



Interactions and ANOVA

[Link to Notebook GitHub](#)

Note: This script is based heavily on Jonathan Taylor's class notes <http://www.stanford.edu/class/stats191/interactions.html>

Download and format data:

```
In [ ]: from __future__ import print_function
from statsmodels.compat import urlopen
import numpy as np
np.set_printoptions(precision=4, suppress=True)
import statsmodels.api as sm
import pandas as pd
pd.set_option("display.width", 100)
import matplotlib.pyplot as plt
from statsmodels.formula.api import ols
from statsmodels.graphics.api import interaction_plot, abline_plot
from statsmodels.stats.anova import anova_lm

try:
    salary_table = pd.read_csv('salary.table')
except: # recent pandas can read URL without urlopen
    url = 'http://stats191.stanford.edu/data/salary.table'
    fh = urlopen(url)
    salary_table = pd.read_table(fh)
    salary_table.to_csv('salary.table')

E = salary_table.E
M = salary_table.M
X = salary_table.X
S = salary_table.S
```

Take a look at the data:

```
In [ ]: plt.figure(figsize=(6,6))
symbols = ['D', '^']
colors = ['r', 'g', 'blue']
factor_groups = salary_table.groupby(['E', 'M'])
for values, group in factor_groups:
    i, j = values
    plt.scatter(group['X'], group['S'], marker=symbols[j], color=colors[i-1],
                s=144)
plt.xlabel('Experience');
plt.ylabel('Salary');
```

Fit a linear model:

```
In [ ]: formula = 'S ~ C(E) + C(M) + X'
lm = ols(formula, salary_table).fit()
print(lm.summary())
```

Have a look at the created design matrix:

```
In [ ]: lm.model.exog[:5]
```

```
=====
                        OLS Regression Results
=====
Dep. Variable:          S      R-squared:                0.957
Model:                  OLS      Adj. R-squared:         0.953
Method:                 Least Squares      F-statistic:       226.8
Date:                   Mon, 20 Jul 2015      Prob (F-statistic): 2.23e-27
Time:                   17:43:41      Log-Likelihood:    -381.63
No. Observations:       46      AIC:                  773.3
Df Residuals:           41      BIC:                  782.4
Df Model:                4
Covariance Type:        nonrobust
=====
```

	coef	std err	t	P> t	[95.0% Conf. Int.]
Intercept	8035.5976	386.689	20.781	0.000	7254.663 8816.532
C(E)[T.2]	3144.0352	361.968	8.686	0.000	2413.025 3875.045
C(E)[T.3]	2996.2103	411.753	7.277	0.000	2164.659 3827.762

C(M)[T.1]	6883.5310	313.919	21.928	0.000	6249.559	7517.503
X	546.1840	30.519	17.896	0.000	484.549	607.819
=====						
Omnibus:		2.293	Durbin-Watson:			2.237
Prob(Omnibus):		0.318	Jarque-Bera (JB):			1.362
Skew:		-0.077	Prob(JB):			0.506
Kurtosis:		2.171	Cond. No.			33.5
=====						

Warnings:
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Or since we initially passed in a DataFrame, we have a DataFrame available in

```
In [ ]: lm.model.data.orig_exog[:5]
```

We keep a reference to the original untouched data in

```
In [ ]: lm.model.data.frame[:5]
```

Influence statistics

```
In [ ]: infl = lm.get_influence()
        print(infl.summary_table())
```

or get a dataframe

```
In [ ]: df_infl = infl.summary_frame()
```

	obs	endog	fitted value	Cook's d	student. residual	hat diag	dffits internal	ext.stud. residual	dffits
	0	13876.000	15465.313	0.104	-1.683	0.155	-0.722	-1.723	-0.739
	1	11608.000	11577.992	0.000	0.031	0.130	0.012	0.031	0.012
	2	18701.000	18461.523	0.001	0.247	0.109	0.086	0.244	0.085
	3	11283.000	11725.817	0.005	-0.458	0.113	-0.163	-0.453	-0.162
	4	11767.000	11577.992	0.001	0.197	0.130	0.076	0.195	0.075
	5	20872.000	19155.532	0.092	1.787	0.126	0.678	1.838	0.698
	6	11772.000	12272.001	0.006	-0.513	0.101	-0.172	-0.509	-0.170
	7	10535.000	9127.966	0.056	1.457	0.116	0.529	1.478	0.537
	8	12195.000	12124.176	0.000	0.074	0.123	0.028	0.073	0.027
	9	12313.000	12818.185	0.005	-0.516	0.091	-0.163	-0.511	-0.161
	10	14975.000	16557.681	0.084	-1.655	0.134	-0.650	-1.692	-0.664
	11	21371.000	19701.716	0.078	1.728	0.116	0.624	1.772	0.640
	12	19800.000	19553.891	0.001	0.252	0.096	0.082	0.249	0.081
	13	11417.000	10220.334	0.033	1.227	0.098	0.405	1.234	0.408
	14	20263.000	20100.075	0.001	0.166	0.093	0.053	0.165	0.053
	15	13231.000	13216.544	0.000	0.015	0.114	0.005	0.015	0.005
	16	12884.000	13364.369	0.004	-0.488	0.082	-0.146	-0.483	-0.145
	17	13245.000	13910.553	0.007	-0.674	0.075	-0.192	-0.669	-0.191
	18	13677.000	13762.728	0.000	-0.089	0.113	-0.032	-0.087	-0.031
	19	15965.000	17650.049	0.082	-1.747	0.119	-0.642	-1.794	-0.659
	20	12336.000	11312.702	0.021	1.043	0.087	0.323	1.044	0.323
	21	21352.000	21192.443	0.001	0.163	0.091	0.052	0.161	0.051
	22	13839.000	14456.737	0.006	-0.624	0.070	-0.171	-0.619	-0.170
	23	22884.000	21340.268	0.052	1.579	0.095	0.511	1.610	0.521
	24	16978.000	18742.417	0.083	-1.822	0.111	-0.644	-1.877	-0.664
	25	14803.000	15549.105	0.008	-0.751	0.065	-0.199	-0.747	-0.198
	26	17404.000	19288.601	0.093	-1.944	0.110	-0.684	-2.016	-0.709
	27	22184.000	22284.811	0.000	-0.103	0.096	-0.034	-0.102	-0.033
	28	13548.000	12405.070	0.025	1.162	0.083	0.350	1.167	0.352
	29	14467.000	13497.438	0.018	0.987	0.086	0.304	0.987	0.304
	30	15942.000	16641.473	0.007	-0.705	0.068	-0.190	-0.701	-0.189
	31	23174.000	23377.179	0.001	-0.209	0.108	-0.073	-0.207	-0.072
	32	23780.000	23525.004	0.001	0.260	0.092	0.083	0.257	0.082
	33	25410.000	24071.188	0.040	1.370	0.096	0.446	1.386	0.451
	34	14861.000	14043.622	0.014	0.834	0.091	0.263	0.831	0.262
	35	16882.000	17733.841	0.012	-0.863	0.077	-0.249	-0.860	-0.249
	36	24170.000	24469.547	0.003	-0.312	0.127	-0.119	-0.309	-0.118
	37	15990.000	15135.990	0.018	0.878	0.104	0.300	0.876	0.299
	38	26330.000	25163.556	0.035	1.202	0.109	0.420	1.209	0.422
	39	17949.000	18826.209	0.017	-0.897	0.093	-0.288	-0.895	-0.287
	40	25685.000	26108.099	0.008	-0.452	0.169	-0.204	-0.447	-0.202
	41	27837.000	26802.108	0.039	1.087	0.141	0.440	1.089	0.441
	42	18838.000	19918.577	0.033	-1.119	0.117	-0.407	-1.123	-0.408
	43	17483.000	16774.542	0.018	0.743	0.138	0.297	0.739	0.295
	44	19207.000	20464.761	0.052	-1.313	0.131	-0.511	-1.325	-0.515
	45	19346.000	18959.278	0.009	0.423	0.208	0.216	0.419	0.214

```
In [ ]: df_infl[:5]
```

Now plot the reiduals within the groups separately:

```
In [ ]: resid = lm.resid
plt.figure(figsize=(6,6));
for values, group in factor_groups:
    i,j = values
    group_num = i*2 + j - 1 # for plotting purposes
    x = [group_num] * len(group)
    plt.scatter(x, resid[group.index], marker=symbols[j], color=colors[i-1],
               s=144, edgecolors='black')
plt.xlabel('Group');
plt.ylabel('Residuals');
```

Now we will test some interactions using anova or f_test

```
In [ ]: interX_lm = ols("S ~ C(E) * X + C(M)", salary_table).fit()
print(interX_lm.summary())
```

Do an ANOVA check

```
In [ ]: from statsmodels.stats.api import anova_lm

table1 = anova_lm(lm, interX_lm)
print(table1)

interM_lm = ols("S ~ X + C(E)*C(M)", data=salary_table).fit()
print(interM_lm.summary())

table2 = anova_lm(lm, interM_lm)
print(table2)
```

```
=====
                        OLS Regression Results
=====
Dep. Variable:          S      R-squared:          0.961
Model:                  OLS      Adj. R-squared:      0.955
Method:                 Least Squares      F-statistic:      158.6
Date:                   Mon, 20 Jul 2015      Prob (F-statistic):      8.23e-26
Time:                   17:43:41      Log-Likelihood:      -379.47
No. Observations:       46      AIC:              772.9
Df Residuals:           39      BIC:              785.7
Df Model:                6
Covariance Type:        nonrobust
=====
                        coef      std err          t      P>|t|      [95.0% Conf. Int.]
-----
Intercept      7256.2800      549.494      13.205      0.000      6144.824  8367.736
C(E)[T.2]      4172.5045      674.966       6.182      0.000      2807.256  5537.753
C(E)[T.3]      3946.3649      686.693       5.747      0.000      2557.396  5335.333
C(M)[T.1]      7102.4539      333.442      21.300      0.000      6428.005  7776.903
X               632.2878       53.185      11.888      0.000       524.710  739.865
C(E)[T.2]:X    -125.5147       69.863      -1.797      0.080      -266.826  15.796
C(E)[T.3]:X    -141.2741       89.281      -1.582      0.122      -321.861  39.313
=====
Omnibus:              0.432      Durbin-Watson:      2.179
Prob(Omnibus):        0.806      Jarque-Bera (JB):      0.590
Skew:                 0.144      Prob(JB):              0.744
Kurtosis:             2.526      Cond. No.              69.7
=====
```

Warnings:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

The design matrix as a DataFrame

```
In [ ]: interM_lm.model.data.orig_exog[:5]
```

	df_resid	ssr	df_diff	ss_diff	F	Pr(>F)
0	41	43280719.492876	0	NaN	NaN	NaN
1	39	39410679.807560	2	3870039.685316	1.914856	0.160964

```
=====
                        OLS Regression Results
=====
Dep. Variable:          S      R-squared:          0.999
Model:                  OLS      Adj. R-squared:      0.999
Method:                 Least Squares      F-statistic:      5517.
Date:                   Mon, 20 Jul 2015      Prob (F-statistic):      1.67e-55
Time:                   17:43:41      Log-Likelihood:      -298.74
No. Observations:       46      AIC:              611.5
Df Residuals:           39      BIC:              624.3
Df Model:                6
Covariance Type:        nonrobust
=====
                        coef      std err          t      P>|t|      [95.0% Conf. Int.]
-----
Intercept      9472.6854       80.344     117.902      0.000      9310.175  9635.196
C(E)[T.2]      1381.6706       77.319     17.870      0.000      1225.279  1538.063
C(E)[T.3]      1730.7483      105.334     16.431      0.000      1517.690  1943.806
C(M)[T.1]      3981.3769      101.175     39.351      0.000      3776.732  4186.022
C(E)[T.2]:C(M)[T.1] 4902.5231      131.359     37.322      0.000      4636.825  5168.222
```

```

C(E)[T.3]:C(M)[T.1]    3066.0351    149.330    20.532    0.000    2763.986    3368.084
X                      496.9870     5.566     89.283    0.000    485.728    508.246
=====
Omnibus:                74.761    Durbin-Watson:                2.244
Prob(Omnibus):          0.000    Jarque-Bera (JB):          1037.873
Skew:                   -4.103    Prob(JB):                4.25e-226
Kurtosis:               24.776    Cond. No.                79.0
=====

```

Warnings:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

```

      df_resid      ssr df_diff      ss_diff      F      Pr(>F)
0         41  43280719.492876      0         NaN         NaN         NaN
1         39  1178167.864864      2  42102551.628012  696.844466  3.025504e-31

```

The design matrix as an ndarray

```
In [ ]: interM_lm.model.exog
interM_lm.model.exog_names
```

```
In [ ]: infl = interM_lm.get_influence()
resid = infl.resid_studentized_internal
plt.figure(figsize=(6,6))
for values, group in factor_groups:
    i,j = values
    idx = group.index
    plt.scatter(X[idx], resid[idx], marker=symbols[j], color=colors[i-1],
                s=144, edgecolors='black')
plt.xlabel('X');
plt.ylabel('standardized resid');
```

Looks like one observation is an outlier.

```
In [ ]: drop_idx = abs(resid).argmax()
print(drop_idx) # zero-based index
idx = salary_table.index.drop(drop_idx)

lm32 = ols('S ~ C(E) + X + C(M)', data=salary_table, subset=idx).fit()

print(lm32.summary())
print('\n')

interX_lm32 = ols('S ~ C(E) * X + C(M)', data=salary_table, subset=idx).fit()

print(interX_lm32.summary())
print('\n')

table3 = anova_lm(lm32, interX_lm32)
print(table3)
print('\n')

interM_lm32 = ols('S ~ X + C(E) * C(M)', data=salary_table, subset=idx).fit()

table4 = anova_lm(lm32, interM_lm32)
print(table4)
print('\n')
```

Replot the residuals

```
In [ ]: try:
        resid = interM_lm32.get_influence().summary_frame()['standard_resid']
    except:
        resid = interM_lm32.get_influence().summary_frame()['standard_resid']

plt.figure(figsize=(6,6))
for values, group in factor_groups:
    i,j = values
    idx = group.index
    plt.scatter(X[idx], resid[idx], marker=symbols[j], color=colors[i-1],
                s=144, edgecolors='black')
plt.xlabel('X[~[32]]');
plt.ylabel('standardized resid');
```

32

OLS Regression Results

```

=====
Dep. Variable:          S    R-squared:                0.955
Model:                OLS    Adj. R-squared:            0.950
Method:             Least Squares    F-statistic:            211.7
Date:               Mon, 20 Jul 2015    Prob (F-statistic):      2.45e-26
Time:               17:43:42    Log-Likelihood:         -373.79

```

No. Observations:	45	AIC:	757.6			
Df Residuals:	40	BIC:	766.6			
Df Model:	4					
Covariance Type:	nonrobust					
=====						
	coef	std err	t	P> t	[95.0% Conf. Int.]	

Intercept	8044.7518	392.781	20.482	0.000	7250.911	8838.592
C(E)[T.2]	3129.5286	370.470	8.447	0.000	2380.780	3878.277
C(E)[T.3]	2999.4451	416.712	7.198	0.000	2157.238	3841.652
C(M)[T.1]	6866.9856	323.991	21.195	0.000	6212.175	7521.796
X	545.7855	30.912	17.656	0.000	483.311	608.260
=====						
Omnibus:	2.511	Durbin-Watson:		2.265		
Prob(Omnibus):	0.285	Jarque-Bera (JB):		1.400		
Skew:	-0.044	Prob(JB):		0.496		
Kurtosis:	2.140	Cond. No.		33.1		
=====						

Warnings:
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

OLS Regression Results						
=====						
Dep. Variable:	S	R-squared:	0.959			
Model:	OLS	Adj. R-squared:	0.952			
Method:	Least Squares	F-statistic:	147.7			
Date:	Mon, 20 Jul 2015	Prob (F-statistic):	8.97e-25			
Time:	17:43:42	Log-Likelihood:	-371.70			
No. Observations:	45	AIC:	757.4			
Df Residuals:	38	BIC:	770.0			
Df Model:	6					
Covariance Type:	nonrobust					
=====						
	coef	std err	t	P> t	[95.0% Conf. Int.]	

Intercept	7266.0887	558.872	13.001	0.000	6134.711	8397.466
C(E)[T.2]	4162.0846	685.728	6.070	0.000	2773.900	5550.269
C(E)[T.3]	3940.4359	696.067	5.661	0.000	2531.322	5349.549
C(M)[T.1]	7088.6387	345.587	20.512	0.000	6389.035	7788.243
X	631.6892	53.950	11.709	0.000	522.473	740.905
C(E)[T.2]:X	-125.5009	70.744	-1.774	0.084	-268.714	17.712
C(E)[T.3]:X	-139.8410	90.728	-1.541	0.132	-323.511	43.829
=====						
Omnibus:	0.617	Durbin-Watson:	2.194			
Prob(Omnibus):	0.734	Jarque-Bera (JB):	0.728			
Skew:	0.162	Prob(JB):	0.695			
Kurtosis:	2.468	Cond. No.	68.7			
=====						

Warnings:
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

	df_resid	ssr	df_diff	ss_diff	F	Pr(>F)
0	40	43209096.482552	0	NaN	NaN	NaN
1	38	39374237.269069	2	3834859.213483	1.850508	0.171042

	df_resid	ssr	df_diff	ss_diff	F	Pr(>F)
0	40	43209096.482552	0	NaN	NaN	NaN
1	38	171188.119937	2	43037908.362615	4776.734853	2.291239e-46

Plot the fitted values

```
In [ ]: lm_final = ols('S ~ X + C(E)*C(M)', data = salary_table.drop([drop_idx])).fit()
mf = lm_final.model.data.orig_exog
lstyle = ['-','--']

plt.figure(figsize=(6,6))
for values, group in factor_groups:
    i,j = values
    idx = group.index
    plt.scatter(X[idx], S[idx], marker=symbols[j], color=colors[i-1],
                s=144, edgecolors='black')
    # drop NA because there is no idx 32 in the final model
    plt.plot(mf.X[idx].dropna(), lm_final.fittedvalues[idx].dropna(),
            ls=lstyle[j], color=colors[i-1])
plt.xlabel('Experience');
plt.ylabel('Salary');
```

From our first look at the data, the difference between Master's and PhD in the management group is different than in the non-management group. This is an interaction between the two qualitative variables management,M and education,E. We can visualize this by first removing the effect of experience, then plotting the means within each of the 6 groups using interaction.plot.

```
In [ ]: U = S - X * interX_lm32.params['X']
```

```
plt.figure(figsize=(6,6))
interaction_plot(E, M, U, colors=['red','blue'], markers=['^','D'],
                markersize=10, ax=plt.gca())
```

Minority Employment Data

```
In [ ]: try:
        minority_table = pd.read_table('minority.table')
    except: # don't have data already
        url = 'http://stats191.stanford.edu/data/minority.table'
        minority_table = pd.read_table(url)
```

```
factor_group = minority_table.groupby(['ETHN'])

fig, ax = plt.subplots(figsize=(6,6))
colors = ['purple', 'green']
markers = ['o', 'v']
for factor, group in factor_group:
    ax.scatter(group['TEST'], group['JPERF'], color=colors[factor],
               marker=markers[factor], s=12**2)
ax.set_xlabel('TEST');
ax.set_ylabel('JPERF');
```

```
In [ ]: min_lm = ols('JPERF ~ TEST', data=minority_table).fit()
        print(min_lm.summary())
```

```
In [ ]: fig, ax = plt.subplots(figsize=(6,6));
        for factor, group in factor_group:
            ax.scatter(group['TEST'], group['JPERF'], color=colors[factor],
                       marker=markers[factor], s=12**2)

        ax.set_xlabel('TEST')
        ax.set_ylabel('JPERF')
        fig = abline_plot(model_results = min_lm, ax=ax)
```

```
In [ ]: min_lm2 = ols('JPERF ~ TEST + TEST:ETHN',
                      data=minority_table).fit()

        print(min_lm2.summary())
```

```
In [ ]: fig, ax = plt.subplots(figsize=(6,6));
        for factor, group in factor_group:
            ax.scatter(group['TEST'], group['JPERF'], color=colors[factor],
                       marker=markers[factor], s=12**2)

        fig = abline_plot(intercept = min_lm2.params['Intercept'],
                           slope = min_lm2.params['TEST'], ax=ax, color='purple');
        fig = abline_plot(intercept = min_lm2.params['Intercept'],
                           slope = min_lm2.params['TEST'] + min_lm2.params['TEST:ETHN'],
                           ax=ax, color='green');
```

```
In [ ]: min_lm3 = ols('JPERF ~ TEST + ETHN', data = minority_table).fit()
        print(min_lm3.summary())
```

```
In [ ]: fig, ax = plt.subplots(figsize=(6,6));
        for factor, group in factor_group:
            ax.scatter(group['TEST'], group['JPERF'], color=colors[factor],
                       marker=markers[factor], s=12**2)

        fig = abline_plot(intercept = min_lm3.params['Intercept'],
                           slope = min_lm3.params['TEST'], ax=ax, color='purple');
        fig = abline_plot(intercept = min_lm3.params['Intercept'] + min_lm3.params['ETHN'],
                           slope = min_lm3.params['TEST'], ax=ax, color='green');
```

```
In [ ]: min_lm4 = ols('JPERF ~ TEST * ETHN', data = minority_table).fit()
        print(min_lm4.summary())
```

```
In [ ]: fig, ax = plt.subplots(figsize=(8,6));
        for factor, group in factor_group:
            ax.scatter(group['TEST'], group['JPERF'], color=colors[factor],
                       marker=markers[factor], s=12**2)

        fig = abline_plot(intercept = min_lm4.params['Intercept'],
                           slope = min_lm4.params['TEST'], ax=ax, color='purple');
        fig = abline_plot(intercept = min_lm4.params['Intercept'] + min_lm4.params['ETHN'],
                           slope = min_lm4.params['TEST'] + min_lm4.params['TEST:ETHN'],
                           ax=ax, color='green');
```

```
In [ ]: # is there any effect of ETHN on slope or intercept?
        table5 = anova_lm(min_lm, min_lm4)
        print(table5)
```

```
In [ ]: # is there any effect of ETHN on intercept
```

```
table6 = anova_lm(min_lm, min_lm3)
print(table6)
```

```
In [ ]: # is there any effect of ETHN on slope
table7 = anova_lm(min_lm, min_lm2)
print(table7)
```

```
In [ ]: # is it just the slope or both?
table8 = anova_lm(min_lm2, min_lm4)
print(table8)
```

One-way ANOVA

```
In [ ]: try:
        rehab_table = pd.read_csv('rehab.table')
    except:
        url = 'http://stats191.stanford.edu/data/rehab.csv'
        rehab_table = pd.read_table(url, delimiter=",")
        rehab_table.to_csv('rehab.table')

fig, ax = plt.subplots(figsize=(8,6))
fig = rehab_table.boxplot('Time', 'Fitness', ax=ax, grid=False)
```

```
In [ ]: rehab_lm = ols('Time ~ C(Fitness)', data=rehab_table).fit()
table9 = anova_lm(rehab_lm)
print(table9)

print(rehab_lm.model.data.orig_exog)
```

	df	sum_sq	mean_sq	F	PR(>F)
C(Fitness)	2	672	336.000000	16.961538	0.000041
Residual	21	416	19.809524	NaN	NaN
Intercept					
0	1		0	0	
1	1		0	0	
2	1		0	0	
3	1		0	0	
4	1		0	0	
5	1		0	0	
6	1		0	0	
7	1		0	0	
8	1		1	0	
9	1		1	0	
10	1		1	0	
11	1		1	0	
12	1		1	0	
13	1		1	0	
14	1		1	0	
15	1		1	0	
16	1		1	0	
17	1		1	0	
18	1		0	1	
19	1		0	1	
20	1		0	1	
21	1		0	1	
22	1		0	1	
23	1		0	1	

```
In [ ]: print(rehab_lm.summary())
```

OLS Regression Results						
=====						
Dep. Variable:	Time	R-squared:	0.618			
Model:	OLS	Adj. R-squared:	0.581			
Method:	Least Squares	F-statistic:	16.96			
Date:	Mon, 20 Jul 2015	Prob (F-statistic):	4.13e-05			
Time:	17:43:46	Log-likelihood:	-68.286			
No. Observations:	24	AIC:	142.6			
Df Residuals:	21	BIC:	146.1			
Df Model:	2					
Covariance Type:	nonrobust					
=====						
	coef	std err	t	P> t	[95.0% Conf. Int.]	

Intercept	38.0000	1.574	24.149	0.000	34.728	41.272
C(Fitness)[T.2]	-6.0000	2.111	-2.842	0.010	-10.390	-1.610
C(Fitness)[T.3]	-14.0000	2.404	-5.824	0.000	-18.999	-9.001
=====						
Omnibus:	0.163	Durbin-Watson:	2.209			
Prob(Omnibus):	0.922	Jarque-Bera (JB):	0.211			
Skew:	-0.163	Prob(JB):	0.900			
Kurtosis:	2.675	Cond. No.	3.80			
=====						

Warnings:
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Two-way ANOVA

```
In [ ]: try:
        kidney_table = pd.read_table('./kidney.table')
    except:
        url = 'http://stats191.stanford.edu/data/kidney.table'
        kidney_table = pd.read_table(url, delimiter=" *")

/Users/tom.augspurger/Envs/py3/lib/python3.4/site-packages/pandas/io/parsers.py:648: ParserWarning: Falling back to
ParserWarning)
```

Explore the dataset

```
In [ ]: kidney_table.groupby(['Weight', 'Duration']).size()
```

Balanced panel

```
In [ ]: kt = kidney_table
plt.figure(figsize=(8,6))
fig = interaction_plot(kt['Weight'], kt['Duration'], np.log(kt['Days']+1),
                      colors=['red', 'blue'], markers=['D', '^'], ms=10, ax=plt.gca())
```

You have things available in the calling namespace available in the formula evaluation namespace

```
In [ ]: kidney_lm = ols('np.log(Days+1) ~ C(Duration) * C(Weight)', data=kt).fit()

table10 = anova_lm(kidney_lm)

print(anova_lm(ols('np.log(Days+1) ~ C(Duration) + C(Weight)',
                  data=kt).fit(), kidney_lm))
print(anova_lm(ols('np.log(Days+1) ~ C(Duration)', data=kt).fit(),
              ols('np.log(Days+1) ~ C(Duration) + C(Weight, Sum)',
                  data=kt).fit()))
print(anova_lm(ols('np.log(Days+1) ~ C(Weight)', data=kt).fit(),
              ols('np.log(Days+1) ~ C(Duration) + C(Weight, Sum)',
                  data=kt).fit()))
```

	df_resid	ssr	df_diff	ss_diff	F	Pr(>F)
0	56	29.624856	0	NaN	NaN	NaN
1	54	28.989198	2	0.635658	0.59204	0.556748
	df_resid	ssr	df_diff	ss_diff	F	Pr(>F)
0	58	46.596147	0	NaN	NaN	NaN
1	56	29.624856	2	16.971291	16.040454	0.000003
	df_resid	ssr	df_diff	ss_diff	F	Pr(>F)
0	57	31.964549	0	NaN	NaN	NaN
1	56	29.624856	1	2.339693	4.422732	0.03997

Sum of squares

Illustrates the use of different types of sums of squares (I,II,III) and how the Sum contrast can be used to produce the same output between the 3.

Types I and II are equivalent under a balanced design.

Don't use Type III with non-orthogonal contrast - ie., Treatment

```
In [ ]: sum_lm = ols('np.log(Days+1) ~ C(Duration, Sum) * C(Weight, Sum)',
                    data=kt).fit()

print(anova_lm(sum_lm))
print(anova_lm(sum_lm, typ=2))
print(anova_lm(sum_lm, typ=3))
```

	df	sum_sq	mean_sq	F	PR(>F)
C(Duration, Sum)	1	2.339693	2.339693	4.358293	0.041562
C(Weight, Sum)	2	16.971291	8.485645	15.806745	0.000004
C(Duration, Sum):C(Weight, Sum)	2	0.635658	0.317829	0.592040	0.556748
Residual	54	28.989198	0.536837	NaN	NaN
	sum_sq	df	F	PR(>F)	
C(Duration, Sum)	2.339693	1	4.358293	0.041562	
C(Weight, Sum)	16.971291	2	15.806745	0.000004	
C(Duration, Sum):C(Weight, Sum)	0.635658	2	0.592040	0.556748	
Residual	28.989198	54	NaN	NaN	
	sum_sq	df	F	PR(>F)	
Intercept	156.301830	1	291.153237	2.077589e-23	
C(Duration, Sum)	2.339693	1	4.358293	4.156170e-02	
C(Weight, Sum)	16.971291	2	15.806745	3.944502e-06	
C(Duration, Sum):C(Weight, Sum)	0.635658	2	0.592040	5.567479e-01	
Residual	28.989198	54	NaN	NaN	

```
In [ ]: nosum_lm = ols('np.log(Days+1) ~ C(Duration, Treatment) * C(Weight, Treatment)',
                      data=kt).fit()
```



```
print(anova_lm(nosum_lm))
print(anova_lm(nosum_lm, typ=2))
print(anova_lm(nosum_lm, typ=3))
```

	df	sum_sq	mean_sq	F	PR(>F)
C(Duration, Treatment)	1	2.339693	2.339693	4.358293	0.041562
C(Weight, Treatment)	2	16.971291	8.485645	15.806745	0.000004
C(Duration, Treatment):C(Weight, Treatment)	2	0.635658	0.317829	0.592040	0.556748
Residual	54	28.989198	0.536837	NaN	NaN
		sum_sq	df	F	PR(>F)
C(Duration, Treatment)		2.339693	1	4.358293	0.041562
C(Weight, Treatment)		16.971291	2	15.806745	0.000004
C(Duration, Treatment):C(Weight, Treatment)		0.635658	2	0.592040	0.556748
Residual		28.989198	54	NaN	NaN
		sum_sq	df	F	PR(>F)
Intercept		10.427596	1	19.424139	0.000050
C(Duration, Treatment)		0.054293	1	0.101134	0.751699
C(Weight, Treatment)		11.703387	2	10.900317	0.000106
C(Duration, Treatment):C(Weight, Treatment)		0.635658	2	0.592040	0.556748
Residual		28.989198	54	NaN	NaN