Differentially Private Bayesian Inference

1 Preliminary

 θ : The parameter vector of multinomial distribution, $\theta \in [0,1]^k$.

 \mathbf{x} : Observed dataset. $\mathbf{x} \in \mathcal{X}^n, |\mathcal{X}| = k$

 $\mathsf{Dir}(\alpha)$: Dirichlet distribution. The prior or posterior distribution over θ .

 $\mathsf{DirP}(\alpha, \theta, \mathbf{x})$: Posterior distribution over θ from Bayesian inference given prior distribution $\mathsf{Dir}(\alpha)$

and observed data set \mathbf{x} .

 $\mathcal{H}(\cdot,\cdot) \qquad : \text{ Hellinger Distance between two distributions. } \mathcal{H}(\mathsf{Dir}(\boldsymbol{\alpha}_1),\mathsf{Dir}(\boldsymbol{\alpha}_2)) = \sqrt{1 - \frac{\mathsf{B}(\frac{\boldsymbol{\alpha}_1 + \boldsymbol{\alpha}_2}{2})}{\sqrt{\mathsf{B}(\boldsymbol{\alpha}_1)\mathsf{B}(\boldsymbol{\alpha}_2)}}}$

 $u(\mathbf{x},r)$: Scoring function given $\boldsymbol{\alpha}$ and $\boldsymbol{\theta}$ for candidate r. $u(\mathbf{x},r) = -\mathcal{H}(\mathsf{DirP}(\boldsymbol{\alpha},\boldsymbol{\theta},\mathbf{x}),r)$

GS : Global sensitivity of Hellinger distance. $GS = \sqrt{1 - \pi/4}$

 $LS(\mathbf{x})$: Local sensitivity of Hellinger distance for \mathbf{x} .

 $LS(\mathbf{x}) = \max_{\mathbf{x}' \in \mathcal{X}^n, \mathbf{adj}(\mathbf{x}, \mathbf{x}'), r \in \mathcal{R}_{\alpha}} |\mathcal{H}(\mathsf{DirP}(\mathbf{x}, \alpha), r) - \mathcal{H}(\mathsf{DirP}(\mathbf{x}', \alpha), r)|$

 $S(\mathbf{x}) : \gamma - \text{smooth sensitivity. } S(\mathbf{x}) = \max_{\mathbf{x}' \in \mathcal{X}^n} \left\{ \frac{1}{\frac{1}{LS(\mathbf{x}')} + \gamma \cdot Himming(\mathbf{x}, \mathbf{x}')} \right\}$

2 Private Mechanisms

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Algorithm 1 LSZhang[1] - Calibrating noise w.r.t. \ell_1 norm and Dimension
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\begin{split} \mathbf{x} &\in \mathcal{X}^n, \, \mathsf{Dir}(\boldsymbol{\alpha}) \\ &\mathbf{let} \,\, \boldsymbol{\alpha}' = \mathsf{DirP}(\boldsymbol{\alpha}, \boldsymbol{\theta}, \mathbf{x}) \\ &\mathbf{Initialize} \,\, \mathbf{a} \,\, \mathsf{vector} \,\, \tilde{\boldsymbol{\alpha}} = (0, \dots, 0) \in \mathbb{N}^{|\mathcal{X}|} \\ &\mathbf{For} \,\, i = 1 \dots |\mathcal{X}| - 1 \colon \\ &\mathbf{let} \,\, \tilde{\alpha_i} = \alpha_i + \lfloor (\alpha_i' - \alpha_i) + \mathsf{Lap}(0, \frac{2|\mathcal{X}|}{\epsilon}) \rfloor_0^n \\ &\mathbf{return} \,\, \tilde{\boldsymbol{\alpha}} \end{split}
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Algorithm 2 LSDim - Calibrating noise w.r.t. ℓ_1 norm

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\begin{aligned} \mathbf{x} &\in \mathcal{X}^n, \, \mathsf{Dir}(\boldsymbol{\alpha}) \\ & \mathbf{let} \,\, \boldsymbol{\alpha}' = \mathsf{DirP}(\boldsymbol{\alpha}, \boldsymbol{\theta}, \mathbf{x}) \\ & \mathbf{Initialize} \,\, \mathbf{a} \,\, \mathrm{vector} \,\, \tilde{\boldsymbol{\alpha}} = (0, \dots, 0) \in \mathbb{N}^{|\mathcal{X}|} \\ & \mathbf{For} \,\, i = 1 \dots |\mathcal{X}| - 1: \\ & \tilde{\alpha_i} = \alpha_i + \lfloor (\alpha_i' - \alpha_i) + \mathsf{Lap}(0, \frac{|\mathcal{X}|}{\epsilon}) \rfloor_0^n \\ & \tilde{\alpha}_{|\mathcal{X}|} = \alpha_{|\mathcal{X}|} + \lfloor n - \sum_{i=1}^{|\mathcal{X}|-1} \lfloor (\alpha_i' - \alpha_i) + \eta_i \rfloor_0^n \rfloor_0^n \\ & \mathbf{return} \,\, \tilde{\boldsymbol{\alpha}} \end{aligned}
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Algorithm 3 LSHist - Calibrating noise w.r.t. histogram sensitivity

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\begin{split} & \text{input } \mathbf{x} \in \mathcal{X}^n, \, \mathsf{Dir}(\boldsymbol{\alpha}) \\ & \text{let } \boldsymbol{\alpha}' = \mathsf{DirP}(\boldsymbol{\alpha}, \boldsymbol{\theta}, \mathbf{x}) \\ & \text{let } k = \left\{ \begin{array}{ll} 1 & \text{if} & |\mathcal{X}| = 2 \\ 2 & \text{otherwise} \end{array} \right. \\ & \textbf{Initialize } \text{a vector } \tilde{\boldsymbol{\alpha}} = (0, \dots, 0) \in \mathbb{N}^{|\mathcal{X}|} \\ & \textbf{For } i = 1 \dots |\mathcal{X}| - 1 \text{:} \\ & \text{let } \eta \sim \mathsf{Lap}(0, \frac{k}{\epsilon}) \\ & \tilde{\alpha}_i = \alpha_i + \lfloor (\alpha_i' - \alpha_i) + \eta \rfloor_0^n \\ & \tilde{\alpha}_{|\mathcal{X}|} = \alpha_{|\mathcal{X}|} + \lfloor n - \sum_{i=1}^{|\mathcal{X}|-1} \lfloor (\alpha_i' - \alpha_i) + \eta_i \rfloor_0^n \rfloor_0^n \\ & \textbf{return } \tilde{\boldsymbol{\alpha}} \end{split}
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${f Algorithm~4~EHD}$ - Instantiation of the exponential mechanism

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observed data set \mathbf{x} \in \mathcal{X}^n, prior: \mathsf{Dir}(\boldsymbol{\alpha}), \epsilon let \mathsf{Dir}(\boldsymbol{\alpha}') = \mathsf{DirP}(\mathbf{x}, \boldsymbol{\alpha}).

let GS be the global sensitivity for \mathbf{x}.

set z = r with probability \frac{\exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x}, \boldsymbol{\alpha}), r')}{2 \cdot GS})}{\sum_{r' \in \mathcal{R}_{\boldsymbol{\alpha}}} \exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x}, \boldsymbol{\alpha}), r')}{2 \cdot GS})}
return z
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Algorithm 5 EHDL - Instantiation of the exponential mechanism with local sensitivity

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 \begin{aligned} & \text{input observed data set } \mathbf{x} \in \mathcal{X}^n, \text{ prior: } \mathsf{Dir}(\boldsymbol{\alpha}), \, \epsilon \\ & \text{let } \mathsf{Dir}(\boldsymbol{\alpha}') = \mathsf{Dir}\mathsf{P}(\mathbf{x}, \boldsymbol{\alpha}). \\ & \text{let } LS(\mathbf{x}) \text{ be the local sensitivity for } \mathbf{x}. \\ & \text{set } z = r \text{ with probability } \frac{\exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{Dir}\mathsf{P}(\mathbf{x}, \boldsymbol{\alpha}), r)}{2 \cdot LS(\mathbf{x})})}{\sum_{r' \in \mathcal{R}_{\boldsymbol{\alpha}}} \exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{Dir}\mathsf{P}(\mathbf{x}, \boldsymbol{\alpha}), r')}{2 \cdot LS(\mathbf{x})})} \\ & \text{return } z \end{aligned}
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Algorithm 6 EHDS - Instantiation of the exponential mechanism with γ -smooth sensitivity

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observed data set \mathbf{x} \in \mathcal{X}^n, prior: \mathsf{Dir}(\boldsymbol{\alpha}), \epsilon let \mathsf{Dir}(\boldsymbol{\alpha}') = \mathsf{DirP}(\mathbf{x}, \boldsymbol{\alpha}).
let S(\mathbf{x}) be the smooth sensitivity for \mathbf{x}.
set z = r with probability \frac{\exp(\frac{\epsilon \cdot u(\mathbf{x},r)}{2(1+\gamma)S(\mathbf{x})})}{\sum_{r' \in \mathcal{R}_{\boldsymbol{\alpha}}} \exp(\frac{\epsilon \cdot u(\mathbf{x},r')}{2(1+\gamma)S(\mathbf{x})})}
return z
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3 Privacy Analysis

Theorem 3.1. The LSDim, LSHist, EHD and EHDS are ϵ -differentially private.

The proofs are in the Arxiv version.

4 Accuracy Analysis

Theorem 4.1. To prove the optimality of Laplace mechanism, we are showing

$$\frac{ELap(\boldsymbol{x})}{(\epsilon \times LS(\boldsymbol{x}))}$$

is O(1), considering $n = |\mathbf{x}| \ge 2$ being the parameter.

Where $LS(\cdot)$ is the local sensitivity, and where $ELap(\cdot)$ is the measure of the error of the Laplace mechanism, defined in this way:

$$ELap(\textbf{\textit{x}}) = \arg \big(\min_t \{ Pr[\mathsf{H}(\mathsf{DirP}(\textbf{\textit{x}}), \mathsf{LSHist}(\textbf{\textit{x}})) < t] \geq 1 - \gamma \big).$$

[[Jiawen:

Theorem 4.2. For $\gamma = e^{O(\epsilon)}$, $\frac{ELap(x)}{(\epsilon \times LS(x))}$ is $O(\epsilon)$

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Proof. Let $t = LS(\mathbf{x})$, we have following by p.d.f. of Laplace distribution:

$$Pr[\mathsf{H}(\mathsf{DirP}(\mathbf{x}),\mathsf{LSHist}(\mathbf{x})) < t] \geq 1 - \frac{1}{2}(e^{-\epsilon} + e^{-2\epsilon}) > 1 - e^{-\epsilon}$$

Then we can get when $\gamma = e^{-\epsilon}$,

$$\frac{ELap(\mathbf{x})}{(\epsilon \times LS(\mathbf{x}))} = \frac{1}{\epsilon}$$

Theorem 4.3. In order to prove the optimality of Laplace mechanism, instead of prove $\frac{ELap(x)}{(\epsilon \times LS(x))}$ is O(1), we prove a constant upper bound on following equations:

$$\leq \frac{\underset{t}{\arg\min}\left\{\Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(x)) < t] \geq 1 - \gamma\right\}}{\max\limits_{\substack{|k| \leq \frac{\lg(\frac{1}{\gamma})}{\epsilon}}}{\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{Beta}(\alpha+k,n-\lfloor\alpha+k\rfloor))}} \\ \leq \frac{\sum\limits_{\substack{|k| \leq \frac{\lg(\frac{1}{\gamma})}{\epsilon}}}{\mathsf{L}S(x)}}{\mathsf{L}S(x)}$$

[[Jiawen:

Theorem 4.4. For $\gamma = e^{O(k)\epsilon}$, it is proved that $O(\frac{\lg \frac{1}{\gamma}}{\epsilon})$ is bounded by O(k).

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Proof. By Laplace distribution, we have:

$$\begin{array}{lcl} \Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(\mathbf{x})) < t] & = & \Pr[\{|\mathsf{Lap}(0,\frac{1}{\epsilon})| < O(k)|\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{Beta}(\alpha+k,n-\lfloor\alpha+k\rfloor)) < t\}] \\ & \leq & 1 - e^{-O(k)\epsilon} \end{array}$$

Then we have:

$$\gamma = e^{-O(k)\epsilon}$$

So we can get:

$$O(\frac{\lg \frac{1}{\gamma}}{\epsilon}) = O(\frac{\lg \frac{1}{e^{-O(k)\epsilon}}}{\epsilon}) = O(k)$$

[[Jiawen:

Corollary 4.4.1. For $-1 \le k < 2$, it is proved that $O(\frac{\lg \frac{1}{\gamma}}{\epsilon})$ is bounded by O(1).

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Proof. Given $-1 \le k < 2$, we have:

$$\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{Beta}(\alpha+k,n-|\alpha+k|)) \le LS(\mathbf{x})$$
 (1)

For any ϵ , $k \sim \mathsf{Lap}(0, \frac{1}{\epsilon})$ from Laplace mechanism, we have:

$$\Pr[|k| \le \frac{b}{\epsilon}] = 1 - \exp(-b)$$

Then we can get:

$$\Pr[-1 \le k < 2] = 1 - \frac{\exp(-\epsilon) + \exp(-2\epsilon)}{2} \tag{2}$$

By Equation (1) and (2), we can get:

$$\Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{Beta}(\alpha+k,n-\lfloor\alpha+k\rfloor)) \leq LS(\mathbf{x})] \geq 1 - \frac{\exp(-\epsilon) + \exp(-2\epsilon)}{2}$$

i.e.,

$$\begin{split} & \frac{\arg\min_{t} \Big\{\Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(\mathbf{x})) < t] \geq 1 - \frac{\exp(-\epsilon) + \exp(-2\epsilon)}{2} \Big\}}{LS(\mathbf{x})} \\ \leq & O\Big(\frac{\lg(\frac{2}{\exp(-\epsilon) + \exp(-2\epsilon)})}{\epsilon}\Big) \\ < & O\Big(\frac{\lg(\frac{2}{\exp(-2\epsilon)})}{2} = 2 \end{split}$$

[[Jiawen:

Theorem 4.5. Let k = |k'| be the largest integer that satisfying $\mathsf{H}(\mathsf{Beta}(\alpha, \beta), \mathsf{Beta}(\alpha + k', n - |\alpha + k'|)) < t$, we have:

$$\begin{split} \Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSZhang}(\pmb{x})) < t] &= (1 - \frac{1}{2}(e^{-k\epsilon/2} + e^{-(k+1)\epsilon/2}))^2 \geq (1 - e^{-k\epsilon/2})^2 \\ &\quad \Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(\pmb{x})) < t] = 1 - \frac{1}{2}(e^{-k\epsilon} + e^{-(k+1)\epsilon}) \geq 1 - e^{-k\epsilon} \\ &\quad \frac{2ke^{-\epsilon} + 1}{n} < \Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHD}(\pmb{x})) < t] < \frac{2k + 1}{ne^{-\epsilon}}. \\ &\quad \frac{2k\exp(\frac{-\epsilon}{2 \cdot LS(\pmb{x})}) + 1}{n} < \Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHDL}(\pmb{x})) < t] < \frac{2k + 1}{n\exp(\frac{-\epsilon}{2 \cdot LS(\pmb{x})})}. \end{split}$$

Proof. Given k = |k'| be the largest integer that satisfying $\mathsf{H}(\mathsf{Beta}(\alpha, \beta), \mathsf{Beta}(\alpha + k', n - \lfloor \alpha + k' \rfloor)) < t$, by the post-processing of Laplace distribution and p.d.f. of Laplace distribution, we have:

$$\Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(\mathbf{x})) < t] = \Pr[-k < \mathsf{LSHist}(\mathbf{x}) \leq k+1] = 1 - \frac{1}{2}(e^{-k\epsilon} + e^{-(k+1)\epsilon}) \geq 1 - e^{-k\epsilon}$$

By definition of EHD and $GS = \sqrt{1 - \pi/4}$, we have:

$$\begin{split} Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHD}(\mathbf{x})) < t] &= \sum_{\substack{c \geq -t}} \frac{\exp(\frac{\epsilon \cdot c}{2 \cdot GS})}{\sum_{r' \in \mathcal{R}_\alpha} \exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x},\alpha),r')}{2 \cdot GS})} \\ &\leq \frac{2k \exp(-\frac{0\epsilon}{2 \cdot GS}) + 1}{n \exp(\frac{2\epsilon}{2 \cdot GS})} \\ &= \frac{2k + 1}{n \exp(\frac{-\epsilon}{2 \cdot GS})} \\ &< \frac{2k + 1}{n \exp(-\epsilon)} \end{split}$$

$$Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHD}(\mathbf{x})) < t] &= \sum_{\substack{c \geq -t}} \frac{\exp(\frac{\epsilon \cdot c}{2 \cdot GS})}{\sum_{r' \in \mathcal{R}_\alpha} \exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x},\alpha),r')}{2 \cdot GS})} \\ &\geq \frac{2k \exp(\frac{-\epsilon \cdot c}{2 \cdot GS}) + 1}{n \exp(\frac{2\epsilon \cdot c}{2 \cdot GS})} \\ &> \frac{2k \exp(\frac{2\epsilon \cdot c}{2 \cdot GS})}{2 \cdot (1 - \pi / 4}) + 1} \\ &> \frac{2k \exp(\frac{2\epsilon \cdot c}{2 \cdot (1 - \pi / 4}) + 1}{n} \\ &> \frac{2k e^{-\epsilon} + 1}{n} \end{split}$$

By definition of EHDL, we have:

$$\begin{split} Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHDL}(\mathbf{x})) < t] &= \sum_{\substack{c \geq -t}} \frac{\exp(\frac{c\epsilon}{2 \cdot LS(\mathbf{x})})}{\sum_{r' \in \mathcal{R}_{\boldsymbol{\alpha}}} \exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x},\boldsymbol{\alpha}),r')}{2 \cdot LS(\mathbf{x})})} \\ &\leq \frac{2k \exp(\frac{-0\epsilon}{2 \cdot LS(\mathbf{x})}) + 1}{n \exp(\frac{-\epsilon}{2 \cdot LS(\mathbf{x})})} \\ &= \frac{2k + 1}{n \exp(\frac{-\epsilon}{2 \cdot LS(\mathbf{x})})} \end{split}$$

$$Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHD}(\mathbf{x})) < t] &= \sum_{\substack{c \geq -t}} \frac{\exp(\frac{\epsilon \cdot c}{2 \cdot LS(\mathbf{x})})}{\sum_{r' \in \mathcal{R}_{\boldsymbol{\alpha}}} \exp(\frac{-\epsilon \cdot \mathcal{H}(\mathsf{DirP}(\mathbf{x},\boldsymbol{\alpha}),r')}{2 \cdot LS(\mathbf{x})})} \\ &\geq \frac{2k \exp(\frac{-t\epsilon}{2 \cdot LS(\mathbf{x})}) + 1}{n \exp(\frac{-0\epsilon}{2 \cdot LS(\mathbf{x})})} \\ &> \frac{2k \exp(\frac{-t\epsilon}{2 \cdot LS(\mathbf{x})}) + 1}{n} \end{split}$$

Since $\lim_{n\to\infty} LS(\mathbf{x})\to 0$, we have $\lim_{n\to\infty} \frac{-1}{LS(\mathbf{x})}\to -\infty$. So $\exp(\frac{-\epsilon}{2\cdot LS(\mathbf{x})})$ can only be bounded by 0. We cannot found a tighter lower bound.

[[Jiawen:

Corollary 4.5.1. For a reasonable small t, we have when data size $n=|x|>O(\frac{(2k+1)e^{\epsilon}}{1-e^{-\epsilon}})$, the accuracy of LSHist is higher than EHD.

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Proof. Based on Theorem 2.5, let:

$$\frac{2k+1}{n\exp(-\epsilon)} \le 1 - e^{-k\epsilon},$$

we can have:

$$\Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(\mathbf{x})) < t] > \Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHD}(\mathbf{x})) < t].$$

By simplification, we have
$$n > \frac{2k+1}{e^{-\epsilon}(1-e^{-k\epsilon})} \sim O(\frac{(2k+1)e^{\epsilon}}{1-e^{-\epsilon}})$$
.

[[Jiawen:

Corollary 4.5.2. For a reasonable small t, we have when data size $n=|x|< O(\frac{(2ke^{-\epsilon}+1)}{1-\frac{1}{2}(e^{-k\epsilon}+e^{-(k+1)\epsilon})})$, the accuracy of EHD is better than LSHist.

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Proof. Applying the Theorem 2.5, let:

$$\frac{2ke^{-\epsilon} + 1}{n} > 1 - \frac{1}{2}(e^{-k\epsilon} + e^{-(k+1)\epsilon}),$$

we can have:

$$\Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{LSHist}(\mathbf{x})) < t] < \Pr[\mathsf{H}(\mathsf{Beta}(\alpha,\beta),\mathsf{EHD}(\mathbf{x})) < t].$$

By simplification, we have
$$n < \frac{(2ke^{-\epsilon}+1)}{1-\frac{1}{2}(e^{-k\epsilon}+e^{-(k+1)\epsilon})}$$
.

[[Jiawen:

Corollary 4.5.3. Let R_g be the good output set where $\forall r \in R$, $\mathsf{H}(\mathsf{DirP}(x), r) \leq LS(x)$, we have:

$$Pr[\mathsf{LSHist}(\mathbf{x}, \epsilon) \in R_g] > Pr[\mathsf{EHD}(\mathbf{x}, \epsilon) \in R_g]$$

for data size $n = |\mathbf{x}| > O(\frac{e^{\epsilon}}{1 - e^{-\epsilon}})$

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Proof. simply apply the Theorem 2.5 and corollary 2.5.1, we can get this conclusion.

Let R_g be the good output set where $\forall r \in R$, $\mathsf{H}(\mathsf{DirP}(\mathbf{x}), r) \leq LS(\mathbf{x})$, we have:

$$Pr[\mathsf{LSHist}(\mathbf{x}) \in R_g] \ge 1 - \frac{1}{2}(e^{-\epsilon} + e^{-2\epsilon}) > 1 - e^{-\epsilon}$$

By definition of EHD and $GS = \sqrt{1 - \pi/4}$, we have:

$$\begin{split} Pr[\mathsf{EHD}(\mathbf{x}) \in R_g] &= \sum_{\substack{c \geq -LS(\mathbf{x})}} \frac{\exp(\frac{\epsilon \cdot c}{2 \cdot GS})}{\sum_{r' \in \mathcal{R}_{\boldsymbol{\alpha}}} \exp(\frac{-\epsilon \cdot H(\mathsf{DirP}(\mathbf{x}, \boldsymbol{\alpha}), r')}{2 \cdot GS})} \\ &\leq \frac{2 \exp(-\frac{\epsilon LS(\mathbf{x})}{2 \cdot GS}) + 1}{n \exp(\frac{-\epsilon}{2 \cdot GS})} \\ &\leq \frac{3}{n \exp(\frac{-\epsilon}{2\sqrt{1-\pi/4}})} \\ &\leq \frac{3}{n \exp(-\epsilon)} \end{split}$$

Let $c=2\sqrt{1-\pi/4}$, we have when $n>\frac{3}{e^{-\epsilon/c}(1-e^{-\epsilon})}\sim O(\frac{e^\epsilon}{1-e^{-\epsilon}})$ LSHist performs better than EHD.

References

[1] Zuhe Zhang, Benjamin IP Rubinstein, Christos Dimitrakakis, et al. On the differential privacy of bayesian inference. In AAAI, pages 2365–2371, 2016.