# Spring 2018: STA 6448 Advanced Probability and Inference II Lecture 17

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Random matrices and covariance estimation

# Cumulant function of sum of independent matrices

The cumulant function of sum of independent matrices does not decompose additively, because **matrix products need not commute**.

Fortunately, for independent random matrices, it is possible to establish an upper bound in terms of the trace of the cumulant generating functions.

#### Lemma

Let  $Q_1, \ldots, Q_n$  be independent symmetric random matrices whose cumulant functions exists for all  $\lambda \in I$ . Then the sum  $S_n = \sum_{i=1}^n Q_i$  satisfies

$$\operatorname{Tr}\left(e^{\Phi_{S_n}(\lambda)}\right) \leq \operatorname{Tr}\left(e^{\sum_{i=1}^n \Phi_{\mathcal{Q}_i}(\lambda)}\right) \quad \textit{for all } \lambda \in I.$$

A proof uses Lieb's theorem: for any fixed  $H \in \mathcal{S}^{d \times d}$ , the following function is concave:

$$A \mapsto \operatorname{Tr}\left(e^{H+\log(A)}\right)$$
.

#### Tail bounds for sub-Gaussian matrices

#### Theorem (Hoeffding bound for random matrices)

Let  $Q_1, \ldots, Q_n$  be independent symmetric random matrices that are sub-Gaussian with parameters  $V_1, \ldots, V_n$ . Then for any  $\delta > 0$ , we have

$$\mathbb{P}\Big[\|\sum_{i=1}^n Q_i\|_{op} \geq \delta\Big] \leq 2 d e^{-\frac{n\delta^2}{2\sigma^2}},$$

where 
$$\sigma^2 = |||n^{-1} \sum_{i=1}^n V_i|||_{op}$$
.

This inequality also implies an analogous bound for general independent but potentially non-symmetric and/or non-square matrices in  $\mathbb{R}^{d_1 \times d_2}$ , with d replaced by  $d_1 + d_2$  (why?).

# Example: Looseness/Sharpness of leading factor d

- Let n = d, and  $E_i$  denote the diagonal matrix with 1 in position (i, i) and 0 elsewhere.
- Let  $D_i = g_i E_i$  where  $g_i$  are i.i.d. sub-Gaussian with parameter 1.
- ▶ We showed  $D_i$  is sub-Gaussian with  $V_i = E_i$ , and hence  $\sigma^2 = \|d^{-1}I_d\|_{\infty} = 1/d$ . Therefore,

$$\mathbb{P}\Big[\|\!\| \frac{1}{d} \sum_{i=1}^d Q_i \|\!\|_{\mathsf{op}} \geq \delta\Big] \leq 2 \, d \, e^{-\frac{d\delta^2}{2\sigma^2}},$$

implying  $\|\frac{1}{d}\sum_{i=1}^d Q_i\|_{op} \leq \frac{\sqrt{2\log(2d)}}{d}$  with high probability.

▶ On the other hand, if  $g_i$  are Rademacher variables, then  $\|\frac{1}{d}\sum_{i=1}^d Q_i\|_{\text{op}} = \frac{1}{d}$  and the concentration inequality is off by the order d; if  $g_i$  are standard Gaussians, then  $\|\frac{1}{d}\sum_{i=1}^d Q_i\|_{\text{op}} \approx \frac{\sqrt{2\log(2d)}}{d}$  and the inequality cannot be

improved.

# Bernstein-type bounds for random matrices

#### Theorem (Matrix Bernstein concentration inequality)

Let  $Q_1, \ldots, Q_n$  be a sequence of independent, zero-mean, symmetric random matrices that satisfy the Bernstein condition with parameter b > 0. Then

$$\mathbb{P}\Big[\|\sum_{i=1}^n Q_i\|_{op} \geq \delta\Big] \leq 2 d \exp\Big\{-\frac{n\delta^2}{2(\sigma^2 + b\delta)}\Big\},\,$$

where 
$$\sigma^2 = |||n^{-1} \sum_{i=1}^n \text{Var}(Q_i)|||_{op}$$
.

This inequality can also be generalized to non-symmetric matrices  $A_i \in \mathbb{R}^{d_1 \times d_2}$ , as long as we use

$$\sigma^{2} = \max \Big\{ \| \frac{1}{n} \sum_{i=1}^{n} \mathbb{E}[A_{i} A_{i}^{T}] \|_{\text{op}}, \| \frac{1}{n} \sum_{i=1}^{n} \mathbb{E}[A_{i}^{T} A_{i}] \|_{\text{op}} \Big\},$$

and replace d by  $d_1 + d_2$ .

# Example: Tail bounds in matrix completion

- ► Consider an i.i.d. sequence of matrices of the form  $A_i = \xi_i X_i \in \mathbb{R}^{d \times d}$ .
- $\xi_i$  is symmetric around zero, satisfying Bernstein condition with parameter b and variance  $\nu^2$ .
- ►  $X_i$  is independent from  $\xi_i$ , with a single entry d in a position chosen uniformly at random from all  $d^2$  entries.
- Define a symmetric version

$$Q_i = \begin{bmatrix} 0_{d \times d} & A_i \\ A_i^T & 0_{d \times d} \end{bmatrix}$$

- ▶  $\| \sum_{i=1}^n A_i \|_{op} = \| \sum_{i=1}^n Q_i \|_{op}$ ,  $Q_i$  satisfies the Bernstein condition with parameter bd, and  $\sigma^2 = \nu^2 d$ .
- Then we have

$$\mathbb{P}\Big[\|\sum_{i=1}^{n} A_i\|_{\mathsf{op}} \geq \delta\Big] \leq 4 d \exp\Big\{-\frac{n\delta^2}{2d(\nu^2 + b\delta)}\Big\}.$$

# Example: Tail bounds in matrix completion

Now we try to reduce the symmetric assumption on the distribution of  $\xi_i$ . We achieve this via the symmetrization technique:

$$\mathbb{E}\Big[\exp\Big\{\lambda\gamma_{\max}(\sum_{i=1}^{n}Q_{i})\Big\}\Big] = \mathbb{E}\Big[\exp\Big\{\lambda\sup_{\|u\|_{2}=1}u^{T}\Big(\sum_{i=1}^{n}Q_{i}\Big)u\Big\}\Big]$$

$$\leq \mathbb{E}\Big[\exp\Big\{2\lambda\sup_{\|u\|_{2}=1}u^{T}\Big(\sum_{i=1}^{n}\varepsilon_{i}Q_{i}\Big)u\Big\}\Big]$$

$$= \mathbb{E}\Big[\exp\Big\{2\lambda\gamma_{\max}(\sum_{i=1}^{n}\varepsilon_{i}Q_{i})\Big\}\Big],$$

where  $\varepsilon_i$  are i.i.d. Rademacher variables, and the second step follows by the symmetrization theorem with  $\Phi(t) = e^{\lambda t}$ .

Therefore, we may consider the symmetrized version  $\varepsilon_i Q_i$  with the loss of a constant factor.

# Applications to covariance matrices

#### Corollary (Sample Covariance concentration)

Let  $X_i$  be i.i.d. zero-mean random vectors with covariance  $\Sigma$ , such that  $||x_i||_2 \leq \sqrt{b}$  almost surely. Then for all  $\delta > 0$ ,

$$\mathbb{P}\big[\|\widehat{\Sigma} - \Sigma\|_{op} \ge \delta\big] \le 2d \exp\Big(-\frac{n\delta^2}{2b(\|\Sigma\|_{op} + \delta)}\Big).$$

Proof: Apply matrix Bernstein concentration inequality to  $Q_i = x_i x_i^T - \Sigma$ .

$$|||Q_i|||_{\text{op}} \le ||x_i||_2^2 + |||\Sigma|||_{\text{op}} \le 2b.$$

Moreover,

$$\operatorname{Var}(Q_i) \leq \mathbb{E}[(x_i x_i^T)^2] \leq b\Sigma.$$

# Example: Random vectors uniform on sphere

 $x_i$  are chosen uniformly from the sphere  $S^{d-1}(\sqrt{d})$ , so that  $||x_i||_2 = \sqrt{d}$ .

By construction,  $\mathbb{E}[x_i x_i^T] = \Sigma = I_d$ , and  $||\Sigma||_{op} = 1$ . Therefore,

$$\mathbb{P}\big[|\!|\!|\!|\widehat{\Sigma} - \Sigma|\!|\!|_{\mathsf{op}} \geq \delta\big] \leq 2d \exp\Big(-\frac{n\delta^2}{2d(1+\delta)}\Big),$$

which implies the high probability bound

$$\|\widehat{\Sigma} - \Sigma\|_{\text{op}} \lesssim \sqrt{\frac{d \log d}{n}} + \frac{d \log d}{n}.$$

This bound is off by a factor of  $\log d$ , since we can directly apply the matrix sub-Gaussian concentration inequality ( $x_i$  is sub-Gaussian with parameter c for some universal constant c > 0).

#### Example: "Spiked" random vectors

 $x_i$  is uniformly chosen from  $\{\sqrt{d}e_1,\ldots,\sqrt{d}e_d\}$ , where  $e_j\in\mathbb{R}^d$  is the canonical basis vector with 1 in position j.

As before, we have  $||x_i||_2 = \sqrt{d}$ , and  $\mathbb{E}[x_i x_i^T] = I_d$ . Therefore, the same bound applies:

$$\|\widehat{\Sigma} - \Sigma\|_{\text{op}} \lesssim \sqrt{\frac{d \log d}{n}} + \frac{d \log d}{n}.$$

This time, this bound is sharp (up to constant factors).

# Structured covariance estimation: sparsity and thresholding

- ▶ Suppose  $\Sigma$  is known to be sparse, but the positions of non-zero entires are unknown.
- Motivates estimators based thresholding.
- ▶ Given a tuning parameter  $\lambda > 0$ , define the *hard* thresholding operator  $T_{\lambda} : \mathbb{R} \to \mathbb{R}$  by

$$T_{\lambda}(u) = u \mathbb{I}[|u| > \lambda].$$

- ▶ For a matrix M, we define  $T_{\lambda}(M)$  by applying  $T_{\lambda}$  to each element.
- We will study the property of the estimator  $T_{\lambda_n}(\widehat{\Sigma})$ , where  $\lambda_n > 0$  is a suitably chosen parameter.

# Sparsity and thresholding

- Let  $A \in \mathbb{R}^{d \times d}$  denote the adjacency matrix, where  $A_{ij} = \mathbb{I}(\Sigma_{ij} \neq 0)$ .
- ▶  $|||A|||_{op}$  provides a measure of sparsity: if  $\Sigma$  has at most s non-zero entries per row, then  $|||A|||_{op} \leq s$ .

#### **Theorem**

 $x_i$  are independent zero-mean sub-Gaussian with parameter at most  $\sigma^2$ . If  $n \geq \log d$ , then for any  $\delta > 0$  and  $\lambda_n/\sigma^2 = 8\sqrt{\frac{\log d}{n}} + \delta$ ,

$$\mathbb{P} \Big[ |\!|\!|\!| T_{\lambda_n}(\widehat{\Sigma}) - \Sigma |\!|\!|\!|_{\mathrm{op}} \geq 2 \, |\!|\!|\!| A |\!|\!|\!|\!|\!|_{\mathrm{op}} \lambda_n \Big] \leq 8 e^{-\frac{n}{16} \min\{\delta, \delta^2\}}.$$

#### Corollary

Suppose  $\Sigma$  has at most s non-zero entries per row, then

$$\mathbb{P}\Big[ |\!|\!| T_{\lambda_n}(\widehat{\Sigma}) - \Sigma |\!|\!|_{\mathrm{op}}/\sigma^2 \geq 16s\sqrt{\frac{\log d}{n}} + 2\delta \Big] \leq 8e^{-\frac{n}{16}\min\{\delta,\delta^2\}}.$$

#### Example: Sparsity and adjacency matrices

- In certain cases, the two bounds discussed before coincide.
- ▶ Consider any graph with maximum degree s − 1 that contains a s-clique
- For any such graph, we have

$$||A||_{op} = s - 1.$$

- In general, the bound with  $||A||_{op}$  can be substantially sharper.
- ▶ Consider a hub-and-spoke graph, in which one central node known as the hub is connected to s of the remaining d-1 node.
- ► For this graph, we have

$$||A||_{\mathsf{op}} = \sqrt{s}.$$