# CSOR W4231: Analysis of Algorithms (sec.001) - Problem Set #2

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# Problem 1

We assume probability of  $X_1 = [R \text{ contains at least 1 element of T}]$  as  $P_1$ 

$$P_1 = 1 - (\frac{n-1}{n})^n$$

The probability of  $X_2 = [R \text{ contains no element of } S]$  as  $P_2$  is

$$P_2 = \left(\frac{n-1}{n}\right)^n$$

Since  $X_1$  and  $X_2$  is independent events, the probability that the sample is good is

$$P = P_1 * P_2 = \left(1 - \left(\frac{n-1}{n}\right)^n\right) * \left(\frac{n-1}{n}\right)^n$$

where

$$\frac{\partial (\frac{n-1}{n})^n}{\partial n} = \frac{(1-\frac{1}{n})^{n-1}}{n} > 0 \quad (when \quad n > 1)$$

from which we know that the  $(\frac{n-1}{n})^n$  has min value when n=2, which is  $\frac{1}{4}$  According to the hint

$$(1 - (1 - \frac{1}{n})^n) \ge 1 - e^{-1}$$

SO

$$P = (1 - (\frac{n-1}{n})^n) * (\frac{n-1}{n})^n \ge (1 - e^{-1}) * \frac{1}{4}$$

Then we have the statement that the probability our sample is good is larger than some positive constant, such as any number less then  $(1-e^{-1})*\frac{1}{4}$ , which is about 0.158.

(a)

#### Pseudocode:

## $\overline{\mathbf{replace}(a)}$

$$item = a[1]$$

for k = 2 to Length(a) do

$$item = \begin{cases} a[i] & \text{with probability } \frac{1}{k} \\ item & \text{with probability } 1 - \frac{1}{k} \end{cases}$$

end for

#### Correctness:

**Base case:** k = 2, the probabilities of storing a[1] and a[2] are both 1/2 then the statement is true.

**Induction hypothesis:** assume that the statement is true for  $k \geq 2$ .

*Inductive step:* show it true for case k+1

Since each of a[1], a[2], ..., a[k] will be stored with a probability of 1/k, when a[k+1] appears, there is a probability of  $1 - \frac{1}{k+1}$  that the stored item will fall in the previous range (a[1], a[2], ..., a[k]), which means each of a[1], a[2], ..., a[k] will be stored with a probability of

$$\frac{1}{k} * (1 - \frac{1}{k+1}) = \frac{1}{k}$$

There is also another probability of  $\frac{1}{k+1}$  that the stored item will be replace with a[k+1]. To conclude, the item has the property that it is uniformly distributed over all the items we have already seen, so the statement is also true in case k+1.

**Conclusion:** it follows that the statement is true for  $k \geq 2$ .

# Running Time:

The algorithm has only a for loop, so the running complexity is O(Length(a)), although we do not know the length in the process of the algorithm.

## Spaces

The space complexity is O(1).

(b)

It follows a distribution that when k-th items have been seen

$$p(a[i]) = \begin{cases} \frac{1}{2^{k-i+1}} & i \neq 1\\ \frac{1}{2^{k-1}} & i = 1 \end{cases}$$

#### Correctness:

**Base case:** k = 2, the probabilities of storing a[1] and a[2] are both 1/2 then the statement is true.

**Induction hypothesis:** assume that the statement is true for  $k \geq 2$ .

*Inductive step:* show it true for case k+1

When k-th item appears, item  $a[i](i \neq 0)$  will be stored with a probability of

$$\frac{1}{2^{k-i+1}}$$

and the probability for item a[1] is  $\frac{1}{2^{k-1}}$ .

When a[k+1] appears, there is a probability of  $\frac{1}{2}$  that the stored item will fall in the previous range (a[1], a[2], ..., a[k]), which means item  $a[i] (i \neq 0)$  will be stored with a probability of

$$\frac{1}{2} * \frac{1}{2^{k-i+1}} = \frac{1}{2^{k-i+2}}$$

and the probability for item a[1] is  $\frac{1}{2} * \frac{1}{2^{k-1}} = \frac{1}{2^k}$ .

There is also another probability of  $\frac{1}{2}$  that the stored item will be replace with a[k+1]. To conclude, the item has the property that it follows a distribution that when k-th items have been seen

$$p(a[i]) = \begin{cases} \frac{1}{2^{k-i+1}} & i \neq 1\\ \frac{1}{2^{k-1}} & i = 1 \end{cases}$$

so the statement is also true in case k+1.

**Conclusion:** it follows that the statement is true for  $k \geq 2$ .

#### **Running Time:**

The algorithm is very similar to the one in (a), so the running complexity is O(Length(a)).

### Space:

The space complexity is O(1).

#### Pseudocode:

### replace(a)

$$item = a[1]$$

for k = 2 to Length(a) do

$$item = \begin{cases} a[i] & \text{with probability } \frac{1}{2} \\ item & \text{with probability } \frac{1}{2} \end{cases}$$

#### end for

(a)

If n is old, we set  $k = \lfloor n/2 \rfloor + 1$  and use k-th order statistic (S,k) to find the median. If n is even, we set k = n/2 and k = n/2 + 1 separately and use k-th order statistic (S,k) to find the two items, then calculate the mean of the two items, which is the median of S.

(b)

### Upper bound the expected time of a subproblem of type j:

The expected running time of a subproblem of type j excluding the time on recurives calls is O(Length(S)), here  $n*(\frac{3}{4})^j \leq S \leq n*(\frac{3}{4})^{j+1}$ .

## Upper bound the expected time until a good item $a_i$ is selected:

If we want to select an item  $a_i$  that can help us throw out at least  $\frac{1}{4}$  of the input, we can judge it by looking at the length of  $|S^-|$  to assure the size of the part we can throw up, which takes O(Length(S)) time. Also we have a probability of  $\frac{1}{2}$  to pick an item  $a_i$  that we want, so we select a good  $a_i$  by spending O(Length(S)) time in total.

## Upper bound the expected time of the entire algorithm:

For the entire algorithm, the depth of j is  $\log_{\frac{3}{4}} \frac{1}{n}$ , so we have

$$O(\frac{3}{4}n) + O(\frac{3}{4}n) + O((\frac{3}{4})^2n) + O((\frac{3}{4})^2n) + \dots + O((\frac{3}{4})^{\log_{\frac{3}{4}}\frac{1}{n}}n) + O((\frac{3}{4})^{\log_{\frac{3}{4}}\frac{1}{n}}n) = O(n)$$

Also we can derive the answer according to master theorem,

$$T(n) = T(\frac{3}{4}n) + O(n)$$

The expected running time is O(n).

(a)

Suppose we have k balls, for each ball the probability that it will be discarded is

$$p = (1 - \frac{1}{n})^{k-1}$$

so the expected number of balls that should be discarded in this round is

$$E(D_k) = k * p = k * (1 - \frac{1}{n})^{k-1}$$

Now we have  $k = \epsilon n$ , so the expected balls that we have for next round is

$$E = k - E(D_k) = \epsilon n (1 - (1 - \frac{1}{n})^{\epsilon n - 1})$$

$$\leq \epsilon n (1 - e^{-\frac{1}{n} * (\epsilon n - 1)})$$

$$< \epsilon n (1 - e^{-\epsilon})$$

$$\leq \epsilon n (1 - (1 - \epsilon))$$

$$= \epsilon^2 n = k^2 / n$$

We expect to have at most  $\epsilon^2 n$  balls.

(b)

From part (a), we know that if we have k balls in some round, the expected number of balls left will be  $k^2/n$ , so suppose we have  $a_i$  balls after i-th round we can induct the equation

$$a_k = a_{k-1}^2/n = a_{k-2}^4/n^3 = \dots$$

Also from part (a), we have  $a_k < \epsilon n(1 - e^{-\epsilon})$ . Inducted by this equation, we can get that after the first round we have at most  $n(1 - e^{-1})$  balls and after the second round we have at most  $n(1 - e^{-1})^2 < 1/2$  balls, which means  $a_2 < 1/2$ . Back to the previous equation, we can now finish it

$$a_k = a_2^{2^{k-2}}/n^{1+2+\dots+2^{k-3}}$$

$$< (n/2)^{2^{k-2}}/n^{2^{k-2}-1}$$

$$= \frac{1}{2^{2^{k-2}}n}$$

If we want the experiment ends at round k, we need  $a_k < 1$ , which means

$$\frac{1}{2^{2^{k-2}n}} < 1$$
$$k > 2 + \log\log n$$

So as long as  $k > 2 + \log \log n$ , the experiment must end no later than k-th round. Equivalently, we would discard all the balls within  $O(\log \log n)$  rounds.

(a)

If the other n-1 people are not trying to access the computer and person i is trying to access, then the person i could succeed. The probability is  $p * (1-p)^{n-1}$ , which we will use  $P_1$  to denote.

(b)

We can calculate the first order derivative and makes it equal to 0.

$$\frac{\partial P_1}{\partial p} = (1-p)^{n-1} - p * (n-1) * (1-p)^{n-2} = 0$$
$$p = \frac{1}{n}$$

so we need to set  $p = \frac{1}{n}$  and probability  $P_1$  will be  $\frac{1}{n} * (1 - \frac{1}{n})^{n-1}$ .

(c)

The probability that person i fails in one step is  $1 - \frac{1}{n} * (1 - \frac{1}{n})^{n-1}$ . If the person i did not succeed to access in any of the first t = en steps, the probability will be  $(1 - \frac{1}{n} * (1 - \frac{1}{n})^{n-1})^{en}$ , which we will use  $P_2$  to denote.

$$P_2 = \left(1 - \frac{1}{n} * \left(1 - \frac{1}{n}\right)^{n-1}\right)^{en}$$

$$\leq \left(1 - \frac{1}{n} * \frac{n}{n-1} * e^{-1}\left(1 - \frac{1}{n}\right)\right)^{en}$$

$$= \left(1 - \frac{1}{en}\right)^{en}$$

$$< e^{-1}$$

so the upper bound is  $e^{-1}$ .

(d)

From part (c), we have  $(1 - \frac{1}{en})^t$  is the approximation of  $P_2$ , and if t = en,  $P_2$  should be upper bounded by  $e^{-1}$ .

Now we make  $t = ken \log n$ , then we have

$$P_2 \le e^{-k \ln n} \le \frac{1}{n^k}$$

so t should be more than  $ken \ln n$ .

(e)

We know  $1 - P_2$  is the probability that one person will succeed to access the computer in the first t steps, so if all people want to succeed, the probability is  $(1 - P_2)^n$ . We need the probability greater than  $1 - \frac{1}{n}$  then we have

$$(1-P_2)^n \ge 1 - \frac{1}{n}$$

We make  $t = -en \ln(1 - e^{-\frac{1}{n^2}})$ , so we have

$$P_2 \le e^{\ln(1-e^{-\frac{1}{n^2}})} = 1 - e^{-\frac{1}{n^2}}$$

$$1 - p_2 \ge e^{-\frac{1}{n^2}}$$

since  $e^{-1} \ge (1 - \frac{1}{n})^n$ ,  $e^{-\frac{1}{n}} \ge 1 - \frac{1}{n}$ 

$$(1 - P_2)^n \ge e^{-\frac{1}{n}} \ge 1 - \frac{1}{n}$$

so t should be more than  $-en \ln(1-e^{-\frac{1}{n^2}})$ , which is roughly  $\ln \frac{3}{2}en \approx 1.102n$  when n is large enough.