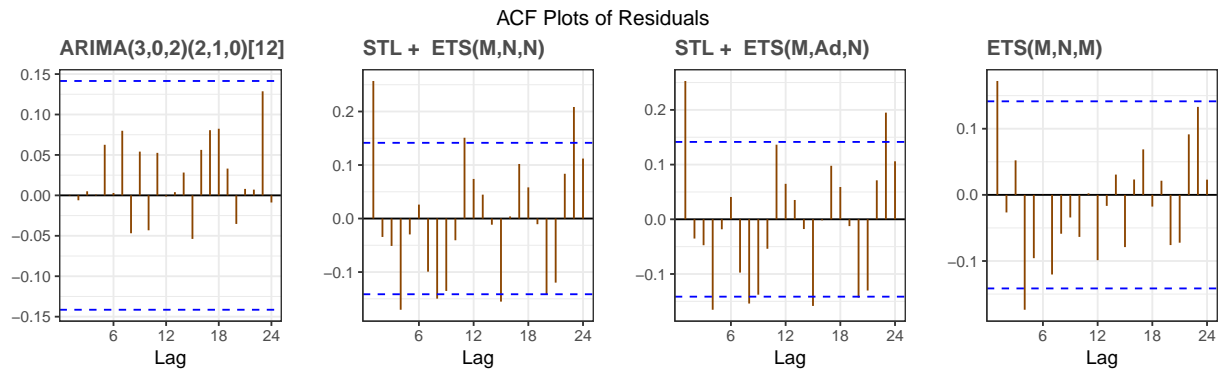
## Diagnostic Plots



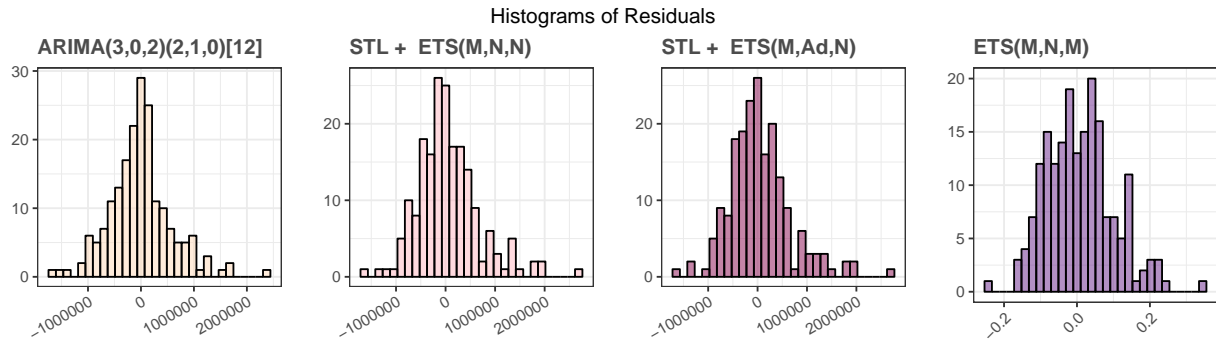
## 2.3 Modeling

We built four different models using ARIMA, STL (with and without dampening), and ETS methods. By checking residuals we can make some preliminary observations on these models' reliability.

The residual ACF plots show residual autocorrelations for each of our models. Model 1 (ARIMA) has less autocorrelation than the other three; it is also well within the 95% limits (indicated by dotted blue lines).

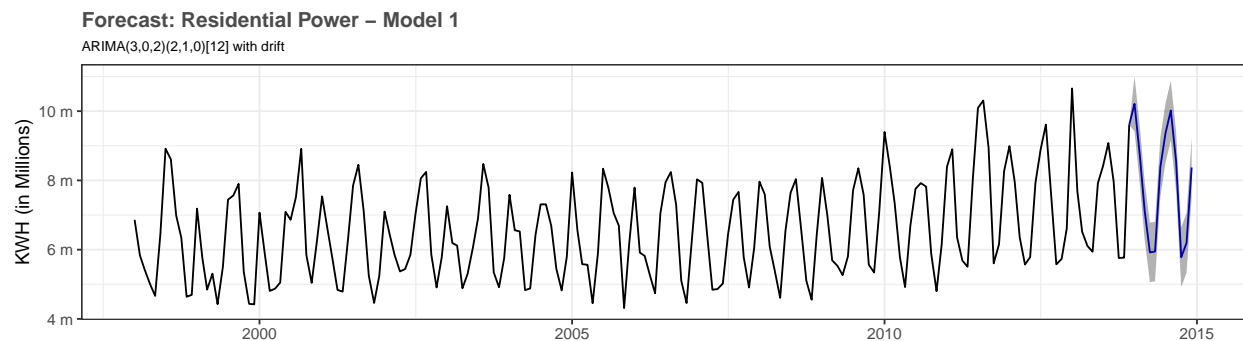


The residuals for each of our models do not deviate substantially from normality. While the residuals of Model 1 (ARIMA) do not have an extended number of bins and this distorts the normality proximity, we can regard the distribution as normal.



A Ljung-Box test yields a p-value  $> 0.05$  for Model 1 (ARIMA), implying that the residuals from other models are not independent, hence not white noise. We will continue with this model for forecasting; a full summary for this and other models attempted is included in the appendix.

## 2.4 Forecast



The `auto.arima()` function performs cross validation on hyperparameter tuning to find the best model with parameters of order and seasonal that minimize AIC. This approach yielded **arima\_model**: ARIMA(3, 0, 2)(2, 1, 0)12 with drift resulting in an of AIC = 5332.24. As other models failed the Ljung-Box test, we develop forecasts based only on the reliable ARIMA model; forecasted values are included in the appendix.

## 2.5 Summary

We implemented a cross-validation method of testing for  $h=12$ , randomly choosing 12 points over the fitted model to measure and take the average of RMSEs. By definition, a lower RMSE on test set indicates a better forecast of the test data.

Using time series cross-validation, we compute RMSE on the test set ( $h=12$ ). If other models had not failed the Ljung-Box test, we use the lowest RMSE as a criterion of selection. Cross-validation test of the seasonal ARIMA model produces an RMSE on test set of around 720k, and on training set of around 589k. We conclude the model is not necessarily overfitted. This finding is consistent with the MAPE on the training set that is less than 7.

[1] "RMSE - Train: 589381.7 ; RMSE - Test: 725175"

### 3 Part C: Waterflow

**Instructions:** Part C consists of two data sets. These are simple 2 columns sets, however they have different time stamps. Your optional assignment is to time-base sequence the data and aggregate based on hour (example of what this looks like, follows). Note for multiple recordings within an hour, take the mean. Then to test appropriate assumptions and forecast a week forward with confidence bands (80 and 95%). Add these to your existing files above - clearly labeled.

#### 3.1 Exploration

Because of the disparities in the data some grooming was necessary:

##### Pipe one

1. 1000 Observations
2. No missing values
3. Multiple reading within each hour
4. 9-days of data

##### Pipe Two

1. 1000 Observations
2. No missing values
3. Single reading on the hour
4. 41-days of data

Pipe One represents 9 days of water flow rate measurements with multiple samples per hour. In order to align with hourly readings from Pipe Two, a mean of all Pipe One rates in a given hour was taken and labeled by its 'floor' (i.e. 9 for mean of all times between 9:00 and 9:59 -inclusive of both bounds). After aggregating, this yielded 236 observations (spanning nine days) for Pipe One and 1000 observations (spanning 41days) for Pipe Two.

The two data sets posed an interesting conundrum. We considered two possible approaches:

(1) Merge the files and use only 236 observations.

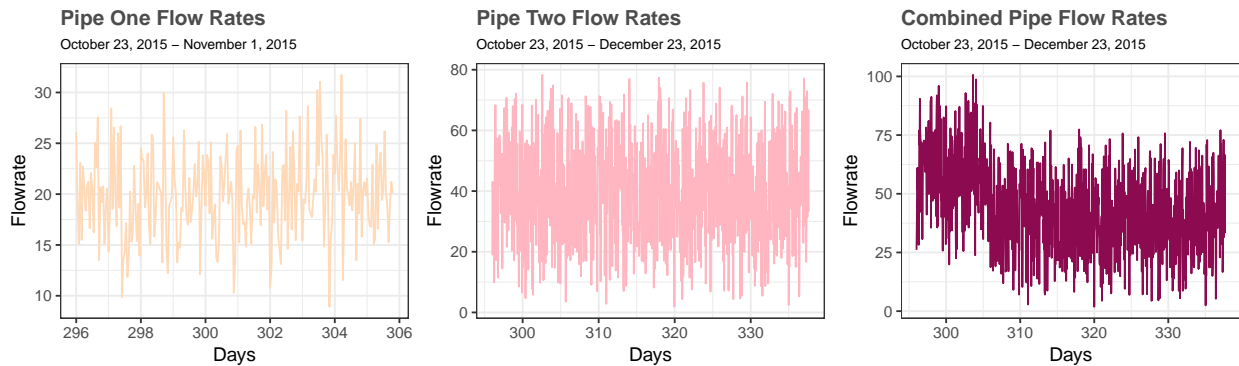
- All forecasts would be based on the combined data.
- This would mean making 168 forecasts (*7daysx24hours*) with only 236 data-points prior.
- All forecasts would start November 1 rather than December 3 (the end of the most recent time series).

(2) Merge the files and use the whole set to make predictions.

- We would have 1000 observations to model prior to forecasts.
- 236 of the observations would be different from the remaining 764, which could both alter the model type and forecast.
- We would forecast from the natural ending of the Pipe Two readings.

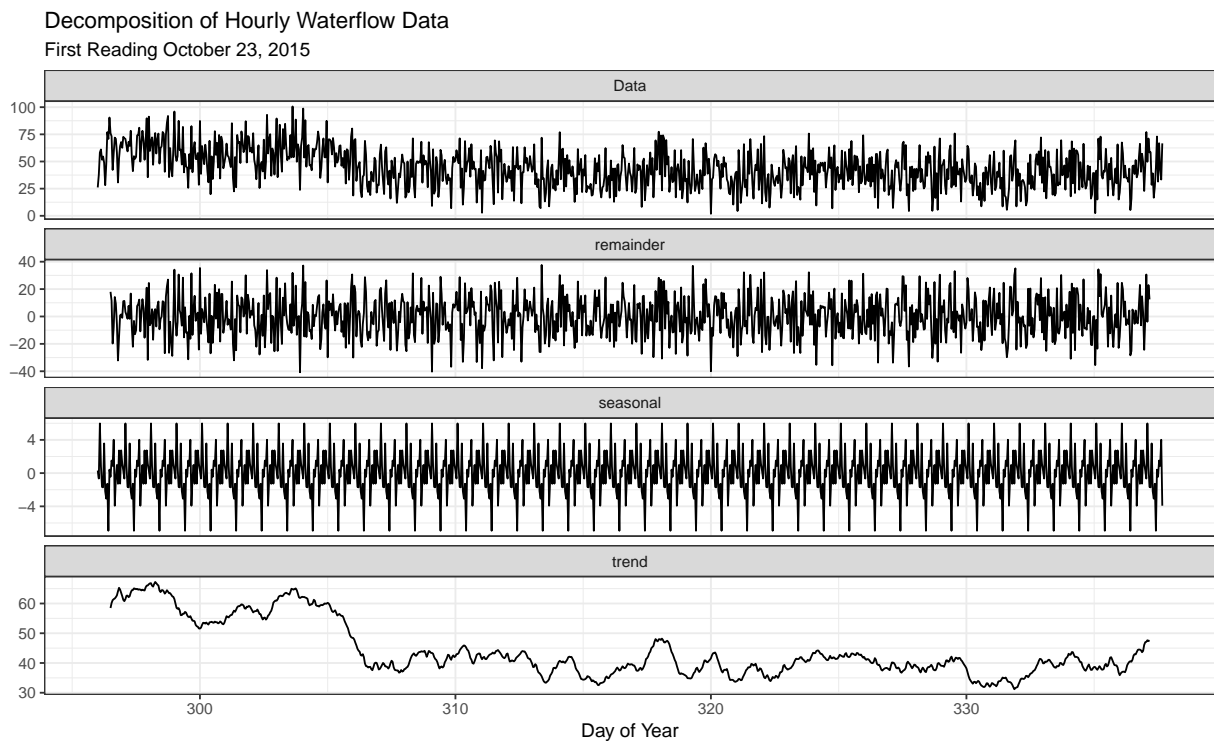
In the end, it made the most sense to model the combined sets in their entirety, so method two was adopted. Because daily periodicity is conceivable for this data, it was important to use a frequency of 24 in converting this data. This entailed numbering by day of year and grooming the time series to start on the 7081 hour (which aligns with October 23 01:00 AM our first merged observation).

## 3.2 Evaluation



### 3.2.1 Decomposition

It is clear from the combined plot that there is a pretty notable change in the trend when the readings from Pipe One wane. We examined the decomposed series for insight into a good model.



From the decomposition, there appears to be a seasonal component, which is in agreement with the prior assessment that there might be a daily flowrate periodicity. Also, as expected, around day 306, where Pipe One flow rates go silent, there is a downward trend followed by a rolling plateau thereafter.

### 3.2.2 Estimating Stationarity

Number of Estimated Differences using `ndiff()`: 1

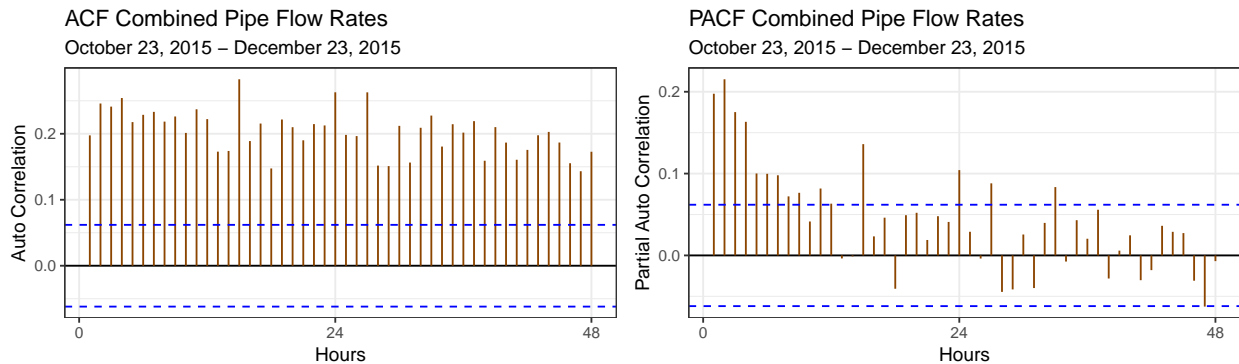
## Augmented Dickey-Fuller Test

```
data: ws
Dickey-Fuller = -6.4409, Lag order = 9, p-value = 0.01
alternative hypothesis: stationary
```

Here we encounter contradictory estimates: `ndiffs()` suggests a difference of 1, and the augmented dicky fuller test suggests that we are stationary as-is. An `auto.arima()` may provide insight into a reasonable starting place.

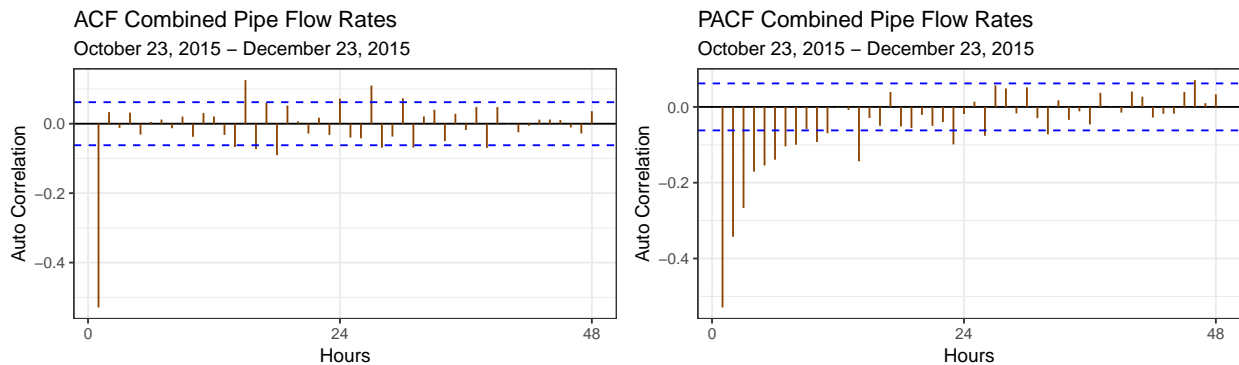
### 3.2.3 Estimating Orders for ARIMA

#### 3.2.3.1 Interpreting the ACF and PACF



As the ACF remains wholly above the critical threshold the series will likely require differencing as suggested by the `ndiffs()`. There is a spike at 24 on both PACF and ACF suggesting a daily period or season that needs to be accounted for in our forecast.

#### 3.2.3.2 Differenced ACF



We examined a final ACF of the differenced data to ensure that a second first-order difference was not needed; while we assume  $d = 1$ , the appropriate value of  $q$  is not so clear, and seasonal orders were in question, so we use `auto.arima()` to help iterate up on the best starting place.

## 3.3 Modeling

The `auto.arima()` function was used in model selection. Using a Box-Cox lambda value to normalize the data yields a  $\lambda = .931552$ . Because models can vary a lot based on the selection criterion, both BIC and AIC models were run using lambda

to estimate a good starting place. We included the transformations in the model (as opposed to outside the model) because we are using the ARIMA function to difference the data automatically for more consistency and flexibility in testing other model orders.

The AICc chose a seasonal ARIMA of the following order:

*ARIMA*(1, 1, 3)(0, 0, 1)[24] AIC=7359.84 AICc=7359.9 BIC=7384.38

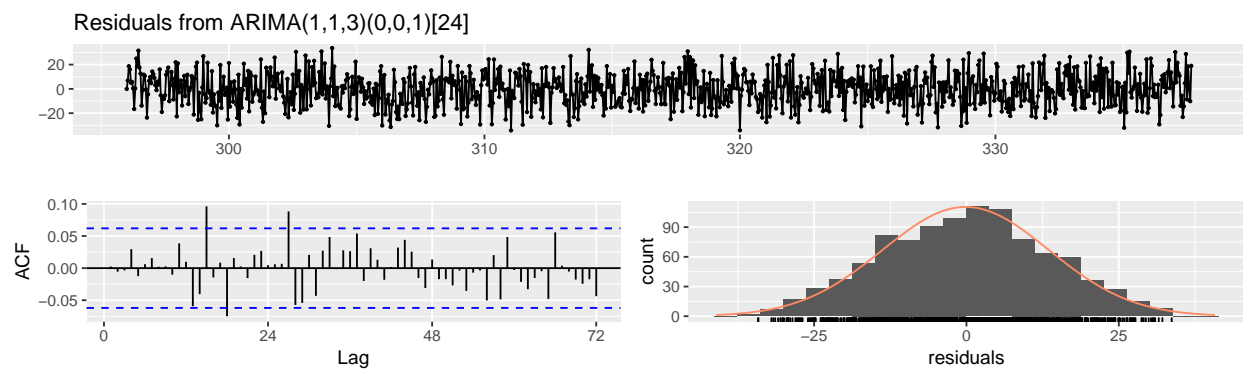
The BIC chose a non-seasonal ARIMA model as follows:

*ARIMA*(2, 1, 1) AIC=8082.22 AICc=8082.26 BIC=8101.85

In both cases, the `auto.arima()` estimated that there needed to be differencing `m`, which was supported by `ndiffs()` and our ACF and PACF plots.

While both models' forecasts degrade pretty quickly towards the series mean, the AICc model generates forecast that consider the variation better before it levels out. Accordingly, we decided to explore and attempt to tune this model to provide more robust predictions.

#### AIC *ARIMA*(1, 1, 3)(0, 0, 1)[24] Residual Plots

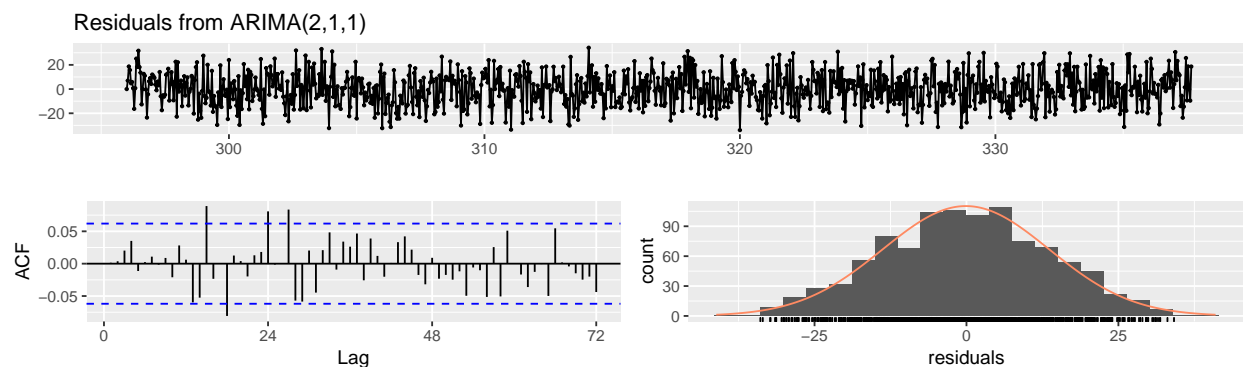


#### Ljung-Box test

data: Residuals from *ARIMA*(1,1,3)(0,0,1)[24]  
 Q\* = 57.362, df = 43, p-value = 0.07027

Model df: 5. Total lags used: 48

#### BIC *ARIMA*(2, 1, 1) Residual Plots



#### Ljung-Box test

data: Residuals from *ARIMA*(2,1,1)

Q\* = 64.403, df = 45, p-value = 0.03029

Model df: 3. Total lags used: 48

### 3.3.1 Interpreting auto.arima()

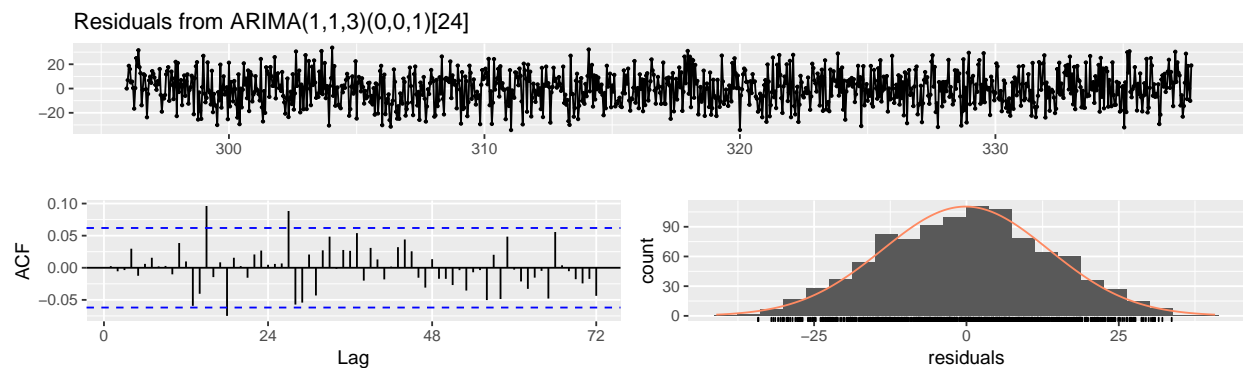
Both the AICc and BIC ARIMA models appear relatively 'white-noisy', with no autocorrelation on the first and 24th observations as well as relatively normal residuals. However, examining the Ljung-Box test for independence made clear that the Seasonal  $ARIMA(1, 1, 3)(0, 0, 1)[24]$  is independent while the  $ARIMA(2, 1, 1)$  is not. This confirmed our suspicion of unaccounted for seasonal variation in the model, which required a seasonal  $MA(1)$  to rectify. To ensure that the best model had been found, we varied p, q, and Q to determine if slight modifications could improve the performance of the model.

### 3.3.2 Manual ARIMA testing

```
Series: ws
ARIMA(1,1,3)(0,0,1)[24]
Box Cox transformation: lambda= 0.9531552
```

```
Coefficients:
      ar1      ma1      ma2      ma3      sma1
      0.7602 -1.7578  0.8286 -0.0614  0.0833
s.e.  0.1857  0.1874  0.1886  0.0324  0.0320
```

```
sigma^2 estimated as 187: log likelihood=-4033.28
AIC=8078.56 AICc=8078.64 BIC=8108
```



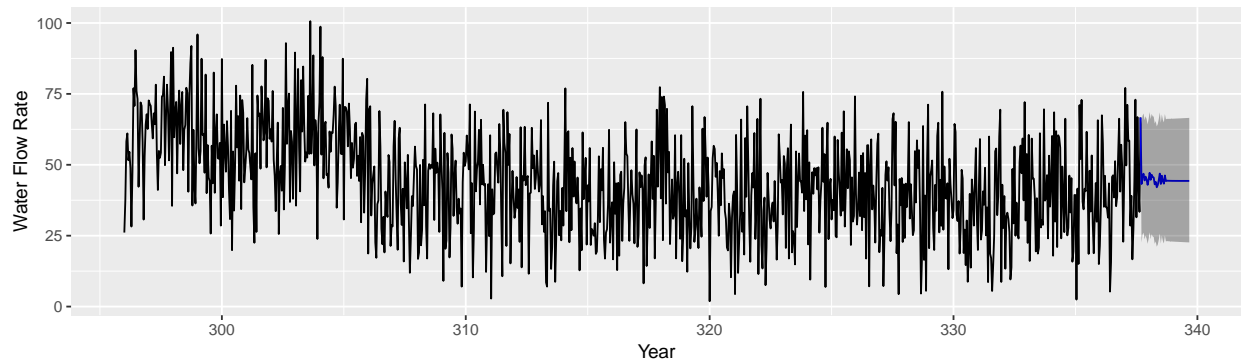
Ljung-Box test

```
data: Residuals from ARIMA(1,1,3)(0,0,1)[24]
Q* = 47.142, df = 31, p-value = 0.03174
```

Model df: 5. Total lags used: 36

## 3.4 Forecast

### 3.4.1 $ARIMA(1, 1, 3)(0, 0, 1)[24]$



### 3.4.2 $ARIMA(2, 1, 3)(0, 0, 1)[24]$

Series: ws

$ARIMA(2, 1, 3)(0, 0, 1)[24]$

Box Cox transformation:  $\lambda = 0.9531552$

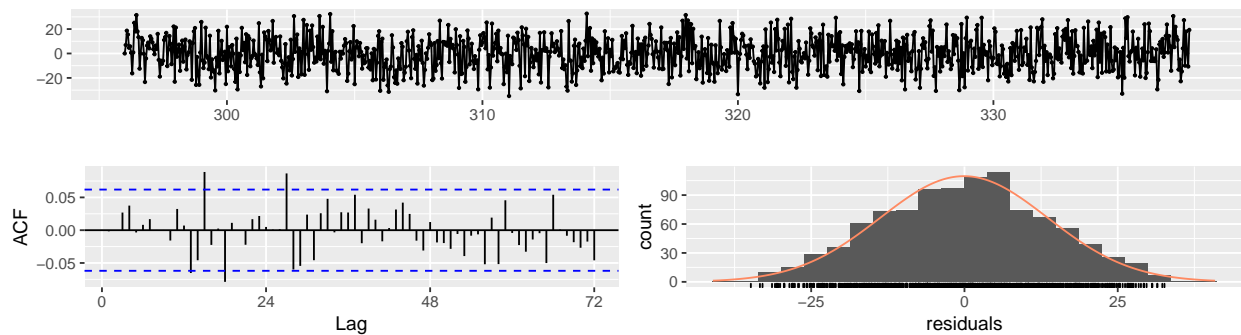
Coefficients:

	ar1	ar2	ma1	ma2	ma3	sma1
	-0.1435	0.1884	-0.8478	-0.2709	0.1621	0.0798
s.e.	NaN	0.5408	NaN	0.6069	0.5320	0.0318

$\sigma^2$  estimated as 187.5: log likelihood=-4034.02

AIC=8082.05 AICc=8082.16 BIC=8116.4

Residuals from  $ARIMA(2, 1, 3)(0, 0, 1)[24]$



Ljung-Box test

data: Residuals from  $ARIMA(2, 1, 3)(0, 0, 1)[24]$

$Q^* = 48.506$ ,  $df = 30$ ,  $p\text{-value} = 0.01764$

Model df: 6. Total lags used: 36

This Ljung-Box test shows unexplained variances in the residuals, indicating that this model is not yet fully realized and inferior to the Seasonal  $ARIMA(1, 1, 3)(0, 0, 1)[24]$ .

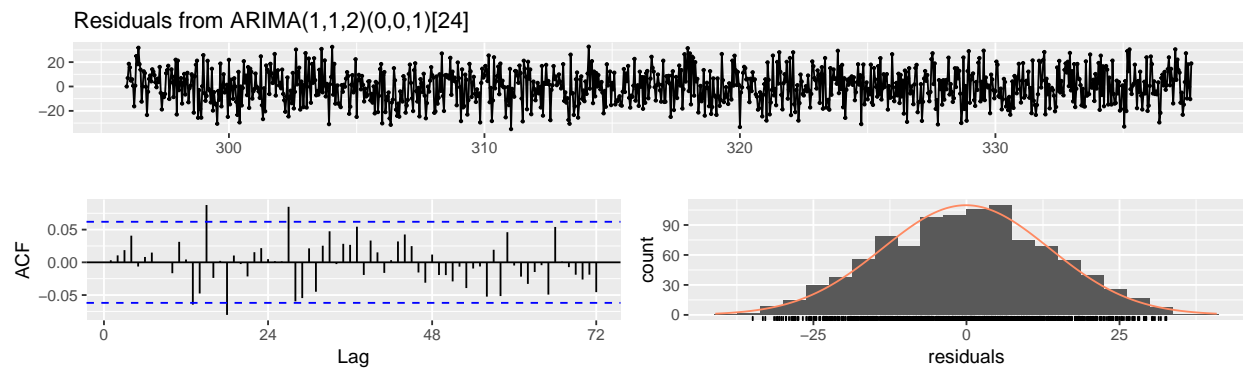


Series: ws  
 ARIMA(1,1,2)(0,0,1)[24]  
 Box Cox transformation: lambda= 0.9531552

Coefficients:

	ar1	ma1	ma2	sma1
	-0.2655	-0.7307	-0.2104	0.0790
s.e.	0.9490	0.9533	0.9121	0.0318

sigma<sup>2</sup> estimated as 187.1: log likelihood=-4034.08  
 AIC=8078.16 AICc=8078.22 BIC=8102.7



Ljung-Box test

data: Residuals from ARIMA(1,1,2)(0,0,1)[24]  
 Q\* = 47.963, df = 32, p-value = 0.03467

Model df: 4. Total lags used: 36

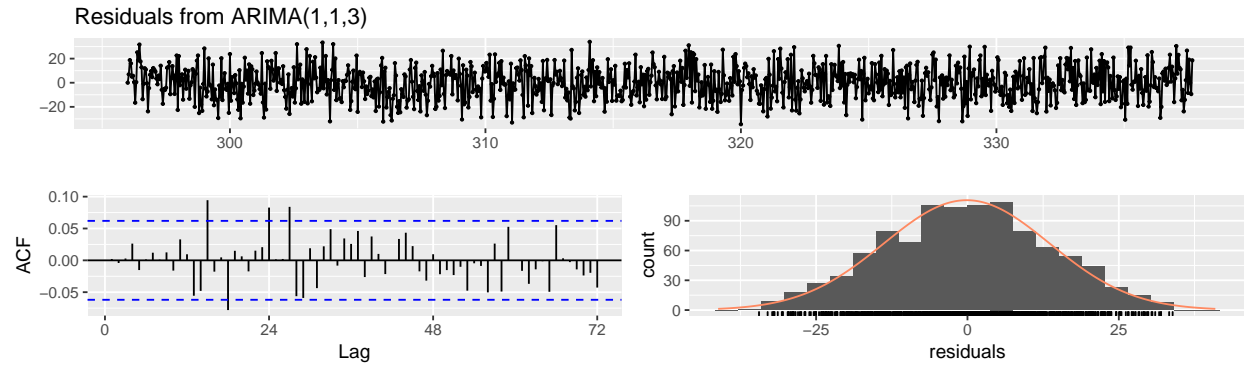
This Ljung-Box also shows unexplained variances in the residuals, indicating that this model is not yet fully realized and inferior to the Seasonal  $ARIMA(1, 1, 2)(0, 0, 1)[24]$ .

Series: ws  
 ARIMA(1,1,3)  
 Box Cox transformation: lambda= 0.9531552

Coefficients:

	ar1	ma1	ma2	ma3
	0.6792	-1.6742	0.7437	-0.0553
s.e.	0.2923	0.2930	0.2903	0.0330

sigma<sup>2</sup> estimated as 188.1: log likelihood=-4036.63  
 AIC=8083.27 AICc=8083.33 BIC=8107.81



Ljung-Box test

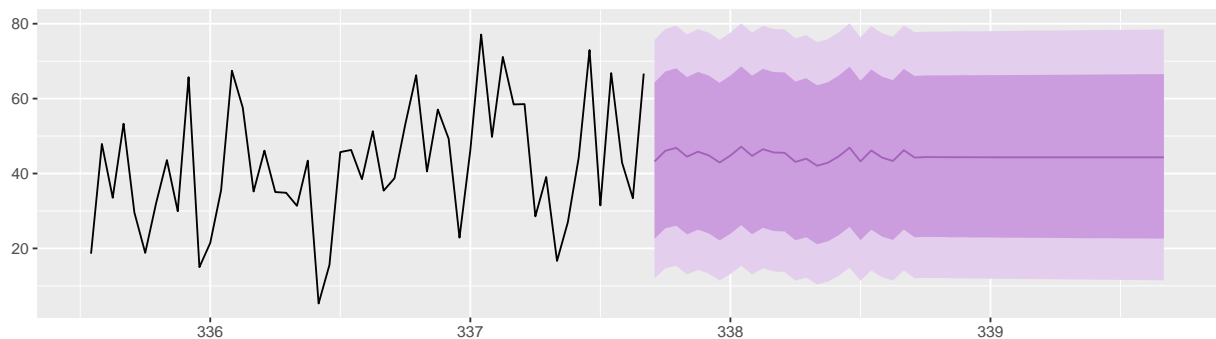
```
data: Residuals from ARIMA(1,1,3)
Q* = 53.61, df = 32, p-value = 0.009708
```

```
Model df: 4. Total lags used: 36
```

This Ljung-Box also shows unexplained variances in the residuals, indicating that this model is not yet fully realized and inferior to the Seasonal  $ARIMA(1, 1, 3)$ .

### 3.4.3 Accepting the `auto.arima()`

Given that the other models show unexplained variance in the residuals, we made our final predictions using the AICc recommended model of  $ARIMA(1, 1, 3)(0, 0, 1)_{24}$ .



### 3.4.4 Forecast Accuracy

	ME	RMSE	MAE	MPE	MAPE	MASE	ACF1
Training set	0.0015679	16.27402	13.23093	-28.76247	50.34448	0.7489308	0.0014339

## 3.5 Summary

Ultimately, we assess that the Seasonal  $ARIMA(1, 1, 3)$  model is marginally useful given its Mean Absolute Percentage of Error. This measure indicates that on average each forecast differs from the actual value on percentage basis by around 50%.

As is visible in the above graph, which depicts the last 150 points in the time series as well as our forecasts, predictions modulate around the mean and deteriorate to it pretty quickly.

The original decomposition revealed very little trend, a lot of seasonality, and a substantial amount of random noise. The extensive random noise component, is assumed to be responsible for the majority of the error, as white noise is never predictable.

# Appendix A

## Summary Statistics

ATM	n	mean	sd	median	trimmed	mad	min	max	range	skew	kurtosis	se
ATM1	365	84.10	36.60	91.0	86.86	25.20	1	180	179	-0.72	0.21	1.92
ATM2	364	62.46	38.90	66.5	62.09	49.67	0	147	147	-0.03	-1.10	2.04
ATM3	365	0.72	7.94	0.0	0.00	0.00	0	96	96	10.93	118.38	0.42
ATM4	365	86.84	65.52	91.0	86.86	25.20	1	1123	1122	10.67	168.66	3.43

## ARIMA Model Summary

ATM1:

Series: ATM1\_ts

ARIMA(0,0,2)(0,1,1)[7]

Box Cox transformation: lambda= 0.2584338

Coefficients:

	ma1	ma2	sma1
	0.1085	-0.1089	-0.6425
s.e.	0.0524	0.0521	0.0431

sigma<sup>2</sup> estimated as 1.726: log likelihood=-606.1

AIC=1220.2 AICc=1220.32 BIC=1235.72

ATM2:

Series: ATM2\_ts

ARIMA(2,0,2)(0,1,1)[7]

Box Cox transformation: lambda= 0.661752

Coefficients:

	ar1	ar2	ma1	ma2	sma1
	-0.4238	-0.8978	0.4766	0.7875	-0.7064
s.e.	0.0592	0.0473	0.0883	0.0608	0.0417

sigma<sup>2</sup> estimated as 38.94: log likelihood=-1162.96

AIC=2337.93 AICc=2338.17 BIC=2361.21

ATM4:

Series: ATM4\_ts

ARIMA(0,0,2)(0,1,1)[7]

Box Cox transformation: lambda= 0.2328582

Coefficients:

	ma1	ma2	sma1
	0.1095	-0.1088	-0.6474
s.e.	0.0524	0.0523	0.0420

sigma<sup>2</sup> estimated as 1.439: log likelihood=-573.5

AIC=1154.99 AICc=1155.11 BIC=1170.52

## Point Forecasts

Table 3.2: ATM Mean Point Forecast

Date	ATM1	ATM2	ATM3	ATM4
2010-05-01	87	66	88	87
2010-05-02	101	71	88	101
2010-05-03	74	11	88	74
2010-05-04	4	2	88	4
2010-05-05	100	98	88	100
2010-05-06	79	89	88	79
2010-05-07	86	66	88	86
2010-05-08	87	66	88	87
2010-05-09	100	71	88	100
2010-05-10	74	11	88	74
2010-05-11	4	2	88	4
2010-05-12	100	98	88	100
2010-05-13	79	89	88	79
2010-05-14	86	66	88	86
2010-05-15	87	66	88	87
2010-05-16	100	71	88	100
2010-05-17	74	11	88	74
2010-05-18	4	2	88	4
2010-05-19	100	98	88	100
2010-05-20	79	89	88	79
2010-05-21	86	66	88	86
2010-05-22	87	66	88	87
2010-05-23	100	71	88	100
2010-05-24	74	11	88	74
2010-05-25	4	2	88	4
2010-05-26	100	98	88	100
2010-05-27	79	89	88	79
2010-05-28	86	66	88	86
2010-05-29	87	66	88	87
2010-05-30	100	71	88	100
2010-05-31	74	11	88	74

## R Script

```
# Load data
atm_data <- read_excel("data/ATM624Data.xlsx")

# clean dataframe
atm <- atm_data %>%
  # create wide dataframe
  spread(ATM, Cash) %>%
  # remove NA column using function from janitor package
  remove_empty(which = "cols") %>%
  # filter unobserved values from May 2010
  filter(Date < as.Date("2010-05-01")) %>% arrange(Date)

atm$ATM2[is.na(atm$ATM2)] <- mean(atm$ATM2, na.rm = TRUE) ## remove NA
atm$ATM4[which.max(atm$ATM4)] <- mean(atm$ATM4, na.rm = TRUE) ## remove outlier

# create TS with weekly frequency & subset data
atm_ts <- atm %>% select(-Date) %>% ts(start=1, frequency = 7)
ATM1_ts <- atm_ts[,1]; ATM2_ts <- atm_ts[,2]; ATM3_ts <- atm_ts[,3]; ATM4_ts <- atm_ts[,4]

#unit root test:
ATM1_ur <-ur.kpss(ATM1_ts); ATM2_ur <-ur.kpss(ATM2_ts); ATM4_ur <-ur.kpss(ATM4_ts)
ATM1d_ur <-ur.kpss(diff(ATM1_ts, lag=7)); ATM2d_ur <-ur.kpss(diff(ATM2_ts, lag=7))
ATM4d_ur <-ur.kpss(diff(ATM4_ts, lag=7))

# AUTO.ARIMA function; set D=1 for seasonal differencing
ATM1_AA <-auto.arima(ATM1_ts, D = 1, lambda = "auto", approximation = F, stepwise = T)
ATM2_AA <-auto.arima(ATM2_ts, D = 1, lambda = "auto", approximation = F, stepwise = T)
ATM4_AA <-auto.arima(ATM4_ts, D = 1, lambda = "auto", approximation = F, stepwise = T)

# Forecast Results
ATM1_fc <- forecast(ATM1_AA,h=31); ATM2_fc <- forecast(ATM2_AA,h=31)
ATM3_fc <- meanf(ATM3_ts[ATM3_ts > 0], h=31); ATM4_fc <- forecast(ATM4_AA,h=31)

# Prepare dataframe for ATM3 mean forecast plotting
ATM3_plotdata_fc <- cbind(seq(from = 366, to = 396), ATM3_fc[[5]], ATM3_fc[[6]],
  ATM3_fc[[7]]) %>% as.data.frame()

colnames(ATM3_plotdata_fc) <- c('Date', 'Point Forecast',
  'Lo 80', 'Lo 95', 'Hi 80', 'Hi 95')
ATM3_plotdata <- ATM3_ts %>% fortify() %>% select(-Index) %>% rename(Cash = Data) %>%
  mutate(Date = as.numeric(row.names(.))) %>% select(Date, Cash) %>%
  full_join(ATM3_plotdata_fc, by = 'Date')

#Revert results back into original form
date <- as.character(seq(as.Date('2010-05-01'), length.out=31, by=1))
ATM_FC <- cbind("Date"=date, "ATM1"=ATM1_fc$mean, "ATM2"=ATM2_fc$mean,
  "ATM3"=ATM3_fc$mean, "ATM4"=ATM4_fc$mean) %>%
  as.data.frame() %>% gather("ATM", "cash", -Date) %>%
  mutate(Date = as.Date(as.character(Date)), Cash = round(as.numeric(cash))) %>%
  select(-cash)
```

```
write_csv(ATM_FC, path = "forecasts/ATM_all_forecast.csv")
```

## Appendix B

### Model Summary

ARIMA:

Series: ts\_data\_o

ARIMA(3,0,2)(2,1,0)[12] with drift

Coefficients:

	ar1	ar2	ar3	ma1	ma2	sar1	sar2	drift
	-0.5606	-0.2216	0.3284	0.8902	0.4827	-0.7249	-0.4152	9018.405
s.e.	0.3992	0.3382	0.0960	0.4120	0.4551	0.0797	0.0841	3027.685

sigma<sup>2</sup> estimated as 387762785879: log likelihood=-2657.12

AIC=5332.24 AICc=5333.3 BIC=5360.97

Training set error measures:

	ME	RMSE	MAE	MPE	MAPE	MASE
Training set	-8455.077	589381.7	427752.5	-0.7944782	6.475365	0.6904053
	ACF1					
Training set	0.0006090194					

STL - MNN:

Forecast method: STL + ETS(M,N,N)

Model Information:

ETS(M,N,N)

Call:

```
ets(y = x, model = etsmodel, allow.multiplicative.trend = allow.multiplicative.trend)
```

Smoothing parameters:

alpha = 0.1159

Initial states:

l = 6317745.8917

sigma: 0.097

AIC	AICc	BIC
6139.631	6139.758	6149.403

Error measures:

	ME	RMSE	MAE	MPE	MAPE	MASE
Training set	56926.03	633571.7	460713.4	-0.03288687	6.945185	0.7436052
	ACF1					
Training set	0.2570241					



# Forecasts:

	Point Forecast	Lo 80	Hi 80	Lo 95	Hi 95
Jan 2014	8992609	8049591	9935628	7550387	10434831
Feb 2014	7908116	6958724	8857508	6456146	9360086
Mar 2014	7079434	6123709	8035158	5617779	8541088
Apr 2014	6435209	5473193	7397225	4963933	7906486
May 2014	6161593	5193326	7129860	4680756	7642430
Jun 2014	7728705	6754226	8703185	6238368	9219043
Jul 2014	8837980	7857327	9818633	7338201	10337759
Aug 2014	9376841	8390053	10363630	7867678	10886004
Sep 2014	8601001	7608114	9593888	7082511	10119490
Oct 2014	6670419	5671470	7669368	5142658	8198180
Nov 2014	6035845	5030870	7040821	4498868	7572822
Dec 2014	7189087	6178120	8200053	5642947	8735226

## STL - MAdN:

Forecast method: STL + ETS(M,Ad,N)

## Model Information:

ETS(M,Ad,N)

## Call:

```
ets(y = x, model = etsmodel, damped = TRUE, allow.multiplicative.trend = allow.multiplicative.trend)
```

## Smoothing parameters:

alpha = 0.1233

beta = 0.0001

phi = 0.8

## Initial states:

l = 5615471.7851

b = 173606.4508

sigma: 0.0972

AIC	AICc	BIC
6143.452	6143.906	6162.997

## Error measures:

	ME	RMSE	MAE	MPE	MAPE	MASE
Training set	54337.68	631081.9	458777.5	-0.07364717	6.937249	0.7404807
ACF1						
Training set	0.2528558					

## Forecasts:

	Point Forecast	Lo 80	Hi 80	Lo 95	Hi 95
Jan 2014	9007707	8060947	9954467	7559763	10455651
Feb 2014	7923348	6969325	8877372	6464295	9382401
Mar 2014	7094774	6133536	8056011	5624687	8564860
Apr 2014	6450635	5482232	7419038	4969591	7931680
May 2014	6177088	5201569	7152607	4685160	7669016
Jun 2014	7744256	6761668	8726843	6241518	9246993

Jul 2014	8853574	7863967	9843182	7340100	10367048
Aug 2014	9392471	8395890	10389052	7868332	10916609
Sep 2014	8616658	7613151	9620166	7081926	10151391
Oct 2014	6686100	5675711	7696488	5140843	8231356
Nov 2014	6051544	5034319	7068769	4495832	7607255
Dec 2014	7204799	6180782	8228817	5638700	8770899

ets - MNM:

Forecast method: ETS(M,N,M)

Model Information:

ETS(M,N,M)

Call:

ets(y = ts\_data\_o)

Smoothing parameters:

alpha = 0.1428

gamma = 0.2119

Initial states:

l = 6189149.8743

s = 0.8984 0.7596 0.938 1.2229 1.2597 1.2396

1.0059 0.7638 0.8078 0.8864 1.0269 1.191

sigma: 0.0967

	AIC	AICc	BIC
	6144.033	6146.760	6192.895

Error measures:

	ME	RMSE	MAE	MPE	MAPE	MASE
Training set	45241.77	628252.5	481520.9	-0.04000239	7.277118	0.7771892
	ACF1					
Training set	0.1927438					

Forecasts:

	Point Forecast	Lo 80	Hi 80	Lo 95	Hi 95
Jan 2014	9917654	8689211	11146096	8038913	11796394
Feb 2014	8522973	7456477	9589469	6891908	10154038
Mar 2014	7012478	6126191	7898765	5657019	8367937
Apr 2014	6208601	5416196	7001006	4996722	7420480
May 2014	5928833	5164834	6692832	4760398	7097269
Jun 2014	7840532	6820624	8860440	6280717	9400347
Jul 2014	9115823	7919004	10312642	7285446	10946200
Aug 2014	9648549	8370229	10926869	7693527	11603571
Sep 2014	8553364	7409986	9696742	6804718	10302010
Oct 2014	6266745	5421655	7111835	4974291	7559199
Nov 2014	5938289	5130560	6746017	4702975	7173603
Dec 2014	8020901	6920610	9121192	6338151	9703651

## R Script

```
library(readxl)
library(tidyverse)
library(forecast)
library(imputeTS)
library(tsoutliers)

# load data
power_data <- read_excel("data/ResidentialCustomerForecastLoad-624.xlsx")
# Time Series
ts_data <- ts(power_data$KWH, frequency = 12, start = c(1998,1))
# Missing value imputation
ts_data <- na_interpolation(ts_data)
# STL decomposition
stl1 <- stl(ts_data, s.window = 'periodic')
# Handling outlier
outlier_func <- tsoutliers(ts_data, iterate = 2, lambda = "auto")

# Time Series - After outlier and imputation handled
ts_data_o <- ts_data # Let's treat outlier handled data separately for Modelling part.
ts_data_o[outlier_func$index] <- outlier_func$replacements

# Model#1: ARIMA
arima_auto <- auto.arima(ts_data_o)
arima_fc <- forecast(arima_auto, h=12)

# Model #2: STL (no-damped) - MNN
stl_ndemp <- stlf(ts_data_o, s.window = "periodic", robust=TRUE, h = 12)

# Model #2-2: STL (damped) - MADN
stl_demp <- stlf(ts_data_o, damped=TRUE, s.window = "periodic", robust=TRUE, h = 12)

# Model #3: ets - MNM
ets_auto <- ets(ts_data_o)
ets_model <- forecast(ets_auto, h=12)

# tsCV - ARIMA -> it takes so much time. I got the results and saved them
##arima_cv <- function(x, h){forecast(Arima(x, order = c(3, 0, 2),
## seasonal = c(2, 1, 0), include.drift = TRUE), h=h)}
##e <- tsCV(ts_data_o, arima_cv, h=12)

# RMSEs -> tsCV takes lot of time to process so just saved the output
#rmse_train_arima <- arima_auto[2]
#rmse_test_arima <- sqrt(mean(e^2, na.rm=TRUE))
rmse_train_arima <- 589381.7
rmse_test_arima <- 725175

# Save output
write.csv(arima_fc, file="forecasts/POWER_ARIMA_FC.csv")
```

## Appendix C

### Sample Forecasts

Table 3.3: First few predictions in the set

DateTime	Point Forecast	Lo 80	Hi 80	Lo 95	Hi 95
2015-12-03 17:00:00	43.21837	22.59441	64.33311	12.00034	75.65243
2015-12-03 18:00:00	46.07958	25.37341	67.24682	14.70394	78.58795
2015-12-03 19:00:00	46.85016	26.06919	68.08732	15.35468	79.46457
2015-12-03 20:00:00	44.49638	23.73897	65.73546	13.06315	77.11903
2015-12-03 21:00:00	45.83029	25.00018	67.13008	14.27275	78.54342
2015-12-03 22:00:00	44.85032	24.01864	66.16308	13.30217	77.58566
2015-12-03 23:00:00	42.92705	22.12687	64.23068	11.45169	75.65293
2015-12-04 00:00:00	44.79836	23.91958	66.16114	13.18081	77.61089
2015-12-04 01:00:00	47.17329	26.20684	68.60103	15.39770	80.08059
2015-12-04 02:00:00	44.70609	23.78935	66.10979	13.03325	77.58190
2015-12-04 03:00:00	46.48881	25.50281	67.94446	14.69153	79.44061
2015-12-04 04:00:00	45.62158	24.64210	67.08023	13.84406	78.57995
2015-12-04 05:00:00	45.52709	24.53307	67.00208	13.72907	78.51085
2015-12-04 06:00:00	43.10639	22.16724	64.55376	11.42233	76.05335
2015-12-04 07:00:00	43.96360	22.98208	65.44444	12.20441	76.96005
2015-12-04 08:00:00	42.07391	21.13451	63.53552	10.40558	75.04543
2015-12-04 09:00:00	42.87840	21.89785	64.37233	11.13642	75.89768
2015-12-04 10:00:00	44.60108	23.55269	66.14421	12.73392	77.69198
2015-12-04 11:00:00	46.89847	25.76875	68.50006	14.88240	80.07419
2015-12-04 12:00:00	43.23698	22.19835	64.78723	11.40350	76.34217
2015-12-04 13:00:00	46.15105	25.01095	67.77192	14.12809	79.35815
2015-12-04 14:00:00	44.24754	23.14728	65.84951	12.30812	77.42997
2015-12-04 15:00:00	43.35173	22.26320	64.95292	11.44256	76.53515
2015-12-04 16:00:00	46.23353	25.04461	67.90461	14.13691	79.51781
2015-12-04 17:00:00	44.25878	22.97866	66.04954	12.05224	77.73211
2015-12-04 18:00:00	44.38901	23.08956	66.19841	12.15194	77.89075
2015-12-04 19:00:00	44.37188	23.05269	66.20224	12.10576	77.90597
2015-12-04 20:00:00	44.35886	23.02043	66.20961	12.06437	77.92439
2015-12-04 21:00:00	44.34896	22.99166	66.21966	12.02661	77.94527
2015-12-04 22:00:00	44.34144	22.96555	66.23177	11.99162	77.96801

## R-Script

```
library(tidyverse)
library(readxl)
library(fpp2)
library(forecast)
library(lubridate)
library(psych)
#library(xlsx)
options(scipen = 999)

# Reading Data
waterflow_1 <- read_excel("data/Waterflow_Pipe1.xlsx")
waterflow_2 <- read_excel("data/Waterflow_Pipe2.xlsx")

# Writing original data to submission file
#file = 'forecasts/water-pipes.xlsx'
#write.xlsx(waterflow_1, file = file, sheetName = "Waterflow Pipe 1",
#col.names = TRUE, row.names = TRUE, append = FALSE)
#write.xlsx(waterflow_2, file=file, sheetName = "Waterflow Pipe 2",
#col.names = TRUE, row.names = TRUE, append = TRUE)

# Grooming, aligning dates and aggregating Data
waterflow_1 <- waterflow_1 %>%
  mutate(DateTime = as.POSIXct(DateTime)) %>%
  group_by(hour=floor_date(DateTime, "hour")) %>%
  summarize(WaterFlow=mean(WaterFlow))

waterflow_2 <- waterflow_2 %>%
  mutate(DateTime = as.POSIXct(DateTime)) %>%
  group_by(hour=floor_date(DateTime, "hour")) %>%
  summarize(WaterFlow=mean(WaterFlow))

# Creating a combined data set
waterflow_all <- merge(waterflow_1, waterflow_2, by = 'hour', all = TRUE) %>%
  mutate(waterflow = rowSums(.[,c("WaterFlow.y", "WaterFlow.x")], na.rm = TRUE)) %>%
  select(hour, waterflow)

# Converting all Three Data Sets to Time Series
w1 <- ts(waterflow_1$WaterFlow, start=c(1,7081), frequency=24)
w2 <- ts(waterflow_2$WaterFlow, start=c(1,7081), frequency=24)
ws <- ts(waterflow_all$waterflow, start=c(1,7081), frequency=24)

#Decomposition of Time Series
ws_decomp <- ws %>%
  decompose() %>%
  autoplot() +
  labs(title = "Decomposition of Hourly Waterflow Data",
       subtitle = 'First Reading October 23, 2015',
       x = 'Day of Year') +
  theme_bw()
```

```

# Checking Differences
ws_diffs<- ws%>%
  ndiffs() #1

# Testing Stationarity
dickie<-tseries::adf.test(ws)

# ACF & PACF

ws_acf <- ggAcf(ws, color = 'darkorange4')+
  labs(title = "ACF Combined Pipe Flow Rates",
        subtitle = 'October 23, 2015 - December 23, 2015',
        y="Auto Correlation", x="Hours")+
  theme_bw()+ theme()

ws_pacf <- ggPacf(ws, color = 'darkorange4')+
  labs(title = "PACF Combined Pipe Flow Rates",
        subtitle = 'October 23, 2015 - December 23, 2015',
        y="Partial Auto Correlation", x="Hours")+
  theme_bw()+ theme()

# Differencesd ACF & PACF

ws_acf_diff <-ggAcf(diff(ws,lag = 1), color = 'darkorange4')+
  labs(title = "ACF Combined Pipe Flow Rates",
        subtitle = 'October 23, 2015 - December 23, 2015',
        y="Auto Correlation", x="Hours")+
  theme_bw()+ theme()

ws_pacf_diff <-ggPacf(diff(ws,lag = 1), color = 'darkorange4')+
  labs(title = "PACF Combined Pipe Flow Rates",
        subtitle = 'October 23, 2015 - December 23, 2015',
        y="Auto Correlation", x="Hours")+
  theme_bw()+ theme()

#Establishing a lambda value for ARIMA transformations
lambda <- BoxCox.lambda(ws)
#Lambda = 0.9531552

# Auto arima's including season components for AICc and BIC
aic<- auto.arima(ws, seasonal = TRUE, ic = 'aicc', lambda = lambda)

bic<-auto.arima(ws, seasonal = TRUE, ic = 'bic', lambda = lambda )

# Plots of auto.arimas
aic_plot <- auto.arima(ws, seasonal = TRUE, ic = 'aicc', lambda = lambda)%>%
  forecast(h=24*7)%>%
  autoplot() +
  labs(title = "AIC selected ARIMA(1,1,3)(0,0,1)[24] ",

```

```

        subtitle = 'October 23, 2015 - December 23, 2015',
        y="Flowrate", x="Days")+
theme_bw()+ theme()

bic_plot<-auto.arima(ws, seasonal = TRUE, ic = 'bic', lambda = lambda )%>%
  forecast(h=24*7)%>%
  autoplot()+
  labs(title = "BIC selected ARIMA(2,1,1)  ",
        subtitle = 'October 23, 2015 - December 23, 2015',
        y="Flowrate", x="Days")+
  theme_bw()+ theme()

# Final AIC from AICc and predictions
final_ws <- Arima(ws, order=c(1,1,3), seasonal=c(0,0,1),lambda=lambda)

preds_ws <-as.data.frame(forecast(final_ws, h = 168))

#Renaming fields for output data
waterflow_all <-waterflow_all%>%
  rename( DateTime = hour,
          WaterFlow = waterflow)
# Formatting forecasts for output data
preds_ws<-preds_ws%>%
  mutate(DateTime = seq(from=as.POSIXct("2015-12-3 17:00", tz="UTC"),
                        to=as.POSIXct("2015-12-10 16:00", tz="UTC"),
                        by="hour") )%>%
  select(DateTime, `Point Forecast`, `Lo 80`, `Hi 80`, `Lo 95`, `Hi 95`)

# Writing forecasts and final data to the 'XLSX' file
#write.xlsx(waterflow_all, file = file, sheetName = "Combined Waterflow",
#col.names = TRUE, row.names = FALSE, append = TRUE)
#write.xlsx(preds_ws, file = file , sheetName = "Forecasts",
#col.names = TRUE, row.names = FALSE, append = TRUE)

```