

STA 4103/5107 Computational Methods in Statistics II

Department of Statistics
Florida State University

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Next Five Classes

• Thursday, 02/23:

Point Process Filter

• Tuesday, 02/28:

Review for the Midterm

• Thursday, 03/02:

No Class

• Tuesday, 03/07:

Midterm Presentation, Group I

• Thursday, 03/09:

Midterm Presentation, Group II



7.4 Sequential Monte Carlo Methods



Nonliniear Filtering

- In many problems it is not possible to assume linearity of system evolution and observation maps.
- Such problems require tools to handle nonlinearity of the underlying systems.
- An important consequence of nonlinear models is that the posterior density at any time may not be Gaussian, and hence its representation using a mean and a covariance is no longer sufficient.
- In such cases one resorts to Monte Carlo idea and generates a large number of samples from the posterior at every time *t*.



Recursive Estimation

- Those samples are then used to estimate system state, the estimation error, or any other parameter of interest.
- Since the posterior changes at every observation time, one would have to generate samples at every time.
- However, using the relationship between evolving posteriors, one seeks an efficient algorithm for generating samples from posterior at time t + 1, using samples from the posterior at time t, in a recursive fashion.
- In this section, we describe the method of **sequential Monte** Carlo to accomplish this task.



Goal

- Let $f(x_t|y_1, ..., y_t)$ be the posterior density function at time t, and let x_t^i be a sample from this posterior. Define the sample set $S_t = \{x_t^1, x_t^2, ..., x_t^n\}$.
- Our goal is to use elements of S_t , and the new observation y_{t+1} and the system-observation models, to generate the sample set S_{t+1} .
- This task is performed in two steps (analogous to Kalman Filtering).
- The first step generates prediction and the second step updates, or corrects, these predictions using the data y_{t+1} .



Prediction Step

• 1. Prediction Step: This step uses samples from the posterior density at time t to generate samples from the prediction density $f(x_{t+1}|y_1, \ldots, y_t)$. That is,

$$f(x_{t+1} | y_1, ..., y_t) = \int_{x_t} f(x_{t+1} | x_t) f(x_t | y_1, ..., y_t) dx_t$$

• Since x_t^i is a sample from $f(x_t|y_1, \ldots, y_t)$, one can use a sample from the conditional density

$$f(x_{t+1} \mid x_t^i)$$

• Let $\widetilde{x}_{t+1}^i \sim f(x_{t+1} \mid x_t^i)$. Then

$$\widetilde{x}_{t+1}^i \sim f(x_{t+1} | y_1, y_2, ..., y_t)$$



Prediction Set

We will call the set

$$\widetilde{S}_{t+1} = {\{\widetilde{x}_{t+1}^1, \widetilde{x}_{t+1}^2, ..., \widetilde{x}_{t+1}^n\}},$$

as the **prediction set**.

- This one-step-prior comes from a state equation involving an update function $F(\cdot)$ and a random term.
- In this situation, sampling from this prior is straightforward, according to: $\widetilde{x}_{t+1}^i = F(x_t^i) + w_t^i,$

where w_t^i is a random sample from its given density function. In examples where w_t is multivariate normal this simulation is rather simple.



Update Step

- 2. Update Step: This step uses the prediction set to generate samples from the posterior density $f(x_{t+1}|y_1, \ldots, y_{t+1})$.
- We are interested in estimating a parameter

$$\theta_{t+1} = \int g(x_{t+1}) f(x_{t+1} | y_1, ..., y_{t+1}) dx_{t+1},$$

for a given function g. In case of minimum mean square estimation, the function is g is identity and we are interested in estimating the posterior mean as an estimate of the state x_{t+1} .

- The function *g* is always known beforehand.
- These two goals are accomplished using **importance sampling** and **resampling** as described next.



Importance Sampling

• Using ideas from importance sampling we can rewrite the definition of θ_{t+1} as follows:

$$\theta_{t+1} = \int g(x_{t+1}) f(x_{t+1} | y_1, ..., y_{t+1}) dx_{t+1}$$

$$= \int g(x_{t+1}) \frac{f(y_{t+1} | x_{t+1}) f(x_{t+1} | y_1, ..., y_t)}{f(y_{t+1} | y_1, ..., y_t)} dx_{t+1}$$

$$= \int \frac{g(x_{t+1}) f(y_{t+1} | x_{t+1})}{f(y_{t+1} | y_1, ..., y_t)} f(x_{t+1} | y_1, ..., y_t) dx_{t+1}$$

- If we have samples from $f(x_{t+1}|y_1, \ldots, y_t)$, we can use them to estimate θ_{t+1} .
- However, notice that the quantity $f(y_{t+1}|y_1, \ldots, y_t)$ is still unknown.



Sum of Weights

• We have to estimate this quantity. This is done by considering:

$$f(y_{t+1} \mid y_1, y_2, ..., y_t) = \int f(y_{t+1} \mid x_{t+1}) f(x_{t+1} \mid y_1, y_2, ..., y_t) dx_{t+1}$$

- Once again, since we have samples from $f(x_{t+1}|y_1, \ldots, y_t)$, we can use them, and the likelihood function $f(y_{t+1}|x_{t+1})$ to estimate the left side.
- Define the weights

$$w_{t+1}^{i} = f(y_{t+1} \mid \widetilde{x}_{t+1}^{i}),$$

and estimate $f(y_{t+1}|y_1, \dots, y_t)$ using

$$\frac{1}{n}\sum_{i=1}^n w_{t+1}^i.$$



Estimation

Using Monte-Carlo Method,

$$\hat{\theta}_{t+1} = \frac{1}{n} \sum_{i=1}^{n} \frac{g(\widetilde{x}_{t+1}^{i}) w_{t+1}^{i}}{\frac{1}{n} \sum_{j=1}^{n} w_{t+1}^{j}}.$$

Define the normalized weights

$$\widetilde{W}_{t+1}^{i} = \frac{W_{t+1}^{i}}{\sum_{j=1}^{n} W_{t+1}^{j}},$$

we can restate the estimator of θ_{t+1} as

$$\hat{\theta}_{t+1} = \sum_{i=1}^{n} g(\widetilde{x}_{t+1}^{i}) \widetilde{w}_{t+1}^{i}.$$



Resampling

- Next, we return to the task of generating samples from the posterior $f(x_{t+1}|y_1, \ldots, y_{t+1})$.
- It can be shown that if we resample from $\widetilde{S}_{t+1} = {\{\widetilde{x}_{t+1}^1, \widetilde{x}_{t+1}^2, ..., \widetilde{x}_{t+1}^n\}}$, with probabilities given by

$$\{\widetilde{w}_{t+1}^{1}, \widetilde{w}_{t+1}^{2}, ..., \widetilde{w}_{t+1}^{n}\},\$$

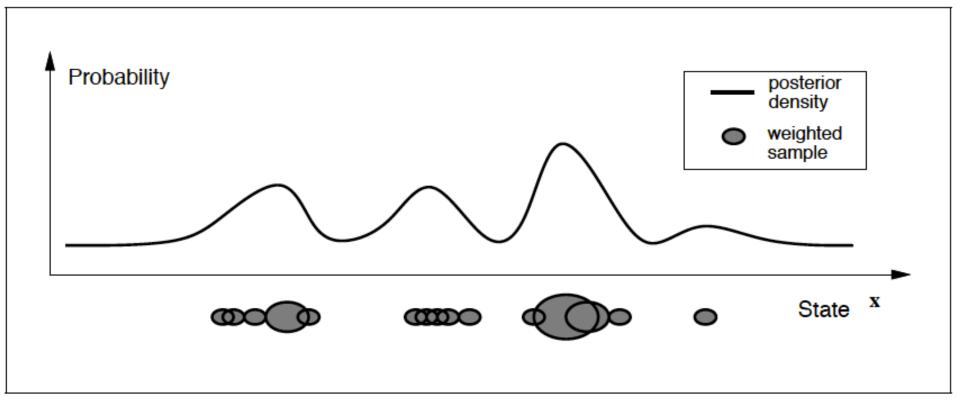
then, the resulting values are approximately samples from the posterior (when n is large).

• Generating n such resamples, and calling them x_{t+1}^i , we obtain the sample set

$$S_{t+1} = \{x_{t+1}^1, x_{t+1}^2, \dots, x_{t+1}^n\}.$$



Factor Sampling



CONDENSATION: conditional density propagation for visual tracking

International Journal of Computer Vision, 1998



Sequential Monte Carlo Algorithm

- Algorithm 40 (Classical SMC)
 - 1. Generate *n* samples $x_0^i \sim f(x_0)$. Set t = 0.
 - 2. **Prediction:** Generate the prediction set using:

$$\widetilde{x}_{t+1}^i \sim f(x_{t+1} \mid x_t^i), i = 1, 2, ..., n.$$

- 3. **Update:** Compute the weights $w_{t+1}^i = f(y_{t+1} \mid \widetilde{x}_{t+1}^i)$, and normalize them using $\widetilde{w}_{t+1}^i = w_{t+1}^i / \sum_{j=1}^n w_{t+1}^j$.
- (a) Estimate θ_{t+1} using $\hat{\theta}_{t+1} = \sum_{i=1}^{n} g(\widetilde{x}_{t+1}^{i}) \widetilde{w}_{t+1}^{i}.$
- (b) Resample from the set $\{\widetilde{x}_{t+1}^1, \widetilde{x}_{t+1}^2, ..., \widetilde{x}_{t+1}^n\}$ with probabilities $\{\widetilde{w}_{t+1}^1, \widetilde{w}_{t+1}^2, ..., \widetilde{w}_{t+1}^n\}$ *n* times to obtain the samples $x_{t+1}^i, i = 1, 2, ..., n$.
- 4. Set t = t + 1, and return to Step 2.



Illustration

