Uncertainty In Cognitive Science

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# Abstract

Trusting papers introducing models without formally testing them is like trusting a mathematician to measure a 1m stick because of a phd in mathematics or physics instead of having him measure the stick and compare it to the 1m stick.

# Acknowledgement

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# Introduction

Most scientific inquiry revolves around measurements of the physical world; may that be the time it takes for a cup to fall to the ground or the time it takes for a person to react to a visual stimulus on a computer screen. There will be uncertainty associated with such measurements as repeatedly measuring the same thing, will result in different measurement. This means that one of the most basic things that scientific inquiry and theories rests on is bounded on uncertainty and it is one of the roles of science to quantify and account for such uncertainties. In this thesis I will investigate shortcomings in uncertainty handling in cognitive science, while providing ways to properly account for these uncertainties when analyses are conducted. To do this I will rely on Monte Carlo simulations which provides a robust method for accounting for uncertainties when analyses and models are deployed. In particular the thesis will introduce a partially novel way of testing and validating the parameters of models in cognitive science. It will be demonstrated that using these simulation based methods, one can achieve a more accurate representation of the probabilities of rejecting a null hypothesis given the assumptions of an experiment. The thesis will demonstrate this with a focus on the psychometric function. It will be shown that the parameters of the model and their uncertainties can be used reduced by several different interventions and assumptions including optimization of the experimental design, but also by incorporating additional information already available in most experiments like reaction times.

To further reiterate the utility of jointly modeling additional variables the thesis will re-analyze published data using a psychometric function. Here it will also be demonstrated how carefully using the known structure of the data can greatly improve session by session correlation between parameters, i.e. again minimizing uncertainty. Lastly using this reanalysis, the thesis will highlight opportunities to conduct power analyses utilizing a novel modeling framework that accounts for the uncertainty in model parameters as well as sampling variability of the effect investigated. Comparison will be made to popular tools such as G\*power highlighting the need for more rigorous methods, when conducting power analyses.

## *Uncertainties in science and examples from physics*

Science is a systematic way to organize knowledge in hierarchies, leading to testable hypotheses. Knowledge can be hard to define, but most often it is something that is achieved though experience. Imaging a cup being dropped, one will have the knowledge that it will fall towards the ground and reach our foot at a particular speed, because of our previous experiences with dropping a cup. This is to say that knowledge is the relationships that we believe to be true with differing amounts of certainty. The reality is that even though we might say that we are completely certain of events, i.e. know, that the cup will fall towards the ground and reach it at a particular speed. This is still an assumption that is true most of the time, but given that the natural world is bounded on probabilities, complete certainty is unwarranted, both in the assumption of the cup hitting our foot, but especially the speed at which it hits our foot. Most of the time this probabilistic nature of the natural world stems from the uncertainties during measurement or perhaps unseen events.The interest here is not in the unseen events but instead in the predictability and (un)certainty of the expected. Taking the falling cup as an example, we would normally not be interested in the probability that the cup will hit out foot, but instead in the acceleration of the cup and the uncertainty in this estimate. What scientists have shown is that objects dropped on earth will accelerate towards the ground with an acceleration of ([Johannes & Smilde, 2009](#ref-johannes_fundamentals_2009)). However, this number does not mean anything without an estimate of the uncertainty, while also accounting for the assumptions that are entailed with these numbers. The first proposition is well studied and the 95% confidence interval of the value of ([Johannes & Smilde, 2009](#ref-johannes_fundamentals_2009)). The second proposition is also quite well studied as we know that the density of the medium that the cup is dropped in, is important but also the shape and weight of the cup if dropped outside a vacuum. In order to estimate this constant acceleration, measurements have to be made of the distance a falling cup travels and the time it takes to reach the ground. With these measures of distance and time, uncertainty is introduced and propagated to get the an estimate for the acceleration, but also the uncertainty associated with it.

There are 2 main points of the example which this thesis will explore, firstly uncertainties are organized in hierarchies and are just as important as beliefs as without one, the other is meaningless. Secondly, taking these uncertainties seriously and herein estimating and propagating them should not be a choice or something that can be avoided, but a necessity of all scientific endeavors, where they can be quantified or at least approximated. Taking its outset in the published literature, the thesis will establish issues regarding uncertainty handling and propagation and use simulations to highlight the problems with neglecting a proper account of uncertainty in statistical models. After highlighting these potential issues, the thesis will provide possible ways of dealing with the shortcomings by means of simulations. The goal of this thesis is to illuminate the often-overlooked uncertainties in the data collected on human behavior and cognition while providing ways of accounting for it, such that the uncertainty reported in the published literature more accurately reflects the (un)certainty we should have in the results. The thesis will argue that accounting for uncertainties is more important than ever, especially in research of complex systems such as humans as computational resources have made it possible to easily develop more sophisticated analyses and models that have dependencies on lower-level analyses. The dependent structure makes the need for proper uncertainty handling even more imperative, as without propagating uncertainties the resulting estimate will be overly confident. Furthermore, these computational resources do allow for uncertainty propagation without understanding the underlying mathematics, making it accessible to most researchers with some coding experience. To effectively communicate both the statistical models as well as the underlying uncertainties associated with doing computations on data, the thesis will start by exploring different types of uncertainty in cognitive science. Next the thesis will be investigating a particular cognitive model used in many subfields of cognitive science and examine how validation of such cognitive models have been done and how proper uncertainty handling can improve these validation steps.

## *Levels of uncertainty and uncertainty propagation*

I will here broadly define 3 different types of uncertainty, measurement, estimation, and test-retest reliability uncertainty see figure 1 and 2 for a visualization. These definitions are not exhaustive and will be centered around how experimental studies in cognitive science are conducted, from data collection to data analysis. The first aspect of uncertainty is to acknowledge that uncertainties can be defined in hierarchies and that uncertainty propagates through these hierarchies. This uncertainty propagation means that as you do calculations based on measures with uncertainty, the uncertainty propagates to the results of the calculations. In this thesis I will be using simulations to show how uncertainty propagation can be understood and handled without a need for rigorous mathematical proofs. For a more mathematical treatment see ([Saccenti et al., 2020](#ref-saccenti_corruption_2020)).

The lowest level of uncertainty is in the measurements themselves i.e. measurement uncertainty. Measurement uncertainty reflects the uncertainty in how well one can for instance measure the reaction time on a computer or the time it took the falling cup to reach the ground. This level of uncertainty is often neglected in cognitive science when applying statistical models, because they are thought to be minuscule as in the case of reaction time tasks, which may or may not be true given the experiment setup ([Crocetta & Andrade, 2015](#ref-crocetta_problem_2015); [Holden et al., 2019](#ref-holden_accuracy_2019); [Ohyanagi & Sengoku, 2010](#ref-ohyanagi_solution_2010)). This is not to say that cognitive scientists do not care about them, as moving towards more sophisticated measurement methods is an ongoing endeavor. For instance, using better and more sophisticated computers to measure reaction times commonly found in cognitive science experiments would decrease the uncertainty in the measurements themselves ([Crocetta & Andrade, 2015](#ref-crocetta_problem_2015)). Minimizing this kind of uncertainty most often revolves around getting better tools to measure the variable(s) of interest. However, there might also be other avenues where a more explicit quantification of measurement uncertainty might be appropriate. One of the main instances coming to mind is the use of questionnaires. Many of these questionnaires try to quantify a latent construct such as mental health conditions i.e. depression or anxiety. Many of the main questionnaires used to assess depression, anxiety and stress use several questions that are then added together to give a score of the mental health condition without a quantification of the uncertainty on the latent construct because no uncertainty is quantified on each question ([Cohen, 1994](#ref-cohen_perceived_1994); [Johnson et al., 2019](#ref-johnson_psychometric_2019); [Kroenke et al., 2001](#ref-kroenke_phq-9_2001); [Xiao et al., 2023](#ref-xiao_psychometric_2023)).

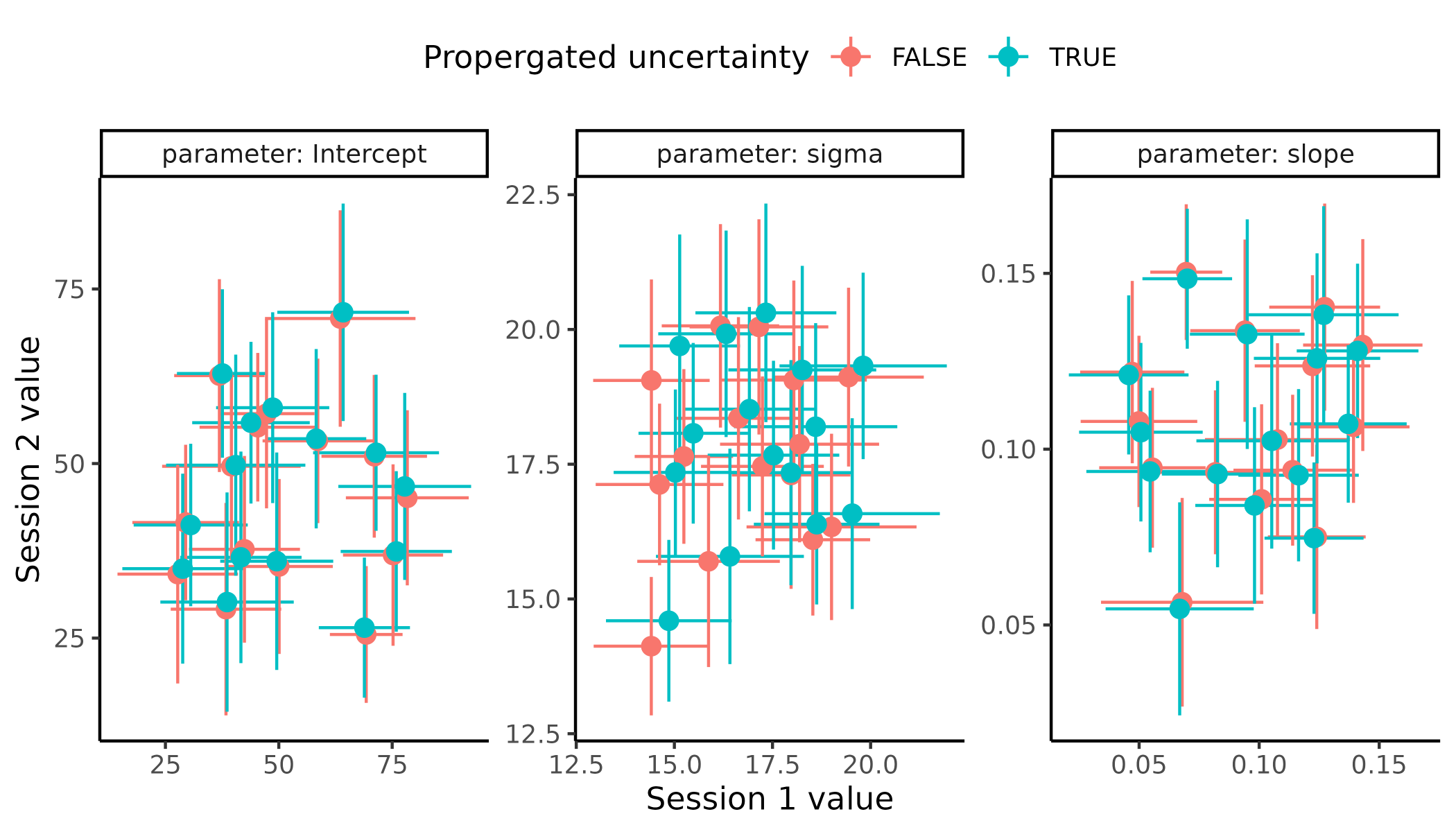
In order to contextualize these forms of uncertainty i will consider an example of a researcher wanting to understand the relationship between reaction times and stress. To do this the researcher conducts an experiment where participants are measured several times under different conditions to introduce stress. A questionnaire is used to access the stress of the participant and a computer based experiment is used to determine their reaction time. In this example both of these measures have associated uncertainty, measurement uncertainty, see individual datapoints in figure 1.



**Figure 1 Measurement and Estimation uncertainty;** the figure displays a linear regression between two measurements for instance reaction time and stress with measurement uncertainty depicted as vertical and horizontal error bars on individual points. The mean of the regression line with and without propagated uncertainty is highlighted in grey and dark green respectively. Lastly a prediction interval is depicted as the shaded area around the mean of the regression line with and without propagated uncertainty again in grey and green respectively.

The next level of uncertainty is when a particular model is fit to some data, or more broadly when calculations are done on data with uncertainty. In cognitive science we achieve parameter estimates from our collected data and these parameter estimates have uncertainty associated with them, this uncertainty will be referred to as estimation uncertainty. Estimation uncertainty is most often quantified by the statistical model be that the standard error of a regression coefficient or the width of a posterior distribution of a parameter in a Bayesian framework. Taking the example of the researcher investigating reaction times and stress this is the uncertainty in the parameter estimates achieved by fitting a linear model to the data see figure 1. Minimizing this estimation uncertainty is what most scientists care about, as inevitably most cognitive science experiments revolve around null hypothesis testing, which in most cases will involve testing whether the parameter estimate includes a particular value mostly, 0. To minimize this type of uncertainty the standard approach is to get more data, given they are from the same population and behave similarly. In cognitive science this might include more trials or subjects to get a more precise estimate of interest i.e. minimizing estimation uncertainty. In cognitive science the minimizing of estimation uncertainty is however not free or free of uncertainty itself. Firstly, increasing the number of trials in a cognitive task might even increase the estimation uncertainty itself. This can happen for several reasons, but boredom, habituation, fatigue and lack of engagement can become big contributors when experimental tasks become very long ([Jeong et al., 2023](#ref-jeong_exhaustive_2023); [Meier et al., 2024](#ref-meier_is_2024)). Secondly for some cognitive science experiments massively increasing the number of trials could make subjects more prone to switching between cognitive strategies and if not properly accounted for in the analysis might be interpreted as additional noise by the model. Next increasing the number of subjects included in a study will many times decrease estimation uncertainty on the population level estimates, if the sample population is homogeneous. The trade off between subjects and trials in an experiment is therefore quite important to minimize estimation uncertainty, but also minimize the overuse of resources. However, there are many times also other ways to minimize estimation uncertainty ([Baldi Antognini et al., 2023](#ref-baldi_antognini_new_2023); [Stone, 2014](#ref-stone_using_2014)). For instance changing the task design such that responses will give more information on parameter values of interest. This optimization strategy involves individualizing the task design such that each presented stimulus is the most informative. This task design optimization is frequently used in psychophysical experiments where adaptive algorithms are used to select the upcoming stimuli such that it minimizes the uncertainty in the estimated parameter values. See for example algorithms like PSI, QUEST and ADOPY ([Prins, 2013](#ref-prins_psi-marginal_2013); [Watson, 2017](#ref-watson_quest_2017); [Yang et al., 2021](#ref-yang_adopy_2021)). From the example with reaction time and stress this might invovle selecting interventions that will produce varying levels of stress to increase the precision of the parameters of the model. A practical example of how these algorithms work to minimize uncertainty in the parameter estimates see the section about Adaptive design optimization.

The next level of uncertainty stems from the fact that these parameter estimates will vary over time, as humans vary over time. This variation stems from both behavioral factors like learning, but also psychological factors such as mood and arousal ([Schurr et al., 2024](#ref-schurr_dynamic_2024)). This type of uncertainty will be referred to as test-retest uncertainty. Again with offset in the example, participants in the researchers study on reaction times and stress might be tested twice on different days to understand how stable the relationship is over time. As the relationship is measured by the parameters of the model the stability of the relationship is measured by the stability of the parameters. One might imagine that the amount of sleep acquired before the experimental day could influence both measures of the task i.e. reaction time and suseptibility to stress and perhaps even their relationship. Figure 2 displays how the parameter estimates of the same model, Figure 1 with and without accounting for uncertainty propagation. As can be seen from figure 1 accounting for the measurement uncertainty does not change much the prediction made by the model, however when propagating these extra uncertainties into the next analysis of the parameters from session to session in figure 2 the change in results become more pronounced. The main effect for the current linear model is that the residual variance is underestimated without error propagation and the slope parameter is overestimated.



**Figure 2 Test retest uncertainty;** displays the results of fitting the linear regression in Figure 1 twice with and without accounting for measurement uncertainty. Each facet represents one of the three parameters of the linear model, the intercept the residual uncertainty and the slope respectively from left to right. Colors represented weather the measurement uncertainty was proporgated or not.

The main message here is that to get reliable estimates and, in the end, to make reliable inference one needs to account for all these sorts of uncertainties and the lower in the hierarchy you move the more fundamental and important they become. Having a parameter estimate that is stable over time won’t matter if you cannot estimate it or measure it reliably in the first place.

## *Investigating measurement uncertainty*

To keep a consistent theme, I will throughout the thesis be demonstrating how computational resources have made the need for analytic solutions involving tedious assumptions sometimes irrelevant. This is highly relevant as closed-form-problems where an analytic solution is known or even attainable are becoming less and less frequent with the surge in popularity of more and more complex models, see section about modeling definitions for further elaboration. In order to explore measurement uncertainty in examples related to cognitive science the thesis will here investigate the relationship between correlation coefficients and measurement uncertainty. This will be done, as a non trivial part of the published litterature in cognitive science revovles around conducting correlational analyses on measures that have quantifiable uncertainties such as estimated parameters or even structural properties of the brain like the myelination or grey matter volume in a region of interest ([Berker et al., 2016](#ref-de_berker_computations_2016); [Luijcks et al., 2015](#ref-luijcks_influence_2015); [J. Wu et al., 2021](#ref-wu_neurobiological_2021)).

In this section I will demonstrate how using simulations to both understand and explore how adding measurement uncertainty will change the strength of interpretation of doing correlational analyses this will serve as an abstract representation of how correlation coefficient estimate change under different sizes of measurement uncertainty. In order to use simulations to include measurement uncertainty firstly an understanding of the uncertainty of the correlation coefficient itself is needed. Analytical solutions exist to calculate the uncertainty of such statistics, which is incorporated in most statistical softwares ([Makowski et al., 2023](#ref-R-correlation)), however another way to find and understand this uncertainty comes from resampling. Below is an explanation of using resampling to evalute uncertainty in a general case and thereafter its implementation for adding measurement uncertainty.

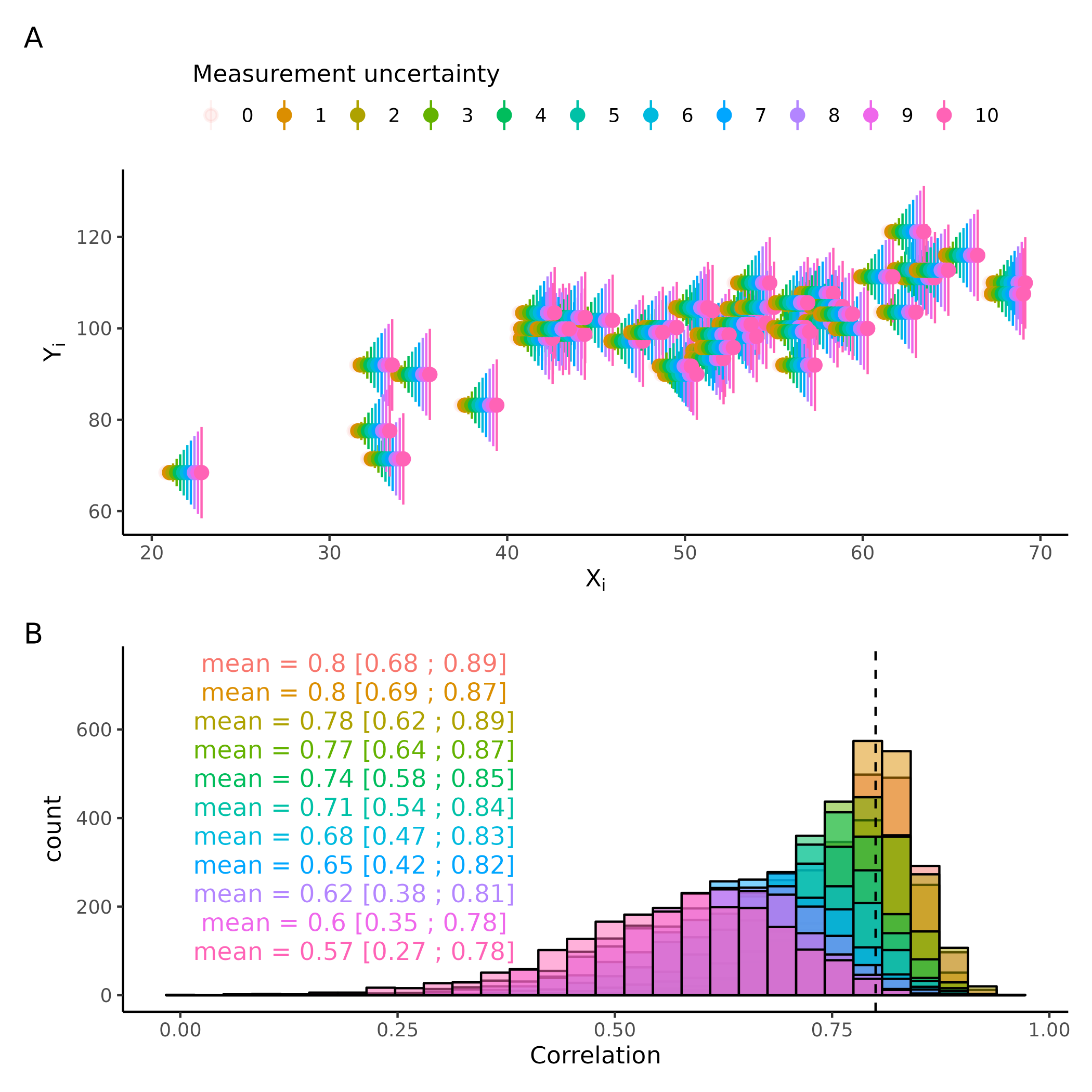
The way to estimate the uncertainty in the correlation coefficient is to re sample the collected data with replacement i.e. bootstrapping and then recalculate the test statistic of interest ([Efron, 1983](#ref-efron_estimating_1983)). Iterating this process gives a distribution of test statistics which with enough iterations will converge towards the analytic solution with recommendations of having at least 30 data points to begin with ([Efron, 1983](#ref-efron_estimating_1983); [Efron & Tibshirani, 1994](#ref-efron_introduction_1994); [C. F. J. Wu, 1986](#ref-wu_jackknife_1986)). For the simplest case of recalculating the correlation coefficient and its uncertainty might seem somewhat tedious compared to taking the direct analytic solution, as this is already implemented in most statistical softwares and packages ([Makowski et al., 2022](#ref-correlationPackage)), however once setup and understood this approach allows for adding not only measurement uncertainty, but a more general way of thinking about the uncertainty of statistical metrics. One of the advantages of having an analytic solution to this simple case of recalculating the uncertainty of the correlation coefficient is to ensure that the code and scripts are properly set up. This therefore serves as a validation step before exploring territories where analytic solutions are scarce or nonexistent.

The first step is therefore to show that the two approaches of simulating and analytically estimating the uncertainty of the correlation coefficient is identical across different ranges of correlations and sample sizes. To do this, I’ve simulated data from a multivariate normal distribution with the following parameters.

Where

The multinormal distribution produces random variables with a means a standard deviation and crucially with a correlation coefficient between all random variables , here the subscript indicates that there are x random variables being sampled together. This distribution is perfect for understanding how the correlation coefficient changes as it is a parameter of the distribution. Now demonstrating that bootstrapping and the analytic solution implemented in R are identical, I simulate correlation coefficients ranging from -0.9 to 0.9 in increments of 0.1 with the total number of samples per random variable being between 50 and 500 in increments of 50 ([Makowski et al., 2022](#ref-correlationPackage); [R Core Team, 2024](#ref-R2024)). See supplementary material and supplementary Figure 1 for demonstration of the simularity of these two approaches.

Having shown that the two approaches are identical (or close to) we can add measurement uncertainty to each observation. This has again been analytically solved and solutions exist to calculate the correlation coefficient under these circumstances ([Saccenti et al., 2020](#ref-saccenti_corruption_2020)). To add measurement uncertainty to the measurements we can instead of randomly re sampling pairs of data points from the original data, as done for the simplest case above, one re samples these pairs as means of an error distribution where the uncertainty (standard deviation) of this distribution is the measurement uncertainty. A mindless choice of error distribution would be the normal distribution which would reflect the fact that the directionality of the uncertainty is assumed to be bidirectional i.e. with no preferred direction. Of note here is that one might re sample the original data from other error distributions for instance if values are strictly positive or bounded in other ways then simulating from a truncated normal or strictly positive distributions like a lognormal, would be preferred to avoid sampling values that cannot be obtained i.e. negative reaction time values. For this demonstration of adding measurement uncertainties to observed data, normally distributed noise is simulated, which means simulating new “observed values” from a normal distribution with a mean of the observed observation and a standard deviation equal to the measurement uncertainty. This can be seen in Figure 3, here uncertainty is added to just the x values in increasing amounts (A), with the resulting correlation coefficient distribution obtained by bootstrapping displayed in (B). It should be noted that the correlation coefficient simulated in this case was 0.8, as indicated by the vertical line in figure 3 (B). it is clear the estimated correlation coefficient using bootstrapping is being attenuated in size but also that the width of the correlation coefficient distribution is increasing with increasing measurement uncertainty, mimicking what can be shown using the analytical solutions ([Saccenti et al., 2020](#ref-saccenti_corruption_2020)).



**Figure 3 Measurement uncertainty on correlation coefficient** (A) Displays a scatterplot with varying amounts of measurement uncertainty. (B) displays how the correlation coefficient distribution, obtained through bootstrapping changes with increasing measurement uncertainty. Vertical line is the simulated correlation coeficient without uncertainty.

## *Modeling definitions*

This thesis will revolve around building, refining, testing, and designing models of cognition. To do this cognitive modelling will be deployed. Here cognitive modelling is meant as an intermediate level in a hierarchy of computational models on top, and statistical models in the bottom. The distinction between these concepts can be found in their flexibility, assumptions, and scope of investigation. It should be noted that all these types of models have many things in common such as being mathematical representations of a data generating process and that these are working definitions with fuzzy boundaries ([Durstewitz et al., 2016](#ref-durstewitz_computational_2016)).

*Statistical models* are the models primarily used in medical, social, and educational sciences, these models mostly consist of linear and generalized linear (mixed) models. What these models have in common is that they are linear combinations of independent variables which are sometimes transformed (making them generalized) to a particular domain such that this linear combination maps to a dependent variable. The mathematical representation of such models are as follows:

Where y is a vector of dependent variables of N elements, F is a link function that maps the conditional mean unto a particular space, common link function are the logit and log transformations which maps unto domains of [0 ; 1] and [0 ; ∞] respectively, which makes predictions on probabilities and strictly positive values like reaction times possible. is a vector of regression coefficients of P predictors which gets estimated, X Is a matrix of independent variables of size [N, P]. Lastly is a vector of N elements containing the errors of the model predictions on the dependent variables. The benefit of these regression models is that maximum likelihood estimators are available meaning that parameters estimates can be calculated using a frequentists statistical framework, making the estimation process fast and efficient. However, the downfall of these models is that they put quite big constraints on the types of models that can be fit, i.e. there must be a linear mapping between all independent variable and the dependent variable in a domain that can be mapped with a link function. This constraint will in many instances make theories hard or impossible to test as human behavior and cognition is highly nonlinear in many ways ([Ivanova et al., 2022](#ref-ivanova_beyond_2022)). It should be noted that the correlation coefficient examined in the previous section, can be thought of as a special case of this linear model where is a single value and y and x are z-transformed vectors, see supplementary figure 2.

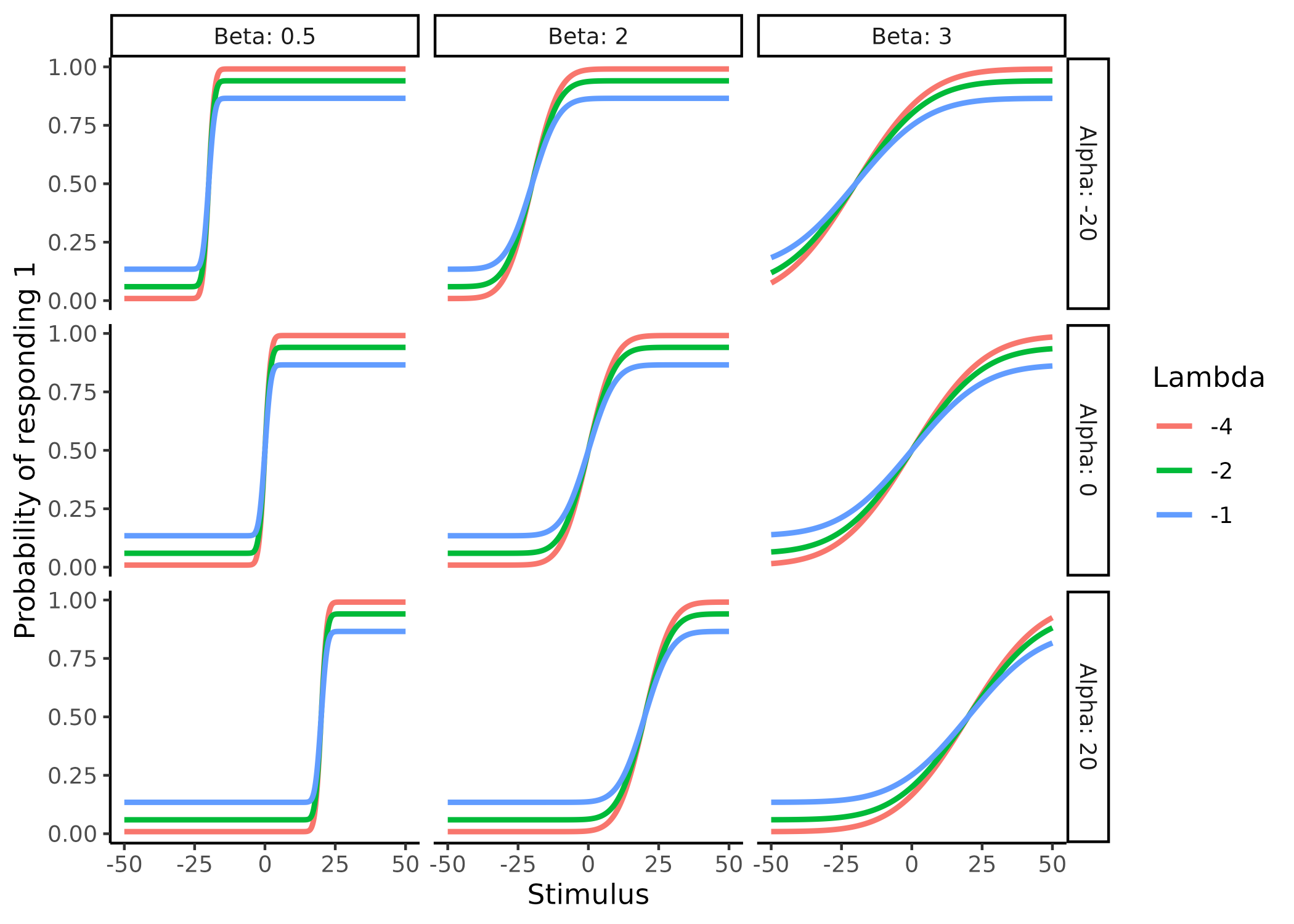
*Cognitive models* are models that are meant to resemble the generative processes of human behavior more closely. These models are generally more theoretically driven as the constraint of linear combinations is avoided, by employing different optimization schemes that sometimes use sampling algorithms to obtain the results. In many cases cognitive models are estimated in a Bayesian framework due to the flexibility with which models can be specified. The main advantage of these models, in this context, is the added freedom in model specification, but see discussion for other advantages.

*Computational models* are the upper most level of the hierarchy which here will be used to refer to the generalization of cognitive models to other scientific domains, such as physics, biology chemistry etc. These models are outside the scope of this thesis.

These three categories are arbitrary, and many methods and models will fall between them, with this vague definition. However many times these arbitrary definitions do add value in communicating what general framework we are working in and thereby what methods are used. The next section will describe a particular cognitive model which will be the the focal point for the rest of the thesis.

## *Model descriptions*

In this thesis the psychometric function (PF), will be investigated as this has been a stable corner stone in the cognitive science literature across many different sub fields ([Bahrami et al., 2012](#ref-bahrami_what_2012); [Coates & Chung, 2014](#ref-coates_changes_2014); [Courtin et al., 2023](#ref-courtin_spatial_2023); [Ma et al., 2024](#ref-ma_memorability_2024)). The psychometric function is a continuous function that maps real or positive inputs into probabilities, i.e. the domain is whereas the range is . In most cases the PF used is like a logistic regression in statistical modeling and is commonly used in perceptual research where the inputs are stimulus intensities, and the probabilities are then converted into binary forced choices through a Bernoulli or binomial distribution. The mapping of inputs to probabilities is usually done through a cumulative density function such as the cumulative logistic or normal distribution, which amounts to conducting a logistic or probit regression in the statistical framework. The main difference between the statistical and cognitive framework of the PF is the number of parameters. The least number of parameters used to describe the PF is 2 the threshold and the slope (,). These two parameters describe the center of the curve, with being the intensity of the stimulus at probability 0.5 and being the steepness of the function around this value. In the cognitive modeling framework one or two more parameters are typically introduced the lapse and guess rates (, ). These two parameters together handle the tails (i.e. the far ends) of the psychometric functions and essentially makes the probability in the two ends of the psychometric no deterministic i.e. the upper and lower bounds become 𝛾 and 𝜆 instead of 0 and 1, see figure 4. These parameters help with fitting the PF to data where sometimes attentional slips or wrong button presses happen and it can be shown that including these parameters will greatly improve the estimation of the slope of the PF if lapses and or guesses are present in the data ([Wichmann & Hill, 2001](#ref-wichmann_psychometric_2001)). This also makes intuitive sense as the function cannot predict deterministic (i.e. probabilities of 0 or 1) if there are responses at a high stimulus level which was caused by a lapse. Figure 4 depicts how all these parameters change the shape of the PF. For the sake of this thesis, I’ll be using the cumulative normal distribution to map stimulus values to probabilities with a single lapse rate. This single lapse rate will govern the distance between the upper and lower bound, essentially making it equally likely to have an erroneous response for high and low stimulus values. This mathematical formulation of the function is as follows:



**Figure 4 Psychometric parameters.** Displays how the parameters alpha (), beta () and lambda () of the psychometric fucntion changes its shape. Columns display how the beta parameters changes the slope of the function. Rows show how alpha changes the location of the center of the function changes. Lastly, colors in the plot depict how lambda changes the asympotes in extreme stimulus (x) values.

## *Model validation.*

In the same vein of validating the bootstrapping approach with the analytical solution in the previous section about measurement uncertainty, cognitive models themselves are many times validated to ensure that at least in principle the parameters of the model can be estimated with increasing accuracy with increasing number of trials. This section will highlight some of the emerging ways in which computational models in the cognitive science literature are being tested and validated and takes offset in the seminal paper from Wilson and Collins ([Wilson & Collins, 2019](#ref-wilson_ten_2019)) describing 10 simple rules of computational modeling, which is commonly cited when validation of computational models are described ([Hess et al., 2024](#ref-hess_bayesian_2024)).

There are at least three main challenges when building and validation cognitive models which are particularly relevant when writing novel models. How do we know that our models do what we think they do (identifiability). How do we know that they accurately estimate the parameters of interest (Internal validity)? And lastly how do we know that we can distinguish between competing models (external validity). The last challenge is beyond the scope of the current thesis and is well covered elsewhere ([Wilson & Collins, 2019](#ref-wilson_ten_2019)).

The answer to the first two challenges must be found in simulations when our models become more and more complex and analytical solutions are sparse. This simulation practice revolves around selecting an appropriate range of parameter and using these to simulate data from our models and then refitting the data to then see how well the model approximates the simulated parameter values. It should therefore come as no surprise that ensuring that in these simulations, we would like our models to perform well, such that we can have faith in them when the real underlying process is unknown, i.e. analyzing real world data. An appropriate range of parameter values for a particular model can be difficult to select as is exactly the problem of identifiability. However, several lines of information can help gauge this. Firstly, looking at mathematical constraints of the model formulations can reduce the possible ranges of parameter values. For the case of the psychometric function this amounts to ensuring that the slope is strictly positive as this ensures that increasing levels of stimuli (when x increases) will produce greater probabilities of responding 1, but also ensure that the standard deviation of the underlying probability density function is strictly positive. The lapse rate of the psychometric will be constrained between 0 and 0.5 again to ensure that the shape of the psychometric. lapse rates below 0 and above 1 will produce probability values outside the [0; 1] range and values above 0.5 will flip the shape of the psychometric, as negative slope values will. Not containing the PF in this way could lead to two distinct solutions to a given problem as negative slope values and lapse rates above 0.5 would be able to produce the same mathematical transformation of stimulus values to probabilities making the solution non unique.

From a more theoretical level an appropriate range of parameter values can be narrowed down by looking at the function of interest and investigating whether the observed behavior (given the parameter values) is physically or biologically plausible and which values we would expect are more frequent. For the PF we might expect a few of our participants to not be particularly interested in the task and therefore just respond at random, which would amount to having a lapse rate of 0.5 or really shallow slopes, however this behavior is quite unlikely and expecting only few lapses in the experiment, given that it is conducted in a quiet environment is likely. Lastly using empirical knowledge from the literature at large helps narrow the parameter space further. For the sake of argument, one might investigate the detection threshold for cold stimulation on the skin. Just given this information alone we can narrow down the threshold for the cold detection to being below the skin temperature of around 30-34 degrees ([Courtin et al., 2023](#ref-courtin_spatial_2023)) and -273 degrees. however knowledge from the scientific literature would suggest that thresholds between 28 and 33 would capture most of the population ([Lithfous et al., 2020](#ref-lithfous_accurate_2020)). These same arguments would apply for the slope. This practice of investigating the assumptions of the used parameter values is closely related to those of prior predictive checks when doing Bayesian inference. Prior predictive checks serve as a check of the model, without having seen any data. This check also revolves around simulating data from just the priors of the model and then investigating whether these conform with both what is physically and theoretically plausible, but also serves as a tool to investigate that the model can capture the behavior that is expected from the given experiment ([Kruschke, 2021](#ref-kruschke_bayesian_2021)).

The next challenge is about internal validity i.e. can our model estimate the parameter values that was used to simulate the data that the model estimates the parameter values on. To test and validate our models, we simulate data from pre-specified parameter values which have been deemed to be appropriate using the first step. We then feed our models with this simulated data and investigate how well the model can estimate the latent simulated parameters. This exercise of simulating behavior and then re-estimating the parameter values from the simulated behavior is commonly known as parameter recovery. Generally if this procedure succeeds, then the parameters are said to be recovered. The satisfactory criterion often refers to some correlation coefficient, between the estimated and simulated parameter values ([Wilson & Collins, 2019](#ref-wilson_ten_2019)).

Parameter recovery can thus be thought of as an internal validation of a model, which if done properly should increase the faith in the parameter estimates when the model is fit to real world data. This is because if we had known the parameters values beforehand (i.e. simulated them) then we know that they are somewhat close to the estimated parameter values we got from fitting our model to the data. The assumption is thus; if our model recovers the parameter values well in a simulated setting then it must also do so when fitted to real world data where the underlying parameters are unknown. This assumption is of cause not necessarily true and rests on auxiliary assumptions such as that the underlying generative cognitive model is the same or at least close to the same as the one used to model the data. Because the process of doing parameter recovery assumes that we know the underlying generative model, which is not the case when fitting real world data. To further elucidate this point we imagine using the 3 parameter PF described above, we find that it recovers its parameters well using simulated parameters from the same model. However, if we instead of simulating data from the same underlying model, instead simulated data where the underlying cumulative distribution was the logistic or another cumulative probability density function, we might find that our model cannot well recover the parameters. This is of course nonsensical from the beginning, as how might our model recover parameters of another model, but many times the differences in our model space (i.e. the models that we think underlie the generative process) are similar and the parameters have similar meanings as they come from the same or similar underlying theory, meaning they can be compared. This last point of ensuring that we are selecting the right generative model is the challenge of external validity. The challenge is that infinitely many generative models exist that are also compatible with the observed behavior. This challenge cannot easily be solved as ensuring that we are using the right generative model would entail testing all generative models and being able to compare them, while ensuring that all these models are distinguishable. What is therefore commonly done in the cognitive science litterature, is to use the theoretical framework(s) to build competing models which contain different assumptions of the underlying generative process and then compare this subset of the entire model space, as these are the models that our theories deem relevant. This highlights two important aspects, firstly our models reflect our theories and are therefore at best as good as our theories and secondly, we are surely missing the real generative model in most cases. In practice what is commonly done is that models are fit to real world data and then compared on how well they can describe the data using statistical metrics such as information criteria. The problem with this approach is whether we can accurately distinguish the the particular models that we are testing. This challenge has been addressed using model recovery, which is the act of simulating data from all tested models and then refitting all models to the data simulated by all individual models. Going back to the example of the PF we might have two competing theories of how stimulus values are translated into binary choices, one involving the lapse rate and one without, further we want to ensure that we can distinguish between the normal and logistic cumulative distributions which transform stimulus values into probabilities in different fashions. In this practical example the model space consists of 4 models i.e. two or three parameters for each of the two types of PFs. One would therefore simulate data from these 4 distinct models and fit them all individually to each of the 4 simulated datasets and lastly determine which of the 4 models describe the data the best in each case. The result of such model recovery is a N times N matrix with N being the number of models, the rows being which model was used for the simulation and columns being which model was used to fit the data. The entries of the matrix are commonly depicted as the probability of choosing a particular model given the data simulating model. An identity matrix therefore represents that the models are completely distinguishable and anything else would indicate that in some of the simulations the best fitting model was not the model that simulated the data ([Wilson & Collins, 2019](#ref-wilson_ten_2019)).

## *Limitations of current model validation steps*

The model validation steps above should ideally serve to increase our faith in our models, their parameters, and the comparison between them. However, the metrics used to access these different types of validations are flawed. Firstly, the metrics used can be easily manipulated (faithfully or not) to show good model validation when masking the actual poor or terrible validation. This problem can thus introduce false faith in the model and overconfidence in the inference made based on it. Next and perhaps more importantly the metrics used are not sensitive or specific enough to give the person building the model information about how and were in the model space the models perform well, thereby leaving valuable insights hidden. In this section I will highlight the metrics commonly used in the literature for model validation which are described in Wilson & Collins ([2019](#ref-wilson_ten_2019)), focusing one of the challenges described above; internal recoverability or parameter recovery.

As mentioned above internal recoverability of computational models are accessed by simulating data from a model given a set of parameters. This behavioral data is then fitted to the model which then optimizes for the parameters, given the data. What is commonly done is then estimating the correlation coefficient between the estimated and simulated parameters. In their seminal paper Wilson & Collins ([2019](#ref-wilson_ten_2019)) describes that in a perfect world the estimated and simulated parameters should be tightly correlated without any bias, and that a weak correlation could mean bugs in the code or an underpowered study i.e. few trials. They also reiterate that plotting simulated vs estimated parameters should be done to access if ranges of parameter values are problematic and whether there might be biases. I will here argue that the correlation coefficient is an inappropriate metric and that a version of an intra class correlation (ICC) is better suited for the task. Acknowledging two important things; neither metric is perfect, and visually inspecting the simulated vs estimated parameter scatterplot is crucial. The importance in using the right metric is therefore as a precautionary step given that some literatures are starting to just report correlation coefficients without this crucial scatter plot, which arguably in some cases would make the correlation coefficient meaningless, see problems with the correlation coefficient below. These precautionary steps are crucial to enforce, in the development stages of new statistical, cognitive or computational models as they will serve as the basis of model validation and if not sensitive or specific enough many resources might be used in using a model that in reality cannot be properly identified. This would therefore serve as a roadblock for scientific progress as years might pass before someone realizes that the model used in the field is not behaving properly.

## *Current problems with internal recoverability of models (parameter recovery)*

The first and perhaps biggest problem of internal recoverability of computational models Is that it is not universally done, which from a readers perspective makes it hard or even impossible to know if the generative model in question can be trusted. The second, almost ubiquitous problem in the literature using parameter recovery is that interactions between parameters are either neglected or disregarded. This is less of a concern for individuals using an established cognitive models wanting to ensure that given their experimental design and ranges of parameters are sensible, but a big concern in highly cited method papers describing and formalizing the models themselves. A prime example of this is the Hierarchical Gaussian filter paper ([C. Mathys et al., 2011](#ref-mathys_bayesian_2011) ; [C. D. Mathys et al., 2014](#ref-mathys_uncertainty_2014)). Where after having laid out the equations of the model, two of the most crucial parameters of the model are held constant when performing parameter recovery. Even in much more simple models such as with the PF described above, I will show that there are tradeoffs and interchangeability between parameters. The last problem with parameter recovery is the metric used to access it. As has been suggested elsewhere, the correlational approach to parameter recovery is at best insufficient and at worst misleading ([Schurr et al., 2024](#ref-schurr_dynamic_2024)). The three most obvious problems with using correlation coefficients to examine internal recoverability is as follows;

Correlation coefficients are invariant to linear transformations, making two sets of variables i.e. [1,2,3] and [1,2,3] have the same correlation after transforming on one of the sets with linear transformation. y=2\*x+3 (or report as a matrix idk) Resulting in the sets [1,2,3] and [5,7,9]. This invariance to linear transformations does not make sense for parameter recovery as we want a metric that penalizes this behavior.

The domain of correlations is between -1 and 1. This directionality also does not make sense given that a correlation coefficient of -1 would mean perfect parameter recovery, with a negative sign of the simulated or estimated parameter meaning that you do recover the value (or the linear transformed value) just not the sign. Ideally, we would want a metric that goes from no recovery to perfect recovery.

Lastly, the interpretation of the correlation coefficient in terms of parameter recovery is difficult. What is a sufficiently large correlation coefficient for the parameter to be said to be recovered and what types of uncertainty is causing the correlation to be less than ideal. Authors have tried to make such distinction without much traction ([White et al., 2018](#ref-white_testing_2018)). All these issues are similar to what researchers face when wanting to estimate the stability and or test -retest reliability of different metrics over time, where the solution has been to use the ICC as the metric instead of simple correlation coefficients ([Schurr et al., 2024](#ref-schurr_dynamic_2024)).

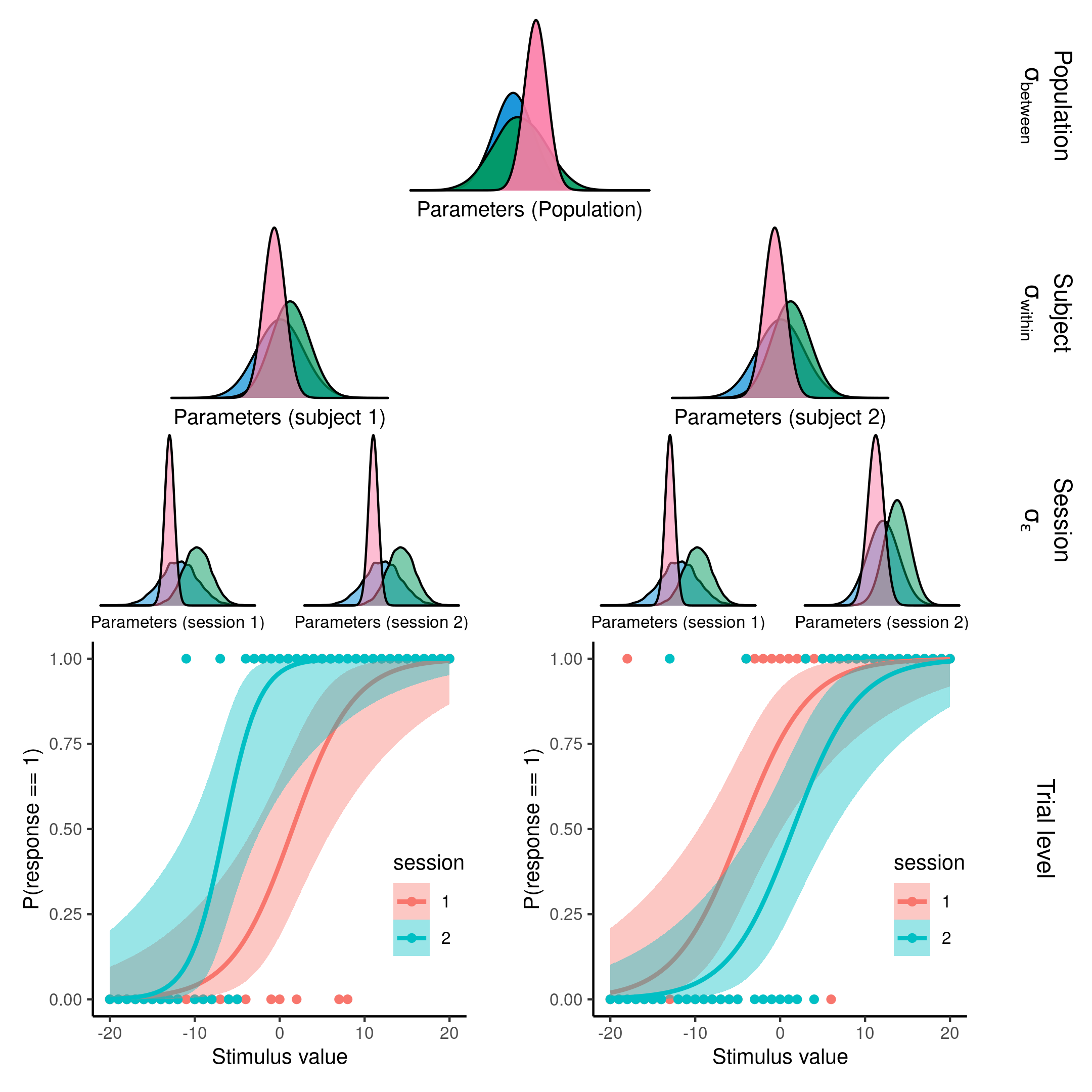
## *ICC Parameter recovery*

Given that the idea of using the ICC as a metric for parameter recovery is relatively new and to the authors knowledge has only been suggested and not been used anywhere in the literature ([Schurr et al., 2024](#ref-schurr_dynamic_2024)) I will here outline what the ICC is and how it can overcome some of the shortcomings of the correlation coefficient. The ICC is in its simplest form a ratio of irreducible variances (uncertainties) to the total variance in the data. In practical terms revovling around a cognitive science experiment the irreducible uncertainty is the uncertainty between subjects, whereas the total uncertainty can have several parts. In order to calculate the ICC, a model is needed that can properly account for these different types of variance and the typical approach are hierarchical models, where known structure of the data is added to the model.

Taking an example, we imagine a researcher doing a test-retest reliability study on a parameter of a cognitive model. His subjects are coming in for x sessions and doing the same cognitive task each time. We will now assume that all subjects come from the same underlying distribution of say humans (i.e. the population), this is the highest level in the hierarchy and is governed by a population mean and a population variance, i.e. the between subject variance. The next level in the hierarchy is the subject level, here each subject has their own means and variances (within subject variances), their means are drawn from the population distribution. Now for each session that the subject is in, a parameter value is drawn from this subject level distribution which then governs the participants’ behavioral responses. This nested hierarchical structure is demonstrated in figure 5, where each of the levels are governed by the levels above and each of the levels has a variance associated with it, where the between subject variance is the variance of the population level distribution and the within subject variance is the variance of each of the participant level distributions. The ICC as mentioned above is the ratio between within and between subject variances.

Where is the variance between the subjects’ parameter estimates and is the within subject variance. Given that we are interested in the performance of the model we can simulate agents that have no within subject variance i.e. the same true parameter values for each session and then see how the number of subjects and or trials of the cognitive task will influence the model’s ability to pick up on this association. This approach has one clear problem it does not necessarily tell us something about how well the model estimates the true parameter values for each participant at each session, as it just looks at how close each parameter is to itself between sessions. To capture this, one might use the mean squared errors (MSE) between the simulated and estimated parameter values, which serves as a residual error of the model. Including this into the ICC formulation is easy, as this is just another source of variance which can be added into the denominator, highlighting the fact that the ICC is a partitioning of variance in the model. This partitioning of variance is exactly what we are interested in when building and validating models, as this tells us where the model fails and where it might excel. In figure 5 this amounts to the difference between the estimated parameter value of a particular subject at a particular session and the simulated value. Formally we add the MSE into the equation and get.

Where is the MSE. This conceptualization allows us to put parameter recovery for a model into a single value for each parameter that ranges from 0 to 1 which is going to be trial and subject level dependent, but also dependent on the simulated ranges of parameter values.



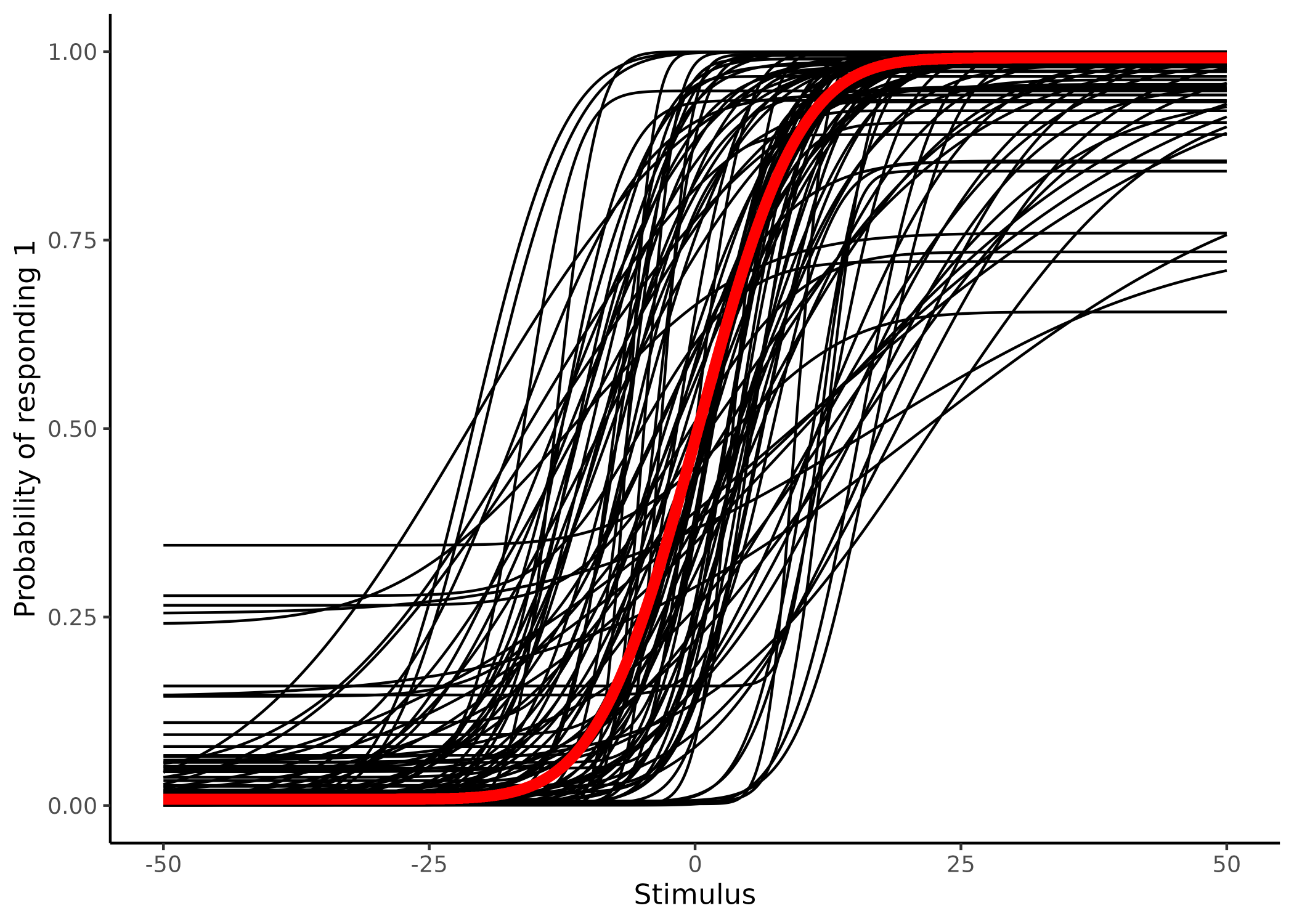
**Figure 5. Visualization of a nested hierarchical model with sessions nested in subjects in a population.**

## *Standard parameter recovery.*

The model and task used to demonstrate and investigate parameter recovery in this thesis is the 3 parameter PF described above which is widely used in the cognitive science literature from perception to decision making ([Courtin et al., 2023](#ref-courtin_spatial_2023); [Gold & Ding, 2013](#ref-gold_how_2013)).

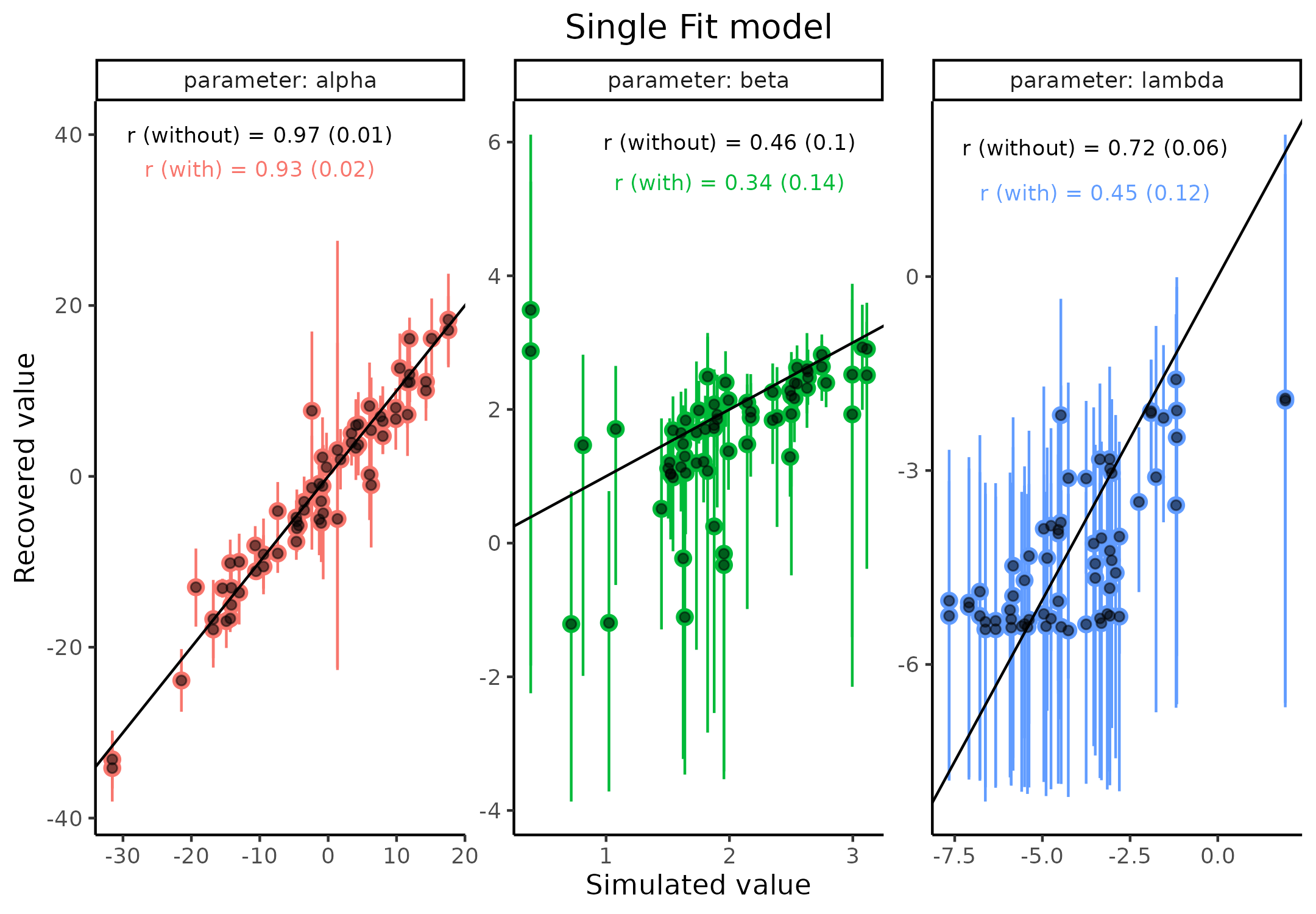
After having specified the model, we can simulate data from different ranges of parameters to select appropriate ranges of parameter values. Firstly, parameter ranges are selected and simulated in accordance to table 1 and figure 6. Using the probabilistic programming language Stan and its interface with R, Rstan ([Gabry et al., 2024](#ref-R-cmdstanr)), it is possible to invert the model from the data to obtain estimates of the latent parameters which were used to simulate the data in the first place. The fitted psychometric function is first fitted on a single subject level and the parameters and their associated estimation uncertainty extracted. Note that for all models displayed and estimated their convergence was accessed by ensuring rhat values were below 1.03 and no divergent transitions were present, ideally all chains would have been inspected, but given the vast simulation approach presented throughout the thesis, visual inspection of each model was infeasiable and summary diagnotistic were used. Furthermore all priors for all models were weakly informed, meaning that most of the prior distributions were set as normal distributions with means of 0 and standard deviations of 3-5 in the unconstrained space. Readers are referred to to the supplementary material or the github for a list of all the priors used.

| Parameter | Mean | Sd | Transformation |
| --- | --- | --- | --- |
| Alpha | 0 | 10.0 | x |
| Beta | 2 | 0.6 | Log(x) |
| Lapse | -4 | 2.0 | Logit^-1(x) / 2 |



**Figure 6. Displaying 100 samples of the parameters of the psychometric function from table1.** Visualization of the implications of the simulated parameters of table1. Black lines depicting individual subjects, while the red line depicts the group mean.

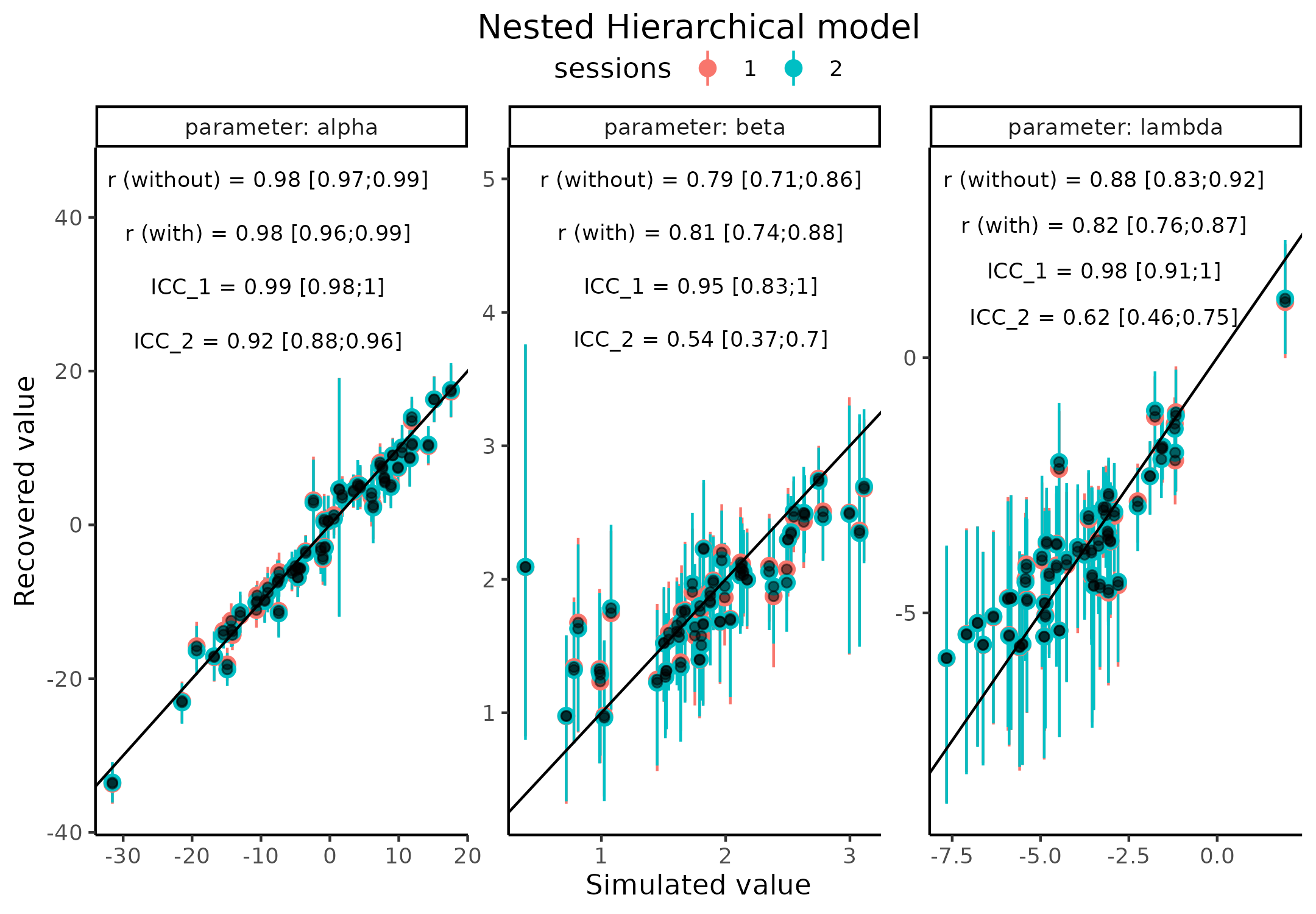
For the sake of argument, I have plotted the pairwise scatter plot of estimated vs simulated parameter values in figure 7 with the added estimation uncertainty. This particular simulation is done for 100 data sets i.e. subjects over 100 trials each. Figure 7 also displays how adding the estimation uncertainty (of the parameters) to the correlation coefficient changes the resulting size and uncertainty estimate of the correlation coefficient (i.e. its estimation uncertainty). This addition of the estimation uncertainty again highlights how deceptive these estimates can be if uncertainty is not properly propagated as they are all inflated. This inflation is however not necessarily always the case, as if a couple of points fall way off the identity line with high uncertainties, they will have less weight when accessed with uncertainty compared to without, meaning that adding estimation uncertainty could in principle also increase the correlation coefficient, highlighting the non trivial and linear link when uncertainties are non linear aswell.



**Figure 7. Parameter recovery for the three parameters of the psychometric function.** Scatter plot of Simulated vs recovered parameter values, with error bars displaying the 95 highest density interval for that parameter on that simulation.

Next the purposed ICC metric will be tested on the same data set as above, crucially the data set above was simulated using only 50 simulations that were duplicated, making it eligible to compare the above standard parameter recovery with the ICC. This simulation therefore implies that there is no within subject variation. One particular difference between the above single fit models and the proposed model depicted in figure 8 is the hierarchical structure embedded in the model. The hierarchical structure of the model serves to shrink parameter estimates in relation to their distance and uncertainty from the mean of the higher level which they are drawn from, which in the end has been shown to improve predictive capacity. Hierarchical models are becoming corner stones in most cognitive science experiments ([Bates et al., 2015](#ref-bates_fitting_2015); [Boekel, 2021](#ref-van_boekel_pool_2021); [Gomes, 2022](#ref-gomes_should_2022)).

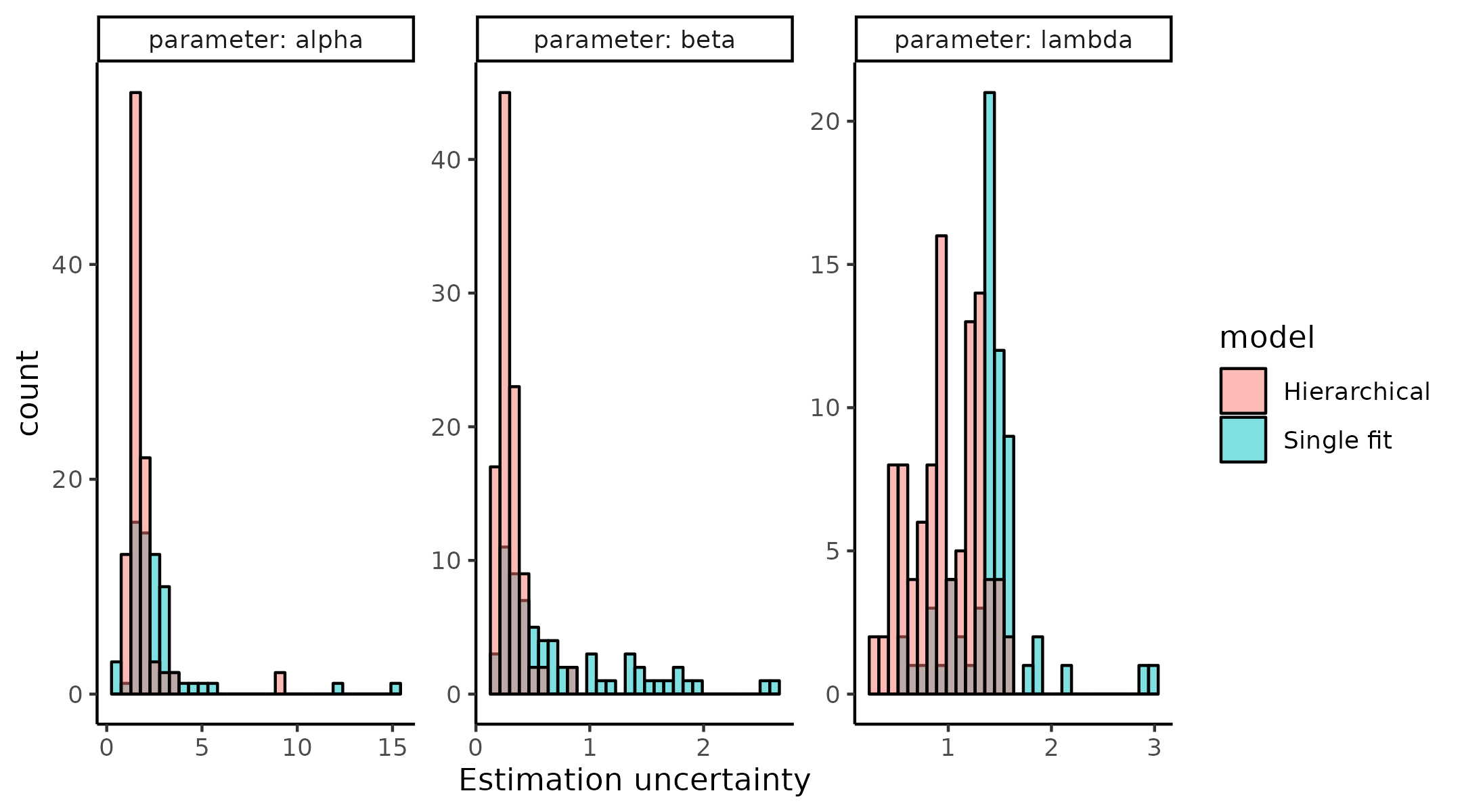
Therefore to fairly test and compare the two internal validity metrics, correlation coefficient and the ICC, each of the metrics were calculated from this hierarchical model. the Two ICC values were calculated as described in equation 1 and 2 now referred to as and .



**Figure 8. Parameter recovery for the three parameters of the psychometric function using the hierarchical model.** Scatter plot of Simulated vs recovered parameter values, with error bars displaying the 95 highest density interval for that parameter on that simulation.

As can be seen this hierarchical fit does improve the parameter recovery, both from a visual inspection (points falling closer to the identity line with less estimation uncertainty) and by looking at the correlation estimates, this is both with and without accounting for estimation uncertainty in the correlation coefficient. It should here be noted that a single simulation like this would not be enough to ensure that the parameters are nicely recovered as a good example of this is the lambda parameter. The pairwise scatter plot of the nested hierarchical model seems to suggest that this parameter is quite well recovered but, if we back calculate a lambda value of -5 corresponds to a lapse rate of 1.3%, which obviously is hard to find when there are 100 trials for each subject. See supplementary material XYZ for a deeper explanation on this.

(supplementary it is even lower than the 1.3% for a lapse of -5 as the stimulus value also needs to be far enough away from the underlying psychometric to actually matter in the estimation). Turning the attention to the ICC values, we firstly observe that on each of the 3 parameters has an upper bound at the maximum value of one, which is somewhat confirmed looked at the scatter plot as one would assume that all the session one estimates are hidden behind the session two estimates with only a few estimates deviating slightly. The estimate is crucially the lowest for all three parameters as all parameters have upper bounds of 1. Visual inspection of the pairwise scatter plot makes this clear as this metric is penalized by both the degree to which the points fall away from the identity line but also by the estimation uncertainty associated with these points. This therefore also explains why the alpha parameter is close to being asymptotic at 1, but with a little to be desired for simulated values between 0 and 10. An interesting observation in Figure 8 is that the difference between lambda and beta is minute in all the metrics used. In the next section it will be shown how it is possible to do better by reducing the estimation uncertainty. See Figure 9 for the distribution of estimation uncertainty for the three parameters in the single and hierarchical fit. This Figure also helps explain why hierarchical models in general are prefered to single fit models as the partial pooling improves estimation of the parameters ([Bates et al., 2015](#ref-bates_fitting_2015); [Boekel, 2021](#ref-van_boekel_pool_2021); [Gomes, 2022](#ref-gomes_should_2022)).

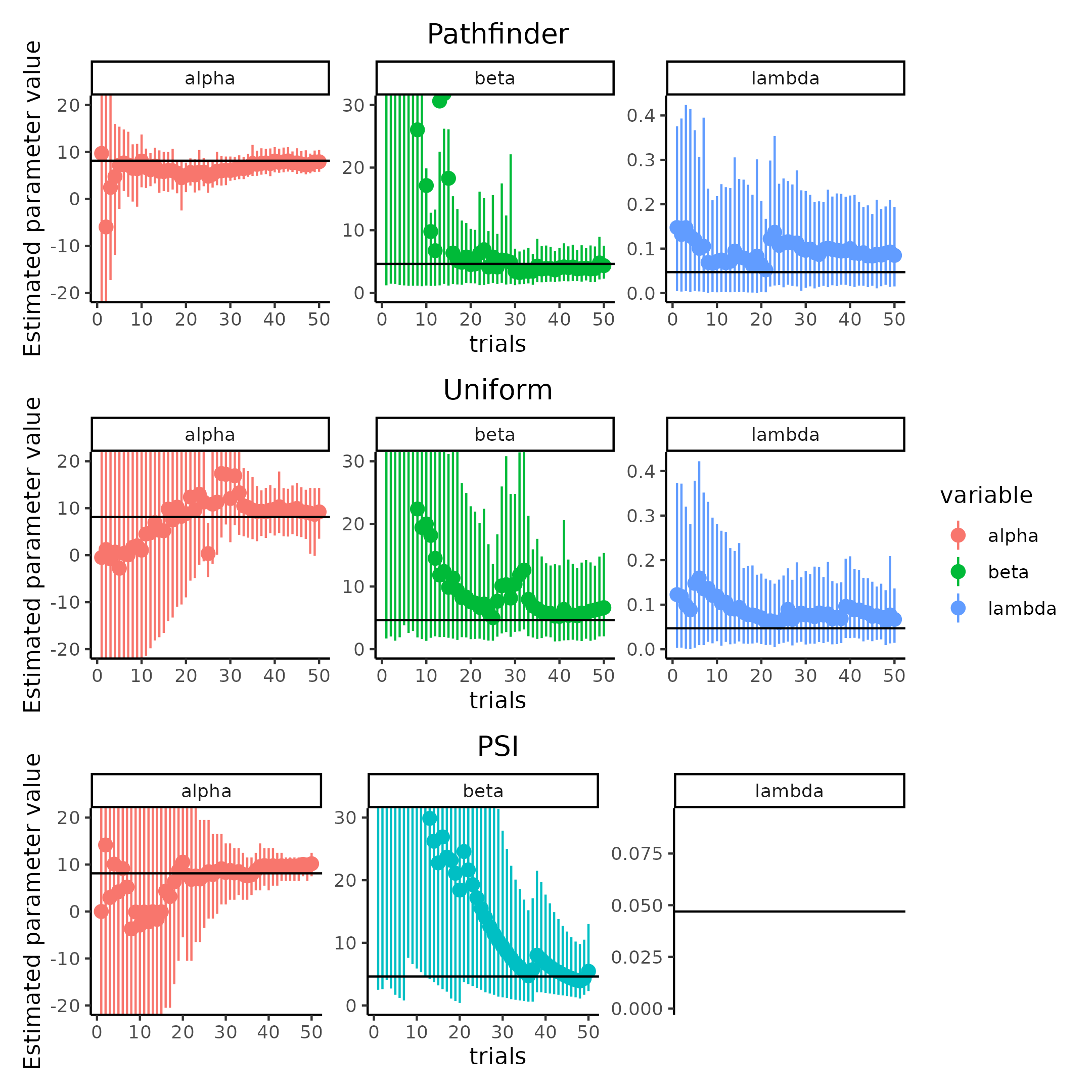


**Figure 9 Estimation uncertainty for each parameter for both single and hierarchical fit models** Each panel represents one of the three parameters of the psychometric function with the estimated uncertainty depicted as histograms. The color of the histogram shows whether the model was fit using the single fit or hierarchical model.

## *Adaptive design optimization*

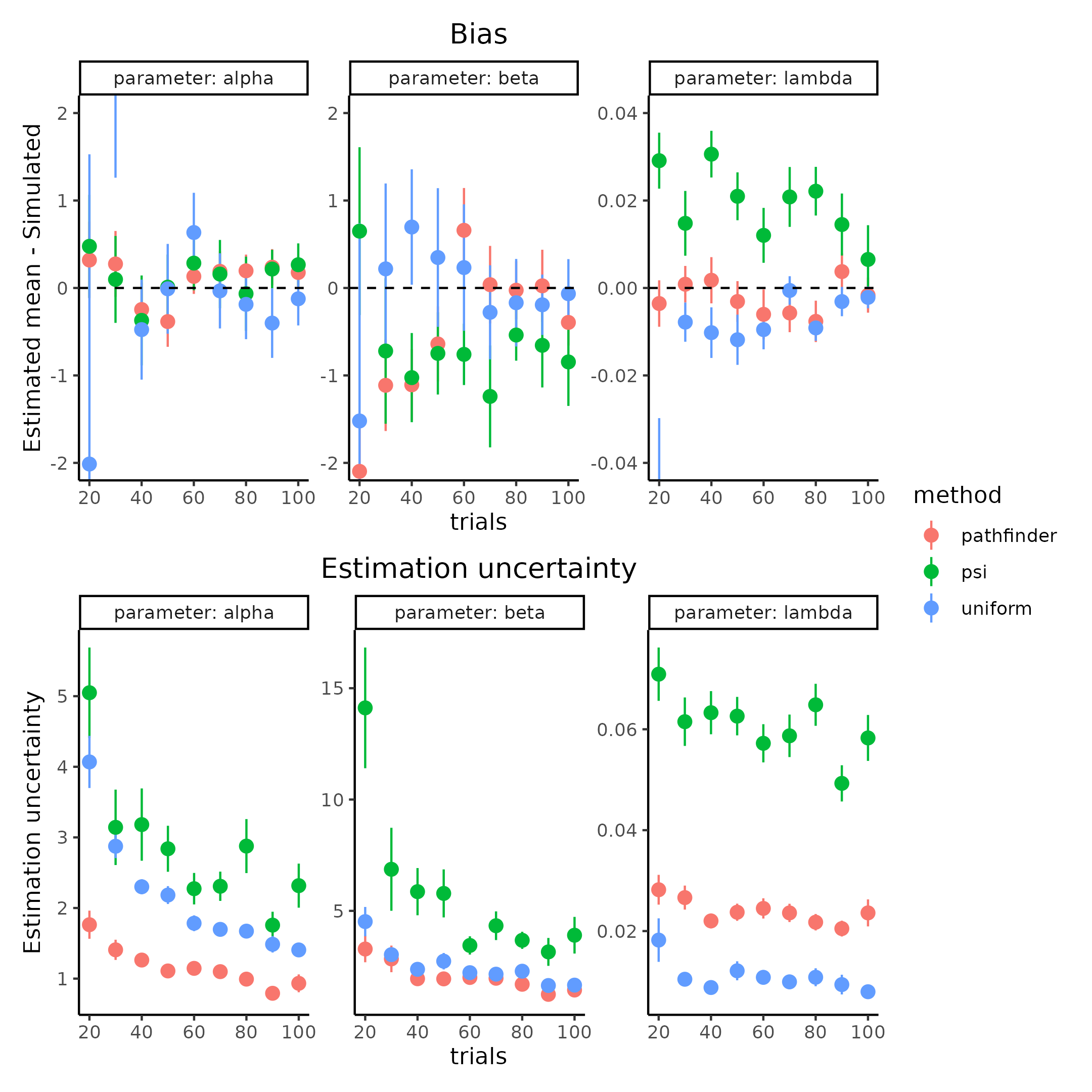
An import consideration of the parameter recovery that was left out in the parameter recovery analysis described above. What is the design of the experiment that the simulated agent goes through? Looking back at figure 6 providing stimulus values in the far ends of the psychometric functions i.e. in the ranges of [-50 ; -25] and [25 ; 50] will in most cases for most agents give next to no information on the shape of the psychometric and therefore the parameters we mostly care about i.e. alpha and beta, as on average the agents’ psychometric functions are monotonically increasing in the interval of [-25 ; 25]. Therefore, selecting stimuli (inputs) in this interval must be better for decreasing the estimation uncertainty in the two parameters we care about, compared to randomly or uniformly exploring the input space. We might even go a step further and instead of selecting inputs that are more appropriate for the mean of the population we could individualize each experiment to the agent or subject. This practice of individualizing the experiment of interest is called adaptive design optimization (ADO) and has quite a big literature behind it and revolves around selecting inputs that are optimal given a specific criterion ([Prins, 2013](#ref-prins_psi-marginal_2013); [Watson, 2017](#ref-watson_quest_2017)). Many of these criterions exists such as minimizing entropy, minimizing the posterior variance or mutual information, but what they all have in common is that they do decrease estimation uncertainty of either all or certain parameters to a meaningful degree. In order to keep in the same theme as the rest of the thesis I will instead of utilizing the few available packages that exist for doing ADO for psychometric functions I will show how utilizing the single fit model which was built for conducting the simple single subject parameter recovery can be utilized together with the knowledge that the most informative stimuli for determining the shape of a single agents’ psychometric function is somewhere in the middle region of that psychometric. One of the main challenges of utilizing ADO is that because the experiment is updated and individualized an algorithm determining the next stimulus must run in tandem with the experiment. This puts quite a high constraint on computation time of the algorithm, this issue has partly been solved in the existing packages by before conducting the experiment mapping out a grid at a particular resolution of parameter values at a current trial and then what the optimal stimulus value to present is. This clever solution puts the heavy computation time before the experiment and ensures that when the experiment is run only a single look up is needed to provide the next stimulus value. This approach works great for psychophysical experiments or other experiments where each trial is independent of the next. This is because then only a single optimization step is required for each trial, whereas if trials were mutually dependent as in a learning experiment, then the algorithm would need to calculate all possible lines of stimuli and responses to a certain point which given the combinatorics can become a daunting task.

To provide something that is more generalizable and can be continuously updated on a trial-by-trial basis while the experiment is run, other approaches might be more appropriate. Illustrating such an approach can easily be done with the same model used to fit individual subjects when using the R and Stan, the quick estimation of the posterior distribution is then done using a variational inference algorithm, particularly pathfinder ([Zhang et al., 2022](#ref-zhang_pathfinder_2022)). Figure 5 shows how the posterior distribution of the 3 parameters of the PF varies as a function of trials in both the uniform and pathfinder approach to selecting stimulus values. As can be seen both approaches makes the parameters converge towards the real simulated values (black line), however the speed at which this happens is clearly very different, especially for the two parameters we are the most interested in i.e. alpha and beta. For these two parameters after just 20 trials of pathfinder the optimization has found the simulated parameter value and decreased the estimation uncertainty (posterior variance) to close to 0 whereas even after 50 trials the uniform approach still has a bit of a bias in the estimation, the individual points are not on the black line, but also a substantial estimation uncertainty associated with it. For completeness a PSI-algorithm was also used to compare to ensure that the pathfinder algorithm was not too slow or bad ([Kontsevich & Tyler, 1999](#ref-kontsevich_bayesian_1999)).



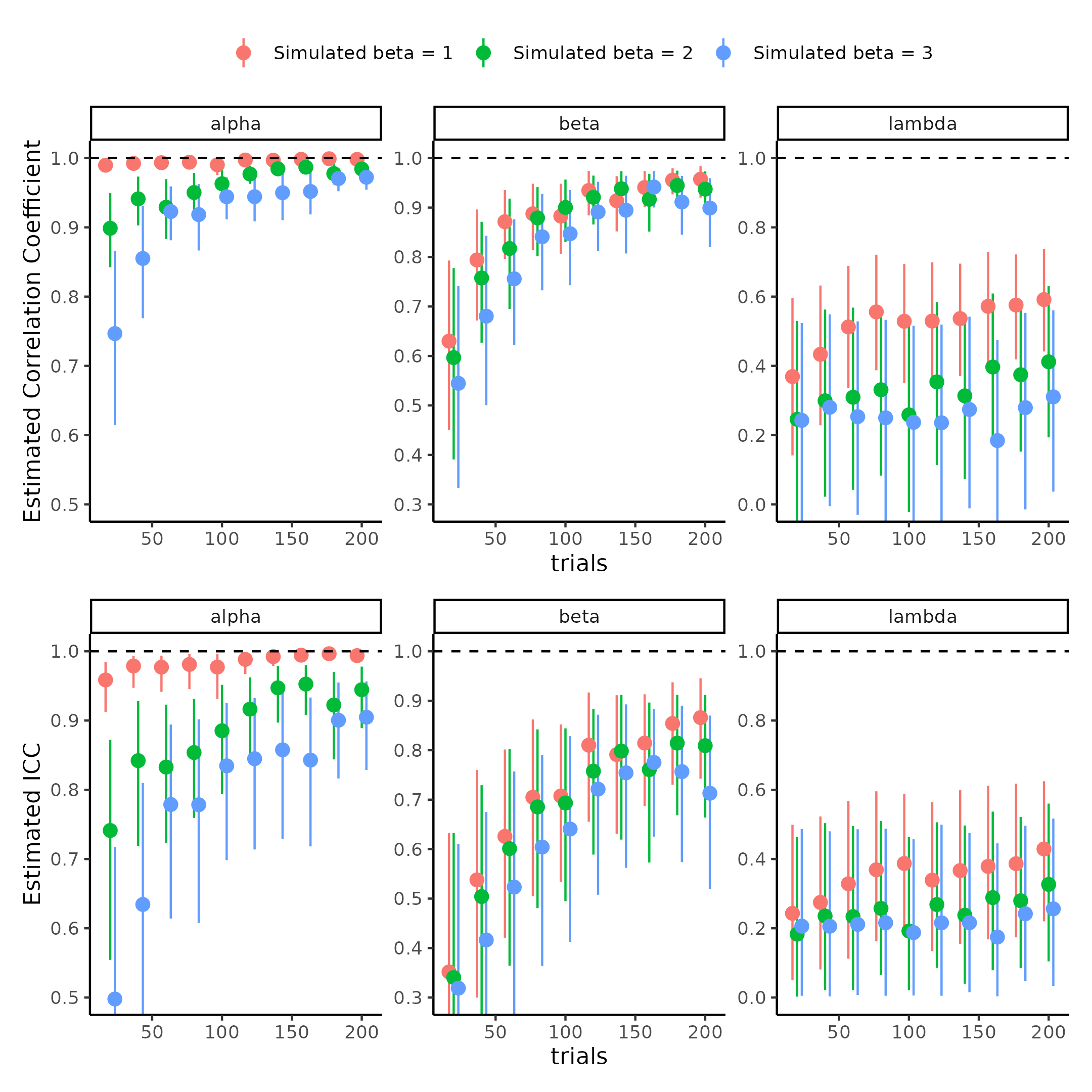
**Figure 10 comparison of algorithms to obtain stimulus values of the psychometric function**

To show the improvement more rigorously in reduced estimation uncertainty especially across a range of trial numbers, the Pathfinder, Uniform and PSI algorithms were run 100 times for trials ranging from 20 to 100 in a sequence of 10 trials, in order to make the comparison as fair as possible each of the algorithms were only used to generate the stimulus sequence, meaning that all three types were refitted using the same single fit Bayesian model. For complete details on the fitting and optimization strategy see supplementary material XX.



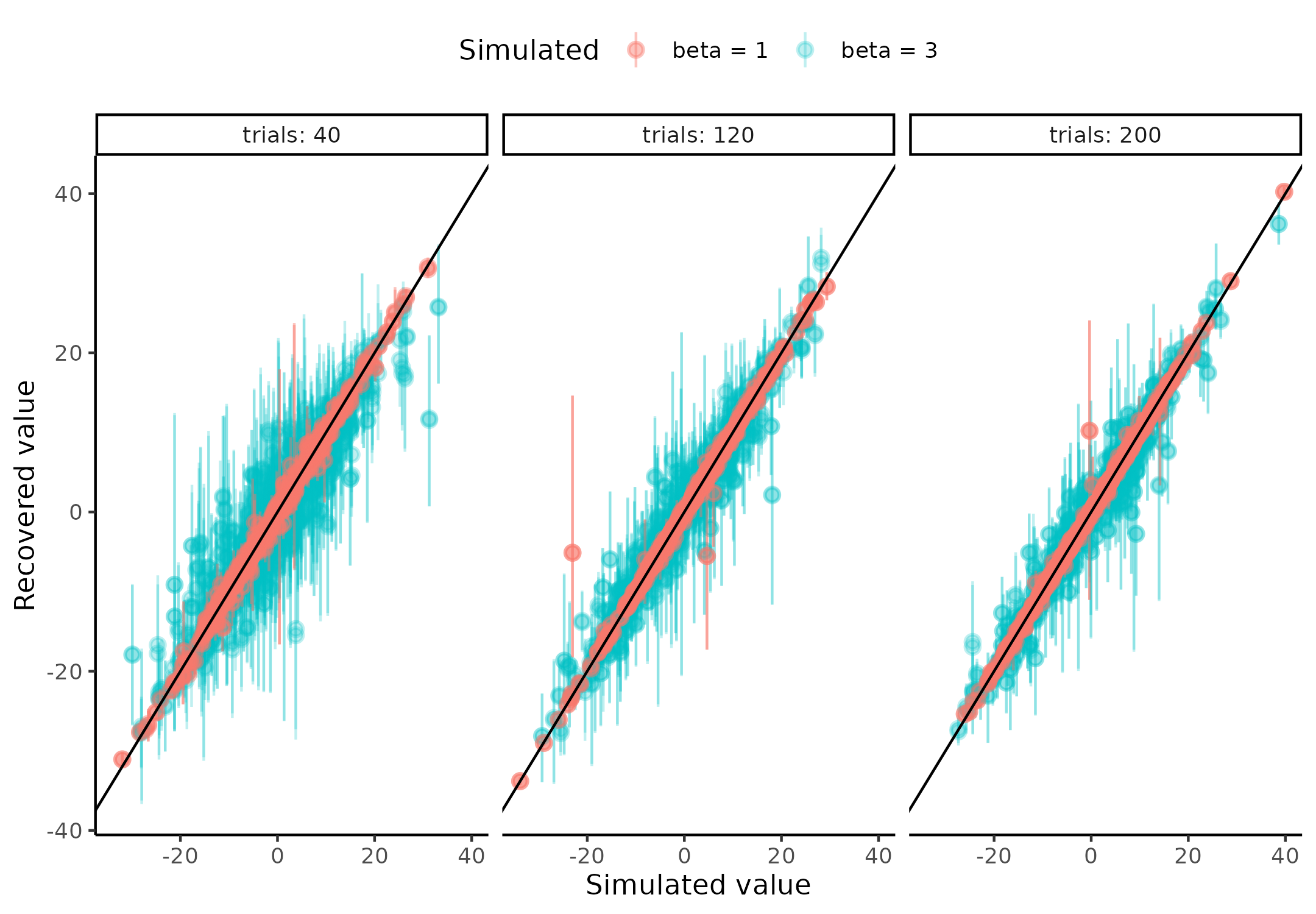
**Figure 11.** shows how the estimation uncertainty and bias changes according to the number of trials and parameter value estimated with the different methods.

Using this pathfinder algorithm, we can now examine three other focal points, subjects, trials and the influence on the mean simulated slope value. The last point is less obvious than the two others but stems from the fact that increasing the slope (decreasing the steepness) of the PF will make it harder to estimate, but also influence the estimation of the threshold, which will become clear below. For this purpose, I will simulate trials ranging from 20 to 200 in increments of 20, subjects being between 10, 30 and 50 and lastly mean slope values of 1,2 and 3 in the unconstrained space, all other parameter values being identical to table 1. To guard against simulations that are not representative due to either bad convergences in the ADO or in the fitting procedure, each combination was run 5 times. Figure 12 shows how the correlation approach with added uncertainty to parameter recovery fairs (for the standard approach of no uncertainty see supplementary XXX), the lower panel shows how the fairs on the same simulated data sets (for the analysis with only the within subject variance see supplementary QQQ).



**Figure 12 comparison of parameter recovery metrics.** First row depicts how the estimate of the correlation coefficent between simulated and estimated means change as a function of trials (x-axis) and the simulated mean slope (color) for each parameter of the psychometric function (columns). The bottom row shows how the estimate of changes based on the same metrics as the correlation coefficient. Note that the correlation coeffecient has been uncertainty propergated using bootstrapping.

What seems to be the main difference between the two approaches is the in the lower number of trials and especially in the comparison between the high simulated slopes (lowest panel) for the threshold as both approaches seem to suggest that in high number of trials (> 100) and in steep slopes (beta >= 2) that the threshold is fully recovered. The difference is clearly in the lowest panel where the ICC approach suggest that there is still variance left unexplained, to investigate this we can plot the pairwise scatter plot of the high simulated slopes (beta = 3) and low simulated slopes (beta = 1) on different trials and subjects. Figure 13 clearly shows why there is such a difference between the two approaches, the ICC metric is much more stringent on the higher-level estimation uncertainty when the simulated slope is less steep. Turning the attention to the slope itself, there also seems to be a difference. What is present is again that the ICC metric has lower values in general and is not asymptotic at one with the configurations used here (to see the pairwise scatter plots see supplementary material XXX). Lastly both approaches suggest that the lapse rate is well below acceptable ranges, but still with the ICC being more conservative.



**Figure 13.** Showing the pairwise scatter plots of simulated vs recovered threshold () parameter when the simulated beta value is low (beta = 1) and high (beta = 3) for subjects (rows) and trials (columns).

As conveyed by the pairwise scatter plots the conservative ICC metric capture the fact that estimation uncertainty is a source of variability that can still be reduced even when the correlation coefficient (also with the estimation uncertainty propagated) might indicate perfect fit. This is exactly the behavior one would like to have when trying to understand their model as this information is much more sensitive, furthermore the values also have a natural interpretation. An ICC value of 0.8 means that 80% of the variance in the model is governed by the between subject level and only 20 % is in the estimation or test -retest uncertainty, the ICC could of cause be further decomposed into what proportion of variance of the 20% is from estimation and what is from test retest uncertainty, however for this particular model it seems like most if not all is from estimation uncertainty (see supplementary XYXX). This straightforward interpretation is not present for the correlation coefficient, especially because of the arguments laid forth in the “current problems with internal recovery” section.

Many if not all papers describing the test retest reliability of cognitive models in the literature finds that hierarchical models are better but also argue some something around the correlation coefficient, here the ICC metric is again really helpful as the estimation uncertainty is going to be soaked up somewhere in the model and that is going to go towards the within subject variance or at least increase it such that in reality the ICC might have been higher than observed but because you did not have enough trials the uncertainty around the estimate i.e. the estimation uncertainty is what is causing the low test retest reliability.

## **Increasing the information in the cognitive model**

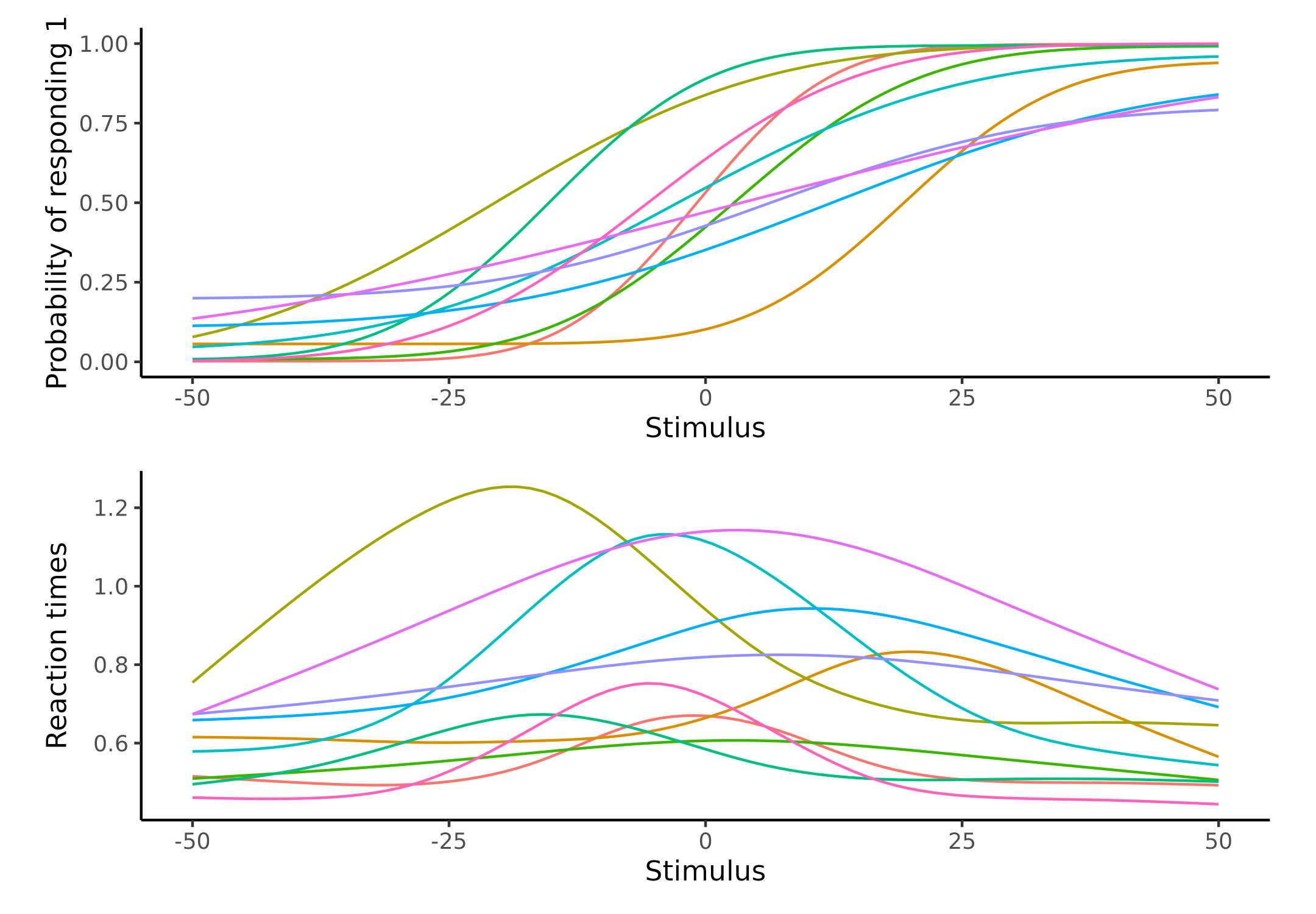
Now I’ve hopefully convinced that the ICC approach to parameter recovery is superior to both the standard and the uncertainty propagated correlational approach in that it better reflects our expectations given the pairwise scatter plots of simulated vs estimated parameter values. The question now becomes given our more nuanced view of parameter recovery what we can do to improve it. The obvious answer given the plots above seems to be increasing the number of trials. However mindlessly increasing trials to gain a certain recovery and or statistical power can be troublesome in non-obvious ways. The obvious problems with increasing the trials number are resources costs, both in terms of money to the participants completing the experiment, the experimenter, but also espeically the time investment from the participants’ side. However, the most problematic aspect becomes more obvious if we take a step back and think carefully about what we are studying. We are studying a complex system that has its own goals, desires and motivations and it is not trivial to how this participant will behave if the task is double the length. Firstly, will the participant employ a different strategy knowing that the experiment is going to take X time longer, or will they halfway through employ a different strategy. Even if the participant keeps the same underlying cognitive strategy that we are trying to model, then one reasonable assumption would be that attentional lapses and engagement in the task will decreasing, making each additional trial after a certain point less “valuable”. I will here argue that in many of the cognitive science paradigms there might be no need for increasing trial counts to increase the recovery of parameters, but to utilize the data that the participants has already provided in better and more sophisticated ways. For the sake of this thesis, I will look at incorporating the reaction times of the agents’ responses as sources of information about the underlying psychometric function of their binary choices. I will be focusing on the reaction times as these have a long and rigorous history in cognitive science literature, but more importantly are present in most experiments conducted today ([MacLeod & Dunbar, 1988](#ref-macleod_training_1988); [Pirolli & Anderson, 1985](#ref-pirolli_role_1985); [Sternberg, 1969](#ref-sternberg_memory-scanning_1969)). In order to incorpurate the reactions times into the current formulation of generative structure of the task, it is helpful to think of the output of PF as a probability of responding a particular value, say 1. This therefore means that in either end of the tail of the PF the certainty with which you respond is the highest and the midpoint between the extremes (the threshold) is the most uncertain. This descriptive formulation is what the variance of the Bernoulli distribution highlights:

Using this information together with the assumption that participants will respond slower when more uncertain and faster while more certain the reactions times of each trial can be modeled as a linear function of this Bernoulli variance, which is calculated from the psychometric function and an intercept to account for the individual differences in mean reaction time.

where represents the intercept and represents the degree to which the uncertainty from the psychometric function influences the reaction times, see Figure 14 for this mapping. The current implementation of the reaction times was done using a shifted log normal distribution introducing two more variables, a non decision time () and a standard deviation () for the log normal distribution. The non decision time parameter helps account for the fact that reaction times below a certain threshold are impossible without the trial being erroneous, as information of the stimuli has to travel to the brain before a response can be made ([Jain et al., 2015](#ref-jain_comparative_2015); [Ranger et al., 2020](#ref-ranger_modeling_2020)).

To show how these reaction times help the recovery of the parameters of interest i.e. the threshold and slope of the psychometric, I have chosen to simulate agents with the parameter values displayed in table 2. To understand the influence of the size of coupling between the binary responses and the reaction times I have chosen to simulate this coupling parameter being 1.5 with the other parameters being as in table 2 with the slope of the psychometric () being 3. Again, showing and understanding what these parameter values mean we simulate the parameters and display the behavior. This can be seen in figure 14 where 10 simulated subjects are visualized.

| Parameter | Mean | Sd | Transformation |
| --- | --- | --- | --- |
| Alpha | 0 | 10.0 | x |
| Beta | [1 ; 3] | 0.6 | Log(x) |
| Lapse | -4 | 2.0 | Logit^-1(x) / 2 |
| Intercept | -2 | 0.5 | X |
| BetaRT | [1 ; 1.5] | 0.3 | Log(x) |
| Sigma | -1 | 0.5 | Log(x) |
| Non-decision-time | -1 | 0.5 | Logit^-1(x) \* minRT |



**Figure 14. Visualization of the psychometric function with Reaction times.** Upper panel depicts 10 psychometric functions where parameters were drawn from table2. Lower Panel depicts the assumed relationship between the stimulus value (x) and the reaction times (y), which as can be seen is dependent on the shape of the psychometric function in the upper panel. The reaction time functions peak around the psychometric threshold and tapers off when the psychometric function asympotes at 1 or 0.

With these simulations we can now visualize what including the reaction times into the modeling means for the parameter recovery for the influence on the other metrics like the correlation coefficient see supplementary material XXX. Figure 15 displays the results of this analysis, showing the means and 95% confidence intervals of the for the simulation for the three parameters of the PF together.

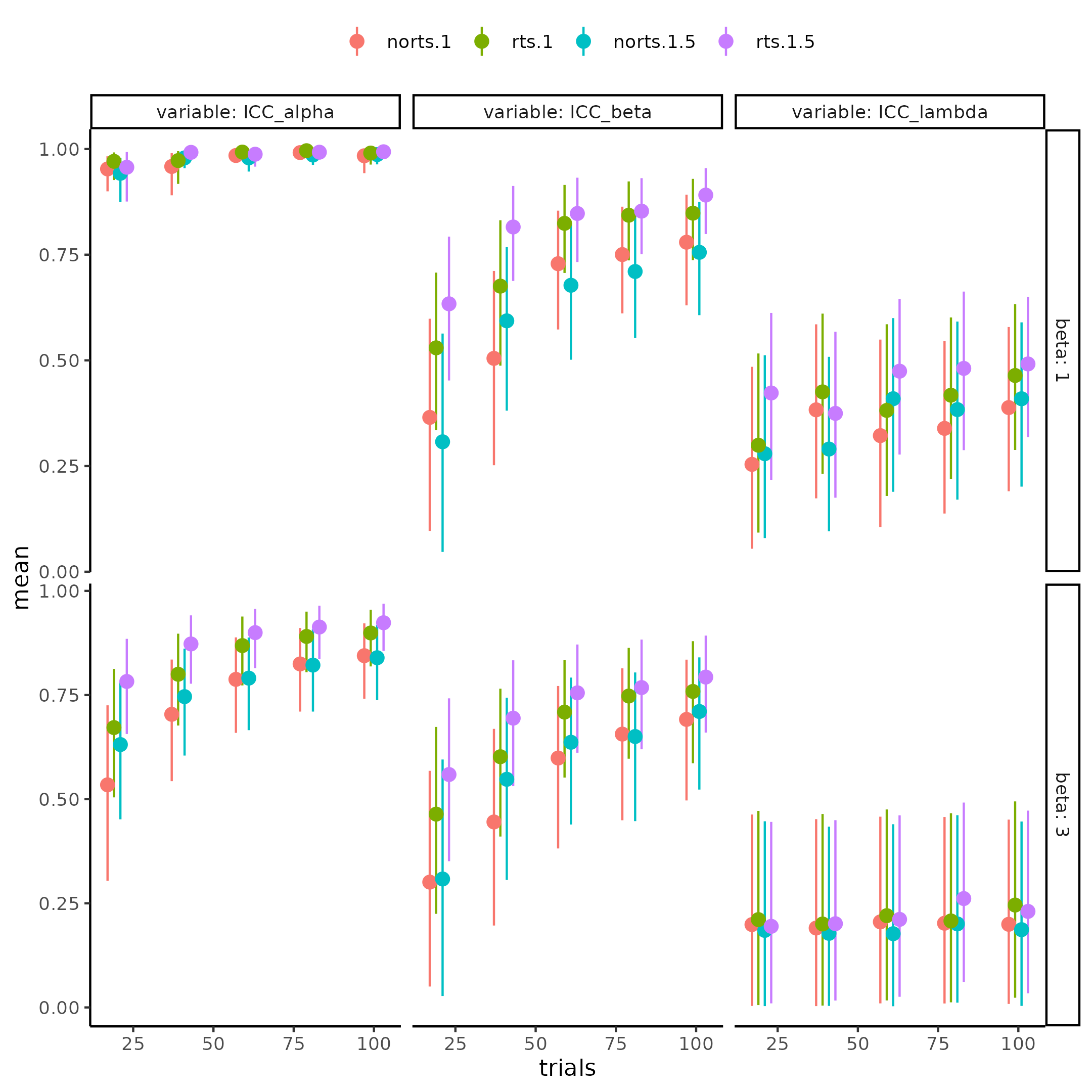


Figure 15 histogram of the mean difference between the ICC value obtained from using the reaction times or not, colors display the simulated level of coupling between the underlying psychometric function and the reaction times. Stronger coupling is associated with bigger parameter recovery effects for both threshold and slope, but not lapse rate.

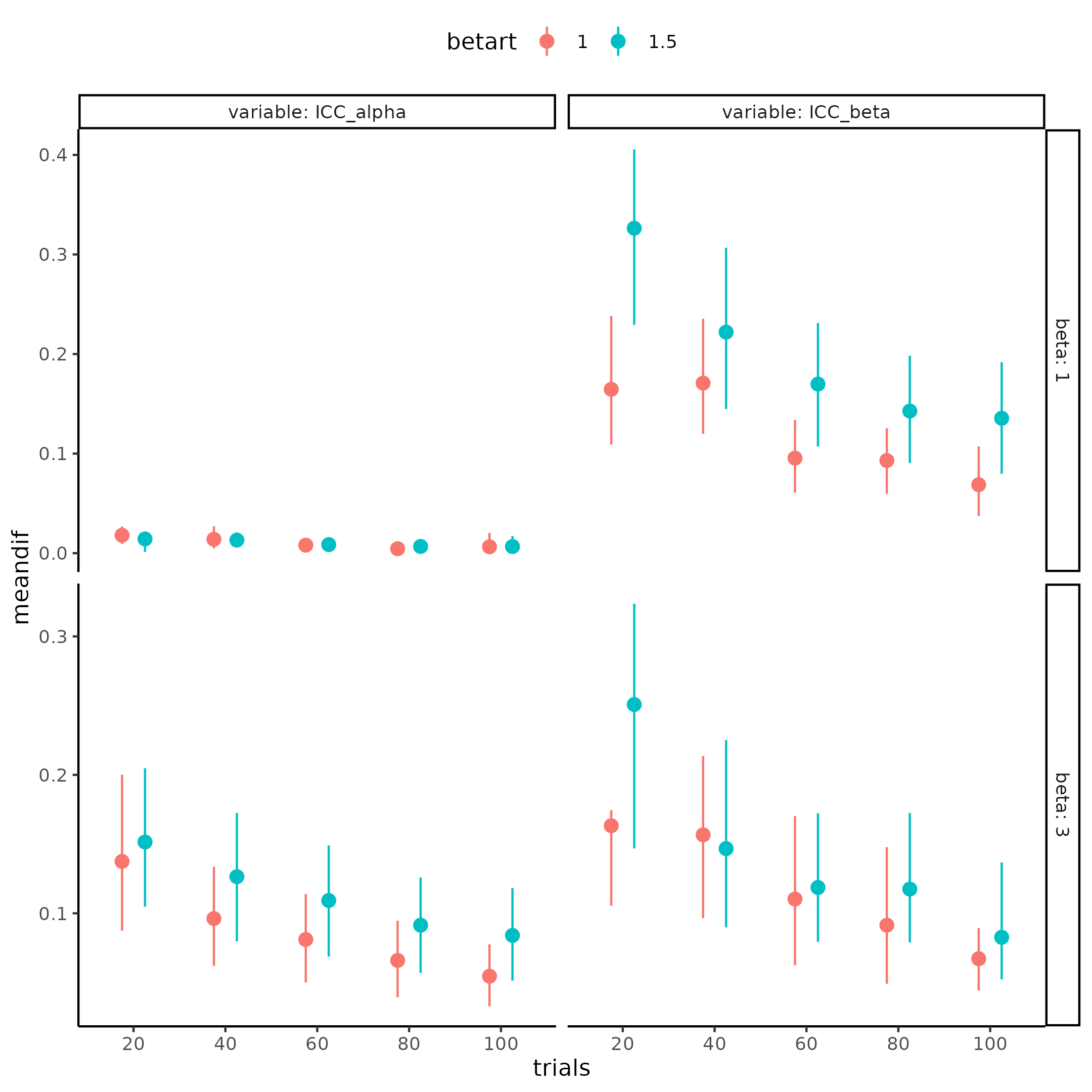


Figure 141 histogram of the mean difference between the ICC value obtained from using the reaction times or not, colors display the simulated level of coupling between the underlying psychometric function and the reaction times. Stronger coupling is associated with bigger parameter recovery effects for both threshold and slope, but not lapse rate.

## *Experimental data*

Having rigorously investigated how the psychometric function behaves and how the certainty of the parameters interacts with each other but also with the number of trials for each subject, one can now turn to real data. I’ll in this section introduce the published dataset that I will re-analysis utilizing the psychometric functions introduced above. The goal with this re-analysis is 2-fold. Firstly, it reiterates the fact that the assumptions about the structure of the data can make quite a difference in the parameter estimates and their uncertainty. Secondly, it will serve as a starting point to understand why the internal model validity using the ICC can be helpful as a metric to gauge how trials and subjects interact on the statistical power of a model to reject a hypothesis. This last aspect of testing hypotheses will tie together how these validity steps help determine the ability of a model to do what researchers are many times interested in i.e. hypothesis testing. The last point of the thesis is going revolve around conducting a thorough power analysis of the current model, utilizing the published dataset described below here I will compare the ICC metric for the model to its ability to reject hypotheses at certain trial and subject numbers. In this regard of conducting a power analysis I will again highlight where uncertainty creeps in and how we can deal with and account for these, as common practices are insufficient.

## *Heart rate discrimination task*

The article where the dataset was published is Legrand et al. ([2022](#ref-legrand_heart_2022)) and is an interoceptive task. Here the authors collected 223 participants who came in twice to complete a heart rate discrimination (HRD) task within 6 weeks between visits. The HRD task is comprised of two distinct tasks, a comparison and an interoceptive task. Here I’ll focus on the Interceptive task where participants were asked to internalize their own heart rate for 5 seconds. Meanwhile the participant attends to their own heart rate, the heart rate is monitored and calculated in real time. Next based on the observed heart rate participants will hear five auditory tones in a frequency (not the internal frequency of the tone, but the frequency of how fast the tones is presented) that is either faster or slower than their own objective heart rate. The amount this auditory tone frequency was faster or slower was determined by the PSI procedure introduced in the Adaptive design optimizing paragraph. This means that the stimulus value for the psychometric function is the difference between the external tones frequency and the observed heart rate of the participant in the current trial and the responses are given by faster or slower with faster being coded as 1 and slower being coded as 0. This means that a participant might have a heart rate of 50beats per minute (BPM) at a particular trial and then hear tones in a frequency of 40 BPM and are asked to respond whether they think this 40BPM is slower or faster than their own heart rate. The authors of the experiment, described above, ended up running single participant level models of each subject, for each session, and then correlating the slope and threshold of the psychometric function. They found a medium correlation between the threshold r = 0.51 p < .001 between sessions and a negligible correlation r = 0.1, p = .15 for the slope. In the next section I will show how this reliability might change given different model assumptions and different models, to demonstrate that thinking hard about what model is fitted is worthwhile.

## *The models*

In this section I will describe the models that I’ll fit to this big test-retest dataset to examine the influence of the model fit on the correlation between session one and two. The single fit model is going to be the references and going to be the same as the original authors did. That is estimating each individual for each of the sessions individually without a lapse rate (i.e. a two parameter psychometric function) and then post hoc correlating the estimates between session one and two. I will add the propagated uncertainty to these estimates as they do not seem to be adjusted by the authors. Next, I’ll investigate the same model as above but adding the third lapse parameter. The hierarchical model is going to model the two sessions from the same multivariate normal distribution. This model directly models the correlation between sessions as its included in the variance - covariance matrix of the multivariate normal distribution. The last type of model is the nested hierarchical model, this model assumes that all subjects have a mean level parameter which is drawn from the same multivariate distribution, then each parameter for each session is then drawn from a subject level distribution, identical to the model presented in the ICC parameter recovery section. For this last model the ICC is the statistical metric estimated by the model itself, and the correlation will afterwards be calculated. Additionally, each of these models will be fitted using the reaction times as described in the XXXX section to investigate the influence of adding this additional information. A final full model is going to be fit utilizing even more information already available in the dataset. This model will not only incorporate the reaction times on a trial-by-trial basis, but also the confidence ratings for each trial. These confidence ratings were included in the task of the original experiment to examine the participants’ interoceptive metacognitive abilities. These confidence ratings are going to be modeled in close resemblance to the reaction times, just inverted. This inversion is because in the middle of the psychometric function the uncertainty about the stimulus representation is the highest and therefore reaction times should be their highest as well, but confidence should be at the lowest. Another difference between the reaction times and the confidence ratings is their range of possible values. Confidence ratings were bounded between 0 and 100 indicating complete uncertainty and certainty respectively. A natural likelihood function for such kind of double bounded variables is the beta distribution as its already bounded between 0 and 1. The only problem with using this likelihood function is the edge cases of 0 and 1’s which for the confidence ratings are 0 and 100. One approach is to model these edge values separately using a zero-one-inflated beta distribution. This approach, however, models these edge values as separate processes which does not make sense in this case as the confidence ratings are meant to represent a continuous measure of confidence. I will therefore here subtract 0.1 from the 100 ratings and add 0.1 to the 0 ratings making it possible to use the beta distribution for the full range of confidence ratings. This approach of modeling the bounded ratings between 0 and 100 is tenuous and new mixture methods are slowly being developed for a more holistic approach see Kubinec ([2023](#ref-kubinec_ordered_2023)). Reaction times of the responses were at maximum 8 seconds and can therefore still be modeled by the shifted log normal distribution introduced above.

## *Results*

Table 3 displays the correlation coefficient between the first and second session for the threshold and slope for each model when uncertainty has been propagated using bootstrapping. For a full table of all parameters of all models as well as with and without uncertainty propagation see supplementary table XYX

| alpha | beta | lapse | model | structure |
| --- | --- | --- | --- | --- |
| 0.48 [0.34 ; 0.60] | -0.01 [-0.06 ; 0.05] | FALSE | Binary | Single |
| 0.50 [0.39 ; 0.62] | 0.01 [-0.06 ; 0.10] | TRUE | Binary | Single |
| 0.52 [0.39 ; 0.64] | -0.00 [-0.10 ; 0.06] | FALSE | RT | Single |
| 0.52 [0.39 ; 0.65] | 0.01 [-0.11 ; 0.10] | TRUE | RT | Single |
| 0.52 [0.41 ; 0.64] | 0.00 [-0.14 ; 0.17] | FALSE | RT+Conf | Single |
| 0.52 [0.40 ; 0.62] | 0.03 [-0.13 ; 0.22] | TRUE | RT+Conf | Single |
| 0.51 [0.41 ; 0.59] | 0.23 [0.01 ; 0.45] | TRUE | Binary | Hierarchical |
| 0.49 [0.39 ; 0.58] | 0.26 [0.09 ; 0.41] | TRUE | RT | Hierarchical |
| 0.48 [0.39 ; 0.57] | 0.21 [0.06 ; 0.36] | TRUE | RT+Conf | Hierarchical |
| 0.54 [0.49 ; 0.58] | 0.19 [-0.01 ; 0.42] | TRUE | Binary | Nested Hierarchical |
| 0.54 [0.51 ; 0.57] | 0.27 [0.07 ; 0.45] | TRUE | RT | Nested Hierarchical |
| 0.55 [0.53 ; 0.58] | 0.20 [0.06 ; 0.37] | TRUE | RT+Conf | Nested Hierarchical |

Table 3 clearly highlights the fact that the additional assumptions of the hierarchical models both nested and unnested increases the session-by-session correlation of the slope of the psychometric function, and that additionally including the reactions time increase the correlation even more, however still with overlay in the 95% credibility intervals. The main difference in session-by-session correlation between the two hierarchical models can be found in the threshold as the nested hierarchical model outperforms the non-nested hierarchical model in this regard. A concern of this approach of just looking the correlation coefficients is of cause that a model with a high session by session correlation might not fit the data the best, as the latent underlying stability i.e. correlation might be 0. One approach would therefore be to examine model fit using common metrics such as cross validation, information criterion etc. This difficulty here is that most of the models are incompatible; because they have been fit to differing amounts of subjects in the case of hierarchical vs single fit models, and to differing amounts of dependent variables in the case of within model architecture. The only models being compatible for comparison are the two hierarchical fit models with the same model architecture, resulting in a very limited comparison.

## *Power analysis*

As a last step in this model building, formulating, and testing framework we can easily and accurately inform future work. When researchers are interested in the parameter values of their models, they are many times, especially in computational psychiatry, also interested in how they differ using pharmacological interventions or between healthy controls and patient populations. The question in such a scenario is how many participants and or trials do you need to reliably detect a particular size of effect, between the two conditions? In the next section I outline the idea of power analyses and how it can be conducted within the framework of cognitive modeling. To do so I will utilize the results from the above analysis of the test retest dataset of ([Legrand et al., 2022](#ref-legrand_heart_2022)). Given the due diligence of the model validation steps it will be easy for the person using the model to get an accurate estimate of what effect size can be detected with differing amounts of trials and subjects. For the current analysis I’ll be using the group level estimates from above here focusing on the simplest model for computational efficiency. Before conducting this power analysis a few details about power and power analyses should be explicitly highlighted.

## *What is Power analysis*

The idea about power analysis is simple; we want to calculate a priori to conducting our experiment, the probability that our results are going to be “significant” given that there is some “real” underlying effect. Usually this is depicted in a 2 by 2 matrix with the real latent effect being in one dimension and the model results in the other dimension. The probabilities of landing in either of these 4 categories is usually described as functions of our statistical significance threshold (alpha / p-value) and the power (1-).

|  | Reality (effect) | Reality (no effect) |
| --- | --- | --- |
| Model result (significant) | 1-β | α |
| Model result (in-significant) | β | 1-α |

The latent or underlying effect might be our intervention or a difference between a healthy and a patient population. The framing of power analyses is then to say that results are significant if the p-value is less than a particular value most often 5% and the probability that we detect the effect given that it is there is another arbitrary value with 80% being the standard. Moving to the more practical side of the power analysis our cognitive or statistical models will reject and fail to reject different rates of effects given magnitude of this effect. The table above is therefore quite misleading as in reality the dimension of reality is a continuous variable of size of the effect and our models have a particular probability of rejecting a hypothesis (given subjects and trials) at a particular effect size its tested on. An example of this might be that we want to detect whether there is an effect of gender on height in the human population. We assume that there is an underlying effect and observe X females and Y males and run a statistical analysis to determine whether we can reject the null hypothesis (there are no differences in height in the two genders). Compare this to the hypothesis that there is an effect of age (late adolescents and adults) on height. The former difference might in general be much larger than the former and therefore with all else being equal (trials, subjects, statistical model etc.) this difference will be easier to detect compared to the other difference. What is therefore done when conducting power analyses is that different observed effect sizes are simulated in differing number of trials and subjects and the ability of the statistical model to reject these simulated experiments are then accessed. Usually, this amounts to then counting the number of times, the model achieves “significant” results compared to non-significant results, which is the power of the model at that number of trials subjects and observed effect size. This approach accurately captures how we expect the model to behave when we fit the data to the model after obtaining it. It tells us if we observe a particular effect size, we will with x percent be able call the results significant. The utility of this analysis is therefore to be able to examine how many subjects are needed to obtain a statistical power of usually 80% given that an effect size in the population is present, this effect size in the population might then be informed by previous studies and or meta-analyses. Extra assumptions are then needed to approximate the distribution of effect sizes as these statistical metrics also have uncertainty associated with them. This extra aspect many times disregarded or forgotten will be expanded upon later while showing how to incorporate it.

The power simulations here will be for a repeated measures design interested in a difference in threshold due to some intervention. I therefore simulate subjects, trials, and effect sizes in a variety of combinations see figure 16. Here the effect size chosen was cohens’ which formula can be seen in equation below. This meant that the variance of the second condition (i.e. after the treatment) had to be specified, such that as simulated effect sizes increased the mean difference only increased. The choice here was that second session variances was 1.5 the variances of the first condition. See supplementary XYX for explanation for the choice and why this choice is arbitrary and does not meaningfully influence the interpretations further on. The simulation process followed the following procedure: first two set of agents were simulated from a multivariate normal distribution with group level parameters of the binary nested hierarchical model presented above. The second set of agents had their threshold increased by a random variable that was drawn from the difference distribution defined by 1.5 times the first session variance and the effect size i.e. equations below. To ensure a particular observed effect, size this process was repeated until an observed effect size of the desired value was obtained ( 0.01), this step of resampling for a particular effect size was mainly for visualization purposes later on. Next each agent was put through the pathfinder algorithm to get their trial-by-trial responses and stimulus values. The full trial-by-trial dataset was then fitted using a simple hierarchical model where the threshold was parameterized as a linear combination of intercept and session with dummy coding of session.

Mean and standard deviation of the difference distribution between the two sessions, where is the variance of session 1 is the variance of session 2. is the correlation between the two sessions.

## *Power analysis results*

Given that the space of trial and subject combination is in the extreme, infinite, and at even a practical level quite huge. What I will show here is that the variation in how well the model rejects the null hypothesis, given subject and trial combinations are quite stable over observed effect sizes of the given model, making it possible to give good predictions i.e. extrapolating from the simulations. I will use the decision threshold of saying that a result is significant if less than 5% of the posterior distribution of the difference in the threshold crosses 0 like setting an alpha value of 5% in a frequentist power analysis. This 5% is a reflection of what is currently used in the field as the standard decision threshold but can easily be modified (see github). To properly display the raw results of the power analysis where it’s possible to compare the effects of trials and subjects I will use the beta distribution to properly display a summary of the 100 simulations for each effect size. The beta distribution is a two-parameter distribution that can be parameterized in different ways ([Alshkaki, 2021](#ref-alshkaki_six_2021)). The utility of the beta distribution to display the results of the power analysis is that one parameterization of this distribution involves how many times an event happened that we cared about and the other parameter being how many times this event did not happen. This therefore makes it possible to start with a uniform prior on the probability of rejecting the null hypothesis i.e. Beta(1,1) and then updating this probability density function with the amount of hits and misses here significant or non-significant results. This results in a PDF that contains all the information in each of the 100 binary points (i.e. significant or not). Figure 16 shows each trial subject combination with points representing this prior uniform beta distribution updated by the 100 datapoints that were either deemed significant or non-significant. Three main things are of particular importance. The shape of the points very closely resembles a psychometric function where subjects and trials influence both the steepness and the location of the function. Secondly, Increasing the number of subjects has two important features, it shifts the points towards higher power with lower effect sizes, but it also seems to increase the sensitivity to the effect size, i.e. the slope of the curve is getting steeper with higher number of subjects. The amount that trials for each subject also matters for the shape of the curves, in figure 16 its quite clear that increasing trials if very low i.e. 10, makes a big difference in the shape of the function, but the difference in going from high to very high i.e. 100 to 150, does not matter much. The tendency of the function to be less affected by ever increasing trials is also present for the number of subjects. This observation makes sense if one takes the function to its extremes in trials and subjects. Increasing subject and trials to infinitely many, we would expect, assuming the model has been shown to become increasingly better with increasing trials (like with the ICC metric presented above) that the model would be able to pick up on even the tiniest difference in groups. This would essentially mean that the function would consistently be at y = 1 with x approaching 0 from the positive direction and then jump to (0,0) in the (observed effect size, power) curve as no difference would entail no power. In the other extreme were no subjects or trials are present the curve should approach a flat line at y = 0 entailing no power for any amount of effect size. Essentially reaching a step function in the limit when x goes to 0 and subjects and trials goes to infinity I.e.

These observations are what is going to be used in the next section in order to extrapolate the results from figure 16. Making it possible to construct a model that will map trials, subjects and effect sizes to power.

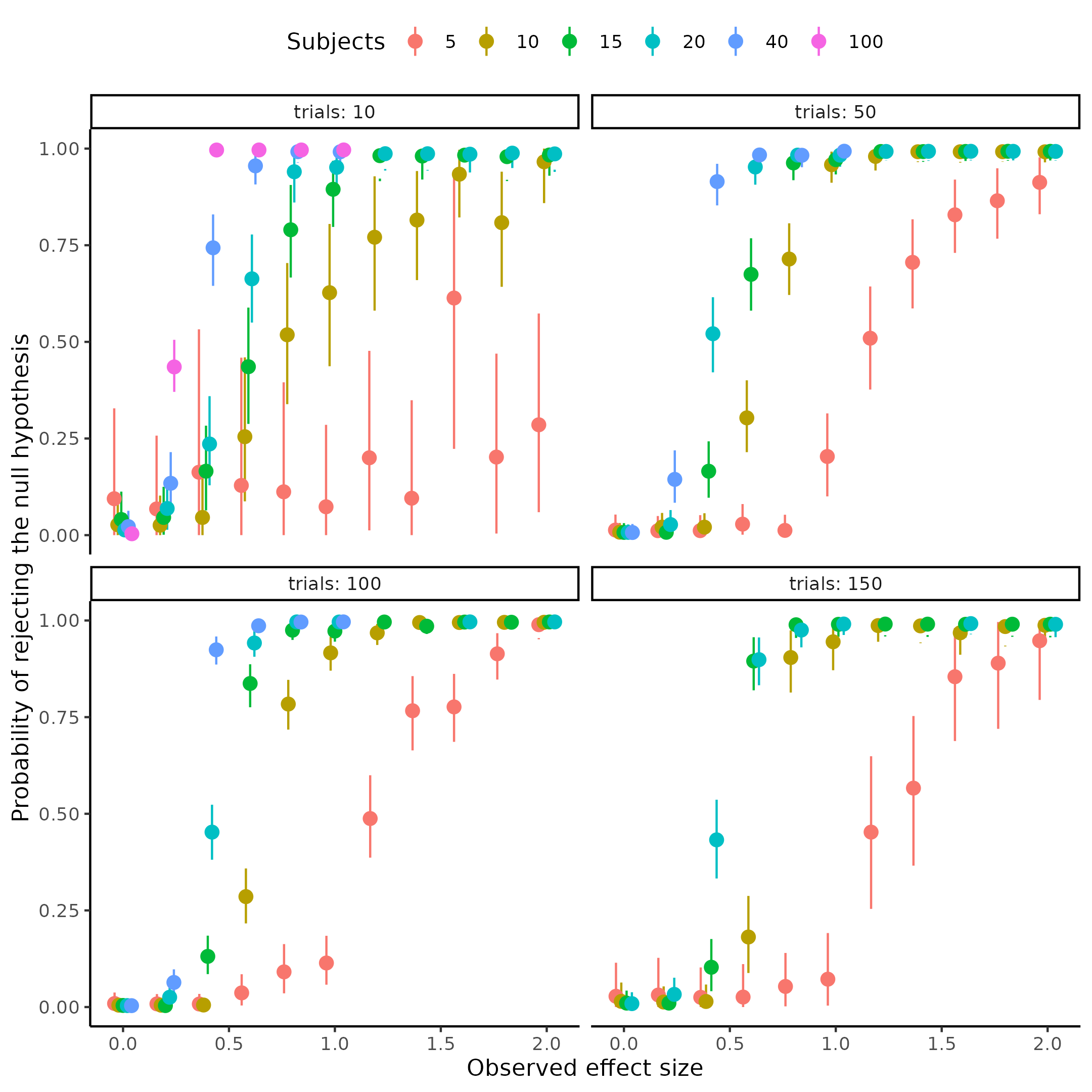


Figure 16, depicts power as a function of observed effectsizes in different combinations of trials and subjects.

## *Modeling of power analysis*

To use the information from above the latent psychometric function describing the relationship between subjects, trials and effect sizes needs to be investigated. Ideally a psychometric function that enforces the curve going through the origin, as no observed effect size should always entail no power. Next the parameters of these psychometric functions i.e. the threshold and slope need to be parameterized by the number of trials and subjects. Before fitting the general case and ensuring that a psychometric function is well fitting to the problem at hand, I start by fitting each set of trials and subject combinations independently to the parameters of the psychometric function. This amounts to fitting trials and subjects as factors in a linear regression framework (see supplementary for further explanation). This will help ensure that the fitted functions do pass through the points depicted in figure 16 and increase the faith in the next type of modeling. For this it is also possible to fit several kinds of psychometric functions and then compare them on their performance of predictability, because that is what we in the end care about here. One way to compare these models is using the Pareto smoothed importance sampling leave one out cross validation (PSIS-LOO-CV) ([Vehtari et al., 2017](#ref-vehtari_practical_2017), [2024](#ref-vehtari_pareto_2024); [Yao et al., 2018](#ref-yao_using_2018)). The three types of psychometric functions that were fit, were the cumulative normal, the cumulative logistic and the cumulative Weibull function. The main differences between the normal and logistic function is that the logistic function has heavier tails than the normal allowing for more disperse observations. The difference between the Weibull and the two other distributions is that the Weibull function is forced through the origin and its shape therefore quite different from the two other functions. The choice of the cumulative normal or logistic function does not necessarily violate the assumptions laid out above because of the way that the parameters are going to be dependent on the trials and subjects. This is clear if one considers an asymptote at y = 0 for the slope and threshold. This exactly matches the observation from above that the psychometric function moves closer and closer to a step-function (as the slope gets closer to 0) and that the location of this step function approaches x = 0 but never reaches it, if the asymptote for the threshold is not zero but close to. The results of this preliminary analysis can be seen in figure 17 where the independently fit logistic psychometric functions are overlaid on the observed datapoints from figure 16. The figure clearly shows well fit for most of the trials and subject combinations.

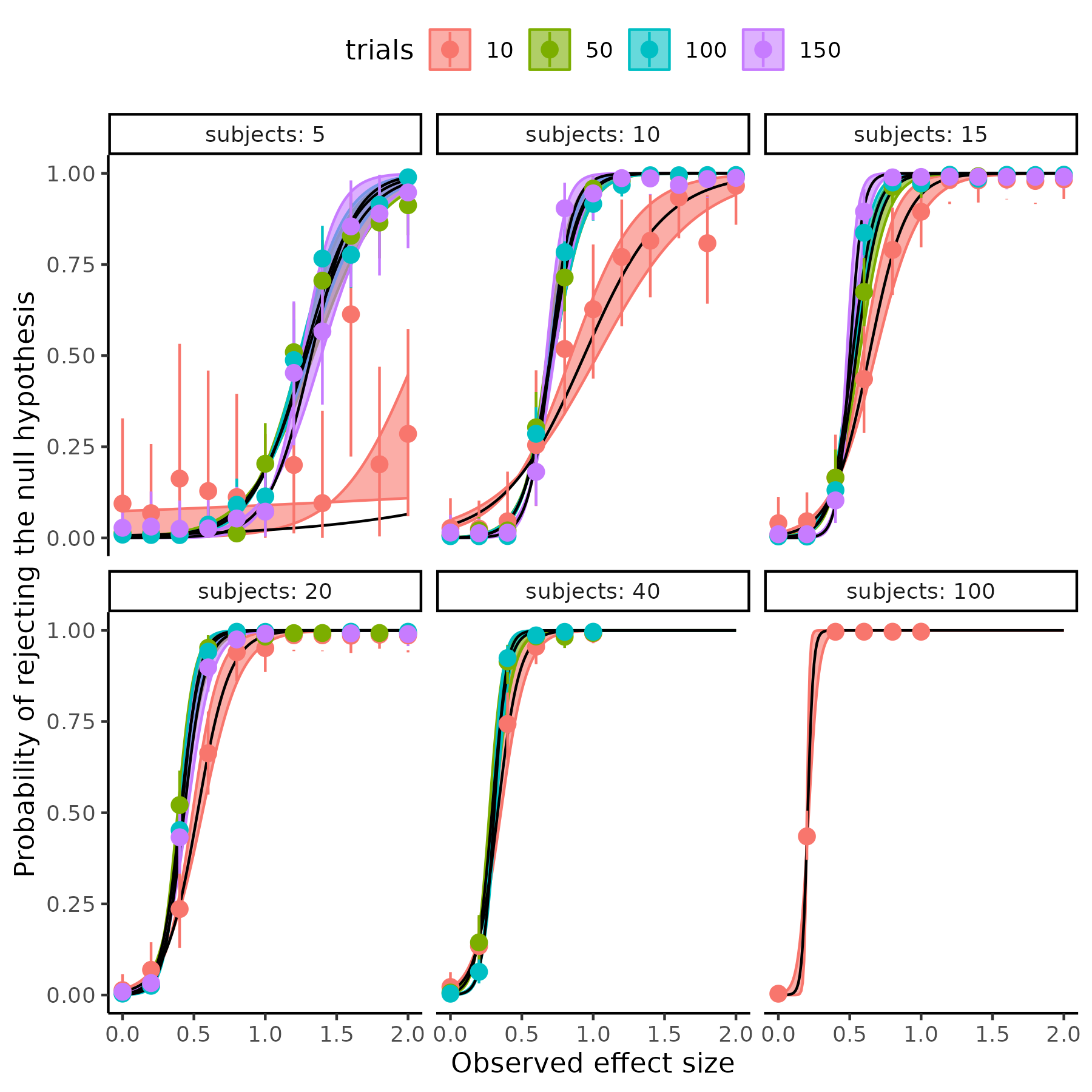


Figure 17; depicts power as a function of observed effectsizes in different combinations of trials and subjects. With lines being independently fit logistic psychometric functions to each trial by subject combination.

## *Continuous mapping of the power analysis*

Moving to the continuous mapping of subjects and trials to the psychometric function that maps observed effect sizes to power, one needs to define the function that relates subjects and trials to these parameters. Given the observations above, that the steepness of the function increases with increasing trials and subjects and that the threshold moves towards 0, a first choice of this mapping would be to model the two parameters as exponentially decreasing by trials, subjects, together with their interaction. An exponentially decreasing function in the complete general case would mean the following relationship.

Where represents the parameters of the psychmetric function i.e. slope and thereshpld, is the value of the parameter when the number of trials and subjects approach infinity. is vector of parameters determining the steepness of the exponential decrease from the covariates in the matrix X, here trials subjects and their interaction. The parameter serve, together with , as the value of the parameter when trials and subjects are 0. Another formulation of the dependency might be a power law equation XX.

Both approaches can produce the observed behavior and their difference depends on the underlying relationship between the parameters and X i.e. trials and subjects and perhaps their interaction. The exponential equation assumes that as trials and subjects increase by a fixed amount the parameters will decrease by a percentage, whereas the power law assumes that as trials and subjects increase by a percentage the parameters will decrease by a percentage. There are several ways of investigating which of these two approaches results in the better fit, firstly plotting the parameters of the independent fits vs trials and or subjects, which was conducted in the “modeling of power analysis” section, on two different coordinate systems either in (log(y),x) or (log(y),log(x)). Which of these produces the best-looking linear fit / line would be the best candidate. Figure 18 displays the three functions fitted independently on each of the two coordinate scales.



Figure 18. Parameters estimates of the individually fit psychometric functions (columns) on trials and subjects. The top row depcits a (log(y),x) coordinate system whereas the bottom row a (log(y), log(x)) coordinate system. A straight line relationship between subjects (trials) and the log of the parameter value (top row) would indicate exponential relationship whereas a straight line in the log;log coordinate system would imply a power law relationship in the native (x,y) space.

Another approach would be to fit both types of models and then compare them on LOO-CV. Doing this, displayed problems with 15, 25 and 3 % of observations for the normal, Weibull and logistic function respectively as the pareto k diagnostic value was above 1 for these percentages of datapoints. This essentially makes it meaningless to compare the functions ([Vehtari et al., 2024](#ref-vehtari_pareto_2024)). Moving forward only the logistic cumulative function was used as this was the only model that produced the least amount of problems with pareto k values when fitting trials and subjects as continuous variables, for a complete set of models including the normal and Weibull see supplementary XXX. The first logistic model was the exponentially decreasing function equation above. Four other models were fit with different ways of parameterizing the power law equation above. These four models displayed different ways of how the trials and subjects interact as there is no straightforward way of combining X and β. The first was an additive model with the following parameterization.

The second was with a combination of additive and multiplicative operations:

The third was a multiplicative model without an interaction.

The last was the multiplicative model with an interaction but defined as the sum of subjects and trials as the normal interaction of multiplying trials and subjects would lead to a similar model of the model without an interaction.

Comparing these four models using loo indicated that the best model was the last model but closely followed by the second model, which can be seen in table 5. Importantly for these reported models the diagnostic values were all below 0.7.

| models | elpd\_diff | se\_diff |
| --- | --- | --- |
| logs\_power | 0.00000 | 0.00000 |
| addititve\_multiplicative | -10.05930 | 4.65188 |
| logs\_power\_noint | -35.91091 | 8.86807 |
| logs\_expo | -85.85955 | 15.73949 |
| additive | -820.01552 | 37.47927 |

This indicates that as the trials and subjects increase by a percentage the parameters of the psychometric decrease by a percentage as the top two models are both variations of the power law. To ensure that the tested models still capture the underlying data, figure XYX displays the winning model imposed on the data with 95 credibility intervals of the mean. As can be seen this closely resembles the individual independent fits, with the most drastic deviation in the 5 subjects 10 trials condition

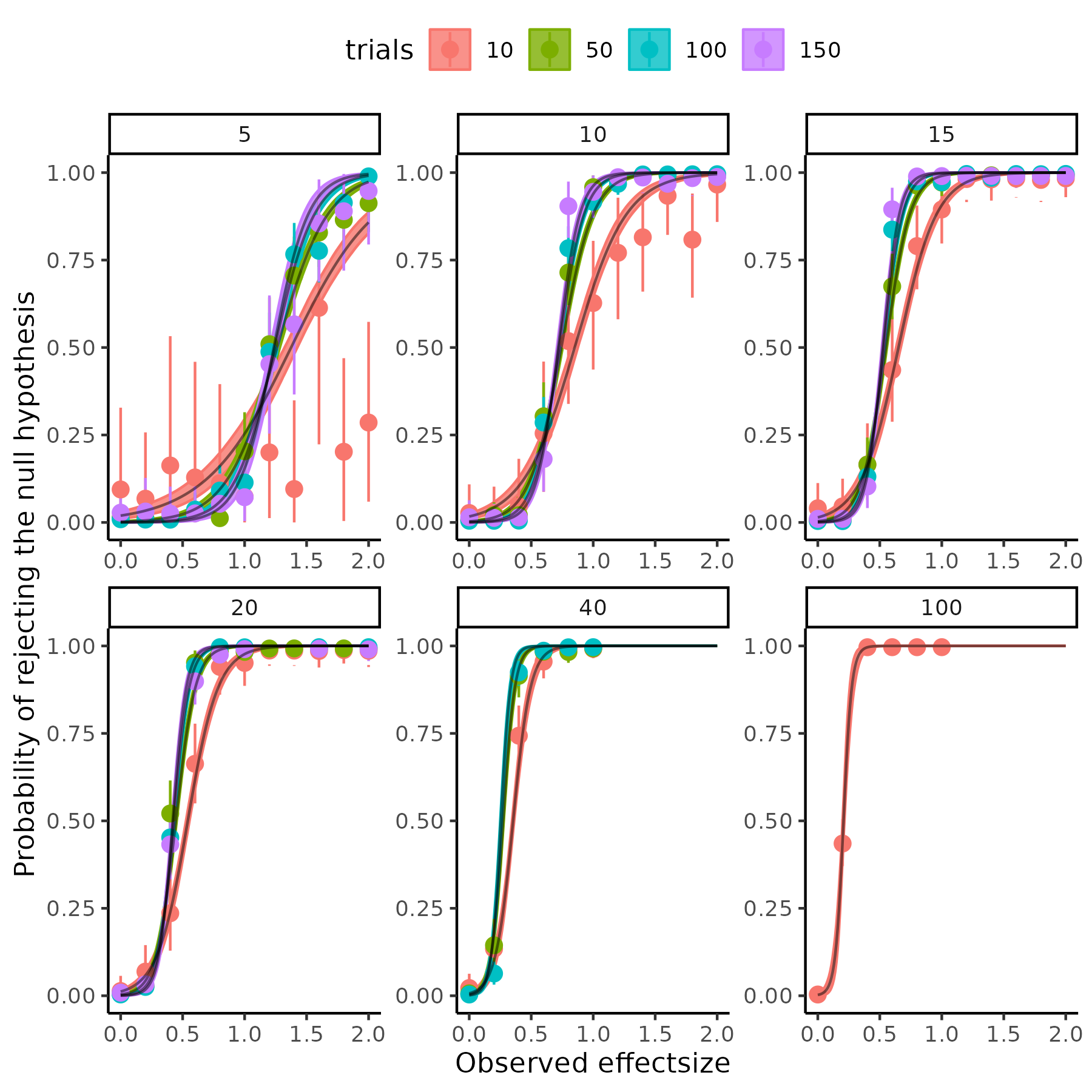


Figure 19; depicts power as a function of observed effectsizes in different combinations of trials and subjects. With lines being the dependently fit logistic psychometric functions to each trial by subject combination.

The marginal posterior distributions of the parameters of the winning model are displayed below in figure 20

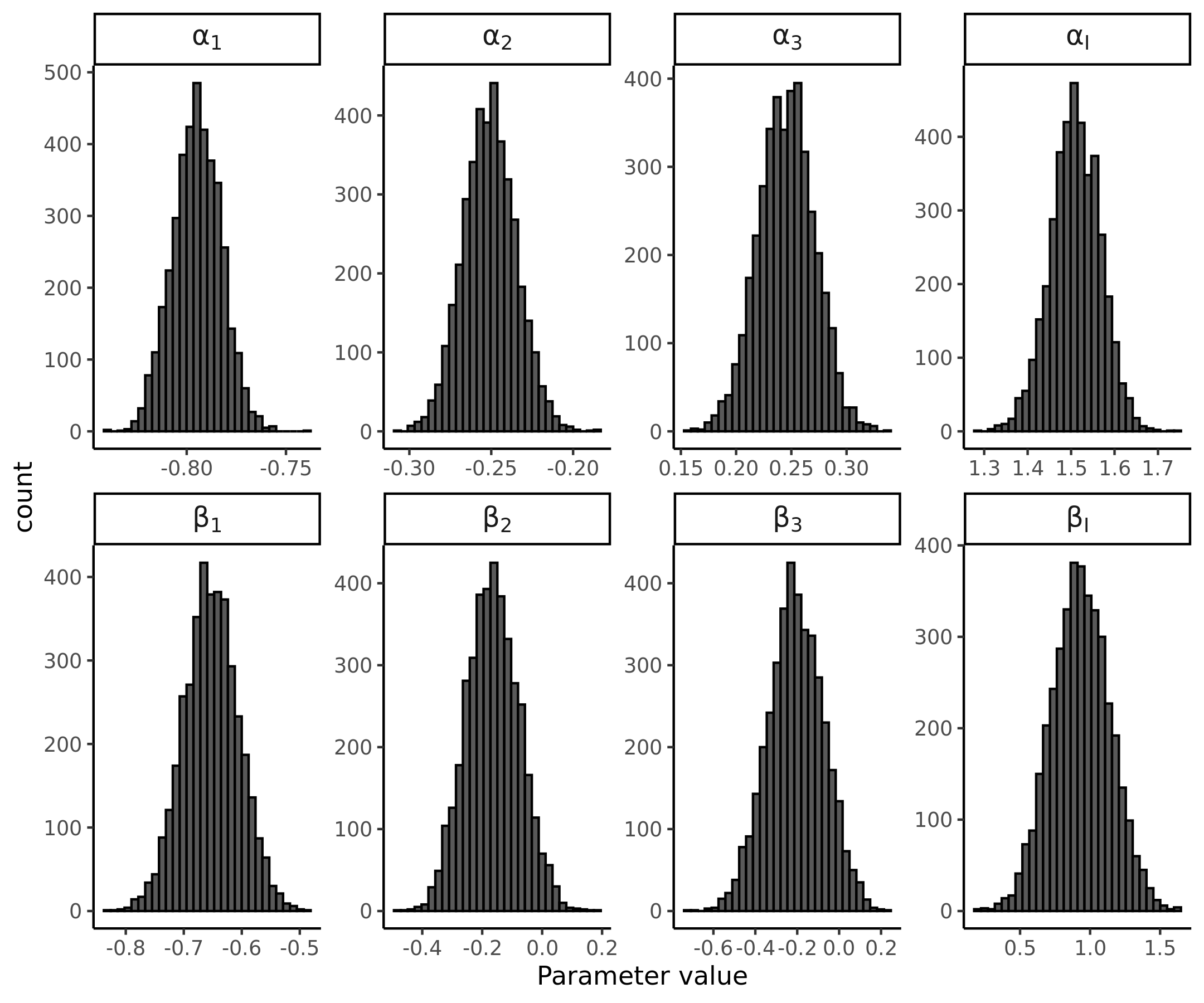


Figure 20. Marginal posterior distributions for the winning model’s parameters

Meaning that the resulting best guess of the underlying function transforming trials, subjects and observed effect sizes into a probability of rejecting the null hypothesis of no differnce in threshold is as follows:

Where

Where each of these parameters are given by the distributions depicted above.

## *Utility of the power analysis*

As alluded to in the beginning section of the power analysis, the work presented here would be able to help independent researchers determine the probability of rejecting a particular observed effect size using this model, given trials and subjects. However, if this researcher wants to know the probability of rejecting a null hypothesis given that they assume a particular effect size in the population further assumptions needs to be made, as the effect size when conducting an experiment is not a fixed quantity. In practice this means that when conducting an experiment, an effect size is drawn from the latent effect size distribution. Mathematically this means that the effect size that is observed in an experiment is given by a probability. The mean and standard deviation of this probability density function is given analytically by Cohen which could also be derived from bootstrapping as was done with the measurement uncertainty ([Goulet-Pelletier & Cousineau, 2018](#ref-goulet-pelletier_review_2018); [Hedges & Olkin, 2014](#ref-hedges_statistical_2014); [Lakens, 2013](#ref-lakens_calculating_2013)).

Assuming that the effect size is normally distributed we get

The probability of rejecting this sampled effect size is given by the function that was obtained above.

What we ideally want to know here is the probability of observing a particular effect size and that we can reject the null hypothesis given this observed effect size. Probability theory and particularly conditional probabilities gives us the relationship between these quantities.

Here represents the probability that we are interested in, i.e. rejecting, and observing a particular effect size.

Integrating over all possible values of the effect size is now necessary to essentially integrating out the effect size, also known as marginalizing.

Which becomes

Instead of trying to analytically solve this integral analytically, we again use the power of the computational resources to approximate the integral by taking draws of the normal distribution of the observed effect size and then putting them through which will give draws from a probability distribution of rejecting the null hypothesis. As a last step it is then possible to calculate the proportion of rejected null hypotheses to the total number of draws giving us the power of the study assuming the mean difference and variance in the two sessions.

## *Practical implementation of the power analysis.*

The above high-level explanation of calculating power for an experiment might be quite difficult to understand and therefore implement for independent researchers. To make this more accessible I will here demonstrate how this can be done using what has been provided up until this point. This section will therefore hopefully provide a practical understanding of what different parts should go into a power analysis and how different factors will influence power. Firstly investigating how the sampling distribution of effect sizes changes based on subjects and session by session correlation. I will here assume that the group mean difference of the threshold in the psychometric function is -5 and the variance in the second session is 1.5 times the variance of the first session, i.e. assuming that the intervention increases variation in the threshold, but that there is a clear effect (for reproducibility and ease of use the GitHub repository provides a function that does what is described below by inputting these assumptions).

Firstly, investigating the assumptions for the choice of mean difference and difference in variance can be visualized by repeated sampling from a multivariate normal distribution with the following parameterization:

Here and are given by the large test-retest reliability analysis (here rounded) and are -8, 8 respectively. Given our assumptions and are therefore -3, 10. We can then vary the number of subjects i.e. draws from this multivariate normal and the correlation coefficient ρ to see the effect on the distribution of effect sizes i.e. p(d\_obs). Highlighting the sampling distribution of the effect size, and the factors influencing it. The results can be seen in figure 21 highlights the fact that both the sample size i.e. subjects, but also the correlation between sessions is vitally important for the variances of the observed effect size.

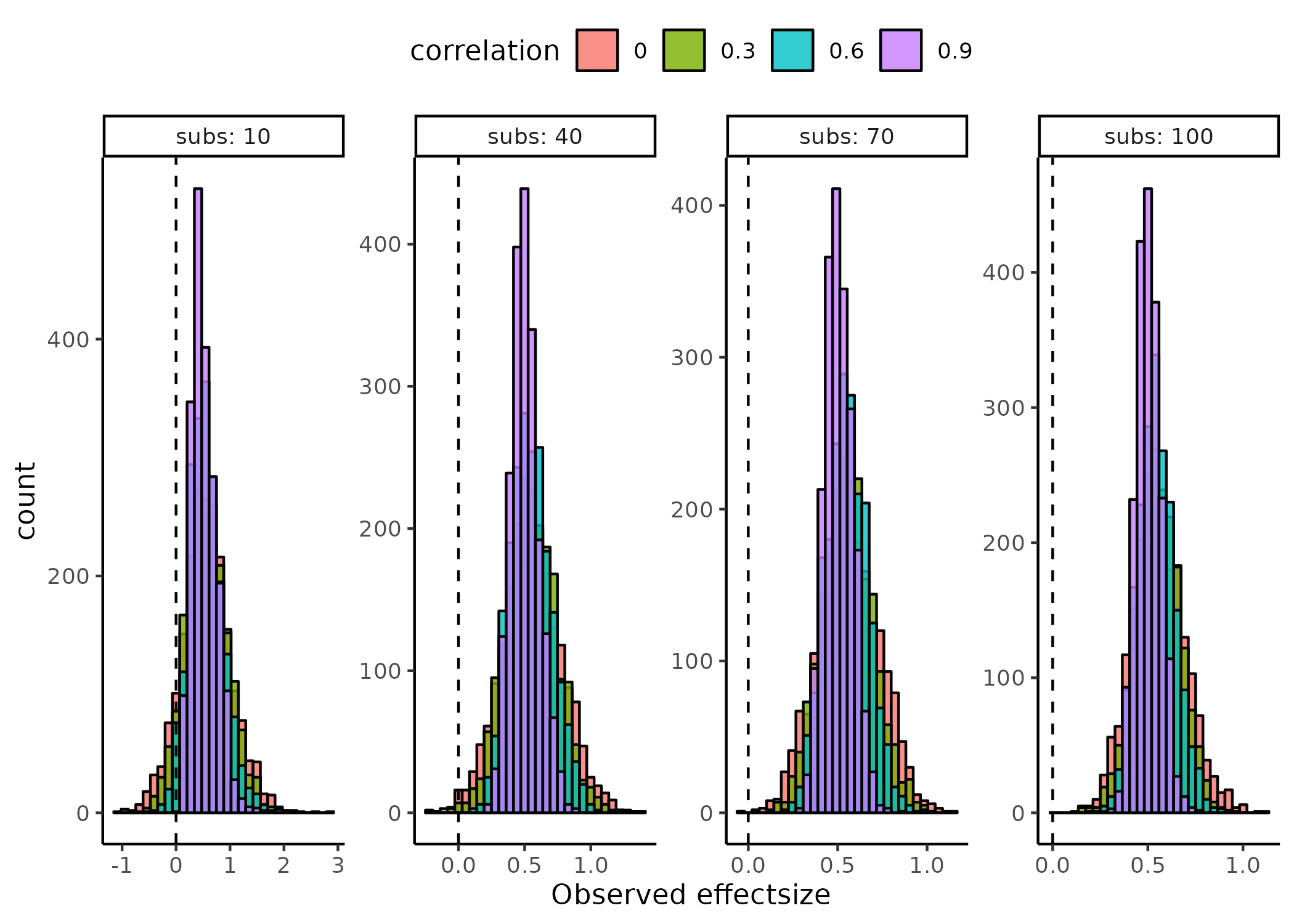


Figure 21. Sampling distributions of effectsizes across subjects (facets) and session by session correlations (colors)

Now we can visualize how these observed effect size distributions fit into the probability of rejecting the null hypothesis i.e.  Note that the observed effect sizes above are not dependent on the number of trials in the experiment i.e. assuming that they are observed with perfect precision, what the function derived from the continuous power analysis function does is that it incorporates this information together with other factors that might change the precision of the parameters. As shown above the implications of the function can be visualized as psychometric functions in a ( , ) coordinate system with trials and subjects being fixed at values. Another more informative way to visualize these implications for the current purpose is to make a 3-dimensional grid of (Subjects OR trials , , ) with facets of the last variable of either subjects or trials. This visualization can also serve the purpose of projecting the above distributions of observing a particular effect-size unto the space of Figure 22 displays the projection of the histograms as ellipse where the vertical width of the ellipse (the major axis) is given by the 95% Highest density interval of the histograms above and the horizontal width (the minor axis) is for visualization purposes, to see the underlying probability of rejection. The correlation is of cause informed by the test-retest reliability study which was found to be 0.54 [0.49; 0.58] for the current model.

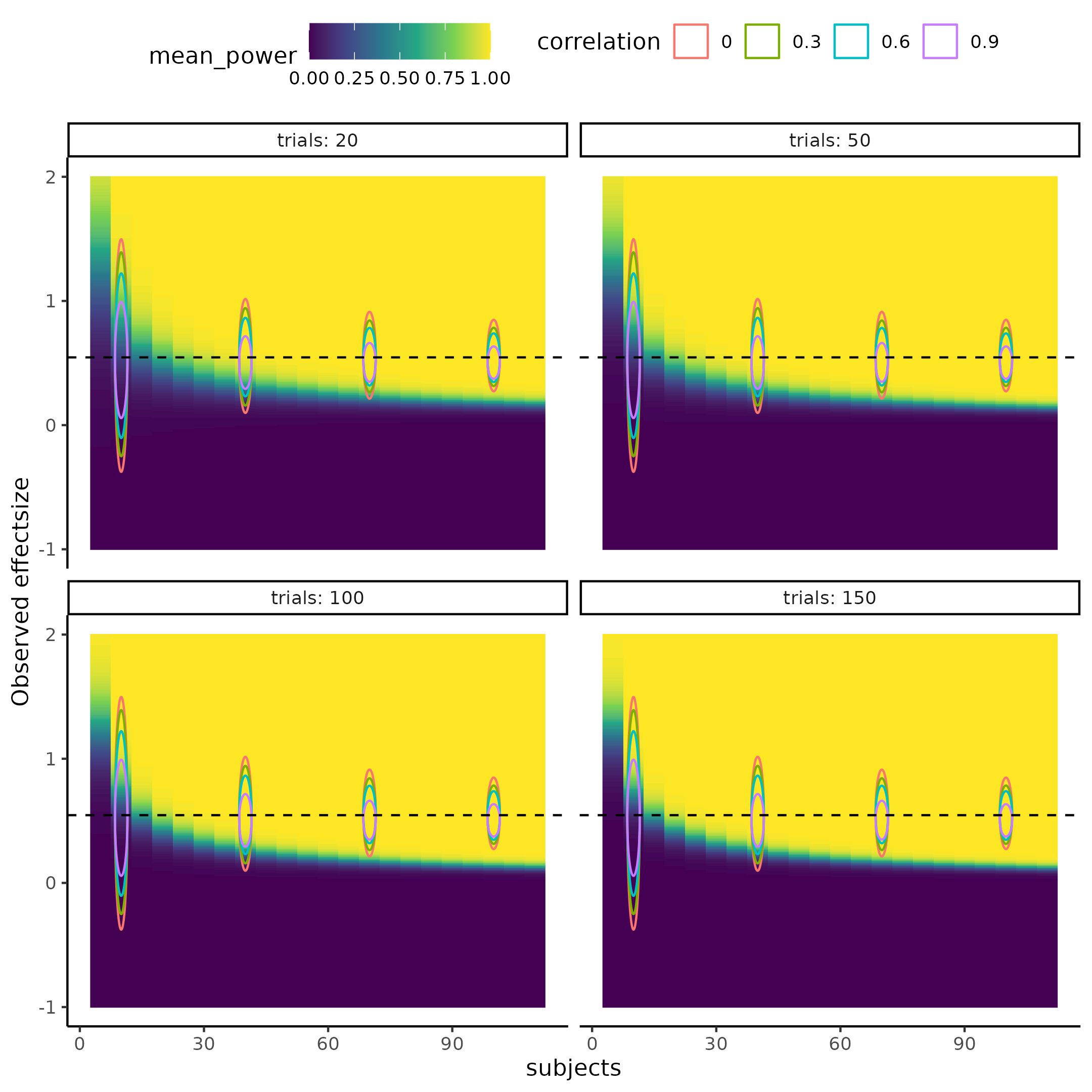


Figure 22. Visualization of how the power of a particular study is informed by the session by session correlation and the number of trials and subjects.

Now turning to the particular example of a researcher wanting to conduct a power analysis uterlizing the information provided. Two assumptions have been made, either a mean effect size or a mean difference of the intervention is assumed, and the variance introduced by the intervention. Here we expect a medium effect size of the intervention of and that the intervention does not increase variability meaning that the variance in both groups should be equal. To fully appreciate the power of this approach one could even imagine sampling these values i.e. 0.5 and the variances of the second session as random variables and not as point estimates, the most obvious case where this could be implemented is when effect sizes from meta-analyses are used for the best guess of an underlying effect size estimate for the study. These effect size estimates from meta-analyses namely come with uncertainties and neglecting this should not be advised!

Using equation above, it is possible to derive the mean difference and therefore simulate observed effect sizes which are then put into equation XXX and the probability of rejecting that draw is calculated. Repeating this process over the 4000 draws of the posterior distribution of and calculating the ratio of rejected to failed rejected null hypotheses gives an estimate of power including all uncertainty. In the case of not including the prior probability of the effect size the effect size estimate is just repeatedly entered as 0.5. As can now be seen in figure 23 the observed effect size has been “integrated” out and a grid of subjects by trial span the space of power to reject the null hypothesis. The left and right column of figure 23 quite clearly display the difference between accounting for the sampling process of effect sizes with the left not accounting for the sampling process. As a reference frame in figure 23 the red dashed line with subjects = 20 depicts the results from plugging the same assumptions i.e. Here and and , into the widely used and cited statistical software tool G\*power ([Faul et al., 2007](#ref-faul_gpower_2007)), which is widely used for power analysis of simpler designs. Further reiterating how and why uncertainty propagation is vital to designing studies of adequate power.

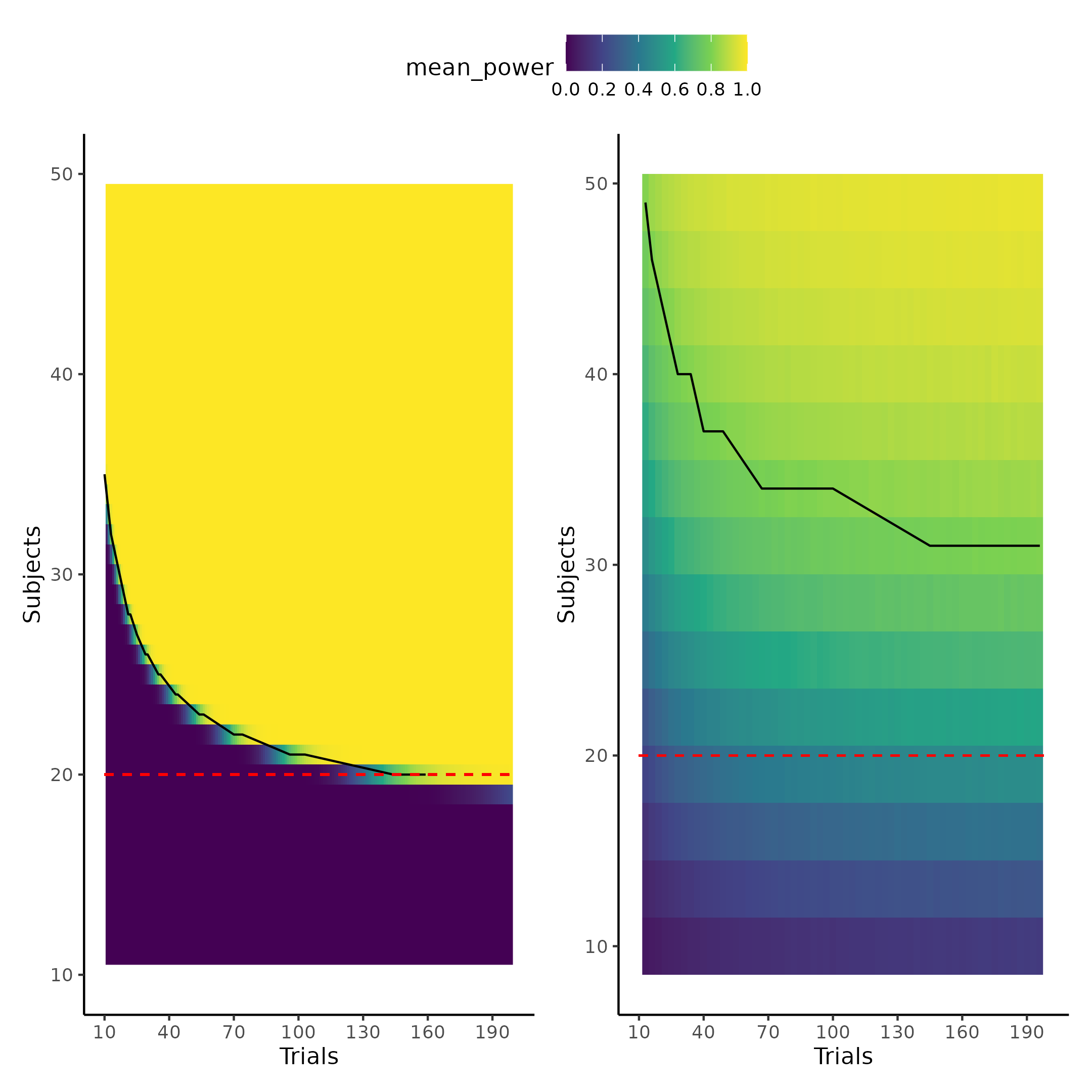


Figure 22.

# Discussion

The current thesis has investigated how the handling of uncertainty in the field of cognitive science and especially in the developing field of cognitive modeling can be improved. The thesis has done this by demonstrating that using computational resources in the form of simulations, a deep mathematical understanding with rigorous closed end solutions is not necessary to get a deeper understanding of how uncertainties on each level can and will influence every statistical metric. The thesis outlined three critical types of uncertainty; measurement uncertainty being the lowest level of uncertainty that is often completely neglected in the field, even though it influences the resulting statistical metrics in very unpredictable ways if the data has influential data points which might be associated with particularly high measurement uncertainty. Researchers should firstly be aware that measurement uncertainty is always present and examining the extent to which it can be safely ignored in their statistical models must be determined. Even in measures like reaction times which is commonly used in cognitive science ([MacLeod & Dunbar, 1988](#ref-macleod_training_1988); [Pirolli & Anderson, 1985](#ref-pirolli_role_1985); [Sternberg, 1969](#ref-sternberg_memory-scanning_1969)), there are measurement uncertainties, which depending on the soft and hard-ware the experiment might be a huge factor ([Crocetta & Andrade, 2015](#ref-crocetta_problem_2015); [Holden et al., 2019](#ref-holden_accuracy_2019); [Ohyanagi & Sengoku, 2010](#ref-ohyanagi_solution_2010)).

The thesis pointed to one aspect of cognitive science literature where measurement uncertainty is of difficulty, i.e. questionnaires, note however that the arguments laid out can be applied to any type of measure. Quite a literature exists on doing correlational or testing differences in populations groups from the results of questionnaires, which begs a question of the certainty of these questionnaire scores. This is especially true when considering some quite important aspects in handling these statistical problems. Firstly, questionnaires are easy, fast, and cheap to conduct when performing a behavioral experiment, subjecting many questionnaires to be collected in order “to see if something is there”. This curiosity is sometimes what derives sciences forwards, however in cases like these it will inevitably lead to false positive findings that cannot be replicated as the pressure for publishing results might hide the multiple comparison correction that should have been made when finalizing a manuscript. Including some uncertainty into the questionnaires would serve to push the significance barrier higher and make it harder to find significant results in these types of analyses. Perhaps a reasonable comprise would be that the added uncertainty on questionnaire scores should be proportional or just in general related to the internal consistency, measured by ICC, Cronbach alpha, correlation coefficient etc. of the questionnaire itself i.e. its test re-test reliability uncertainty.

Estimation uncertainty was introduced as the uncertainty associated with doing computations and is often displayed as standard error of statistical metrics. The main focus of the thesis was to use this understanding of uncertainties in the field of cognitive modeling and revise some of the methods and metrics used to validate these cognitive models. It was shown that using correlations, which has been used in many previous studies ([Schurr et al., 2024](#ref-schurr_dynamic_2024)), between simulated and recovered parameters values was not a particularly sensible metric to determine the extent of internal model validity. Two important things about the correlation coefficient made it insensible for internal model validity, the decision of choosing what size and uncertainty of correlation coefficient to deem model parameters sensible is not straightforward, because the interpretation of the correlation coefficient itself in the regard of model validation is not straightforward. This is particularly true when highlighting that the correlation coefficient is invariant of a linear transformations. The other reason was that in instances where the simulated and recovered parameter values did show good dependency the correlation coefficient rapidly approached the asymptote at 1 even when more information could be gained by increasing the number of trials, due to its limited inclusion of estimation uncertainty. The study therefore suggests that using a variant of the intra class correlation coefficient (ICC) as the statistical metric for examining internal model validity which has recently been suggested in the literature ([Schurr et al., 2024](#ref-schurr_dynamic_2024)). The thesis found that this metric was much more sensitive to estimation uncertainty in the parameters, with a sensible interpretation of the ratio between desirable to undesirable variance / uncertainty. With this new metric the thesis explored ways to decrease the undesirable variance and thereby increase the ICC metric, by incorporating smart experimental designs that are optimized for each individual on a trial-by-trial basis. Furthermore, the study showed how thinking generatively about the origins of the responses given in an experiment can increase the ICC metric without the need for extra trials, the concrete example given in the thesis was incorporating reaction times into the cognitive model describing how stimulus intensities are transformed to binary forced choices. This highlights an idea of jointly modeling several dependent variables and their interactions, that has been around for quite some time but only now is slowly gaining traction in cognitive science literature ([Hess et al., 2024](#ref-hess_bayesian_2024); [Pedersen et al., 2017](#ref-pedersen_drift_2017); [Stone, 2014](#ref-stone_using_2014)). The thesis highlights how all the above implementations and considerations do not necessarily have to rest on heavy mathematical understandings and proofs as computational resources and especially simulations has made it possible for people with coding experience to gain these insights by the power of (re) sampling; some of the implications of this will be discussed below. Lastly the thesis investigated and used data from a test-retest reliability study and showed that a reanalyzed could achieve better test-retest reliability by incorporating knowledge about the structure of how the data was gathered together with incorporating information already represent in the data. This data set was then used as an example of how power analyses of cognitive models could be conducted. This was done by first simulating and then fitting the cognitive model to many different simulated effect sizes in different trials and subject combinations. This approach allowed modelling the latent power curve that relates observed effect size trials and subjects to the probability of rejecting a null hypothesis in an experiment. Using posterior predictive checks and leave one out cross validation a particular power law related the parameters of the power curve to subjects and trials with good predictive abilities. Using this equation together with the many times overlooked aspect of sampling variability in the observed effect sizes when conducting power analyses, it was shown that incorporating sampling variability greatly increased the need for more subjects to achieve the same amount of power. This section also highlighted why and where the test re-test reliability of these metric matters as increasing test re-test reliability shrinks the influence of sampling variability in the observed effect sizes. Lastly the complete uncertainty propagated power analysis was compared to not accounting for sampling variability or estimation uncertainty on the parameters by using the widely used statistical software tool G\*power in a concrete example ([AARTS et al., 2015](#ref-aarts_estimating_2015); [Faul et al., 2007](#ref-faul_gpower_2007); [Ioannidis, 2005](#ref-ioannidis_why_2005)) . This comparison showed how G\*power estimation of sample size was equivalent to having close to infinitely many trials and not accounting for sampling variability i.e. disregarding much of the uncertainty inherent in the experiment.

## **Power analyses, certainty and replication crisis.**

In recent years many scientific fields, and especially psychology, social science and medicine has been under scrutiny due to a lack of and failure of replication of previous studies. ([Forbes et al., 2023](#ref-forbes_chapter_2023); [Wiggins & Christopherson, 2019](#ref-wiggins_replication_2019)). Many contributing factors has been laid out such as publication bias, questionable research practices such as doing statistical analyses until significant (p-hacking) or hypothesizing after the results are known (HARKING) ([Head et al., 2015](#ref-head_extent_2015); [Kerr, 1998](#ref-kerr_harking_1998)). However a quite paradoxical aspect of this replication crisis and the use of power analyses is that many times it is advised as one of the ways to increase the replicability of studies, as analyses of power of detecting small to medium effect size in social sciences have been found to be low to very low ([Felix Singleton & Fidler, 2023](#ref-felix_singleton_statistical_2023)). It is therefore argued that a reason for such low replicability, might also be due to very low probability of being able to detect the underlying effect i.e. low statistical power. The argument is sound as long as the analysis of power is accurate or accurate to a certain degree. What this thesis has highlighted is that the use of very popular tools like G\*power for conducting these types of power analyses will underestimate the number of subjects by a large margin. The problem is therefore that with the confidence of having done a power analysis there will be a false sense of certainty, just like the false sense of certainty about the measurements assumed by these popular softwares or measurements in cognitive science in general. Therefore instead of increasing replicability and certainty in the effects observed, utilizing these tools might paradoxically decrease them, as researchers might be tricked into conducting less powered studies due to the recommendations of the software.

Interestingly, quite a large number of scientists have suggested that moving the arbitrary statistical significance threshold to 0.005 instead of the commonly used 0.05 could be oan approach used to combat this replication crisis ([Benjamin et al., 2018](#ref-benjamin_redefine_2018)). Interestingly, lowing of the statistical threshold for significance would in practice lead to the conclusions drawn from this thesis of including and propagating uncertainty in most cases, of cause depending on the stucture and uncertainty measures of the data. These two approaches however have a very different reason to making these adjustments as the lowering of the significance threshold would be a means to an end, instead of addressing the underlying problem, which the authors also do acknowledge ([Benjamin et al., 2018](#ref-benjamin_redefine_2018)).

Another interesting idea that coincides with the general theme of the thesis to combat the replication crisis is that of preregistration, registered reports, and blind analyses.([Chambers & Tzavella, 2022](#ref-chambers_past_2022); [Evans et al., 2023](#ref-evans_improving_2023); [Klein & Roodman, 2005](#ref-klein_blind_2005); [MacCoun & Perlmutter, 2015](#ref-maccoun_blind_2015)). What all these types of interventions have in common is that they acknowledge the subjectivity in not only the data collection but also in the data analysis pipeline of the scientific inquiry. This subjectivity is both what introduces biases, but also what drives novel ideas, and a tradeoff between exploration and exploitation might be necessary to fully guard against unwanted subjectivity. What these interventions try to do is to have the analysis pipeline either fixed before data collection or have the data scrambled such that the results of the analyses are not known when producing the analysis pipeline. The rigorous checking, testing and validating of cognitive models the thesis outlined is not at stake with these interventions but facilitate them as they build on simulations. However, there are still considerations when analyzing the experimental data especially on the model convergence side, where in or excluding covariate or reparameterization of the models might be necessary. This is where the blind analysis intervention might be a valuable insight from physics, where experimental data is scrambled in various ways such that models and analysis pipelines can be done on data that resembles the collected data, but without being able to know the results before the data is unblinded. Decisions are therefore made on scientific justifications instead of completely subjective decisions to either make the experimental results fit a research paradigm or perhaps even worse, produce significant results. The distinction between decisions based on scientific justification and subjective nonsensical rationale is fuzzy and narrow, however keeping incentives, such as publishing pressure, fitting into a hypothesis or research paradigm, out of the equation can help with this distinction. This might even give rise to more rigorous methods and analysis pipelines because it hinders arbitrarily stopping the development of this pipeline when the results fit the preconceived notions of the scientific paradigm. Instead it forces researchers to stop when they are satisfied with the assumptions and implementations made. This process might also help researchers understand the uncertainty that is associated with many of the methods or practices commonly used in the literature, which are taken as either ground truths or good approximations when they are at best noisy estimates. This could for instance be the difference between taking an observed effect size from a previous study instead of relying or even incorporating some further scientifically justified assumptions that the researchers hold in their domain, that might give a much better approximation for the size of the underlying effect.

## *Why and how computational tools are becoming vital in science.*

Perhaps cognitive or even computational modeling is the fresh start that is needed in sciences, but especially sciences that have notoriously been relying on statistical models such as linear or generalized linear models. These more sophisticated models might be the steppingstone to engage in more theoretically driven modeling, however for this movement to succeed, it is essential that rigorous metrics are enforced from the beginning such that those models without any even provable, in principle, parameters or behaviors are discarded from the beginning. Example might arise where rigorous mathematical formulation of theories are developed but that in practice the formulation is not tractable from a computational perspective, it would be a shame to spent years investigating this model and its assumption in a field of research, just to discover that it in fact is intractable in practice ([Ho & Griffiths, 2021](#ref-ho_cognitive_2021); [McClelland, 2009](#ref-mcclelland_place_2009); [Zuidema et al., 2020](#ref-zuidema_five_2020)). One might think that a necessity of these more complicated models is a need for deeper mathematical understanding. What this thesis has argued is that this is not necessarily the case as simulations allow researchers to observe the implications of their assumptions as well as investigate when they break. A counter point to such argument would also be the increase in adaption of sophisticated hierarchical / multi-level models which are mathematically much more complex than single level models in cognitive science ([Dedrick et al., 2009](#ref-dedrick_multilevel_2009)). This is not to say that a better understanding of the machinery itself would not be helpful for researchers of various fields, but perhaps that instead of giving researchers and students a flowchart of when to use a particular statistical model, they should be getting the tools to understand and reflect on these statistical models and therefore also the tools to understand when they break. In the same way that a good scientific program does not teach students the right theories or hypotheses, it teaches them to think in a scientific way such that the individual can decide and test these themselves. Here the tools for understanding reflecting and experimenting with statistical models and concepts could be programming experience in statistics to conduct the types of data simulations presented in the current thesis. This would allow the user to understand the assumptions that are being made when they go wrong. This approach also requires to think more generatively about the process of how the data has been generated, because in order to simulate it is necessarily to have a model of how it was generated, which might even help spark new scientific ideas. This approach would have researchers more closely engaged in the statistical process of analyzing the data, instead of just picking an off the shelf model from a flowchart.

## Standing on the shoulders of giants

All of the models used in the current paper were fitted using Stan with the cmdstanr interface, which uses full Bayesian statistical inference with Markov chain monte carlo (MCMC) sampling ([Gabry et al., 2024](#ref-R-cmdstanr)). As described in the introduction section about modeling definitions, fitting and building models in this framework is extremely flexible as the sampling algorithm essentially serve as the optimizing process for the parameters of interest. This also means that essentially the code for simulating the generative process is close to identical in nature, to the code that specifies the model, making it easy for users with a generative framework to code up these types of models. The additional benefits to using especially Stan and the its Hamiltonian Monte Carlo (HMC) algorithm is that when issues arise the algorithm will complain and let you know, reducing the risk for erroneous inference due to the sampling algorithm ([Vehtari et al., 2021](#ref-vehtari_rank-normalization_2021)). The thesis used Bayesian inference and Stan, mainly due to its flexibility in model formulation and not because of the inherent differences between Bayesian and frequentist statistics, however Bayesian inference does allow for a more optimistic way to interpret the replication crisis discussion. This interpretation is that perhaps instead of starting each experimental analysis from the perspective that nothing is known, perhaps incorporating information from previous studies would be beneficial. This is in essences what science is about, a hierarchical organization of knowledge, where each step rests on the step below, i.e. on auxiliary assumptions as put by the Duhem–Quine thesis ([Ariew, 1984](#ref-ariew_duhem_1984)).

This view on science also matches that of uncertainties as these are also hierarchically organized and when doing analyses on data with uncertainties this uncertainty has to be accounted for. So in the same way that the results of a scientific theory is only as strong as its auxiliary assumptions; the strength of an analysis is also only as strong as the certainty with which the data is measured with. What the Bayesian framework of inference allows, is that prior information from similar studies can be used in the modelling allowing for researchers to not start their scientific studies from scratch, but pick up where others left off. This would essentially mean that instead of having to collect a larger number of subjects to achieve the actual desired power of the study, this could be done by two independent laboratories, the second using the information provided by the first. This essentially is already was is being done when conducting meta-analyses of different fields or sub fields. This approach incentives publications of all types of finding as they serve as the stepping stones for the next researcher, making the problem of publication bias where null findings are unpublished less incentives ([Laitin et al., 2021](#ref-laitin_reporting_2021)).

## Limitations

The current study investigated how the correlation coefficient is an inappropriate metric to internal model validation of cognitive models and purposed the modified intra class correlation (ICC) as a more sensible metric. As is clearly the case in Figure 11, the correlation coefficient quickly becomes asymptotic with quickly diminishing uncertainties whereas the granularity in the ICC is much better with quite high uncertainties. It might be argued that this increased granularity however is not important as the shape of the curve is similar, with increases in correlation and ICC with increasing trials with an asymptote at 1. This interpretation is sensible and comparing Figure 11(with simulated beta of 2 as the population mean in the power analysis) and the power analysis results in Figure 22, it seems like the pattern for the correlation coefficient better follows the shape of the curve of the power analysis, however with no straightforward way to quantitatively compare these its impossible to say. What is however known is that the invariance to linear transformation of the correlation coefficient makes it non sensical from a theoretical standpoint to access how well a model can recover simulated parameter values. Future work should investigate the link between how the internal validation metric behaves as a function of trials and how the power of detecting a difference in the particular parameter estimates change as a function of trials. A thorough investigation of this link would mean that the somewhat arbitrary choice of trials when designing an experiment would no longer be arbitrary, but informed by how estimation uncertainty in the parameters of interest changes based on the number of trials ([Miller, 2024](#ref-miller_how_2024)).

Another limitation of the current study is the quite limited power analysis done, future investigations should expand upon this power analysis for the other parameters and especially the slope of the psychometric function. This would not only help elucidate the question posed above about the relationship between the internal model validation metric, trials and power but could also be done while implementing the reaction time and or confidence informed models. The reasoning for only conducting the single power analysis on the threshold in the current thesis highlights one of the main hurdles of the framework purposed, computational resources. Firstly fitting models using HMC and Bayesian inference is both more time and computational resource intensive compared to frequentist inference in packages such as lme4, lmertest or GAMLSS to name a few quite flexible models fitting packages in R ([Bates et al., 2015](#ref-bates_fitting_2015); [Kuznetsova et al., 2017](#ref-kuznetsova_lmertest_2017); [Stasinopoulos & Rigby, 2008](#ref-stasinopoulos_generalized_2008)).

This added time for doing the optimization of posterior distributions of parameters has drawbacks in a need for access to bigger machines to necessitate the need for parallelization of the computational burden especially when several chains are needed to ensure convergence. Fortunately the access to bigger machine both privately but also on a institutation level is something that is growing in accessibility and already available to many universities or centers of research and has been correlated with research competitiveness ([Apon et al., 2010](#ref-apon_high_2010)). The current thesis was supported by Ucloud (see acknowledgement).

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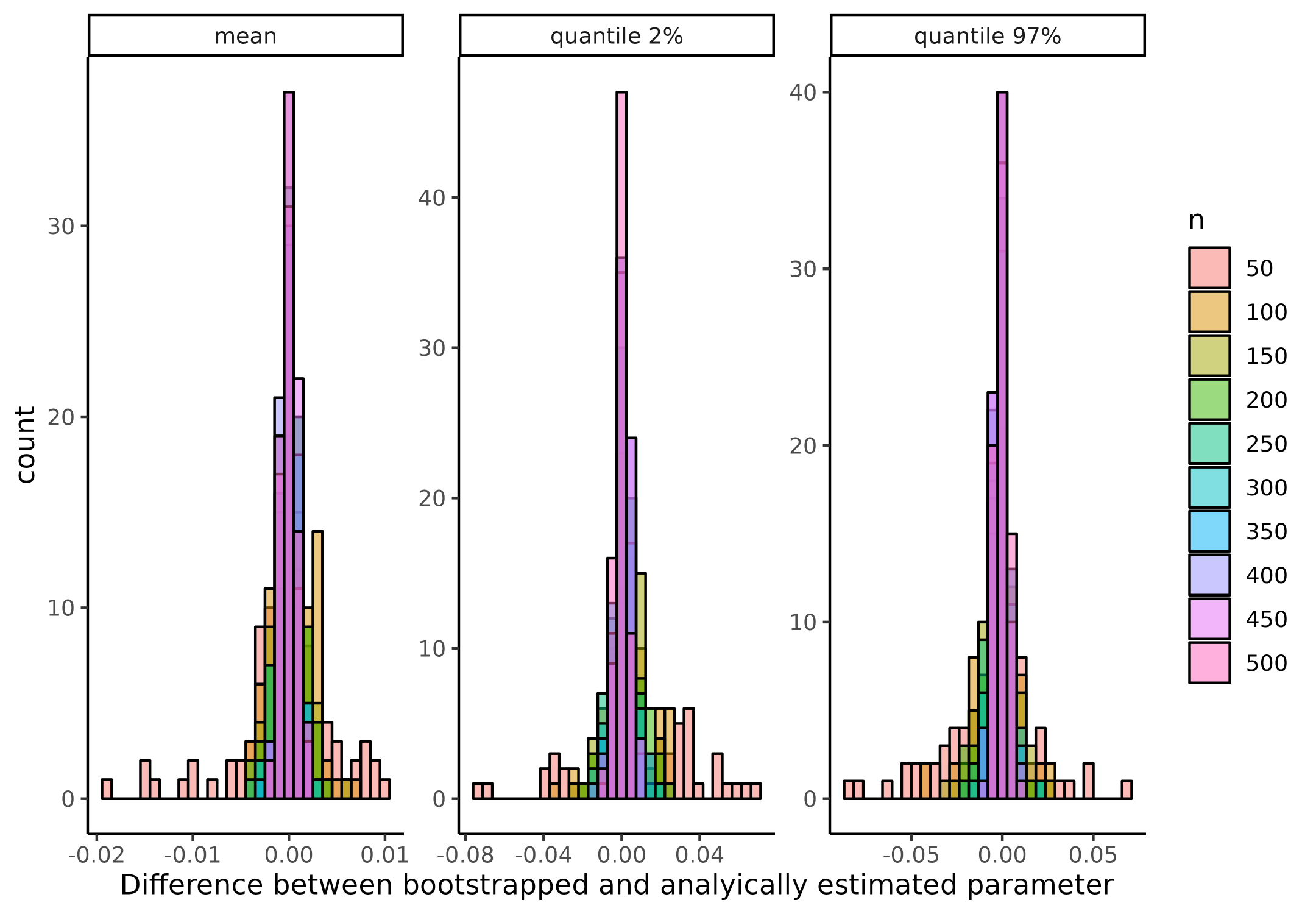
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# Supplementary material

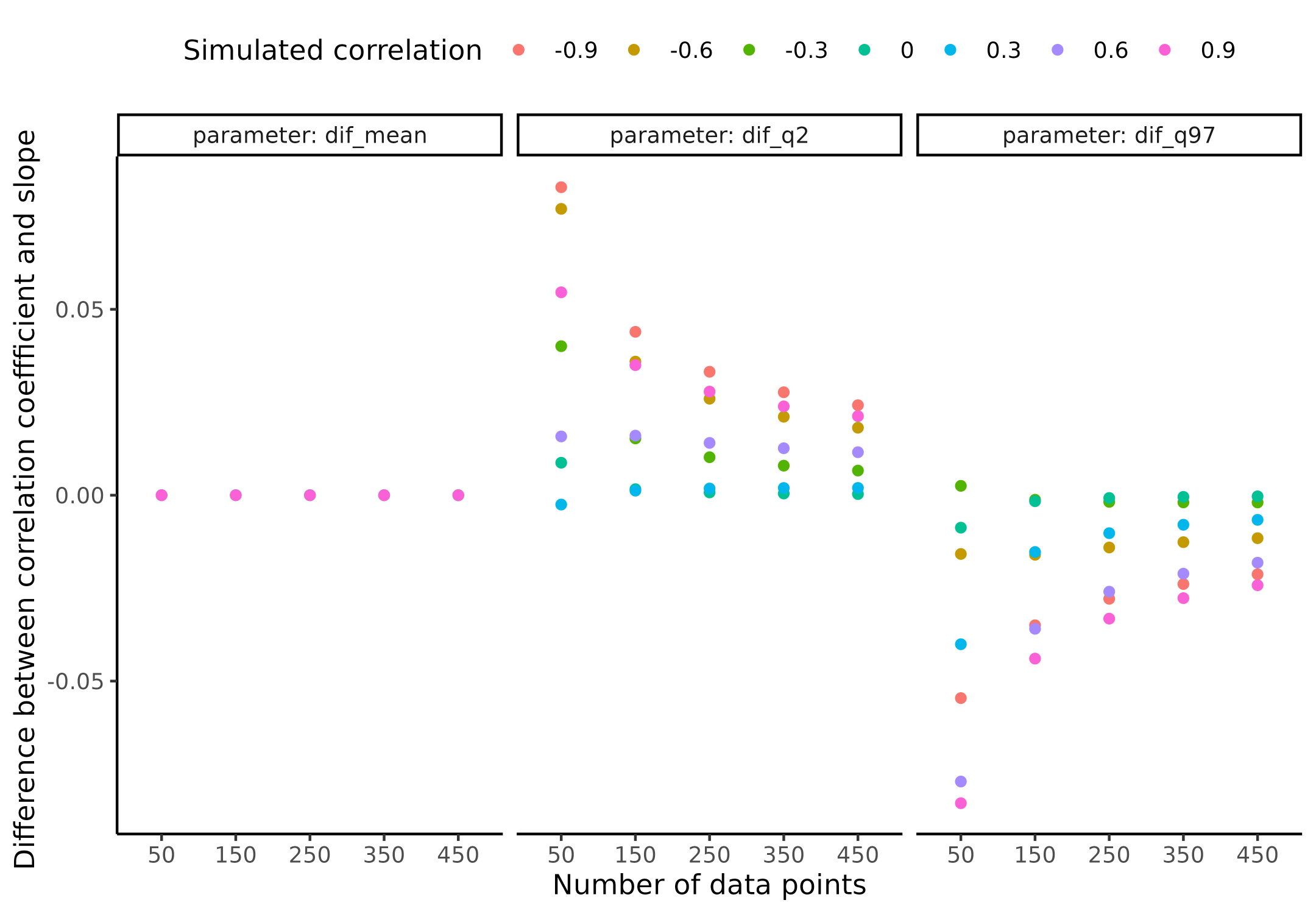
## **Supplementary Figures**

### Supplementary figure 1



**Supplemenary figure 1 Comparison of analytical and bootstrapping correlation coefficient.** Histograms of the difference between different metrics of the correlation coefficient when bootstrapping and analytically calculating the correlation coefficient. Facets shows the different used metrics when evaluating the correlation coefficient i.e. the mean, the 2% quantile and the 97% quantile. colors represent the sample size, i.e. the number of datapoints the simulated correlation coefficient was based on.

### Supplementary figure 2



**Supplemenary figure 2 Comparison of linear regression and correlation coefficient.** Scatter plot of the difference between the mean, quantile 2 and quantile 97 of the slope of a standardized linear regression and the correlation coefficient. Facets show the different statistical metric i.e. mean, q2 and q97 and the colors display the simulated correlation coefficient of the multinomial distribution. The x-axis is the number of datapoints simulated from the multinomial distribuion.

## Supplementary Notes:

### Priors

In the section for Standard parameter recovery a single fit psychometric function was used to demonstrate the standard parameter recovery approach to internal validity of the psychometric function. The priors for that model were as follows for the threshold,slope and lapse rate respectively.

Next the priors for the nested hierarchical model was as follows: