

§4.7. Some Useful Facts

P4.61

S.3.1

I. Independence:

Easy to see (intuitively)

if X and Y are independent r.v.s

$g(\cdot)$ and $h(\cdot)$ are two functions,

then the random variables $U = g(X)$

and $V = h(Y)$ are also independent

In particular,

if X and Y are independent

then $aX+b$ and $cY+d$ are

also independent where a, b, c, d

are constants ($a \neq 0, c \neq 0$).

More generally, if $\{x_1, x_2, \dots, x_m, x_{m+1}, \dots, x_n\}$ are mutual independent rv's and

$g(x_1, \dots, x_m)$ and $h(y_1, \dots, y_{n-m})$ are two ordinary functions, then the rv's

$g(x_1, \dots, x_m)$ and $h(x_{m+1}, \dots, x_n)$ are also independent.

For example, if $\{x_1, x_2, x_3, x_4, x_5\}$ are independent, then

$$2x_2 + e^{x_3} - x_1 x_2^2 x_3 \quad \text{and}$$

$x_5 x_4 + \frac{5x_4}{e^{x_5}}$, say are also independent.

The essential thing here is that there is no common v.v. in both functions

For example, if $\{Z_1, Z_2, \dots, Z_n\}$ are independent random variables, (even if not in the same type) then, for ex. $Z_1 + Z_2 + \dots + Z_{n-1}$ and Z_n are independent. But you can not say $Z_1 + Z_2 + \dots + Z_n$ and Z_n are independent.

II. Mixing case.

Suppose X is a discrete r.v. but

Y is a continuous r.v.

Then joint cdf can be similarly defined as $F(x, y) = P\{X \leq x, Y \leq y\}$ for all $(-\infty < x < \infty, -\infty < y < \infty)$.

Then, again, they are independent if joint cdf is the product of two marginal cdfs. i.e.

$$F(x, y) = F_X(x) \cdot F_Y(y)$$

§3.5

We can also define the so-called
"joint pdf-pmf" function

$$f(x, y) \quad \left[\begin{array}{l} x = \{x_1, x_2, \dots\} \\ y: (-\infty, +\infty) \end{array} \right]$$

such that for any real numbers a, b
we have

$$F(a, b) = \sum_{x \leq a} \int_{-\infty}^b f(x, y) dy$$

(Similar for the mixing case for
more than 2 random variables)
It can be proved that X and Y
are independent if and only if

$$\text{"joint pdf-pmf"} = (\text{Marginal pdf}) \cdot (\text{Marginal pmf})$$

III Is a constant random variable?

For example, is the constant 10 a random variable?

Usually not, "we emphasise ^qr.v. ⁸depends upon the change!!" (This difference is important!!)

However, for convenience, some times we may view a constant as a "random variable".

Questions. If the constant, 10, say is viewed as a random variable X , (i.e. $X(\omega) \equiv 10$), then

① discrete type or continuous type?

(Discrete type, since only taking one value)

② "pmf" ? "cdf" ?

pmf: simply, $P\{X=10\}=1$.

cdf: For any $x \in (-\infty, +\infty)$

$$F_X(x) = \Pr\{X \leq x\}$$

Hence, if $x < 10$

$$\text{Then } F_Z(x) = P\{Z \leq x\} = 0$$

$$(\because Z \equiv 10, \Rightarrow \{Z \leq x\} = \{10 \leq x\} = \phi) \xrightarrow{\text{Impossible event.}}$$

If $x \geq 10$,

$$\text{then } F_Z(x) = P\{Z \leq x\} = 1$$

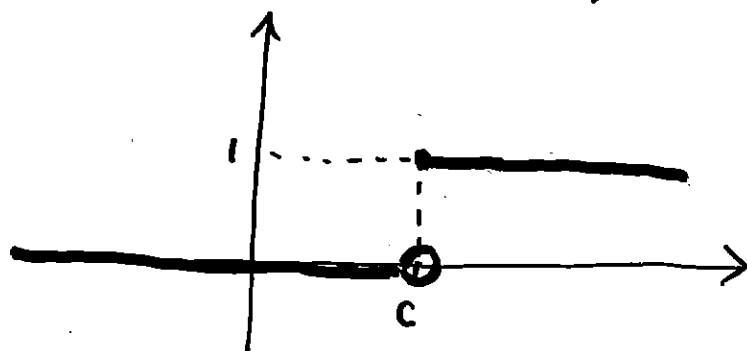
($\because Z \equiv 10 \Rightarrow \{Z \leq x\} = \{10 \leq x\}$ must be true and thus certain event!)

In general, for any constant c ,

(if viewed as a random variable Z) then

$$\text{pmf: } P\{Z = c\} = 1 \quad (\text{discrete type})$$

$$\text{cdf: } F_Z(x) = P\{Z \leq x\} = \begin{cases} 1 & \text{if } x \geq c \\ 0 & \text{if } x < c \end{cases}$$



③ Independent with other random variable?

If $X = 10$ (constant) is viewed as a r.v.

and Y is another r.v. (either type?)

" X and Y are independent?"

Conclusion: Yes!

Proof. $F(x, y) \neq F_X(x) F_Y(y)$ (for all x , and y)

If $x < 10$, then $F(x, y) = P\{X \leq x, Y \leq y\} = P\{10 \leq x, Y \leq y\}$
 $= P(\emptyset) = 0$

but $F_X(x) F_Y(y) = P\{X \leq x\} P\{Y \leq y\} = P\{10 \leq x\} P\{Y \leq y\}$
 $= P(\emptyset) \cdot P\{Y \leq y\} = 0 \times P\{Y \leq y\} = 0$

If $x \geq 10$, then $F(x, y) = P\{X \leq x, Y \leq y\} = P\{10 \leq x, Y \leq y\}$
 $= P\{Y \leq y\}$ (Think why here!)

but $F_X(x) F_Y(y) = P\{X \leq x\} P\{Y \leq y\} = P\{10 \leq x\} P\{Y \leq y\}$
 $= P(1) P\{Y \leq y\} = P\{Y \leq y\}$

Hence always true. (Similar for any constant c)

II. Normal distribution:

1. If $Z \sim N(\mu, \sigma^2)$

then $E(Z) = \mu$, $\text{Var}(Z) = \sigma^2$.

2. If $Z \sim N(\mu, \sigma^2)$

then $a \cdot Z + b \sim N(a\mu + b, a^2\sigma^2)$

i.e. $E(aZ + b) = a\mu + b$

$$\text{Var}(a \cdot Z + b) = a^2 \sigma^2$$

Indeed, $E(aZ + b) = a E(Z) + b = a\mu + b$

$$\begin{aligned} \text{Var}(a \cdot Z + b) &= \text{Var}(aZ) = a^2 \cdot \text{Var}(Z) \\ &= a^2 \cdot \sigma^2 \end{aligned}$$

3. If $Z \sim N(\mu_1, \sigma_1^2)$, $Y \sim N(\mu_2, \sigma_2^2)$
and Z and Y are independent

then $Z + Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$.

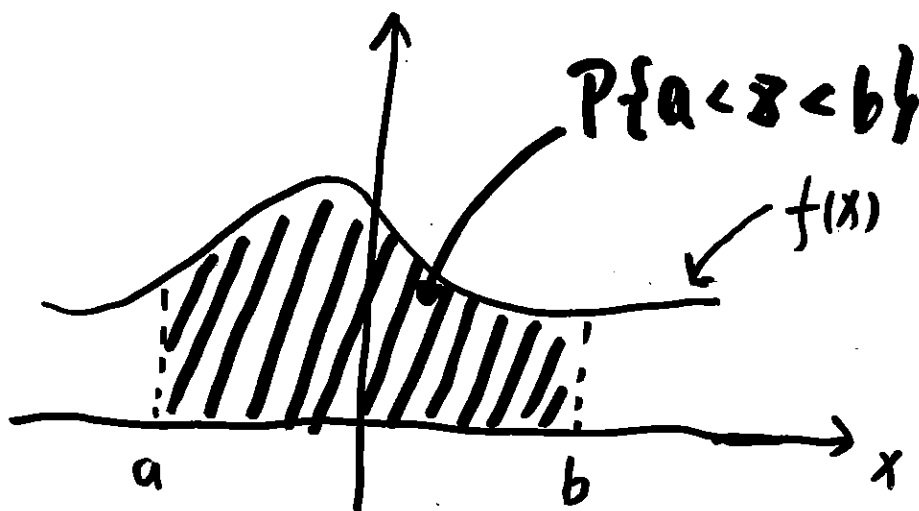
III. Other Continuous Random Variables:

1. Geometric meaning: (True for any continuous r.v.)

If X is a continuous r.v. with c.d.f. $F(x)$ and p.d.f. $f(x)$, then

$$\begin{aligned}\forall a < b, \quad P\{a < X < b\} &= F(b) - F(a) \\ &= \int_a^b f(x) dx\end{aligned}$$

the geometric meaning of $\int_a^b f(x) dx$ is the area between a and b under the curve $f(x)$ (and so is the probability that $\{a < X < b\}$).



(a, b can be infinities !!)

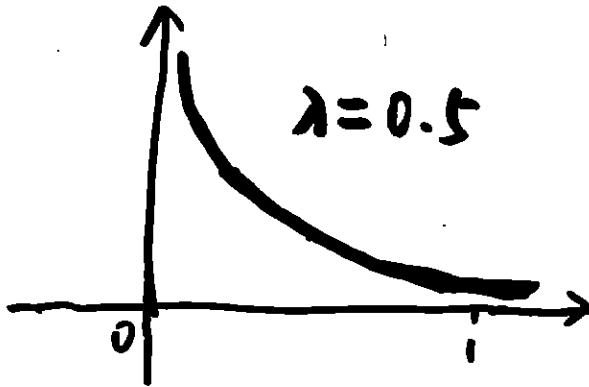
2. Exponential Distribution:

$$\text{P.d.f.} \quad f(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x \geq 0 \\ 0 & \text{if } x < 0 \end{cases}$$

$$\text{c.d.f.} \quad F(x) = \begin{cases} 0 & \text{if } x < 0 \\ x & \text{if } 0 \leq x \leq 1 \\ 1 & \text{if } x > 1 \end{cases}$$

$$E(X) = \frac{1}{\lambda} \quad \text{Var}(X) = \frac{1}{\lambda^2}$$

Graph of $f(x)$ (P.d.f)



So, X can not take negative values.

§2. Distributions Derived from the Normal Distribution

The following three distributions play extremely important role in Statistics. (See Sec 6.2 of the Reference book)

3. χ^2 Distribution.

① Def: If X_1, X_2, \dots, X_n are i.i.d. (independent, identically distributed) with common standard normal distribution, i.e. $X_i \sim N(0,1)$ ($\forall i$) and $\{X_1, X_2, \dots, X_n\}$ are mutually independent.

then
$$Y = X_1^2 + X_2^2 + \dots + X_n^2$$

is called a χ^2 random variable with degree freedom of n .

② Notation: $Y \sim \chi^2(n)$

③ P.d.f:

$$f(x) = \begin{cases} \frac{1}{2^{\frac{n}{2}} \cdot \Gamma(\frac{n}{2})} x^{\frac{n}{2}-1} \cdot e^{-\frac{x}{2}}, & x \geq 0 \\ 0 & x < 0 \end{cases}$$

(The following three continuous distributions are useful in Statistics and will be discussed in another course)

3. χ^2 Distribution:

① Def: If X_1, X_2, \dots, X_n are i.i.d. (independent, identically distributed) with common standard normal distribution, i.e. $X_i \sim N(0,1)$ ($\forall i$) and $\{X_1, X_2, \dots, X_n\}$ are mutually independent.

then $Y = X_1^2 + X_2^2 + \dots + X_n^2$

is called a χ^2 random variable with degree freedom of n .

② Notation: $Y \sim \chi^2(n)$

③ P.d.f:

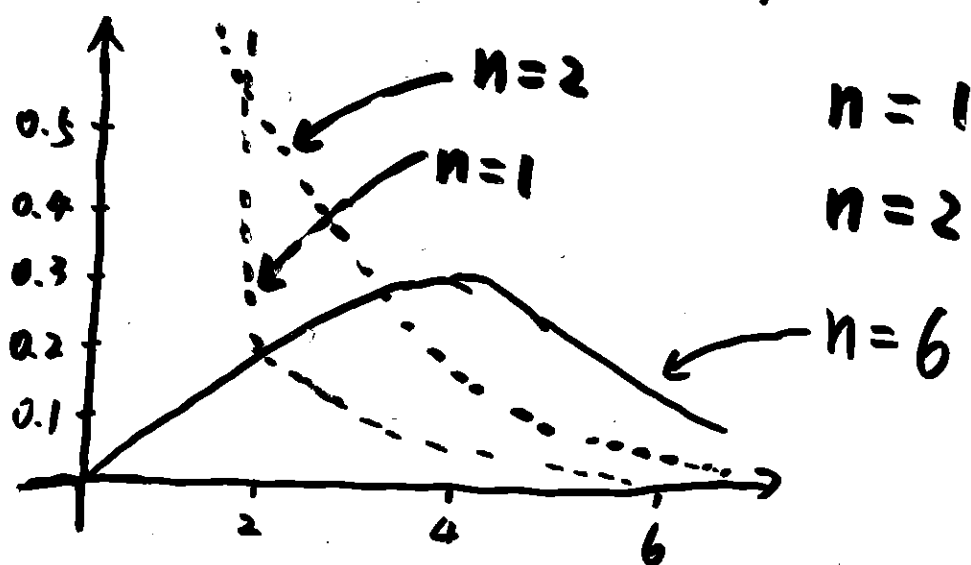
$$f(x) = \begin{cases} \frac{1}{2^{\frac{n}{2}} \cdot \Gamma(\frac{n}{2})} x^{\frac{n}{2}-1} \cdot e^{-\frac{x}{2}}, & x \geq 0 \\ 0 & x < 0 \end{cases}$$

④ Expectation and Variance:

If $Y \sim \chi^2(n)$ (\Rightarrow moment generating function $M(t) = (1-2t)^{-\frac{n}{2}}$)

then $E(Y) = n$, $\text{Var}(Y) = 2n$

⑤ Graph of P.d.f. (depend upon the parameter n where n is a positive integer)



⑥ $P(\frac{n}{2})$:

if n is even then $P(\frac{n}{2}) = [(\frac{n}{2}) - 1]!$

for example, $n=4$, $P(\frac{4}{2}) = [(\frac{4}{2}) - 1]! = 1! = 1$

$n=10$, $P(\frac{10}{2}) = [(\frac{10}{2}) - 1]! = 4! = 24$

if n is odd then $P(\frac{n}{2}) = \frac{n-2}{2} \cdot \frac{n-4}{2} \cdots \frac{1}{2} \cdot \sqrt{\pi}$

e.x., $P(\frac{5}{2}) = \frac{5-2}{2} \cdot \frac{5-4}{2} \cdot \sqrt{\pi} = \frac{3}{2} \sqrt{\pi}$

4. t-distribution (student-distribution)

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P5.9

① Def: If $Z \sim N(0,1)$, $Y \sim \chi^2(n)$

Z and Y are independent

$$\text{then } T = \frac{Z}{\sqrt{\frac{Y}{n}}}$$

is called to obey the t-distribution with n degree of freedom. (Also a parameter)

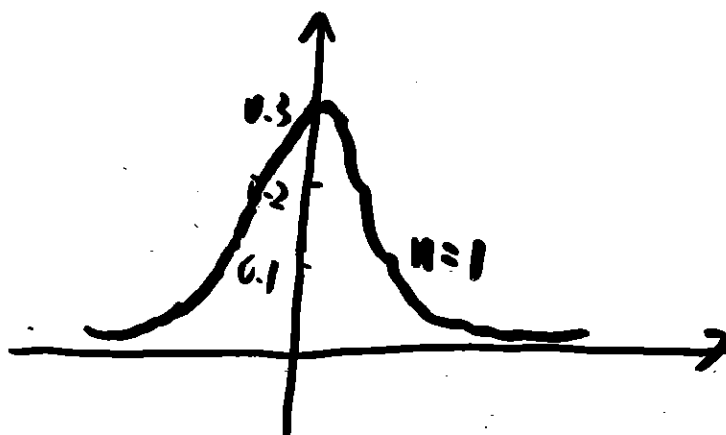
② Notation: $T \sim t(n)$

③ P.d.f.
$$f(x) = \begin{cases} \frac{\Gamma(\frac{n+1}{2})}{\Gamma(\frac{n}{2}) \sqrt{n\pi}} (1+x^2/n)^{-\frac{n+1}{2}} \\ -\infty < x < +\infty \end{cases}$$

So, $f(x)$ is symmetric with $x=0$

④ Property: $t(n) \rightarrow N(0,1)$ ($n \rightarrow \infty$)

⑤ Graph



5. F-distribution:

① Def: If $Z \sim \chi^2(m)$, $Y \sim \chi^2(n)$

Z and Y are independent

then $F = \frac{Z/m}{Y/n}$ is called to obey

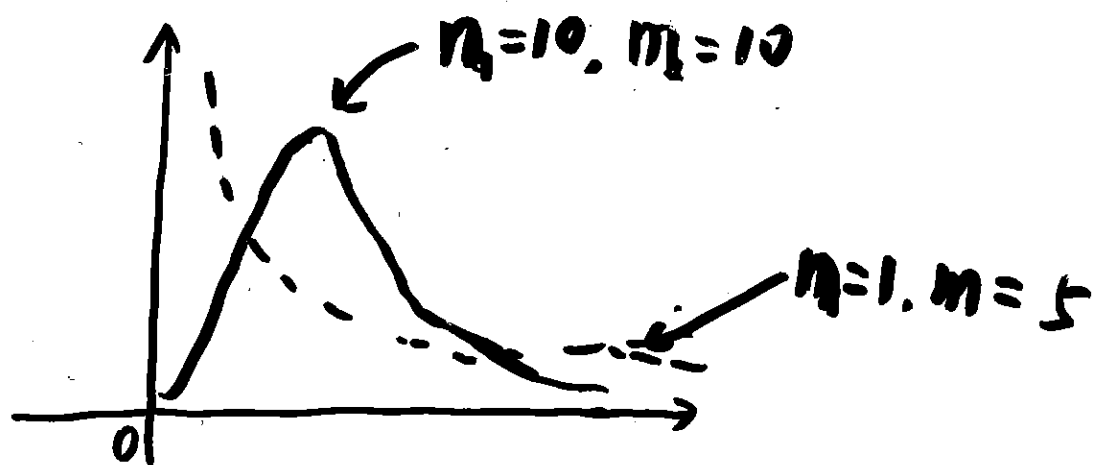
the F-distribution with (m, n) degree of freedom.

② Notation: $F \sim F(m, n)$ (Two parameter)

③ P.d.f:

$$f(x) = \begin{cases} \frac{P(\frac{m+n}{2})}{P(\frac{m}{2}) \cdot P(\frac{n}{2})} \left(\frac{m}{n}\right) \left(\frac{m}{n}x\right)^{\frac{m}{2}-1} \left(1+\frac{m}{n}x\right)^{-\frac{m+n}{2}} & \text{if } x > 0 \\ 0 & \text{if } x \leq 0 \end{cases}$$

④ Graph:



⑤ Property: If $X \sim F(m, n)$, then $X^{-1} \sim F(n, m)$.

If $T \sim t(n)$, then $T^2 \sim F(1, n)$

IV. Discrete Random Variables:

1. Bernoulli: P.m.f:

x	0	1
$p(x)$	$1-p$	p

where $0 \leq p \leq 1$

$$E(\bar{x}) = p$$

$$\text{Var}(\bar{x}) = p q$$

2. Binomial: $\bar{x} \sim B(n, p)$

$$\text{P.m.f: } P(k) = P\{\bar{x} = k\} = \binom{n}{k} p^k q^{n-k}$$

$$(k = 0, 1, 2, \dots, n)$$

$$\text{where } 0 \leq p \leq 1 \text{ and } p + q = 1$$

$$\text{If } \bar{x} \sim B(n, p)$$

$$\text{then } E(\bar{x}) = np$$

$$\text{Var}(\bar{x}) = npq$$

3. Poisson: $X \sim \text{Poisson}(\lambda)$

$$\text{P.m.f.: } P(k) = P\{X=k\} = e^{-\lambda} \cdot \frac{\lambda^k}{k!}$$

$$(k=0, 1, 2, 3, \dots)$$

$\lambda > 0$ is a constant.

If $X \sim \text{Poisson}(\lambda)$

$$\text{then } E(X) = \lambda, \quad \text{Var}(X) = \lambda.$$

4. Geometric distribution:

P.m.f.: all possible values: $\{1, 2, \dots\}$

$$P_k = P\{X=k\} = p \cdot q^{k-1} \quad (k=1, 2, \dots)$$

where $p > 0, q > 0, p+q=1$

$$E(X) = \frac{1}{p} \quad \text{Var}(X) = \frac{q}{p^2} = \frac{1-p}{p^2}$$

Summary of Chapter 4

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I. Basic Concepts:

1. Expected value: $E(\bar{X})$ (Weighted average)

2. Variance: $\text{Var}(\bar{X}) = E[(\bar{X} - E(\bar{X}))^2]$

3. Standard Deviation: $\sqrt{\text{Var}(\bar{X})}$

4. Covariance

$$\text{Cov}(\bar{X}, Y) = E[(\bar{X} - E(\bar{X}))(Y - E(Y))]$$

5. Correlation

$$\rho(\bar{X}, Y) = \frac{\text{Cov}(\bar{X}, Y)}{\sqrt{\text{Var}(\bar{X}) \cdot \text{Var}(Y)}}$$

6. Function of random variables:

$g(\bar{X})$, $g(\bar{X}, Y)$ etc.

II. Basic Properties:

1. Expectation:

$$\textcircled{1} E(c) = c \quad \text{for constant } c.$$

$$* \textcircled{2} E(a \cdot X + b \cdot Y) = a \cdot E(X) + b \cdot E(Y)$$

(a, b: constants) (linear property!)

* $\textcircled{3}$ If X and Y are independent, then

$$E(X \cdot Y) = E(X) \cdot E(Y)$$

2. Variance:

$$\textcircled{1} \text{Var}(c) = 0 \quad \text{for constant } c.$$

$$\textcircled{2} \text{Var}(a \cdot X + b) = a^2 \cdot \text{Var}(X)$$

(a, b: constants)

$\textcircled{3}$ If X and Y are independent, then

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y).$$

3. Covariance:

① $\text{Cov}(a, b) = 0$ for constants a and b

② $\text{Cov}(a_1 X + b_1 Y + c_1, a_2 U + b_2 V + c_2)$

$$= a_1 a_2 \text{Cov}(X, U) + a_1 b_2 \text{Cov}(X, V)$$

$$+ a_2 b_1 \text{Cov}(Y, U) + a_2 b_2 \text{Cov}(Y, V)$$

for constants $a_1, b_1, c_1, a_2, b_2,$ and c_2
and random variables $X, Y, U,$ and V .

③ If X and Y are independent

then $\text{Cov}(X, Y) = 0$

(but the converse is not true)

(for bivariate normal the converse holds true)

4. Correlation: for any two r.v.s

$$-1 \leq \rho(X, Y) \leq 1$$

III. Calculation:

1. Function of random variables

$$E[g(x)] = \int_{-\infty}^{+\infty} g(x) \cdot f(x) dx$$

$$E[g(x)] = \sum_i g(x_i) \cdot p(x_i)$$

$$E[g(x, y)] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} g(x, y) f(x, y) dx dy$$

$$E[g(x, y)] = \sum_j \sum_i g(x_i, y_j) \cdot f(x_i, y_j)$$

2. Variance:

$$\text{Var}(x) = E(x^2) - (E(x))^2$$

3. Covariance:

$$\text{Cov}(x, y) = E(x \cdot y) - E(x) \cdot E(y)$$

IV. Some important facts:

1. Normal distribution:

① If $Z \sim N(\mu, \sigma^2)$ then $E(Z) = \mu$

$$\text{Var}(Z) = \sigma^2$$

② If $Z \sim N(\mu, \sigma^2)$, a, b are constants, $a \neq 0$,
then $a \cdot Z + b \sim N(a\mu + b, a^2\sigma^2)$

③ If $Z \sim N(\mu_1, \sigma_1^2)$, $Y \sim N(\mu_2, \sigma_2^2)$
and Z and Y are independent

then $Z + Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$

④ If $Z \sim N(\mu, \sigma^2)$,

$$\text{let } Y = \frac{Z - \mu}{\sigma}$$

then $Y \sim N(0, 1)$

2. Poisson distribution

If $X \sim \text{Poisson}(\lambda)$ then $E(X) = \lambda$

$$\text{Var}(X) = \lambda$$

3. Binomial distribution

If $X \sim B(n, p)$ then $E(X) = np$

$$\begin{aligned}\text{Var}(X) &= npq \\ &= np(1-p)\end{aligned}$$

4. Exponential distribution:

If $X \sim \text{Expon}(\lambda)$

$$\text{then } E(X) = \frac{1}{\lambda} \quad \text{Var}(X) = \frac{1}{\lambda^2}$$

$$\text{P.d.f. } f(x) = \begin{cases} \lambda e^{-\lambda x} & x \geq 0 \\ 0 & x < 0 \end{cases}$$

V. Other useful facts:

See §4.6