TA section 10

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Homework 5



§8.2 #5

- 5. 設 X_1, \dots, X_{36} 為一組由 $\mathcal{N}(\theta, 36)$ 分佈所產生之隨機樣本。欲檢定 $H_0: \theta = 10$, vs. $H_a: \theta = 12$ 。拒絕城取為 $\{\overline{X} > 11\}$ 。試決定型 I 及型 I 錯誤之機率 α, β 。(解. $\alpha = \beta \doteq 0.1587$)
- $\bar{X} \sim \mathcal{N}(\theta, 1)$.
- $\alpha = \mathbb{P}(\bar{X} > 11 | \theta = 10) = \mathbb{P}(Z > 1) = 1 \Phi(1) \approx 0.1587.$
- $\beta = \mathbb{P}(\bar{X} \le 11 | \theta = 12) = \mathbb{P}(Z \le -1) = \Phi(-1) = 1 \Phi(1) \approx 0.1587.$



§8.2 #16

- 16. 設X有 $\mathcal{B}(10,\theta)$ 分佈, $\theta\in\Omega=\{1/4,1/2\}$ 。基於X、微檢定 $H_0:\theta=1/2$ 、vs. $H_a:\theta=1/4$ 。拒絕域取爲 $\{X\leq 3\}$ 。試求此檢定之檢力函數。(解. $K(1/4)=11/64,\,K(1/2)=31\cdot 3^8/4^9)$
- $K(\theta) = \mathbb{P}(\text{reject } H_o \mid \theta) = \mathbb{P}(X \leq 3 \mid \theta \in \{1/4, 1/2\}).$
- $K(1/4) = \sum_{k=0}^{3} {10 \choose k} (1/4)^k (3/4)^{10-k} \approx 31.3^8/4^9$.
- $K(1/2) = \sum_{k=0}^{3} {10 \choose k} (1/2)^{10} = 11/64.$



§8.2 #20

20. 設 X_1, \dots, X_n 爲一組由 $U(0, \theta)$ 分佈所產生之隨機樣本。欲檢定 H_0 :

$$\theta \leq \theta_0, \, \mathrm{vs.} \,\, H_a : \theta > \theta_0$$
。取拒絕域爲 $\{X_{(n)} \geq c\}$ 。

- (i) 試求檢力函數 $K(\theta)$, 並證明此爲 θ 之一增函數;
- (ii) 設 $\theta_0 = 1/2$, 試求c之值, 使顯著水準爲0.05;
- $f_{X_{(n)}}(t) = nt^{n-1}/\theta^n$
- $K(\theta) = \mathbb{P}(X_{(n)} > c|\theta) = 1 \theta^{-n} \int_0^c nx^{n-1} dx = 1 (c/\theta)^n$.
- $dK(\theta)/d\theta = nc^n/\theta^{n+1} > 0$ for all $0 < \theta_0 < c < \theta$.
- $\lim_{n\to\infty} K(\theta) = 1$ as $c < \theta$.
- $K(\theta_0 = 1/2) = 1 (2c)^n = 0.05 \Rightarrow c^* = \frac{(0.95)^{1/n}}{2}$.



- (iii) 試決定n要多大, 使得對(ii)中之檢定, K(0.75) = 0.98;
- (iv) 設n = 20, 且 $X_{(n)} = 0.48$, 試求此時之p-値。

•
$$K(\theta = 0.75) = 1 - \frac{(0.95)/2^n}{(0.75)^n} = 0.98 \Rightarrow n \approx 10.$$

• p-value:
$$p(\boldsymbol{x}) = \mathbb{P}(T(\boldsymbol{X}) > T(\boldsymbol{x}) \mid H_o) = \mathbb{P}(X_{(n)} > 0.48 \mid H_o: \theta = \theta_0) = 1 - (0.48/0.5)^{20} \approx 0.558.$$



§8.3 #7

- 7. 設 X_1, \dots, X_{30} 為一組由 $N(\mu, \sigma^2)$ 分佈所產生之隨機樣本, μ 設為未 知。做檢定 $H_0: \sigma^2 = 8$, vs. $H_a: \sigma^2 < 8$ 。假設樣本變異數 $S^2 =$ 6.4。試在 $\alpha = 0.05$ 之下, 作一檢定。(解. K = 23, 不能接受 H_0)
- Let a test statistic $K(\boldsymbol{X}) := (n-1)S^2/\sigma^2 \sim \chi^2_{n-1} = \chi^2_{29}.$
- Under H_o , a testing rule:

$$\phi(\mathbf{X}) = \mathbf{I}(K(\mathbf{X}) < \chi^2_{0.05, 29}),$$

such that $0.05 = \sup_{\sigma^2 = 8} \mathbb{P}(K(X) < \chi^2_{0.05,29} \mid H_o).$

- $K = (29) \times (6.4)/(8) \approx 23.2 > \chi^2_{0.05,29} = 17.71 \Rightarrow$ do not reject H_o under the level $\alpha = 0.05$.
- level: 事前設定的允許型一誤差之最大機率上限; size: 在所有原假設範圍內實際 達到的可能出現最大型一誤差機率, 通常等於或小於 level。若為簡單假設, size = level.

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§8.4 #6. #8

- 6. 設 X_1, \dots, X_n 爲一組由 $U[0, \theta]$ 分佈所產生之隨機樣本。欲檢定 H_0 : $\theta = \theta_0$, vs. $H_a: \theta \neq \theta_0$ 。試給一 α 下之LRT。(解. 拒絕域爲 $\{X_{(n)} < \theta_0\}$ $\theta_0 \alpha^{1/n}$
- likelihood $L(\theta) = \theta^{-n} I(0 < x_i < \theta), \forall i = 1, 2, \dots, n. \ \Omega := \{\theta : \theta > 0\},$ $\widehat{\theta}_{ML} = X_{(n)} \Rightarrow L(\widehat{\theta}) = (1/X_{(n)})^n.$
- Under $\Omega_0 = \{\theta : \theta = \theta_0\}, \ \ddot{\theta}_{ML} = \theta_0 \Rightarrow L(\ddot{\theta}) = (1/\theta_0)^n$.
- LRT statistic: $\lambda(X) = L(\ddot{\theta})/L(\hat{\theta}) = (X_{(n)}/\theta_0)^n \in (0,1]$, for $X_{(n)} \leq \theta_0$,
- a testing rule: $\phi(X) = I(\lambda(X) < c) \iff$

$$\phi(\mathbf{X}) = \mathbf{I}(X_{(n)} \le c' = c^{1/n}\theta_0),$$

such that $c \in (0,1)$ satisfying $\alpha = \sup_{\theta = \theta_0} \mathbb{P}(\lambda(X) \le c | H_o : \theta = \theta_0)$.

- $\alpha = \theta_0^{-n} \int_0^{c^{1/n}\theta_0} nx^{n-1} dx \Rightarrow c = \alpha \text{ and } c' = \alpha^{1/n}\theta_0.$
- So, the rejection region $C = \{X_{(n)} \leq \alpha^{1/n}\theta_0\}$.



JERRY C. Mathematical Statistics II 8. 設 X_1, \dots, X_n 爲一組由 $Be(\theta, 1)$ 分佈所產生之隨機樣本。欲檢定 H_0 : $\theta = \theta_0$, vs. $H_a: \theta \neq \theta_0$ a a sign of a constant $\theta = \theta_0$ vs. $\theta = \theta_0$ and $\theta = \theta_0$ vs. $\theta = \theta_0$ vs. $\theta = \theta_0$ and $\theta = \theta_0$ vs. $\theta = \theta_0$ vs. $\theta = \theta_0$ and $\theta = \theta_0$ vs. $\theta =$ 理4.2. 得到近似的拒絕域。

- Under $H_a: \theta \in \Omega$, $\widehat{\theta} = -n/\sum_{i=1}^n \log X_i$; under $H_o: \theta \in \Omega_o$, $\widehat{\theta} = \theta_0$.
- (i) Let $T(\mathbf{X}) = -\sum_{i=1}^{n} \log X_i$, a S.S. of θ ,
 - LRT statistic:

$$\lambda(T(\boldsymbol{x})) := L(\ddot{\theta})/L(\widehat{\theta}) = (\theta_0 T/n)^n \exp(n - \theta_0 T)$$

• $\log \lambda(T(x)) = n \log \theta + n \log T - n \log n + n - \theta_0 T$.

$$d\log \lambda(T(\boldsymbol{x}))/dT = n/T - \theta_0 \begin{cases} >0, & [T < n/\theta_0] \\ =0, & [T = n/\theta_0] \\ <0, & [T > n/\theta_0], \end{cases}$$

so, a testing rule: $\phi(X) = I(\lambda(X) < c) \iff$

$$\phi(\mathbf{X}) = \mathbf{I}(\{T < c_1\} \cup \{T > c_2\}),$$

such that $c_1, c_2 \in (0,1)$ with

$$\alpha = \sup_{\theta = \theta_0} \mathbb{P}(\{T < c_1\} \cup \{T > c_2\} | H_o : \theta = \theta_0).$$

 $\begin{array}{l} \bullet \, \because 2\theta_0 T \sim \Gamma(n,2) = \chi^2_{(2n)} \,\, \text{under} \,\, H_o, \\ \text{choose} \,\, \alpha/2 = \mathbb{P}(2\theta_0 T < c_1' = \chi^2_{\alpha/2,2n}) \,\, \text{and} \\ \alpha/2 = \mathbb{P}(2\theta_0 T > c_2' = \chi^2_{1-\alpha/2,2n}). \,\, \text{Thus, a testing rule:} \end{array}$

$$\phi(\mathbf{X}) = \mathbf{I}(\{T < \chi^2_{\alpha/2,2n}/(2\theta_0)\} \cup \{T > \chi^2_{1-\alpha/2,2n}/(2\theta_0)\})$$

is a size α LRT for $H_o: \theta = \theta_0$.

(ii) Since that $-2 \log \lambda \xrightarrow{\mathsf{d}} \chi_1^2$, as $n \to \infty$, the rejection of region of an approximate size α LRT can be:

$$C^* = \{-2\log \lambda > \chi^2_{1-\alpha/2,1} \vee -2\log \lambda < \chi^2_{\alpha/2,1}\}.$$