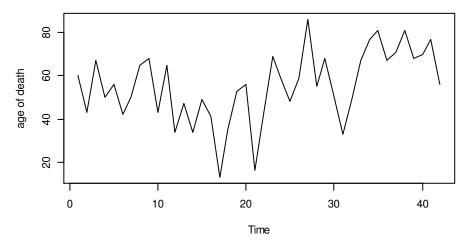
## First load forecast package

```
x=c(60,43,67,50,56,42,50,65,68,43,65,34,47,34,49,41,13,35,
+ 53,56,16,43,69,59,48,59,86,55,68,51,33,49,67,77,81,67,71,81,
+ 68,70,77,56)

x( age of death of 42 successive kings in England)
60 43 67 50 56 42 50 65 68 43 65 34 47 34 49 41 13 35 53 56 16 43 69 59 48
59 86 55 68 51 33 49 67 77 81 67 71 81 68 70 77 56
```

y=ts(x)# create a time series object

## Age of death of 42 successive kings in England & Wales



■ From the diagram, we can see that the time series is not stationary in mean.

#### Let us now look at the acf and pacf values and the correlograms

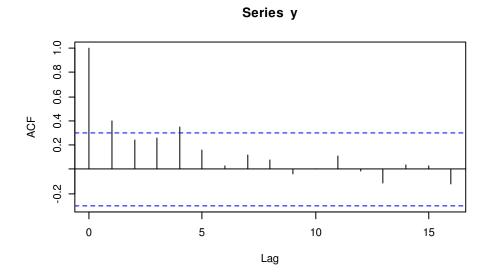
acf(x,lag.max=15,plot=F)#give the autocorrelation values

Autocorrelations of series 'y', by lag

0 1 2 3 4 5 6 7 8 9 10 1.000 0.401 0.238 0.260 0.348 0.161 0.031 0.115 0.078 -0.036 -0.001 11 12 13 14 15 0.111 -0.010 -0.117 0.035 0.025 pacf(x,lag.max=15,plot=F) # gives the partial autocorrelation values Partial autocorrelations of series 'y', by lag

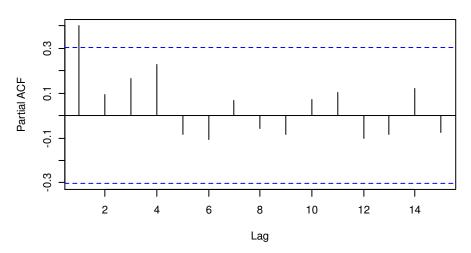
- The top rows are lags and the bottom rows are correlation values.
- We can see from the observed values of the correlations(acf/pacf) that the series is not stationary

acf(y,lag.max=15)# gives the correlogram



pacf(x,lag.max=15,plot=F)# gives the partial autocorrelation values

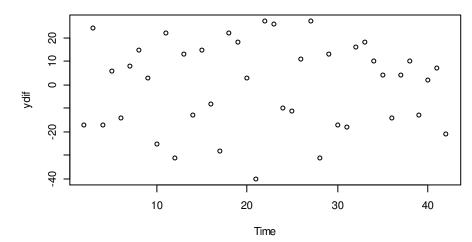




## Take the first difference

```
ydif=diff(y,differences=1)
plot(ydif,type="p",main="First differences of age of death")
```

## First differences of age of death



- The series of the first differences appears to be stationary in mean and variance.
- The first differences have removed the trend component in the series.

• We can now examine whether there are correlations between successive terms of the irregular component; if so, would aid us to make predictive model for the ages at death of the kings.

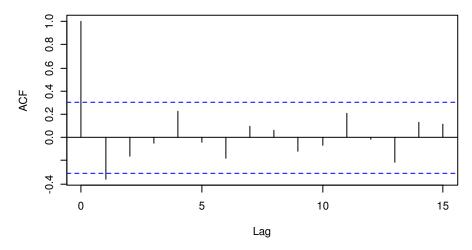
acf(ydif,lag.max=15,plot=F)

Autocorrelations of series 'ydif', by lag

• From the first differences, correlation values after lag 1 is relatively small as lag increases.

acf(ydif,lag.max=15)

## Series ydif



- From the correlogram the autocorrelation at lag 1 (-0.360) exceeds the significance bounds, whereas all others are within. This suggests that the series is stationary.
- To do a comparison of the acfs of the original series and the first differences. Then we should ensure that the both y-axes and x-axes of both series are similar.

To do this:- acf(y,lag.max=15,ylim=c(-0.4,0.4))#ylim adjust the y-axes, while xlim acf(ydif,lag.max=15,ylim=c(-0.4,0.4))# will adjust the x-axes

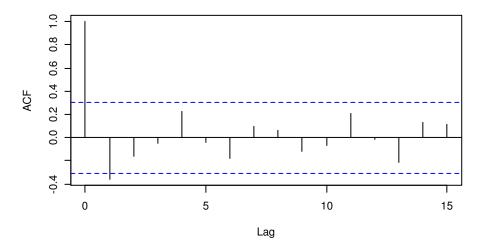
## Selecting candidate ARIMA model

- If your time series is stationary, or you have transformed it to be a stationary time series.
- ullet The next step is to select the appropriate ARIMA model, which means finding the most appropriate values of p and q for an ARIMA (p,d,q) model.
- To do this, we examine the acf and pacf of the stationary time series.

```
acf(ydif,lag.max=15,plot=F)
acf(ydif,lag.max=15)
```

Autocorrelations of series 'ydif', by lag

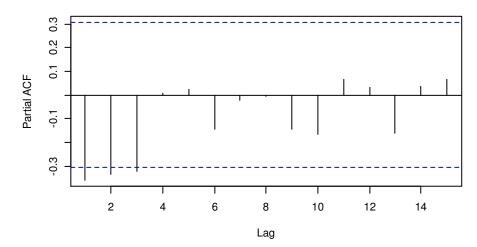
#### Series ydif



```
pacf(ydif,lag.max=15,plot=F)
pacf(ydif,lag.max=15)
```

Partial autocorrelations of series 'ydif', by lag

## Series ydif



- Notice that pacf of the first differences, shows that the first three lags exceed the significance bounds and are negative, and are slowly decreasing in magnitude as the lag increases. Thus, the pacf tails off to zero after lag 3
- This is evident from the partial correlogram

#### How do we choose a model to fit our data:-

(i) We will use the acf and pacf of our stationary series

Since the correlogram is zero after lag 1, and the partial correlogram tails off to zero after lag 3. This means that the following ARMA models are possible for the data:-

- ARMA(3,0) that is, AR(3) since the pacf tails to zero after lag 3, and the acf tails to zero
- ARMA(0,1) that is MA(1) since the acf is zero after lag 1 and the pacf tails off to zero
- (ii) We use the principle of parsimony to decide which model is best; that is, we assume that the model with the fewest parameters is best.
- (iii) In addition, our selection criteria; i.e, the lowest AICc and/or BICc values
- (iv) Using the principle of parisomony, then ARMA(0,1) will be selected since it has the 1 parameter compared to ARMA(3,0) which has 3 parameters. Or one could choose from a family of ARIMA

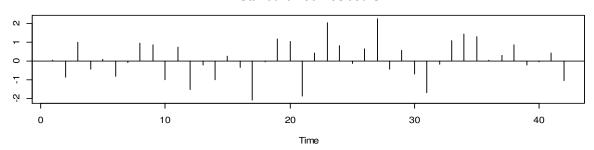
	General Behaviour o	of the ACF and PACF for	ARMA Models
	AR(p)	MA (q)	ARMA(p,q), p>0 & q>0
ACF	Tails off	Cuts off after lag q	Tails off
PACF	Cuts off after lag p	Tails off	Tails off

Now ARIMA(p,d,q):- autoregressive integrated moving average where 'p' is for AR, 'q' for MA and 'd' for the differencing

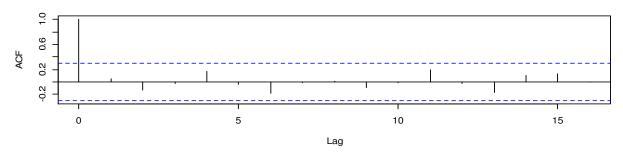
To fit both models in R, we use the original series of the ages of death and model the time series as an ARIMA (3,1,0) or ARIMA(0,1,1).

```
mod=Arima(y, order=c(0, 1, 1))
summary(mod)
Series: y
ARIMA(0,1,1)
Coefficients:
         ma1
     -0.7218
s.e. 0.1208
sigma^2 estimated as 230.4: log likelihood=-170.06
AIC=344.13 AICc=344.44 BIC=347.56
Training set error measures:
ME RMSE MAE MPE MAPE MASE
Training set 0.9712931 14.99836 11.92162 -10.40664 29.5176 0.9047383
Alternative way:-
m1=Arima(ydif,order=c(0,0,1),include.mean=0)
summary(m1)
Series: ydif
ARIMA(0,0,1) with zero mean
Coefficients:
         ma1
     -0.7218
s.e. 0.1208
sigma^2 estimated as 230.4: log likelihood=-170.06
AIC=344.13 AICc=344.44 BIC=347.56
Training set error measures:
                       RMSE MAE MPE MAPE MASE
                  ME
Training set 0.9935174 15.18016 12.21093 89.97649 100.7174 0.4953723 0.05200044
```

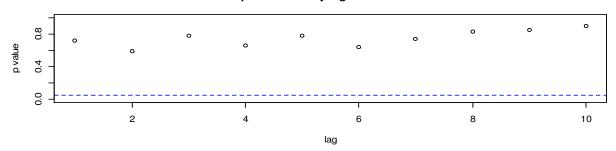
#### Standardized Residuals



#### **ACF of Residuals**



## p values for Ljung-Box statistic



- Inspection of the time plot of the standardized residuals shows no obvious patterns.
- $\bullet$  The ACF of the standardized residuals shows no apparent departure from the model assumptions.
- The Q-statistic is insignificant at the lags shown.

```
Acf(residuals(mod))
```

Box.test(mod\$residuals,lag=20,type="Ljung-Box")

Box-Ljung test

data: mod\$residuals

X-squared = 13.5844, df = 20, p-value = 0.8509

## ■ The test suggests that the model is specified correctly

mod.forecast=forecast.Arima(mod,h=5)

mod.forecast

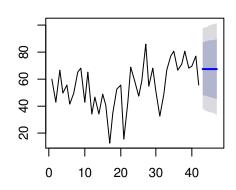
	Point	Forecast	Lo 80	Hi 80	Lo 95	Hi 95
43		67.75063	48.29647	87.20479	37.99806	97.50319
44		67.75063	47.55748	87.94377	36.86788	98.63338
45		67.75063	46.84460	88.65665	35.77762	99.72363
46		67.75063	46.15524	89.34601	34.72333	100.77792
47		67.75063	45.48722	90.01404	33.70168	101.79958

plot(mod.forecast)
acf(mod.forecast\$residuals,lag.max=15)
Box.test(mod.forecast\$residuals,lag=20,type="Ljung-Box")# same results as above

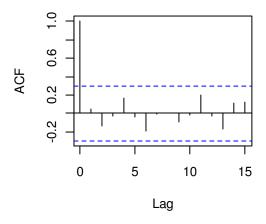
## Series: residuals(mod)

# 

## Forecasts from ARIMA(0,1,1)



## Series mod.forecast\$residuals



```
mod1=Arima(x, order=c(3, 1, 0))
summary(mod1)
Series: x
ARIMA(3,1,0)
Coefficients:
              ar2 ar3
         ar1
     -0.6063 -0.4904 -0.3284
s.e. 0.1489 0.1551 0.1477
sigma^2 estimated as 222: log likelihood=-169.3
AIC=346.59 AICc=347.7 BIC=353.45
Training set error measures:
                  ME RMSE
                                MAE
                                        MPE MAPE
                                                         MASE
Training set 0.4680342 14.7227 11.90333 -9.9101 29.0211 0.9033502
mod2=Arima(x, order=c(3, 1, 1))
summary(mod2)
Series: x
ARIMA(3,1,1)
Coefficients:
         ar1
                ar2 ar3
                                 ma1
     -0.5805 \quad -0.4778 \quad -0.3196 \quad -0.0293
s.e. 0.4646 0.2669 0.2140 0.4985
sigma^2 estimated as 222: log likelihood=-169.29
AIC=348.59 AICc=350.3 BIC=357.16
Training set error measures:
                  ME
                         RMSE
                                MAE MPE MAPE
Training set 0.4792116 14.72198 11.90275 -9.899733 29.00887 0.7484857
                  ACF1
Training set -0.01216043
```