

# Differentially Private Bayesian Inference

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## Abstract

### 1. Setting up

The Bayesian inference process is denoted as  $\text{BayesInfer}(x, \text{prior})$  taking an observed data set  $x \in \mathcal{X}^n$  and a prior distribution as input, outputting a posterior distribution *posterior*. For conciseness, when prior is given, we use  $\text{BayesInfer}(x)$ .

For now, we already have a prior distribution *prior*, an observed data set  $x$ .

#### 1.1. Exponential Mechanism with Global Sensitivity

##### 1.1.1. MECHANISM SET UP

In exponential mechanism, candidate set  $R$  can be obtained by enumerating  $y \in \mathcal{X}^n$ , i.e.

$$R = \{\text{BayesInfer}(y) \mid y \in \mathcal{X}^n\}.$$

Hellinger distance  $\text{Hlg}$  is used here to score these candidates. The utility function:

$$u(x, r) = -\text{Hlg}(\text{BayesInfer}(x), r); r \in R. \quad (1)$$

Exponential mechanism with global sensitivity selects and outputs a candidate  $r \in R$  with probability proportional to  $\exp(\frac{\epsilon u(x, r)}{2\Delta_g u})$ :

$$P[r] = \frac{\exp(\frac{\epsilon u(x, r)}{2\Delta_g u})}{\sum_{r' \in R} \exp(\frac{\epsilon u(x, r')}{2\Delta_g u})},$$

where global sensitivity is calculated by:

$$\Delta_g u = \text{Hlg}(\text{BayesInfer}(x'), r) - \text{Hlg}(\text{BayesInfer}(y'), r) \mid \max_{\{x', y' \mid \leq 1; x', y' \in \mathcal{X}^n\}} \max_{\{r \in R\}}.$$

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#### 1.1.2. SECURITY ANALYSIS

It can be proved that exponential mechanism with global sensitivity is  $\epsilon$ -differentially private. We denote the  $\text{BayesInfer}$  with privacy mechanism as  $\text{PrivInfer}$ . For adjacent data set  $\|x, y\|_1 = 1$ :

$$\begin{aligned} & \frac{P[\text{PrivInfer}(x, u, R) = r]}{P[\text{PrivInfer}(y, u, R) = r]} \\ &= \frac{\exp(\frac{\epsilon u(x, r)}{2\Delta_g u})}{\sum_{r' \in R} \exp(\frac{\epsilon u(x, r')}{2\Delta_g u})} \\ &= \frac{\exp(\frac{\epsilon u(y, r)}{2\Delta_g u})}{\sum_{r' \in R} \exp(\frac{\epsilon u(y, r')}{2\Delta_g u})} \\ &= \left( \frac{\exp(\frac{\epsilon u(x, r)}{2\Delta_g u})}{\exp(\frac{\epsilon u(y, r)}{2\Delta_g u})} \right) \cdot \left( \frac{\sum_{r' \in R} \exp(\frac{\epsilon u(y, r')}{2\Delta_g u})}{\sum_{r' \in R} \exp(\frac{\epsilon u(x, r')}{2\Delta_g u})} \right) \\ &= \exp\left(\frac{\epsilon(u(x, r) - u(y, r))}{2\Delta_g u}\right) \\ &\quad \cdot \left( \frac{\sum_{r' \in R} \exp(\frac{\epsilon u(y, r')}{2\Delta_g u})}{\sum_{r' \in R} \exp(\frac{\epsilon u(x, r')}{2\Delta_g u})} \right) \\ &\leq \exp\left(\frac{\epsilon}{2}\right) \cdot \exp\left(\frac{\epsilon}{2}\right) \cdot \left( \frac{\sum_{r' \in R} \exp(\frac{\epsilon u(x, r')}{2\Delta_g u})}{\sum_{r' \in R} \exp(\frac{\epsilon u(x, r')}{2\Delta_g u})} \right) \\ &= \exp(\epsilon). \end{aligned}$$

Then,  $\frac{P[\text{PrivInfer}(x, u, R) = r]}{P[\text{PrivInfer}(y, u, R) = r]} \geq \exp(-\epsilon)$  can be obtained by symmetry.

#### 1.2. Exponential Mechanism with Local Sensitivity

##### 1.2.1. MECHANISM SET UP

Exponential mechanism with local sensitivity share the same candidate set and utility function as it with global sensitivity. This outputs a candidate  $r \in R$  with probability proportional to  $\exp(\frac{\epsilon u(x, r)}{2\Delta_l u})$ :

$$P[r] = \frac{\exp(\frac{\epsilon u(x, r)}{2\Delta_l u})}{\sum_{r' \in R} \exp(\frac{\epsilon u(x, r')}{2\Delta_l u})},$$

where local sensitivity is calculated by:

$$\Delta_I u(x) = \text{Hlg}(\text{BayesInfer}(x), r) - \text{Hlg}(\text{BayesInfer}(y'), r) | \max_{\{x, y' | \leq 1; y' \in \mathcal{X}^n\}} \max_{\{r \in R\}}.$$

### 1.2.2. SECURITY ANALYSIS

We will then prove that exponential mechanism with local sensitivity is non-differentially private.

$$\begin{aligned} & \frac{P[\text{PrivInfer}(x, u, R) = r]}{P[\text{PrivInfer}(y, u, R) = r]} \\ &= \exp\left(\frac{\epsilon u(x, r)}{2\Delta_I u(x)} - \frac{\epsilon u(y, r)}{2\Delta_I u(y)}\right) \cdot \left(\frac{\sum_{r' \in R} \exp(\frac{\epsilon u(y, r')}{2\Delta_I u(y)})}{\sum_{r' \in R} \exp(\frac{\epsilon u(x, r')}{2\Delta_I u(x)})}\right) \\ &= \frac{\sum_{r' \in R} \exp(\frac{\epsilon u(x, r)}{2\Delta_I u(x)} + \frac{\epsilon u(y, r')}{2\Delta_I u(y)})}{\sum_{r' \in R} \exp(\frac{\epsilon u(y, r)}{2\Delta_I u(y)} + \frac{\epsilon u(x, r')}{2\Delta_I u(x)})}. \end{aligned}$$

Without loss of generality, we consider the case that  $\Delta_I u(y) < \Delta_I u(x)$ ,  $r = \arg(\max_{r' \in R} \{u(x, r')\}) = \arg(\min_{r' \in R} \{u(y, r')\})$  and  $\Delta_I u(y) = u(x, r) - u(y, r)$ . We have:

$$\begin{aligned} & \frac{\sum_{r' \in R} \exp(\frac{\epsilon u(x, r)}{2\Delta_I u(x)} + \frac{\epsilon u(y, r')}{2\Delta_I u(y)})}{\sum_{r' \in R} \exp(\frac{\epsilon u(y, r)}{2\Delta_I u(y)} + \frac{\epsilon u(x, r')}{2\Delta_I u(x)})} \\ & > \frac{\sum_{r' \in R} \exp(\frac{\epsilon(u(x, r) + u(y, r'))}{2\Delta_I u(x)})}{\sum_{r' \in R} \exp(\frac{\epsilon(u(y, r) + u(x, r'))}{2\Delta_I u(y)})} \\ & > \frac{|R| \exp(\frac{\epsilon(u(x, r) + u(y, r))}{2\Delta_I u(x)})}{|R| \exp(\frac{\epsilon(u(y, r) + u(x, r))}{2\Delta_I u(y)})} \\ &= \exp(\frac{\epsilon}{2}(\frac{u(x, r) + u(y, r)}{\Delta_I u(x)} - \frac{u(x, r) + u(y, r)}{\Delta_I u(y)})). \end{aligned}$$

From Eq. 1,  $\{u(x, r') \leq 0 | r' \in R\}$  and  $\{u(y, r') \leq 0 | r' \in R\}$ , we can infer that  $r = \arg(\max_{r' \in R} \{u(x, r')\}) = \arg(\min_{r' \in R} \{u(y, r')\})$ . From  $\Delta_I u(y) = u(x, r) - u(y, r)$ , we can also infer that  $\Delta_I u(y) = -u(y, r)$ . Then, the following relationship between  $u(x, r)$ ,  $u(y, r)$ ,  $\Delta_I u(x)$  and  $\Delta_I u(y)$ :

$$\begin{aligned} & -\Delta_I u(x) < \Delta_I u(y) \\ & \Delta_I u(x) - \Delta_I u(y) < 2\Delta_I u(x) \\ & -\Delta_I u(y)(\Delta_I u(y) - \Delta_I u(x)) < 2\Delta_I u(x)\Delta_I u(y) \\ & u(y, r)(\Delta_I u(y) - \Delta_I u(x)) < 2\Delta_I u(x)\Delta_I u(y) \\ & \frac{u(x, r) + u(y, r)}{\Delta_I u(x)} - \frac{u(x, r) + u(y, r)}{\Delta_I u(y)} > 2. \end{aligned}$$

holds.

Then we can have:

$$\begin{aligned} & \exp(\frac{\epsilon}{2}(\frac{u(x, r) + u(y, r)}{\Delta_I u(y)} - \frac{u(x, r) + u(y, r)}{\Delta_I u(x)})) \\ & > \exp(\frac{\epsilon}{2} * 2) \\ &= \exp(\epsilon), \end{aligned}$$

i.e.

$$\frac{P[\text{PrivInfer}(x, u, R) = r]}{P[\text{PrivInfer}(y, u, R) = r]} > \exp(\epsilon).$$

Since there are cases where exponential mechanism with local sensitivity's privacy loss is greater than  $e^\epsilon$ . we can say it is non-differentially private.

## 1.3. Exponential Mechanism of Varying Sensitivity

### 1.3.1. MECHANISM SETTING UP

### 1.3.2. SECURITY ANALYSIS

## 1.4. Exponential Mechanism of Smooth Sensitivity

### 1.4.1. MECHANISM SETTING UP

### 1.4.2. SECURITY ANALYSIS

## 2. Privacy Fix

### 2.1. Propositions

Assume we have a prior distribution  $\text{beta}(1, 1)$ , an observed data set  $x \in \{0, 1\}^n$ ,  $n > 0$ . We use the  $x + 1$  and  $x - 1$  to denote:

if  $\text{BayesInfer}(x) = \text{beta}(a_1 + 1, b_1 + 1)$

then  $\text{BayesInfer}(x + 1) = \text{beta}((a_1 + 1) + 1, (b_1 - 1) + 1)$

$\text{BayesInfer}(x - 1) = \text{beta}((a_1 - 1) + 1, (b_1 + 1) + 1)$ ,

$x_0$  to denote:

if  $n$  is even

then  $\text{BayesInfer}(x_0) = \text{beta}(\frac{n}{2} + 1, \frac{n}{2} + 1)$

else  $\text{BayesInfer}(x_0) = \{\text{beta}(\frac{n+1}{2} + 1, \frac{n-1}{2} + 1), \text{beta}(\frac{n-1}{2} + 1, \frac{n+1}{2} + 1)\}$

$\text{beta}(\alpha, \beta)$  is the beta function with two arguments  $\alpha$  and  $\beta$ .

Then, we have the following three statements, and proofs of the statements.

$$\begin{aligned} & \text{I } \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x + 1)) < \\ & \text{Hlg}(\text{BayesInfer}(x + 1), \text{BayesInfer}(x + 2)) \forall x \geq x_0; \end{aligned}$$

or  $\text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x + 1)) > \text{Hlg}(\text{BayesInfer}(x + 1), \text{BayesInfer}(x + 2)) \forall x \leq x_0$ .

II  $\Delta_l u(x) = \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x + 1)), \forall x \geq x_0$ ;  
 $\Delta_l u(x) = \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x - 1)), \forall x \leq x_0$ .

III  $\forall x \neq x_0 : \Delta_l u(x) > \Delta_l u(x_0)$ .

## 2.2. proof

### 2.2.1. STATEMENT I

We use the MI (Mathematical Induction) method to prove the first statement.

*Proof.* Since the Hellinger distance is symmetric, if we prove the  $\text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x + 1)) < \text{Hlg}(\text{BayesInfer}(x + 1), \text{BayesInfer}(x + 2)) \forall x \geq x_0$ , the other part when  $\forall x \leq x_0$  also holds.

1. if  $x = x_0$ ,  $\text{Hlg}(\text{BayesInfer}(x_0), \text{BayesInfer}(x_0 + 1)) < \text{Hlg}(\text{BayesInfer}(x_0 + 1), \text{BayesInfer}(x_0 + 2))$  holds:

$$\begin{aligned} & \text{Hlg}(\text{beta}(\frac{n}{2} + 1, \frac{n}{2} + 1), \text{beta}(\frac{n}{2} + 1 + 1, \frac{n}{2} + 1 - 1)) < \text{Hlg}(\text{beta}(\frac{n}{2} + 1 + 1, \frac{n}{2} + 1 - 1), \text{beta}(\frac{n}{2} + 1 + 2, \frac{n}{2} + 1 - 2)) \\ & \sqrt{1 - \frac{\text{beta}(\frac{\frac{n}{2}+1+\frac{n}{2}+1+1}{2}, \frac{\frac{n}{2}+1+\frac{n}{2}+1-1}{2})}{\sqrt{\text{beta}(\frac{n}{2}+1, \frac{n}{2}+1)\text{beta}(\frac{n}{2}+1+1, \frac{n}{2}+1-1)}}} < \sqrt{1 - \frac{\text{beta}(\frac{\frac{n}{2}+1+1+\frac{n}{2}+1+2}{2}, \frac{\frac{n}{2}+1-1+\frac{n}{2}+1-2}{2})}{\sqrt{\text{beta}(\frac{n}{2}+1+1, \frac{n}{2}+1-1)\text{beta}(\frac{n}{2}+1+2, \frac{n}{2}+1-2)}}} \\ & \sqrt{1 - \frac{\text{beta}(\frac{n+3}{2}, \frac{n+1}{2})}{\sqrt{\text{beta}(\frac{n}{2}+1, \frac{n}{2}+1)\text{beta}(\frac{n}{2}+2, \frac{n}{2})}}} < \sqrt{1 - \frac{\text{beta}(\frac{n+5}{2}, \frac{n-1}{2})}{\sqrt{\text{beta}(\frac{n}{2}+2, \frac{n}{2})\text{beta}(\frac{n}{2}+3, \frac{n}{2}-1)}}} \\ & \frac{\text{beta}(\frac{n+3}{2}, \frac{n+1}{2})}{\sqrt{\text{beta}(\frac{n}{2}+1, \frac{n}{2}+1)\text{beta}(\frac{n}{2}+2, \frac{n}{2})}} > \frac{\text{beta}(\frac{n+5}{2}, \frac{n-1}{2})}{\sqrt{\text{beta}(\frac{n}{2}+2, \frac{n}{2})\text{beta}(\frac{n}{2}+3, \frac{n}{2}-1)}} \\ & \frac{\text{beta}(\frac{n+3}{2}, \frac{n-1}{2})^{\frac{\frac{n-1}{2}-1}{\frac{n-1}{2}+\frac{n+3}{2}}}}{\sqrt{\text{beta}(\frac{n}{2}+1, \frac{n}{2}-1)^{\frac{\frac{n-1}{2}-1}{\frac{n-1}{2}+\frac{n+3}{2}}}}} > \frac{\text{beta}(\frac{n+3}{2}, \frac{n-1}{2})^{\frac{\frac{n+3}{2}-1}{\frac{n+3}{2}+\frac{n-1}{2}}}}{\sqrt{\text{beta}(\frac{n}{2}+1, \frac{n}{2}-1)^{\frac{\frac{n+3}{2}-1}{\frac{n+3}{2}+\frac{n-1}{2}}}}} \\ & \frac{\frac{n-1}{2}}{\sqrt{(\frac{n}{2}-1)(\frac{n}{2})}} > \frac{\frac{n+3}{2}}{\sqrt{(\frac{n}{2}+1)(\frac{n}{2}+2)}} \\ & (n-1)^2(n+2)(n+4) > (n+3)^2n(n-2) \\ & n > -1. \end{aligned}$$

Since  $n > 0$ , it always holds.

2. if  $x = x_0 + m$  holds, then also  $x = x_0 + m + 1$  holds:

i.e  $\text{Hlg}(\text{beta}(\frac{n}{2} + 1 + m, \frac{n}{2} + 1 - m), \text{beta}(\frac{n}{2} + 1 + m + 1, \frac{n}{2} + 1 - m - 1)) < \text{Hlg}(\text{beta}(\frac{n}{2} + 1 + m + 1, \frac{n}{2} + 1 - m - 1), \text{beta}(\frac{n}{2} + 1 + m + 2, \frac{n}{2} + 1 - m - 2))$  is what we know:

$$\begin{aligned} & \sqrt{1 - \frac{\text{beta}(\frac{\frac{n}{2}+1+m+\frac{n}{2}+1+m+1}{2}, \frac{\frac{n}{2}+1-m+\frac{n}{2}+1-m-1}{2})}{\sqrt{\text{beta}(\frac{n}{2}+1+m, \frac{n}{2}+1-m)\text{beta}(\frac{n}{2}+2+m, \frac{n}{2}-m)}}} \\ & < \sqrt{1 - \frac{\text{beta}(\frac{\frac{n}{2}+1+m+1+\frac{n}{2}+1+m+2}{2}, \frac{\frac{n}{2}+1-m-1+\frac{n}{2}+1-m-2}{2})}{\sqrt{\text{beta}(\frac{n}{2}+2+m, \frac{n}{2}-m)\text{beta}(\frac{n}{2}+3+m, \frac{n}{2}-m-1)}}} \\ & \frac{\text{beta}(\frac{n+2m+3}{2}, \frac{n-2m+1}{2})}{\sqrt{\text{beta}(\frac{n}{2}+1+m, \frac{n}{2}+1-m)\text{beta}(\frac{n}{2}+2+m, \frac{n}{2}-m)}} \\ & > \frac{\text{beta}(\frac{n+2m+5}{2}, \frac{n-2m-1}{2})}{\sqrt{\text{beta}(\frac{n}{2}+2+m, \frac{n}{2}-m)\text{beta}(\frac{n}{2}+3+m, \frac{n}{2}-m-1)}} \\ & \sqrt{\text{beta}(\frac{n}{2}+2+m, \frac{n}{2}-m)\text{beta}(\frac{n}{2}+3+m, \frac{n}{2}-m-1)^{\frac{n-2m}{n+2m+2}}} \\ & > \sqrt{\text{beta}(\frac{n}{2}+2+m, \frac{n}{2}-m)\text{beta}(\frac{n}{2}+3+m, \frac{n}{2}-m-1)^{\frac{n-2m-2}{n+2m+6}}} \end{aligned}$$

## 2.2.3. STATEMENT III

*Proof.* From Statement I and Statement II, we can conclude that:

when  $x > x_0$

$$\begin{aligned} & \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x+1)) \\ & > \text{Hlg}(\text{BayesInfer}(x_0), \text{BayesInfer}(x_0+1)); \\ & \text{i.e. } \Delta_{lu}(x) > \Delta_{lu}(x_0) \end{aligned}$$

when  $x < x_0$

$$\begin{aligned} & \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x-1)) \\ & > \text{Hlg}(\text{BayesInfer}(x_0), \text{BayesInfer}(x_0-1)); \\ & \text{i.e. } \Delta_{lu}(x) > \Delta_{lu}(x_0). \end{aligned}$$

$$\begin{aligned} & \frac{\text{beta}(\frac{n+2m+5}{2}, \frac{n-2m-1}{2})}{\sqrt{\text{beta}(\frac{n}{2}+2+m, \frac{n}{2}-m)\text{beta}(\frac{n}{2}+3+m, \frac{n}{2}-m-1)}} \\ & > \frac{\text{beta}(\frac{n+2m+7}{2}, \frac{n-2m-3}{2})}{\sqrt{\text{beta}(\frac{n}{2}+2+m, \frac{n}{2}-m)\text{beta}(\frac{n}{2}+3+m, \frac{n}{2}-m-1)}} \\ & < \sqrt{1 - \frac{\text{beta}(\frac{n+2m+5}{2}, \frac{n-2m-1}{2})}{\sqrt{\text{beta}(\frac{n}{2}+2+m, \frac{n}{2}-m)\text{beta}(\frac{n}{2}+3+m, \frac{n}{2}-m-1)}}} \\ & < \sqrt{1 - \frac{\text{beta}(\frac{n+2m+7}{2}, \frac{n-2m-3}{2})}{\sqrt{\text{beta}(\frac{n}{2}+2+m, \frac{n}{2}-m)\text{beta}(\frac{n}{2}+3+m, \frac{n}{2}-m-1)}}} \\ & \text{Hlg}(\text{beta}(\frac{n}{2}+2+m, \frac{n}{2}-m), \text{beta}(\frac{n}{2}+3+m, \frac{n}{2}-1-m)) \\ & < \text{Hlg}(\text{beta}(\frac{n}{2}+m+3, \frac{n}{2}-1-m), \text{beta}(\frac{n}{2}+m+4, \frac{n}{2}-m-2)) \end{aligned}$$

i.e.  $x = x_0 + m + 1$  also holds when  $x = x_0 + m$  is valid.  $\square$

## 3. Experimental Evaluations

We got some results from these mechanisms.

## References

## 2.2.2. STATEMENT II

*Proof.*

$$\begin{aligned} \therefore \Delta_{lu}(x) &= |\text{Hlg}(\text{BayesInfer}(x), r) - \text{Hlg}(\text{BayesInfer}(y'), r)|, \\ & \max_{\{|x, y'| \leq 1; y' \in \mathcal{X}^n\}} \max_{\{r \in R\}}; \\ \therefore \text{Hlg}(\text{BayesInfer}(x), r) - \text{Hlg}(\text{BayesInfer}(y'), r) \\ & \leq \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(y')); \\ \therefore \Delta_{lu}(x) &= \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(y')), \\ & \max_{\{|x, y'| \leq 1; y' \in \mathcal{X}^n\}}; \\ \therefore \Delta_{lu}(x) &= \max\{\text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x+1)), \\ & \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x-1))\}; \end{aligned}$$

According to Statement I :

$$\begin{aligned} & \text{if } x > x_0 \\ & \text{then } \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x-1)) \\ & < \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x+1)); \\ & \text{then } \Delta_{lu}(x) = \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x+1)); \\ & \text{if } x < x_0 \\ & \text{then } \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x-1)) \\ & > \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x+1)); \\ & \text{then } \Delta_{lu}(x) = \text{Hlg}(\text{BayesInfer}(x), \text{BayesInfer}(x-1)); \\ & \text{else } \Delta_{lu}(x_0) = \text{Hlg}(\text{BayesInfer}(x_0), \text{BayesInfer}(x_0-1)) \\ & = \text{Hlg}(\text{BayesInfer}(x_0), \text{BayesInfer}(x_0+1)). \end{aligned}$$

From above, we can conclude the Statement II.  $\square$