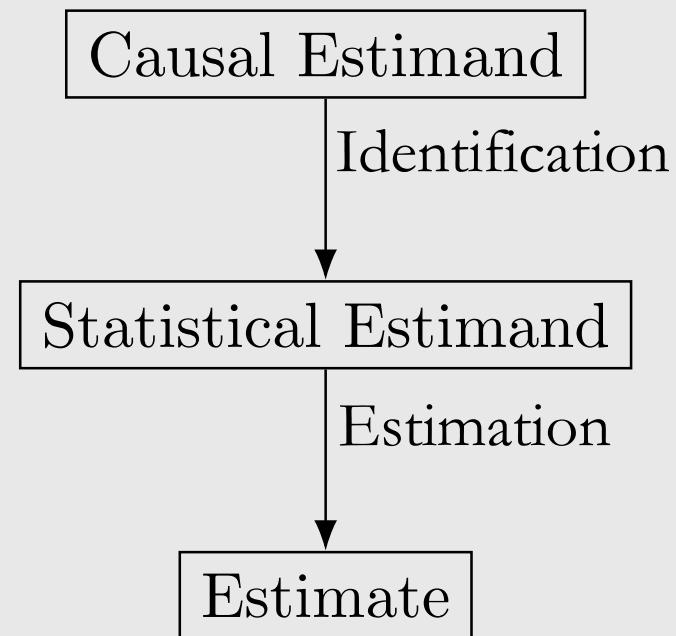


Causal Models

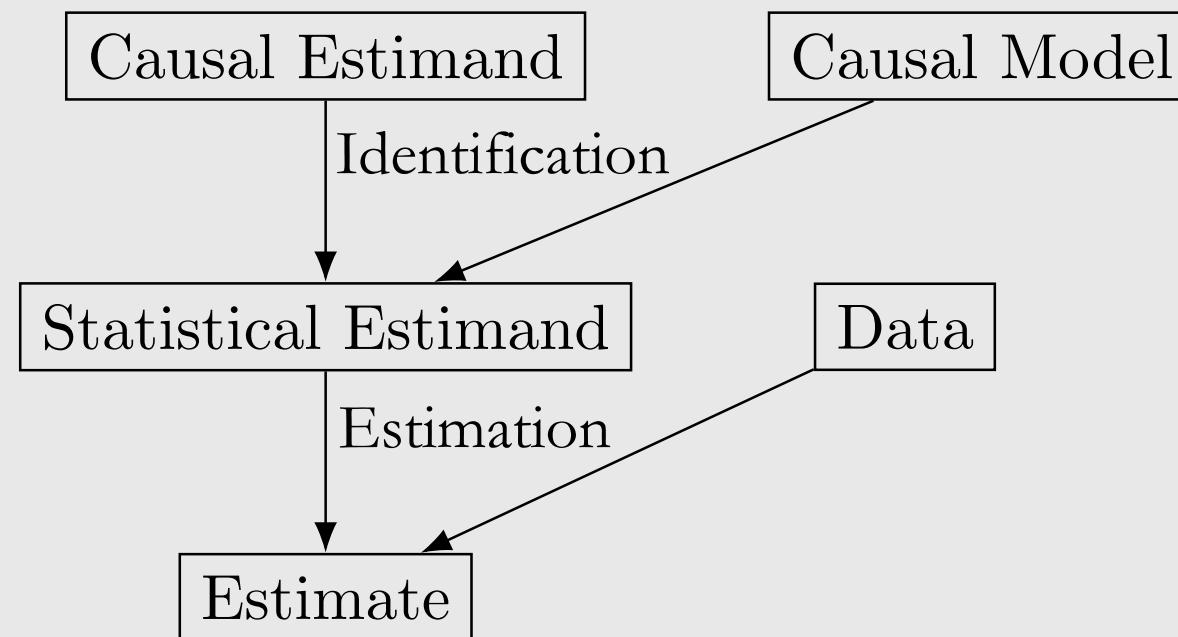
Brady Neal

causalcourse.com

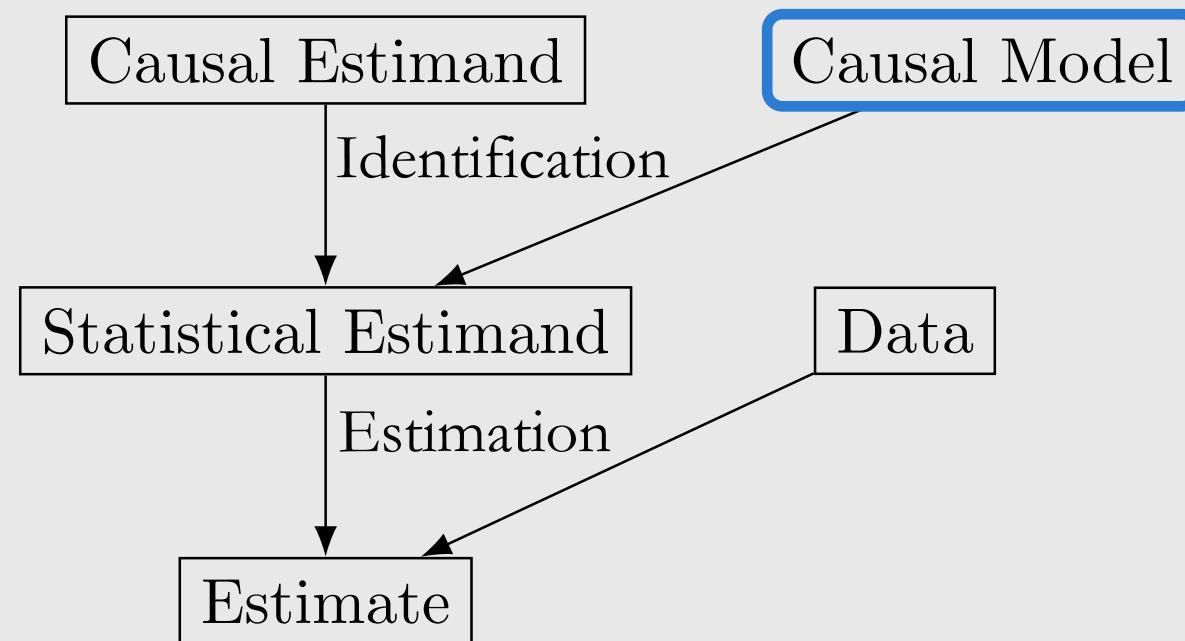
The Identification-Estimation Flowchart



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The Identification-Estimation Flowchart



The *do*-operator

Main assumption: modularity

Backdoor adjustment

Structural causal models

A complete example with estimation

The *do*-operator

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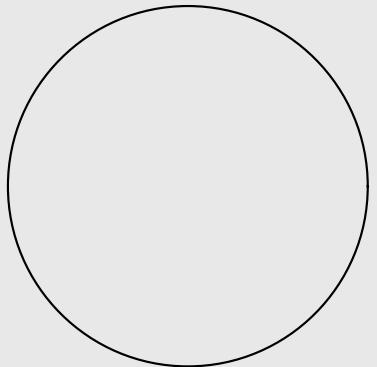
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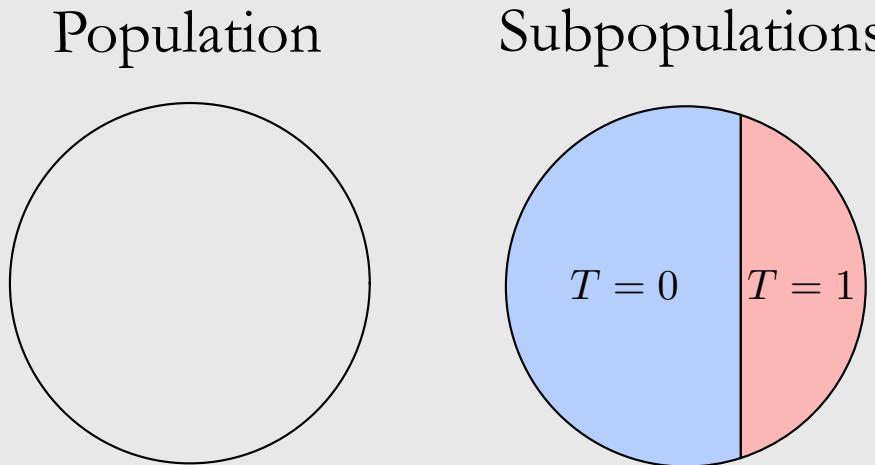
Conditioning vs. intervening

Conditioning vs. intervening

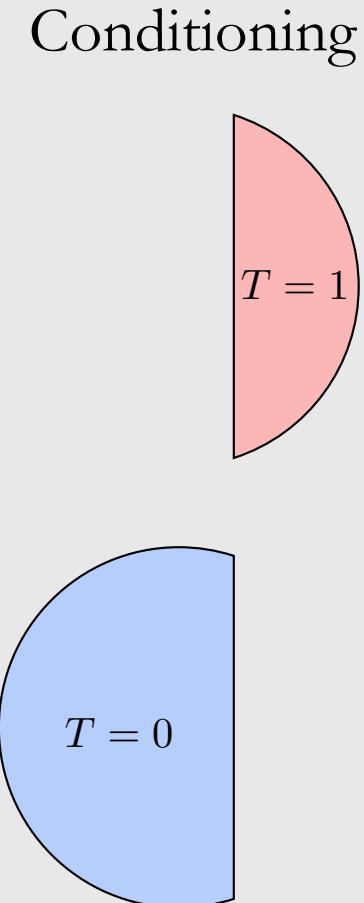
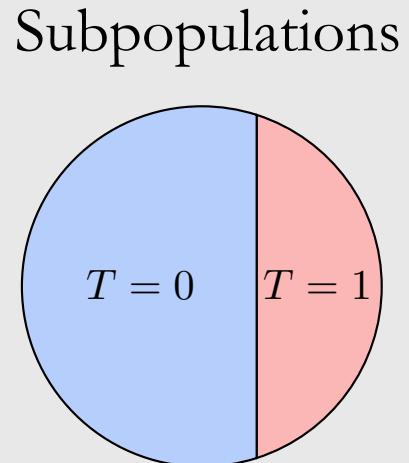
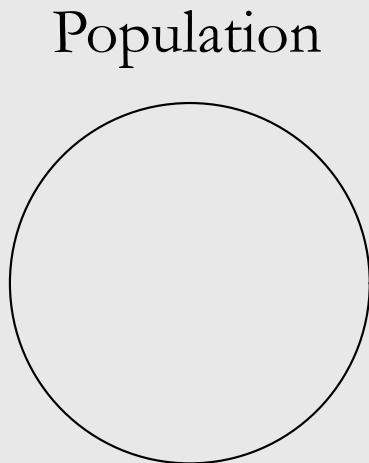
Population



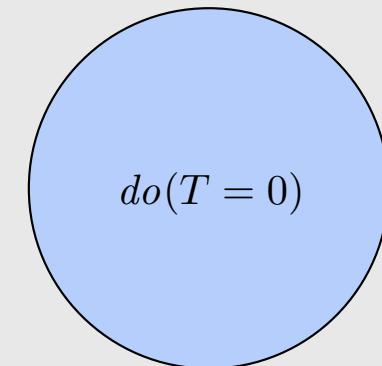
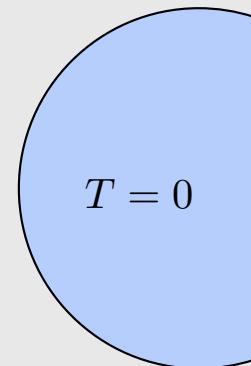
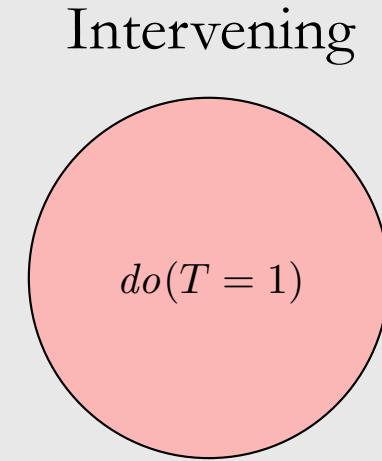
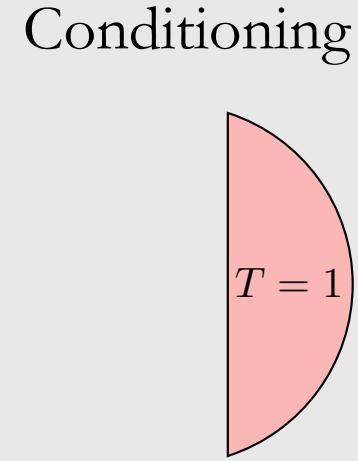
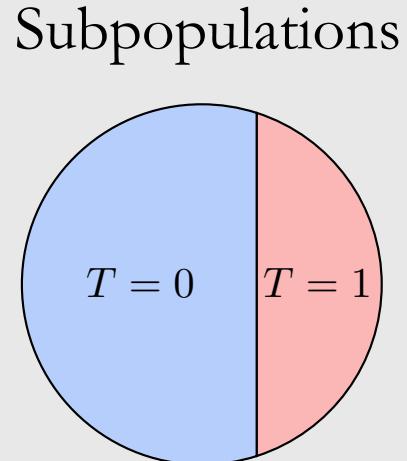
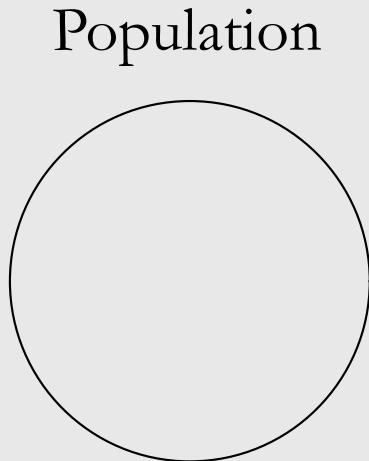
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Some notation and terminology

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Interventional distributions:

$$P(Y(t) = y)$$

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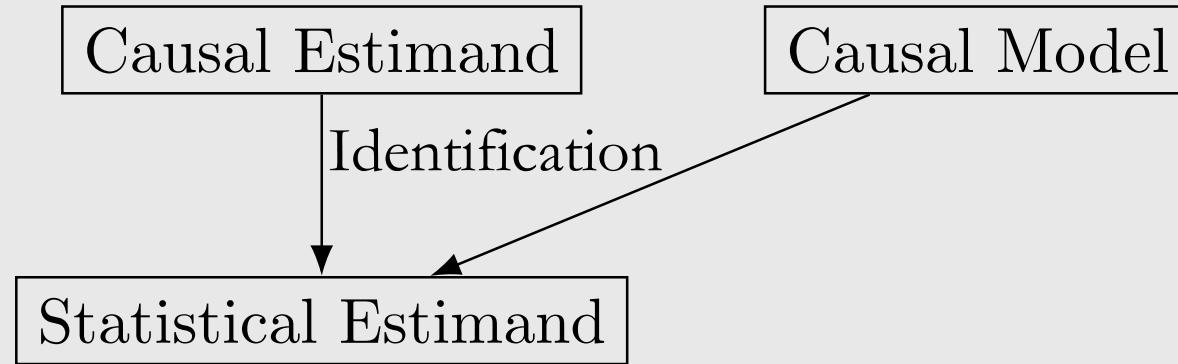
$$P(Y \mid T = t)$$

Interventional

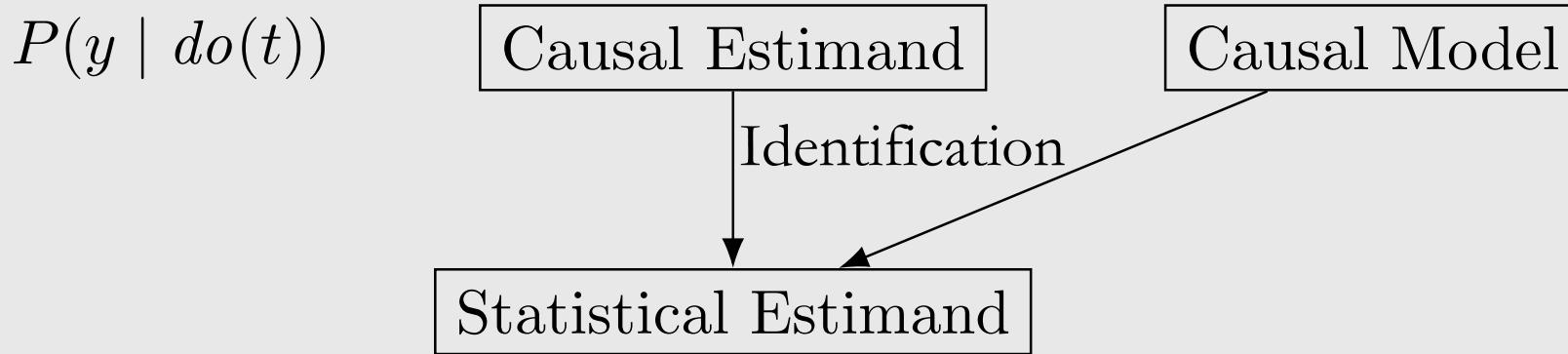
$$P(Y \mid do(T = t))$$

$$P(Y \mid do(T = t), X = x)$$

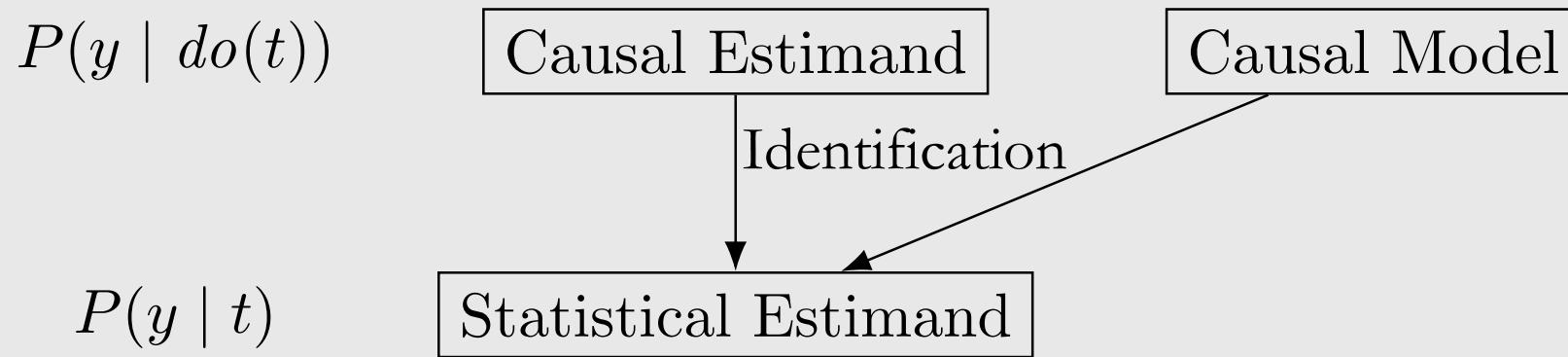
Identifiability



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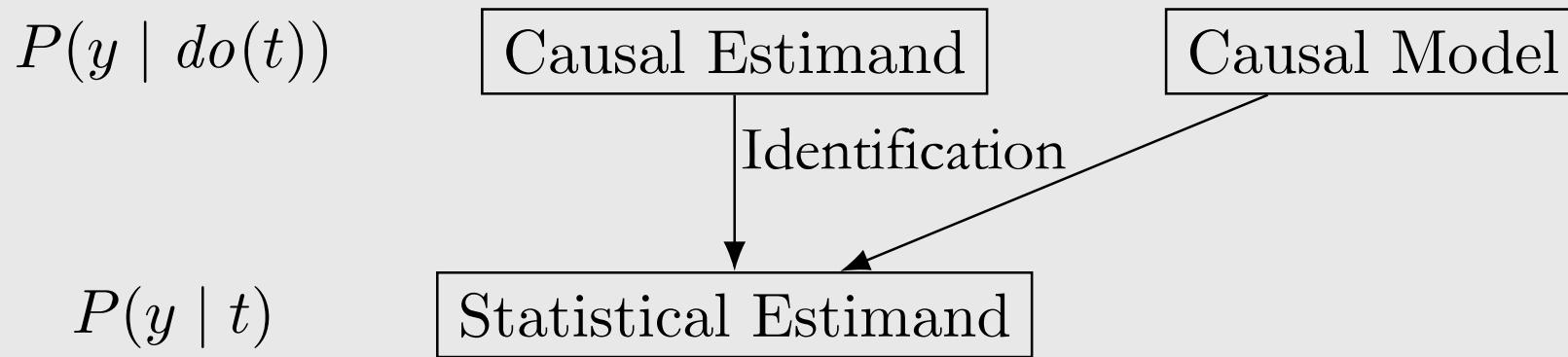


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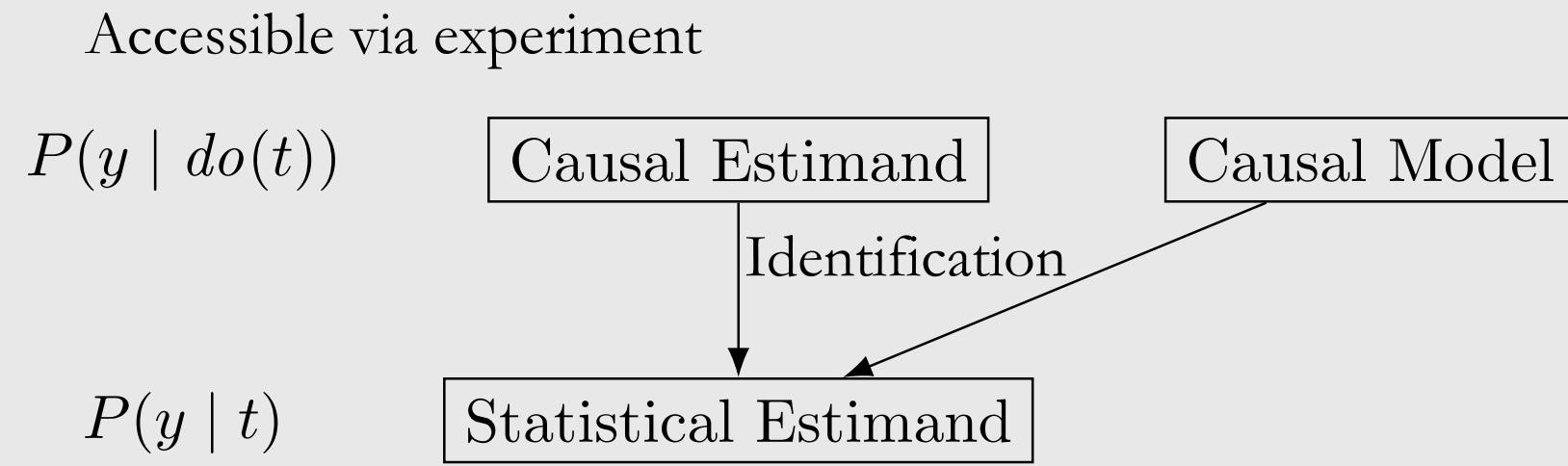


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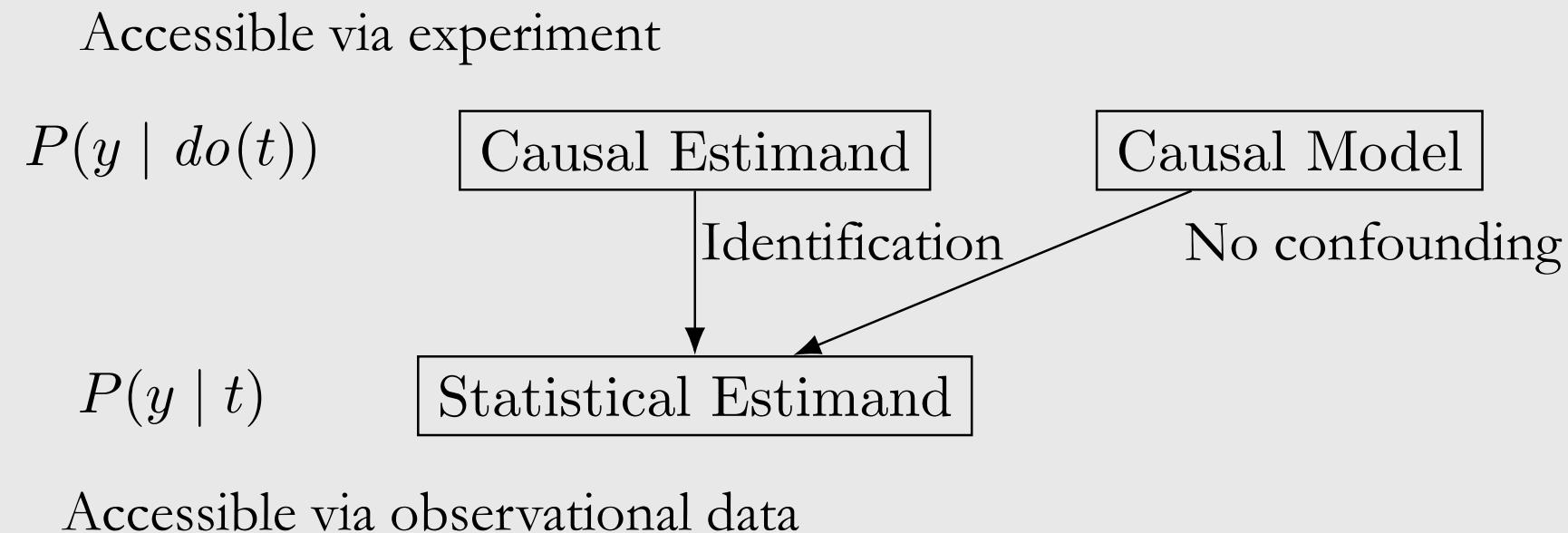
Accessible via experiment



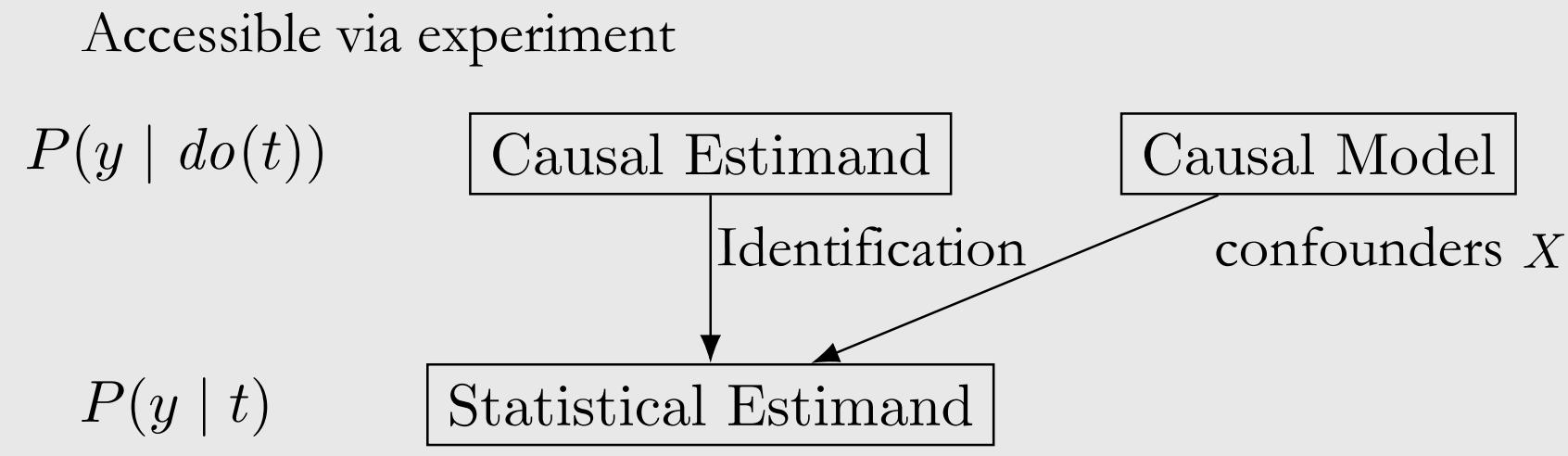
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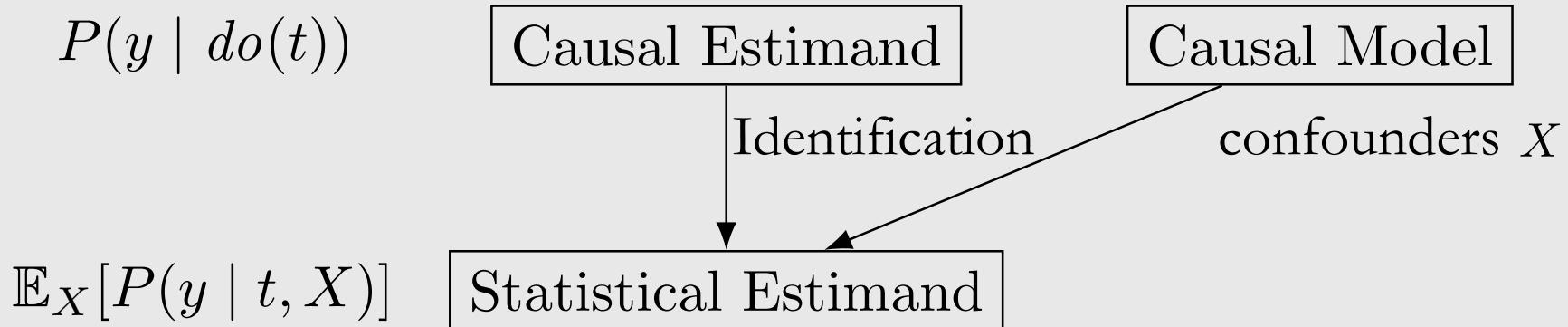
Identifiability



Accessible via observational data

Identifiability

Accessible via experiment



Accessible via observational data

The *do*-operator

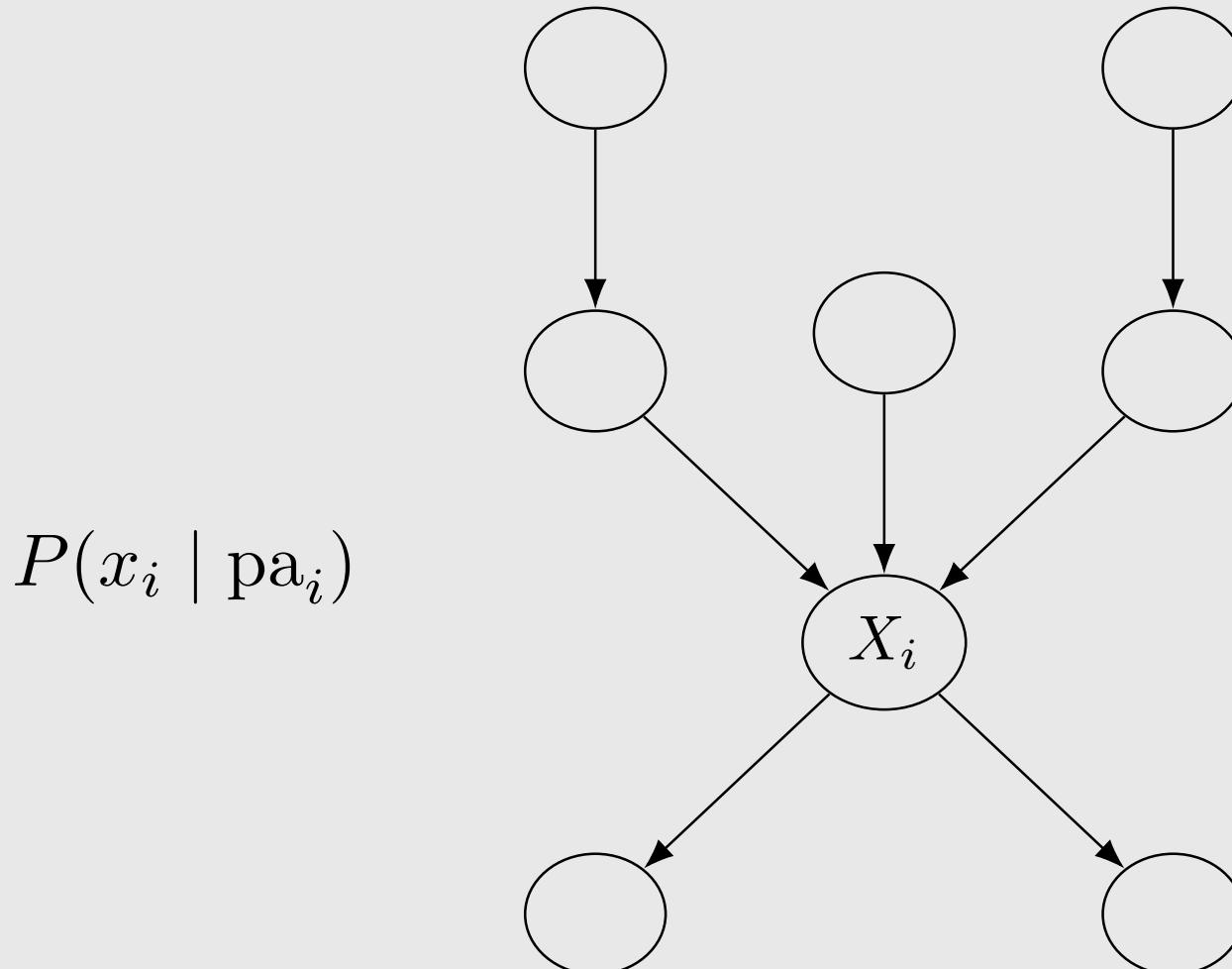
Main assumption: modularity

Backdoor adjustment

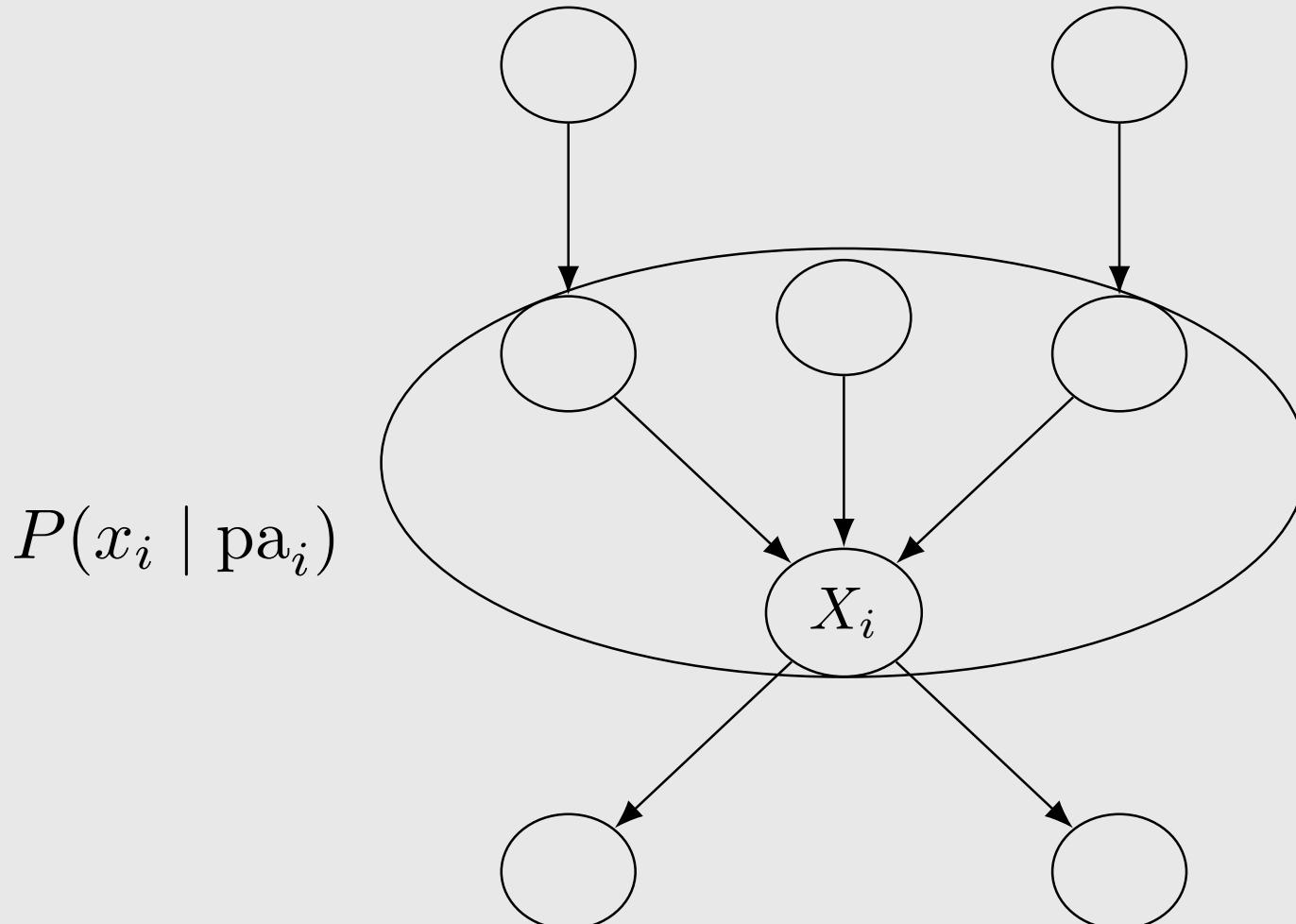
Structural causal models

A complete example with estimation

Causal mechanism



Causal mechanism



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If we intervene on a node X_i , then only the mechanism $P(x_i | \text{pa}_i)$ changes. All other mechanisms $P(x_j | \text{pa}_j)$ where $i \neq j$ remain unchanged.

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Many names: independent mechanisms, autonomy, invariance, etc.

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1. If $i \notin S$, then $P(x_i | \text{pa}_i)$ remains unchanged.
2. If $i \in S$, then $P(x_i | \text{pa}_i) = 1$ if x_i is the value that X_i was set to by the intervention; otherwise, $P(x_i | \text{pa}_i) = 0$.

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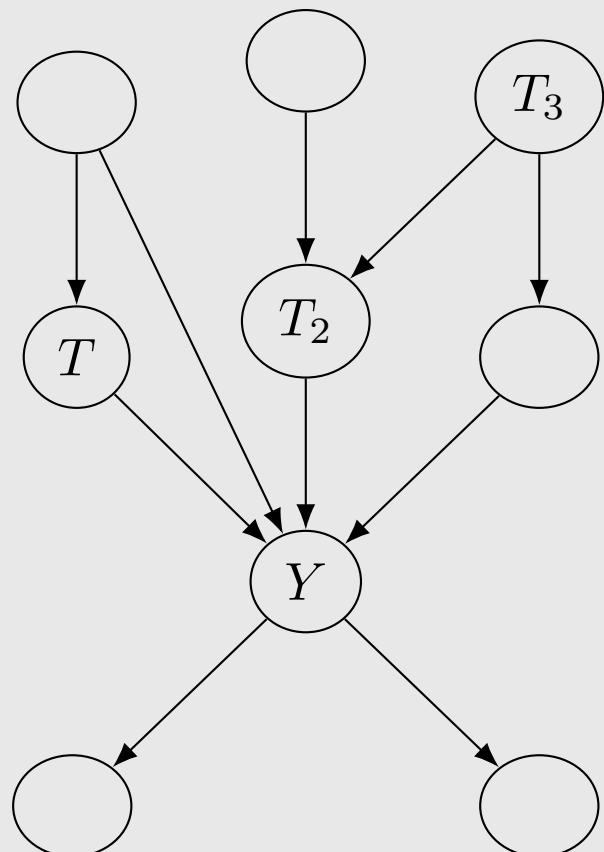
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consistent with
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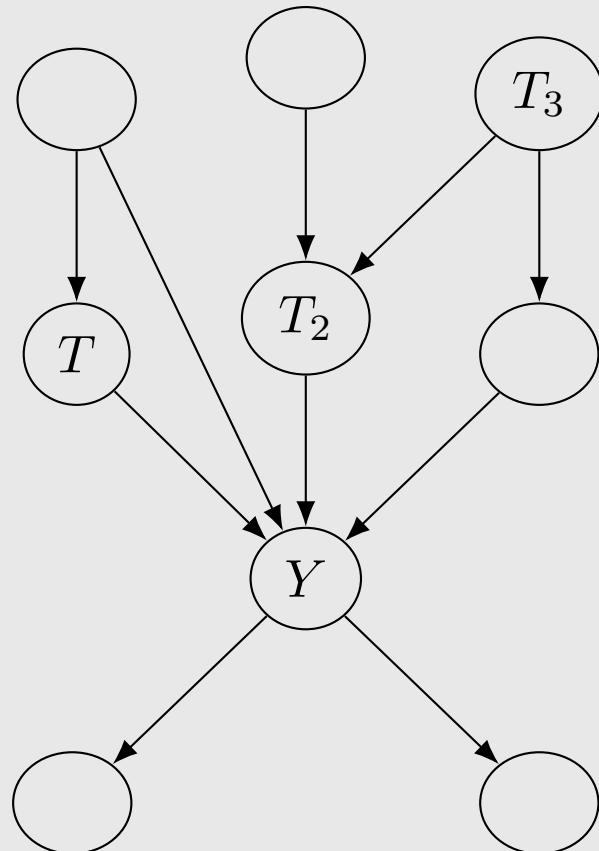
Manipulated graphs

Observational data

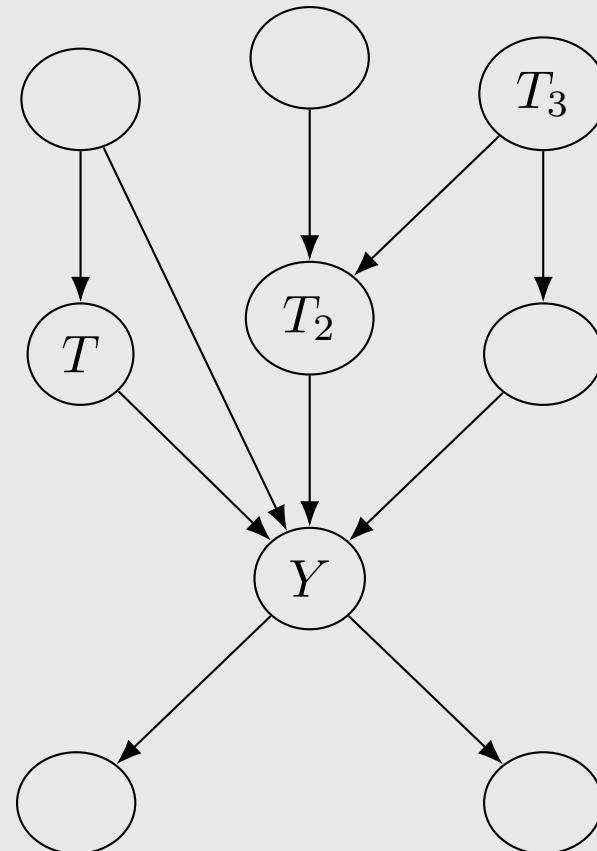


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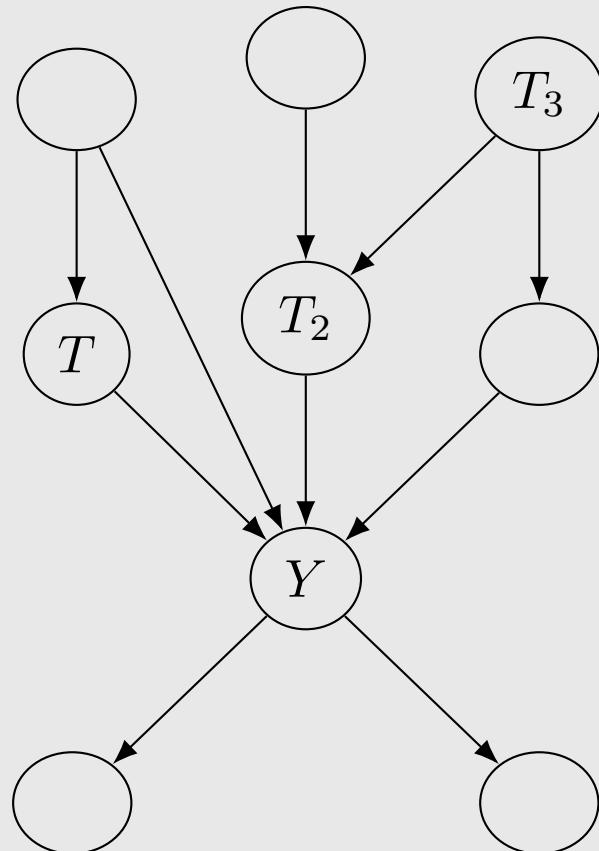


Interventional data

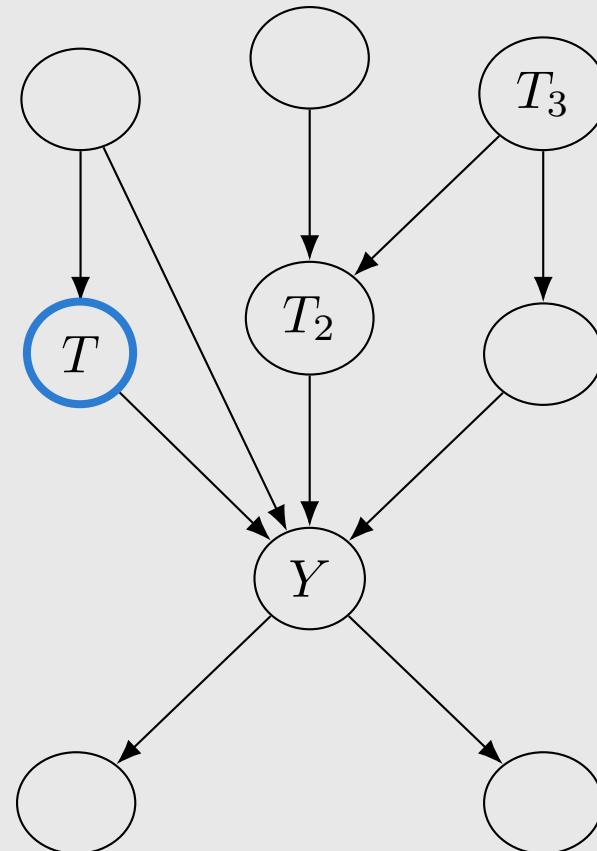


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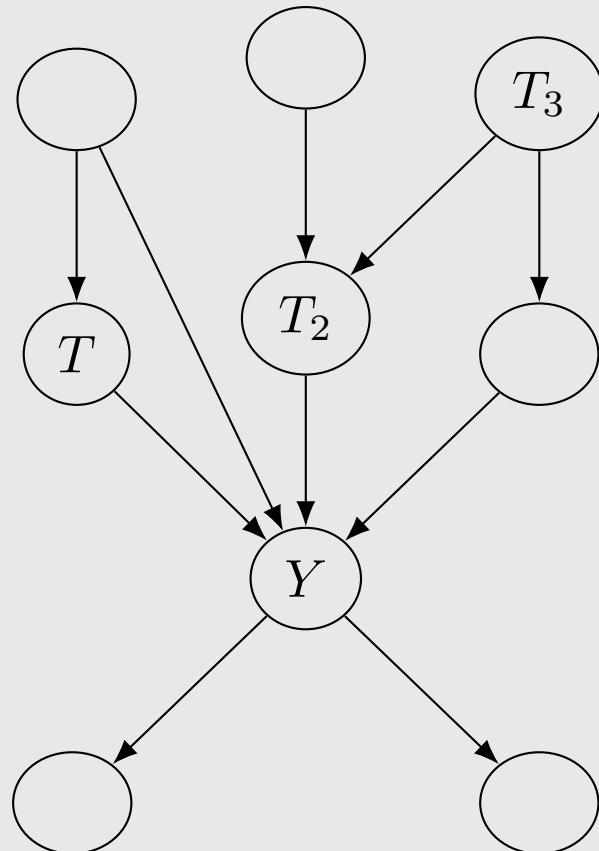


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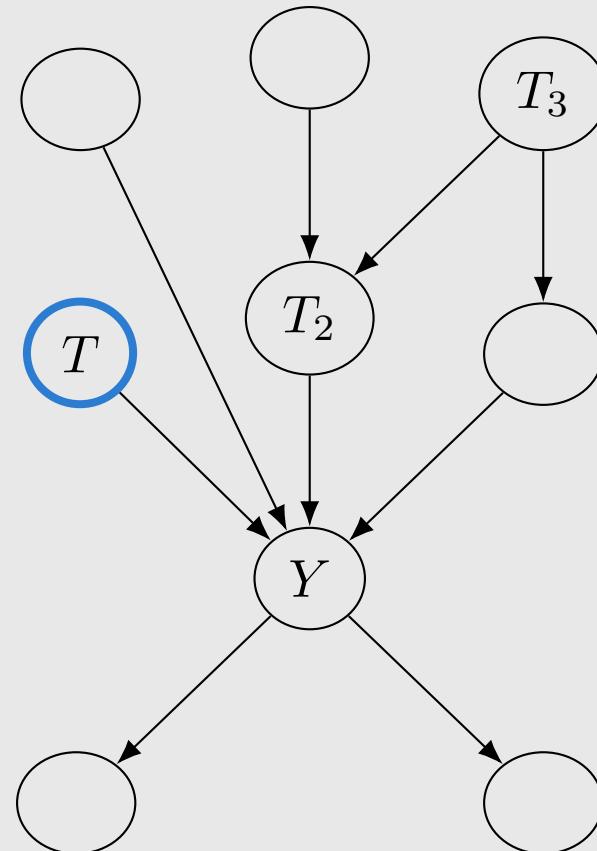


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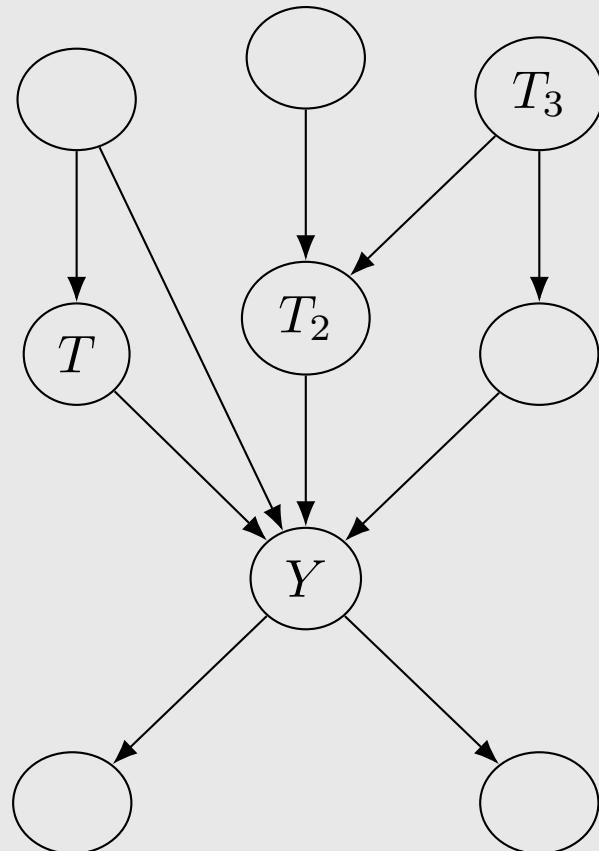


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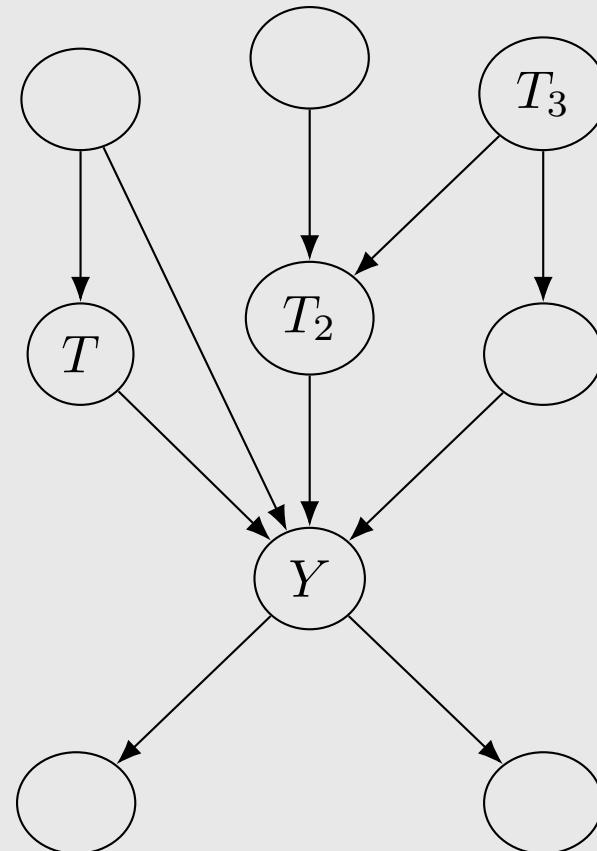


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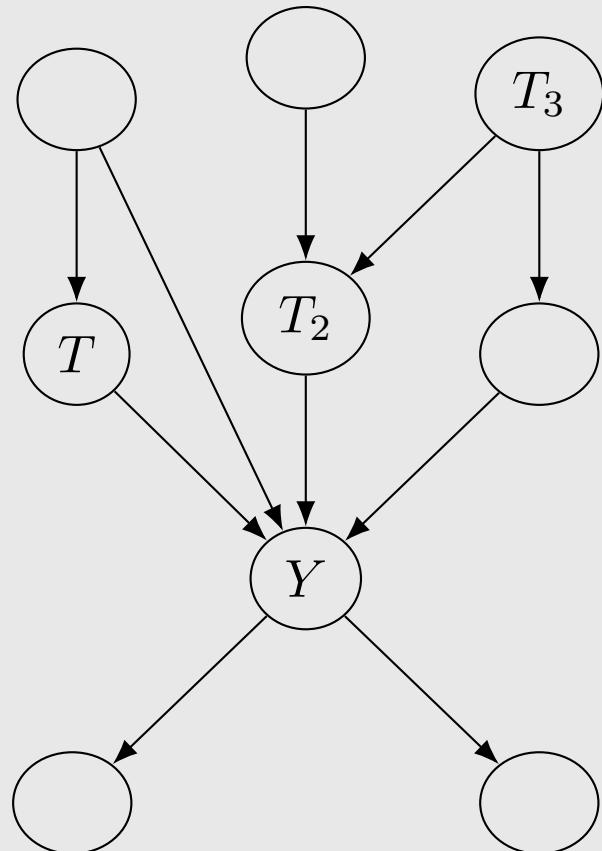


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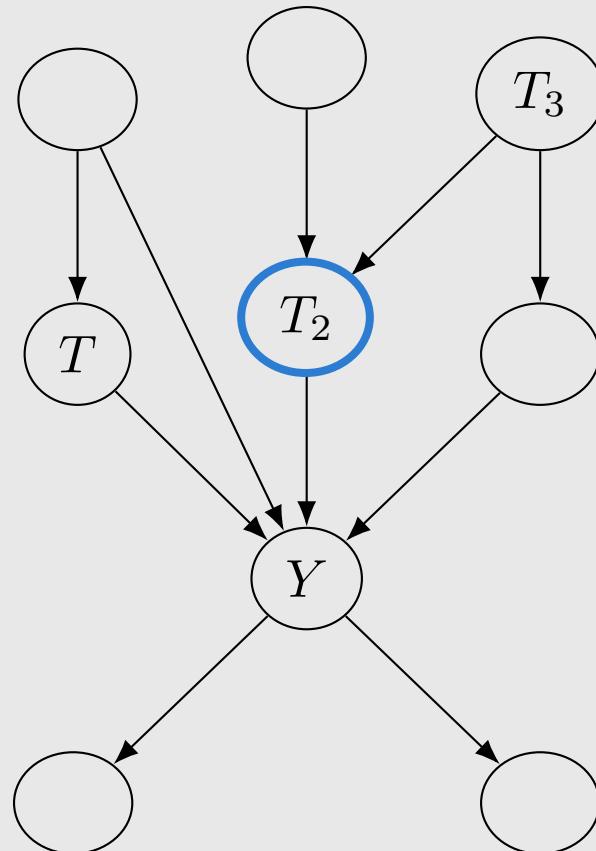


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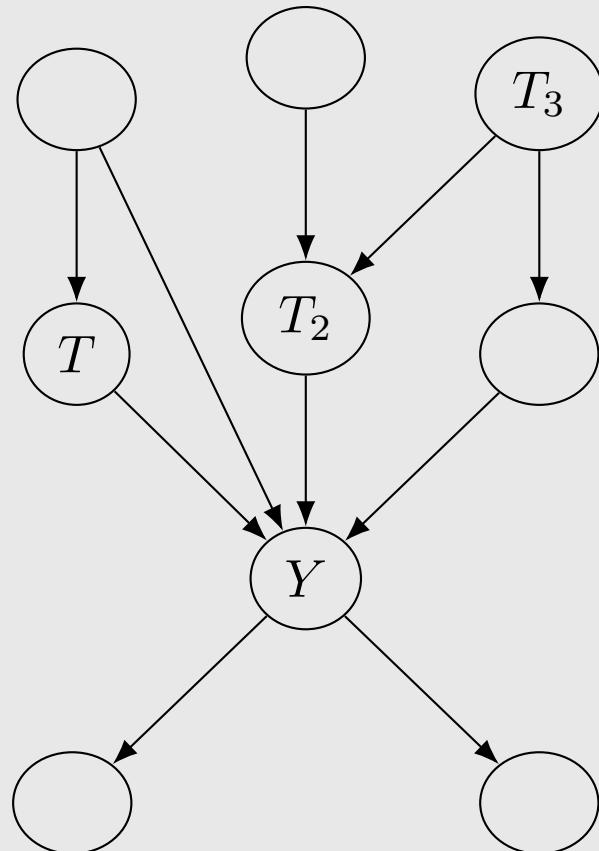


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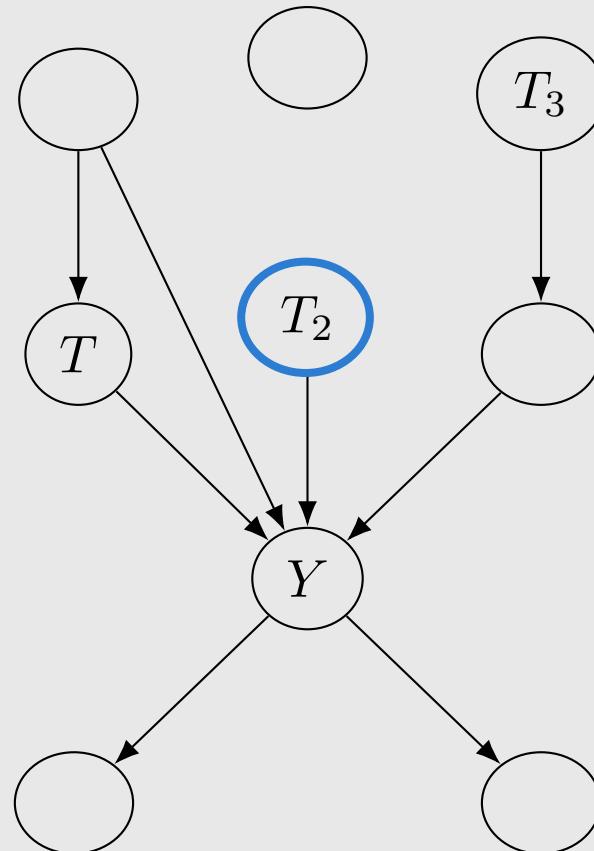


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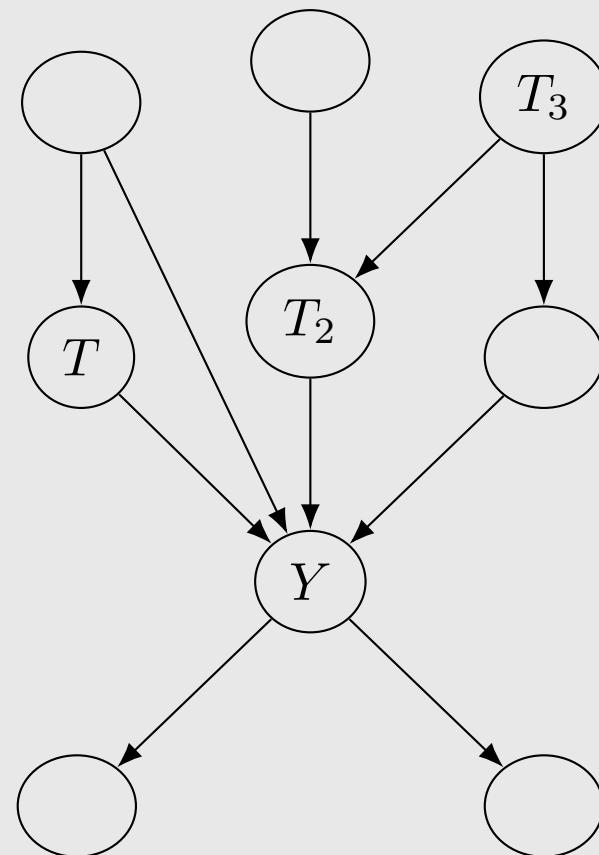
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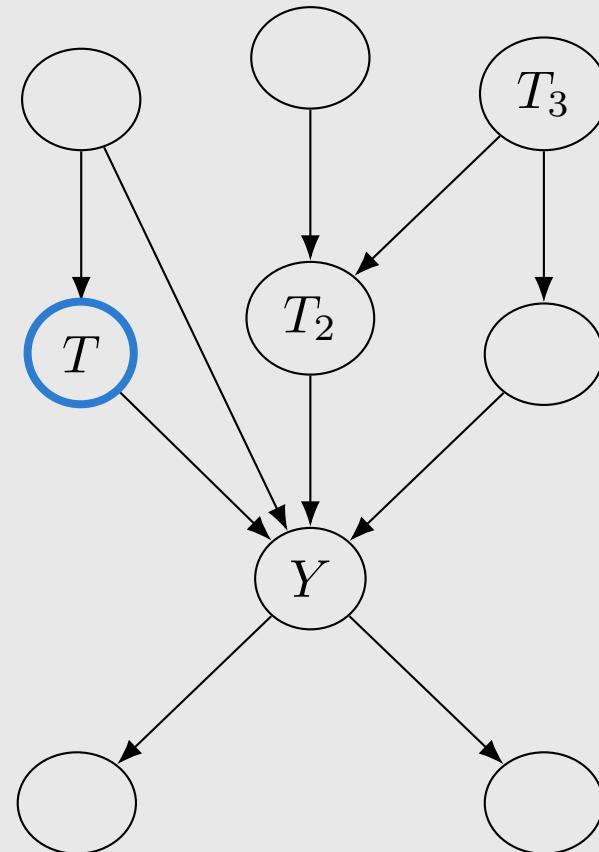


What would it mean if modularity is violated?



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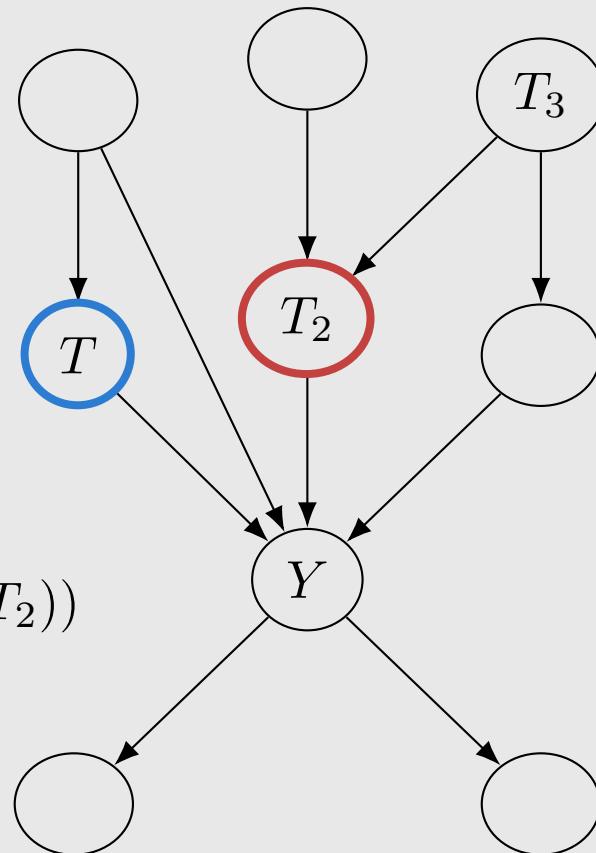
Intervention on T not
only changes $P(T | \text{pa}(T))$



What would it mean if modularity is violated?

Intervention on T not
only changes $P(T | \text{pa}(T))$

but also changes other
mechanisms such as $P(T_2 | \text{pa}(T_2))$



Truncated factorization

Recall the Bayesian network factorization:

$$P(x_1, \dots, x_n) = \prod_i P(x_i \mid \text{pa}_i)$$

Truncated factorization

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$$P(x_1, \dots, x_n \mid do(S = s)) = \prod_i P(x_i \mid \text{pa}_i)$$

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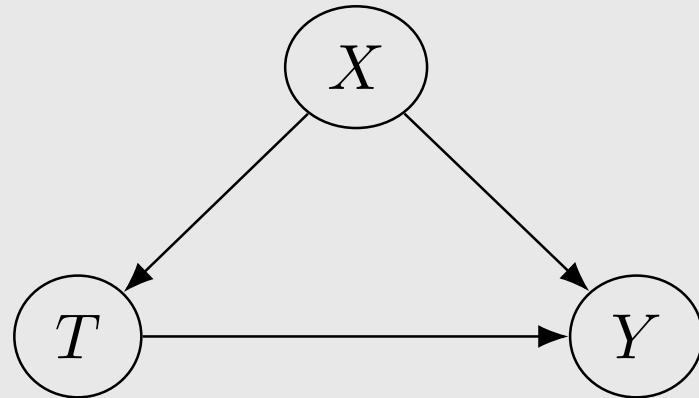
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Otherwise,

$$P(x_1, \dots, x_n \mid do(S = s)) = 0$$

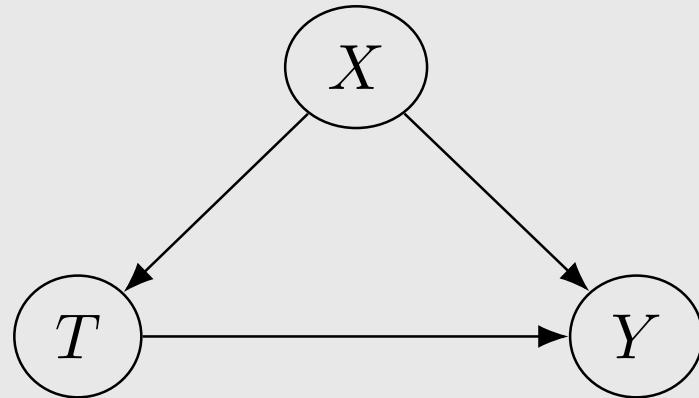
Simple identification via truncated factorization

Goal: identify $P(y \mid do(t))$



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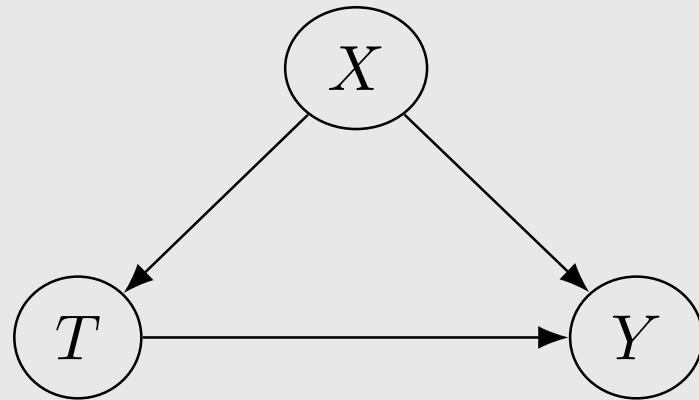
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Bayesian network factorization: $P(y, t, x) = P(x) P(t \mid x) P(y \mid t, x)$

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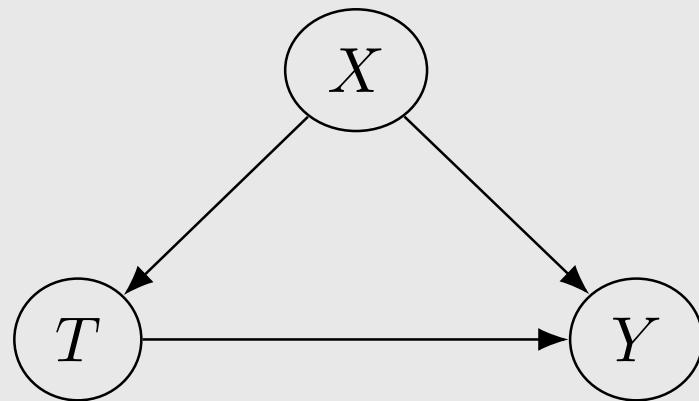


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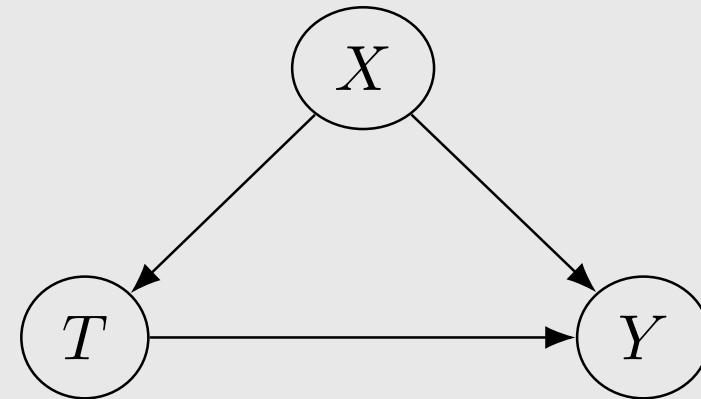
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Marginalize: $P(y \mid do(t)) = \sum_x P(y \mid t, x) P(x)$

Association vs. causation revisited

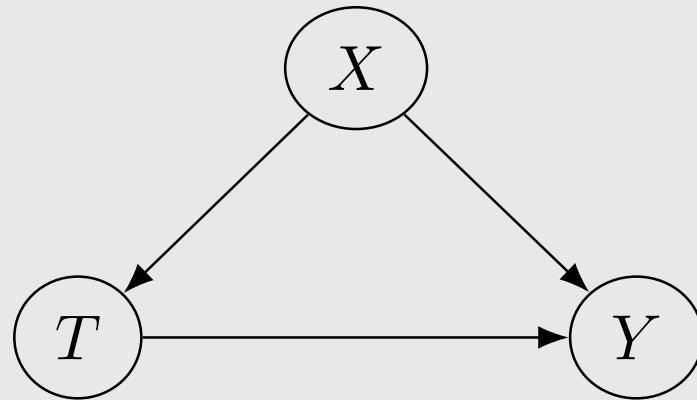
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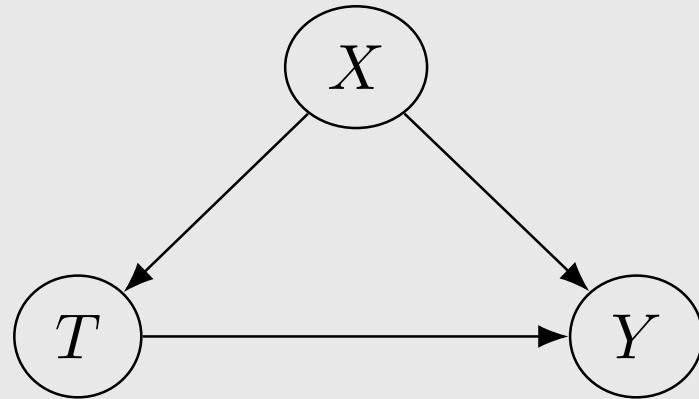
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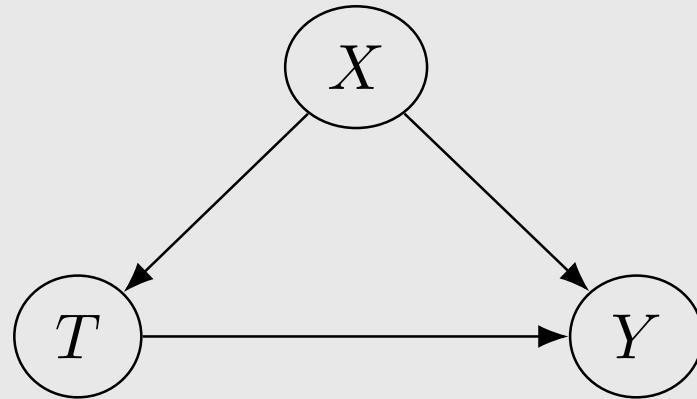


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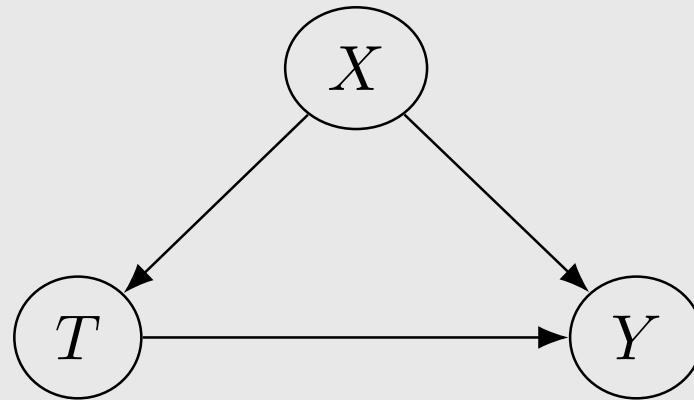


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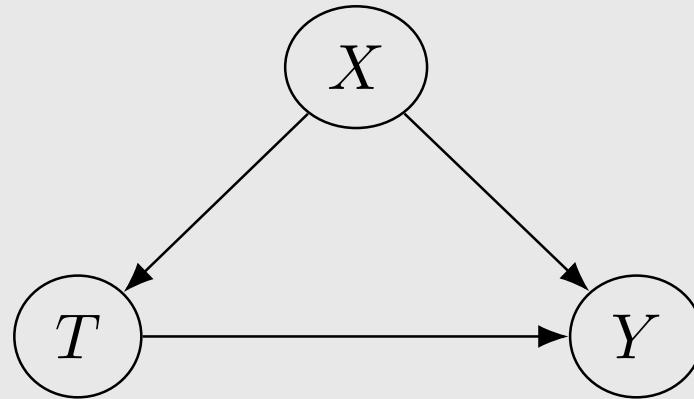


$$\sum_x P(y \mid t, x) P(x \mid t) = \sum_x P(y, x \mid t)$$

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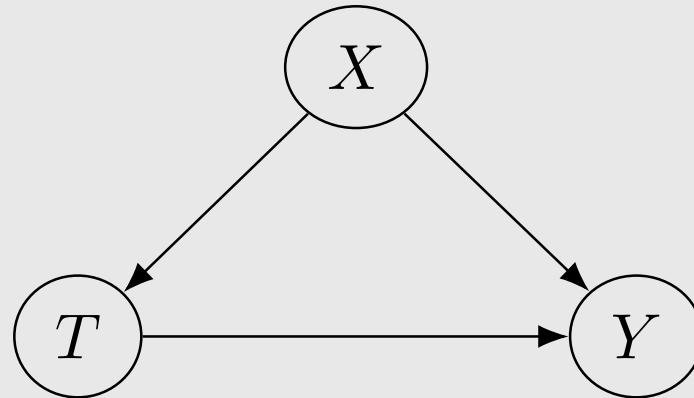


$$\begin{aligned} \sum_x P(y \mid t, x) P(x \mid t) &= \sum_x P(y, x \mid t) \\ &= P(y \mid t) \end{aligned}$$

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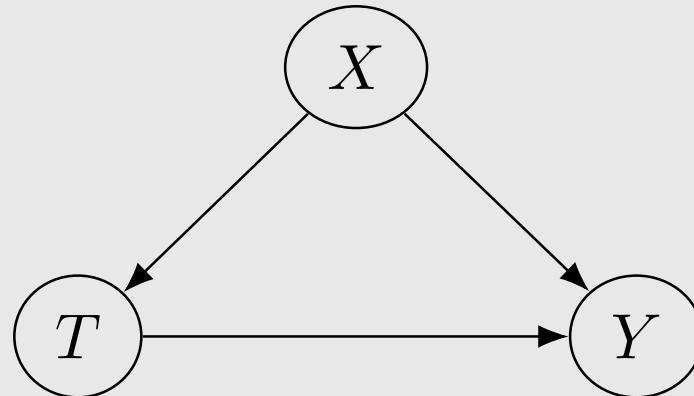


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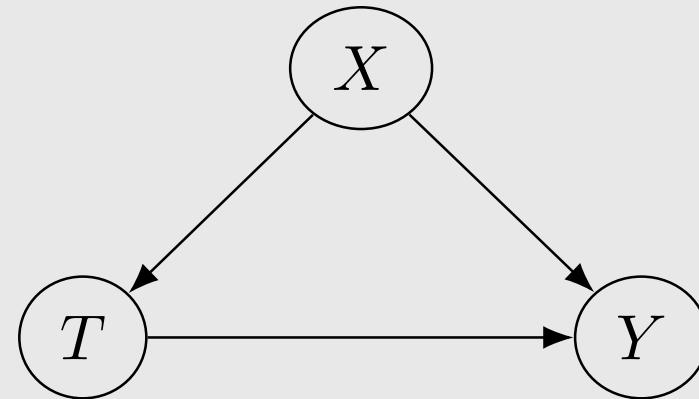


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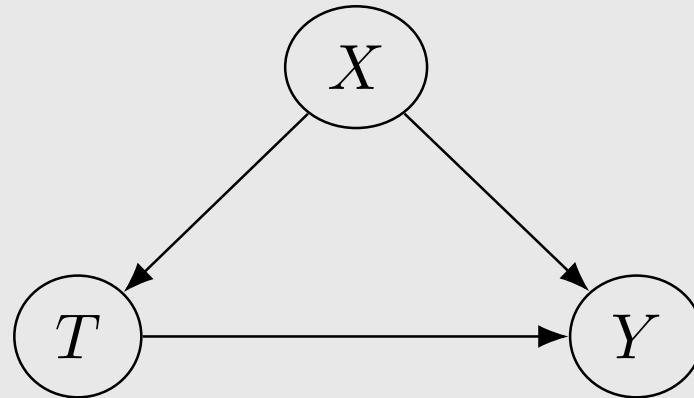


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$$\begin{aligned} \sum_x P(y \mid t, x) \underline{P(x \mid t)} &= \sum_x P(y, x \mid t) \\ &= \underline{P(y \mid t)} \end{aligned}$$

The *do*-operator

Main assumption: modularity

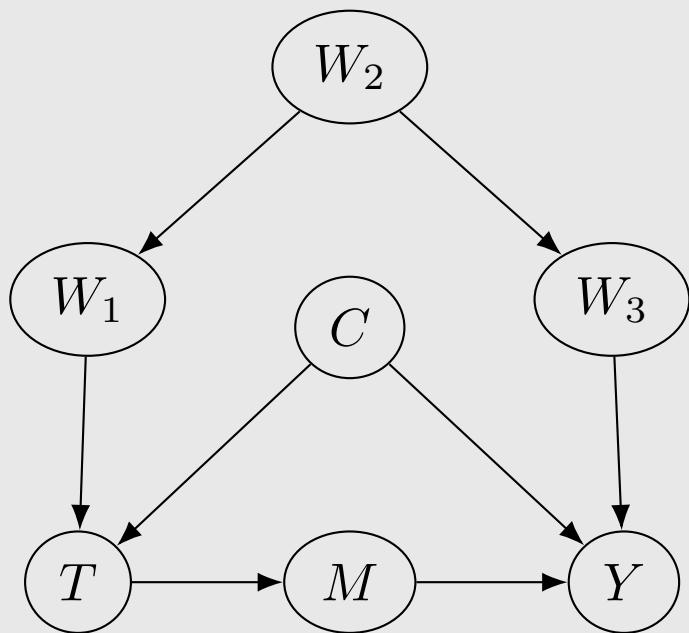
Backdoor adjustment

Structural causal models

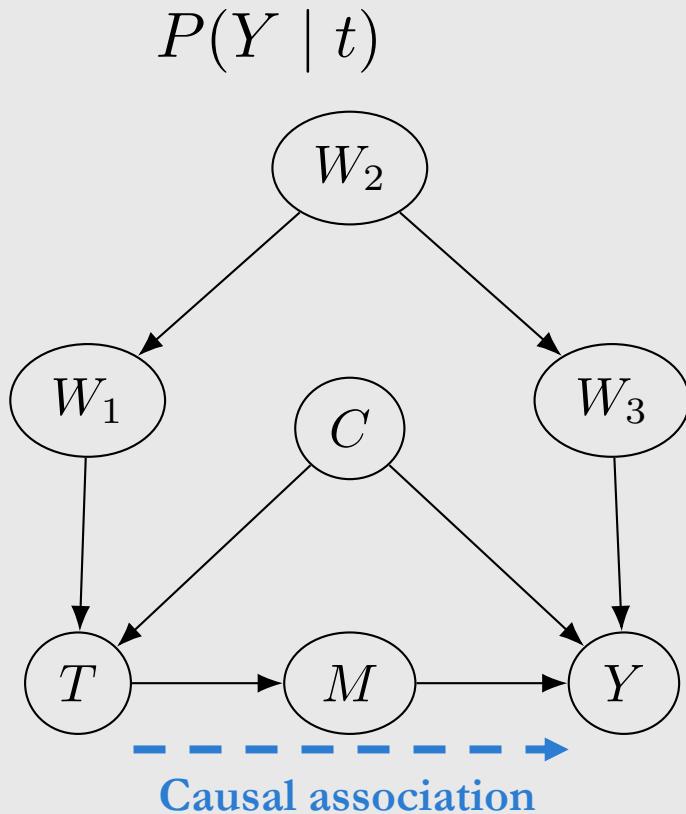
A complete example with estimation

Blocking backdoor paths

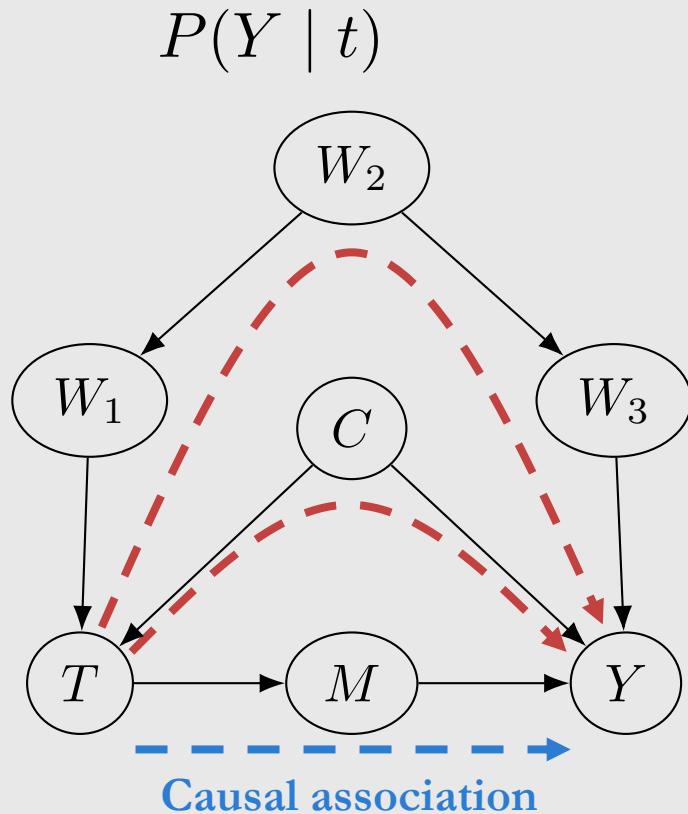
$$P(Y \mid t)$$



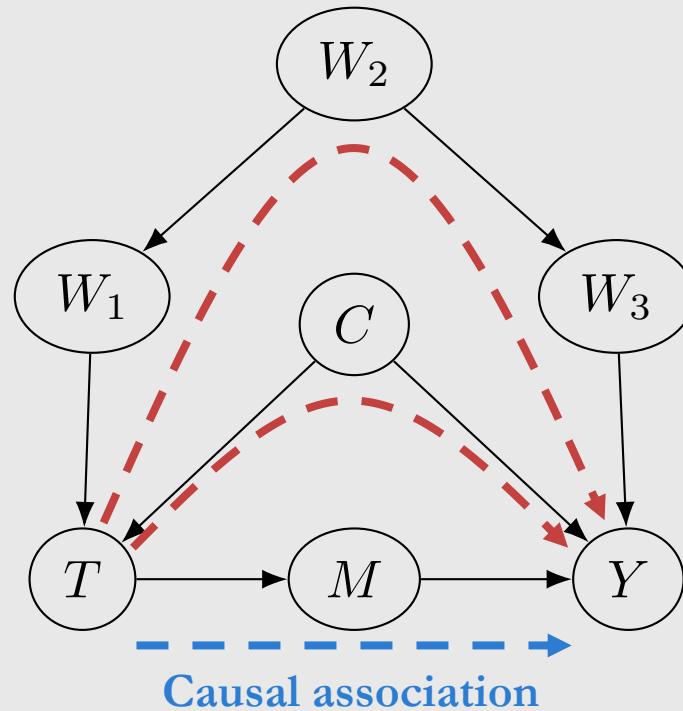
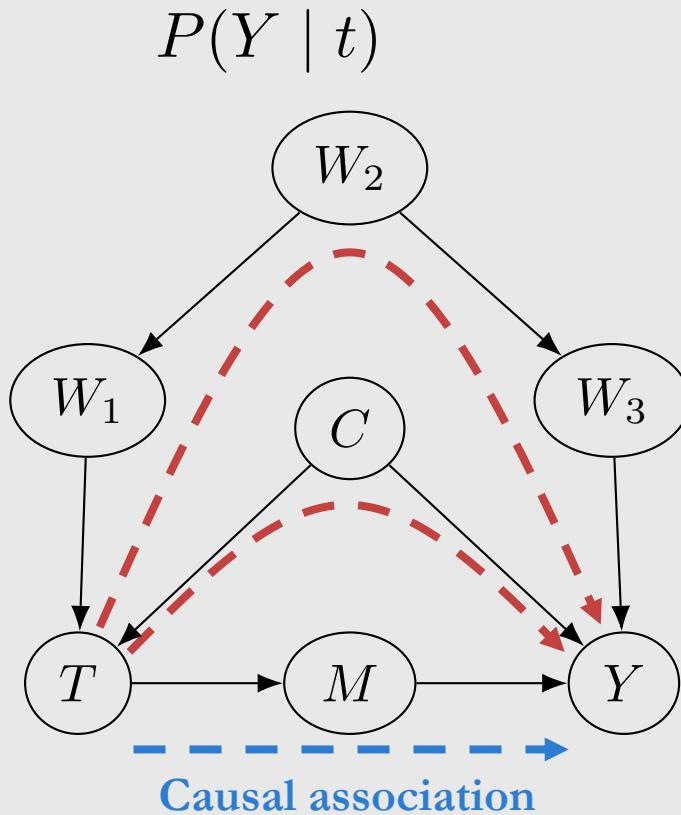
Blocking backdoor paths



Blocking backdoor paths

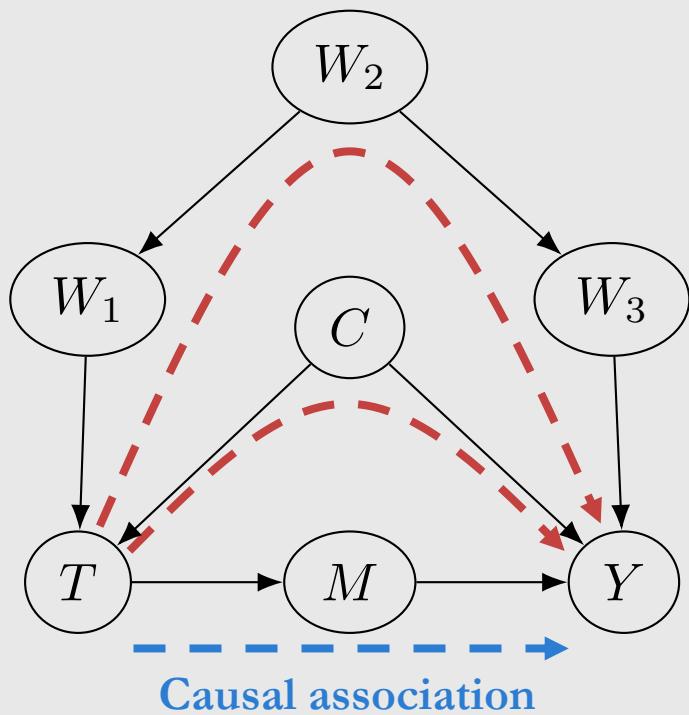


Blocking backdoor paths

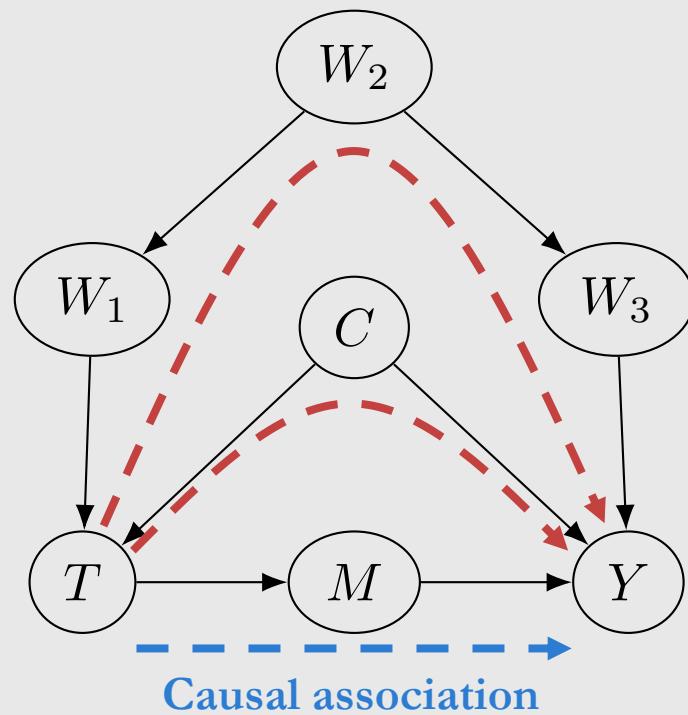


Blocking backdoor paths

$P(Y | t)$

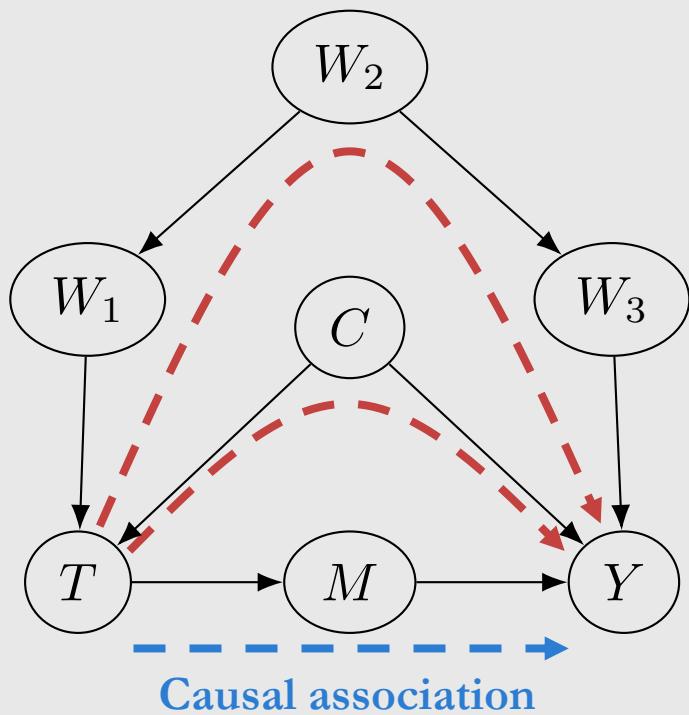


$P(Y | do(t))$

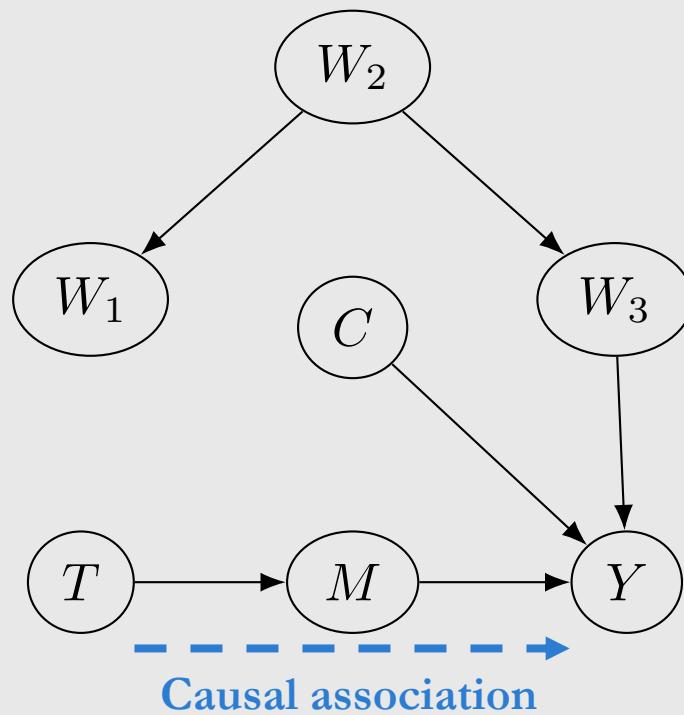


Blocking backdoor paths

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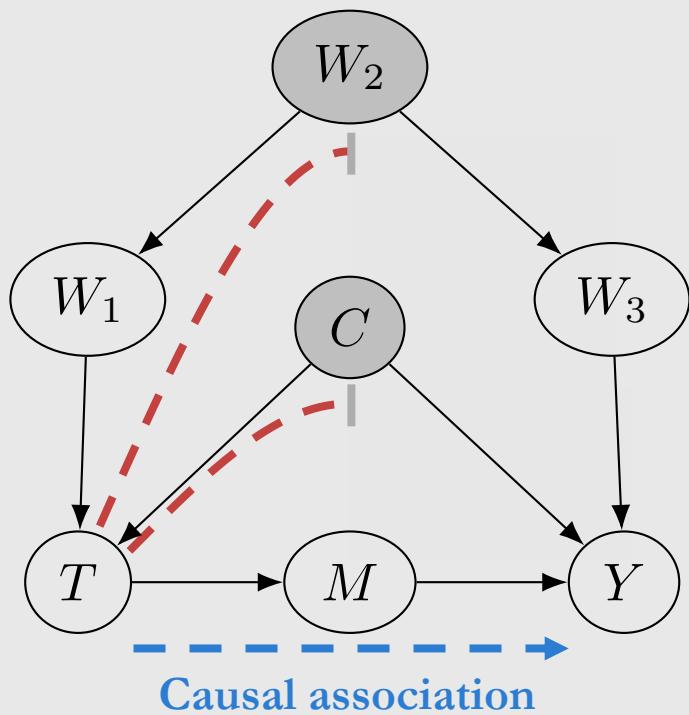


$P(Y | do(t))$

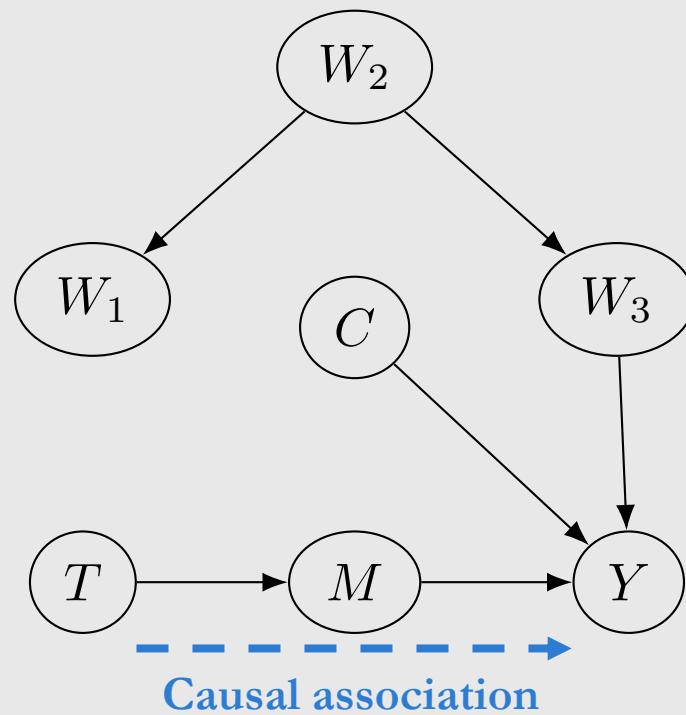


Blocking backdoor paths

$$P(Y \mid t, c, w_2)$$

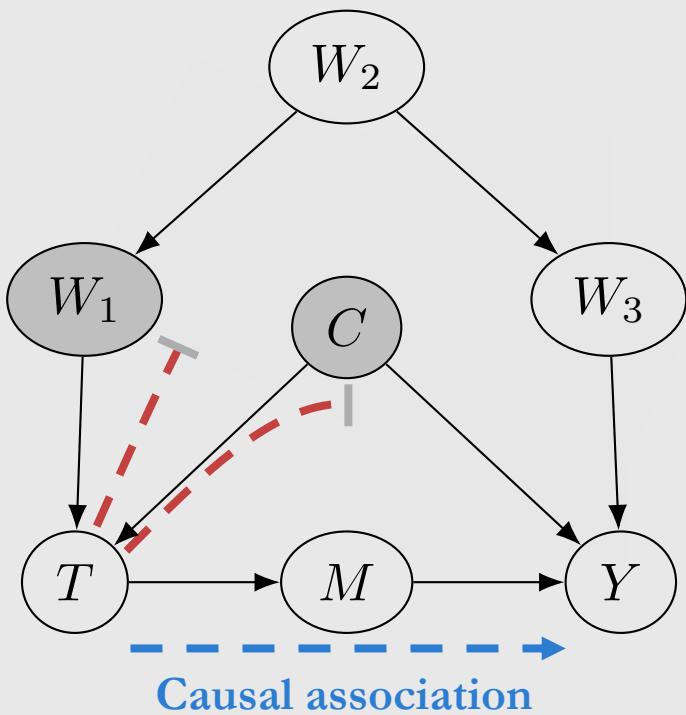


$$P(Y \mid do(t))$$

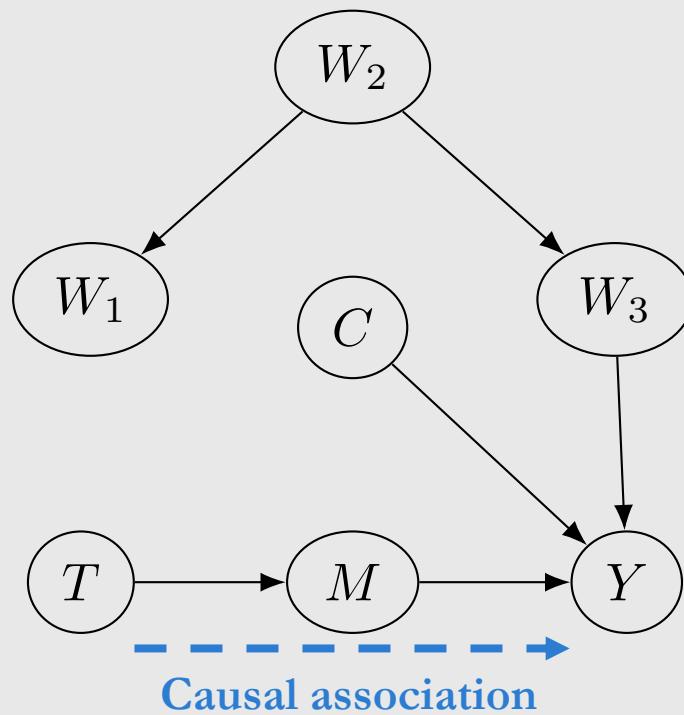


Blocking backdoor paths

$$P(Y \mid t, c, w_1)$$

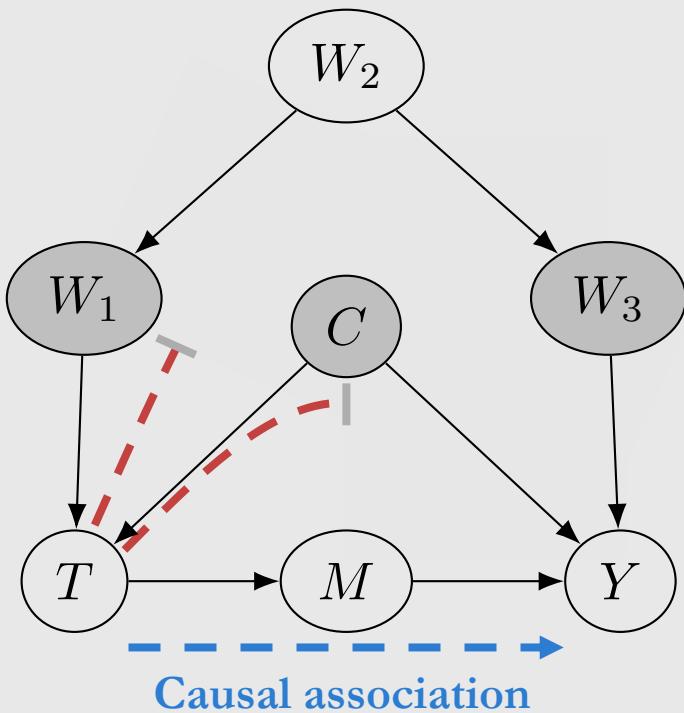


$$P(Y \mid do(t))$$

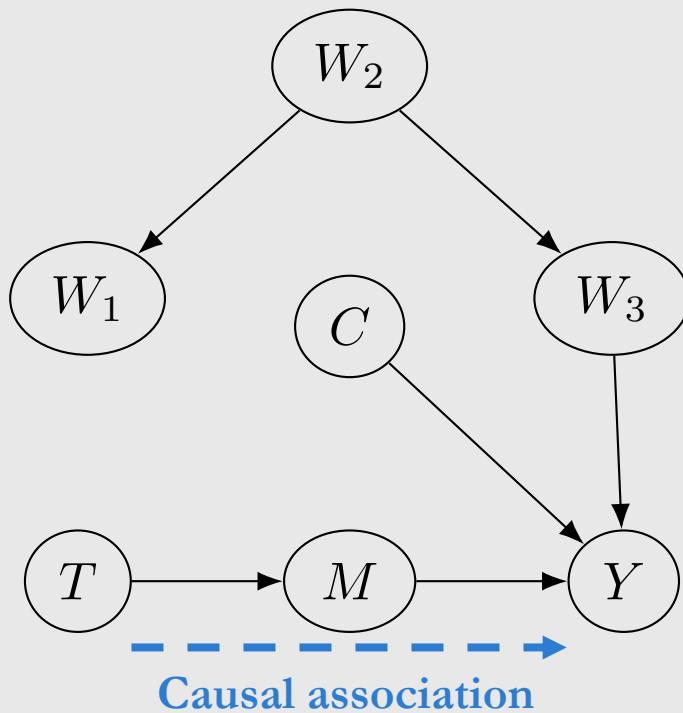


Blocking backdoor paths

$$P(Y \mid t, c, w_1, w_3)$$



$$P(Y \mid do(t))$$



Backdoor criterion and backdoor adjustment

Backdoor criterion and backdoor adjustment

A set of variables W satisfies the backdoor criterion relative to T and Y if the following are true:

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Backdoor criterion and backdoor adjustment

A set of variables W satisfies the backdoor criterion relative to T and Y if the following are true:

1. W blocks all backdoor paths from T to Y
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Given the modularity assumption and that W satisfies the backdoor criterion, we can identify the causal effect of T on Y :

$$P(y \mid do(t)) = \sum_w P(y \mid t, w) P(w)$$

Proof of backdoor adjustment

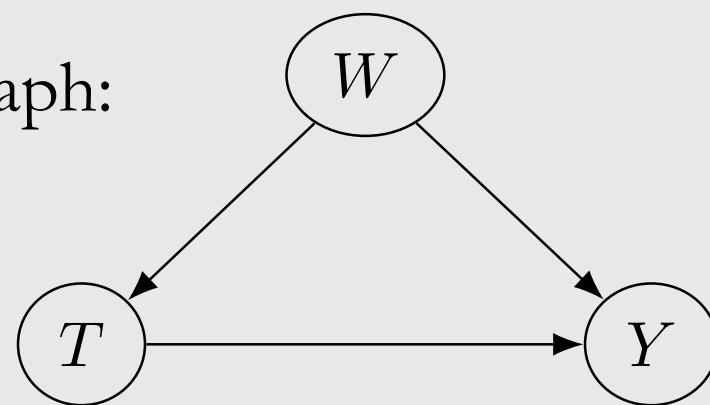
$$P(y \mid do(t))$$

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Proof of backdoor adjustment

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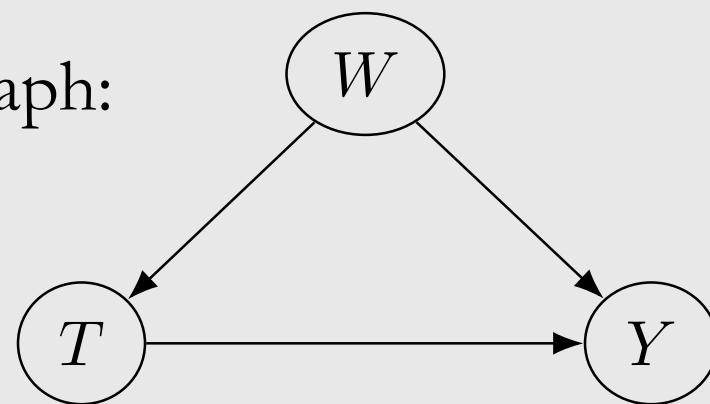
Example graph:



Proof of backdoor adjustment

$$P(y \mid do(t)) = \sum_w P(y \mid do(t), w) P(w \mid do(t))$$

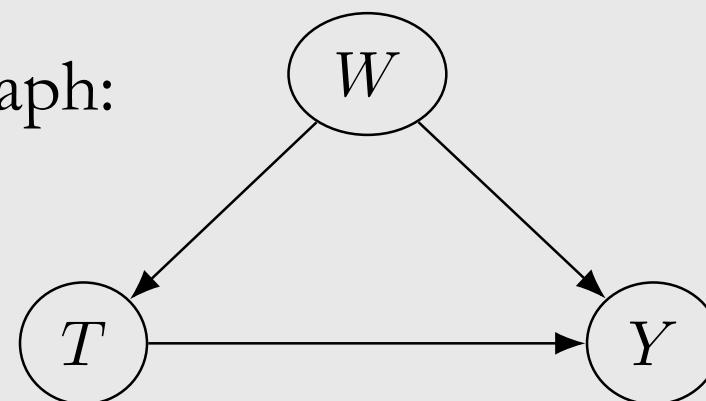
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Proof of backdoor adjustment

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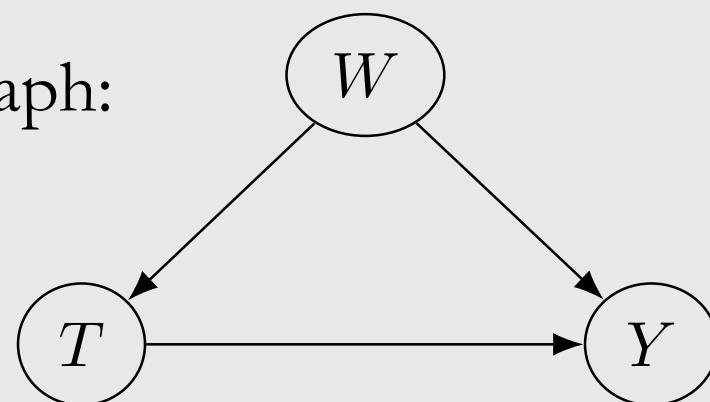
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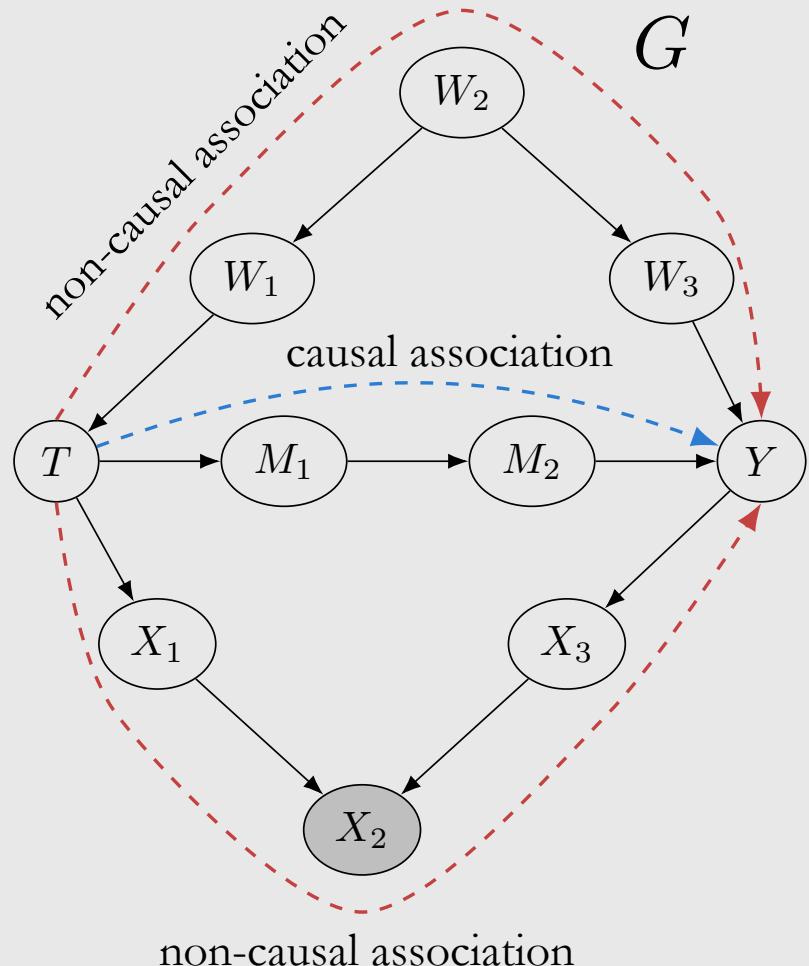
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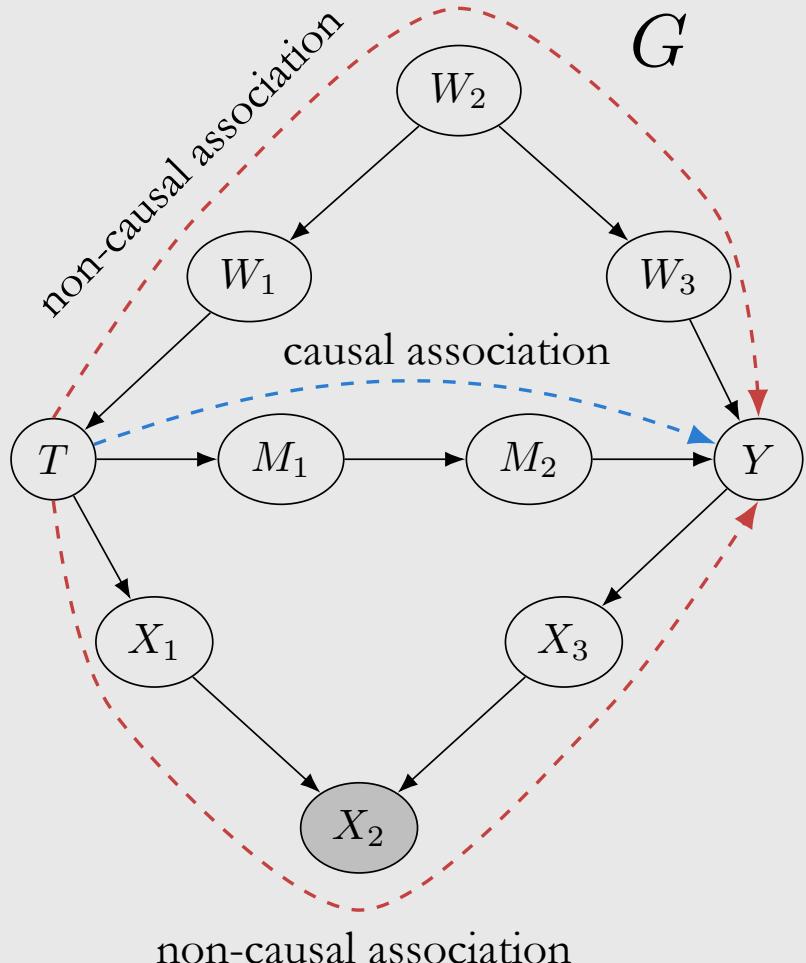
Example graph:



Backdoor criterion as d-separation

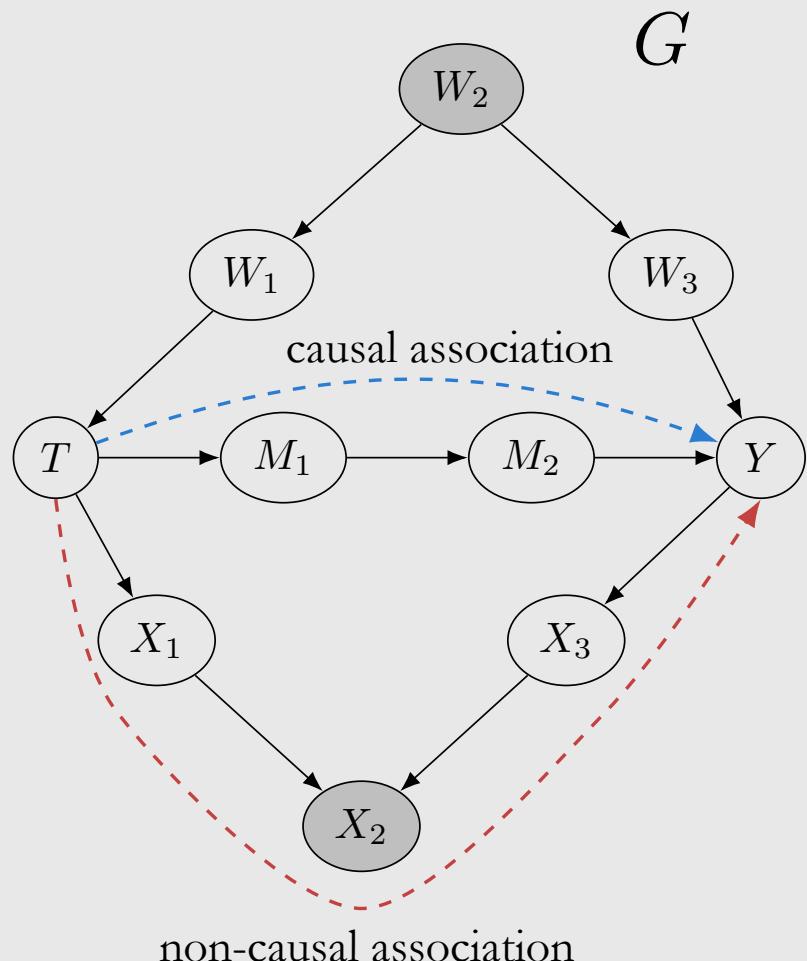


Backdoor criterion as d-separation



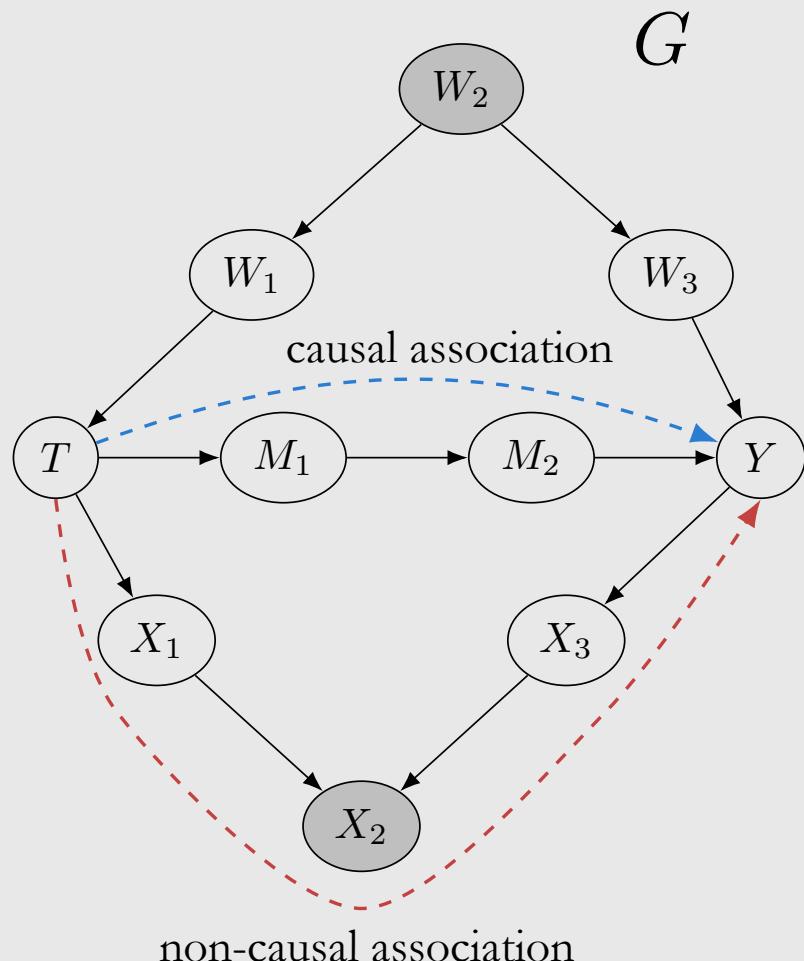
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Backdoor criterion as d-separation



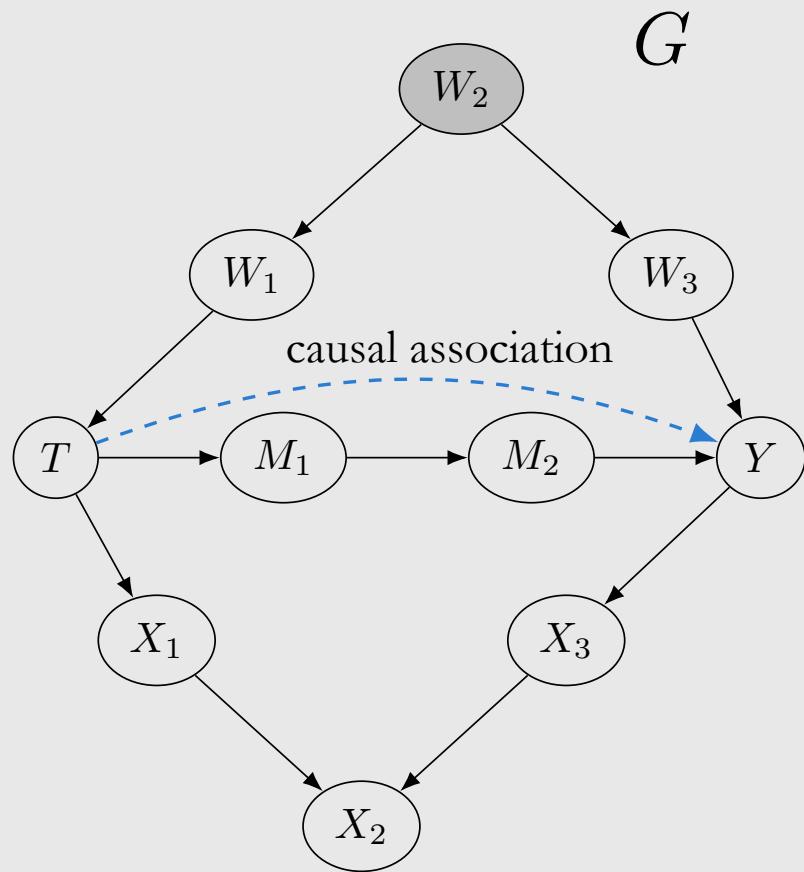
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Backdoor criterion as d-separation



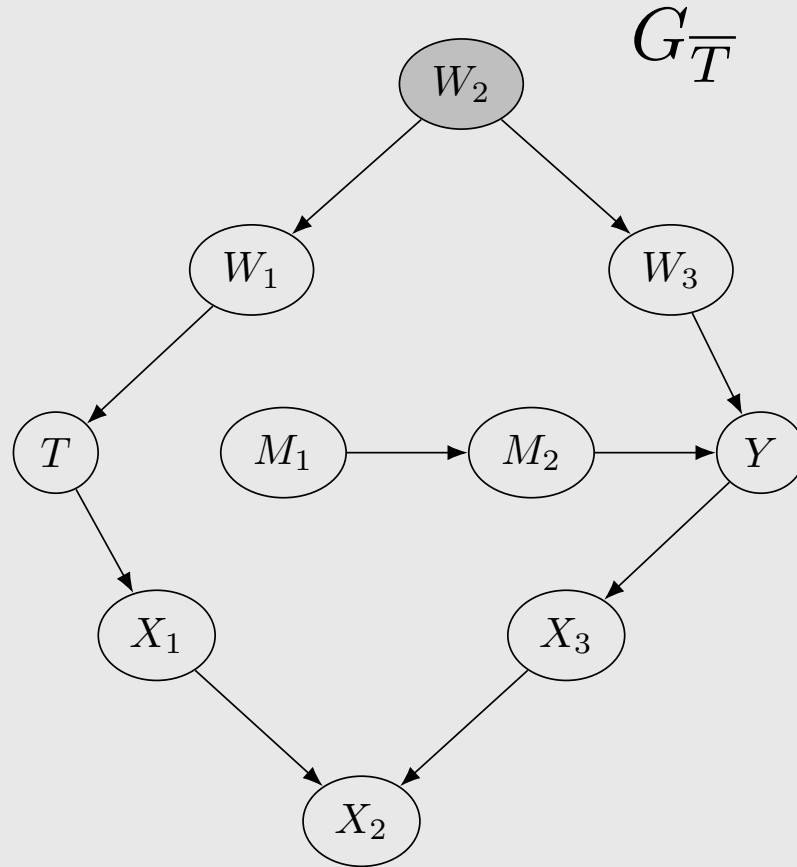
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Backdoor criterion as d-separation



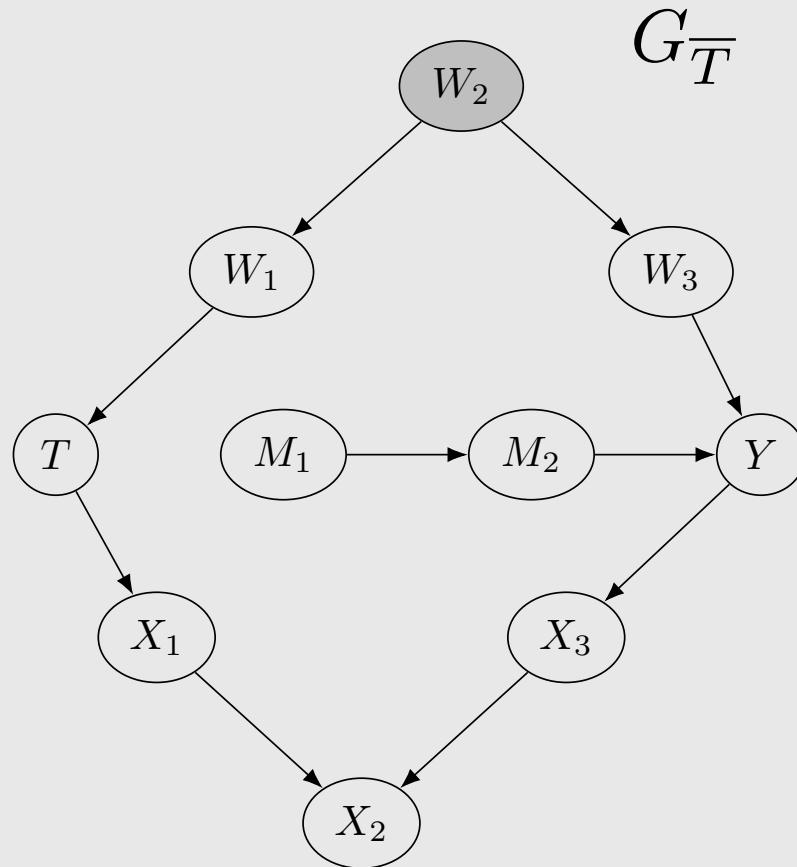
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Backdoor criterion as d-separation



1. W blocks all backdoor paths from T to Y
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$$Y \perp\!\!\!\perp_{G_{\overline{T}}} T \mid W$$

Question:

How does this backdoor adjustment relate to the adjustment formula we saw in the potential outcomes lecture?

Backdoor adjustment:

$$P(y \mid do(t)) = \sum_w P(y \mid t, w) P(w)$$

Adjustment formula from before:

$$\mathbb{E}[Y(1) - Y(0)] = \mathbb{E}_W [\mathbb{E}[Y \mid T = 1, W] - \mathbb{E}[Y \mid T = 0, W]]$$

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How does this backdoor adjustment relate to the adjustment formula we saw in the potential outcomes lecture?

Section 4.4.1 of the ICI book

Backdoor adjustment:

$$P(y \mid do(t)) = \sum_w P(y \mid t, w) P(w)$$

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The *do*-operator

Main assumption: modularity

Backdoor adjustment

Structural causal models

A complete example with estimation

Structural equations

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Structural equation for A as a cause of B:

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Structural equation for A as a cause of B:

$$B := f(A)$$

$$B := f(A, U)$$

Structural equations

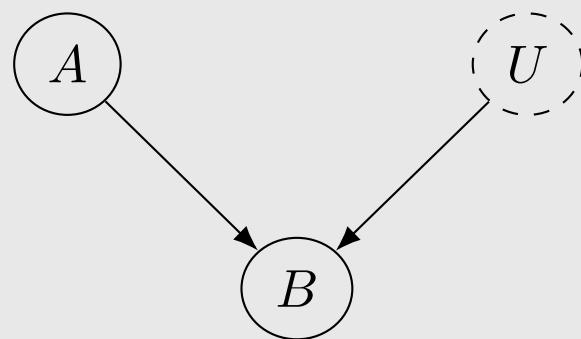
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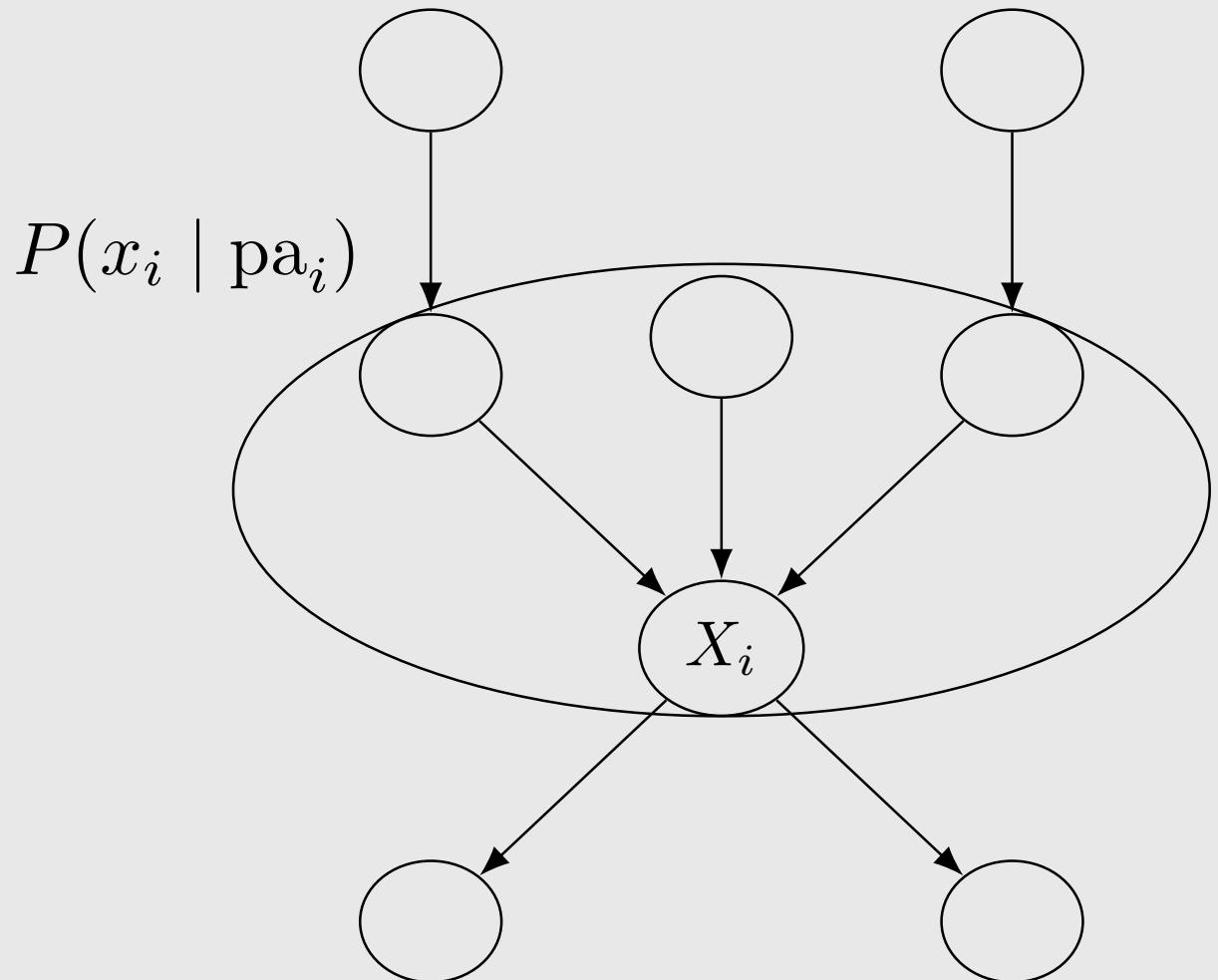
Structural equation for A as a cause of B:

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Causal mechanisms and direct causes revisited

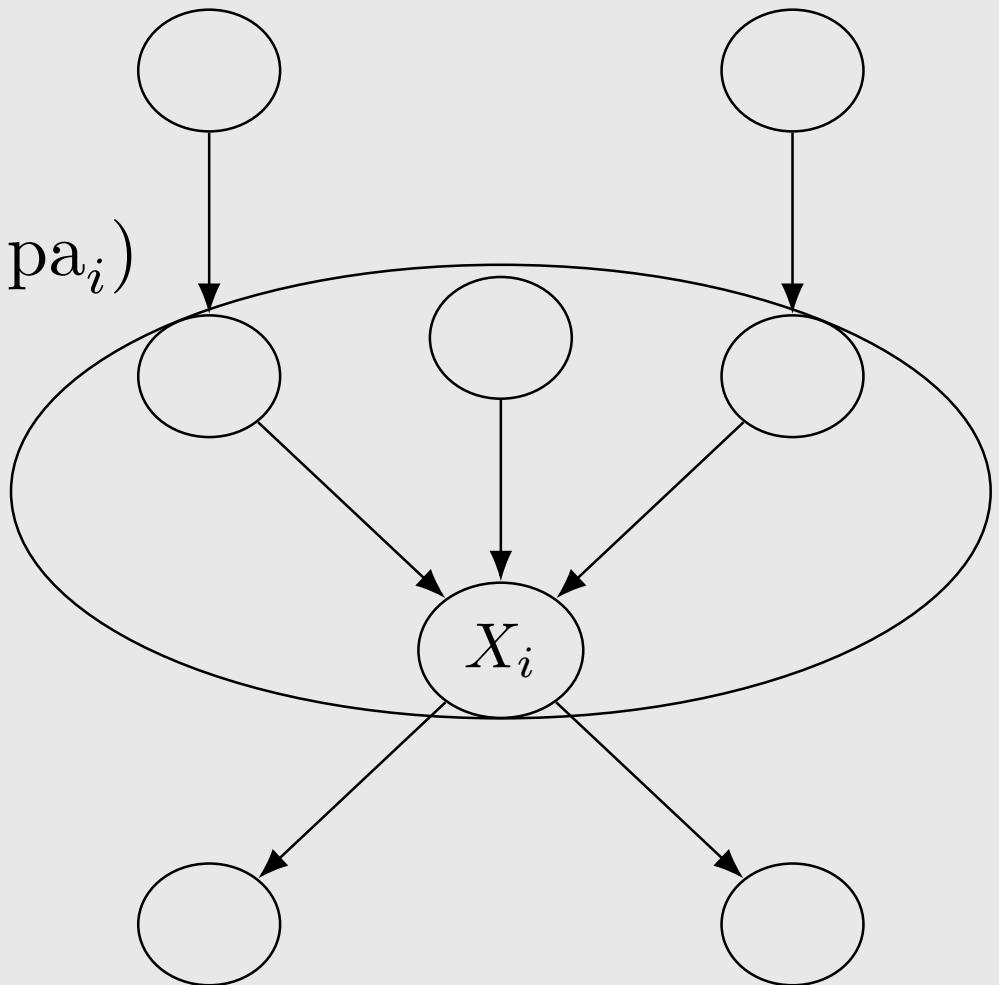


Causal mechanisms and direct causes revisited

Causal mechanism for X_i

$$X_i := f(A, B, \dots)$$

$$P(x_i | \text{pa}_i)$$



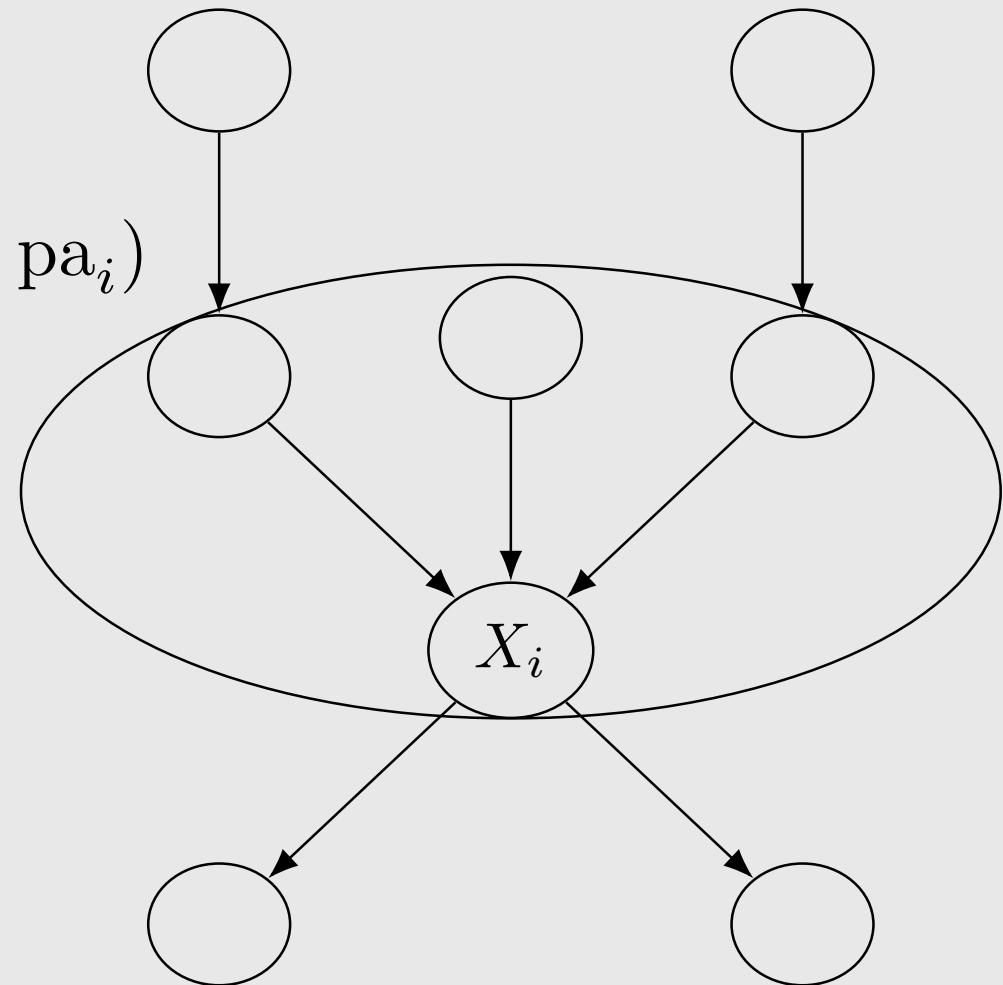
Causal mechanisms and direct causes revisited

Causal mechanism for X_i

$$X_i := f(A, B, \dots)$$

Direct causes of X_i

$$P(x_i | \text{pa}_i)$$



Structural causal models (SCMs)

$$B := f_B(A, U_B)$$

$$M : \quad C := f_C(A, B, U_C)$$

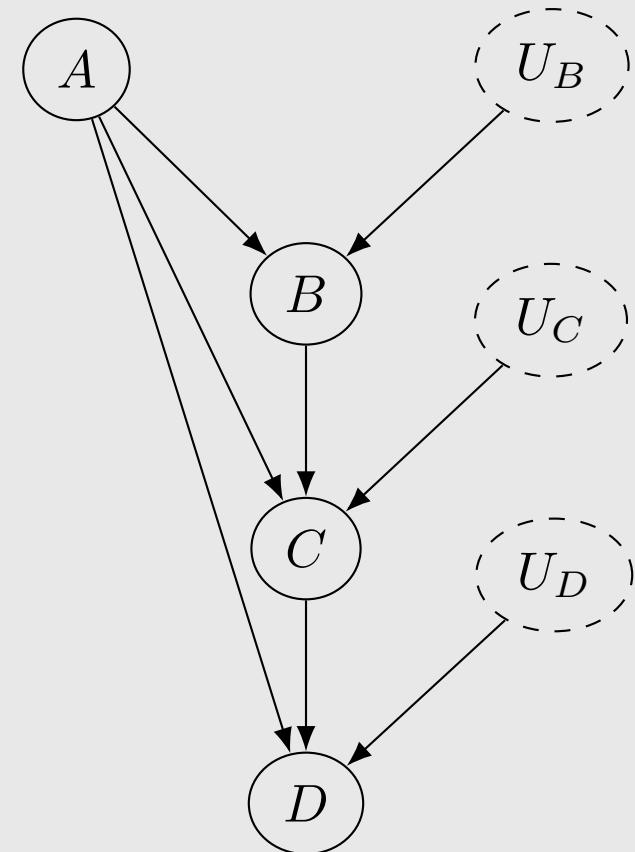
$$D := f_D(A, C, U_D)$$

Structural causal models (SCMs)

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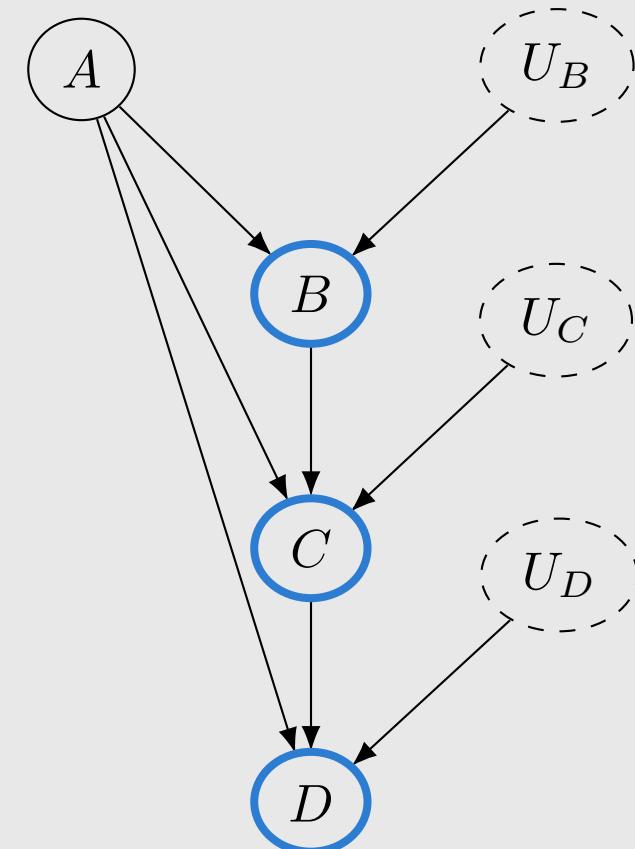


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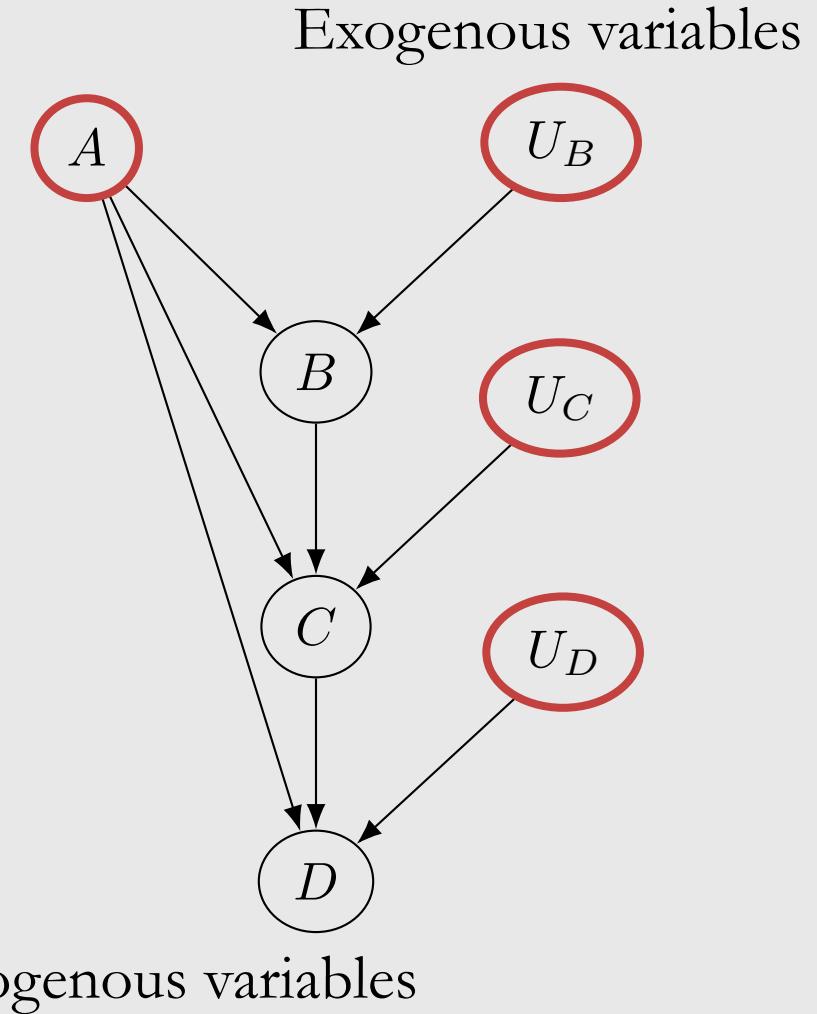
Endogenous variables

Structural causal models (SCMs)

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Structural causal models (SCMs)

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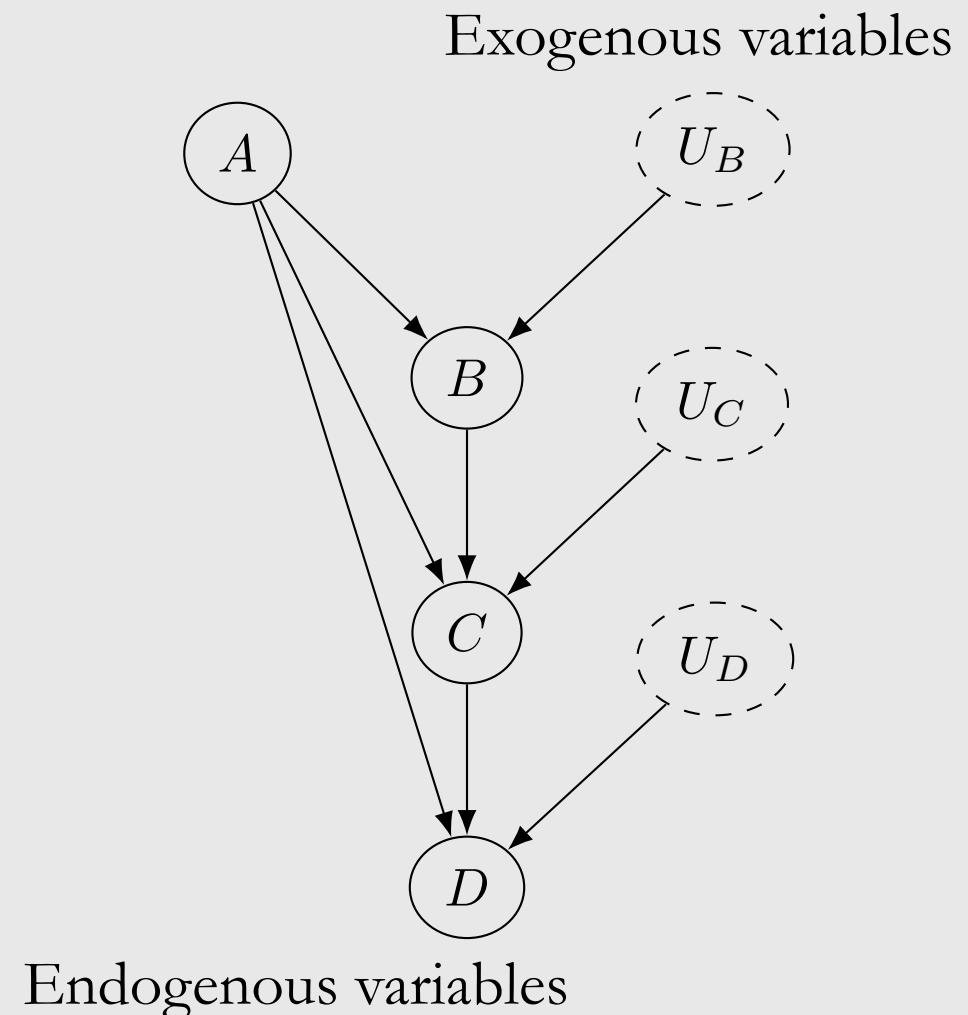
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SCM Definition

A tuple of the following sets:

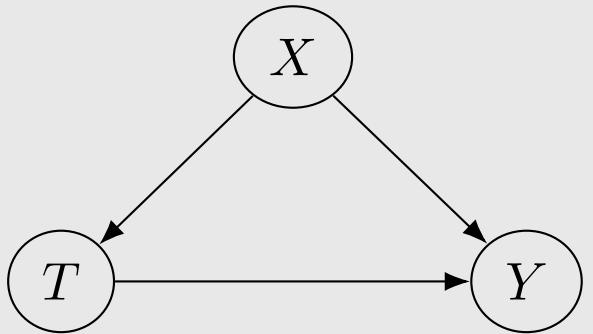
1. A set of endogenous variables
2. A set of exogenous variables
3. A set of functions, one to generate each endogenous variable as a function of the other variables



Interventions

SCM
(model)

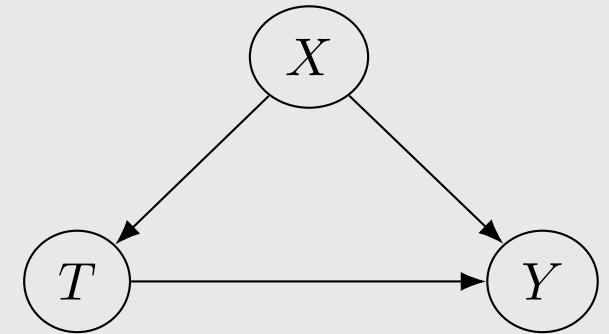
$$M : \quad \begin{aligned} T &:= f_T(X, U_T) \\ Y &:= f_Y(X, T, U_Y) \end{aligned}$$



Interventions

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Interventional SCM
(submodel)

$$M_t : \quad \begin{aligned} T &:= t \\ Y &:= f_Y(X, T, U_Y) \end{aligned}$$

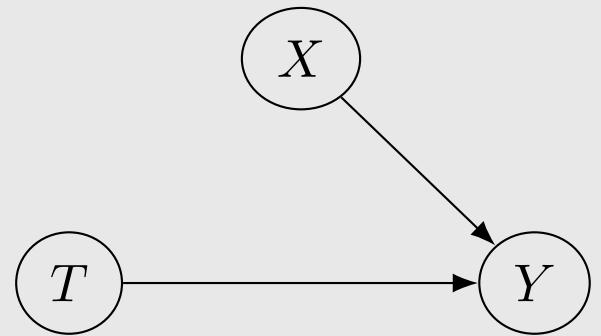
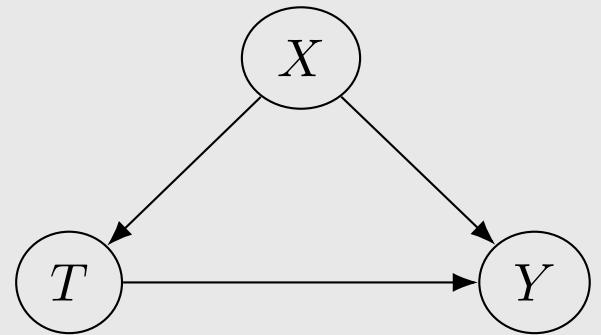
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Modularity assumption for SCMs

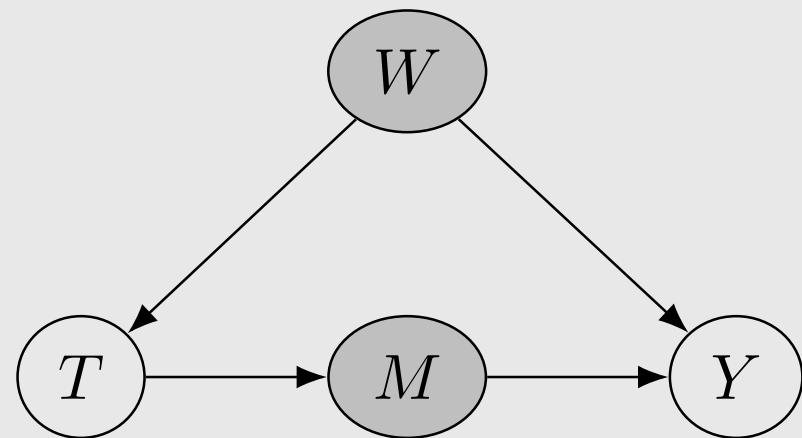
Consider an SCM M and an interventional SCM M_t that we get by performing the intervention $do(T = t)$. The modularity assumption states that M and M_t share all of their structural equations except the structural equation for T , which is $T := t$ in M_t .

Modularity assumption for SCMs

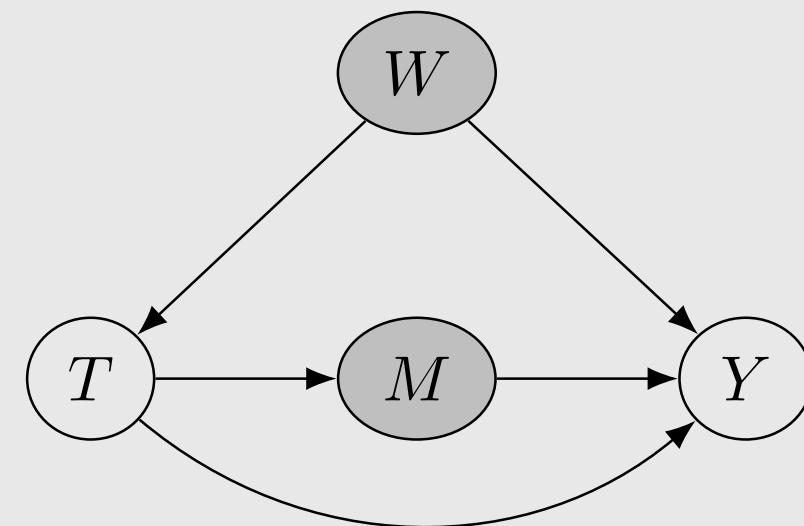
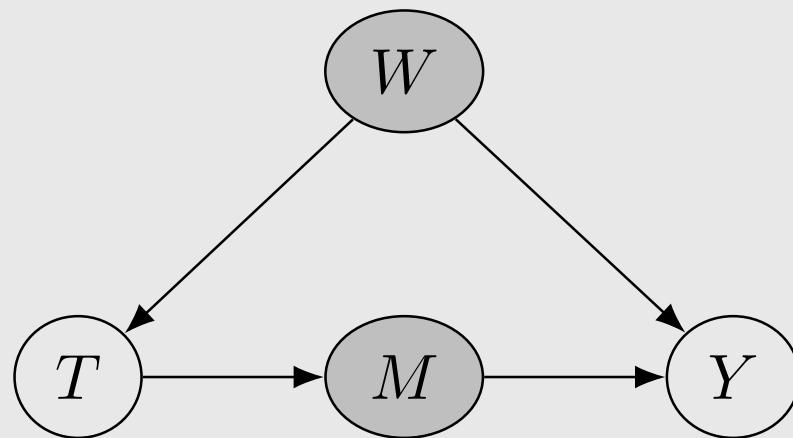
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$$\begin{array}{ll} M : & \begin{aligned} T &:= f_T(X, U_T) \\ Y &:= f_Y(X, T, U_Y) \end{aligned} & M_t : & \begin{aligned} T &:= t \\ Y &:= f_Y(X, T, U_Y) \end{aligned} \end{array}$$

Why not condition on descendants of treatment:
blocking causal association

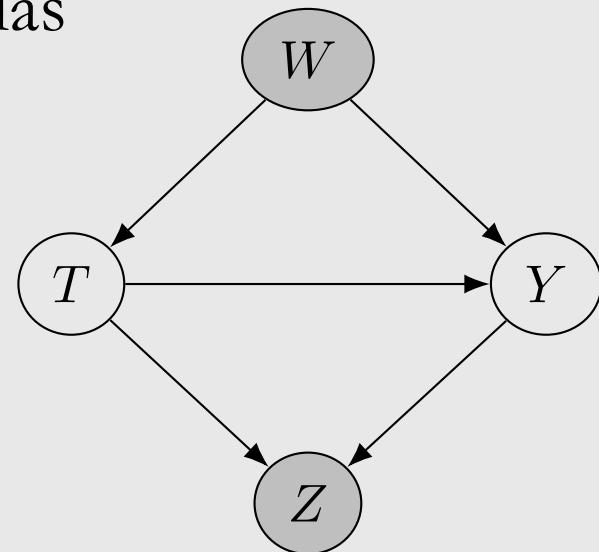


Why not condition on descendants of treatment:
blocking causal association



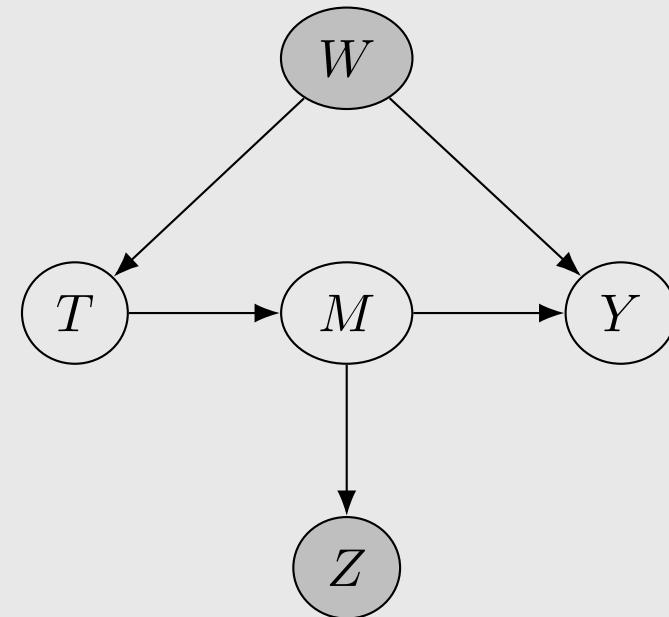
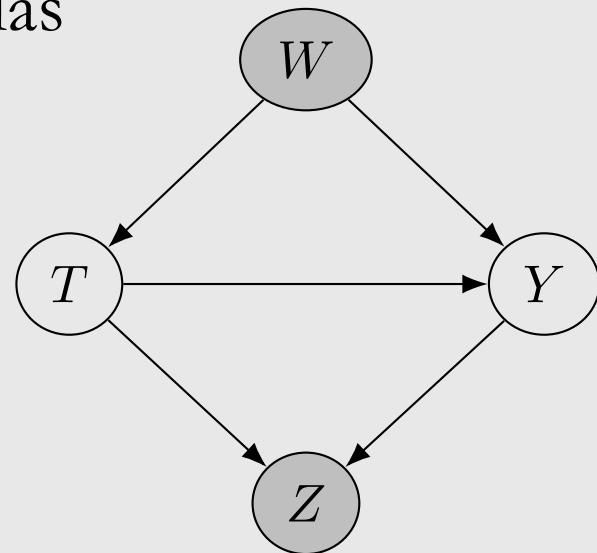
Why not condition on descendants of treatment:
inducing new post-treatment association

Collider bias



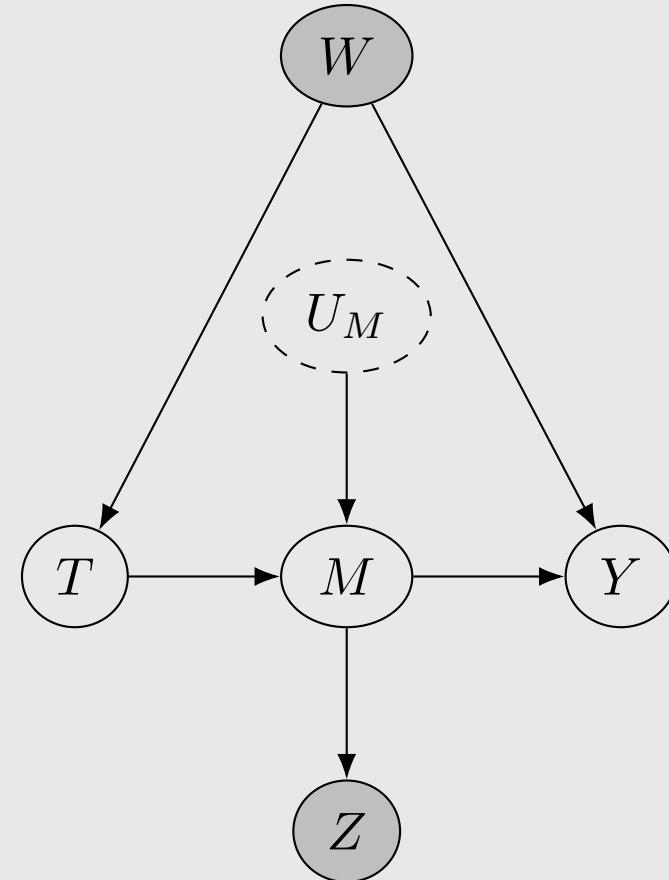
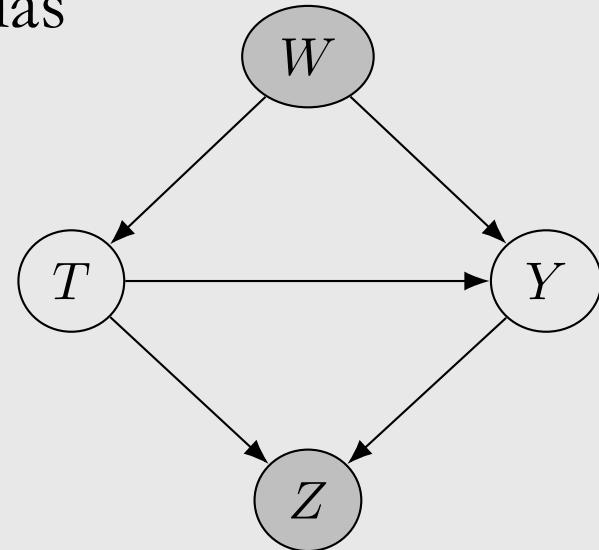
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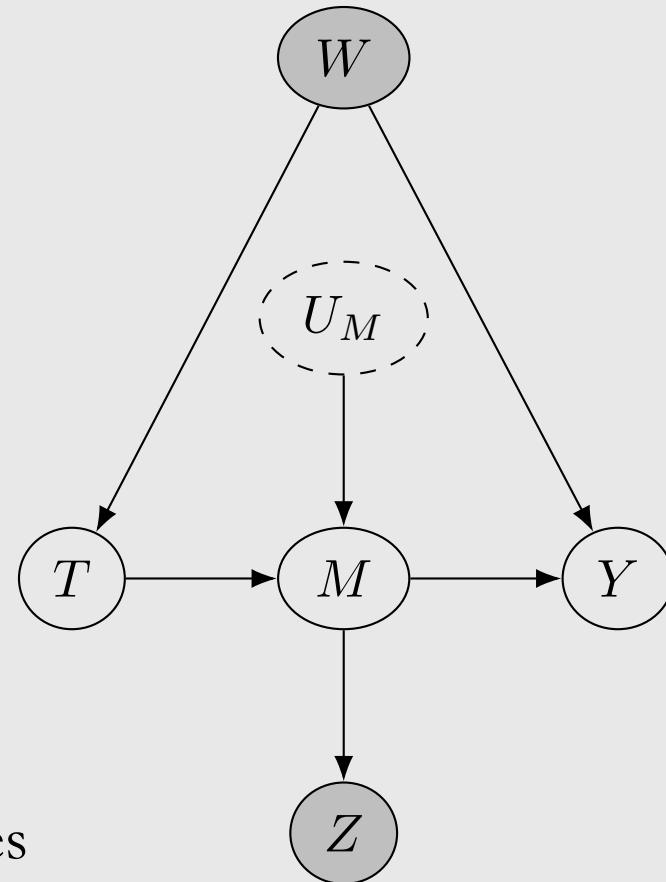
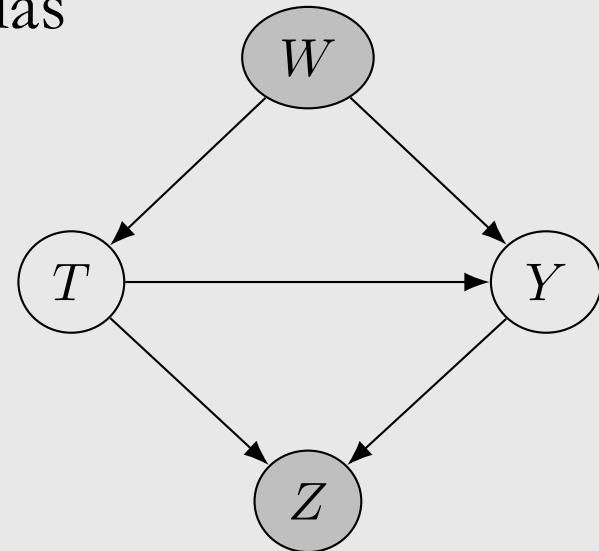
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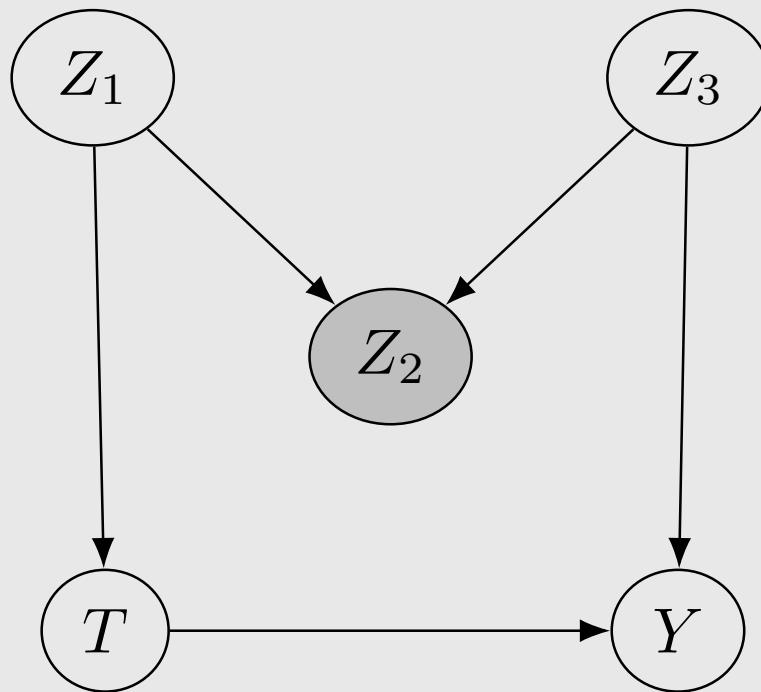
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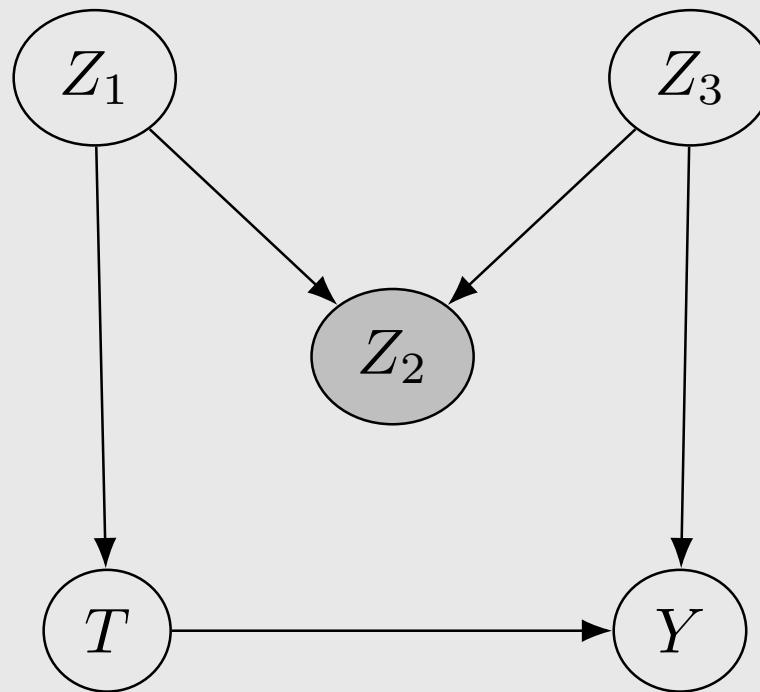


Rule: don't condition on post-treatment covariates

Inducing new pretreatment association (M-bias)



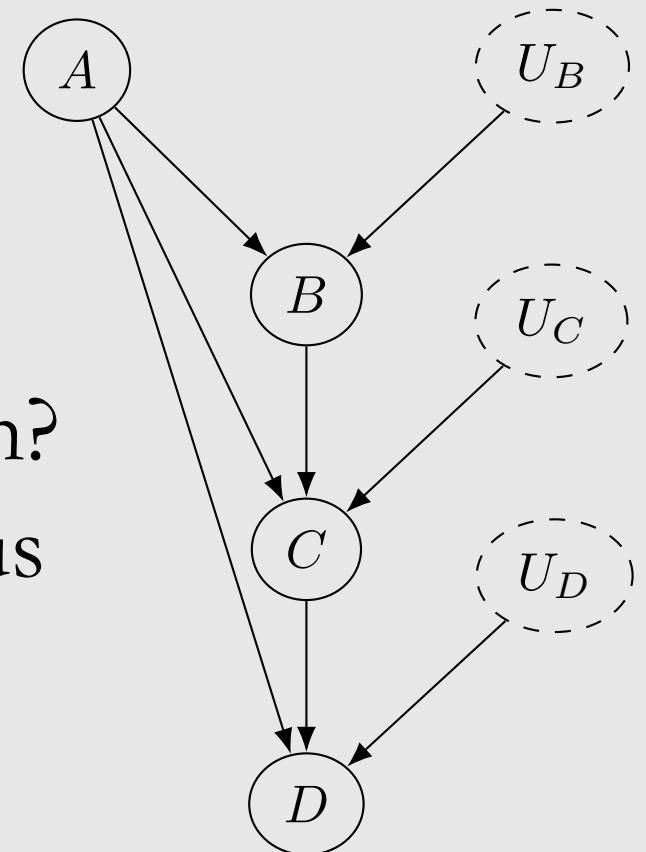
Inducing new pretreatment association (M-bias)



See [Elwert & Winship \(2014\)](#) for many real examples of collider bias

Questions:

1. What are the nonparametric structural equations for this causal graph?
2. What are the endogenous and exogenous variables in this causal graph?
3. What is collider bias?



The *do*-operator

Main assumption: modularity

Backdoor adjustment

Structural causal models

A complete example with estimation

Problem: effect of sodium intake on blood pressure

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Motivation: 46% of Americans have high blood pressure and high blood pressure is associated with increased risk of mortality

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Data:

- Epidemiological example taken from Luque-Fernandez et al. (2018)

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- Outcome Y: (systolic) blood pressure (continuous)

Problem: effect of sodium intake on blood pressure

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Data:

- Epidemiological example taken from Luque-Fernandez et al. (2018)
- Outcome Y: (systolic) blood pressure (continuous)
- Treatment T: sodium intake (1 if above 3.5 mg and 0 if below)

Problem: effect of sodium intake on blood pressure

Motivation: 46% of Americans have high blood pressure and high blood pressure is associated with increased risk of mortality

Data:

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- Outcome Y: (systolic) blood pressure (continuous)
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- Covariates
 - W age
 - Z amount of protein excreted in urine

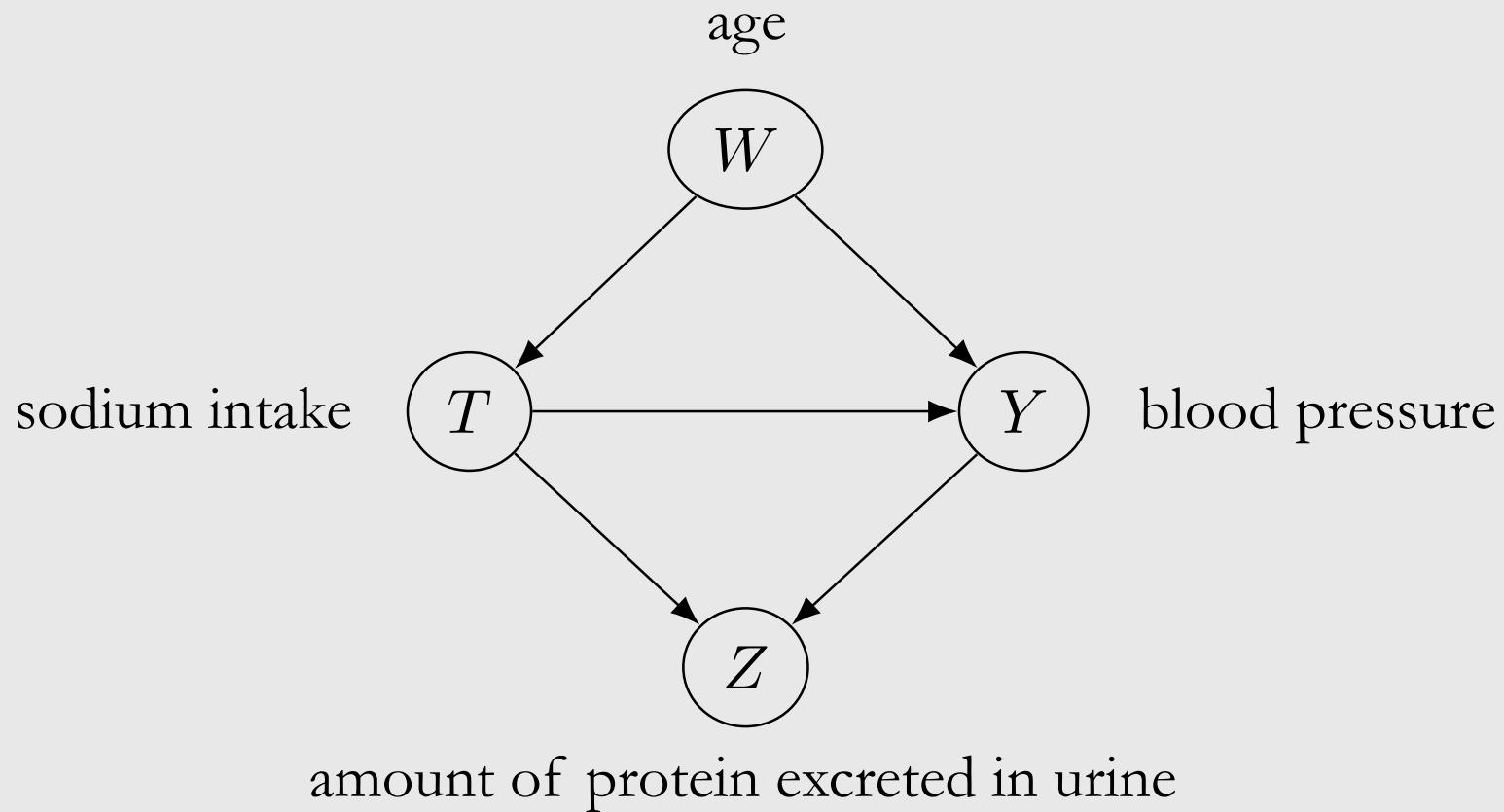
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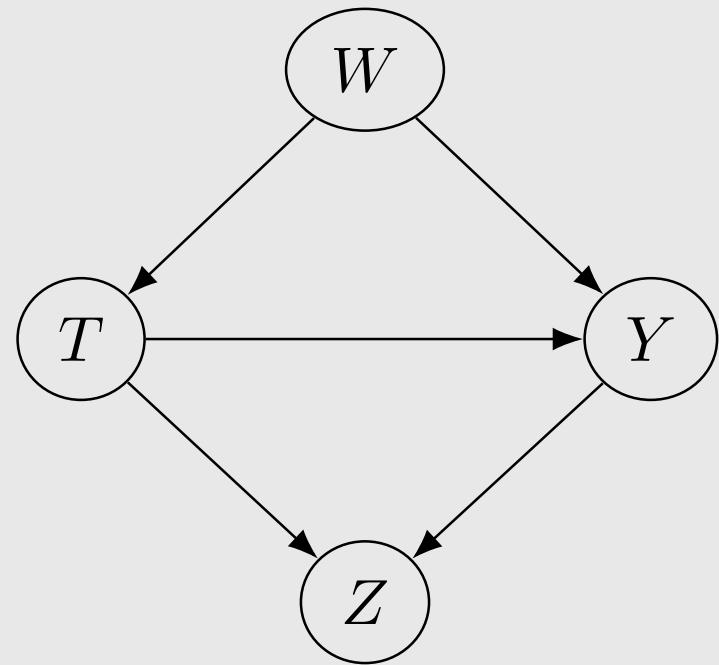
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 - W age
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- Simulation: so we know the “true” ATE is 1.05

The causal graph

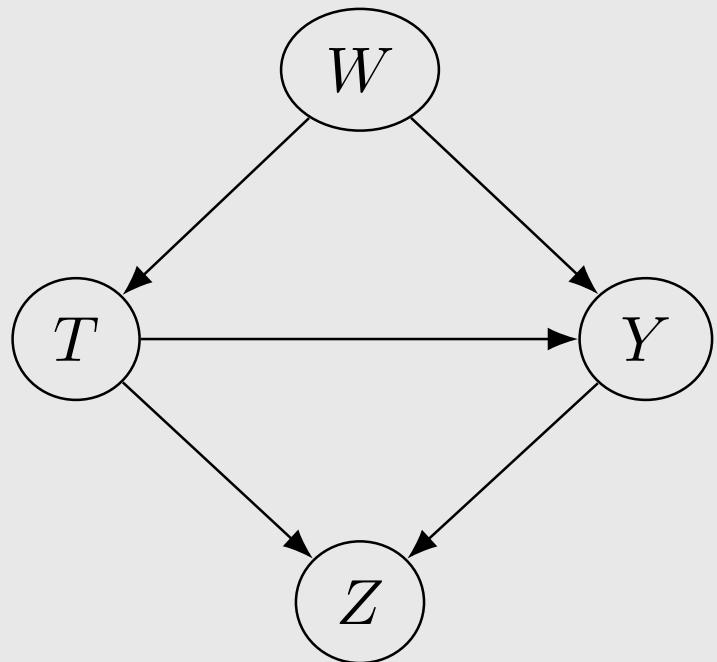


Identification

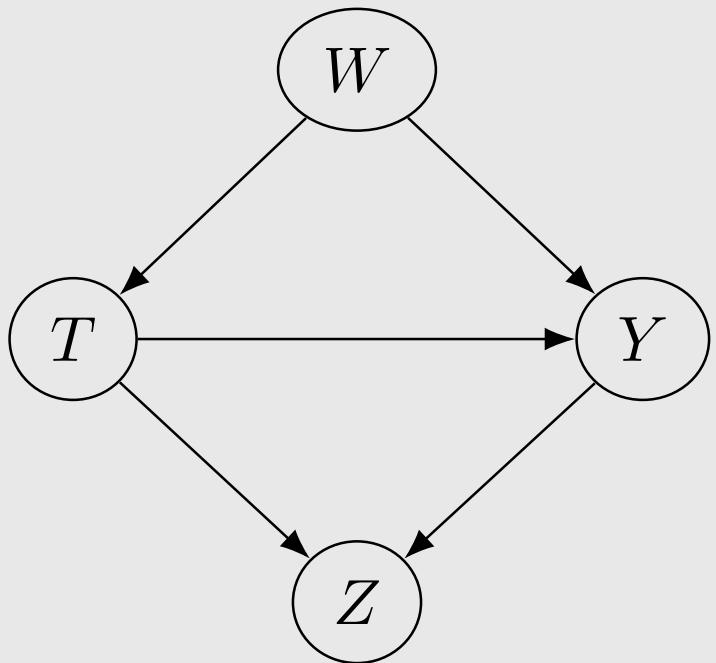


Identification

Causal estimand: $\mathbb{E}[Y \mid do(t)]$



Identification



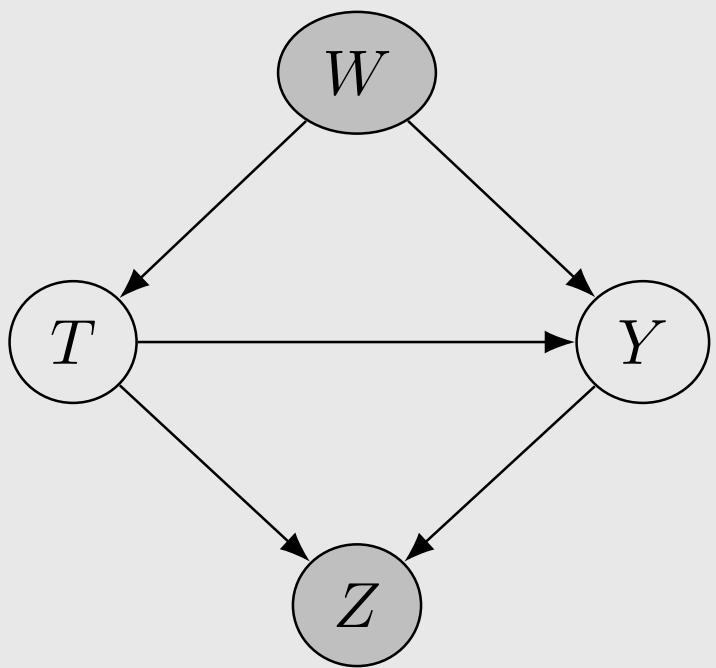
Causal estimand:

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Statistical estimand
from last week:

$$\mathbb{E}_{W,Z}\mathbb{E}[Y \mid t, W, Z]$$

Identification



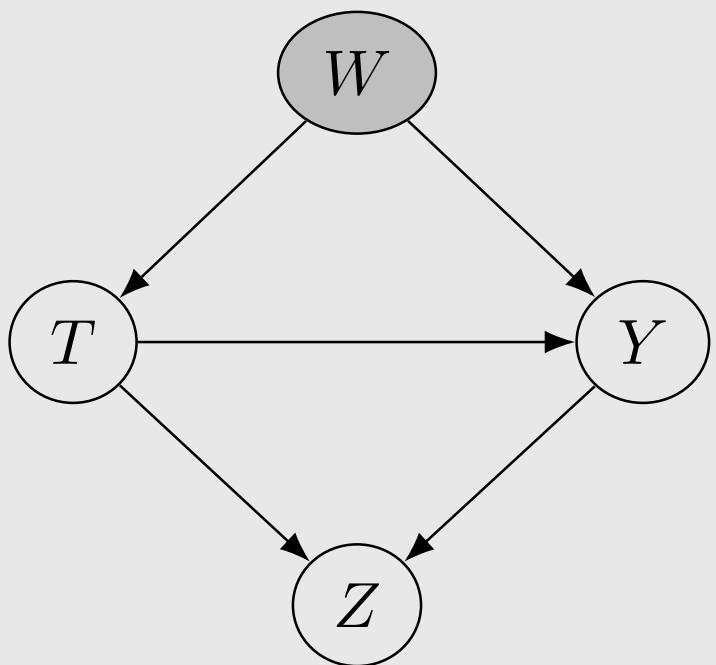
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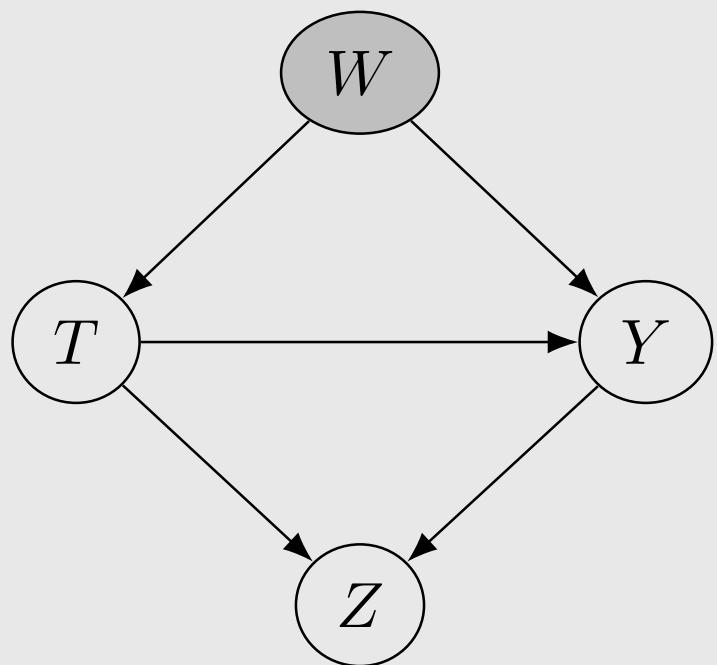
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Statistical estimand
from causal graph:

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Estimation of ATE

True ATE: $\mathbb{E}[Y(1) - Y(0)] = 1.05$

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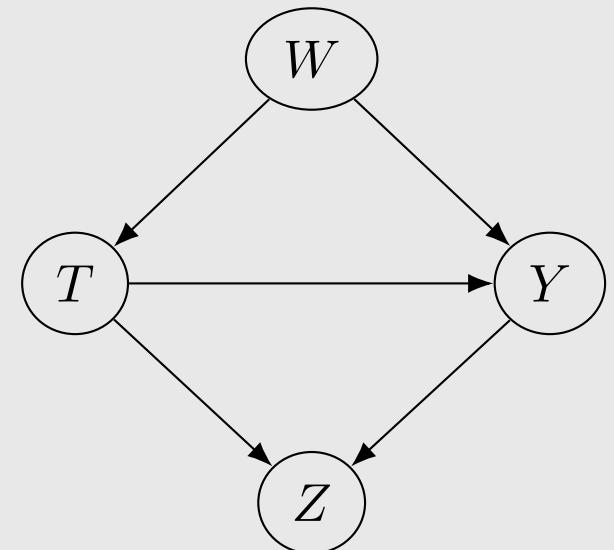
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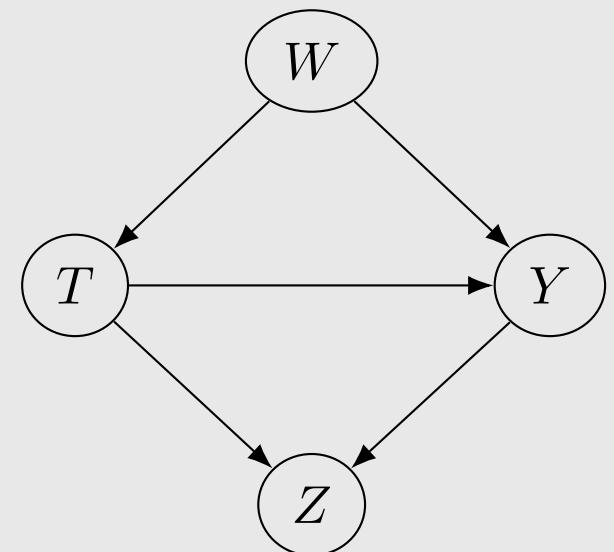
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Estimates:

$X = \{\}$ (naive): 5.33



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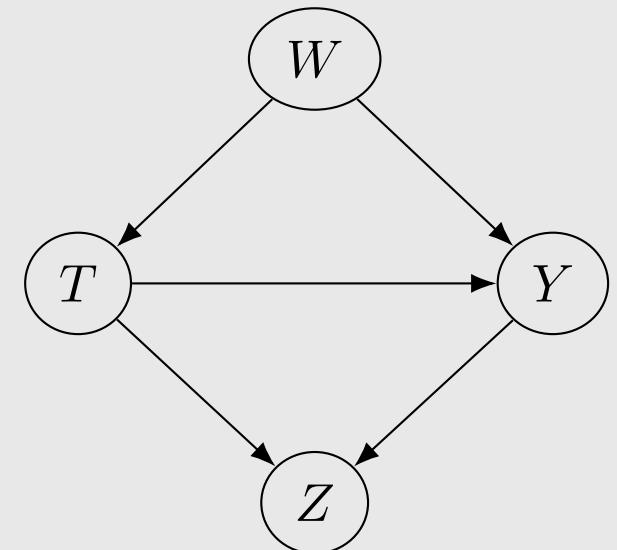
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$$\frac{|5.33 - 1.05|}{1.05} \times 100\% = 407\% \text{ error}$$



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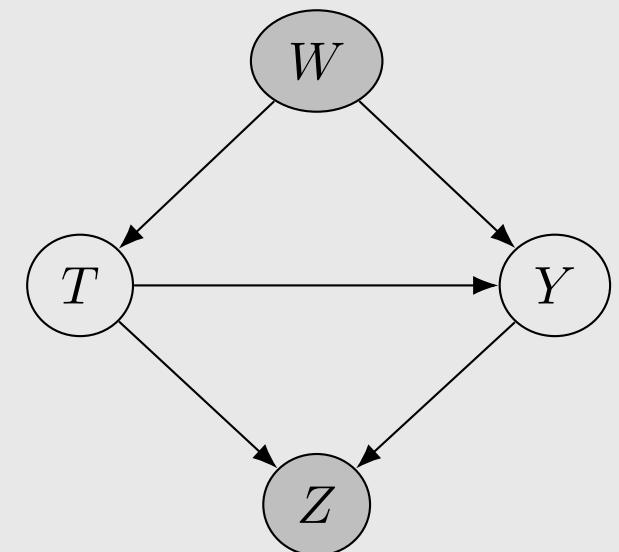
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$$X = \{\} \text{ (naive): } 5.33 \quad \frac{|5.33 - 1.05|}{1.05} \times 100\% = 407\% \text{ error}$$

$$X = \{W, Z\} \text{ (last week): } 0.85$$



Estimation of ATE

True ATE: $\mathbb{E}[Y(1) - Y(0)] = 1.05$

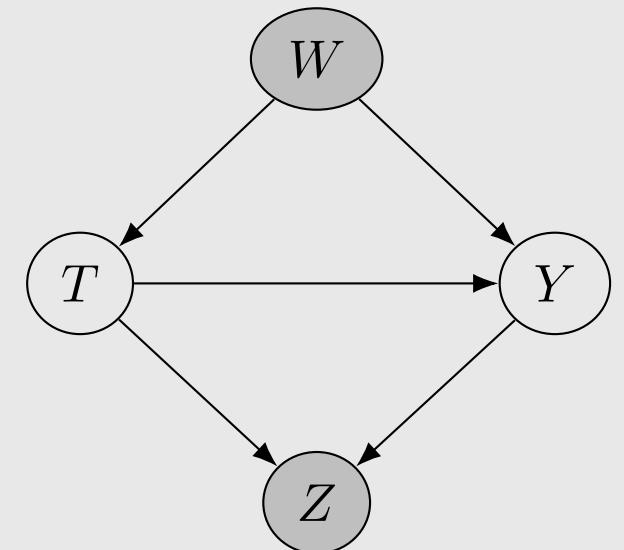
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$$X = \{\} \text{ (naive): } 5.33 \quad \frac{|5.33 - 1.05|}{1.05} \times 100\% = 407\% \text{ error}$$

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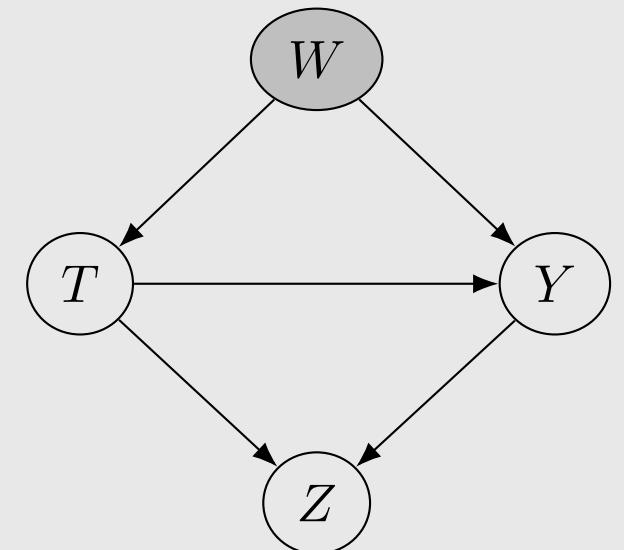
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Estimates:

$$X = \{\} \text{ (naive): } 5.33 \quad \frac{|5.33 - 1.05|}{1.05} \times 100\% = 407\% \text{ error}$$

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$$X = \{W\} \text{ (unbiased): } 1.0502$$



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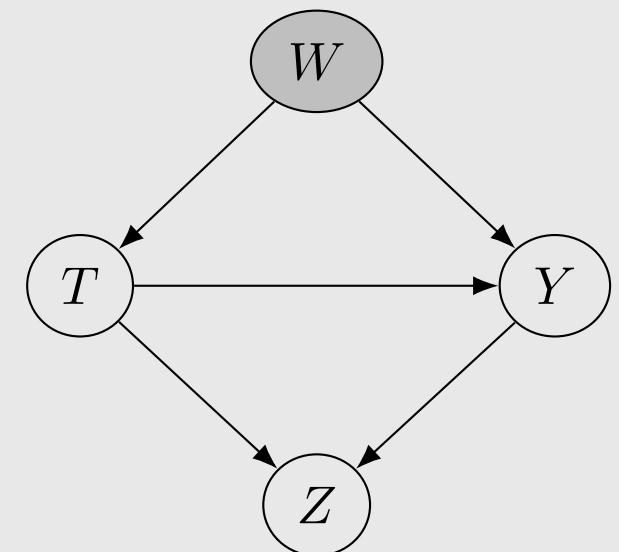
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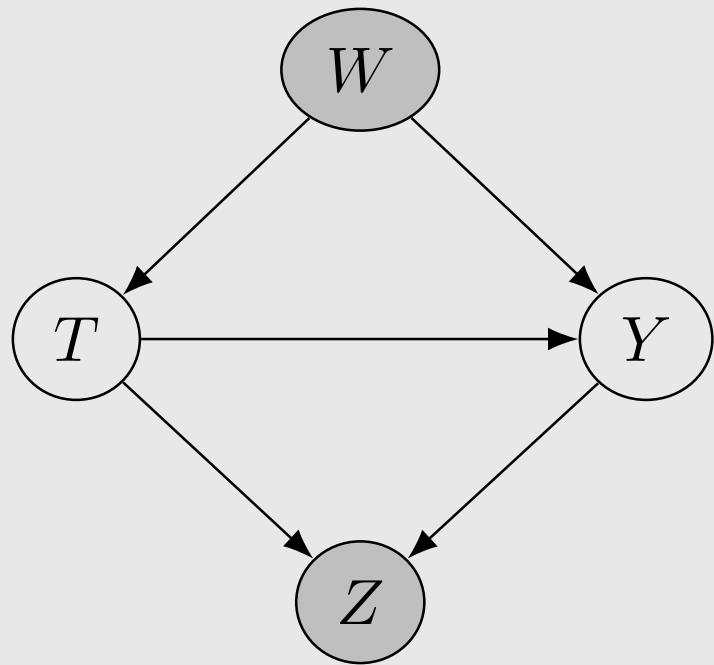
$$X = \{\} \text{ (naive): } 5.33 \quad \frac{|5.33 - 1.05|}{1.05} \times 100\% = 407\% \text{ error}$$

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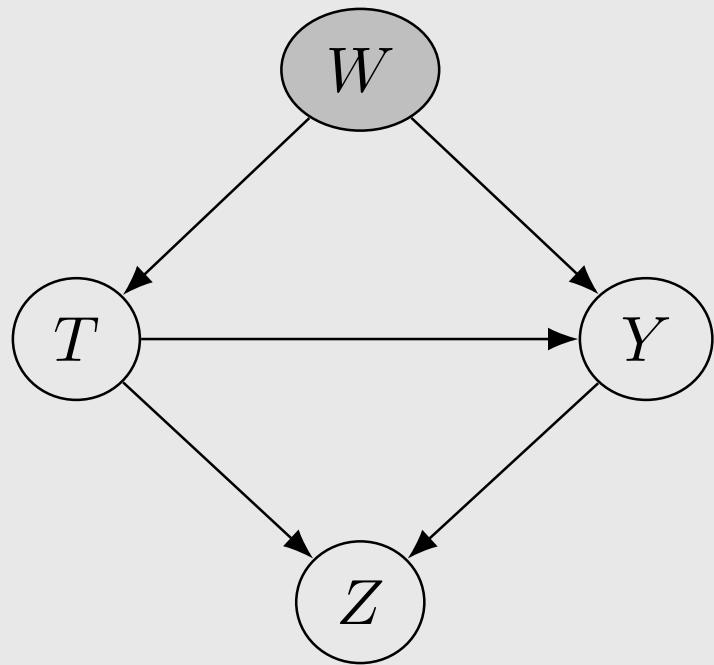
$$X = \{W\} \text{ (unbiased): } 1.0502 \quad 0.02\% \text{ error}$$



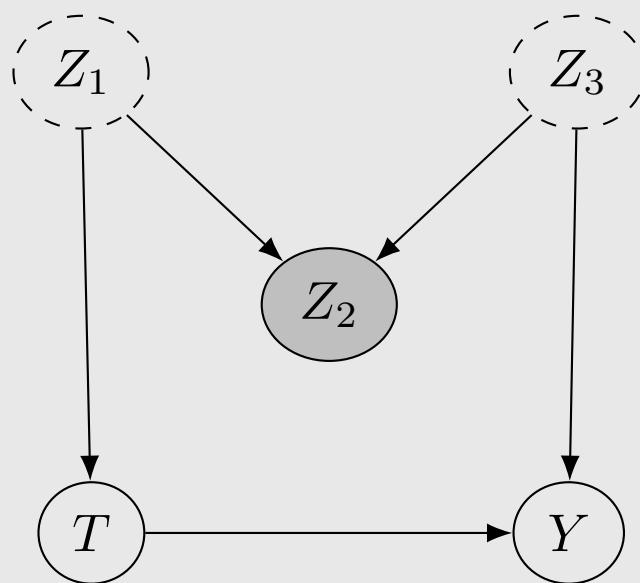
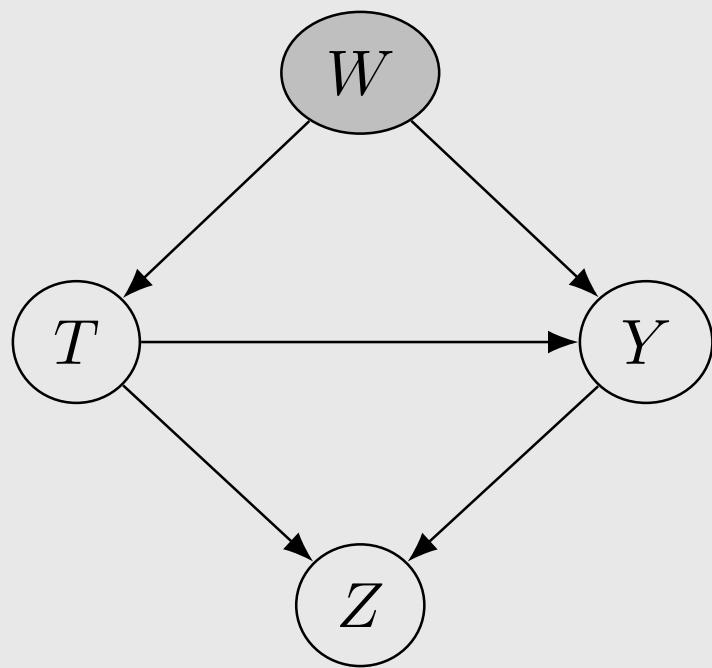
M-bias



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