

Chapter 1. Probability and Distribution

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用有效的方式收集数据的问题的研究，构成了数理统计学中两个分支，其一叫做抽样理论，其二叫做实验设计（试验设计）。

1 Some note

The number of permutations of n distinct objects taken r at a time is

$${}_nP_r = n(n-1)\cdots(n-r+1) = \frac{n!}{(n-r)!}, r=0, 1, 2, \dots, n.$$

The number of combinations of n distinct objects taken r at a time is

$$\binom{n}{r} = \frac{n(n-1)\cdots(n-r+1)}{r!} = \frac{n!}{r!(n-r)!}, r=0, 1, 2, \dots, n.$$

The binomial coefficient of the term of $x^r y^{n-r}$ in the expansion of

$$(x+y)^n = \sum_{r=0}^n \binom{n}{r} x^r y^{n-r}$$

is $\binom{n}{r}$, where n is a positive integer and r is a non-negative less than or equal to n .

The number of ways in which a set of n distinct objects can be partitioned into k subsets with n_1 objects in the first subset, n_2 objects in the second subset,...,and n_k objects in the k -th subset is

$$\binom{n}{n_1, \dots, n_k} = \frac{n!}{n_1! \cdots n_k!},$$

which is the multinomial coefficient of the term of $x_1^{n_1} \cdots x_k^{n_k}$ in the expansion of $(x_1 + \cdots + x_k)^n$, where $n_1 + \cdots + n_k = n$.

Here are some useful formulae

$$\bullet \quad \binom{x}{r} = \binom{n-1}{r} + \binom{n-1}{r-1}$$

2 Sample Space

An experiment is a process of observation or measurement. The results obtained from an experiment are called the *outcomes* of the experiment. The set of all possible outcomes of an experiment is called the *sample space* denoted by \mathbb{S} . Each outcome in a sample space is called an *elements* or a *sample point*. An *event* is a subset of a sample space.

According to the number of elements they obtain, sample space can be classified into *discrete* sample and *continuous* sample space. A sample space is discrete, if the number of elements is finite or countable. A sample space is continuous, if the sample space consists of a continuum.

Events has operation as complement, union and intersection.

3 Properties of probability

Defintion 1.1 (Probability of a set). Let A be a subset of the sample space S , then $\Pr(A)$ is said to be the probability of A if

- i. $\Pr(A) \geq 0$ and $\Pr(S) = 1$;
- ii. If A_1, A_2, \dots is a sequence of mutually exclusive of S , then

$$\Pr\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} \Pr(A_i)$$

4 Conditional Probability

Definition (Conditional probability of two sets). If A and B are two events in the sample space S , the conditional probability of B given A is defined by

$$\Pr(B|A) = \frac{\Pr(A \cap B)}{\Pr(A)}$$

where $\Pr(A) > 0$.

Defintion (Independency of two events). Two events A and B are said to be *independent*, denoted by $A \perp B$, if

$$\Pr(A \cap B) = \Pr(A) \times \Pr(B)$$

Theorem (Independcy). Let $A \perp B$, then $A \perp B'$ and $A' \perp B'$.

Definition (Mutual independency). Event A_1, \dots, A_n are said mutually independent, if the probability of the intersection of any 2, 3, ..., or n of these events equals the product of their respective probabilities.

Definition (Partition). A partition of the sample space S is a collection of mutually exclusive sets B_1, \dots, B_n , such that $S = \bigcup_{i=1}^n B_i$.

Bayes formula. Let B_1, \dots, B_n be a partition of sample space S and A be an evnet, then

- i. Law of total probability:

$$\Pr(A) = \sum_{i=1}^n \Pr(A|B_i) \times \Pr(B_i)$$

- ii. Bayes formula:

$$\Pr(B_j|A) = \frac{\Pr(A|B_j) \times \Pr(B_j)}{\sum_{i=1}^n \Pr(A|B_i) \times \Pr(B_i)} \quad \text{for } j = 1, \dots, n.$$

5 Probility Distribution

Defintion (Random variable)s. A *random variable* (r.v.) is a funtion from a sample space S into the real numbers. An r.v. is *discrete* if it takes values in a finite or countable set. An r.v. is *continus* if it take values over some interval. (Just like we trans angle to radian)

Definition (Probability mass function). If X is a discrete r.v., the function denoted by

$$p(x) = \Pr(X = x)$$

for each x within the range of X is called the *probability mass function* (pmf) of X .

Definition (Probability density function). Let X be a continuous r.v.. A non-negative function $f(x)$ is called the *probability density function*(pdf) of X , if

$$\Pr(A) = \int_A f(x) dx$$

In other word:

$$\Pr(a \leq X \leq b) = \int_a^b f(x) dx$$

Definition(Cumulative density function). The *cumulative distribution function*(cdf) of an r.v. X is defined by

$$F(x) = \Pr(X \leq x) = \begin{cases} \sum_{t \leq x} p(t) & , \text{ if } X \text{ is discrete,} \\ \int_{-\infty}^x f(x) dx & , \text{ if } X \text{ is continuous.} \end{cases}$$

6 Bivariate Distributions

Definition (Bivariate pmf). If X and Y are two discrete r.v.'s, the function defined by

$$p(x, y) = \Pr(X = x, Y = y)$$

for each pair of values (x, y) within the range of X and Y is called the joint pmf of X and Y .

Similarly, a bivariate function $f(x, y)$ is called a joint pdf of the continuous r.v.'s X and Y if

$$\Pr\{(X, Y) \in A\} = \int \int_A f(x, y) dx dy$$

for a region A in the domain of (X, Y) .

Then the joint distribution(or joint cdf) of r.v.'s (X, Y) is defined by

$$\begin{aligned} F(x, y) &= \Pr(X \leq x, Y \leq y) \\ &= \begin{cases} \sum_{s \leq x, t \leq y} p(s, t) & , \text{ if } X \text{ and } Y \text{ are discrete,} \\ \int_{-\infty}^x \int_{-\infty}^y f(x, y) dx dy & , \text{ if } X \text{ and } Y \text{ are continuous.} \end{cases} \end{aligned}$$

6.1 Marginal and conditional distributions

Let $p(x, y)$ be the joint pmf of discrete r.v.'s (X, Y) . The *marginal* pmfs of X and Y are defined by

$$p(x) = \sum_y p(x, y) \quad \text{and} \quad p(y) = \sum_x p(x, y),$$

respectively. The *conditional* pmfs of X given $Y = y$ and Y given $X = x$ are defined by

$$p(x|y) = \frac{p(x, y)}{p(y)}, p(y) \neq 0 \quad \text{and} \quad p(y|x) = \frac{p(x, y)}{p(x)}, p(x) \neq 0$$

respectively.

6.2 Independency of two random variables

Let $f(x, y)$ denote the joint pdf of r.v.'s (X, Y) , and $f(x)$ and $f(y)$ be their marginal pdfs. The r.v.'s X and Y are said to be *independent*, denoted by $X \perp Y$, if

$$\begin{aligned} f(x, y) &= f(x) \times f(y), \forall (x, y) \in \mathcal{S}_{(X, Y)}, \quad \text{or} \\ F(x, y) &= F(x) \times F(y), \forall (x, y) \in \mathcal{S}_{(X, Y)}. \end{aligned}$$

where $\mathcal{S}_{(X, Y)} \triangleq \{(x, y): f(x, y) > 0\}$ denotes the joint *support* of (X, Y) .

6.3 Expectation, Variance and Moments

The expectation of $g(X)$ is defined by

$$E\{g(X)\} = \begin{cases} \sum g(x)p(x) & , \text{if } X \text{ is discrete} \\ \int_{-\infty}^{+\infty} f(x)g(x)dx & , \text{if } X \text{ is continuous.} \end{cases}$$

When $g(X) = X$, the expectation of X , measure the *central location* of the pdf of X . (how about expectation of multi-r.v.) Let $\mu = E(X)$, then

$$\sigma^2 = \text{Var}(X) = E(X - \mu)^2 = E(X^2) - \mu^2$$

is a measure of the *dispersion* of the pdf of X . $\sigma = \sqrt{\text{Var}(X)}$ is called the standard deviation. We also define covariance as

$$\text{Cov}(X_1, X_2) = E\{(X_1 - \mu_1)(X_2 - \mu_2)\}, \text{ where } \mu_i = E(X_i), i = 1, 2.$$

Covariance also can be calculated by $\text{Cov}(X_1, X_2) = E(X_1 X_2) - E(X_1)E(X_2)$.

The r -th moment of the r.v. X is defined by $\mu'_r = E(X^r)$. The r -th central moment of the r.v. X is defined by $\mu_r = E[(X - \mu)^r]$. $\mu_3 = E[(X - \mu)^3]$ is a measure of *asymmetry* of the pdf of X . The fourth central moment $\mu_4 = E[(X - \mu)^4]$ is a measure of kurtosis (峰态), which is the *degree* of flatness of a density near its center.

7 Moment Generating Function

For an r.v. X , if $E(e^{tX})$ exists for any $t \in (-\varepsilon, \varepsilon)$ for some $\varepsilon > 0$, then

$$M_X(t) = E(e^{tX})$$

is called the *moment generating function*(mgf) of X . By using Maclaurin's expansion, we have

$$M_X(t) = E\left\{\sum_{n=0}^{\infty} \frac{(tX)^n}{n!}\right\} = \sum_{n=0}^{\infty} E\left(\frac{t^n}{n!} X^n\right) = \sum_{n=0}^{\infty} \frac{t^n}{n!} E(X^n)$$

Thus, the n -th moment can be obtained by

$$\mu'_n = E(X^n) = \left. \frac{d^n M_X(t)}{dt^n} \right|_{t=0}.$$

8 Useful Distribution

8.1 Bivariate Normal Distribution

It is well known that X is normally distributed with mean μ and variance σ^2 , denoted by $X \sim N(\mu, \sigma^2)$, if its pdf is

$$N(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\}, -\infty < x < +\infty$$

To introduce the bivariate normal distribution, first of all, we define the *correlation coefficient* of X_1 and X_2 by

$$\rho = \text{Corr}(X_1, X_2) = \frac{\text{Cov}(X_1, X_2)}{\sigma_1 \sigma_2}.$$

A random vector $\mathbf{x} = (X_1, \dots, X_d)^T$ is said to follow a d -dimensional normal distribution, if its joint pdf is given by

$$N_d(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{2\pi}|\boldsymbol{\Sigma}|^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}(\mathbf{x} - \boldsymbol{\mu})\right\}$$

for $\mathbf{x} \in \mathbb{R}^d$, where the mean vector $\boldsymbol{\mu} \in \mathbb{R}^d$ and the covariance matrix $\boldsymbol{\Sigma}$ is positive definite, denoted by $\boldsymbol{\Sigma} > 0$. We will write $\mathbf{x} \sim N_d(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ or $\mathbf{x} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$.

8.2 Beta distribution and Gamma distribution

The gamma function

$$\Gamma(\alpha) = \int_0^{\infty} x^{\alpha-1} e^{-x} dx$$

which is well defined for $\alpha > 0$.

$$X \sim \text{Beta}(x|a, b) = \frac{x^{a-1}(1-x)^{b-1}}{B(a, b)}, 0 \leq x \leq 1.$$

where $B(a, b) = \Gamma(a)\Gamma(b)/\Gamma(a+b)$ is the beta function. $E(X) = \frac{a}{a+b}$. The beta distribution is the conjugate prior for the binomial likelihood.

$$X \sim \text{Gamma}(x|\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x}, x \geq 0.$$

$$E(X) = \frac{\alpha}{\beta},$$

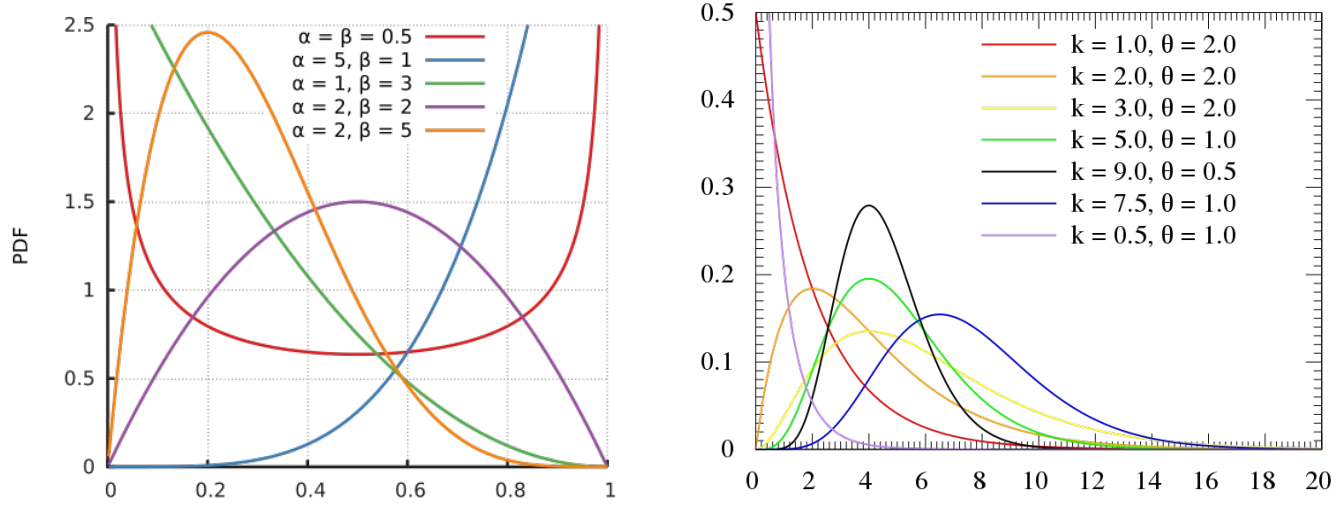


Figure 1. The beta(left) and gamma(right) distribution with different parameter.

8.3 Categorical Distribution

8.4 Zero-inflated Poisson Distribution

9 Inverse Byes Formulae

1. Point-wise formula

$$\begin{aligned}
 f_Y(y) &= \frac{f_{(Y|X)}(y|x)}{f_{(X|Y)}(x|y)} f_X(x) \\
 \Rightarrow \int_{S_Y} f_Y(y) dy &= \int_{S_Y} \frac{f_{(Y|X)}(y|x)}{f_{(X|Y)}(x|y)} f_X(x) dy = f_X(x) \int_{S_Y} \frac{f_{(Y|X)}(y|x)}{f_{(X|Y)}(x|y)} dy = 1 \\
 \Rightarrow f_X(x) &= \left\{ \int_{S_Y} \frac{f_{(Y|X)}(y|x)}{f_{(X|Y)}(x|y)} dy \right\}^{-1}, \text{ for any } x \in S_X
 \end{aligned}$$

2. Function-wise formula

Substituing point-wise formula to $f_Y(y) = \frac{f_{(Y|X)}(y|x)}{f_{(X|Y)}(x|y)} f_X(x)$, we can get(symmetry)

$$f_X(x) = \left\{ \int_{S_X} \frac{f_{(X|Y)}(x|y_0)}{f_{(Y|X)}(y_0|x)} dx \right\}^{-1} \frac{f_{(X|Y)}(x|y_0)}{f_{(Y|X)}(y_0|x)}$$

for all $x \in S_X$ and an arbitrarily fixed $y_0 \in S_Y$.

3. Sampling-wise formula

By dropping the normalizing constant in functione-wise formula, we obtain

$$f_X(x) \propto \frac{f_{(X|Y)}(x|y_0)}{f_{(Y|X)}(y_0|x)}$$

For discrete r.v.,the point-wise formula and smaple-wise formula for all $x \in S_Y$ and fiexed $y_0 \in S_Y$.

$$\Pr(X = x) = \left\{ \sum_{y \in S_Y} \frac{\Pr(Y = y|X = x)}{\Pr(X = x|Y = y)} \right\}^{-1}, \Pr(X = x) \propto \frac{\Pr(X = x|Y = y_0)}{\Pr(Y = y_0|X = x)}$$

10 References

- https://en.wikipedia.org/wiki/Central_moment
- https://en.wikipedia.org/wiki/Beta_distribution