## Differentially Private Bayesian Inference

#### March 9, 2020

#### **Preliminary** 1

 $\theta$ 

: the parameter vector of multinomial distribution,  $\boldsymbol{\theta} \in [0,1]^k$ .

: Observed dataset. $\mathbf{x} \in \mathcal{X}^n$  $\mathbf{x}$ 

: Dirichlet distribution. The prior or posterior distribution over  $\theta$ .  $Dir(\alpha)$ 

 $\mathsf{DirP}(\alpha, \theta, \mathbf{x})$ : posterior distribution over  $\theta$  from Bayesian inference given prior distribution  $\mathsf{Dir}(\alpha)$ 

and observed data set  $\mathbf{x}$ .

: Hellinger Distance between two distributions.  $\mathcal{H}(\mathsf{Dir}(\alpha_1),\mathsf{Dir}(\alpha_2)) = \sqrt{1 - \frac{\mathsf{B}(\frac{\alpha_1 + \alpha_2}{2})}{\sqrt{\mathsf{B}(\alpha_1)\mathsf{B}(\alpha_2)}}}$  $\mathcal{H}(\cdot,\cdot)$ 

 $u(\mathbf{x}, r)$ : scoring function given  $\alpha$  and  $\theta$  for candidate r.  $u(\mathbf{x}, r) = -\mathcal{H}(\mathsf{DirP}(\alpha, \theta, \mathbf{x}), r)$ 

GS: Global sensitivity of Hellinger distance.  $GS = \sqrt{1 - \pi/4}$ 

 $LS(\mathbf{x})$ : Local sensitivity of Hellinger distance for  $\mathbf{x}$ .

 $LS(\mathbf{x}) = \max_{\mathbf{x}' \in \mathcal{X}^n, \mathbf{adj}(\mathbf{x}, \mathbf{x}'), r \in \mathcal{R}_{\alpha}} |\mathcal{H}(\mathsf{DirP}(\mathbf{x}, \alpha), r) - \mathcal{H}(\mathsf{DirP}(\mathbf{x}', \alpha), r)|$   $: \gamma - \mathsf{smooth \ sensitivity.} \ S(\mathbf{x}) = \max_{\mathbf{x}' \in \mathcal{X}^n} \left\{ \frac{1}{\frac{1}{LS(\mathbf{x}')} + \gamma \cdot Himming(\mathbf{x}, \mathbf{x}')} \right\}$  $S(\mathbf{x})$ 

### Private Mechanisms

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Algorithm 1 LSDim - Calibrating noise w.r.t. \ell_1 norm
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\mathbf{x} \in \mathcal{X}^n, \mathsf{Dir}(\boldsymbol{\alpha})
         let \alpha' = \mathsf{DirP}(\alpha, \theta, \mathbf{x})
         Initialize a vector \tilde{\boldsymbol{\alpha}} = (0, \dots, 0) \in \mathbb{N}^{|\mathcal{X}|}
         For i = 1 ... |\mathcal{X}| - 1:
        \begin{split} & \det \, \eta \sim \mathsf{Lap}(0,\frac{|\mathcal{X}|}{\epsilon}) \\ & \tilde{\alpha}_i = \alpha_i + \lfloor (\alpha_i' - \alpha_i) + \eta \rfloor_0^n \\ & \tilde{\alpha}_{|\mathcal{X}|} = \alpha_{|\mathcal{X}|} + \lfloor n - \sum_{i=1}^{|\mathcal{X}|-1} \lfloor (\alpha_i' - \alpha_i) + \eta_i \rfloor_0^n \rfloor_0^n \end{split}
return \tilde{\alpha}
```

#### Algorithm 2 LSHist - Calibrating noise w.r.t. histogram sensitivity

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\begin{split} & \text{input } \mathbf{x} \in \mathcal{X}^n, \, \mathsf{Dir}(\boldsymbol{\alpha}) \\ & \text{let } \boldsymbol{\alpha}' = \mathsf{DirP}(\boldsymbol{\alpha}, \boldsymbol{\theta}, \mathbf{x}) \\ & \text{let } k = \left\{ \begin{array}{ll} 1 & \text{if} & |\mathcal{X}| = 2 \\ 2 & \text{otherwise} \end{array} \right. \\ & \textbf{Initialize } \text{a vector } \tilde{\boldsymbol{\alpha}} = (0, \dots, 0) \in \mathbb{N}^{|\mathcal{X}|} \\ & \textbf{For } i = 1 \dots |\mathcal{X}| - 1 \text{:} \\ & \text{let } \eta \sim \mathsf{Lap}(0, \frac{k}{\epsilon}) \\ & \tilde{\alpha}_i = \alpha_i + \lfloor (\alpha_i' - \alpha_i) + \eta \rfloor_0^n \\ & \tilde{\alpha}_{|\mathcal{X}|} = \alpha_{|\mathcal{X}|} + \lfloor n - \sum_{i=1}^{|\mathcal{X}|-1} \lfloor (\alpha_i' - \alpha_i) + \eta_i \rfloor_0^n \rfloor_0^n \\ & \textbf{return } \tilde{\boldsymbol{\alpha}} \end{split}
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#### Algorithm 3 EHD - Instantiation of the exponential mechanism

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observed data set \mathbf{x} \in \mathcal{X}^n, prior: \mathsf{Dir}(\boldsymbol{\alpha}), \epsilon let \mathsf{Dir}(\boldsymbol{\alpha}') = \mathsf{DirP}(\mathbf{x}, \boldsymbol{\alpha}).

let GS be the global sensitivity for \mathbf{x}.

set z = r with probability \frac{\exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x}, \boldsymbol{\alpha}), r')}{2 \cdot GS})}{\sum_{r' \in \mathcal{R}_{\boldsymbol{\alpha}}} \exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x}, \boldsymbol{\alpha}), r')}{2 \cdot GS})}
return z
```

# Algorithm 4 EHDL - Instantiation of the exponential mechanism with local sensitivity

# **Algorithm 5** EHDS - Instantiation of the exponential mechanism with $\gamma$ -smooth sensitivity

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observed data set \mathbf{x} \in \mathcal{X}^n, prior: \mathsf{Dir}(\boldsymbol{\alpha}), \epsilon let \mathsf{Dir}(\boldsymbol{\alpha}') = \mathsf{Dir}\mathsf{P}(\mathbf{x}, \boldsymbol{\alpha}).
let S(\mathbf{x}) be the smooth sensitivity for \mathbf{x}.
set z = r with probability \frac{\exp(\frac{\epsilon \cdot u(\mathbf{x}, r)}{4 \cdot S(\mathbf{x})})}{\sum_{r' \in \mathcal{R}_{\boldsymbol{\alpha}}} \exp(\frac{\epsilon \cdot u(\mathbf{x}, r')}{4 \cdot S(\mathbf{x})})}
return z
```

### 3 Accuracy Analysis

Theorem 3.1. To prove the optimality of Laplace mechanism, we are showing

$$\frac{ELap(\boldsymbol{x})}{(\epsilon \times LS(\boldsymbol{x}))}$$

is O(1), considering  $n = |\mathbf{x}| \ge 2$  being the parameter.

Where  $LS(\cdot)$  is the local sensitivity, and where  $ELap(\cdot)$  is the measure of the error of the Laplace mechanism, defined in this way:

$$ELap(\textbf{\textit{x}}) = \arg \big( \min_t \{ Pr[\mathsf{H}(\mathsf{DirP}(\textbf{\textit{x}}), \mathsf{LSHist}(\textbf{\textit{x}})) < t] \geq 1 - \gamma \big).$$

[[Jiawen:

**Theorem 3.2.** For  $\gamma = e^{O(\epsilon)}$ ,  $\frac{ELap(x)}{(\epsilon \times LS(x))}$  is  $O(\epsilon)$ 

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*Proof.* Let  $t = LS(\mathbf{x})$ , we have following by p.d.f. of Laplace distribution:

$$Pr[\mathsf{H}(\mathsf{DirP}(\mathbf{x}),\mathsf{LSHist}(\mathbf{x})) < t] \geq 1 - \frac{1}{2}(e^{-\epsilon} + e^{-2\epsilon}) > 1 - e^{-\epsilon}$$

Then we can get when  $\gamma = e^{-\epsilon}$ ,

$$\frac{ELap(\mathbf{x})}{(\epsilon \times LS(\mathbf{x}))} = \frac{1}{\epsilon}$$

**Theorem 3.3.** In order to prove the optimality of Laplace mechanism, instead of prove  $\frac{ELap(x)}{(\epsilon \times LS(x))}$  is O(1), we prove a constant upper bound on following equations:

$$\leq \frac{\underset{t}{\arg\min}\left\{\Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(x)) < t] \geq 1 - \gamma\right\}}{\max\limits_{\substack{|k| \leq \frac{\lg(\frac{1}{\gamma})}{\epsilon}}} \mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{Beta}(\alpha+k,n-\lfloor\alpha+k\rfloor))}{LS(x)} } \\ \leq \frac{C(\frac{\lg\frac{1}{\gamma}}{\epsilon})}{C(\frac{\lg\frac{1}{\gamma}}{\epsilon})}$$

[[Jiawen:

**Theorem 3.4.** For  $\gamma = e^{O(k)\epsilon}$ , it is proved that  $O(\frac{\lg \frac{1}{\gamma}}{\epsilon})$  is bounded by O(k).

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*Proof.* By Laplace distribution, we have:

$$\begin{array}{lcl} \Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(\mathbf{x})) < t] & = & \Pr[\{|\mathsf{Lap}(0,\frac{1}{\epsilon})| < O(k)|\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{Beta}(\alpha+k,n-\lfloor\alpha+k\rfloor)) < t\}] \\ & \leq & 1 - e^{-O(k)\epsilon} \end{array}$$

Then we have:

$$\gamma = e^{-O(k)\epsilon}$$

So we can get:

$$O(\frac{\lg \frac{1}{\gamma}}{\epsilon}) = O(\frac{\lg \frac{1}{e^{-O(k)\epsilon}}}{\epsilon}) = O(k)$$

[[Jiawen:

Corollary 3.4.1. For  $-1 \le k < 2$ , it is proved that  $O(\frac{\lg \frac{1}{\gamma}}{\epsilon})$  is bounded by O(1).

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*Proof.* Given  $-1 \le k < 2$ , we have:

$$\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{Beta}(\alpha+k,n-|\alpha+k|)) \le LS(\mathbf{x})$$
 (1)

For any  $\epsilon$ ,  $k \sim \mathsf{Lap}(0, \frac{1}{\epsilon})$  from Laplace mechanism, we have:

$$\Pr[|k| \le \frac{b}{\epsilon}] = 1 - \exp(-b)$$

Then we can get:

$$\Pr[-1 \le k < 2] = 1 - \frac{\exp(-\epsilon) + \exp(-2\epsilon)}{2} \tag{2}$$

By Equation (1) and (2), we can get:

$$\Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{Beta}(\alpha+k,n-\lfloor\alpha+k\rfloor)) \leq LS(\mathbf{x})] \geq 1 - \frac{\exp(-\epsilon) + \exp(-2\epsilon)}{2}$$

i.e.,

$$\begin{split} & \frac{\arg\min_{t} \Big\{\Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(\mathbf{x})) < t] \ge 1 - \frac{\exp(-\epsilon) + \exp(-2\epsilon)}{2}\Big\}}{LS(\mathbf{x})} \\ \le & O\Big(\frac{\lg(\frac{2}{\exp(-\epsilon) + \exp(-2\epsilon)})}{\epsilon}\Big) \\ < & O\Big(\frac{\lg(\frac{2}{\exp(-2\epsilon)})}{2} = 2 \end{split}$$

[[Jiawen:

**Theorem 3.5.** Let k = |k'| be the largest integer that satisfying  $\mathsf{H}(\mathsf{Beta}(\alpha, \beta), \mathsf{Beta}(\alpha + k', n - \lfloor \alpha + k' \rfloor)) < t$ , we have:

$$\begin{split} \Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(\pmb{x})) < t] &\geq 1 - e^{-k\epsilon} \\ \frac{2ke^{-\epsilon} + 1}{n} < \Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHD}(\pmb{x})) < t] < \frac{2k + 1}{ne^{-\epsilon}}. \\ \frac{2k\exp(\frac{-\epsilon}{2\cdot LS(\pmb{x})}) + 1}{n} < \Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHDL}(\pmb{x})) < t] < \frac{2k + 1}{n\exp(\frac{-\epsilon}{2\cdot LS(\pmb{x})})}. \end{split}$$

*Proof.* By the p.d.f. of Laplace distribution, we have:

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$$\Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(\mathbf{x})) < t] = \Pr[\mathsf{LSHist}(\mathbf{x}) \le k] \ge 1 - e^{-k\epsilon}.$$

By definition of EHD and  $GS = \sqrt{1 - \pi/4}$ , we have:

$$\begin{split} Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHD}(\mathbf{x})) < t] &= \sum_{c \geq -t} \frac{\exp(\frac{\epsilon \cdot c}{2 \cdot GS})}{\sum_{r' \in \mathcal{R}_\alpha} \exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x},\alpha),r')}{2 \cdot GS})} \\ &\leq \frac{2k \exp(-\frac{0\epsilon}{2 \cdot GS}) + 1}{n \exp(\frac{2\epsilon}{2 \cdot GS})} \\ &= \frac{2k + 1}{n \exp(\frac{-\epsilon}{2\sqrt{1 - \pi/4}})} \\ &< \frac{2k + 1}{n \exp(-\epsilon)} \end{split}$$
 
$$&< \frac{2k + 1}{n \exp(-\epsilon)} \end{split}$$
 
$$& Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHD}(\mathbf{x})) < t] = \sum_{s \geq -t} \frac{\exp(\frac{\epsilon \cdot c}{2 \cdot GS})}{\sum_{t \leq n} \exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x},\alpha),r')}{2 \cdot GS})} \end{split}$$

$$\begin{array}{ll} Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHD}(\mathbf{x})) < t] & = & \sum\limits_{c \geq -t} \frac{\exp(\frac{e \cdot c}{2 \cdot G S})}{\sum_{r' \in \mathcal{R}_{\pmb{\alpha}}} \exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x},\alpha),r')}{2 \cdot G S})} \\ & \geq & \frac{2k \exp(\frac{-t\epsilon}{2 \cdot G S}) + 1}{n \exp(\frac{-0\epsilon}{2 \cdot G S})} \\ & > & \frac{2k \exp(\frac{-\epsilon}{2 \cdot G S})}{n} \\ & > & \frac{2k \exp(\frac{-\epsilon}{2 \cdot G S}) + 1}{n} \\ & > & \frac{2ke^{-\epsilon} + 1}{n} \end{array}$$

By definition of EHDL, we have:

$$\begin{split} Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHDL}(\mathbf{x})) < t] &= \sum_{\substack{c \geq -t \\ \sum_{r' \in \mathcal{R}_\alpha} \exp(\frac{-c \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x},\alpha),r')}{2 \cdot LS(\mathbf{x})})}} \\ &\leq \frac{2k \exp(\frac{-0c}{2 \cdot LS(\mathbf{x})}) + 1}{n \exp(\frac{-c}{2 \cdot LS(\mathbf{x})})} \\ &= \frac{2k + 1}{n \exp(\frac{-c}{2 \cdot LS(\mathbf{x})})} \end{split}$$

$$\begin{split} Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHD}(\mathbf{x})) < t] &= \sum_{\substack{c \geq -t}} \frac{\exp(\frac{\epsilon \cdot c}{2 \cdot LS(\mathbf{x})})}{\sum_{r' \in \mathcal{R}_{\alpha}} \exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x},\alpha),r')}{2 \cdot LS(\mathbf{x})})} \\ &\geq \frac{2k \exp(\frac{-t}{2 \cdot LS(\mathbf{x})}) + 1}{n \exp(\frac{-0\epsilon}{2 \cdot LS(\mathbf{x})})} \\ &> \frac{2k \exp(\frac{-\epsilon}{2 \cdot LS(\mathbf{x})}) + 1}{n} \end{split}$$

Since  $\lim_{n\to\infty} LS(\mathbf{x})\to 0$ , we have  $\lim_{n\to\infty} \frac{-1}{LS(\mathbf{x})}\to -\infty$ . So  $\exp(\frac{-\epsilon}{2\cdot LS(\mathbf{x})})$  can only be bounded by 0. We cannot found a tighter lower bound.

#### [[Jiawen:

**Corollary 3.5.1.** For a reasonable small t, we have when data size  $n=|x|>O(\frac{(2k+1)e^{\epsilon}}{1-e^{-\epsilon}})$ , the accuracy of LSHist is higher than EHD.

11

Proof. Based on Theorem 2.5, let:

$$\frac{2k+1}{n\exp(-\epsilon)} \le 1 - e^{-k\epsilon},$$

we can have:

$$\Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(\mathbf{x})) < t] > \Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHD}(\mathbf{x})) < t].$$

By simplification, we have 
$$n > \frac{2k+1}{e^{-\epsilon}(1-e^{-k\epsilon})} \sim O(\frac{(2k+1)e^{\epsilon}}{1-e^{-\epsilon}})$$
.

[[Jiawen:

Corollary 3.5.2. Let  $R_g$  be the good output set where  $\forall r \in R$ ,  $\mathsf{H}(\mathsf{DirP}(x), r) \leq LS(x)$ , we have:

$$Pr[\mathsf{LSHist}(\mathbf{x}, \epsilon) \in R_q] > Pr[\mathsf{EHD}(\mathbf{x}, \epsilon) \in R_q]$$

for data size  $n = |\mathbf{x}| > O(\frac{e^{\epsilon}}{1 - e^{-\epsilon}})$ 

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Proof. simply apply the Theorem 2.5 and corollary 2.5.1, we can get this conclusion.

Let  $R_g$  be the good output set where  $\forall r \in R$ ,  $\mathsf{H}(\mathsf{DirP}(\mathbf{x}), r) \leq LS(\mathbf{x})$ , we have:

$$Pr[\mathsf{LSHist}(\mathbf{x}) \in R_g] \ge 1 - \frac{1}{2}(e^{-\epsilon} + e^{-2\epsilon}) > 1 - e^{-\epsilon}$$

By definition of EHD and  $GS = \sqrt{1 - \pi/4}$ , we have:

$$\begin{array}{lcl} Pr[\mathsf{EHD}(\mathbf{x}) \in R_g] & = & \sum\limits_{\substack{c \geq -LS(\mathbf{x})}} \frac{\exp(\frac{\epsilon \cdot c}{2 \cdot GS})}{\sum_{r' \in \mathcal{R}_{\mathbf{\alpha}}} \exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x}, \mathbf{\alpha}), r')}{2 \cdot GS})}) \\ & \leq & \frac{2 \exp(-\frac{\epsilon LS(\mathbf{x})}{2 \cdot GS}) + 1}{n \exp(\frac{-\epsilon}{2 \cdot GS})} \\ & \leq & \frac{3}{n \exp(\frac{-\epsilon}{2\sqrt{1 - \pi/4}})} \\ & \leq & \frac{3}{n \exp(-\epsilon)} \end{array}$$

Let  $c=2\sqrt{1-\pi/4}$ , we have when  $n>\frac{3}{e^{-\epsilon/c}(1-e^{-\epsilon})}\sim O(\frac{e^\epsilon}{1-e^{-\epsilon}})$  LSHist performs better than EHD.