

Assignment IV

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a) Clearly distinguish the following terms

i. Stochastic Process and Deterministic Data

A stochastic process is a sequence of values of a variable where its future values cannot be predicted with certainty.

Deterministic data is a sequence of values whose values can be predicted with a higher degree of certainty. That is the deviation between their mean line and each of the data points observations is negligible.

ii. Discrete and Continuous stochastic process.

A discrete stochastic process is one where a random variable evolves over time at discrete fixed or random intervals (ie well defined time intervals), while a continuous stochastic process is a function of its time or index parameter.

iii. σ algebra and Filtration

Filtration is the history of a stochastic process until time t for martingales. while σ algebra is a non-empty set of sets that is closed under countable unions, countable intersections and complements

iv. Random Walk and Gambler's Ruin Problem

Let X_1, X_2, \dots, X_n be identically independently distributed (iid) random variables. $S_n = X_1 + X_2 + \dots + X_n$ with the initial condition of $S_0 = 0$. The sequence S_n is a general random walk.

Assume a gambler starts with an initial wealth of 1. The gamble stakes are either win or lose 1 with probabilities p and q respectively. The gamblers aim is to get to the total wealth of N . If the gambler gets to this level of wealth, he has won the game, however if he loses all his wealth that is wealth $= 0$, the gambler is ruined. Thus the aim of the gambler's ruin problem is to get the probability of the gambler being ruined.

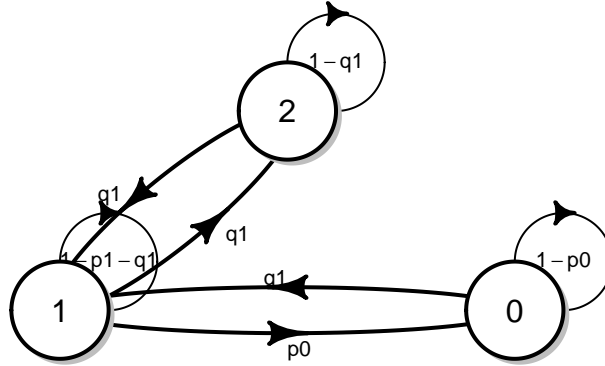
v. Counting Process and Branching Process

Counting processes deal with the occurrences of an event over time, e.g, the number of arrivals at a hospital while a branching process is one that models the growth of populations for example the spread of a disease

vi. Birth and Death Process

It is a markov chain whose diagram looks as follows.

```
## Loading required package: shape
```



with the transition matrix shown as:

```
print(transition_mat)
```

```
##    0      1      2
## 0 "1-p0" "p0"   "0"
## 1 "q1"   "1-p1-q1" "q1"
## 2 "0"    "q1"    "1-q1"
```

it has a finite number of states but for this case, the state m is 2. At any state i of the Markov chain you have the chance to either go forward one step or back one step or remain at the same point as shown in the chain above. For example at state 1 you can either move to the next forward state 2 with a probability of q_1 or back to state 0 with a probability of p_0 , or remain at the same state 1 with a probability of $1 - p_1 - q_1$.

An example of such a chain is the queuing line at a shop. The line (i) can either grow when a new customer comes in that is state $i+1$ or reduce when the attendant serves a customer that is state $i-1$ or remain the same if no one is served or no new customers enter the line that is i .

vii. Static Simulation Model and Dynamic Simulation Model

A static simulation model is one which models a system at a particular point in time while dynamic simulation models a system as it evolves over time.

viii. Mathematical Model and Simulation Model

A mathematical model describes a system by a set of variables and a set of equations that establish the relationship between the variables, for example volume of a cuboid is given by $l \times w \times h$, while a simulation model is a parameterised model that is solved using computers since they are too complex to solve analytically.

ix. Monte Carlo Simulation and Queuing System

A queuing system is one composed of a queue, arrivals and a system offering a service, thus it is a system where there are servers that provide a service to arriving customers, who form a queue or line while waiting to be served. Monte-carlo simulations however is a model used to determine the probability of different outcomes of an uncertain event.

x. Markov and Poisson processes

A poisson process is one where there is a series of discrete events where the average time between events λ is known but timing in which the events occur is random.

A markov process is one where the future is independent of the past given the present.

xi. Queuing and Renewal processes

A queing process is one where there are servers that provide a service to arriving customers, who form a queue or line while waiting to be served.

A renewal process is a stochastic process where events/arrivals occur randomly in time. The assumption here is that the arrival intervals are iid.

xii. Martingale and Brownian motion

A martingale is a stochastic process X_t which the best estimate of the process at time $t+1$ that is X_{t+1} is X_t that is the best estimate for tomorrow (X_{t+1}) given all the information present until today (F_t) (filtration rate) is the value observed today (X_t).

A Brownian motion is the continuous version of a simple symmetric random walk. For a simple symmetric random walk the random variable X_i takes the value either +1 or -1 with equal probability for the Brownian motion, this step is reduced progressively from 1 until it is infinitesimal.

xiii. Hidden Markov model and semi-Markov process

A hidden markov model is composed of two parts, the markov chain and each state's random emissions, here the chain is unobservable but the emissions are observable, while a semi markov process are generalizations of markov processes where the transition time from each state to another is a random variable.

b) Collins bought a share of stock for \$12, and it is believed that the stock price moves (day by day) as a simple random walk with $p = 0.58$. What is the probability that Collins' stock reaches the high value of \$35 before the low value of \$8?

solution

$$\frac{1 - \left(\frac{q}{p}\right)^b}{1 - \left(\frac{q}{p}\right)^{a+b}}$$

where a is the up movement and b is the down movement

```
origin = 12
p = 0.58
q = 1-p
up = 35
down = 8
```

```

a = up-origin
b = origin-down
p_35 = (1-(q/p)^b)/(1-(q/p)^(a+b))
print(p_35)

```

```
## [1] 0.7251491
```

c) Explain clearly the difference between the following terms as used in Markov Chains

i. Communicating class and absorption state

An absorption state i is one which is impossible to leave.

Points i and j in a markov chain are said to communicate if i is reachable to j and j is reachable to i . This makes the markov chain irreducible and thus a communicating class.

ii. Recurrence and nonrecurrence state

A recurrent state i is recurrent if and only if one starts at state i and there is a probability of 1 that one will eventually return to state i .

A non recurrent or transient state is one where if one starts at a state i the probability of returning to that state i is less than 1

iii. Periodicity and aperiodic

A state i is periodic if one starts at state i and the number of steps to get back to that step i is greater than one.

A state i is aperiodic if one starts at state i and the number of steps to get back to that step i is equal to one.

iv. Ergodic chain and transient state

An ergodic chain is one that satisfies three conditions that is the chain is aperiodic, recurrent and irreducible.

A transient state is one where if one starts at a state i the probability of returning to that state i is less than 1.

v. Reducible and irreducible

An irreducible chain is one where it is possible to move from any state to any other state regardless of if the path is direct or indirect.

An irreducible chain is one where it is not possible to move from any state to any other state.

d) Consider an M/M/1 model at steady state, with μ as the service mechanism rate and λ as the arrival rate. Let $P_n(t) = P[n \text{ customers in the system at time } t]$ (Probability that there are n customers at time t). Derive L_q, L_s, W_q and W_s .

Let:

P_n represent the probability that n customers are in a system

L be the expected number of customers in the system

L_q be the expected number of customers in the queue

L_s be the expected number of customers in service

W be the average waiting time in system

W_q be the average waiting time in the queue

W_s be the average waiting time in service

λ be the arrival rate

μ be the rate of departure / service rate

At the steady state what enters the system is the same as what leaves the system, therefore:

$$\lambda P_{n-1} = \mu P_n$$

Solving the differential equation iteratively we get:

$$\lambda P_0 = \mu P_1 = \frac{\lambda}{\mu} P_0, \dots \text{at } n = 1$$

$$\lambda P_1 = \mu P_2 = \frac{\lambda}{\mu} P_1 = \left(\frac{\lambda}{\mu}\right)^2 P_0, \dots \text{for } n=2$$

Which generalizes to yield:

$$P_n = \left(\frac{\lambda}{\mu}\right)^n P_0, \dots \text{for } n = 0, 1, 2, \dots, n$$

Let $\frac{\lambda}{\mu}$ be the fraction of time the server is processing jobs, thus:

$$P_n = \rho^n P_0$$

To get the value of P_0 , P_n must meet the condition:

$$\begin{aligned} \sum_{n=0}^{\infty} P_n &= 1 \Rightarrow \sum_{n=0}^{\infty} \rho^n P_0 \Rightarrow P_0 \sum_{n=0}^{\infty} \rho^n \\ &\Rightarrow P_0 (\rho^0 + \rho^1 + \dots + \rho^n) \end{aligned}$$

for $\rho < 1$ we get:

$$\begin{aligned} P_0 \sum_{n=0}^{\infty} \rho^n &\Rightarrow P_0 \left(\frac{1}{1-\rho}\right) = 1 \\ P_0 &= 1 - \rho \Rightarrow 1 - \frac{\lambda}{\mu} \Rightarrow (1 - \rho) \rho^n \Rightarrow \left(\frac{\lambda}{\mu}\right)^n \left(1 - \frac{\lambda}{\mu}\right) \end{aligned}$$

Getting the average number of customers in the system L_s

$$E(\text{n of pple in system}) = \sum_{n=0}^{\infty} n P_n \Rightarrow \sum_{n=0}^{\infty} n (1 - \rho) \rho^n$$

$$\Rightarrow (1 - \rho) \sum_{n=0}^{\infty} n \rho^n$$

The above is an increasing GP series, thus:

$$\sum_{n=0}^{\infty} n r^n = \frac{\rho}{(1 - \rho)^2}$$

$$L_s \Rightarrow (1 - \rho) \frac{\rho}{(1 - \rho)^2} \Rightarrow \frac{\rho}{(1 - \rho)}$$

Since $\frac{\lambda}{\mu} = \rho$, then :

$$L_s = \frac{\lambda}{(\mu - \lambda)}$$

Getting the average number of entities in the queue L_q

$$E(\text{n of pple in queue}) = L_q = \sum_{n=1}^{\infty} (n - 1) P_n$$

$$\Rightarrow \sum_{n=1}^{\infty} (n - 1) (1 - \rho) \rho^n \Rightarrow \sum_{n=1}^{\infty} n P_n - \sum_{n=1}^{\infty} P_n$$

$$\Rightarrow L_s - (1 - P_o) \Rightarrow \frac{\rho}{1 - \rho} - [1 - (1 - \rho)]$$

$$L_s = \frac{\rho}{1 - \rho}$$

Thus:

$$\frac{\rho}{1 - \rho} - 1 - [(1 - \rho)] \Rightarrow \frac{\rho^2}{1 - \rho}$$

Since $\rho = \frac{\lambda}{\mu}$, we get:

$$L_q \Rightarrow \frac{\lambda^2}{\mu(\mu - \lambda)}$$

According to little's law:

$$\text{average time in system} = W_s = \frac{L_s}{\lambda}$$

$$W_s \Rightarrow \frac{1}{\mu - \lambda}$$

and

$$\text{average time in que} = W_q = \frac{L_q}{\lambda} \Rightarrow \frac{\lambda}{\mu(\mu - \lambda)} \Rightarrow \frac{\rho}{\mu - \lambda}$$

e) Clearly specify five components of a Hidden Markov Model

- A set of finite states S that is (S_1, \dots, S_n)
- Finite set of observations V for each state (V_1, \dots, V_m)
- Initial state distribution π_i which represents the fraction of time one spends at a particular state i
- A state transition probability distribution A , (a_{ij}) which represents the transition probability of moving from a state i to state j .
-

f) Use Chapman Kolmogorov postulates to derive the Poisson Process. Also derive the mean and variance of the Poisson process.

A process is a counting process $\{N_t, t \geq 0\}$ with a rate of $\lambda > 0$ if

- Counting starts at 0 at time 0 that is $N_t = 0$
- The process has both independent and stationary increments.
- $P_r(N_h - N_0 = 1) = \lambda h + o(h)$ that is the probability of observing one event after a small interval h is given as $\lambda h + o(h)$
- $P_r(N_h - N_0 \geq 2) = o(h)$ that is the probability of observing one event after a small interval h is given as $o(h)$

The number of events happening in any interval of length t is poisson distributed with mean $\mu = \lambda t$. That is $P_r(N_{t+s} - N_s = n) = \frac{(\lambda t)^n}{n!} e^{-\lambda t}$ for all $s, t \geq 0$.

$P_n(t) = P_r(N_{t+s} - N_s = n)$ is the distribution of n events in any time interval thus $P_0(t)$ which is the probability of observing 0 events in time interval t and $P_1(n)$ which is which is the probability of observing 1 events in time interval n .

The following 3 steps are applied to show that $P_n(t)$ belongs to the poisson distribution

- Find an ODE for $P_0(t)$ and solve to show that the statement is true for $n=0$
- Find a system of ODEs for $P_n(t) : n > 0$
- Solve the system of ODEs which describes the probabilities for all n .

Step 1: ($n=0$)

$$P_o(t+h) = P_r(N_{t+h} - N_o = 0) = P_r(N_{t+h} - N_t = 0, N_t - N_o = 0)$$

Applying the characteristic of independent increments.

$$P_r(N_{t+h} - N_t = 0, N_t - N_o = 0) = P_r(N_{t+h} - N_t = 0)P_r(N_t - N_o = 0)$$

$$P_r(N_{t+h} - N_t = 0)P_r(N_t - N_o = 0) = P_o(h) \times P_o(t)$$

but $P_o(h) = (1 - p_1(h) - P_r(N_{t+h} - N_t \geq 2))$

Thus:

$$P_o(h) \times P_o(t) = (1 - \lambda h + o(h)) \times P_o(t)$$

Rearranging the above and dividing by h:

$$\frac{P_0(t+h) - P_0(t)}{h} = -\lambda P_0(t) + \frac{o(h)}{h}$$

Taking limits:

$$\lim_{h \rightarrow 0} \frac{P_0(t+h) - P_0(t)}{h} = -\lambda P_0(t) + \lim_{h \rightarrow 0} \frac{o(h)}{h}$$

Solving the ODE:

$$\begin{aligned} \frac{\partial}{\partial t} P_0(t) &= P'_0(t) = -\lambda P_0(t) + 0 \\ \frac{\partial}{\partial t} \log(P_0(t)) &= P'_0(t) = -\lambda \end{aligned}$$

Thus:

$$P_0(t) = e^{-\lambda t + c} = K e^{-\lambda t}$$

If $t = 0$ then the $P_0(0)$ (probability of observing 0 events in 0 time is 0) yields $k=1$ thus:

$$P_0(t) = e^{-\lambda t}$$

step 2: getting system of ODEs for $n > 0$

Let A be the interval $(0, t)$ and B the interval $(t, t+h)$. Then:

$$\begin{aligned} P_0(t+h) &= P_r(N_{t+h} - N_0 = n) = \underbrace{P_r(N_{t+h} - N_t = 0, N_t - N_0 = n)}_{\text{where n in A and 0 in B}} + \underbrace{P_r(N_{t+h} - N_t = 1, N_t - N_0 = n-1)}_{\text{or n-1 in A and 1 in B}} + \\ &\underbrace{\sum_{k=2}^n P_r(N_{t+h} - N_t = k, N_t - N_0 = n-k)}_{\text{or the remaining p terms}} = P_0(h) \times P_n(t) + P_1(h) \times P_{n-1}(t) + o(h) \\ &= \underbrace{(1 - \lambda h + o(h))}_{\text{postulates 3,4}} P_n(t) + \underbrace{(\lambda h)}_{\text{postulate 3}} P_{n-1}(t) + o(h) \end{aligned}$$

Rearrange and divide by h to find:

$$\frac{P_n(t+h) - P_n(t)}{h} = -\lambda P_n(t) + \lambda P_{n-1}(t) + \frac{o(h)}{h}$$

Taking limits of the above:

$$\lim_{h \rightarrow 0} \frac{P_n(t+h) - P_n(t)}{h} = -\lambda P_n(t) + \lambda P_{n-1}(t) + \lim_{h \rightarrow 0} \frac{o(h)}{h}$$

which yields:

$$\frac{\partial}{\partial t} P_n(t) = P'_n(t) = -\lambda P_n(t) + \lambda P_{n-1}(t)$$

for $n = 1, 2, 3, \dots$

Step 3: Getting the MGF associated with the probabilities $P_n(t)$

The Probability Generating Function (PGF), is defined as:

$$G(z) = \sum_{\text{Support}} P_r(X = x) z^x$$

For the random variable $N_t - N_0$, we have:

$$G(z, t) = \sum_{n=0}^{\infty} P_n(t) z^n$$

Given the addition of the time argument (for all $t \geq 0$), for any t the distribution will be the same expression but its parameters depend on t .

The infinite system of ODEs which characterize the probabilities are given as:

$$P'_0(t) = 0 - \lambda P_0(t) P'_1(t) = \lambda P_0(t) - \lambda P_1(t) P'_2(t) = \lambda P_1(t) - \lambda P_2(t) P'_3(t) = \lambda P_2(t) - \lambda P_3(t) \dots P'_n(t) = \lambda P_{n-1}(t) - \lambda P_n(t) \dots$$

Multiplying each equation with a corresponding power of the dummy variable z , we get:

$$\begin{aligned} P'_0(t) &= 0 - \lambda P_0(t) z^0 \\ (P'_1(t) &= \lambda P_0(t) - \lambda P_1(t)) z^1 \\ (P'_2(t) &= \lambda P_1(t) - \lambda P_2(t)) z^2 \\ (P'_3(t) &= \lambda P_2(t) - \lambda P_3(t)) z^3 \\ &\vdots \\ &\vdots \\ &\vdots \\ (P'_n(t) &= \lambda P_{n-1}(t) - \lambda P_n(t)) z^n \\ &\vdots \\ &\vdots \end{aligned}$$

Adding them up we get:

$$\begin{aligned} \sum_{n=0}^{\infty} P'_n(t) z^n &= \sum_{n=1}^{\infty} \lambda P_{n-1}(t) z^n - \sum_{n=0}^{\infty} \lambda P_n(t) z^n \\ &= \lambda z \sum_{n=1}^{\infty} P_{n-1}(t) z^{n-1} - \lambda \sum_{n=0}^{\infty} P_n(t) z^n \dots \dots \dots (\text{taking out } z \text{ and } \lambda) \\ &= \lambda z \sum_{n=1}^{\infty} P_m(t) z^m - \lambda \sum_{n=0}^{\infty} P_n(t) z^n \dots \dots \dots (m = n - 1) \\ G'(s, t) &= \lambda z G(z, t) - \lambda G(z, t) \dots \dots \dots (Def. + \frac{\partial}{\partial t} G(z, t) = G'(z, t)) \end{aligned}$$

The above gives the ODE:

$$G'(z, t) = \lambda(z - 1)G(z, t)$$

subject to the boundary conditions

$$G(1, t) = 1, G(z, 0) = 1$$

On solving the ODE we get:

$$G(z, t) = \exp(\lambda t(z - 1))$$

Comparing the above to the Poisson distribution which is:

$$G(z) = \exp(\mu(z - 1))$$

One notes that by the uniqueness of the PGF that:

$$N_t - N_0 \text{ Poiss}(\mu = \lambda t)$$

g) A certain stock price has been observed to follow a pattern. If the stock price goes up one day, there's a 25% chance of it rising tomorrow, a 35% chance of it falling, and a 40% chance of it remaining the same. If the stock price falls one day, there's a 25% chance of it rising tomorrow, a 50% chance of it falling, and a 25% chance of it remaining the same. Finally, if the price is stable on one day, then it has a 50-50 change of rising or falling the next day.

i. Generate the transition matrix

solution

```
library(markovchain)
```

```
## Package: markovchain
```

```
## Version: 0.9.0
```

```
## Date: 2022-07-01
```

```
## BugReport: https://github.com/spedygiorgio/markovchain/issues
```

```
transition_mat=matrix(nrow = 3,ncol=3,c(0.25,0.25,0.5,0.35,0.5,0.5,0.4,0.25,0))
```

```
statesNames <- c("Up", "Down", "Stable")
```

```
markovB <- new("markovchain", states = statesNames, transitionMatrix =transition_mat, name = "A markovchain")
```

```
print(markovB)
```

```
##           Up Down Stable
```

```
## Up      0.25 0.35  0.40
```

```
## Down    0.25 0.50  0.25
```

```
## Stable  0.50 0.50  0.00
```

ii. Draw the Markov chain using R

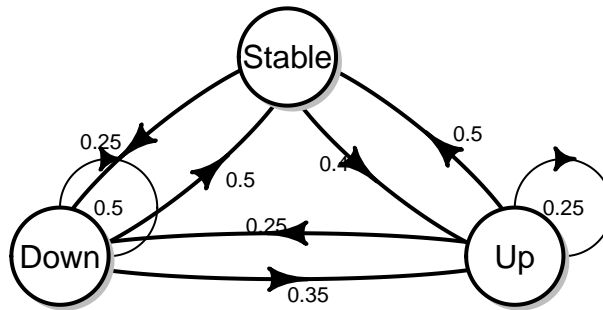
solution

```
library(diagram)
```

```
rownames(transition_mat) = statesNames
```

```
colnames(transition_mat) = statesNames
```

```
plotmat(transition_mat,relsize = .65,cex=0.7)
```



iii. Determine if the chain is Ergodic

solution

The chain is ergodic since it is recurrent, aperiodic and irreducible.

iv. Find the limiting distribution of the transition matrix

solution

```
steadyStates(markovB)
```

```
##           Up      Down   Stable
## [1,] 0.3092784 0.4536082 0.2371134
```

h) A telephone attendant receives 110 calls during the busy hour. Each call takes, on average, 2.1 minutes to process.

a) What percentage of the attendant's time is devoted to answering calls?

$$\text{call time} = \rho = \frac{\lambda}{\mu} \lambda = \frac{1}{110} \mu = \frac{2.1}{60} \rho = \frac{1}{110} \times \frac{2.1}{60} = 0.25$$

b) How long must people wait, on average, before their call is processed?

$$\text{waiting time in que} = W_q = \frac{L_q}{\lambda} = \frac{\lambda}{\mu(\mu - \lambda)} W_q = \frac{\rho}{\mu - \lambda} = \frac{0.25}{\frac{2.1}{60} - \frac{1}{110}} = 9.65$$

i) Jobs arrive to a computer system (consisting of a CPU and an I/O device) according to a Poisson process with rate 8 jobs per minute. Once in the system, a job requires on average 30 seconds of CPU time and 9 minutes of I/O time, in which the CPU and I/O time required by the jobs are exponentially distributed.

i. What is the probability that a job will have to wait before being processed by the devices? (Hint: replace the CPU and I/O subsystem as equivalent to single server)

ii. What proportion of time is the system busy?

iii. On average, how many jobs are waiting in line to be processed?

iv. On average, how long will a job spend in the system?

v. What is the probability that exactly 10 jobs arrive to the system in one minute?