

② Notar que:  $f(x|\theta) = \frac{\Gamma(\theta+1)}{\Gamma(\theta)\Gamma(1)} x^{\theta-1} (1-x)^{\theta-1} = \theta x^{\theta-1}$

④ Ahora,  $\left( \frac{-1}{3\ln(x)}, \frac{-2}{3\ln(x)} \right) \subset C$ , pero la Prob. de cubrimiento está dada por

$P\left( \frac{-1}{3\ln(x)} < \theta < \frac{-2}{3\ln(x)} \right)$  para  $0 < x < 1$ ,  $\theta > 0$ , pero como  $\ln(x) < 0$ ,  $\ln(x)$  es un # negativo.

$\therefore P\left( \frac{-1}{3\theta} > \ln(x) > \frac{-2}{3\theta} \right) \Rightarrow$  Recordemos que, si  $a, b$  son #s en  $\mathbb{R}$ ...

para  $a > b$   
 $a > b \Leftrightarrow e^a > e^b \Leftrightarrow e^{-a} < e^{-b}$  por decrecimiento exponencial.

$\therefore P\left( e^{-\frac{1}{3\theta}} < x < e^{-\frac{2}{3\theta}} \right) \Leftrightarrow P\left( e^{\frac{-2}{3\theta}} < x < e^{\frac{-1}{3\theta}} \right)$ , calculando

$\downarrow$   $\downarrow$   $\downarrow$   
 $a$   $b$   $\rightarrow$  Prob. de cubrimiento

$\int_a^b \theta x^{\theta-1} dx = x^\theta \Big|_a^b = e^{-\frac{1}{3}} - e^{-\frac{2}{3}} \approx 0,2031 \dots$

$\inf_{\theta} \left( \theta \in \left( \frac{-1}{3\ln(x)}, \frac{-2}{3\ln(x)} \right) \right) = \inf_{\theta} \left( e^{-\frac{1}{3}} - e^{-\frac{2}{3}} \right) \approx 0,2031$  Será el coeficiente de confianza

③ ①  $H_0: \beta = \beta_0$  vs  $H_1: \beta \neq \beta_0$

$\therefore L(\beta|x, \lambda) = \prod_{i=1}^n \frac{1}{\Gamma(\lambda)\beta^\lambda} x_i^{\lambda-1} e^{-\frac{x_i}{\beta}} = \left( \frac{1}{\Gamma(\lambda)\beta^\lambda} \right)^n \left( \prod_{i=1}^n x_i^{\lambda-1} \right) \cdot e^{-\frac{1}{\beta} \sum x_i}$

$l(\beta) = -n \ln(\Gamma(\lambda)) - n\lambda \ln(\beta) + (\lambda-1) \sum \ln(x_i) - \frac{1}{\beta} \sum x_i$

$l'(\beta) = \frac{-n\lambda}{\beta} + \frac{1}{\beta^2} \sum x_i - \frac{1}{\beta} \left[ \frac{\sum x_i}{\beta} - n\lambda \right] - \frac{1}{\beta} \left[ \frac{\sum x_i - n\lambda\beta}{\beta} \right]$

$\therefore \hat{\beta} = \frac{\sum x_i}{n\lambda} = \frac{\bar{x}}{\lambda}$ , luego  $S(\beta) = \frac{\sum x_i - n\lambda\beta}{\beta^2} \Rightarrow$  (II)

$\ln(\beta) = -E \left[ \frac{\partial}{\partial \beta} \left( \frac{-n\lambda}{\beta} + \frac{1}{\beta^2} \sum x_i \right) \right]$



$$-E\left[\frac{n\lambda}{\beta^2} - \frac{2}{\beta^3} \sum x_i\right] = \frac{-n\lambda}{\beta^2} + \frac{2}{\beta^3} \sum E(x_i) \rightarrow x_i \sim \text{Gamma}$$

$$= \frac{-n\lambda}{\beta^2} + \frac{2n\lambda\beta}{\beta^3} = \frac{-n\lambda + 2n\lambda}{\beta^2} = \frac{n\lambda}{\beta^2} \quad (II)$$

$\therefore E(x_i) = \lambda\beta$ , Así entonces

$$Z_0 = \frac{S(\beta_0)}{\sqrt{I_n(\beta_0)}} = \frac{\sum x_i - n\lambda\beta_0}{\sqrt{\frac{n\lambda}{\beta_0^2}}}$$

Bajo  $H_0$

$$\uparrow = \frac{n\bar{x} - n\lambda\beta_0}{\beta_0^2} = \frac{\sqrt{n} \sqrt{\lambda}}{\beta_0}$$

y Tomando

$$n\bar{x} = \sum x_i$$

$$= \frac{n(\bar{x} - \beta_0\lambda)}{\sqrt{n} \sqrt{\lambda} \beta_0} = \frac{\sqrt{n}(\bar{x} - \lambda\beta_0)}{\sqrt{\lambda} \beta_0} \xrightarrow{d} N(0,1)$$

Así, la región de rechazo bajo  $H_0$  será  $|Z_0| > z_{\frac{\alpha}{2}}$

y la región de aceptación para hallar un IC approx al  $(100)(1-\alpha)\%$  para  $\beta$  será

$$1-\alpha \cong P(Z_0 | |Z_0| < z_{\alpha/2})$$

$$1-\alpha \cong P\left(\left|\frac{\sqrt{n}(\bar{x} - \lambda\beta_0)}{\sqrt{\lambda} \beta_0}\right| \leq z_{\frac{\alpha}{2}}\right)$$

(II)

Tomando (II)  $\therefore -z_{\frac{\alpha}{2}} \leq \frac{\sqrt{n}(\bar{x} - \lambda\beta_0)}{\sqrt{\lambda} \beta_0} \leq z_{\frac{\alpha}{2}}$

$$\frac{-z_{\frac{\alpha}{2}} \sqrt{\lambda}}{\sqrt{n}} \leq \frac{\bar{x}}{\beta_0} - \frac{\lambda\beta_0}{\beta_0} \leq \frac{z_{\frac{\alpha}{2}} \sqrt{\lambda}}{\sqrt{n}}$$

$$\lambda - \frac{z_{\frac{\alpha}{2}} \sqrt{\lambda}}{\sqrt{n}} \leq \frac{\bar{x}}{\beta_0} \leq \lambda + \frac{z_{\alpha/2} \sqrt{\lambda}}{\sqrt{n}}$$

$$\frac{\sqrt{n}}{\sqrt{n}\lambda + z_{\alpha/2} \sqrt{\lambda}} \leq \frac{\beta_0}{\bar{x}} \leq \frac{\sqrt{n}}{\sqrt{n}\lambda - z_{\alpha/2} \sqrt{\lambda}}$$



11b

11a

$$\frac{\sqrt{n}(\bar{X} - \mu_0)}{\sqrt{\frac{\sigma^2}{n}}} \sim N\left(0, 1\right) \quad \text{Por tanto queda}$$

$$\sqrt{n}(\sqrt{n}\bar{X} - \sqrt{n}\mu_0)$$

$$\text{demostrando que}$$

$$\ln IC \approx \ln \left( \frac{1}{2} \right)^n \text{ para } p = 0.5 \quad \left( \frac{1}{2} \right)^n, \left( \frac{1}{2} \right)^n$$

$$\text{b) Tenemos } \bar{X} = 1.07075, \quad \frac{\sum x_i}{n} = 1.96, \quad n = 80, \text{ reemplazando}$$

$$\text{tendremos } \left( \frac{\sqrt{80}(1.07075)}{\sqrt{1.96}} \right)^2, \quad \left( \frac{\sqrt{80}(1.07075)}{\sqrt{1.96(2) - 1.96}} \right)^2 \quad \left( \frac{1}{2} \right)^n, \left( \frac{1}{2} \right)^n$$

$$= (0.4666)^2; 0.6333$$

1) calculamos la verosimilitud.

$$L(p|x) = \prod_{i=1}^n (p^{x_i} (1-p)^{1-x_i}) = p^{\sum x_i} (1-p)^{n - \sum x_i}, \text{ luego, aplicando } \ln$$

$$\ln L(p|x) = \sum_{i=1}^n x_i \ln(p) + (n - \sum x_i) \ln(1-p) =$$

$$= \sum_{i=1}^n x_i \ln(p) + 20 \ln(1-p) - \sum x_i \ln(1-p) = \ln L(p|x)$$

$$\ln L(p|x) = \frac{\sum x_i}{p} + \frac{(n - \sum x_i)}{(1-p)} = 0 = \frac{\sum x_i - p n}{p(1-p)} = 0$$

$$= \sum_{i=1}^n x_i = p n \Rightarrow \hat{p} = \frac{\sum x_i}{n}, \text{ Verificando con 2da derivada}$$

$$\frac{d^2 \ln L(p|x)}{dp^2} = -\frac{1}{p^2} \sum x_i + \frac{(n - \sum x_i)}{(1-p)^2}$$

$$= \frac{\sum_{i=1}^n x_i}{p^2} + \frac{(n - \sum x_i)}{(1-p)^2} < 0, \text{ por tanto } \hat{p} \text{ es el MLE.}$$

$$n = 20$$

∴ LRT sea

$$\lambda(x) = \frac{L(p_0|x)}{L(\bar{x}|x)}$$

$$\text{con } p_0 = 1/2 \Rightarrow \lambda(x) = \frac{1}{2} \frac{\sum x_i (1/2)^{n - \sum x_i}}{(1/2)^n \sum x_i} = (1/2)^n$$

$$\lambda(x) = \frac{1}{2} \frac{\sum x_i (1/2)^{n - \sum x_i}}{(1/2)^n \sum x_i} \text{ para } x, \text{ se usa } -2 \ln(\lambda(x))$$



$$-2 \ln(\lambda(p)) = -2 \ln \left( \frac{(1-\bar{x})^{\sum x_i}}{\bar{x}^{\sum x_i} (1-\bar{x})^n} \right)$$

$$= -2 \ln(1-\bar{x})^{\sum x_i} - 2 \left[ \frac{\sum x_i \ln(\bar{x}) + n \ln(1-\bar{x})}{2} \right]$$

$$= -2 \ln(1-\bar{x})^{\sum x_i} - \sum x_i \ln(\bar{x}) - n \ln(1-\bar{x})$$

Se cumple un  $n$  suficientemente grande entonces para una confianza al 95%, buscamos en la tabla de la  $\chi^2$  cuadrado  $\chi^2_{0.05}(n) = 3,841$ .

Finalmente se rechaza  $H_0 = \frac{1}{2}$  cuando  $-2 \ln(\lambda(p)) > 3,841$

Nota: Como en el ejercicio,  $n = 20$ .

Suponiendo que se cumplen las condiciones de regularidad, entonces,

$$H_0 = \frac{1}{2}, n = 20$$