
Chapter 3: Joint Distributions

1 JOINT DISTRIBUTIONS FOR MULTIPLE RANDOM VARIABLES

- Very often, we are interested in more than one random variables simultaneously.
- For example, an investigator might be interested in both the height (H) and the weight (W) of an individual from a certain population.
- Another investigator could be interested in both the hardness (H) and the tensile strength (T) of a piece of cold-drawn copper.

DEFINITION 1

- *Let E be an experiment and S be a corresponding sample space.*
- *Let X and Y be two functions each assigning a real number to each $s \in S$.*
- *We call (X, Y) a two-dimensional random vector, or a two-dimensional random variable.*

Similarly to one-dimensional situation, we can denote the **range space** of (X, Y) by

$$R_{X,Y} = \left\{ (x, y) \middle| x = X(s), y = Y(s), s \in S \right\}.$$

The definition above can be extended to more than two random variables.

DEFINITION 2

Let X_1, X_2, \dots, X_n be n functions each assigning a real number to every outcome $s \in S$. We call (X_1, X_2, \dots, X_n) an n -dimensional random variable (or an n -dimensional random vector).

We define the discrete and continuous two-dimensional RVs as follows.

DEFINITION 3

1 (X, Y) is a **discrete** two-dimensional RV if the number of possible values of $(X(s), Y(s))$ are finite or countable.

That is the possible values of $(X(s), Y(s))$ may be represented by

$$(x_i, y_j), i = 1, 2, 3, \dots; j = 1, 2, 3, \dots$$

2 (X, Y) is a **continuous** two-dimensional RV if the possible values of $(X(s), Y(s))$ can assume any value in some region of the Euclidean space \mathbb{R}^2 .

REMARK

we can view X and Y separately to judge whether (X, Y) is discrete or continuous.

- If both X and Y are discrete RVs, then (X, Y) is a discrete RV.
- Likewise, if both X and Y are continuous random variables, then (X, Y) is a continuous RV.
- Clearly, there are other cases. For example, X is discrete, but Y is continuous. These are not our focus in this module.

Example 3.1 (Discrete Random Vector)

- Consider a TV set to be serviced.
- Let

$X = \{\text{age to the nearest year of the set}\};$

$Y = \{\text{\# of defective components in the set}\}.$

- (X, Y) is a discrete 2-dimensional RV.
- $R_{X,Y} = \{(x, y) | x = 0, 1, 2, \dots; y = 0, 1, 2, \dots, n\}$, where n is the total number of components in the TV.
- $(X, Y) = (5, 3)$ means that the TV is 5 years old and has 3 defective components.

Joint Probability Function

- We introduce the probability function for the discrete and continuous RVs separately.
- For discrete random vector, similar to the one-dimensional case, we define its probability function by associate a number with each possible value of the RV.

DEFINITION 4 (JOINT PROBABILITY FUNCTION FOR DISCRETE RV)

*Let (X, Y) be a 2-dimensional **discrete** RV, the **joint probability (mass) function** is defined by*

$$f_{X,Y}(x,y) = P(X = x, Y = y),$$

for x, y being possible values of X and Y , or in the other words $(x, y) \in R_{X,Y}$.

The joint probability mass function has the following properties:

(1) $f_{X,Y}(x,y) \geq 0$ for any $(x,y) \in R_{X,Y}$.

(2) $f_{X,Y}(x,y) = 0$ for any $(x,y) \notin R_{X,Y}$.

(3) $\sum_{i=1}^{\infty} \sum_{j=1}^{\infty} f_{X,Y}(x_i, y_j) = \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} P(X = x_i, Y = y_j) = 1;$

or equivalently $\sum \sum_{(x,y) \in R_{X,Y}} f(x,y) = 1.$

(4) Let A be any subset of $R_{X,Y}$, then

$$P((X,Y) \in A) = \sum \sum_{(x,y) \in A} f_{X,Y}(x,y).$$

Example 3.2 Find the value of k such that $f(x,y) = kxy$ for $x = 1, 2, 3$ and $y = 1, 2, 3$ can serve as a joint probability function.

Solution: $R_{X,Y} = \{(x,y) | x = 1, 2, 3; y = 1, 2, 3\}$.

$$\begin{aligned} f(1,1) &= k, & f(1,2) &= 2k, & f(1,3) &= 3k, \\ f(2,1) &= 2k, & f(2,2) &= 4k, & f(2,3) &= 6k, \\ f(3,1) &= 3k, & f(3,2) &= 6k, & f(3,3) &= 9k. \end{aligned}$$

Based on property (3), we have

$$\begin{aligned} 1 &= \sum \sum_{(x,y) \in R_{X,Y}} f(x,y) \\ &= 1k + 2k + 3k + 2k + 4k + 6k + 3k + 6k + 9k, \end{aligned}$$

which results in $k = 1/36$.

DEFINITION 5 (JOINT PROBABILITY FUNCTION FOR CONTINUOUS RV)

Let (X, Y) be a 2-dimensional *continuous* RV; its *joint probability (density) function* is a function $f_{X,Y}(x, y)$ such that

$$P((X, Y) \in D) = \int \int_{(x,y) \in D} f_{X,Y}(x, y) dy dx,$$

for any $D \subset \mathbb{R}^2$. More specifically,

$$P(a \leq X \leq b, c \leq Y \leq d) = \int_a^b \int_c^d f_{X,Y}(x, y) dy dx.$$

The joint probability density function has the following properties:

(1) $f_{X,Y}(x,y) \geq 0$, for any $(x,y) \in R_{X,Y}$.

(2) $f_{X,Y}(x,y) = 0$, for any $(x,y) \notin R_{X,Y}$.

(3) $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx dy = 1;$

or equivalently $\int \int_{(x,y) \in R_{X,Y}} f_{X,Y}(x,y) dx dy = 1.$

Example 3.3 Find the value c such that $f(x,y)$ below can serve as a joint p.d.f. for a RV (X,Y) :

$$f(x,y) = \begin{cases} cx(x+y), & 0 \leq x \leq 1; 1 \leq y \leq 2 \\ 0, & \text{elsewhere} \end{cases}$$

Solution: In order for $f(x,y)$ to be a p.d.f., we need

$$\begin{aligned} 1 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x,y) dy dx = \int_0^1 \int_1^2 cx(x+y) dy dx = c \int_0^1 x \left(x + \frac{1}{2}y^2 \Big|_1^2 \right) dx \\ &= c \int_0^1 x(x+1.5) dx = c \left(\frac{1}{3}x^3 + 1.5 \cdot \frac{1}{2}x^2 \right) \Big|_0^1 = c \cdot \frac{13}{12}, \end{aligned}$$

which implies $c = 12/13$.

2 MARGINAL AND CONDITIONAL DISTRIBUTIONS

DEFINITION 6 (MARGINAL PROBABILITY DISTRIBUTION)

Let (X, Y) be a two-dimensional RV with joint p.f. $f_{X,Y}(x, y)$. We define the marginal distribution for X as follows.

- If Y is a discrete RV, then for any x ,*

$$f_X(x) = \sum_y f_{X,Y}(x, y).$$

- If Y is a continuous RV, then for any x ,*

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy.$$

REMARK

- $f_Y(y)$ for Y is defined in the same way as that of X .
- We can view the marginal distribution as the “projection” of the 2D function $f_{X,Y}(x,y)$ to the 1D function.
- More intuitively, it is the distribution of X by ignoring the presence of Y .

For example, consider a person of a certain community,

- suppose $X = \text{body weight}$, $Y = \text{height}$. (X,Y) has a joint distribution $f_{X,Y}(x,y)$.
- the marginal distribution $f_X(x)$ of X is the **distribution of body weights for all people in the community**.

- $f_X(x)$ should not involve the variable y ; this can be viewed from its definition: y is either summed out or integrated over.
- $f_X(x)$ is a **probability function** so it satisfies all the properties of the probability function.

Example 3.4

- Revisit Example 3.2. The joint p.f. is given by $f(x,y) = \frac{1}{36}xy$ for $x = 1, 2, 3$ and $y = 1, 2, 3$.
- Note that X has three possible values: 1, 2, and 3. The marginal distribution for X is given by
 - for $x = 1$, $f_X(1) = f(1,1) + f(1,2) + f(1,3) = 6/36 = 1/6$.
 - for $x = 2$, $f_X(2) = f(2,1) + f(2,2) + f(2,3) = 12/36 = 1/3$.
 - for $x = 3$, $f_X(3) = f(3,1) + f(3,2) + f(3,3) = 18/36 = 1/2$.
 - for other values of x , $f_X(x) = 0$.

- Alternatively, for each $x \in \{1, 2, 3\}$,

$$\begin{aligned} f_X(x) &= \sum_y f(x, y) = \sum_{y=1}^3 \frac{1}{36} xy \\ &= \frac{1}{36} x \sum_{y=1}^3 y = \frac{1}{6} x. \end{aligned}$$

DEFINITION 7 (CONDITIONAL DISTRIBUTION)

*Let (X, Y) be a RV with joint p.f. $f_{X,Y}(x, y)$. Let $f_X(x)$ be the marginal p.f. for X . Then for any x such that $f_X(x) > 0$, the **conditional probability function of Y given $X = x$** is defined to be*

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x, y)}{f_X(x)}.$$

REMARK

- For any y such that $f_Y(y) > 0$, we can similarly define the **conditional distribution of X given $Y = y$** :

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}.$$

- $f_{Y|X}(y|x)$ is defined only for x such that $f_X(x) > 0$; likewise $f_{X|Y}(x|y)$ is defined only for y such that $f_Y(y) > 0$.
- The practical meaning of $f_{Y|X}(y|x)$: the distribution of Y given that the random variable X is observed to take the value x .

- Considering y as the variable (x as a fixed value), $f_{Y|X}(y|x)$ is a p.f., so it must satisfy all the properties of p.f..
- But $f_{Y|X}(y|x)$ is not a p.f. for x ; this means that there is **NO** requirement $\int_{-\infty}^{\infty} f_{Y|X}(y|x)dx = 1$ for X continuous or $\sum_x f_{Y|X}(y|x) = 1$ for X discrete.
- With the definition, we immediately have
 - If $f_X(x) > 0$, $f_{X,Y}(x,y) = f_X(x)f_{Y|X}(y|x)$.
 - If $f_Y(y) > 0$, $f_{X,Y}(x,y) = f_Y(y)f_{X|Y}(x|y)$.

- One immediate application of the conditional distribution is to compute, for continuous RV,

$$P(Y \leq y|X = x) = \int_{-\infty}^y f_{Y|X}(y|x)dy;$$
$$E(Y|X = x) = \int_{-\infty}^{\infty} yf_{Y|X}(y|x)dy.$$

Their practical meanings are clear: the former is the probability that $Y \leq y$, given $X = x$; the latter is the average value of Y given $X = x$.

For discrete case, the computation is similarly established based on $f_{Y|X}(y|x)$; please fill in the details on your own.

Example 3.5 Revisit Examples 3.2 and 3.4.

- The joint p.f. for (X, Y) is given by $f(x, y) = \frac{1}{36}xy$ for $x = 1, 2, 3$ and $y = 1, 2, 3$.
- The marginal p.f. for X is $f_X(x) = \frac{1}{6}x$ for $x = 1, 2, 3$.
- Therefore, $f_{Y|X}(y|x)$ is defined for any $x = 1, 2$, or 3 :

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} = \frac{(1/36)xy}{(1/6)x} = \frac{1}{6}y,$$

for $y = 1, 2, 3$.

- We can compute

$$P(Y = 2|X = 1) = f_{Y|X}(2|1) = \frac{1}{6} \cdot 2 = 1/3;$$

$$\begin{aligned} P(Y \leq 2|X = 1) &= P(Y = 1|X = 1) + P(Y = 2|X = 1) \\ &= f_{Y|X}(1|1) + f_{Y|X}(2|1) = 1/6 + 1/3 = 1/2; \end{aligned}$$

$$\begin{aligned} E(Y|X = 2) &= 1 \cdot f_{Y|X}(1|2) + 2 \cdot f_{Y|X}(2|2) + 3 \cdot f_{Y|X}(3|2) \\ &= 1 \cdot (1/6) + 2 \cdot (2/6) + 3 \cdot (3/6) = 7/3. \end{aligned}$$

3 INDEPENDENT RANDOM VARIABLES

DEFINITION 8 (INDEPENDENT RANDOM VARIABLES)

- Random variables X and Y are *independent* if and only if for *any* x and y ,

$$f_{X,Y}(x,y) = f_X(x)f_Y(y).$$

- Random variables X_1, X_2, \dots, X_n are *independent* if and only if for *any* x_1, x_2, \dots, x_n ,

$$f_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n) = f_{X_1}(x_1)f_{X_2}(x_2) \cdots f_{X_n}(x_n).$$

REMARK

- The above definition is applicable no matter whether (X,Y) is continuous or discrete.
- The “product feature” in the definition implies one necessary condition for independence: $R_{X,Y}$ needs to be a product space. In the sense that if X and Y are independent, for any $x \in R_X$ and any $y \in R_Y$, we have

$$f_{X,Y}(x,y) = f_X(x)f_Y(y) > 0,$$

implying $R_{X,Y} = \{(x,y)|x \in R_X; y \in R_Y\} = R_X \times R_Y$.

Conclusion: if $R_{X,Y}$ is not a product space, then X and Y are not independent!

Properties of Independent Random Variables

Suppose X, Y are independent RVs.

- (1) If A and B are arbitrary subsets of \mathbb{R} , the events $X \in A$ and $Y \in B$ are independent events in S . Thus

$$P(X \in A; Y \in B) = P(X \in A)P(Y \in B).$$

In particular, for any real numbers x, y ,

$$P(X \leq x; Y \leq y) = P(X \leq x)P(Y \leq y).$$

(2) For arbitrary functions $g_1(\cdot)$ and $g_2(\cdot)$, $g_1(X)$ and $g_2(Y)$ are independent. For example,

- X^2 and Y are independent.
- $\sin(X)$ and $\cos(Y)$ are independent.
- e^X and $\log(Y)$ are independent.

(3) Independence is connected with conditional distribution.

- If $f_X(x) > 0$, then $f_{Y|X}(y|x) = f_Y(y)$.
- Likewise, if $f_Y(y) > 0$, then $f_{X|Y}(x|y) = f_X(x)$.

Example 3.6 The joint p.f. of (X, Y) is given below.

x	y			$f_X(x)$
	1	3	5	
2	0.1	0.2	0.1	0.4
4	0.15	0.3	0.15	0.6
$f_Y(y)$	0.25	0.5	0.25	1

Are X and Y independent?

Solution:

- We need to check that for every x and y combination, whether we have

$$f_{X,Y}(x,y) = f_X(x)f_Y(y).$$

For example, from the table, we have $f_{X,Y}(2,1) = 0.1$; $f_X(2) = 0.4$, $f_Y(1) = 0.25$. Therefore

$$f_{X,Y}(2,1) = 0.1 = 0.4 \times 0.25 = f_X(2)f_Y(1).$$

- In fact, we can check for each $x \in \{2,4\}$ and $y \in \{1,3,5\}$ combination, the equality holds.
- We conclude that X and Y are independent.

4 EXPECTATION AND COVARIANCE

DEFINITION 9 (EXPECTATION)

For any two variable function $g(x, y)$,

- if (X, Y) is a discrete RV,

$$E(g(X, Y)) = \sum_x \sum_y g(x, y) f_{X, Y}(x, y);$$

- if (X, Y) is a continuous RV,

$$E(g(X, Y)) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{X, Y}(x, y) dy dx.$$

If we let

$$g(X, Y) = (X - E(X))(Y - E(Y)) = (X - \mu_X)(Y - \mu_Y),$$

the expectation $E[g(X, Y)]$ leads to the covariance of X and Y .

DEFINITION 10 (COVARIANCE)

The covariance of X and Y is defined to be

$$\text{cov}(X, Y) = E[(X - E(X))(Y - E(Y))]$$

- If X and Y are discrete RVs,

$$\text{cov}(X, Y) = \sum_x \sum_y (x - \mu_X)(y - \mu_Y) f_{X,Y}(x, y).$$

- If X and Y are continuous RVs,

$$\text{cov}(X, Y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x - \mu_X)(y - \mu_Y) f_{X,Y}(x, y) dx dy.$$

The covariance has the following properties.

(1) $\text{cov}(X, Y) = E(XY) - E(X)E(Y).$

(2) If X and Y are independent, then $\text{cov}(X, Y) = 0$. However, $\text{cov}(X, Y) = 0$ does not imply that X and Y are independent.

$$(3) \operatorname{cov}(aX + b, cY + d) = ac \cdot \operatorname{cov}(X, Y).$$

$$(4) V(aX + bY) = a^2V(X) + b^2V(Y) + 2ab \cdot \operatorname{cov}(X, Y).$$

Example 3.7 Given the joint distribution for (X, Y) :

x	y				$f_X(x)$
	0	1	2	3	
0	1/8	1/4	1/8	0	1/2
1	0	1/8	1/4	1/8	1/2
$f_Y(y)$	1/8	3/8	3/8	1/8	1

(a) Find $E(Y - X)$.

(b) Find $\text{cov}(X, Y)$.

Solution:

(a) Method 1:

$$\begin{aligned} E(Y - X) &= (0 - 0)(1/8) + (1 - 0)(1/4) + (2 - 0)(1/8) \\ &\quad + \dots + (3 - 1)(1/8) = 1. \end{aligned}$$

Method 2:

$$E(Y - X) = E(Y) - E(X) = 1.5 - 0.5 = 1,$$

where

$$E(Y) = 0 \cdot (1/8) + 1 \cdot (3/8) + 2 \cdot (3/8) + 3 \cdot (1/8) = 1.5$$

$$E(X) = 0 \cdot (1/2) + 1 \cdot (1/2) = 0.5.$$

(b) We use $\text{cov}(X, Y) = E(XY) - E(X)E(Y)$ to compute. Note that we have computed $E(X)$ and $E(Y)$ in Part (a).

$$\begin{aligned} E(XY) &= (0)(0)(1/8) + (0)(1)(1/4) + (0)(2)(1/8) \\ &\quad + \dots + (1)(3)(1/8) = 1. \end{aligned}$$

Therefore

$$\text{cov}(X, Y) = E(XY) - E(X)E(Y) = 1 - (0.5)(1.5) = 0.25.$$