TA section 2

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HW 1: Part (I)



- 1. 設 X_n 有 $P(n\lambda)$ 分佈, $n \ge 1$, $\lambda > 0$ 。試證 $X_n/n \xrightarrow{p} \lambda$, 且 $(X_n-n\lambda)/(\sqrt{n\lambda}) \xrightarrow{d} \mathcal{N}(0,1)$ 。
- $\mathbb{E}[X_n] = \text{Var}[X_n] = n\lambda$; $\mathbb{E}[X_n/n] = \lambda$, $\text{Var}[X_n/n] = \lambda/n$, and by Chebyshev's inequality,

$$\mathbb{P}(|X_n/n - \mathbb{E}[X_n/n]| > \epsilon) \le \text{Var}[X_n/n]/\epsilon^2, \ \forall \epsilon > 0.$$

Then,

$$\mathbb{P}(|X_n/n - \lambda| > \epsilon) \le \lambda/(n\epsilon^2) \to 0$$

as $n \to \infty$. So, $X_n/n \xrightarrow{\mathsf{p}} \lambda$.

• 另解: $X_n := \xi_1 + \dots + \xi_n \ \xi_i \sim i.i.d.P(\lambda)$ with $\mathbb{E}[\xi_i] = \lambda$. By LLN, $X_n/n \xrightarrow{\mathsf{p}} \mathbb{E}[\xi_i] = \lambda$. So, $X_n/n \xrightarrow{\mathsf{p}} \lambda$.



- Let $X_n := \xi_1 + \dots + \xi_n$, $\xi_i \sim i.i.d.P(\lambda)$ with $\mathbb{E}[\xi_i] = \operatorname{Var}[\xi_i] = \lambda$.
- By CLT,

$$(X_n - n\lambda)/\sqrt{n\lambda} = (X_n/n - \lambda)/\sqrt{\lambda/n} \xrightarrow{\mathsf{d}} N(0, 1)$$

as $n \to \infty$.



- 9. 設 X_n 有 $\Gamma(n,\lambda)$ 分佈, $n\geq 1$ 。試證 X_n/n $\xrightarrow[n\to\infty]{p}$ λ , 並利用中央極限定理, 估計n很大時, $P(X_n\leq x)$ 之値, $x\in R$ 。
- $\mathbb{E}[X_n] = n\lambda$, $\operatorname{Var}[X_n] = n\lambda^2$; $\mathbb{E}[X_n/n] = \lambda$, $\operatorname{Var}[X_n/n] = \lambda^2/n$, and by Chebyshev's inequality,

$$\mathbb{P}(|X_n/n - \mathbb{E}[X_n/n]| > \epsilon) \le \text{Var}[X_n/n]/\epsilon^2, \ \forall \epsilon > 0.$$

Then,

$$\mathbb{P}(|X_n/n - \lambda| > \epsilon) \le \lambda^2/(n\epsilon^2) \to 0$$

as $n \to \infty$. So, $X_n/n \xrightarrow{\mathsf{p}} \lambda$.

• 另解: $X_n := \xi_1 + \dots + \xi_n \ \xi_i \sim i.i.d.\Gamma(1,\lambda)$ with $\mathbb{E}[\xi_i] = \lambda$. By LLN, $X_n/n \xrightarrow{\mathsf{p}} \mathbb{E}[\xi_i] = \lambda$. So, $X_n/n \xrightarrow{\mathsf{p}} \lambda$.



JERRY C. Mathematical Statistics II

• $X_n := \xi_1 + \dots + \xi_n \ \xi_i \sim i.i.d.\Gamma(1,\lambda)$ with $\mathbb{E}[\xi_i] = \lambda$ and $\operatorname{Var}[\xi_i] = \lambda^2$.

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$$\frac{X_n/n - \lambda}{\lambda/\sqrt{n}} = \frac{(X_n - n\lambda)}{\lambda\sqrt{n}}.$$

By CLT,

$$\frac{X_n/n - \lambda}{\lambda/\sqrt{n}} \stackrel{\mathsf{d}}{\longrightarrow} N(0,1)$$

as $n \to \infty$.

• So, ${\rm I\!P}(X_n \le x) pprox \Phi(\frac{x-n\lambda}{\lambda\sqrt{n}})$, as $n \to \infty$.

14. 設
$$X_1 \sim \mathcal{U}[0,1]$$
, 令 $X_n = X_1^n$, $n \geq 1$ 。試證 $X_n \xrightarrow[n \to \infty]{p} 0$ 。

• Given any $\epsilon>0$, $\mathbb{P}(|X_n|>\epsilon)=\mathbb{P}(X_1>\epsilon^{1/n})=1-\epsilon^{1/n}.$

$$\lim_{n \to \infty} \mathbb{P}(|X_n| > \epsilon) = 1 - \lim_{n \to \infty} \epsilon^{1/n} = 0.$$

JERRY C.

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6. 設 X_1, X_2, X_3 爲獨立的隨機變數,且 X_i 有 $\mathcal{N}(i, i^2)$ 分佈,i = 1, 2, 3。試 利用 X_1, X_2, X_3 的函數,分別造出有如下的分佈。 $(i)\chi_3^2, (ii)\mathcal{T}_2, (iii)\mathcal{F}_{1,2}$ 。

(iii)

- Let $Z_i := (X_i i)/i \sim N(0, 1)$, for i = 1, 2, 3.
- : $Z_1^2 \sim \chi^2(1), Z_2^2 + Z_3^2 \sim \chi^2(2), : F = \frac{Z_1^2}{Z_2^2 + Z_3^2} \sim F(1, 2).$
- So, let $U:=Z_1^2=(X_1-1)^2/1, V:=Z_2^2:=(X_2-2)^2/2, W:=(X_3-3)^2/3,$ we have $\frac{U}{V+W}\sim F(1,2).$

10. 設
$$X_1$$
與 X_2 獨立,且皆有 $\mathcal{N}(0,25)$ 分佈。令 $D=\sqrt{X_1^2+X_2^2}$ 。試求 $P(D\leq 12.25)$ 。(解. 約0.95)

- $X_i=5Z_i, Z_i\sim N(0,1)$ for i=1,2. Then, $D=(X_1^2+X_2^2)^{1/2}=5(Z_1^2+Z_2^2)^{1/2}=5(\chi^2(2))^{1/2}$
- $\mathbb{P}(D \le 12.25) = \mathbb{P}(\chi^2(2) \le (12.25/5)^2) = \mathbb{P}(\chi^2(2) \le (2.45)^2) = 1 \exp(-(2.45)^2/2) \approx 0.95.$ (or 查卡方機率值表)
- Note 1: $W:=\chi^2(k)$, $f_W(w)=(1/2)\exp(-w/2)$. Then, $\mathbb{P}(W\leq x)=\int_{w=0}^x f_W(w)dw=1-\exp(-x/2)$.
- Note 2: by CLT,

$$\mathbb{P}(\chi^2(2) \le (2.45)^2) \approx \Phi(((2.45)^2 - 2)/\sqrt{4}) \approx 0.977.$$

不建議用 CLT 近似, 因為自由度只有 2.

